

# Statistical learning: Advanced

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*Lecture 1: Components-based methods*

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## Successive topics of the coming lectures:

1. Component-based analysis:  
PCA, PLS, CCA,...
2. Beyond quantitative data: Categorical data  
Correspondence Analysis (CA), MCA, FAMD
3. Clustering from dissimilarity measures:  
Similarity,  $K$ -means and kernelized  $K$ -means,  
Hierarchical clustering
4. Change-point detection  
Parametric models and non-parametric ones
5. Multiple testing and FDR control

PCA

PLS

CCA

PCA

PLS

CCA

- ▶ PCA: Principal Component Analysis
- ▶ PLS: Partial Least-Squares
- ▶ CCA: Canonical Correlation Analysis

PCA

Unsupervised

PCA algorithm

PLS

CCA

# Principal Component Analysis (PCA)

## Notation

- ▶  $n$  individuals
- ▶ Each individual  $i$  described by  $d$  (quantitative) variables  
 $X_i = (X_i^1, \dots, X_i^d)^\top \in \mathbb{R}^d$ : features for individual  $i$
- ▶  $\mathbf{X}$ :  $n \times d$  “full” design matrix
- ▶  $\mathbf{X}_{\cdot,j} \in \mathbb{R}^n$ :  $j$ th column vector of  $\mathbf{X}$   
( $n$  measurements for variable  $X^j$ )
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## Notation

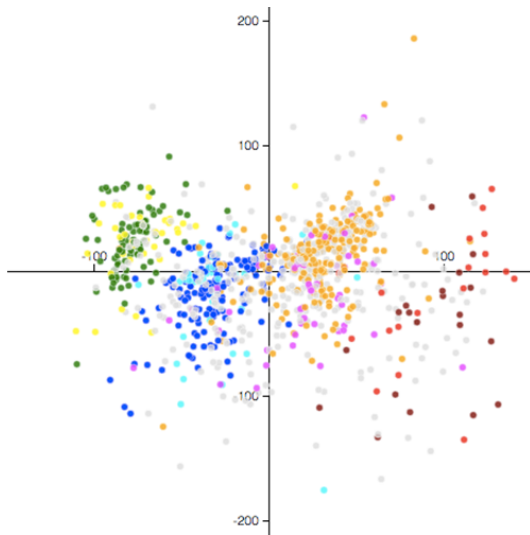
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There is no label  $Y$  in the PCA story!

# Assumption for having PCA work well

## Key assumption

Variability carries information about heterogeneity



# PCA algorithm

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$$\widehat{\text{Var}}(\mathbf{X}^\top \theta) = \frac{1}{n} \sum_{i=1}^n \mathbf{x}_{i\cdot}^\top \theta = \frac{1}{n} (\mathbf{X}\theta)^\top (\mathbf{X}\theta) = \langle \mathbf{X}\theta, \mathbf{X}\theta \rangle$$

## PCA algorithm

### ► Step 1:

► Find

$$\theta^1 \in \text{Arg} \max_{\|\theta\|=1} \{ \langle \mathbf{X}\theta, \mathbf{X}\theta \rangle \}$$



# PCA algorithm

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►  $h^1 = \tilde{\mathbf{X}}_{.1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$ : first principal component

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►  $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$ : first principal component

### ► Step t ( $2 \leq t \leq \text{rank}(\mathbf{X}) = r$ ):

► Find

$$\theta^t \in \text{Arg} \max_{\substack{\|\theta\|=1 \\ \langle \mathbf{X}\theta, h^j \rangle = 0 \\ j = 1, \dots, t-1}} \{ \langle \mathbf{X}\theta, \mathbf{X}\theta \rangle \}$$

$$\widehat{\text{Var}}(\mathbf{X}^\top \theta) = \frac{1}{n} \sum_{i=1}^n \mathbf{x}_i^\top \theta = \frac{1}{n} (\mathbf{X}\theta)^\top (\mathbf{X}\theta) = \langle \mathbf{X}\theta, \mathbf{X}\theta \rangle$$

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►  $h^t = \tilde{\mathbf{X}}_{\cdot t} = \mathbf{X}\theta^t \in \mathbb{R}^n$ :  $t$ th principal component

$$\widehat{\text{Var}}(\mathbf{X}^\top \theta) = \frac{1}{n} \sum_{i=1}^n \mathbf{x}_i^\top \theta = \frac{1}{n} (\mathbf{X}\theta)^\top (\mathbf{X}\theta) = \langle \mathbf{X}\theta, \mathbf{X}\theta \rangle$$

## PCA algorithm

### ► Step 1:

► Find

$$\theta^1 \in \text{Arg} \max_{\|\theta\|=1} \{ \langle \mathbf{X}\theta, \mathbf{X}\theta \rangle \}$$

►  $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$ : first principal component

### ► Step t ( $2 \leq t \leq \text{rank}(\mathbf{X}) = r$ ):

► Find

$$\theta^t \in \text{Arg} \max_{\substack{\|\theta\|=1 \\ \langle \mathbf{X}\theta, h^j \rangle = 0 \\ j = 1, \dots, t-1}} \{ \langle \mathbf{X}\theta, \mathbf{X}\theta \rangle \}$$

►  $h^t = \tilde{\mathbf{X}}_{\cdot t} = \mathbf{X}\theta^t \in \mathbb{R}^n$ : tth principal component

### ► Output:

$$\mathbf{X} \approx \tilde{\mathbf{X}} = [h^1; h^2; \dots; h^r] = \mathbf{X} [\theta^1; \dots; \theta^r]$$

## Step 1

- Find

$$\theta^1 \in \operatorname{Arg} \max_{\|\theta\|=1} \{\langle \mathbf{X}\theta, \mathbf{X}\theta \rangle\}$$

- $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$ : first principal component

## Interpretation

- $\langle \mathbf{X}\theta, \mathbf{X}\theta \rangle = \|\mathbf{X}\theta\|^2 = \theta^\top \mathbf{X}^\top \mathbf{X} \theta = n \widehat{\operatorname{Var}}(\mathbf{X}^\top \theta)$
- $h^1$  is the first (one-dimensional) approximation to  $\mathbf{X}$  for describing each individual  $i$  by  $h_i^1 \in \mathbb{R}$  instead of  $\mathbf{X}_{i\cdot} \in \mathbb{R}^d$

# PCA: Focus on the other steps

Step  $t$  ( $2 \leq t \leq r = \text{rank}(\mathbf{X})$ )

► Find

$$\theta^t \in \underset{\substack{\|\theta\| = 1 \\ \langle \mathbf{X}\theta, h^j \rangle = 0 \\ j = 1, \dots, t-1}}{\text{Arg max}} \{ \langle \mathbf{X}\theta, \mathbf{X}\theta \rangle \}$$

►  $h^t = \tilde{\mathbf{X}}_{\cdot t} = \mathbf{X}\theta^t \in \mathbb{R}^n$ :  $t$ th principal component

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- $h^t = \tilde{\mathbf{X}}_{\cdot t} = \mathbf{X}\theta^t \in \mathbb{R}^n$ :  $t$ th principal component

Interpretation

- $\langle \mathbf{X}\theta, h^j \rangle = \langle \theta, (\mathbf{X}^\top \mathbf{X}) \theta^j \rangle = 0$  for all  $2 \leq j \leq r$
- Any new component must be  $\perp$  previous ones
- $(h^1, \dots, h^r)$ : orthogonal family
- Each individual  $i$  described by:

$$(h_i^1, h_i^2) \in \mathbb{R}^2 \text{ instead of } \mathbf{X}_{i\cdot} \in \mathbb{R}^d$$

(with two components)

# Partial Least-Squares (PLS 1)



## Notation

- ▶  $n$  individuals
- ▶ Each individual  $i$  described by  $d$  (quantitative) variables  $X^1, \dots, X^d$
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- ▶  $\mathbf{X}_{.j} \in \mathbb{R}^n$ :  $j$ th column vector of  $\mathbf{X}$   
( $n$  measurements for variable  $X^j$ )
- ▶  $\mathbf{X}_{i.} \in \mathbb{R}^d$ :  $i$ th row vector of  $\mathbf{X}$
- ▶ Label:  
Each individual  $i$  described by  $Y_i$
- ▶  $\mathbf{Y} = (Y_1, \dots, Y_n)^\top \in \mathbb{R}^n$

PCA

PLS

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PLS algorithm

PLS: Have a new look

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## Notation

- ▶  $n$  individuals
- ▶ Each individual  $i$  described by  $d$  (quantitative) variables  $X^1, \dots, X^d$
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- ▶ Label:

Each individual  $i$  described by  $Y_i$

- ▶  $\mathbf{Y} = (Y_1, \dots, Y_n)^\top \in \mathbb{R}^n$

Unlike PCA, the PLS framework is a supervised one!

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# Comparison PCA/PLS

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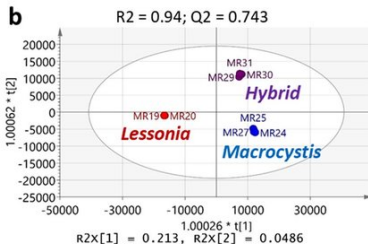
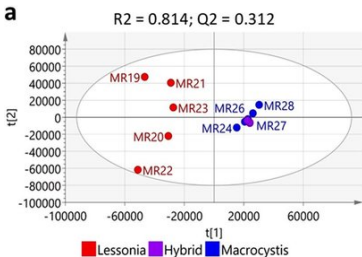
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- **Picture (a):** If one class exhibits more variability than the others, PCA focuses on this class for defining axes  
→ blue and purple points gathered within one cluster

# Comparison PCA/PLS

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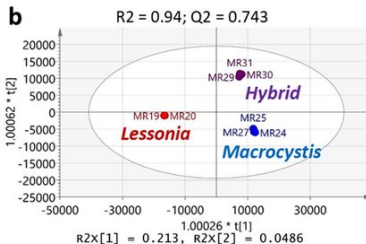
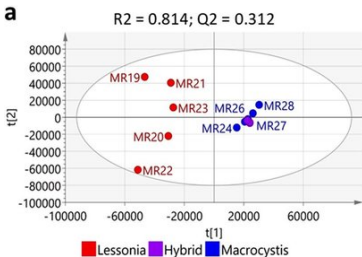
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- ▶ **Picture (a):** If one class exhibits more variability than the others, PCA focuses on this class for defining axes  
→ blue and purple points gathered within one cluster
- ▶ **Picture (b):** With PLS, the axes split labeled classes into well separated clusters

# PLS algorithm: Step 1

$$\text{Cov}(\widehat{X^\top \theta}, Y) = \frac{1}{n} \sum_{i=1}^n \mathbf{x}_{i\cdot}^\top \theta \cdot Y_i = \frac{1}{n} \langle \mathbf{X} \theta, \mathbf{Y} \rangle$$

(columns of  $\mathbf{X}$  centered and reduced)  
(vector  $\mathbf{Y}$  is centered)

# PLS algorithm: Step 1

$$\text{Cov}(\widehat{\mathbf{X}^\top \theta}, \mathbf{Y}) = \frac{1}{n} \sum_{i=1}^n \mathbf{x}_{i\cdot}^\top \theta \cdot Y_i = \frac{1}{n} \langle \mathbf{X}\theta, \mathbf{Y} \rangle$$

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## PLS-1 algorithm: Step 1

- Find

$$\theta^1 \in \text{Arg max}_{\|\theta\|=1} \{ \langle \mathbf{X}\theta, \mathbf{Y} \rangle \}$$

- $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$ : first PLS component

# PLS algorithm: Step 1

$$\text{Cov}(\widehat{X^T \theta}, Y) = \frac{1}{n} \sum_{i=1}^n \mathbf{x}_i^T \theta \cdot Y_i = \frac{1}{n} \langle \mathbf{X} \theta, \mathbf{Y} \rangle$$

(columns of  $\mathbf{X}$  centered and reduced)  
(vector  $\mathbf{Y}$  is centered)

## PLS-1 algorithm: Step 1

- Find

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- $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X} \theta^1 \in \mathbb{R}^n$ : first PLS component

## Remarks:

- $h^1$  maximizes the covariance with  $\mathbf{Y}$  (explicability)

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## Remarks:

- $h^1$  maximizes the covariance with  $\mathbf{Y}$  (explicability)
- Focuses on covariance ( $\neq$  correlation!)
- Not scale invariant  
→ centering and reducing columns of  $\mathbf{X}$



# PLS algorithm: Step 2

PLS-1 algorithm: Step  $t$  ( $2 \leq t \leq r = \text{rank}(\mathbf{X})$ )

► Find

$$\begin{aligned} \theta^t \in \text{Arg} \quad & \max_{\|\theta\| = 1} \quad \{\langle \mathbf{X}\theta, \mathbf{Y} \rangle\} \\ & \langle \mathbf{X}\theta, h^j \rangle = 0 \\ & j = 1, \dots, t-1 \end{aligned}$$

►  $h^t = \tilde{\mathbf{X}}_{.t} = \mathbf{X}\theta^t \in \mathbb{R}^n$ :  $t$ th PLS component

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Interpretation

- $\langle \mathbf{X}\theta, h^j \rangle = \langle \theta, (\mathbf{X}^\top \mathbf{X}) \theta^j \rangle = 0$  for all  $2 \leq j \leq r$
- Any new PLS component must be  $\perp$  previous ones

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(with two components)

# Revisiting PLS algorithm: Step 1

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- Find

$$\theta^1 \in \operatorname{Arg} \max_{\|\theta\|=1} \{\langle \mathbf{X}\theta, \mathbf{Y} \rangle\}$$

- With  $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$

Calculating the solution

$$\theta^1 \in \operatorname{Arg} \max_{\|\theta\|=1} \{\langle \mathbf{X}\theta, \mathbf{Y} \rangle\}$$

$$\Leftrightarrow \theta^1 \in \operatorname{Arg} \max_{\|\theta\|=1} \{\langle \theta, \mathbf{X}^\top \mathbf{Y} \rangle\}$$

$$\Leftrightarrow \theta^1 = \frac{\mathbf{X}^\top \mathbf{Y}}{\|\mathbf{X}^\top \mathbf{Y}\|}$$

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# Revisiting PLS algorithm: Step 1

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Calculating the solution

$$\theta^1 \in \mathop{\text{Arg max}}_{\|\theta\|=1} \{ \langle \mathbf{X}\theta, \mathbf{Y} \rangle \}$$

$$\Leftrightarrow \theta^1 \in \mathop{\text{Arg max}}_{\|\theta\|=1} \left\{ \langle \theta, \mathbf{X}^\top \mathbf{Y} \rangle \right\}$$

$$\Leftrightarrow \theta^1 = \frac{\mathbf{X}^\top \mathbf{Y}}{\|\mathbf{X}^\top \mathbf{Y}\|}$$

→ The first PLS component is  $h^1 = \mathbf{X}\theta^1 = \frac{\mathbf{X}\mathbf{X}^\top \mathbf{Y}}{\|\mathbf{X}^\top \mathbf{Y}\|}$

# Revisiting PLS algorithm: Step 1

- Find

$$\theta^1 \in \text{Arg} \max_{\|\theta\|=1} \{\langle \mathbf{X}\theta, \mathbf{Y} \rangle\}$$

- With  $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$

## Calculating the solution

$$\theta^1 \in \text{Arg} \max_{\|\theta\|=1} \{\langle \mathbf{X}\theta, \mathbf{Y} \rangle\}$$

$$\Leftrightarrow \theta^1 \in \text{Arg} \max_{\|\theta\|=1} \{\langle \theta, \mathbf{X}^\top \mathbf{Y} \rangle\}$$

$$\Leftrightarrow \theta^1 = \frac{\mathbf{X}^\top \mathbf{Y}}{\|\mathbf{X}^\top \mathbf{Y}\|}$$

$$\longrightarrow \text{The first PLS component is } h^1 = \mathbf{X}\theta^1 = \frac{\mathbf{X}\mathbf{X}^\top \mathbf{Y}}{\|\mathbf{X}^\top \mathbf{Y}\|}$$

## Interpretation of $h^1$

$h^1 = \hat{\mathbf{Y}}^1$ : first explanation of  $\mathbf{Y}$  by "the model"  $\mathbf{X}\theta$   
(regression)

# Revisiting PLS algorithm: Step 2 ( $t = 2$ )

Orthogonality constraint

$$(V_1 = \text{Vect}(\mathbf{X}^\top h^1))$$

$$\langle \mathbf{X}\theta, h^1 \rangle = 0 \quad \Leftrightarrow \quad \langle \theta, \mathbf{X}^\top h^1 \rangle = 0$$

$$\Leftrightarrow \quad \theta \in \text{Vect}(\mathbf{X}^\top h^1)^\perp = V_1^\perp$$

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# Revisiting PLS algorithm: Step 2 ( $t = 2$ )

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Orthogonality constraint

$$(V_1 = \text{Vect}(\mathbf{X}^\top h^1))$$

$$\langle \mathbf{X}\theta, h^1 \rangle = 0 \quad \Leftrightarrow \quad \langle \theta, \mathbf{X}^\top h^1 \rangle = 0$$

$$\Leftrightarrow \quad \theta \in \text{Vect}(\mathbf{X}^\top h^1)^\perp = V_1^\perp$$

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Constrained maximization

Under this orthogonality constraint, solving

$$\max_{\|\theta\|=1} \{ \langle \mathbf{X}\theta, \mathbf{Y} \rangle \} \quad \Leftrightarrow \quad \max_{\|\theta\|=1} \left\{ \left\langle \theta, \mathbf{X}^\top \mathbf{Y} - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y}) \right\rangle \right\}$$

where  $\Pi_{V_1}(\cdot)$  denotes the orthogonal projection onto  $V_1$

Reminder:

$$u - \Pi_{V_1}(u) \in V_1^\perp, \quad \forall u \in \mathbb{R}^d$$

# Revisiting PLS algorithm: Step 2 ( $t = 2$ )

## Solution at $t = 2$

$$\theta^2 = \underset{\substack{\|\theta\| = 1 \\ \langle \theta, \mathbf{X}^\top \mathbf{h}^1 \rangle = 0}}{\text{Arg max}} \left\{ \langle \theta, \mathbf{X}^\top \mathbf{Y} \rangle \right\}$$

$$\Leftrightarrow \theta^2 = \underset{\substack{\|\theta\| = 1 \\ \langle \theta, \mathbf{X}^\top \mathbf{h}^1 \rangle = 0}}{\text{Arg max}} \left\{ \langle \theta, \mathbf{X}^\top \mathbf{Y} - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y}) \rangle \right\}$$

$$\Leftrightarrow \theta^2 = \frac{\mathbf{X}^\top \mathbf{Y} - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y})}{\|\mathbf{X}^\top \mathbf{Y} - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y})\|}$$

# Revisiting PLS algorithm: Step 2 ( $t = 2$ )

Solution at  $t = 2$

$$\theta^2 = \underset{\substack{\|\theta\| = 1 \\ \langle \theta, \mathbf{X}^\top \mathbf{h}^1 \rangle = 0}}{\operatorname{Arg} \max} \left\{ \langle \theta, \mathbf{X}^\top \mathbf{Y} \rangle \right\}$$

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Reminder:

$$\Pi_{V_1}(\mathbf{X}^\top \mathbf{Y}) = \left\langle \mathbf{X}^\top \mathbf{Y}, \frac{\mathbf{X}^\top \mathbf{h}^1}{\|\mathbf{X}^\top \mathbf{h}^1\|} \right\rangle \frac{\mathbf{X}^\top \mathbf{h}^1}{\|\mathbf{X}^\top \mathbf{h}^1\|}$$

# PLS-1 algorithm: Summary

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## Pseudo-code

Input:  $\mathbf{X}$ ,  $\mathbf{Y}$

- Initialization  $t = 0$ :

$$\mathbf{R}^0 = \mathbf{X}^\top \mathbf{Y}$$

- Step  $t = 1$ :

- $\theta^1 = \frac{\mathbf{R}^0}{\|\mathbf{R}^0\|}$  and  $h^1 = \mathbf{X}\theta^1$

- $\Pi_{V_1}(\mathbf{X}^\top \mathbf{Y}) = \left\langle \mathbf{X}^\top \mathbf{Y}, \frac{\mathbf{X}^\top h^1}{\|\mathbf{X}^\top h^1\|} \right\rangle \frac{\mathbf{X}^\top h^1}{\|\mathbf{X}^\top h^1\|}$

- $\mathbf{R}^1 = \mathbf{R}^0 - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y})$

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PLS algorithm

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# PLS-1 algorithm: Summary

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- $R^1 = R^0 - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y})$

- For  $t = 2$  to  $\text{rank}(\mathbf{X}) = r$ :

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- $\mathbf{R}^t = \mathbf{R}^{t-1} - \Pi_{V_t}(\mathbf{X}^\top \mathbf{Y})$

Output:  $(h^1, h^2, \dots, h^r)$  such that  $(r = \text{rank}(\mathbf{X}))$

$$\tilde{\mathbf{X}} = [h^1; h^2; \dots; h^r] = \mathbf{X} \cdot [\theta^1; \theta^2; \dots; \theta^r]$$

## Covariance and not correlation

- Covariance is not scale invariante

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  - ▶ **are not unitary vectors!** (unlike the  $\mathbf{X}_{.j}$ s)

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- ▶ By contrast, the  $\mathbf{h}^j$ s (PLS components):
  - ▶ are an orthogonal basis
  - ▶ are not unitary vectors! (unlike the  $\mathbf{X}_{.j}$ s)
- ▶ Can induce some bias in the vizualization

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# Canonical Correlation Analysis (CCA)

## Notation

- ▶  $n$  individuals
- ▶ Each individual  $i$  described by  $d$  (quantitative) variables  $X^1, \dots, X^d$
- ▶  $X_i = (X_i^1, \dots, X_i^d)^\top \in \mathbb{R}^d$ : features for individual  $i$
- ▶  $\mathbf{X}$ :  $n \times d$  “full” design matrix
- ▶  $\mathbf{X}_{.j} \in \mathbb{R}^n$ :  $j$ th column vector of  $\mathbf{X}$   
( $n$  measurements for variable  $X^j$ )
- ▶  $\mathbf{X}_{i.} \in \mathbb{R}^d$ :  $i$ th row vector of  $\mathbf{X}$
- ▶ Label:  
Each individual  $i$  described by  $Y_i$
- ▶  $\mathbf{Y} = (Y_1, \dots, Y_n)^\top \in \mathbb{R}^n$

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Unlike PCA, the CCA framework is a supervised one!

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# CCA algorithm

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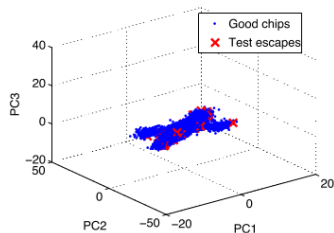
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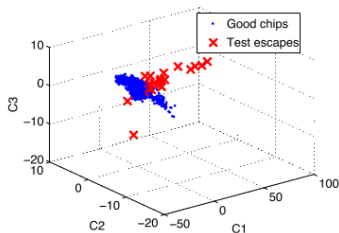
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3-D Principal Component Space



3-D Canonical Space



# CCA-1 algorithm: Step 1

$$\widehat{\text{Corr}}(\mathbf{X}\theta, \mathbf{Y}) = \frac{\langle \mathbf{X}\theta, \mathbf{Y} \rangle}{\|\mathbf{X}\theta\| \cdot \|\mathbf{Y}\|}$$

(columns of  $\mathbf{X}$  centered and reduced)  
(vector  $\mathbf{Y}$  is centered)

## CCA-1 algorithm: Step 1

- Find

$$\theta^1 \in \text{Arg max}_{\theta \in \mathbb{R}^d} \left\{ \frac{\langle \mathbf{X}\theta, \mathbf{Y} \rangle}{\|\mathbf{X}\theta\| \cdot \|\mathbf{Y}\|} \right\}$$

- $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$ : first CCA component

### Remark:

- Focuses on maximizing correlation!
- $h^1$  maximizes the correlation with  $\mathbf{Y}$  (explicability)

# CCA-1 algorithm: Step 2

CCA-1 algorithm: Step  $t$  ( $2 \leq t \leq r = \text{rank}(\mathbf{X})$ )

► Find

$$\theta^t \in \text{Arg} \max_{\substack{\langle \mathbf{X}\theta, h^j \rangle = 0 \\ j = 1, \dots, t-1}} \left\{ \frac{\langle \mathbf{X}\theta, \mathbf{Y} \rangle}{\|\mathbf{X}\theta\| \cdot \|\mathbf{Y}\|} \right\}$$

►  $h^t = \tilde{\mathbf{X}}_{\cdot t} = \mathbf{X}\theta^t \in \mathbb{R}^n$ :  $t$ th CCA component

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# CCA-1 algorithm: Step 2

## CCA-1 algorithm: Step $t$ ( $2 \leq t \leq r = \text{rank}(\mathbf{X})$ )

- Find

$$\theta^t \in \underset{\langle \mathbf{X}\theta, h^j \rangle = 0}{\text{Arg max}} \left\{ \frac{\langle \mathbf{X}\theta, \mathbf{Y} \rangle}{\|\mathbf{X}\theta\| \cdot \|\mathbf{Y}\|} \right\} \\ j = 1, \dots, t-1$$

- $h^t = \tilde{\mathbf{X}}_{\cdot t} = \mathbf{X}\theta^t \in \mathbb{R}^n$ :  $t$ th CCA component

## Interpretation

- $\langle \mathbf{X}\theta, h^j \rangle = \langle \theta, (\mathbf{X}^\top \mathbf{X}) \theta^j \rangle = 0$  for all  $2 \leq j \leq d$
- Any new component must be  $\perp$  previous ones
- $(h^1, \dots, h^r)$ : **orthonormal family**
- Each individual  $i$  described by:

$$(h_i^1, h_i^2) \in \mathbb{R}^2 \text{ instead of } \mathbf{X}_{i\cdot} \in \mathbb{R}^d \\ \text{(with two components)}$$

# Revisiting CCA algorithm: Step 1

► Find

$$\theta^1 \in \operatorname{Arg\,max}_{\theta \in \mathbb{R}^d} \left\{ \frac{\langle \mathbf{X}\theta, \mathbf{Y} \rangle}{\|\mathbf{X}\theta\| \cdot \|\mathbf{Y}\|} \right\}$$

►  $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$

Calculating the solution

# Revisiting CCA algorithm: Step 1

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► Find

$$\theta^1 \in \text{Arg max}_{\theta \in \mathbb{R}^d} \left\{ \frac{\langle \mathbf{X}\theta, \mathbf{Y} \rangle}{\|\mathbf{X}\theta\| \cdot \|\mathbf{Y}\|} \right\}$$

►  $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$

Calculating the solution

$$\begin{aligned} \theta^1 &\in \text{Arg max}_{\theta \in \mathbb{R}^d} \left\{ \frac{\langle \mathbf{X}\theta, \mathbf{Y} \rangle}{\|\mathbf{X}\theta\| \cdot \|\mathbf{Y}\|} \right\} \\ \Leftrightarrow \theta^1 &\in \text{Arg max}_{\|\mathbf{X}\theta\|=1} \{ \langle \mathbf{X}\theta, \mathbf{Y} \rangle \} \\ \Leftrightarrow \theta^1 &\in \text{Arg max}_{\|\mathbf{X}\theta\|=1} \left\{ \left\langle \theta, \mathbf{X}^\top \mathbf{Y} \right\rangle \right\} \\ \Leftrightarrow \theta^1 &= \frac{\mathbf{X}^\top \mathbf{Y}}{\|\mathbf{X}\mathbf{X}^\top \mathbf{Y}\|} \end{aligned}$$

→ The first CCA component is  $h^1 = \mathbf{X}\theta^1 = \frac{\mathbf{X}\mathbf{X}^\top \mathbf{Y}}{\|\mathbf{X}\mathbf{X}^\top \mathbf{Y}\|}$

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# Revisiting CCA algorithm: Step 2 ( $t = 2$ )

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Orthogonality constraint

$$(V_1 = \text{Vect}(\mathbf{X}^\top h^1))$$

$$\begin{aligned}\langle \mathbf{X}\theta, h^1 \rangle = 0 &\Leftrightarrow \langle \theta, \mathbf{X}^\top h^1 \rangle = 0 \\ &\Leftrightarrow \theta \in \text{Vect}(\mathbf{X}^\top h^1)^\perp = V_1^\perp\end{aligned}$$

Constrained maximization

Under this orthogonality constraint, solving

$$\max_{\|\mathbf{X}\theta\|=1} \{ \langle \mathbf{X}\theta, \mathbf{Y} \rangle \} \Leftrightarrow \max_{\|\mathbf{X}\theta\|=1} \left\{ \left\langle \theta, \mathbf{X}^\top \mathbf{Y} - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y}) \right\rangle \right\}$$

where  $\Pi_{V_1}(\cdot)$  denotes the orthogonal projection onto  $V_1$

Reminder:

$$u - \Pi_{V_1}(u) \in V_1^\perp, \quad \forall u \in \mathbb{R}^d$$

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# Revisiting CCA algorithm: Step 2 ( $t = 2$ )

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Solution at  $t = 2$

$$\theta^2 = \underset{\substack{\|\mathbf{X}\theta\| = 1 \\ \langle \theta, \mathbf{X}^\top \mathbf{h}^1 \rangle = 0}}{\operatorname{Arg} \max} \left\{ \langle \theta, \mathbf{X}^\top \mathbf{Y} \rangle \right\}$$

$$\Leftrightarrow \theta^2 = \underset{\substack{\|\mathbf{X}\theta\| = 1 \\ \langle \theta, \mathbf{X}^\top \mathbf{h}^1 \rangle = 0}}{\operatorname{Arg} \max} \left\{ \langle \theta, \mathbf{X}^\top \mathbf{Y} - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y}) \rangle \right\}$$

$$\Leftrightarrow \theta^2 = \frac{\mathbf{X}^\top \mathbf{Y} - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y})}{\|\mathbf{X}(\mathbf{X}^\top \mathbf{Y} - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y}))\|}$$

Reminder:

$$\Pi_{V_1}(\mathbf{X}^\top \mathbf{Y}) = \left\langle \mathbf{X}^\top \mathbf{Y}, \frac{\mathbf{X}^\top \mathbf{h}^1}{\|\mathbf{X}^\top \mathbf{h}^1\|} \right\rangle \frac{\mathbf{X}^\top \mathbf{h}^1}{\|\mathbf{X}^\top \mathbf{h}^1\|}$$

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# CCA-1 algorithm: Summary

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## Pseudo-code

Input:  $\mathbf{X}, \mathbf{Y}$

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$$R^0 = \mathbf{X}^\top \mathbf{Y}$$

- Step  $t = 1$ :

- $\theta^1 = \frac{R^0}{\|\mathbf{X}^\top R^0\|}$  and  $h^1 = \mathbf{X} \theta^1$

- $\Pi_{V_1}(\mathbf{X}^\top \mathbf{Y}) = \left\langle \mathbf{X}^\top \mathbf{Y}, \frac{\mathbf{X}^\top h^1}{\|\mathbf{X}^\top h^1\|} \right\rangle \frac{\mathbf{X}^\top h^1}{\|\mathbf{X}^\top h^1\|}$

- $R^1 = R^0 - \Pi_{V_1}(\mathbf{X}^\top \mathbf{Y})$

- For  $t = 2$  to  $\text{rank}(\mathbf{X}) = r$ :

- $\theta^t = \frac{R^{t-1}}{\|\mathbf{X}^\top R^{t-1}\|}$  and  $h^t = \mathbf{X} \theta^t$

- $\Pi_{V_t}(\mathbf{X}^\top \mathbf{Y}) = \left\langle \mathbf{X}^\top \mathbf{Y}, \frac{\mathbf{X}^\top h^t}{\|\mathbf{X}^\top h^t\|} \right\rangle \frac{\mathbf{X}^\top h^t}{\|\mathbf{X}^\top h^t\|}$

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Output:  $(h^1, h^2, \dots, h^r)$  such that  $(r = \text{rank}(\mathbf{X}))$

$$\tilde{\mathbf{X}} = [h^1; h^2; \dots; h^r] = \mathbf{X} \cdot [\theta^1; \theta^2; \dots; \theta^r]$$

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## Extension to two design matrices

## Notation

- ▶  $n$  individuals
- ▶ First design matrix  $\mathbf{X}$ :
  - ▶  $X_i = (X_i^1, \dots, X_i^d)^\top \in \mathbb{R}^d$ : features for individual  $i$
  - ▶  $\mathbf{X}$ :  $n \times d$  “full” design matrix
  - ▶  $\mathbf{X}_{\cdot,j} \in \mathbb{R}^n$ :  $j$ th column vector of  $\mathbf{X}$
  - ▶  $\mathbf{X}_{i,\cdot} \in \mathbb{R}^d$ :  $i$ th row vector of  $\mathbf{X}$
- ▶ Second design matrix  $\mathbf{W}$ :
  - ▶  $W_i = (W_i^1, \dots, W_i^p)^\top \in \mathbb{R}^p$ : features for individual  $i$
  - ▶  $\mathbf{W}$ :  $n \times p$  “full” design matrix
  - ▶  $\mathbf{W}_{\cdot,j} \in \mathbb{R}^n$ :  $j$ th column vector of  $\mathbf{W}$
  - ▶  $\mathbf{W}_{i,\cdot} \in \mathbb{R}^p$ :  $i$ th row vector of  $\mathbf{W}$

## Examples for $\mathbf{W}$ and $\mathbf{X}$

- ▶ Results from two different experiments  
(Two psychologic tests, bio. exper. at two time-stamps)
- ▶ Wine critic ratings and wine biological features

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# CCA algorithm: Step 1

$$\text{Corr}(\widehat{\mathbf{X}\theta}, \mathbf{W}\gamma) = \frac{\langle \mathbf{X}\theta, \mathbf{W}\gamma \rangle}{\|\mathbf{X}\theta\| \cdot \|\mathbf{W}\gamma\|}$$

(columns of  $\mathbf{X}$ ,  $\mathbf{W}$  centered and reduced)

## CCA algorithm: Step 1

- Find

$$\theta^1, \gamma^1 \in \text{Arg} \max_{\theta \in \mathbb{R}^d, \gamma \in \mathbb{R}^p} \left\{ \frac{\langle \mathbf{X}\theta, \mathbf{W}\gamma \rangle}{\|\mathbf{X}\theta\| \cdot \|\mathbf{W}\gamma\|} \right\}$$

- $h^1 = \tilde{\mathbf{X}}_{\cdot 1} = \mathbf{X}\theta^1 \in \mathbb{R}^n$ : first CCA component for  $\mathbf{X}$
- $g^1 = \tilde{\mathbf{W}}_{\cdot 1} = \mathbf{W}\gamma^1 \in \mathbb{R}^n$ : first CCA component for  $\mathbf{W}$

# CCA algorithm: Step 2

CCA algorithm: Step  $t$  ( $2 \leq t \leq \min(\text{rank}(\mathbf{X}), \text{rank}(\mathbf{W}))$ )

- Find

$$\theta^t, \gamma^t \in \underset{\substack{\langle \mathbf{X}\theta, h^j \rangle = 0 \\ \langle \mathbf{W}\gamma, g^j \rangle = 0 \\ j = 1, \dots, t-1}}{\text{max}} \left\{ \frac{\langle \mathbf{X}\theta, \mathbf{W}\gamma \rangle}{\|\mathbf{X}\theta\| \cdot \|\mathbf{W}\gamma\|} \right\}$$

- $r = \min(\text{rank}(\mathbf{X}), \text{rank}(\mathbf{W}))$

## Interpretation

- $\langle \mathbf{X}\theta, h^j \rangle = \langle \theta, (\mathbf{X}^\top \mathbf{X}) \theta^j \rangle = 0$  for all  $2 \leq j \leq t-1$
- $(h^1, \dots, h^r)$  and  $(g^1, \dots, g^r)$  : **orthonormal families**  
since the above problem is equivalent to

$$\underset{\substack{\|\mathbf{X}\theta\| = 1, \|\mathbf{W}\gamma\| = 1 \\ \langle \mathbf{X}\theta, h^j \rangle = 0 \\ \langle \mathbf{W}\gamma, g^j \rangle = 0 \\ j = 1, \dots, t-1}}{\text{max}} \{ \langle \mathbf{X}\theta, \mathbf{W}\gamma \rangle \}$$

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## Summary

- ▶ Implementations similar to those of PLS and CCA-1 do exist for efficiently solving the CCA optimization problem
- ▶ CCA is used in a lot of applications for combining different types of information ("tables") used for describing "individuals"
- ▶ Extensions of CCA do exist for multiple tables (see Generalized CCA...)