

MiningLeaks

Water Pollution and Child Mortality in Africa

[Job Market Paper]

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Abstract

Industrial mining can be a boon or a bane for communities living in the vicinity of production sites. We assess the effects of mining-induced pollution on health outcomes in Sub-Saharan Africa using the DHS micro-data from 1986-2018 in 26 countries matched with geocoded data of industrial mining sites. Through a staggered difference-in-difference strategy, we exploit the variation of the opening of a mine and the relative topographic position of surrounding villages, comparing upstream and downstream villages. Being downstream of an open mine increases by 2.18 percentage points the 24-month mortality rate, corresponding to a 25% increase. This effect is mainly driven by the consumption of plain water, corroborating the mechanism of water pollution. The effect on mortality is not driven by a change in women's fertility, nor by a change in the access to piped water or other facilities, nor by in-migration. The effect is concentrated while the mine is active and when international mineral prices are high, is larger in areas with high mining density, and fades out with distance. It is robust to the estimator of de Chaisemartin and d'Haultfœuille [2020], to a restriction to a balanced sample, to accounting for measurement errors, and to spatial and temporal randomization inference tests.

Keywords: Africa; Industrial Mining; Health; Water Pollution; Natural Resource; Environmental Degradation.

JEL codes: I1; L72; Q53; Q32; Q33.

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1 Introduction

The increase in commodity prices since 2000, especially in the extractive sector, has intensified investments in areas with abundant resources, from hydrocarbons to minerals. The African geology, which is richly endowed containing 30% of the world's mineral reserves [Chuhan-Pole et al., 2017], remains largely unexplored due to inhospitable terrains and the lack of infrastructure [UN Environment Program , 2022; Africa Bank, 2022] and represents an opportunity for mining investors. Africa is facing a mining boom since the 2000s, attracting foreign investment mainly from China, Canada, Australia, Brazil, and Russia, which raises concerns about environmental degradation [Taylor et al., 2009; Edwards et al., 2013] and health impact on local populations. Human exposure to heavy metals through the consumption of contaminated water is of prior concern in Africa and Sub-Saharan Africa in particular, where only 24% of the population have access to safe drinking water [UNESCO, 2019].

Throughout each stage of a mine's life cycle, its activity can produce and release chemical and mineral pollutants prone to contaminate the surrounding air, water, and soil. Moreover, the ore extraction processes are water-demanding and need access to a water source that very often competes with the local demand, which is all the more alarming in water-stressed areas. Mining activity mostly consists in extracting small concentrations of minerals from huge volumes of rocks and therefore creates a lot of waste, which leaking is hard to avoid. For industrial mines, these wastes are diluted into water and then stored in retention ponds, where they can leak within the local environment and contaminate soil and water bodies. If low concentration levels of heavy metals can be essential for human health, the abnormal quantities found in the environment within the mine's vicinity can cause several health problems. Individuals living nearby industrial mining are exposed to high concentrations of heavy metals through ingestion, dermal contact, and inhalation of soil particles. We mainly focus on the absorption mechanism as we identify the effects of mining activity through water pollution. Exposure to heavy metals plays detrimental effects on human health in general and child health in particular, especially during their first months of development, both in and ex-utero [Coelho and Texeira, 2011]. Children are the most sensitive, even at low concentrations, as they are at a stage of rapid biological development [Dike et al., 2020], but also as they are more exposed, through higher blood concentration linked to incidental ingestion of urban soil and unclean water [He et al., 2020].

In this paper, we focus on Africa and investigate the local impacts of industrial mining

activity on health through water pollution using geocoded micro-data. We examine the 12-month and 24-month mortality as a primary health outcome, as effects on children are the most dramatic, and to capture the effects of heavy metal absorption on early-age biological development. Child mortality is a short-term measure [Greenstone and Hanna, 2014; Do et al., 2018], and is available over a long-time span of four decades and across the majority of African countries. We also look at the effects on other children’s health outcomes such as anthropometric measures and anemia, as well as women’s health and fertility outcomes. We match socio-economic and health data from the Demographic Health Surveys (DHS) with state-of-the-art geolocalized data on industrial mineral resource exploitation from the SNL Metals and Mining database, which provides information on opening dates and mineral types. Our study spans 26 out of the 54 African countries, from 1986 to 2018. We conduct a staggered Difference-in-Difference strategy exploiting the variation of the opening of a mine and the relative topographic position of surrounding villages. We indirectly isolate the mechanism of water pollution by building the treatment and control groups using an upstream-downstream comparison, which is used as a proxy for the exposure to water pollution linked to mining activity.

Our main result shows that being born in a village located downstream of an open mine increases the mortality rates under 24 months by 2.18 percentage points, which corresponds to an increase by 25% of the mortality rate. We find no significant result for the 12-month mortality rate, which suggests a lag in the effect of water pollution on early-childhood health and in the absorption of toxic elements. Our analysis suggests that this could be explained by the protection provided by breastfeeding [VanDerSlice et al., 1994; Fängström et al., 2008], as we find a significant increase in the 12-month mortality rate among children who were given plain water, in comparison to those who were not. We find effects neither on other children’s health outcomes nor on women’s health outcomes and fertility. We exclude many potential mechanisms that could explain our main result on infant mortality and show the robustness to controlling for households’ access to water and facilities such as health facilities and electricity as well as in-migration flows. Our main effect is mainly driven by the mortality of boys, and by individuals living downstream of an open-pit mine that has opened. The effect fades out with distance and increases with surrounding mine density and the intensity of production, proxied by yearly international mineral prices. It seems to be mainly occurring during the mining activity status, as it is robust to restricting the analysis to mines for which we have the closure year. Last but not least, we conduct a battery of robustness checks: the de Chaisemartin and d’Haultfoeuille [2020] estimator, using a balanced sub-sample of DHS repeated cross-sections, correcting

for DHS random displacement, restricting to the sample of mines for which we have the exact coordinates to account for measurement errors, testing for spatial spillovers, and running spatial and temporal randomization inference tests.

The major contributions of this paper are twofold. First, it lies in the construction of the industrial mines dataset, as we checked over 1,700 mines by hand to complete their opening date ¹. Our second contribution is to introduce the topographic heterogeneity of the effects of mining activity on health and to identify the negative effects induced by water pollution, and how they outweigh the positive effects. It nuances the results from the literature looking at average effects by using geographical distance to a mine as a proxy for exposure to mining activity, which finds a reduction in mortality rates [Benshaul-Tolonen, 2018; Cossa et al., 2022].

The remainder of the paper is organized as follows. Section 2 reviews the literature and presents our contribution. Section 3 describes the context and the data. Section 4 details the methodology and the main empirical strategy. Section 5 introduces the main results, while section 6 investigates the mechanisms, and section 7 the heterogeneity of the results. Section 8 looks at the dynamic effects, and section 9 at the intensive margins, digging into the heterogeneity of the results according to the distance of the mine, the mining density, and the production intensity. Section 10 proposes a list of robustness checks and placebo tests. Section 11 discusses the limits of the study and section 12 proposes a policy discussion. Eventually, section 13 concludes.

2 Literature review and contributions

This section first displays the literature on the trade-off of mining activity in developing countries. It then describes the mining-induced pollution literature and the economic literature on the health effects of mines. Thirdly, we discuss the issues emerging from this literature and the solutions we propose to tackle them.

2.0.1 Trade-off of mining activity

Our work is related to the strand of literature analysing the health-wealth trade-off of industrial mining activity in developing countries, which results are still under debate. If

¹Opening dates indicate when production first began. Data available in the SNL database was gathered by SNL from the mining companies' reports, and the hand-check work we made has completed this database by going deeper into archival mining reports

mining can improve health and well-being through local industrial development, it can also damage health through negative externalities such as conflicts, massive migration waves, and exposure to harmful pollution. Determining which of these effects is predominant is still debated in the literature studying the relevance of a natural resource curse [van der Ploeg, 2011; Cust and Poelhekke, 2015; Venables, 2016].

At a broad scale of analysis, Mamo et al. [2019] look at the effects of the discoveries of industrial mining deposits in Sub-Saharan Africa. They find an increase in district-level night light emissions but no significant effects on household wealth ². They find temporary positive effects on public service provisions but a degradation of the sewerage system and piped water supply in the medium and long run. Mining creates also negative effects on the environment and agricultural productivity. Aragón and Rud [2016] find that the expansion of large-scale gold mining in Ghana (1997-2005) is responsible for the agricultural total factor productivity decrease in the vicinity of mines. The use of cross-sectional satellite imagery of NO₂ concentration suggests that air pollution is the main explaining factor. Dietler et al. [2021] analyze a panel of 52 mines in Sub-Saharan Africa using the same DHS and SNL databases. They find improvements in access to modern water and sanitation infrastructures after a mine opens when comparing individuals living within 50 kilometers of an isolated mine. Yet, proxying exposure to mining activity with distance and focusing on areas with low mining density raises many identification issues, that will be largely discussed in Section 2.3.1. Our paper deals with these issues and encompasses a wider sample of mines. Other negative externalities of mining activity are the increase of rapacity and corruption and the trigger of insecurity and conflicts [Berman et al., 2017], migration flows of mine workers fueling the spread of infectious diseases such as HIV [Corno and de Walque, 2012], and discouragement of educational attainment among children [Atkin, 2016; Ahlerup et al., 2020; Malpede, 2021].

Our paper focuses on industrial mining and does not encompass artisanal and small-scale mining (ASM). Few papers have looked at the effects of artisanal and small-scale mining (ASM), mainly due to data limitations. Bazillier and Girard [2020] compare the local spillovers between artisanal and industrial mining sites in Burkina-Faso. They find positive impacts of artisanal mining (labor intensive and managed in common) and an absence of industrial mines' opening (capital intensive and privatized) on household consumption. Our paper focuses on the effects of industrial mining pollution. If ASM

²Household wealth was measured through the dimensions of access to electricity, wealth index, urbanization, mortality, and education.

has severe effects on miners' health due to hazardous working conditions it is likely to be of smaller magnitude than industrial mining which extracts and treats larger volumes³. If ASM is often accused of generating more severe pollution than the industrial sector because of their illegal use of mercury⁴, the latter often use cyanide instead. Both chemicals being highly toxic pollutants, focusing on industrial mining only is a lower-bound analysis of the impacts of mining activity on local populations' health.

2.1 Mining-induced pollutions

Each stage of industrial mining activity produces chemicals and minerals likely to pollute the surrounding air, water, and soil [Coelho and Texeira, 2011]. The exploration and prospecting stage can last several years before a mine is considered economically viable and worthwhile to open. Meanwhile, mining companies conduct mapping and sampling, as well as drilling, boreholes, and excavation that require both physical and chemical measurement methods likely to pollute at the surface and underground, depending on the nature of the deposit in the targeted area. If found financially viable, the company launches the discovery phase where the design and planning of the construction are undertaken, and the feasibility study of the project requires further exploration and engineering studies. Subsequently, the development stage takes place and the mine's infrastructures and processing facilities are constructed. It is only after all these stages that production can start. Once the deposit is exhausted comes the closure and reclamation stage, where the company is supposed to clean, stabilize and rehabilitate the land and isolate contaminated material. Yet, it is common that waste, tailings, or retention dams are just left abandoned without care and maintenance, and this constitutes a potential disaster if the hazardous materials are leaked and discharged into the environment. Figure 28 in the Appendix proposes a scheme to explain the life cycle of a mine. Figure 27 displays satellite images of the different stages of the Essakane mine, an open-pit gold mine in Burkina-Faso.

Throughout all these stages, different types of pollution can be engendered. Air pollutants can be carried by dust over long distances by ore transportation and the wind, they can damage surrounding soils and crops, and be inhaled mostly by mine workers but also by the local population. The leakage of pollutants in the air can also affect water through acid mine drainage that ends up polluting the surface and then groundwater. During the

³Industrial mines are responsible for 80% of the gold production and 75% of the diamond production McQuilken and Perks [2020].

⁴Mercury has been officially banned in over 140 countries (Minamata Convention on Mercury, adopted in 2013).

digging and processing to extract the targeted ore from waste rocks, rocks are crushed and then go through either heap leaching, froth flotation, or smelting. These techniques require the addition of chemicals such as cyanide or acid, that can separate the targeted minerals from waste. Moreover, these processes are water-demanding and need access to a water source in competition with the local demand. Last but not least, even without the use of these chemicals, leaching happens through the contact of water and oxygen with sulfide minerals contained in the extracted rocks , which accelerates the acidification process and modifies the pH levels of water bodies. Pollutants can be released into the environment during the process by spills or after by leaks of humid waste stored in retention dams but also through the erosion and the sedimentation of solid wastes that are piled in the tailings around the mining site and that drain to the soil with rain. The wastes actively pollute during the whole life cycle of the mine, starting from its opening and during production, but also can continue to pollute when a mine closes and is left without maintenance. This is the case when retention ponds are not covered and dry, letting these wastes go directly through the environment.

Few papers have managed to show to what extent industrial mining activity creates negative externalities on the environment. Bialetti et al. [2018] look at the effects of mining industries on deforestation in India, Von der Goltz and Barnwal [2019] have suggested the mechanism of water pollution but without strong empirical evidence (looking at anemia). Yet, in-situ measurements have shown the contamination of water drinking sources by harmful levels of nitrate, turbidity, iron, cadmium, manganese, and arsenic by industrial mining sites [Cobbina et al., 2013]. To our knowledge, we are the first to provide indirect, systematic, and large-scale evidence of the mechanism of water pollution by industrial mining activity.

The main toxic metals released by mining sites are arsenic, cadmium, copper, lead, mercury, and nickel. Depending on their blood level concentration, they can be essential or non-essential for human health [El-Kady and Abdel-Wahhab, 2018]. However, heavy metals released by mining activity are non-biodegradable, have long-term impacts on the environment, and are found at abnormally high concentrations in the vicinity of mines, within the soil, water resources, vegetation, and crops [Oje et al., 2010; Dike et al., 2020]. People living in that environment are exposed to high quantities of heavy metals through ingestion, dermal contact, and inhalation of soil particles, which can cause several implications for their health. High blood metal concentrations are associated with neurological effects (which induce behavioral problems, learning deficits, and memory losses, especially

among children) [Dike et al., 2020], neurodegenerative diseases, cardiovascular effects, gastrointestinal hemorrhages [Obasi et al., 2020], organ dysfunction (kidney, decrease of the production of red and white blood cells, lung irritation) [Briffa et al., 2020], higher probability of cancer development [Madilonga et al., 2021; Obasi et al., 2020], but also a higher probability of infertility, miscarriages for women, and malformation of newborns [Briffa et al., 2020]. Thus, exposure to heavy metals plays detrimental effects on human health in general and child health in particular, especially during their first months of development, both in and ex-utero [Coelho and Texeira, 2011]. Besides, children at an early age are the most sensitive, even at low concentration, as they are at a stage of rapid biological development, but also as they are more exposed, through higher blood concentration linked to incidental ingestion of urban soil and dirty water (less conscious of their environment and danger, playing with polluted soil, eating and drinking without care [He et al., 2020]).

2.2 Health effects of mining activity

The empirical economic literature on the local effects of mining on local communities has been growing during the past decade, yet the debate remains on the costs and benefits, and the positive and negative impacts of industrial mining activity in developing countries. Diverse results have been found on the effects on health, and there is still uncertainty on the direction and the magnitude of the impacts of mines on the local population's health. Besides, if geographical proximity to a mining site is usually used as a proxy for pollution exposure, few papers observe the negative externalities on the environment and its consequences on health.

Papers studying the effects of industrial mines on health proxy the exposition to mining activity by the distance to the mine and can be found in the literature different thresholds and mixed results. Using cross-section data in the state of Orissa in India, Shubhayu et al. [2011] uses the distance to the mine as a proxy to measure environmental effects, and finds that individuals living near a mine report higher respiratory illness and more work days lost due to malaria. Cross-sectional data prevent identifying a clear causal relationship and from adjusting to specific time and spatial confounders. Benshaul-Tolonen [2018] uses a Difference-in-Difference strategy, comparing individuals living within 10 kilometers to those living between 10-100 kilometers of a mine, before and after its opening. The paper finds that large-scale gold mining in nine countries of Sub-Saharan Africa⁵ decreases infant mortality within 10 km during the opening and operating phases, with no effect on

⁵Burkina Faso, Ivory Coast, the Democratic Republic of the Congo, Ghana, Guinea, Ethiopia, Mali, Senegal, and Tanzania between 1987 and 2012

further communities (10-100km). Cossa et al. [2022] uses a similar methodology studying a broader set of countries and finding a decrease in child mortality as well.

Von der Goltz and Barnwal [2019] assess the effects of industrial mines in 44 developing countries from 1988 to 2012. The paper also relies on a Difference-in-Difference strategy, comparing households living within 0 to 5 km to households living between 5 and 20 km before and after the opening of a mine. They find gains in asset wealth, increased anemia among women, and stunting in young children. As anemia and growth deficits are argued to be mainly the consequences of exposure to lead, the observed effects on health are interpreted to be the results of pollution due to metal contamination and lead toxicity. They find that women in mining communities show depressed blood hemoglobin, recover more slowly from blood loss during pregnancy and delivery, and that children in mining communities suffer some important adverse growth outcomes from in-utero exposure (stunting).

2.3 Challenges and contributions

The most common way to proxy exposure to mining activity is to rely on the distance to an active or open mine, however, there is no clear consensus on which threshold to use, and the treatment allocation seems arbitrary. The disparities in the results from the literature could be explained by these differences in terms of empirical strategies and distance choices. Beyond this, using the Euclidian distance to a mine as treatment raises endogeneity concerns. This subsection discusses the main issues arising when studying the local impacts of industrial mining activity on health.

2.3.1 Endogeneity issues

In this section, two challenges are discussed: the endogeneity issues that arise (i) when using the Euclidian distance as a proxy for exposure to mining activity, and those when using (ii) repeated cross-sectional data such as the DHS.

Using the interaction between being close to a mine and the mine's activity status raises endogeneity concerns. For instance, Von der Goltz and Barnwal [2019] use a mine panel and pairs each DHS village to its closest mine. This creates unbalanced treatment and control groups, and such imbalance might be endogenous to socio-economic outcomes or polluting behaviors. As each village is paired to its closest mine, this *de facto* excludes from the control group villages that are in both distance categories (within 5 km of mine

A but 5-20 km of mine B). Thus, there is a higher probability to be treated in areas with high mining density, which is not a random allocation. As a mine fixed-effect identification relies on a within-mine buffer-area comparison, the estimator is driven by mines that have been paired to villages both in the treated and control areas, which is correlated to the mining density of the region. The estimation endogenously selects mines from regions of low or middle mining importance, which might be correlated with the intensity or the type of pollution, the socio-characteristics of the neighboring population, and thus the way health is affected by pollution. To reduce endogeneity issues, Von der Goltz and Barnwal [2019] instrument the mine location with mineral deposit information from S&P, which are *deposits that are being explored or prepared for exploitation*. However, mining exploration is not a random allocation and raises the same concerns as it is directly correlated to mining density. Benshaul-Tolonen [2018] reduces endogeneity issues linked to the pairing by using an administrative district fixed-effect panel and extending the distances (10-100km), but the same concern remains.

A second concern is linked to the nature of the DHS data, which are repeated cross-section surveys. The literature argues that the conditions for an industrial mine to settle are the presence of mineral deposits, which is considered random. However, the presence of a mine and of a declared mineral deposit is correlated to the population density. As mining exploration is labor intensive, it is more likely to occur in dense areas, where DHS is more likely to have surveyed individuals. A treatment allocation based on geographic proximity to the mine is endogenous: treatment groups close to the mine might not be comparable to control groups located further. As district fixed-effect relies on a within-district comparison, the estimation is driven by districts with both control and treated groups, before and after a mine opening, which is correlated to the probability of being surveyed. As DHS renews the surveyed villages at each wave, and as the probability to be surveyed is determined by the population density, the estimation is driven by specific areas. The regression *de facto* and endogenously selects districts that were already dense before the opening and remained after. This might be areas that are more stable, well-off, and where individuals might be less affected by pollution. This might bias the estimation upward (i.e less mortality linked to mining activity), and explain the positive effect of mines that Benshaul-Tolonen [2018] finds on mortality in Africa.

In Appendix section G, we propose a replication analysis of Benshaul-Tolonen [2018], taking advantage of our handwork which extends the SNL database. We find similar results as Benshaul-Tolonen [2018] using the same set of countries and our extended

sample of mines (only gold mines as in the paper). However, we find no longer significant results when applying to our more comprehensive sample, meaning when including other African countries and industrial mines, which suggests that the effects are context and regional-dependent.

2.3.2 Upstream-downstream analyses

Using geographic distance to a mine as treatment allocation raises endogeneity concerns. An upstream-downstream analysis, which relies on a topographic comparison, reduces these concerns as individuals are compared from similar distances.

Few papers have dealt with upstream and downstream at the scale of a continent, since it requires much more computational capacity and a complex pairing methodology. Duflo and Pande [2007] study the productivity and distributional effects of large irrigation dams in India and use river networks and calculate gradients computed from digital elevation maps for India. Do et al. [2018] use river networks and pollution monitoring stations data in India to conduct their upstream-downstream analysis. Unfortunately, it is impossible in our case study due to the absence of water quality data at the scale of Africa. Garg et al. [2018] use river networks in Indonesia and re-calculate the upstream-downstream relationship between village pairs using a 30m resolution Digital Elevation Model. Their very refined level of study is not likely to be undertaken at the scale of the African continent in our case, so we choose secondary data computed by hydrologists (HydroSheds). We use systematic and highly disaggregated data on water sub-basins that enable us to encompass a wider set of countries, overcoming the issue of pairing a mine or a village to the closest river, since there is uncertainty about whether this point is located above or below in altitude compared to the level of the river. Strobl and Strobl [2011] studied the distributional effects of large dams on agricultural productivity at the scale of the African continent, using Pfafstetter level 6 with an average area of 4200 km². Our study takes into account sub-basins at the Pfafstetter level 12, with an average area of 100 km².

3 Data and Context

This section describes the data used for our empirical strategy, and some descriptive statistics on the context of industrial mining and child mortality in Africa.

3.1 Data

In this paper, we match socio-economic data from the Demographic Health Surveys to an industrial mining database provided by SNL Mining and Metals.

3.1.1 Health and socio-economic data

We use all available survey rounds from the Demographic Health Surveys that contain GPS coordinates, from 1986 to 2018, covering 36 out of the 54 African countries. We then select all the countries which have at least two survey waves to be able to implement our Difference-in-Difference strategy with a sufficient time length before and after the opening of a mine and end up with 26 countries⁶, 12,442 clusters and 240,431 children under the age of five. We consider that doing a Difference-in-Difference strategy on the sample of countries that have only one round of the survey, hence a maximum of five years period, will not enable us to capture the longer-term effects of mining activity⁷. Table 22 in the Appendix displays the DHS survey years and countries that we use for the analysis.

We construct the variables of child mortality based on the DHS child recode database which has information on the age and death of children under five years old, whose mothers are aged between 15 and 49 years old. Our dependent variable is the probability of 12-month and 24-month mortality for each DHS cluster (i.e. for each child, we build a dummy equal to 1 if she or he is alive and 0 if not, conditional on having reached 12 and 24 months respectively). We also estimate the effects of mining activity on biomarker variables and other indicators of occurrences of illness (diarrhea, fever, and cough) within two weeks preceding the day of the interview among young children. We extend our analysis to women's fertility behavior and health: current pregnancy, total lifetime fertility, miscarriage, and anemia. Finally, as the aim of this article is to isolate the mechanism of water pollution, we use the questions from the DHS on the main source of drinking water, the presence of flushed toilets, electricity, and the access to health facilities to control for households' sanitary and economic environment.

⁶The list of countries within our sample are: Benin, Burkina Faso, RDC, Burundi, Cote d'Ivoire, Cameroon, Ethiopia, Ghana, Guinea, Kenya, Liberia, Lesotho, Madagascar, Mali, Malawi, Nigeria, Niger, Namibia, Rwanda, Sierra Leone, Senegal, Togo, Tanzania Zambia, and Zimbabwe

⁷Please note that our final sample does not include Egypt which has 7 DHS waves and is a well-known mining country. This is explained by the fact that the SNL database characterized Egypt within the Middle East rather than in Africa and thus was dropped from our sample.

3.1.2 Mineral resource exploitation data

The industrial mining variables come from the SNL Metals and Mining database, which is privately owned by *S&P Global* and on license⁸. The SNL database is the best existing panel of mine production, providing information on the location, the dates of opening and closure, the commodity type, and the yearly production (for some mines). This is a non-exhaustive panel of industrial mines in Africa, yet to our knowledge, it constitutes the most comprehensive sample of mines giving the timing of the industrial activity. This dataset has been intensively used in the literature and argued to be the best product available [Aragón and Rud, 2016; Berman et al., 2017; Kotsadam and Tolonen, 2016; Benshaul-Tolonen, 2018; Von der Goltz and Barnwal, 2019; Mamo et al., 2019]. We emphasize here that this paper focuses on the effects of industrial mining, and that we do not include artisanal mining (ASM) that are not available in the SNL database.

Overall, the SNL database gathers 3,815 industrial mines in Africa from 1981 to 2021, and 2,016 were located within 100 km of a DHS cluster from a country with at least two surveys. For our difference-in-difference strategy, we need information on the timing of the beginning of the mining production. The SNL database gives this information for 278 mines and we retrieved from handwork the start-up year for the 1,738 remaining mines. The hand-check was realized using the information on the mining history available in the SNL database and mine reports (cross-checked with Google Maps and aerial images). We describe this handwork more extensively in the Appendix B.2.

We build three main variables from the SNL Mining and Metals database, relying on the geocoded information and the time of opening. According to the estimation strategy, we will use a variable of proximity (distance to the closest mine), position (whether individual i is upstream or downstream), and a dummy for being open or not. Opening dates that were available in the SNL database were computed by the SNL team, and indicated the actual start-up year of the mine, i.e when production first began. We used the same criteria for our handwork. Finally, we restrict the main analysis that is associated with heavy metal mines (metals with density higher than $5g/cm^3$ [Briffa et al., 2020], which are the metals listed in Table 26 in Section C.2 of the Appendix. We also include coal mines, as their extraction is associated with mercury and arsenic which are highly toxic heavy metals.

⁸We are grateful to CEPREMAP, PjSE, EHESS, and the GPET thematic group of PSE, for their financial support and their help in purchasing the access to the data.

3.1.3 Water basins

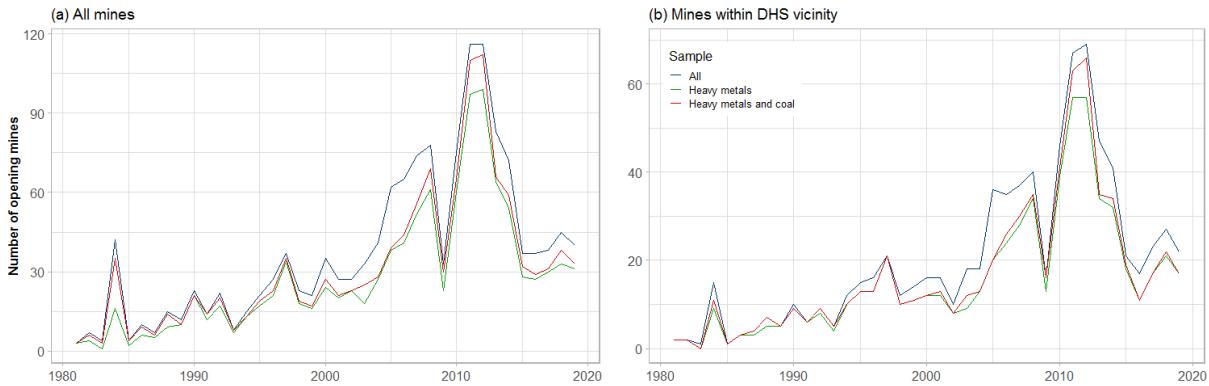
We consider the topographic relationship of water basins where mines and villages are located. A water basin is an area where all the surface water converges towards the same point. We use the HydroBASINS sub-basins geographic information provided by HydroSHEDS, which delineates water basins consistently and subdivides sub-basins into multiple tributary basins to the network of nested sub-basins at different scales. Following the topological concept of the Pfafstetter coding system, each polygon of the sub-basin has a unique direction flow and provides information on the up-and down-stream connectivity. We take the finest Pfafstetter level (12 out of 12) that breaks down sub-basins at an average area of 100 km^2 . See Figure 6a for an example. We conduct our analysis taking into consideration the three closest sub-basins to each industrial mine, meaning that we take each mine's sub-basin A and tag the one just downstream that we call B, the one just downstream of B that we call C, and then the one just downstream of C that we call D. Thus, B, C, and D are the three closest sub-basins of A.

3.2 Descriptive statistics

3.2.1 Mining in Africa

Temporal and spatial variation

Figure 1: Temporal evolution of mine opening



Notes: The Figures plot the number of mines opening each year over the 1981-2019 period, for all mines, heavy metal mines including coal (sample of the main analysis), and only heavy metal mines. Figure (a) displays the temporal evolution of the total mine sample, while Figure (b) of mines that are within the sample of the main analysis, meaning mines that have DHS clusters upstream at most at 100km and DHS clusters downstream within the three closest sub-basins.

Sources: Authors' elaboration on DHS and SNL data.

Figure 1 shows the evolution of the yearly number of mines that opened in Africa over the 1981-2019 period, Figure 1 (a) for the entire mining sample while Figure 1 (b) for the mines that are in the sample of the main analysis. The mining boom since 2000 is captured in the Figures, with the first peak in 2007, in line with the peak in exploration activity that occurred in 2003 [Taylor et al., 2009] (as the exploration phase is on average a couple of years before a mine opens), and the second one in 2012. For instance, around 120 industrial mines opened in 2012 (based on the non-exhaustive SNL database). The Figures also distinguish the evolution according to the mines' characteristics: it distinguishes the pattern for all mines, heavy metal mines, and heavy metals including coal mines. We observe no differences in timing patterns between Figure 1 (a) and (b), neither between mine types.

What is striking in Figure 1 is that the evolution of mine openings follows the same pattern as the evolution of industrial metal prices, as plotted in Figure 29 from Section C.2 in the Appendix. The mining boom since 2000 follows the increase in real prices of Copper, Tin, Lead, Aluminum, Zinc, Nickel, and other heavy metals, while the sharp fall around 2008/2009 corresponds to the financial crisis. Again, the local minimum around 2016 corresponds to the drop in commodity prices in June 2014 [Khan et al., 2016; Glöser et al., 2017]. This similar evolution suggests that heavy metal prices are good Instrument Variables for the variable year of mine opening, such as Berman et al. [2017]; Bazillier and Girard [2020] used in their analysis. In Section 9.3 we will use it as a proxy for production intensity.

Figure 2 (c) shows the map of the number of mines that have opened before 2019, including mines that opened before 1986, averaged at the cell level (160 km cells). Cells in grey represent areas where no mine opened before 2019, but where at least one will open in the future (whether we know from the data that it has opened between 2019-2021, or if the opening is planned further). The main mining countries in the SNL database are Guinea, Sierra Leone, Ivory Coast, Ghana, Niger, Burkina Faso, Zimbabwe, Tanzania, Zambia, and the north of South Africa. Please note that, as we exclude countries with only one DHS wave in our main analysis' sample (cf Tables 21 and 22), to avoid comparing areas with too many differences in terms of temporal variations, we did not undertake the hand work for these countries, which explains why South Africa (which is not in the final sample) does not appear as a major mining country in Figure 2 (c). Figure 3 shows both the temporal and spatial variation of mine opening in Africa (for all the mines sample, and not the restricted one for our main analysis), as it plots the number of mines that opened

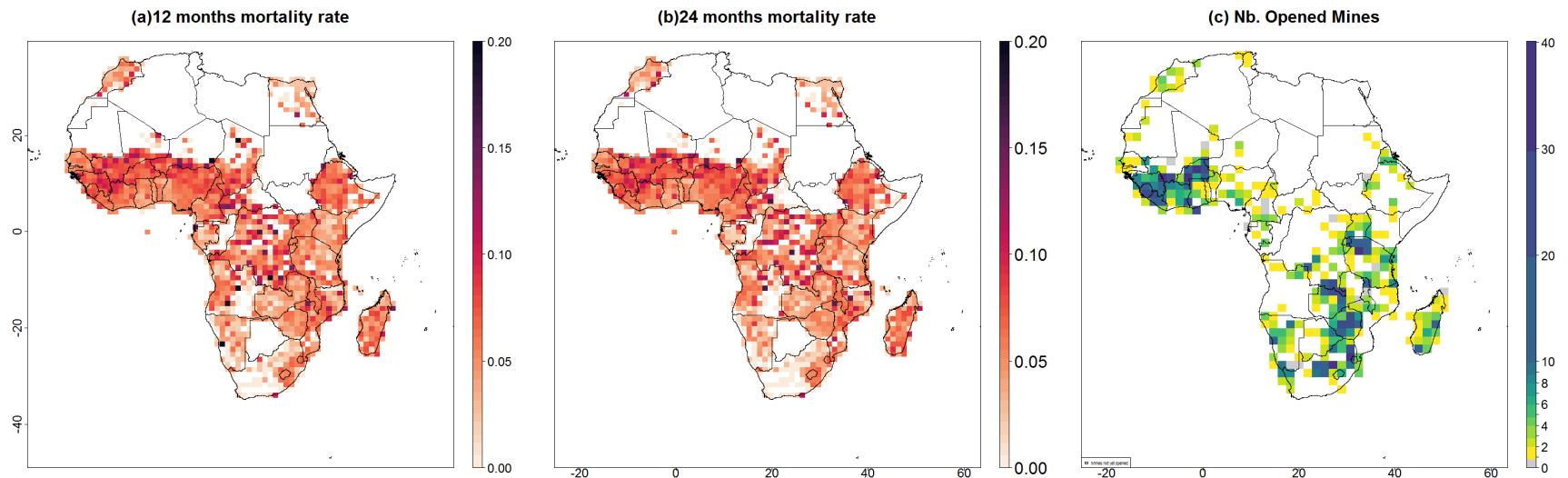
over different periods of our analysis per grid cell. The cells in red are areas where no mines opened during the period, but where at least one mine has opened before, whereas cells in grey are areas where no mines have ever opened while at least one will open in the future. We observe that the increase in mine opening was higher during the third period 2008-2019 (which is in coherence with Figure 1), and was particularly important in West Africa.

3.2.2 Health risks

Africa faces high infant mortality rates, as the average 12 months mortality rate is 6.4 % and the average 24 months mortality rate is 8.3% according to DHS data (cf Table 23). Figures 2 (a) and (b) plot the average mortality rates for all DHS from 1986-2019 averaged at the grid level, and show the spatial variation of mortality rates⁹. Figures 4 and 5 map both spatial and temporal variation of mortality rates as it shows the average mortality rates for the three main periods of our DHS sample. We can observe the global reduction of mortality over the period and also the DHS cluster distribution. Figures 31, 32 and 33 plot the same maps for the sample restricted to the one used in the main regression.

⁹Please note that the higher the DHS cluster density, the more accurate the average. The spatial variation is endogenous to the DHS sample.

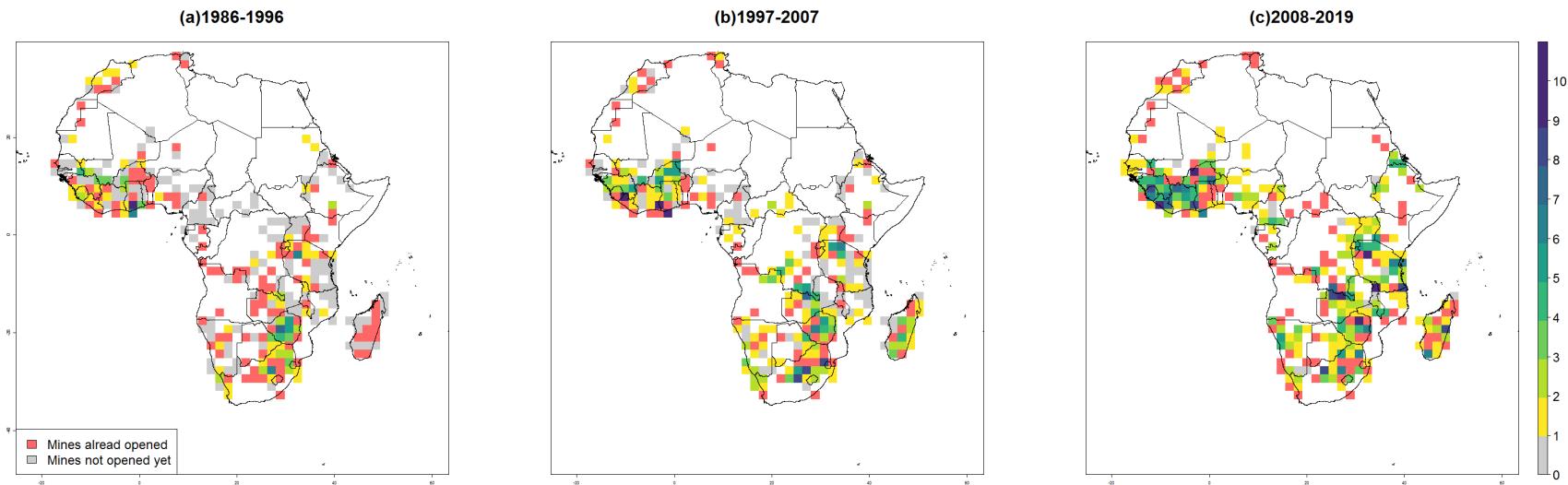
Figure 2: Outcomes spatial distribution



Notes: Figures (a) and (b) represent the means of 12- and 24-month mortality rates for each DHS wave available (listed in table 22), from 1986 to 2019. Means are computed at the grid level (100km mean size). The mortality rates are estimated without the children that did not reach 12/24 months at the time of the survey. Figure (c) displays the stock of mines that opened before 2019 (including mines that opened before 1986). Means are computed at the grid level (100km mean size).

Sources: Authors' elaboration on DHS and SNL data.

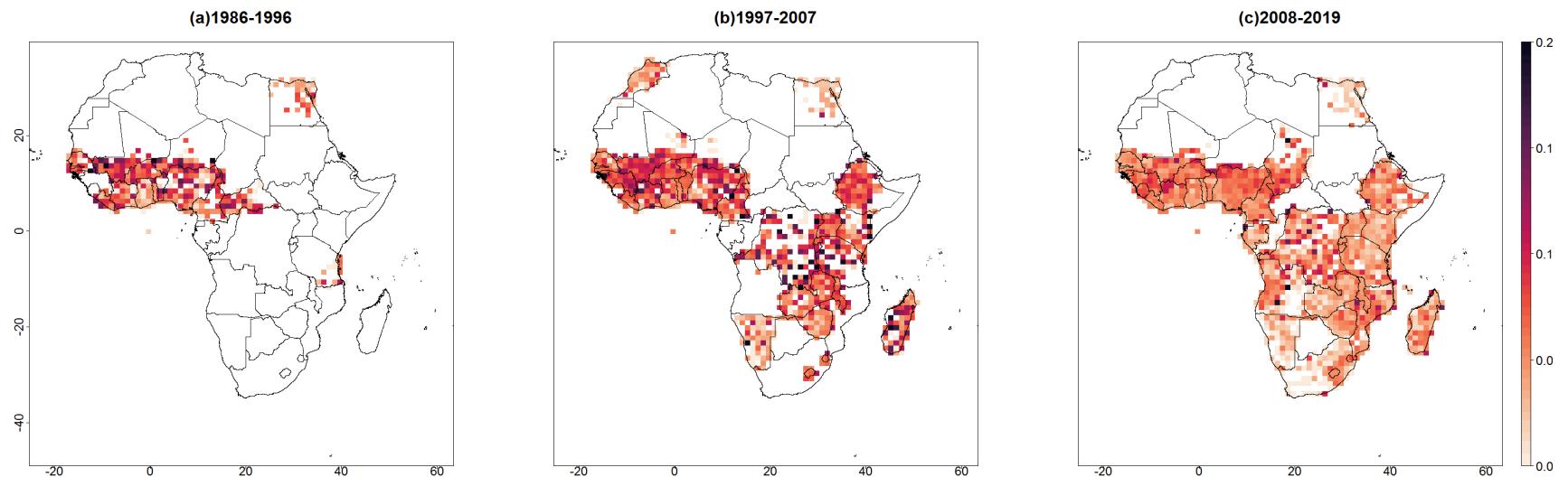
Figure 3: Spatial variation of mine opening per period



Notes: The figures represent the number of mines that opened during the periods over the grid area (160 km on average). A red grid cell represents an area where no mine opened over the period, but where at least one mine open before the period. A grey cell represents an area where no mine opened over the period, but where at least one mine will open in the future.

Sources: Authors' elaboration on SNL data.

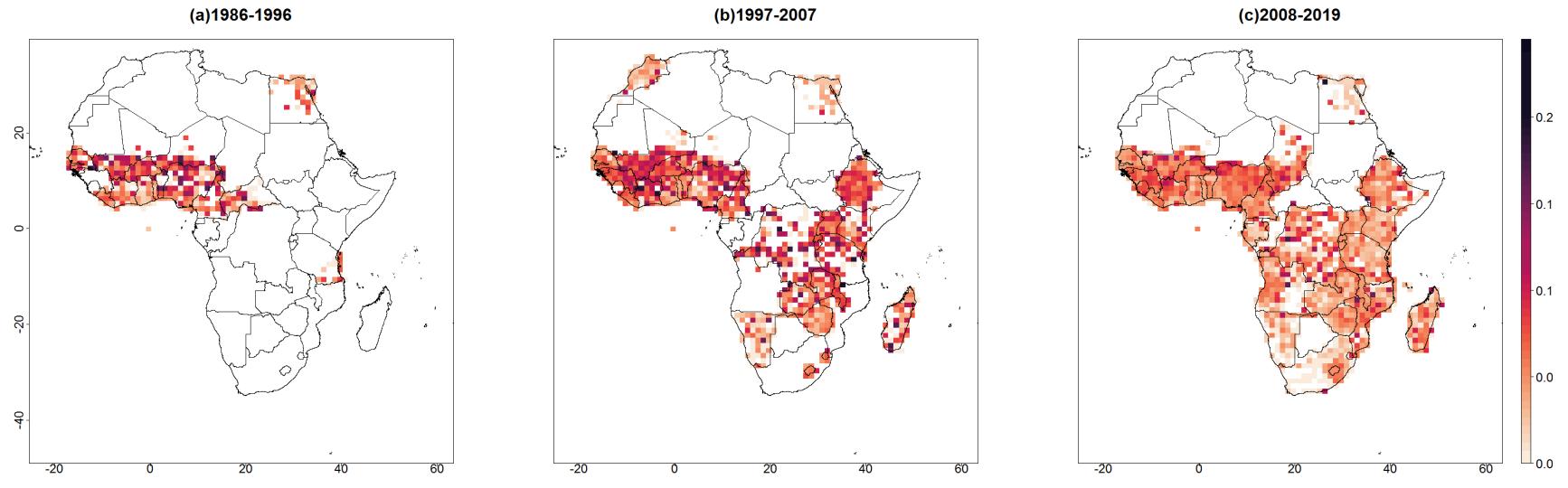
Figure 4: Spatial variation of 12-month mortality rates per period



Notes: The figures represent the means of 12-month mortality rates averaged at the grid level over (a) 1986-1996, (b) 1997-2008, and (c) 2008-2019. The mortality rates are estimated without the children that did not reach 12 months at the time of the survey.

Sources: Authors' elaboration on DHS data.

Figure 5: Spatial variation of 24-month mortality rates per period



Notes: The figures represent the means of 24-month mortality rates averaged at the grid level over (a) 1986-1996, (b) 1997-2008, and (c) 2008-2019. The mortality rates are estimated without the children that did not reach 24 months at the time of the survey.

Sources: Authors' elaboration on DHS data.

4 Empirical strategy

The main empirical strategy of this paper uses the relative topographic position of sub-basins as a proxy for exposure to mining activity pollution. It compares the effects on the health of individuals living downstream to those living upstream of a mine, before and after the opening of at least one site. It is a staggered design Difference-in-Difference analysis with two-way fixed effects at the mine's sub-basin and birth year level. This upstream-downstream strategy intends to identify the mechanism of water pollution.

As seen in Section 2.3.1, this strategy alleviates some endogeneity issues raised by treatments using the Euclidian distance as a proxy for exposure to the mine. First, it reduces the bias linked to unbalanced samples due to endogenous pairing. Second, it breaks the average effects based on distance buffers and highlights the heterogeneity of the effects of mining activity on health, and isolates the negative externalities linked to water degradation.

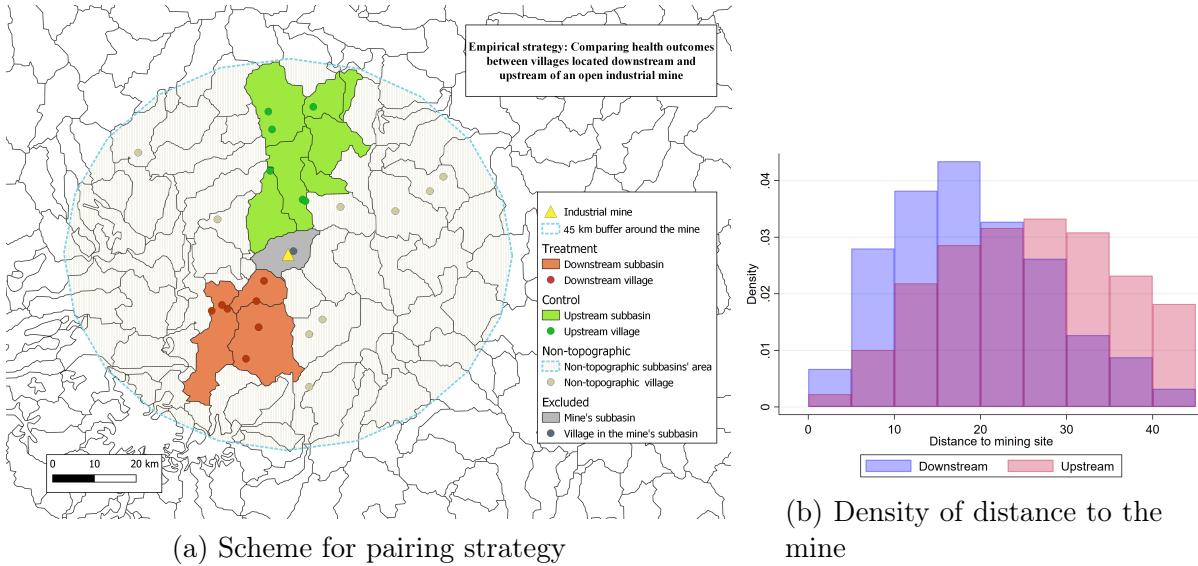
4.1 Measuring exposure to pollution

4.1.1 Pairing strategy

The pairing of DHS clusters to mines represents a significant challenge, as each DHS cluster can be downstream of and close to several industrial sites in major mining areas. It introduces endogeneity in the sample selection and raises the issue of unbalanced samples. In this analysis, we propose the following pairing to overcome this issue and thus be able to measure the exposure to pollution of each DHS cluster.

First, we construct a 100 km buffer around each DHS cluster and register all mines within this buffer (independently of their activity status). We then categorize the topographic position of the DHS cluster relative to the industrial site, using a dummy equal to 1 if the cluster is downstream of the mine and 0 if it is located upstream. This topographic position is defined using the relative position of each sub-basin. As each cluster and sites have GPS coordinates, they lie in a specific sub-basin, and we used the relative position of each sub-basin to classify the DHS according to the paired mine. Through such a process, we also have pairs that are located in the same sub-basin, and for which it is impossible to say exactly whether the cluster is downstream or upstream of the mine. At this stage, for these specific couples, we consider the DHS to be downstream. Please note that, as explained in section 3.1.3, we used the finest Pfafstetter level 12 that breaks

Figure 6: Pairing Strategy



Notes: Figure (a) is a scheme that illustrates the pairing, giving the example of a mine, its main sub-basin, its three closest downstream and upstream sub-basins, and DHS clusters that are in the treatment and control areas within 45 kilometers. Figure (b) plots the density of the distance (in km) to the mining site for DHS clusters across their upstream-downstream position.

Sources: Authors' elaboration on DHS, SNL, and HydroSheds data.

down sub-basins at an average area of 100km^2 (the size of the sub-basin varies according to their shape, cf Figure 6a). At this stage, some villages can be paired with several mines and can have more than one occurrence in the sample. The difficulty of the strategy lies in choosing the mine that will be paired with the cluster.

Second, we restrict the group of downstream DHS clusters to the ones that lie within one of the three closest sub-basins downstream of the mine's sub-basin, to focus on the potentially most contaminated areas. Third, we pair each cluster with only one mine, proceeding as follows. If a DHS was in both groups (i.e downstream a mine A and upstream a mine B), then it is automatically assigned to the downstream group, and it is paired to the mine from which it is downstream (i.e it is paired to mine A), regardless of its activity status. At this stage, some clusters may still be counted twice, as they can be upstream of several mines, or in the three closest sub-basins downstream of several mines. To complete the uniqueness of the pairing, we paired each cluster to the nearest mine, regardless of its activity status as well.

In conclusion, the DHS clusters are attached to the nearest mine from which they are

downstream up to the third sub-basin level, or else attached to the nearest mine upstream up to a radius of 100km. The final remaining problem relates to the clusters that are in the mine’s same sub-basin, which we have so far identified as being downstream. We eliminated from the main analysis all DHS villages which are located in the same sub-basin of the mines from which they were paired. Also, this reduces the noise linked to the random displacement of DHS villages (cf Section 10.3.1) and avoids allocating villages as being downstream whereas they are upstream due to the displacement, as it drops the closest areas around the mine.

Once the pairing is done, we restrict the control group to upstream villages that are within 45 kilometers of the mine, to ensure the comparability of upstream and downstream villages. To choose this distance cut-off, we have calculated the mean of the maximal distance between a mine and the furthest extremity of its third downstream sub-basin, which was 44.7 kilometers. Figure 6b plots the distribution of the distance to the mine for both upstream and downstream villages. As downstream villages are prioritized in the pairing strategy, they are slightly closer to the mine, but the two distributions are comparable.

The pairing is illustrated in Figure 6a. It gives the example of a mine, its main sub-basin (grey), the downstream sub-basins (orange), and upstream sub-basins (green) up to 45 kilometers. The dashed area displays the sub-basins within 45 kilometers, with no topographic relationship to the mine, meaning they are neither downstream nor upstream. In the main strategy, we compare the villages within the green area to those in the orange area. In section 9 we run robustness tests checking whether the results hold allowing for further sub-basins and heterogeneity effects by distance to the mining site. In section 10.2.1 we discuss the results when including the non-topographic sub-basins.

4.2 Identification Strategy

4.2.1 Main estimation

The main analysis relies on a Difference-in-Differences strategy using the topographic position of a DHS cluster relative to a mine deposit to indirectly identify the channel of water pollution. We propose a staggered Difference-in-Difference specification (DiD), with a sub-basin fixed effect panel for each mine. We isolate the mechanism of water pollution by building the treatment and control groups using an upstream-downstream comparison. We restrict our analysis using the pairing strategy explained in the previous section 4.1.1. We compare health outcomes in upstream-downstream areas, both before and after

the opening of the paired mine. The empirical strategy can be formally written as follows:

$$\begin{aligned} Death_{i,v,c,SB} = & \alpha_0 + \alpha_1 Opened_{birthyear,i,v} + \alpha_2 Downstream_{v,SB} \\ & + \alpha_3 Opened_{birthyear,i,v} \times Downstream_{v,SB} + \alpha_4 X_i \\ & + \gamma_{SB} + \gamma_{SB-trend} + \gamma_{c,birthyear} + \epsilon_v \end{aligned} \quad (1)$$

With $Death_{i,v,c,SB}$ a dummy equal to one if child i from DHS village v of country c , has reached the n^{th} month and has died (n being 12 for the 12-month old mortality, same for 24 months). $Opened_{birthyear,i,v}$ is a dummy equal to 1 if the mine, which is located in sub-basin SB , has opened before child i 's year of birth. $Downstream_{v,SB}$ is a dummy of relative position (equal to 1 if village DHS v is located in a sub-basin downstream of the mine sub-basin SB , and 0 if it is upstream), X_i a vector of child and mother level controls (mother's age, age square, years of education, urban residency). Finally, γ_{SB} is a mine sub-basin fixed effect, $\gamma_{SB-trend}$ a mine sub-basin linear birthyear trend and $\gamma_{c,birthyear}$ a country-birthyear fixed effect. This analysis is a staggered design as the treatment shock (mine opening) does not occur at the same time for each DHS cluster.

The main regression is run without the DHS clusters that lie within the same sub-basin as the mine they are coupled with, as discussed in the previous section. The list of countries and survey years used in the main regression are given in Table 22, and the list of metals in Table 26.

4.2.2 Identification assumption

The key assumption of a DiD is that the downstream group would have evolved as the upstream group in the absence of the opening of a mine. As we cannot test those upstream and downstream areas would have followed the same time trends, we test in Section 8 the common trend assumption using pre-treatment data.

However, the fact that pre-treatment data are parallel is neither a necessary nor a sufficient condition for the identification. Past trends can be identical but the upstream group may be affected by a group-specific shock during the period of the treatment. The estimation of this paper relies on the fact that the comparison between downstream and upstream villages is a proxy for exposure to water pollution. The major identification assumption is that the opening of a mine affects differently upstream and downstream areas only through the decrease in water quality. Throughout the paper, we will try to address the concerns of unobservable factors that might not be orthogonal to our treatment and

Section 11 displays a final general discussion on the threats to the identifying assumption, and how they have been solved in the analysis.

4.3 Descriptive statistics

In this section, we describe the balance tables for the outcomes that play a key role in our analysis, out of parsimony.

Table 1: Balance Table

Before Mine Opening						After Mine Opening						Within Up.	Within Dwn.	Within	
Upstream			Downstream			Diff	Upstream			Downstream			Diff		
N	Mean /(SD)	N	Mean /(SD)	(4-2) /(p.v)		N	Mean /(SD)	N	Mean /(SD)	(9-7) /(p.v)		(7-2) /(p.v)	(9-4) /(p.v)	(12-11) /(p.v)	
(1)	(2)	(3)	(4)	(5)		(6)	(7)	(8)	(9)	(10)		(11)	(12)	(13)	
Dth<12															
All	23,547	0.073	7,875	0.074	0.001		12,319	0.055	4,738	0.051	-0.004	-0.018	-0.023	-0.005	
		(0.261)		(0.262)	(0.83)			(0.228)		(0.219)	(0.256)	(0)	(0)	(0.468)	
Mines	244		237				179		183						
Dth<24															
All	17,726	0.096	5,928	0.098	0.002		8,664	0.068	3,330	0.072	0.004	-0.028	-0.026	0.002	
		(0.294)		(0.297)	(0.618)			(0.252)		(0.259)	(0.428)	(0)	(0)	(0.671)	
Mines	244		236				168		168						

Notes: Standard errors and p-values in parentheses. Descriptive statistics of 12- and 24-month mortality outcomes, for villages located upstream and downstream of mining sites, for individuals born before and after the opening of the mine.

Balance Table 1 compares the changes in infant mortality before and after the opening of a mine, for places upstream *vs* downstream of the mining site, following the pairing strategy. It displays also the number of individuals and paired mines in each group of the analysis. On average, upstream and downstream areas have non-significant differences in terms of 12 and 24-month mortality (columns 5 and 10). For both upstream and downstream clusters, the opening of a mine significantly decreases the mortality probability (columns 11 and 12), which is in line with the result of Benshaul-Tolonen [2018], and with the fact that mortality rates decrease over time in Africa, as trends are not included (Figures 31 and 32). Table 1 shows that this reduction is overall slightly more important in upstream areas than in downstream areas for under 24-month mortality (0.002), while it is the contrary for under 12-month mortality (-0.005) but they remain not significant differences (column 13). Table 1 does not include any controls and is only descriptive. Table 27 from Section D.1 in the Appendix replicates this exercise for control variables.

Figure 30 in the Appendix identifies the country with the biggest stock of open mines in our sample (Ghana, Zimbabwe, and Tanzania with the highest density of open mines nearby DHS), as well as insights on the variation in mine opening over the period per country.

5 Main results

This section displays the results of our main analysis. The first section describes the overall effects of mining opening on child mortality among the villages living downstream compared to those living upstream. The second section displays the effects of being downstream of an open mine on other child's health outcomes, while the third section focuses on women's outcomes.

5.1 Child mortality

This section displays the main results of this paper from equation 1 for the 12-month mortality rate and the 24-month mortality rate. Table 2 gathers our main results with mine sub-basin and country-birth-year fixed effects. We also include mine sub-basin and birthyear linear trends, adjusting for spatial and period-specific cofounders and trends, and commodity fixed effects. Columns (1) to (4) give the results for the 12-month mortality rate, while columns (5) to (8) for the 24-month mortality rate. Columns (1), (2), (5), and (6) display the results for the total population while columns (3), (4), (7), and (8) focus on the rural population. Control variables are birth order, mother's age, mother's age square, the mother's years of education, urban, and the intensity of the river network.¹⁰. Even columns include the number of open mines within 45km of the DHS cluster as control, which controls for the mining density.

The results show that being downstream of an open mine increases by 2.18 percentage points (p.p) the 24-month mortality rate¹¹. This corresponds to an increase by 25% as the average 24-month mortality increases from 8.7% to 10.9%. The results are higher in terms of magnitude in rural areas, as being downstream an open mine increases by 3.8 p.p the 24-month mortality rate, which is associated with an increase by 40%, as the mortality increases from 9.4% to 13.2%. This is in line with the fact that rural populations have less access to facilities and infrastructure and are more exposed to unsafe water. The results are not significant concerning the 12-month mortality rate, and are very close to zero, showing no difference between individuals leaving upstream to those leaving downstream. This lag in the effect of water pollution on children's health may be explained by the higher probability of children under 12 months to be breastfed compared to children under 24-month, hence their decreased exposure to contaminated water and limitation of direct ingestion [VanDerSlice et al., 1994; Fängström et al., 2008]. This mechanism explaining the different results on the 12-month mortality *vs* 24-month mortality is explored in Section 6.3.

¹⁰The variable intensity of the river network using the HydroRIVERS product. It is a continuous variable, which takes into account the area of the catchment that contributes directly to a river reach, and the Strahler order of the specific river segment. In our sample, the Strahler spans from 3 to 10.

¹¹95 % Confidence interval: [0.000595; 0.042993]

Table 2: Effects of industrial mining opening on child mortality

	12-month mortality				24-month mortality			
	Total Population		Rural Population		Total Population		Rural Population	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Down×Open	-0.00352 [0.00824]	-0.00506 [0.00831]	0.00517 [0.0102]	0.00510 [0.0102]	0.0231** [0.0105]	0.0218** [0.0108]	0.0379*** [0.0130]	0.0380*** [0.0130]
Downstream	-0.0140** [0.00655]	-0.0152** [0.00665]	-0.0203*** [0.00743]	-0.0204*** [0.00762]	-0.0202*** [0.00731]	-0.0211*** [0.00739]	-0.0287*** [0.00795]	-0.0284*** [0.00810]
Open	0.0121* [0.00722]	0.00963 [0.00754]	0.0106 [0.00858]	0.0102 [0.00952]	-0.00302 [0.00986]	-0.00496 [0.0101]	-0.00650 [0.0115]	-0.00588 [0.0122]
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Nb open mines	No	Yes	No	Yes	No	Yes	No	Yes
Birthmth FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Ctry-bthyr FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
MineSB-bthyr trd	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
N	48,472	48,472	33,231	33,231	35,638	35,638	24,544	24544
R2	0.0378	0.0378	0.0476	0.0476	0.0511	0.0511	0.0633	0.0633
Outcome Mean	0.0666	0.0666	0.0716	0.0716	0.0873	0.0873	0.0945	0.0945

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The variables Downstream and Opened are dummies that indicate whether an individual lives in a village downstream of at least one mining site and whether the site opened before the year of birth. Each DHS village is paired to only one mining site so that each individual appears only once in the regression. Other variables are control variables. The sample focuses on heavy metal mines. Columns (1-3) give the results for the total population while columns (4-6) display the results for rural villages. Columns (2, 4, 6, 8) control for the number of open mines within 45 km. Control variables are birth order number, mother's age, mother's age square, mother's years of education and urban, number of open mines, and a continuous variable indicating the presence of rivers and their order.

5.2 Other health effects

Table 3 represents the effect of industrial mining on other children's health outcomes than mortality. Columns (1-3) display the results on anthropometric measures of children who were still living at the time of the survey and are measured at the time of the survey. A child is affected by stunting if her height-for-age z-score is below minus 2 standard deviations below the mean on the World Health Organization Child Growth Standards. The same definition applies to underweight (weight-for-age) and wasting (weight-for-height). We find a negative and significant effect of industrial mining on underweight but not on stunting and wasting: living downstream of an open mine decreases wasting by 3.9 pp. This result could potentially be explained by the death of the most vulnerable children and the survival of the heaviest ones. We find no results on other diseases among living children: anemia (measured), diarrhea, cough, or fever (reported within the two weeks preceding the interview). We find no effect either of industrial mining on low weight at birth (below 2.5 kg) nor reported size at birth (reported as small or very small by the mother).

Table 3: Effects of industrial mining opening on other child health outcomes

	Surviving children							All births	
	Stunting	Underweight	Wasting	Anemia	Diarrhea	Cough	Fever	< 2.5kg	Small
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Down×Open	-0.0167 [0.0199]	-0.0389** [0.0169]	-0.00479 [0.0109]	-0.0265 [0.0277]	0.00146 [0.0130]	-0.00824 [0.0176]	-0.00219 [0.0154]	-0.00925 [0.0180]	-0.00337 [0.0120]
Downstream	-0.0126 [0.0172]	-0.00293 [0.0152]	0.00330 [0.00897]	0.0428** [0.0188]	0.00287 [0.0108]	-0.0115 [0.0138]	0.0141 [0.0144]	-0.00951 [0.0163]	0.00673 [0.0107]
Open	-0.00618 [0.0162]	0.0257* [0.0146]	0.00769 [0.0106]	0.00582 [0.0246]	-0.00545 [0.0111]	0.00547 [0.0124]	-0.00571 [0.0127]	0.00571 [0.00954]	0.0191 [0.0159]
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Birthmonth FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Ctry-bthy FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
MineSB-bthy trd	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
N	37,393	37,043	37,903	19,331	55,162	54,958	54,955	29,162	58,338
R2	0.155	0.124	0.0895	0.215	0.0900	0.104	0.111	0.0608	0.0532
Outcome mean	0.308	0.246	0.0893	0.660	0.165	0.237	0.246	0.171	0.157

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Columns (1-3) focus only on surviving children (due to variable construction in DHS), while the others encompass all children, including those who died before the survey. The same sample and controls as Table 2 Column 2 apply.

5.3 Women's outcomes

We make sure that the results found on child mortality are not due to a change of fertility among women¹². We find no significant effect of industrial mining neither on whether women ever had a child (Table 4 column 1) nor on the total number of children she had (column 2). We find no effect either on whether women were pregnant during the time of the survey (column 3). Table 4 also displays results on women's other health outcomes: we find no effect of industrial mining on neither whether women ever had a miscarriage or on their anemia status.

Table 4: Effects of industrial mining opening on women outcomes

Outcome	Fertility			Health	
	Ever had a child (1)	Total lifetime fertility (2)	Currently pregnant (3)	Ever had a miscarriage (4)	Anemia (5)
Down × Open	0.0156 [0.00977]	-0.0164 [0.0720]	-0.0171 [0.0111]	-0.00507 [0.0138]	0.000806 [0.0261]
Downstream	0.00886 [0.00952]	0.0841 [0.0731]	0.0160 [0.0118]	0.00894 [0.0134]	-0.00928 [0.0233]
Open	-0.00161 [0.00916]	0.0663 [0.0595]	0.00607 [0.00833]	-0.00136 [0.0119]	-0.00417 [0.0285]
Controls	Yes	Yes	Yes	Yes	Yes
Ctry-survey year FE	Yes	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes	Yes
MineSB-svey year trd	Yes	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes	Yes
N	82,406	82,406	82,373	72,423	31,587
R2	0.510	0.659	0.0422	0.0906	0.122
Outcome mean	0.737	2.912	0.0939	0.136	0.396

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Control variables are birth order number, woman's age, woman's age square, woman's years of education, urban, number of open mines and presence of rivers.

¹²The analysis has been made using DHS Women Recode, which population sample is all women aged 15-49 years old.

6 Mechanisms

6.1 Households' access to water and facilities

We deepen our analysis by studying whether the effects found on children's mortality are indeed due to water pollution downstream of mines and not driven by improved access to water and sanitation or facilities upstream. Under 24-month mortality is still increased significantly by 2 p.p. when adding the triple interaction with several facilities variables: whether a household has piped water as the main drinking source (Table 5 column 1), whether a household has a flushed toilet (column 2), whether it has access to electricity (column 3), and whether the mother had visited health facilities during the 12 months preceding the survey (column 4). We find no significant heterogeneity across the four facilities, which suggests that our result is not explained by an improvement of facilities upstream, which contradicts the findings of Dietler et al. [2021]. Table 28 in Section D.1 of the Appendix looks at the DiD estimator using the access to piped water and electricity as dependent variables and shows no difference after the opening of a mine between upstream and downstream villages.

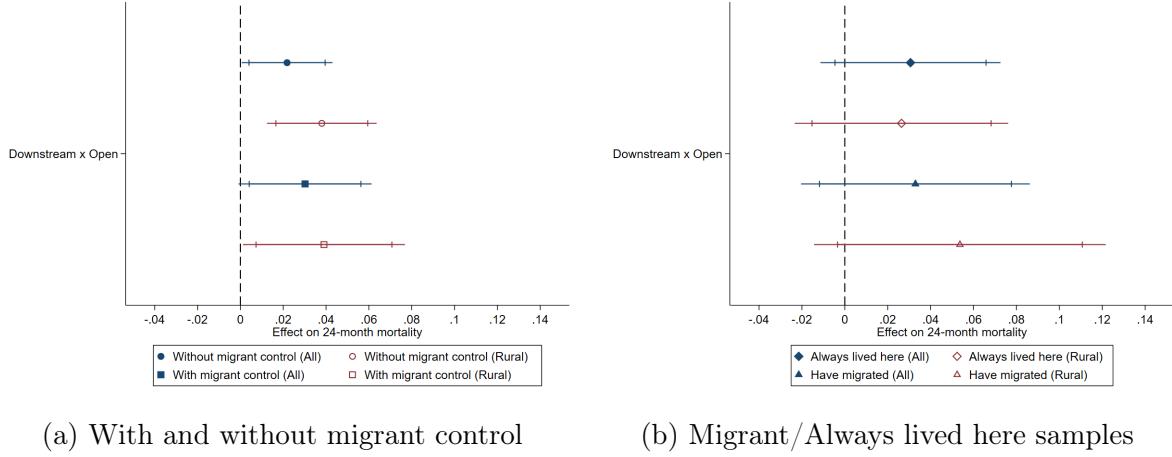
Table 5: Effects of industrial mining opening on access to water, sanitation and facilities

Var.	Outcome	24-month mortality			
		Has piped water		Has flushed toilet	
		(1)	(2)	(3)	(4)
Downstream \times Open \times Var.	0.000283 [0.0195]	-0.0356 [0.0263]	-0.0114 [0.0192]	-0.0140 [0.0165]	
Downstream \times Open	0.0210* [0.0118]	0.0236** [0.0110]	0.0243** [0.0115]	0.0292* [0.0159]	
Var.	-0.00266 [0.00637]	-0.0165* [0.00920]	-0.0157** [0.00695]	-0.00491 [0.00541]	
Downstream	-0.0229*** [0.00770]	-0.0215*** [0.00744]	-0.0223*** [0.00764]	-0.0170* [0.00983]	
Open	-0.000958 [0.0103]	-0.00482 [0.0101]	-0.00715 [0.0103]	-0.00543 [0.0122]	
Controls	Yes	Yes	Yes	Yes	
Birthmonth FE	Yes	Yes	Yes	Yes	
Country-birthyear FE	Yes	Yes	Yes	Yes	
Mine SB FE	Yes	Yes	Yes	Yes	
Mine SB-birthyear trend	Yes	Yes	Yes	Yes	
Commodity FE	Yes	Yes	Yes	Yes	
N	35,638	35,536	35,423	32,018	
R2	0.0512	0.0513	0.0512	0.0512	
Outcome mean	0.0873	0.0873	0.0873	0.0857	

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The same sample and controls as Table 2 Column 2 apply.

6.2 Migration

Figure 7: Migration analysis



Notes: Figure (a) plots the coefficients associated with being downstream of an open mine on under 24-month mortality when controlling for mothers' migration status or not, across the whole rural sample. Figure (b) plots the coefficients of the same interaction but across the sample of mothers who have migrated or always lived here. *Sources:* Authors' elaboration on DHS and SNL data.

We pursue our analysis by making sure that our results on child mortality are not due to migration and do not suffer from selection bias. The migration information is retrieved from the variable indicating whether mothers have ever migrated to the actual place of residency, or if they have always lived there. The information is available among 60% of our sample (cf. Table 24) and controls for in-migration, which is an important effect of the opening of a mine that attracts new working populations (cf Section D.1 for more discussion on bias linked to migration). Figure 7 displays the coefficient associated with the interaction term of being downstream of an open mine. The top coefficient in Figure 7a is our main specification when we do not control for migration, across the whole. We plot the same focusing on the rural sample. The bottom two coefficients are when we control for migration, across our whole and rural sample. We find that all are statistically positive and significant. We further our analysis by splitting the sample across mothers who have ever migrated and mothers who have always lived here (Figure 7b). The estimation suffers from a lack of statistical power (see Appendix Table 6 for the drop of observations) but suggests no differential effect of industrial mining across the two samples.

Table 6: Effects of industrial mining activity, migration analysis

Outcome	Mortality under 24 months							
	Without migrant control		With migrant control		Migrant sample		Always lived here	
Spec.	All	Rural	All	Rural	All	Rural	All	Rural
Sample	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Downstream x Open	0.0218** [0.0108]	0.0380*** [0.0130]	0.0302* [0.0158]	0.0390** [0.0193]	0.0329 [0.0272]	0.0537 [0.0347]	0.0306 [0.0214]	0.0264 [0.0254]
Downstream	-0.0211*** [0.00739]	-0.0284*** [0.00810]	-0.0249** [0.0106]	-0.0332*** [0.0119]	-0.0229 [0.0190]	-0.0304 [0.0226]	-0.0301** [0.0139]	-0.0418*** [0.0147]
Open	-0.00496 [0.0101]	-0.00588 [0.0122]	-0.0255 [0.0159]	-0.0235 [0.0194]	0.0116 [0.0260]	0.0187 [0.0335]	-0.0409** [0.0198]	-0.0425* [0.0256]
Migrant			0.00850* [0.00449]	0.00348 [0.00594]				
N	35638	24544	22231	15060	8658	6007	13503	8982
R2	0.0511	0.0633	0.0634	0.0770	0.112	0.132	0.0797	0.107
Outcome mean	0.0873	0.0945	0.0946	0.104	0.0892	0.102	0.0983	0.107

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The same sample and controls as Table 2 Column 2 apply.

6.3 Early life characteristics

We want to better understand the mechanism behind the results on 12- and 24-month mortality, and focus on children's nutrition and access to health care as the main potential explaining factors. Table 7 gives the results of several triple interactions looking at the heterogeneity of the effect according to early life characteristics.

We find a significant increase in mortality among children living downstream of an open mine and who were given plain water 24 hours before the survey for both the 12-month mortality and 24-month mortality. For individuals living downstream of a mine that opened and who consumed plain water, the 12-month mortality increases by 5.3 p.p. and the 24-month mortality increases by 6.5 p.p. (Columns (1) and (2)). Unfortunately, the DHS variable does not specify the source of plain water, and we cannot show that the given plain water is more polluted downstream than upstream. We find no significant effect of the triple interaction with breastfeeding behaviors on mortality: whether the child was ever breastfed (columns 3 and 4) or number of months during which the child was breastfed (columns 5 and 6). This absence of result can be explained by the low variability as 98 % of the children in the sample were breastfed. There is a larger variability of plain water consumption, as it was the case for 18.7% of children. We thus interpret the fact that drinking plain water as a child is a proxy for having non-exclusive breastfeeding¹³.

We find no significant effect in either of the triple interaction with access to health care: whether the mother received prenatal care (columns 7 and 8) or whether the child was ever vaccinated (columns 9 and 10).

¹³DHS questionnaire asks whether the children have been breastfed and have consumed plain water during the 24 hours before the survey

Table 7: Effects of industrial mining opening on explaining factors 12 vs. 24 months

Var.	Child's nutrition						Child's access to health care			
	Was given plain water		Ever breastfed		Breastfeed months		No prenatal care		Ever vaccinated	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
Mortality outcome	12m	24m	12m	24m	12m	24m	12m	24m	12m	24m
Downstream × Open × Var.	0.0530** [0.0244]	0.0650* [0.0350]	0.000597 [0.0519]	-0.0450 [0.0481]	0.00106 [0.00176]	0.00106 [0.00210]	0.0165 [0.0272]	0.0288 [0.0402]	0.0114 [0.0191]	0.0134 [0.0231]
Downstream	-0.0000982 [0.00760]	-0.00344 [0.0138]	-0.0104 [0.0250]	-0.0161 [0.0305]	-0.0125 [0.0237]	-0.0293 [0.0300]	-0.0140** [0.00613]	-0.0200** [0.00871]	0.0117 [0.0127]	0.00258 [0.0126]
Open	0.0104 [0.00861]	-0.00951 [0.0189]	0.0172 [0.0266]	0.0163 [0.0345]	-0.0785*** [0.0231]	-0.134*** [0.0324]	0.00842 [0.00770]	0.000268 [0.0117]	0.00433 [0.0105]	0.00929 [0.0175]
Downstream × Open	-0.0131 [0.0103]	0.0193 [0.0214]	0.00423 [0.0523]	0.0702 [0.0488]	-0.0213 [0.0373]	-0.000897 [0.0516]	-0.0114 [0.00823]	-0.000663 [0.0121]	-0.0119 [0.0178]	-0.0105 [0.0211]
Var.	0.0338*** [0.0102]	0.0277** [0.0137]	-0.921*** [0.0115]	-0.880*** [0.0134]	-0.0179*** [0.000595]	-0.0201*** [0.000657]	0.0299*** [0.00774]	0.0398*** [0.0110]	-0.0201*** [0.00728]	-0.0257*** [0.00828]
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Birthmonth FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Country-birthyear FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB-birthyear trend	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
N	19,671	10,797	45,168	33,022	29,015	18,323	31,656	19,543	17,372	13,638
R2	0.0735	0.102	0.208	0.174	0.330	0.355	0.0558	0.0822	0.239	0.306
Outcome mean	0.0396	0.0758	0.0493	0.0694	0.0466	0.0768	0.0479	0.0639	0.00835	0.0121

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The same sample and controls as Table 2 Column 2 apply.

7 Heterogeneity

7.1 Individual characteristics

Table 8: Effects of industrial mining opening across children's location and gender

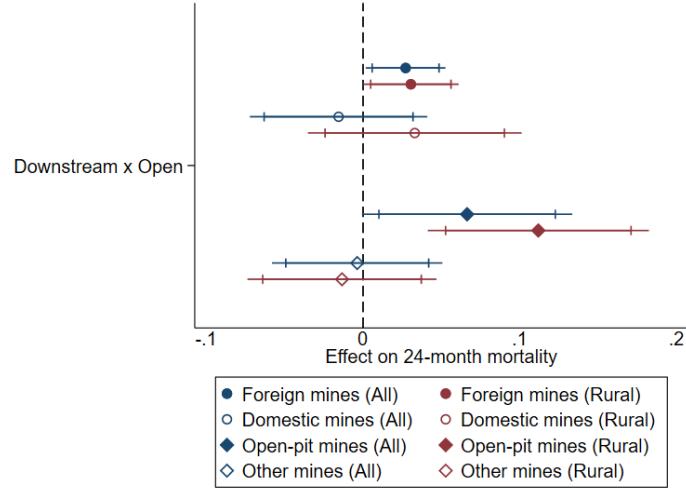
Sample	24-month mortality					
	All: urban and rural			Rural		
	All	Girls	Boys	All	Girls	Boys
	(1)	(2)	(3)	(4)	(5)	(6)
Downstream \times Open	0.0218** [0.0108]	0.0120 [0.0151]	0.0334** [0.0167]	0.0380*** [0.0130]	0.0204 [0.0186]	0.0677*** [0.0199]
Downstream	-0.0211*** [0.00739]	-0.0203* [0.0114]	-0.0206* [0.0111]	-0.0284*** [0.00810]	-0.0292** [0.0126]	-0.0292** [0.0117]
Open	-0.00496 [0.0101]	0.00364 [0.0155]	-0.0178 [0.0143]	-0.00588 [0.0122]	0.00419 [0.0186]	-0.0200 [0.0181]
Controls	Yes	Yes	Yes	Yes	Yes	Yes
Birthmonth FE	Yes	Yes	Yes	Yes	Yes	Yes
Country-birthyear FE	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB-birthyear trend	Yes	Yes	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes	Yes	Yes
N	35,638	17,452	18,142	24,544	12,009	12,481
R2	0.0511	0.0758	0.0762	0.0633	0.0942	0.0972
Outcome mean	0.0873	0.0805	0.0938	0.0945	0.0883	0.101

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The same sample and controls as Table 2 Column 2 apply.

We conduct a heterogeneity analysis across children's location and gender (Table 8). We find that being downstream of an open mine is more critical in rural areas, (Column 4) as it increases the 24-month mortality by 3.8 p.p, which corresponds to a 40 % increase in the mortality rates. The heterogeneity by gender shows that our results are mainly driven by the mortality of males (columns 3 and 6), which remains consistent in rural areas.

7.2 Mining activity's characteristics

Figure 8: Heterogeneity across mines' characteristics



Notes: This graph represents the coefficients associated with the interaction of living downstream of an open mine when splitting the sample between mines that are owned by foreign companies and mines that are owned by at least one domestic company (in blue) and between mines that are open-pit and not open-pit (underground, placer, and in-situ leach) (in red).

Sources: Authors' elaboration.

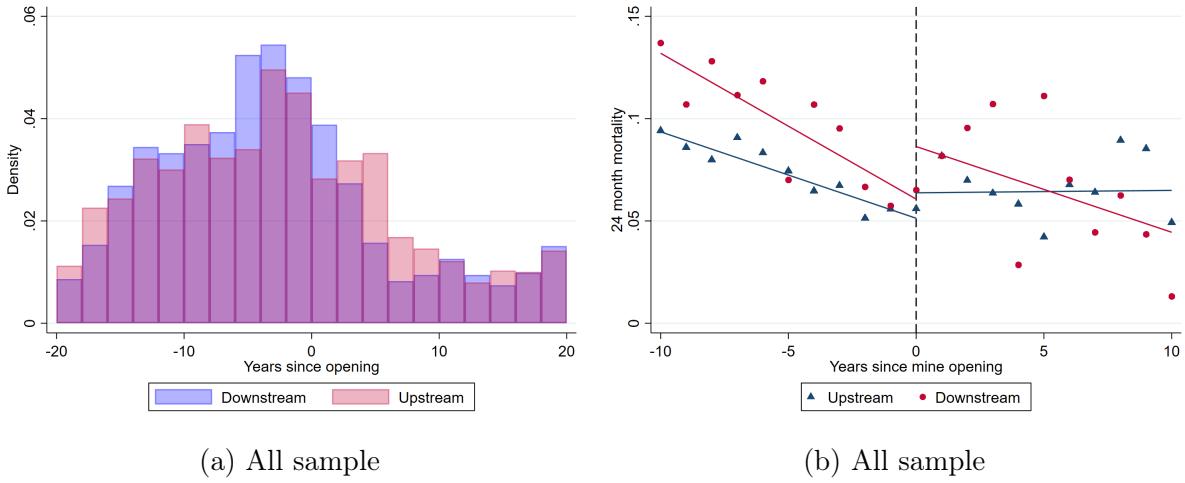
We pursue the heterogeneity analysis across mines' ownership and extraction methods. A mine was considered domestic if at least one of the owning companies is from the same country as the country of location, and they represent 17.8 percent of our mine sample. Figure 8 represents the coefficients associated with the interaction term of living downstream of an open mine. We find no effect of mine opening when we restrict the sample to domestically owned mines whereas our results hold when we restrict to the foreign-owned only mines (in blue). This could potentially be explained by improved management of a mine or a better consideration of the surrounding populations if a national company is involved. We then look at the open-pit nature of the industrial site, which concerns 21.6 percent of our mine sample. We find that our results hold when restricting to open-pit mines but not when we only look at the sample of other extraction methods (underground, placer, and in-situ leach) (in red). This is consistent with the fact that open-pit mines are the most polluting mines due to the generation of large amounts of waste kept in tailing storage facilities.

8 Dynamic effects

In this section, we investigate the dynamic effects of the opening of an industrial mine, looking at pre-trends and at whether the effects on 24-month mortality occur within a short or long time, and during the mining activity.

8.1 Pre-trends and event-study

Figure 9: Linear trends of 24-month mortality



Notes: Figure (a) gives the distribution of the number of observations per opening year. Figure (b) plots the trends of the 24-month mortality rates according to the year of opening. The figures are made for the whole sample and include neither control variables nor fixed effects. The reference point is -1, the year before the mine opening.

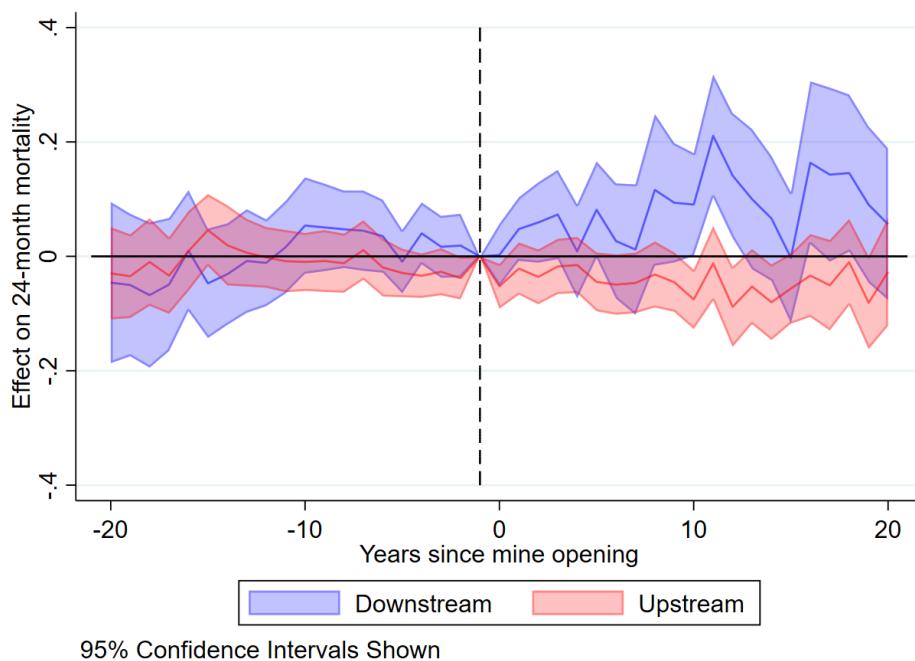
Sources: Authors' elaboration on DHS and SNL data.

The key assumption of the DiD strategy is that the outcome - the 24-month mortality - would follow the same time trend in the absence of the mine opening both in upstream and downstream areas. The common trends assumption cannot be tested. However, we can observe the pre-treatment data and the evolution of mortality rates before each mine opening according to the topographic position. Figure 9b plots the linear trends of the 24-month mortality rates and distinguishes between upstream and downstream DHS clusters, before and after the opening of the paired mine. For each year, it plots the average mortality rates over the sample, with no control nor fixed effect. Figure 9a plots the distribution of the years before and after the mine opening. Figure 9b shows non-exact parallel trends but is only descriptive, and we can see looking at the scatter plot that the downstream and upstream areas seem to follow a similar pattern of decreasing mortality before a mine opens. This decreasing pattern is triggered by temporal trends, as the years closest

to the mine opening are more likely to be recent years, and the mortality rates are overall decreasing over the recent decades (cf Figure 5). This pattern is corrected in Figure 10.

Figure 10 plots the event study of the effect of mine opening, for both the upstream and downstream samples. It includes the same controls and fixed effects as the main analysis (cf Table 2), and thus corrects for previous trends as we include country and mine sub-basin trends. Both upstream and downstream villages do not display any pre-trends, which suggests that the common trend assumption is verified. We observe almost no effect of a mine opening on the mortality rates upstream, a slight decrease which is significant 10 years after the opening. Figure 10 shows that the infant mortality downstream increases once the mine opens. This effect is significant in the medium run, around a decade after the mine opening.

Figure 10: Event study - dynamic effect of mine opening on 24-month mortality



Notes: This Figure plots the event study of the effect of mine opening for downstream and upstream DHS villages, 10 years before the mine opening and 10 years after. The year before the mine opening, -1, is taken as the reference point.

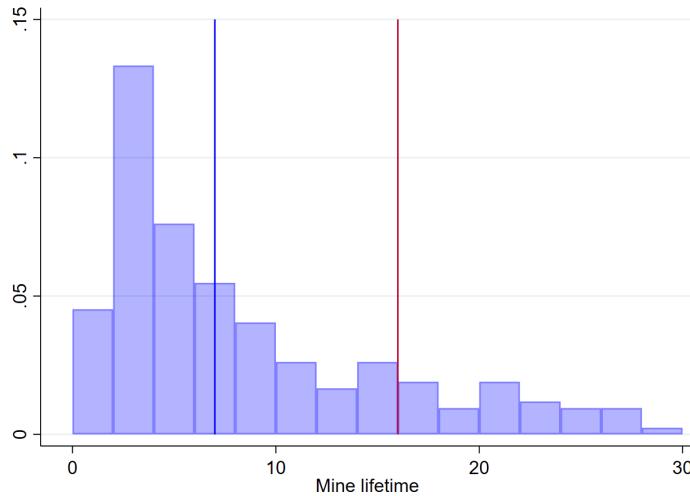
Sources: Authors' elaboration on DHS and SNL data.

8.2 Mine closure

In the main analysis, we focus on mine opening without taking into account the closure year, since it is a piece of information harder to retrieve by hand. In this section, we focus on a restricted sample of mines for which the SNL database provides directly this information, to understand whether the closing date plays a role in our main effect.

Figure 11 gives the distribution of mines' lifetime, for the restricted sample of mines for which closing years are available. On average, a mine lasts 16 years, but the distribution is skewed to the right, and the majority of mines close before 10 years. Please note that the closing date available in the SNL database for this restricted sample is not exact. Over its lifetime, a mine can be put on hold several times, for political or economic reasons. In Table 9 we look at the effects of being downstream of a mine that is active at the year of birth of the child. Columns (1) and (2) show no effect on the 12-month mortality rates. Column (3) shows that being downstream of a mine that is active the year of birth increases the mortality rate by 4 p.p, which corresponds to an increase of 40% in the mortality rate. This result suggests that the harmful effects of mining activity on the individuals living downstream are mainly critical while the mine is active.

Figure 11: Distribution of a mine lifetime



Notes: This figure gives the distribution of mines' lifetime. The red line $y=16$ plots the mean of a mine lifetime, while the blue line $y=7$ plots the median. The maximum of a mine lifetime in our sample is 138 years.

Sources: Authors' elaboration on SNL data.

Table 9: Average effects of mine activity on infantile mortality

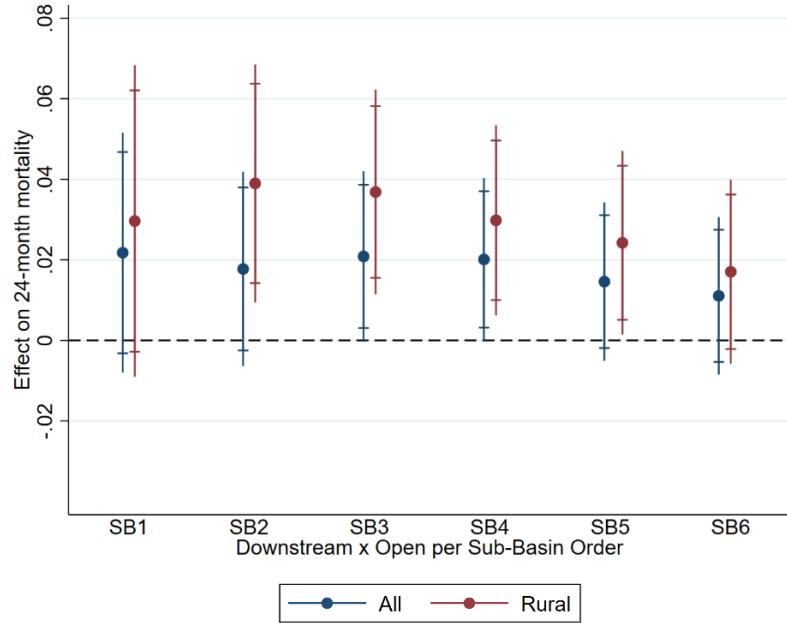
	Mortality under 12 months		Mortality under 24 months	
	All	Rural	All	Rural
	(1)	(2)	(3)	(4)
Downstream×Active	0.0112 [0.0222]	0.0321 [0.0254]	0.0409* [0.0248]	0.0488** [0.0242]
Downstream	-0.0264* [0.0139]	-0.0434** [0.0184]	0.00877 [0.0145]	-0.00721 [0.0184]
Active	0.000600 [0.0140]	0.00420 [0.0173]	-0.00842 [0.0172]	-0.00345 [0.0193]
Controls	Yes	Yes	Yes	Yes
Nb open mines	Yes	Yes	Yes	Yes
Birthmonth FE	Yes	Yes	Yes	Yes
Country-birthyear FE	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes
Mine SB-birthyear trend	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes
N	7,231	5,589	5,270	4,082
R2	0.0899	0.0960	0.0984	0.108
Outcome Mean	0.0756	0.0825	0.0981	0.104

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The variables Downstream and Active are dummies that indicate whether the individual lives in a village downstream of at least one mining site and whether the site is active during the year of birth. The same sample and controls as Table 2 Column 2 apply.

9 Intensive Margin

9.1 Spatial intensive margin

Figure 12: Effect of industrial mine opening according to the downstream sub-basin order

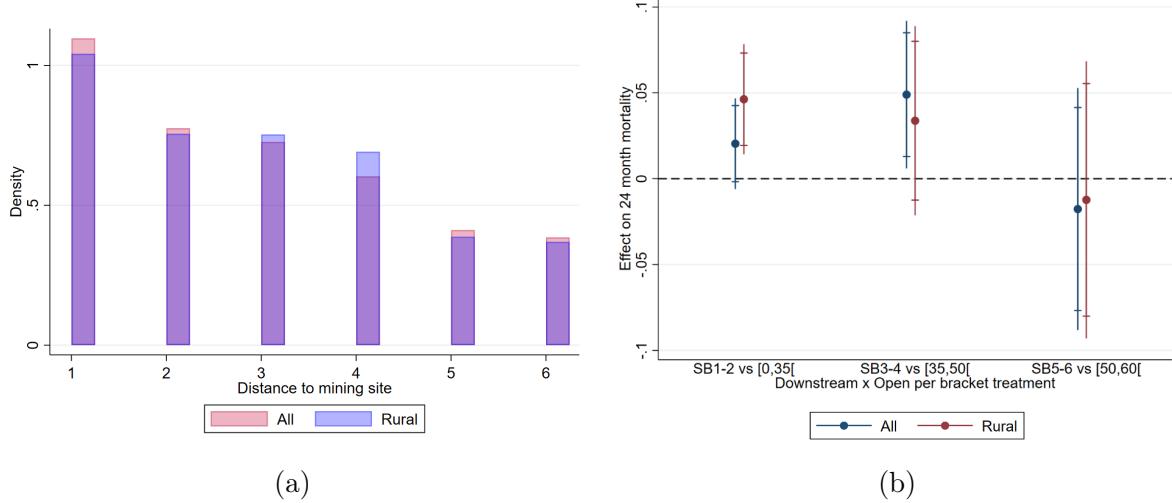


Notes: This Figure plots the treatment effect when changing the treatment group. SB 1 gives the DiD estimator when the control group includes DHS villages up to the first sub-basin, SB 2 up to the second, and then SB6 up to the sixth sub-basin. SB 3 gives the main result of the paper.

Sources: Authors' elaboration on DHS and SNL data.

In this section, we change the cut-off for being treated and test the effect on different orders of downstream sub-basins. In Figure 12, we test whether the effect holds when allowing for further sub-basins downstream. It plots the coefficient on *Downstream* \times *Open* for six different models. SB1 corresponds to the model where the treatment group includes only the first neighboring downstream sub-basin, SB2 up to the second, and SB6 up to the sixth. SB3 gives the main results from Column (6) in Table 2. The Figure shows an attenuation of the magnitude of the effect when including further sub-basins. For all the individuals, the results are significant at the 5% level up to the third sub-basin and up to the fourth, while for all rural areas, it is significant from the second up to the fifth sub-basin. We interpret the non-significance of the result in the first sub-basins as being the consequence of statistical power (as the sample size is relatively low up to SB1 and SB2).

Figure 13: Intensive margin - Effect of the number of mine opening on under 24 months mortality



Notes: Figure (a) plots the distribution of observations downstream per sub-basin order. Figure (b) plots the interaction term on the 24-month mortality rates per distance brackets. The first coefficient on the left gives the effect for individuals living within the first and second sub-basins compared to individuals living up to 35 kilometers. The second coefficient gives the effect for individuals living downstream within the third and fourth sub-basins compared to those living upstream between 35 and 50 kilometers. The third coefficient gives the effect for individuals living downstream within the fifth and sixth sub-basins compared to those living upstream between 50 and 60 kilometers. The distances are chosen based on the mean of the distance between the mine and the extremity of the XXth sub-basin.^a

Sources: Authors' elaboration on DHS and SNL data.

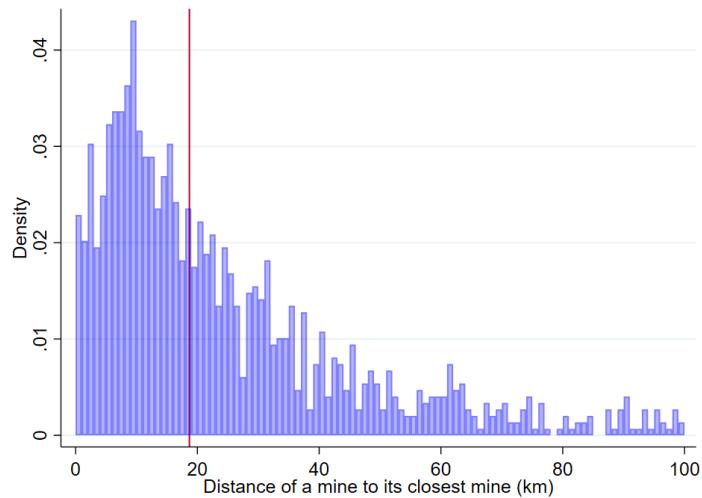
^aThe mean of the distance between the mine and the extremity of the first sub-basin is 27 km. It is 37 km for SB2, 45km for SB3, 46km for SB4, 58km for SB5 and 59km for SB6

Figure 13 looks at the effect per distance brackets. To build the control group for each sub-basin, we determined which distances correspond to which sub-basin order. We calculated the mean of the maximal distance between the mine and the furthest extremity of each sub-basin order. On average, a mine is at 27 kilometers of the furthest extremity of its first sub-basin, at 37 kilometers of its second sub-basin, 45 kilometers of sub-basin 3, 46 kilometers of sub-basin 4, 58 kilometers of sub-basin 5 and 59 kilometers of sub-basin 6. Following this indicator, we compare in Figure 13 the individuals living downstream within the first and second sub-basin to those living upstream within 35 kilometers of the mine (coefficient SB1-2). Then, we compare individuals living downstream within the third and fourth sub-basins to those living upstream within 35 to 50 kilometers of the mine (coefficient SB3-4). Finally, we compare individuals living downstream within the

fifth and sixth sub-basin to those living within 50 to 60 kilometers of the mine. Figure 13 shows that our result is mainly driven by the effect within the third and fourth sub-basins, while in rural areas the effect is only significant within the closest sub-basins. This shows that the effect is critical close to the mine, where the pollution is supposed to be the highest. The difference between the whole and rural samples might be explained by the fact that the location of mines close to urban areas suffers from lower precision (cf section 10.3.2).

9.2 Mine density

Figure 14: Distance of mines to each other



Notes: This Figure gives the distribution of the distance between each mine and its closest industrial mining site, giving insights into how far these mines are located from each other. In red is plotted the median distance (18km), and the graph is given for distances under 100 km (please note that the maximum distance is up to 474 km).

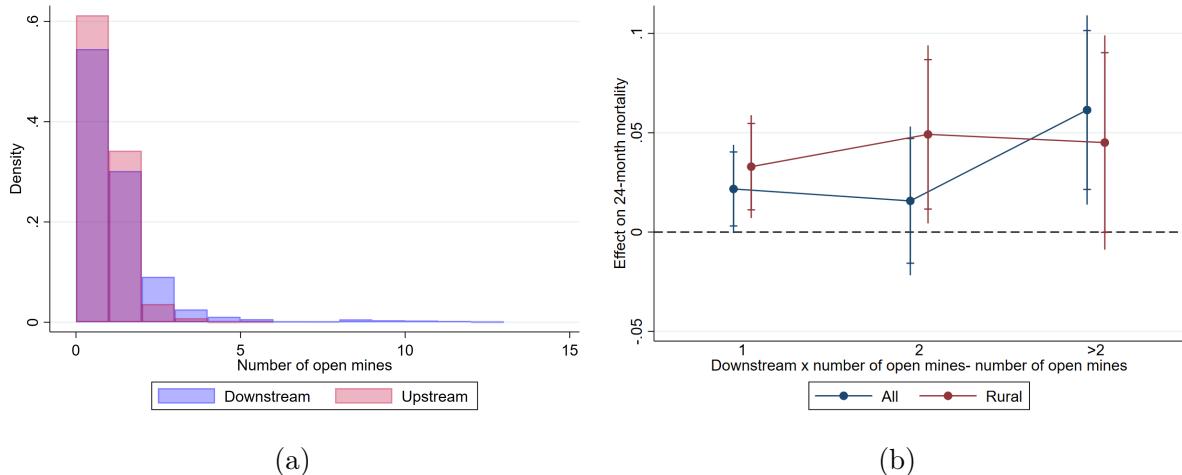
Sources: Authors' elaboration on DHS and SNL data.

In this section, we explore the intensive margin of our result according to the mine density. Figure 14 gives insights into how far these mines are located from each other. It shows the distribution graph of the distance to the closest mine for each mine. It is made with no regard for the activity status of the mining site. On average, a mine is located at least 31 kilometers from its closest mining site, while the median is 18 kilometers. The distribution is skewed to the right, showing that the majority of the mining sites are located in areas with a high density of mining activity. Few mines are isolated, up to 100 kilometers from the closest mining site. This graph shows the necessity first to control for the number of open mines within the area in the main analysis (Table 2), and the

necessity to look at heterogeneous effects according to the intensity of the mining activity within the area.

Figure 15a plots the frequency of the number of open mines within our main sample, both for upstream and downstream areas, within 45 kilometers and up to the third sub-basin. As the number of observations falls starting at 3 open mines, we investigate in Figure 15b and Table 10 the statistical difference between being downstream of one opening site, two or more than two. For the whole sample, being downstream of one open mine increases the 24-month mortality by 2 p.p. The effect increases when the number of open mines increases, as being downstream of more than 2 open mines increases by 6 p.p the mortality rate in comparison to being downstream of only one open mine. Table 10 column (1) gives the DiD interaction term when open becomes a continuous variable and not a dummy, being the number of open mines. It shows that being downstream one additional mine that opens increases the mortality by 1.3 p.p, and by 2 p.p within rural areas.

Figure 15: Intensive margin - Effect of the number of mine openings on 24-month mortality



Notes: Figure (a) plots the distribution of the number of open mines across downstream and upstream villages. Figure (b) plots the interaction variable on the 24-month mortality rates. It gives the average treatment effects of the number of mines open on 24-month mortality.

Sources: Authors' elaboration on DHS and SNL data.

Table 10: Effects of the number mine opening on infantile mortality according to the number of open mine

	24-month mortality			
	All	Rural		
	(1)	(2)	(3)	(4)
Downstream×Nb open	0.0130*** [0.00484]		0.0205*** [0.00744]	
Downstream×Nb open=1		0.0217* [0.0113]		0.0329** [0.0132]
Downstream×Nb open=2		0.0157 [0.0191]		0.0492** [0.0228]
Downstream×Nb open >2		0.0614** [0.0243]		0.0451 [0.0275]
Downstream	-0.0214*** [0.00723]	-0.0208*** [0.00767]	-0.0295*** [0.00823]	-0.0308*** [0.00837]
Nb Open	-0.00419 [0.00526]		-0.00842 [0.00613]	
Nb Open =1		-0.0000134 [0.00785]		-0.00328 [0.00895]
Nb Open =2		-0.00850 [0.0129]		-0.0220 [0.0141]
Nb Open >2		-0.0407** [0.0190]		-0.0329 [0.0219]
Controls	Yes	Yes	Yes	Yes
Birthmonth FE	Yes	Yes	Yes	Yes
Country-birthyear FE	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes
Mine SB-birthyear trend	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes
N	35,638	35,638	24,544	24,544
R2	0.0511	0.0511	0.0633	0.0634

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Columns (1-3) give the results for the total population while columns (4-6) display the results for rural villages. Columns (2, 4, 6, 8) control for the number of open mines within 45 km. The same sample as Table 2 Column 2 applies. Control variables are birth order number, mother's age, mothers' age square, mother's years of education and urban, and a continuous variable indicating the presence of rivers and their order.

9.3 Production intensive margin

Table 11: Effects of industrial mining opening, across each commodity’s price evolution.

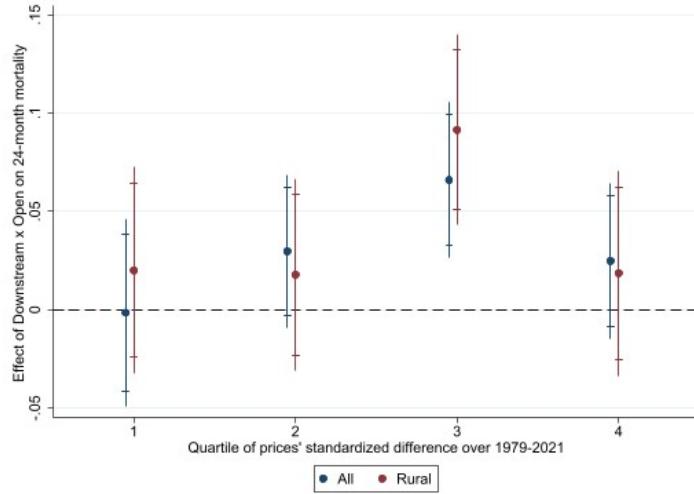
Price var.	Outcome	24-month mortality	
		(1) Standardized difference	(2) Z-score
Downstream × Open × Price var		0.0160* [0.00823]	0.0101** [0.00463]
Downstream		-0.00244 [0.0104]	0.00102 [0.0110]
Controls	Yes	Yes	
Country-survey year FE	Yes	Yes	
Mine SB FE	Yes	Yes	
Mine SB-survey year trend	Yes	Yes	
Commodity FE	Yes	Yes	
N	31,517	31,517	
R ²	0.0509	0.0509	
Outcome mean	0.0907	0.0907	

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Standardized difference and Z-score of each commodity’s price calculated over 1979-2021. The same sample and controls as Table 2 Column 2 apply.

We proxy the production intensity of each mine by using the global price of each mine’s primary commodity as done in Berman et al. [2017] and Girard et al. [2022]. We presume the higher the prices, the more intense the production. Annual prices were retrieved from the SNL data (coal, gold, lead, nickel, platinum, silver, and zinc) and the World bank pink sheet (copper). For each commodity, we calculate the average price over 1971-2021, and for each year we calculate the standardized difference ($\frac{Price_t - \overline{Price}_{[1971-2021]}}{\overline{Price}_{[1971-2021]}}$) and Z-score ($\frac{Price_t - \overline{Price}_{[1971-2021]}}{\sigma_{Price}_{[1971-2021]}}$). We find a positive and significant effect of the triple interaction (Table 11) with both price variables (column 1 for the standardized difference and column 2 for the Z-score).

We then plot the coefficients associated with the interaction term of being downstream of an open mine, across the quartiles of the change in the z-score (Figure 16). For both total and rural samples, we find that an increase in prices (going from the second to the third quartile) leads to an even higher effect of industrial mining on the 24-month mortality.

Figure 16: Effect of living downstream of an open mine across the evolution of the mine's primary commodity's price



Notes: This graph represents the coefficients associated with being downstream of an open mine across the evolution of each mine's primary commodity's price (available for coal, copper, gold, lead, nickel, platinum, silver and zinc). For each commodity, the standardized difference to the mean over the 1979-2021 period was calculated and then split across quartiles to grasp the relative price evolution specific to each type of commodity.

Sources: Authors' elaboration using DHS, SNL, and World Bank pink sheet data.

10 Robustness checks

10.1 Balanced sample and de Chaisemartin and d'Haultfœuille [2020]

10.1.1 Balanced sample

One issue of working with DHS data is dealing with repetitive cross-sections instead of an exact panel. In this section, we define a balanced sample as a restricted sample for which each mine has observations before and after its opening, both upstream and downstream. In this sense, it is a balanced panel of mine, if we consider only two points in time which are (1) the period before the mine opens and (2) the period after its opening. Please note that, in this paper, it is possible to restrict the analysis to a balanced sample by the extension of the mines' sample size, and it underlines the limits of analyses looking at dynamic effects while using a few mines. In this section, we first define the balanced sample and then we replicate the main analysis on this restricted sample.

First, let's define the construction of the balanced sample. A staggered DiD is driven

by changes in mortality rates of switchers, which are observations that change treatment status, in comparison to those that do not change treatment status. In the case of a balanced sample, the design of this paper distinguishes three groups of observations :

- Group 1 "the switchers": the subgroup for which a mine has opened between two different years (for which there are DHS observations) and thus for which the treatment status changes from 0 to 1.
- Group 2 "the always treated": the subgroup of areas for which the mine has always been opened and are thus always treated, i.e. the treatment variable which is an interaction is always equal to 1.
- Group 3 "the never treated": the subgroup of areas where mines have not yet opened. The treatment variable is equal to 0. The third group is made of subgroups for which the mine has not opened yet in 2022 but the opening is planned in the future (the mine is projected to open), but also of mines where no DHS cluster was surveyed after it opened. This group is called "the never treated", but it includes both DHS villages that will never be treated or are not yet treated because they will be treated in the future.

The balanced sample makes it possible to identify the three groups. Formally, it is defined as the following, for each group:

Let's consider observations that can be divided into G groups and T periods, for every $(g, t) \in \{1, \dots, G\} \times \{1, \dots, T\}$, let $N_{g,t}$ denote the number of observations in the group g and period t , and let $N = \sum_{g,t} N_{g,t}$ be the total number of observations. For all $(g, t) \in \{1, \dots, G\} \times \{1, \dots, T\}$, let's call $D_{g,t}$ the *Downstream_{g,t}* variable and $O_{g,t}$ the *Opened_{g,t}* variable.

Definition 1 (Group 1- Balanced sample of "switchers"). *Let's call $G_1 = \{g_0, \dots, g_{n_1}\}$ the set of Group 1. Group 1 is defined as the following:*

For all $g \in G_1, \exists (v_1, v_2, v_3, v_4) \times (t_1, t_2, t_3, t_4) \in g \times T$ such as:

- (i) $N_{v_1, t_1} > 0 \wedge D_{v_1, t_1} = 0 \wedge O_{v_1, t_1} = 0$
- (ii) $N_{v_2, t_2} > 0 \wedge D_{v_2, t_2} = 1 \wedge O_{v_2, t_2} = 0$
- (iii) $N_{v_3, t_3} > 0 \wedge D_{v_3, t_3} = 0 \wedge O_{v_3, t_3} = 1$
- (iv) $N_{v_4, t_4} > 0 \wedge D_{v_4, t_4} = 1 \wedge O_{v_4, t_4} = 1$

In our setting, g is the whole area associated with a mine, including both upstream and downstream observations, and is made of $k \in N$ DHS clusters such as $g = \{v_1, \dots, v_k\}$.

Definition 2 (Group 2- Balanced sample of "always treated"). Let's call $G_2 = \{g_0, \dots, g_{n_2}\}$ the set of Group 2. Group 2 is defined as the following:

For all $g \in G_2, \exists (v_1, v_2) \times (t_1, t_2) \in g \times T$ such as:

- (i) $N_{v_1, t_1} > 0 \wedge D_{v_1, t_1} = 0 \wedge O_{v_1, t_1} = 1$
- (ii) $N_{v_2, t_2} > 0 \wedge D_{v_2, t_2} = 1 \wedge O_{v_2, t_2} = 1$

Definition 3 (Group 3- Balanced sample of "never treated"). Let's call $G_3 = \{g_0, \dots, g_{n_3}\}$ the set of Group 3. Group 3 is defined as the following:

For all $g \in G_3, \exists (v_1, v_2) \times (t_1, t_2) \in g \times T$ such as:

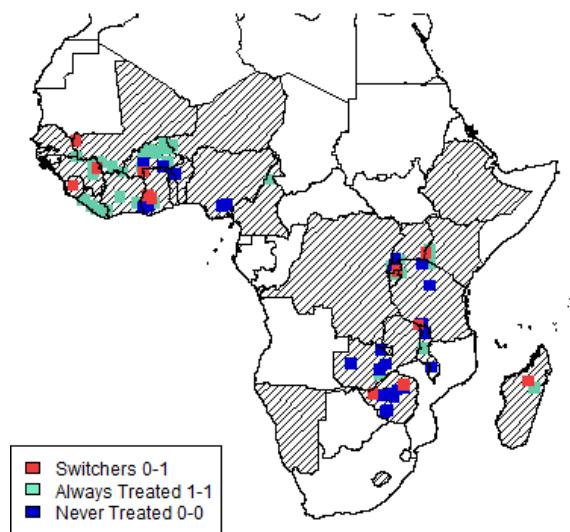
- (i) $N_{v_1, t_1} > 0 \wedge D_{v_1, t_1} = 0 \wedge O_{v_1, t_1} = 0$
- (ii) $N_{v_2, t_2} > 0 \wedge D_{v_2, t_2} = 1 \wedge O_{v_2, t_2} = 0$

Definitions 1, 2, and 3 define the treatment group (Group 1) and the control groups (Group 2+3) in the setting of the mine-balanced sample. For Group 1 switchers, we restrict the sample to areas that have DHS clusters surveyed both downstream and upstream both before and after the opening of the mine. This means that we select areas that have been surveyed at least in four different locations by DHS. For groups 2 and 3, meaning that the surveys have occurred only after the opening (Group 2) or before the opening (Group 3), we restrict to areas that have observations both upstream and downstream.

Figure 17 plots the different groups from the balanced sample and displays the group of switchers, the always-treated and the never treated groups. It displays the areas that are key to the main estimation, which are located in Western Africa, Zimbabwe, Western Kenya, Rwanda, Tanzania, and Madagascar. Table 29 from Appendix plots our main estimator across the African sub-regions and shows that our results are mainly driven by Western Africa, and remain significant in Eastern Africa. Table 12 gives the size of the three groups in the balanced sample, as well as the associated number of mines. It shows that Group 1 switchers account for around 12% of the total balanced sample, and corresponds to the neighborhood of 13 mines. In total, the control groups gather 75 mines. We observe that the average mortality rates have decreased over time, before and after the opening of the mining site in the Switcher Group, linked to the decrease of infant mortality in Africa over time, which is in coherence with the Balance Table 1 and Figures 4 and 5, and which highlights the importance of controlling for trends.

Table 13 displays the balance table for the restricted sample. This table shows the within comparison before and after a mine opens both downstream and upstream. It is only descriptive statistics and neither accounts for control variables nor account for fixed effects. The table shows that there is a significant difference between upstream and downstream areas after a mine opens concerning both the 12 and 24-month mortality rates. From a descriptive point of view, being downstream of a mine increases the 12-month mortality by 2.7 p.p, and the 24-month mortality by 2.4 p.p (column (13)). This difference is explained by a significant decrease in mortality rates within upstream areas after a mine opening (column (11)).

Figure 17: Balanced Panel - Group identification



Notes: The Figure plots the groups' areas across the three groups of the balanced panel, for the 24-month mortality rate.

Sources: Authors' elaboration on DHS and SNL data.

Table 12: Balanced Sample - Descriptive Statistics

	Group 1 : Switchers 0-1						Groups 2+3		Group 2 : 1-1		Group 3 : 0-0							
	All			Before Opening		After Opening		N	Mean (SD)	N	Mean (SD)	N	Mean (SD)					
	N	Mean (SD)	N	Mean (SD)	N	Mean (SD)	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
Dth<12																		
All	1,191	0.07	894	0.069	297	0.071	8,423	0.07	2,368	0.056	6,055	0.076						
Mines	13		13		13		75		31		44							
Dth<24																		
All	1,191	0.089	894	0.091	297	0.084	8,423	0.091	2,368	0.072	6,055	0.099						
Mines	13		13		13		75		31		44							

Notes: Standard errors and p-values are in parentheses. Outcomes' descriptive statistics of under 12-and 24-month mortality, for villages within the Group 1 Switchers for individuals born before and after the opening of the mine, then Group 2 always treated and Group 3 never treated.

Table 13: Balance Table

Before Mine Opening					After Mine Opening					Within Up.	Within Dwn.	Within	
	Upstream	Downstream	Diff		Upstream	Downstream.	Diff						
	N	Mean / (SD)	N	Mean / (SD)	(4-2) / (p.v)	N	Mean / (SD)	N	Mean / (SD)	(9-7) / (p.v)	(7-2) / (p.v)	(9-4) / (p.v)	(12-11) / (p.v)
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)
Dth<12													
All	5,272	0.079	1677	0.063	-0.016	1,812	0.054	853	0.066	0.012	-0.025	0.002	0.027
		(0.27)		(0.243)	(0.025)		(0.226)		(0.248)	(0.248)	(0)	(0.814)	(0.005)
Mines	54		56			38		37					
Dth<24													
All	5,272	0.101	1677	0.088	-0.012	1,812	0.07	853	0.081	0.011	-0.031	-0.007	0.024
		(0.301)		(0.284)	(0.123)		(0.254)		(0.273)	(0.306)	(0)	(0.527)	(0.03)
Mines	54		56			38		37					

Notes: Standard errors and p-values in parentheses. Descriptive statistics of 12-month and 24-month mortality outcomes, for villages upstream and downstream of mining sites, for individuals born before and after the opening of a mine, over the balanced sample.

10.1.2 Heterogeneous treatment effects with two-way fixed effects: de Chaisemartin and d'Haultfœuille [2020]

The main result of this paper estimates the effect of being downstream of an open mine by using standard difference-in-difference designs. However, recent developments in the estimation of difference-in-differences in staggered adoption designs [Borusyak et al., 2021; Goodman-Bacon, 2018; Callaway and Sant'Anna, 2021; de Chaisemartin and d'Haultfœuille, 2020] show that the estimated ATT¹⁴ is a weighted sum of different ATTs with weights that may be negative. The negative weights are an issue when the treatment effect is heterogeneous between groups over time, as one could have the treatment coefficient in those regressions as negative while the treatment effect is positive in every group and time period. Using treated observations as controls creates these negative weights. In our design, the effect on Group 1 Switchers is compared to two control groups, Group 2 always treated and Group 3 never treated. The negative weights might come from the comparison of the effect of the Group 1 switchers to the Group 2 always treated. This biases the DiD estimator as it is an average of local treatment effects. In this section, we use the de Chaisemartin and d'Haultfœuille [2020] estimator which deals with the issue of negative weights in a staggered adoption design.

Table 14 compares the two-way fixed effects (TWFE) used in the main result (odd columns), to the de Chaisemartin and d'Haultfœuille [2020] estimator (dCDH) (even columns)¹⁵. Columns (1, 2, 5, 6) give the results for the entire sample, while columns (3, 4, 7, 8) for the balanced sample, defined in previous Section 10.1.1. Columns (1-4) give the results for the 12-month mortality rates, while columns (5-8) for the 24-month mortality rates.

First, let's look at the 24-month mortality rates. When looking at the TWFE estimator, we see that the results are stable on the balanced sample, even though it only represents 27% of the whole sample. This is coherent with the fact that the balanced sample keeps the villages that drive the main estimation's results. When focusing on the balanced sample, being downstream of an opened mine increases the 24 month-mortality rates by 3.19 p.p, which represents an increase of 36% of the mortality. Column (6) gives the dCDH estimator for the whole sample, while column (4) is for the balanced sample. We observe an increase in terms of the magnitude of the effect when correcting for negative weights, as being downstream of an open mine increases the 24-month mortality

¹⁴Average Treatment on the Treated

¹⁵The Stata command *did_multiplegt* is used to run the dCDH estimator.

rates by 11 p.p, which represents an increase of 129% of the mortality. If these magnitudes seem high, it is reassuring to observe the stability of the direction and significance of our main effect when using the dCDH estimator. Regarding the 12-month mortality rates, we observe that the restriction to the balanced sample displays a 2.8 p.p increase.

Table 14: Effects of industrial mining opening on 24-month mortality de Chaisemartin and d'Haultfœuille [2020]

	12-month mortality				24-month mortality			
	Whole Sample		Balanced Sample		Whole Sample		Balanced Sample	
	TWFE	dCDH	TWFE	dCDH	TWFE	dCDH	TWFE	dCDH
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Downstream×Open	-0.00506 [0.00831]	0.0249 [0.0262]	0.0286* [0.00222]	0.1457 [0.1132]	0.0218** [0.0108]	0.1109** [0.0405]	0.0319** [0.0162]	0.1667* [0.1112]
Downstream	-0.0152** [0.00665]		-0.0242*** [0.00826]		-0.0211*** [0.00739]		-0.0283*** [0.00866]	
Open	0.00963 [0.00754]		0.0347 [0.0320]		-0.00496 [0.0101]		0.0238 [0.0347]	
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Birthmonth FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Country-birthyear FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB-birthyear trend	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
N	48, 472	48, 472	9, 606	9, 606	35, 638	35, 638	9, 606	9, 606
R2	0.0378		0.0523		0.0511		0.0599	

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Column (1) gives the result of the main analysis, the Two Way Fixed Effect (TWFE) for the whole sample, while Column (2) gives the de Chaisemartin and d'Haultfœuille [2020] estimator. Columns (3) and (4) give the TWFE and de Chaisemartin and d'Haultfœuille [2020] estimators for the balanced sample.

10.2 Sensitivity analysis

10.2.1 Including non-topographic sub-basins

In this section, we replicate the main analysis from Table 2, adding within the control group, individuals living in a sub-basin with no topographic relation to the mine sub-basin, within 45 kilometers. This test can have several readings.

First, it strengthens the control for income effects linked to mining activity and enables to isolate more precisely of the channel of water pollution, and excludes other potential mechanisms. Indeed, villages close to the mine but located in a sub-basin with no topographic relationship with the mine, are allegedly less exposed to mining-induced water pollution and would be as exposed to income or labor effects, conflicts, or migration.

Yet, it also leads to the comparison of villages that do not necessarily share the same water resources, and this could blur the interpretation of our estimation. For example, other activities such as more intensive agriculture or livestock farming could aggregate around the mining site and could be responsible for other types of pollution. If these activities are located in a sub-basin with no topographic relationship to the mine, the estimated comparison would display the difference between the pollution of the mine and the pollution of these activities, rather than the pollution of the mine only, and this would lead to a downward bias to our analysis. Moreover, as mining activity is water intensive, the location of these activities might also be endogenous to the location of the mine, and this could induce an even larger downward bias.

Table 15 displays the results when including the non-topographic sub-basins within the control group. As expected, the table suggests that the main results of the 24-month mortality rates are downward biased and only significant at the 10% level for the rural population.

Table 15: Effects of industrial mining opening on infantile mortality - including DHS with non-topographic relationship

	12-month mortality				24-month mortality			
	Total Population		Rural Population		Total Population		Rural Population	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Downstream \times Open	-0.00266 [0.00543]	-0.00227 [0.00544]	0.00341 [0.00632]	0.00381 [0.00632]	0.00746 [0.00752]	0.00804 [0.00753]	0.0172* [0.00887]	0.0165* [0.00888]
Downstream	-0.00490 [0.00445]	-0.00472 [0.00446]	-0.00904* [0.00506]	-0.00884* [0.00507]	-0.00486 [0.00561]	-0.00455 [0.00561]	-0.0103 [0.00637]	-0.0106* [0.00636]
Open	0.00364 [0.00292]	0.00488 [0.00302]	0.00318 [0.00345]	0.00466 [0.00359]	0.00115 [0.00377]	0.00308 [0.00391]	0.00283 [0.00457]	0.000259 [0.00440]
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Nb open mines	No	Yes	No	Yes	No	Yes	No	Yes
Birthmonth FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Country-bthyr FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB-bthyr trd	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
N	168,931	168,931	123,413	123,413	124,670	124,670	91,395	91,395
R2	0.0214	0.0215	0.0252	0.0252	0.0305	0.0305	0.0356	0.0355
Outcome Mean	0.0638	0.0638	0.0670	0.0670	0.0824	0.0824	0.0872	0.0872

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Columns (1-3) give the results for the total population while columns (4-6) display the results for rural villages. The same controls as Table 2 apply. Columns (2, 4, 6, 8) control for the number of open mines within 45 km. The sample includes individuals living in non-topographic sub-basins within 45km.

10.2.2 Dropping fixed effects and other tests

Table 16: Effects of industrial mining opening on 24 months mortality, while dropping fixed-effects.

Outcome	24-month mortality			
	(1)	(2)	(3)	(4)
Downstream \times Open	0.0218** [0.0108]	0.0216** [0.0108]	0.0179* [0.0105]	0.0177* [0.0104]
Downstream	-0.0211*** [0.00739]	-0.0211*** [0.00738]	-0.0218*** [0.00734]	-0.0219*** [0.00733]
Open	-0.00496 [0.0101]	-0.00494 [0.0101]	-0.00459 [0.00972]	-0.00466 [0.00962]
Controls	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes
Country-birthyear FE	Yes	Yes	Yes	Yes
Birthmonth FE	Yes	No	No	No
Commodity FE	Yes	Yes	No	No
Mine SB-birthyear trend	Yes	Yes	Yes	No
N	35,638	35,638	35,638	35,638
R2	0.0511	0.0504	0.0503	0.0491
Outcome mean	0.0873	0.0873	0.0873	0.0873

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The same sample as Table 2 Column 2 apply.

We find stability in our results when dropping fixed effects one by one: birth month primary commodity, and sub-basins birth year trend (Table 16) until keeping the two-way fixed effects (i.e keeping the mine sub-basin fixed effect and the country-birthyear fixed effect).

Section F.1 runs other tests. Table 30 shows that our result is stable when controlling for the hand work. Figure 37 shows that the main results are stable when dropping countries one by one, and Figure 36 when dropping metals one by one.

10.2.3 Spatial correlation

As an additional robustness check, we run our main result's specification while taking into account the spatial correlation of DHS clusters. We estimate the standard errors with a spatial HAC correction following the method developed by ? and using the Stata command introduced by Colella et al. [2019]. Table 17 shows the stability of our results for different cut-off distances of spatial correlation (from 20 km to 200 km). We did not include directly the ? test in the main analysis as it does not allow for several fixed effects.

Table 17 corrects for spatial correlation for the results when using only Mine-Subbasin and country-birthyear fixed effects (result from Table 16 Column (4)).

Table 17: Effects of industrial mining activity, Conley spatial correction (acreg)

Outcome	Mortality under 24 months						
	Conley spatial correction threshold	20 km	45 km	60 km	80 km	100 km	
		(1)	(2)	(3)	(4)	(5)	
Downstream×Open		0.0177* [0.0100]	0.0177* [0.00999]	0.0177* [0.0101]	0.0177* [0.0101]	0.0177* [0.00996]	0.0177* [0.00916]
Downstream		-0.0219*** [0.00709]	-0.0219*** [0.00746]	-0.0219*** [0.00789]	-0.0219*** [0.00847]	-0.0219** [0.00913]	-0.0219** [0.0106]
Open		-0.00466 [0.00941]	-0.00466 [0.00932]	-0.00466 [0.00943]	-0.00466 [0.00960]	-0.00466 [0.00955]	-0.00466 [0.00938]
Controls	Yes	Yes	Yes	Yes	Yes	Yes	
Country-birthyear FE	Yes	Yes	Yes	Yes	Yes	Yes	
Mine SB FE	Yes	Yes	Yes	Yes	Yes	Yes	
N	35,648	35,648	35,648	35,648	35,648	35,648	
R2	0.00262	0.00262	0.00262	0.00262	0.00262	0.00262	

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The variables Downstream and Opened are dummies that indicate whether the individual lives in a village downstream of at least one mining site and whether the site opened before the birth year of the child. Each village DHS is paired to only one mining site so that each individual appears only once in the regression. Other variables are control variables. The sample focuses on heavy metal mines. Control variables are birth order number, mother's age, mother's age square, mother's years of education, urban, number of open mines, and presence of rivers.

10.3 Measurement errors

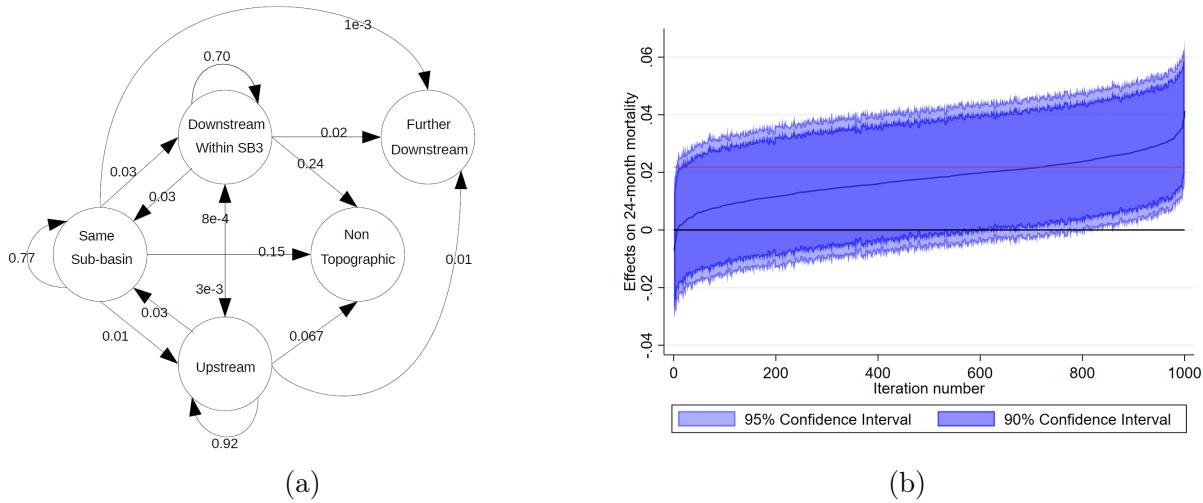
In this section, we deal with measurement errors that come from the nature of the data. Section 10.3.1 tests for the random displacement of DHS villages and section 10.3.2 tests our result according to the precision of the mine location.

10.3.1 DHS random displacement

DHS randomly displaces the GPS coordinates of each village to protect the confidentiality of respondents. Urban locations are displaced within 2 kilometers, while rural clusters are displaced within 5 kilometers, with 1% of rural clusters moved up to 10 kilometers. Displacements are made within administrative districts. This random reshuffling of DHS villages introduces measurement errors in our main estimation, all the more important as our treatment allocation depends on the relative position of the DHS villages to the mine.

First, we randomly displace 1,000 times each DHS village within a buffer of 2 kilometers for urban clusters and 5 kilometers for rural clusters. Thus, each displaced village is located in a new sub-basin, which can be the initial sub-basin or not. Then, we determine the topographic relation of this sub-basin to the sub-basin of the mine, which gives the treatment of the DHS village: whether it falls into a sub-basin upstream, downstream, in the same sub-basin as the mine or in a sub-basin with no topographic relationship with the one of the mine. The topographic relation of the new sub-basin gives the new treatment status of the DHS village. We only reshuffled the position of DHS villages that have a topographic relation with the mine initially, and that were up to the third sub-basin downstream. This means that we can have some DHS villages that exit the main sample, for instance, if their newly assigned sub-basin has no topographic relation, or is downstream in the fourth sub-basin, or if it falls into the same sub-basin as the mine, as these cases are excluded from the main result. The only new observations that come within the sample are DHS villages that were initially within the same sub-basin as the mine and fall upstream, or downstream with the new iteration of the random displacement of their location. Please note that it is possible as well, but very rare, that a DHS village falls into the ocean.

Figure 18: DHS random displacement - 1,000 iterations



Notes: Figure (a) plots the transition probability graph for 1,000 random displacements of DHS clusters. Figure (b) plots the interaction term for 1,000 different regressions, each done for a new sample where DHS GPS coordinates have been randomly displaced. The red line $y=0.0218$ plots the coefficient from our main result. The coefficients are ordered, and we plot the 95% and 90% intervals.

Sources: Authors' elaboration.

Figure 18a gives the probability graph showing the transition probabilities of changing treatment status. For instance, after 1,000 iterations, a DHS village initially downstream within the third sub-basin has 70% chances to remain downstream up to sub-basin three, has 0.3% chances to be upstream the mine, 24% chances to fall into a sub-basin with no topographic relation (and be out of the sample), 3.5% chances to be in the same sub-basin of the mine and finally 2% chances to be downstream further than the third sub-basin (and be out of the sample). In the end, a DHS village treated in our initial sample has 25% chances to leave the sample. Please note that this random reshuffling is not perfect, as DHS villages should be reshuffled within administrative level 2 boundaries as made in the DHS procedure.

Figure 18b plots the interaction term *Downstream* \times *Open* of our main estimation for 1,000 random displacements of DHS GPS coordinates. The coefficients are ordered, and we plot the 95% and 90% intervals. As there is a higher probability that a DHS cluster leaves the sample rather than a new enters it, the number of observations varies for each iteration and is likely to be smaller than our main estimation.

10.3.2 Accuracy of mine location

We further test for potential measurement errors by looking at the precision of the mines' location. The SNL database provides information on the accuracy levels of each mine's GPS coordinates enables us to restrict the analysis to mines with exact coordinates, precise at 1 km. Our main results are positive but no longer significant when restricting to the mines with exact coordinates but hold when focusing on rural households. This hints towards a higher effect of industrial mining activity on child mortality among rural households, and a lack of precision in the location of mines close to urban areas.

Table 18: Effects of industrial mining opening, restriction to exact GPS coordinates.

Outcome		24-month mortality		
Accuracy level	All	Exact coordinates		
Sample	Urban and rural	Urban and rural	Rural	
	(1)	(2)	(3)	
Downstream \times Open	0.0218** [0.0108]	0.0116 [0.0116]	0.0294** [0.0143]	
Downstream	-0.0211*** [0.00739]	-0.0204** [0.00807]	-0.0239*** [0.00888]	
Open	-0.00496 [0.0101]	0.00532 [0.0104]	0.00805 [0.0130]	
Controls	Yes	Yes	Yes	
Birthmonth FE	Yes	Yes	Yes	
Country-birthyear FE	Yes	Yes	Yes	
Mine SB FE	Yes	Yes	Yes	
Mine SB-birthyear trend	Yes	Yes	Yes	
Commodity FE	Yes	Yes	Yes	
N	35,638	29,195	20,172	
R2	0.0511	0.0517	0.0626	
Outcome mean	0.0873	0.0858	0.0920	

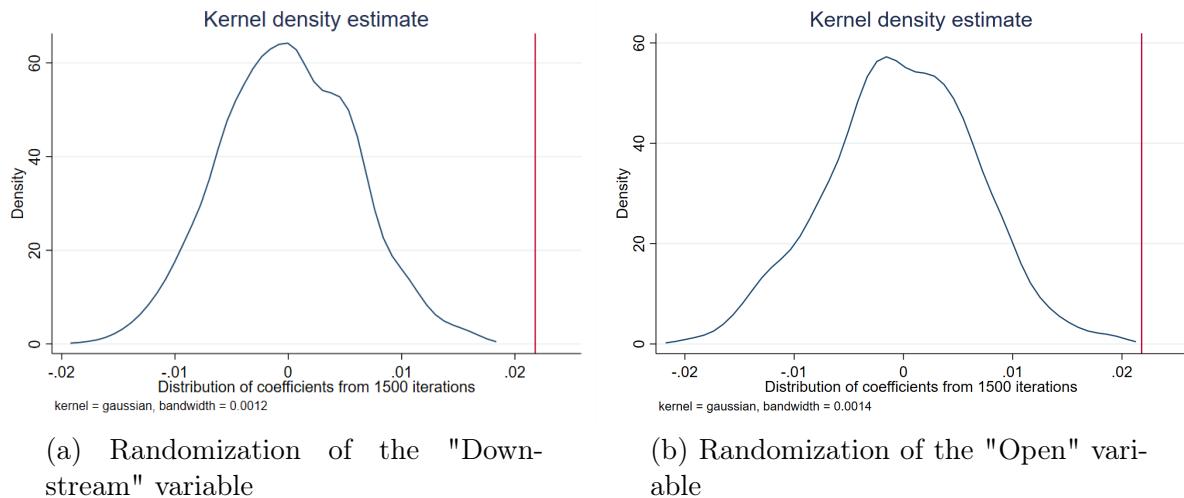
Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The same controls as Table 2 Column 2 apply.

10.4 Placebo tests

10.4.1 Randomization inference

To make sure that the assignment of each village to its topographic position relative to the mine is indeed what drives our result on child mortality, we run a randomization inference test. We draw randomly 1,500 permutations of the "Downstream" variable without changing the start-up year and 1,500 permutations of the "Open" variable without changing the downstream position¹⁶. The simulations show that the distribution of treatment effects ($\text{Downstream} \times \text{Open}$) are shifted around zero (Figure 19). The red line represents the initial treatment effect using our main specification: we are sure at the 1 percent level that our main model is not misspecified.

Figure 19: Spatial and temporal randomization inference tests



Notes: The two figures represent the distribution of coefficients associated with the interaction term of being downstream of an open mine and its effect on under 24-month mortality when conducting 1,500 permutations of the "Downstream" position of each DHS sub-basin (Figure (a)) and 1,500 permutations of the "Open" variable (Figure (b)). The red line represents the initial treatment effect using our main specification.

Sources: Authors' elaboration using the Stata *ritest* command.

10.4.2 Placebo diseases

We conduct a placebo test on other potential diseases that could affect women's and thus children's mortality. We do not find significant industrial mining on the infection of any sexually transmitted disease among women living downstream of an open mine (Table 19

¹⁶The randomization inference of the "Downstream" and "Open" treatment are within the sub-basin level, and are clustered at the DHS village level.

column 1) or among awareness of tuberculosis (column 2). This absence of differential results on women's health across upstream and downstream villages is reassuring for our identification of the water pollution channel.

Table 19: Effects of industrial mining opening on women, placebo diseases.

Outcome	(1)	(2)
	Any sexually transmitted infection	Heard of tuberculosis
Downstream × Open	0.00332 [0.00923]	-0.0387 [0.0259]
Downstream	0.00751 [0.00772]	0.0277 [0.0210]
Open	0.00766 [0.00913]	0.00469 [0.0314]
Controls	Yes	Yes
Country-survey year FE	Yes	Yes
Mine SB FE	Yes	Yes
Mine SB-survey year trend	Yes	Yes
Commodity FE	Yes	Yes
N	66,653	14,750
R2	0.0888	0.186
Outcome mean	0.0501	0.938

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The same sample and controls as Table 4 apply.

11 Discussion and limits

11.1 Selection issues

Our study includes the best available data on mining and child health we have at the scale of a continent, but of course, is neither exhaustive nor represents the whole African continent. First, as we limited our sample to countries with at least two waves of DHS, we had to drop many countries with only one wave among which South Africa which has intense mining activity. Future work will be possible once other survey waves will have been conducted. Our 26 countries represent about two-thirds of the total population of the continent¹⁷.

Second, our mining data is also limited to industrial sites, and do not include artisanal or small-scale mining which information is much harder to retrieve at the scale of the continent. One possibility for another project is to use data on the suitability of artisanal

¹⁷The proportion is stable between 1981 and 2020.

gold mining [Girard et al., 2022] and to compare their environmental impacts to industrial mining.

One remaining concern is about the exhaustivity of the SNL, and the heterogeneity of the sampling selection across countries. We tried our best to evaluate the exhaustivity of the SNL data by comparing it with other mining data and social sources (Ministries, USGS, mining website. . .) but it is out of our feasible means to get an exact proportion of representativity and to compete with the business-oriented activity of SNL.

11.2 Threats to the identifying assumption

11.2.1 Type of pollution

First, our study mainly focuses on water pollution through the lens of water subbasins while controlling for rivers, i.e. surface water. Yet, further work could also control for groundwater as their pollution may follow different dynamics than surface water: villages located in areas with low-depth groundwater could pump the water more easily and with more affordable water pumps than in areas with deeper groundwater. These former villages could therefore be more exposed to mining-induced water pollution than the latter ones. Moreover, groundwater could take more time to be contaminated by mining-induced pollution than surface water, but its contamination could also last more permanently.

This paper does not directly examine mining-induced air pollution. The main hypothesis is that wind direction is less correlated to the topographic position of the village than water pollution and that the comparison between upstream and downstream villages should exclude the effect of air pollution. Besides, as discussed in Section 2, the effects of air pollution seem to concern the mine workers more than the surrounding population, even though fine particles can be displaced over long distances. However, our main result is not entirely net off the impacts of air pollution. The best control included so far is adding the sub-basins with no topographic relationship to the mine, as they are allegedly exposed to air pollution only, while sub-basins with a topographic relationship would be exposed to both water and air pollution. It is beyond the scope of the current paper to take into account the direction of the wind to disentangle both sources of pollution, but empirically feasible for another paper. Our study is most likely an underestimation of the total pollution induced by industrial mining activity.

The same concern remains for soil pollution. An additional heterogeneity analysis would be to take into account areas prone to subsistence agriculture or livestock, as mining-induced water pollution could also contaminate soils and cattle in the long run and the subsequent food produced. We assume more harmful effects of industrial mining on local population health if both water and food are polluted. In another paper, one could study the heterogeneous effects across the global agroecological zones (GAEZ) and crop suitability, and test whether villages located near industrial mining sites in high-yield crop areas are more affected than villages with less suitable soils.

11.2.2 Threats to identification

A major threat to identification is that the opening of a mine may not be orthogonal to unobservable factors that affect health and water quality, in different ways for downstream and upstream areas.

Migration is a major methodological concern, as we show in Table 28 of Appendix D.1 that migrants significantly settle downstream after a mine opening. Section 6.2 shows that our main result is robust when controlling for in-migration. A main violation of the identifying assumption would be if downstream villages anticipate the mine opening and strategically out-migrate within upstream areas to avoid pollution. In this case, there would be a selection bias, as the individuals surveyed downstream after the mine opening would be those that were not able to migrate or anticipate the pollution. Controlling for in-migration in DHS villages, we show that our result is robust to this specific strategic behavior. However, we cannot control for strategic out-migration outside of the study area, meaning individuals out-migrating to avoid pollution elsewhere than the upstream area. In this paper, we made the choice not to use mother fixed-effects and retrospective questions on birth history, to limit endogenous selection due to out-migration, and to account for children born up to five years prior to the year of the survey.

Accordingly, a threat to the identification would be a differed improved access to infrastructure associated with the opening of a mine between upstream and downstream areas. Table 28 shows no difference in terms of access to electricity and piped water between downstream and upstream areas after a mine opening, and Section 6 shows that our result is robust controlling for improved access to facilities.

An important omitted variable in our current study is the increased presence of con-

flicts and violence around areas with mining activity, as shown by Berman et al. [2017] and that could also explain the increase in child mortality in the vicinity of mines. We do not directly control for conflict, but there would be an upward bias of our estimation only if conflicts systematically happen more downstream than upstream. As our results hold when including non-topographic subbasins in rural areas, it is a first-step approximation that water pollution is indeed the main explaining factor of increased child mortality. Further work could include the ACLED data to exclude this mechanism.

Another concern is that other industries could aggregate around the mining industry and be partly responsible for the pollution. More than a bias, this could be a threat to identification if the location of the industry is correlated to the topographic position of the mine. In another paper, we could look at the correlation between mining activity and other industry implementations. Controlling for them could enable us to isolate the pollution linked to the mining activity solely.

12 Policy discussion

In this section, we first try to compute how many deaths were related to the water pollution linked to industrial mining activity in the 26 countries of our sample. Then, we try to assess whether the Extractive Industries Transparency Initiative, a global standard for good governance in the extractive sector, has been successful to reduce this mortality.

12.1 Back-of-the-envelope calculation

In this section, we compute a back-of-the-envelope calculation to grasp how many deaths could have been averted had there been policies implemented to limit water pollution, over the 1981-2020 period and within the 26 Sub-Saharan countries of our analysis.

First, we consider that as DHS is representative at the national level, it is feasible to calculate the proportion of individuals living within 45 kilometers of a mine, the proportion of those living downstream, etc. Here are the probabilities computed using the DHS database:

- $x = 28\%$: Proportion of individuals living within 45 kilometers of a mine¹⁸

¹⁸Please note that, exactly, this is the proportion of individuals living within 45 kilometers of a mine

- $x = x_d + x_u + x_{nt} + x_{sb}$, with
 - $x_d = 1.94\%$: Proportion of individuals living downstream ^{¹⁹}
 - $x_u = 5.25\%$: Proportion of individuals living upstream
 - $x_{nt} = 17.93\%$: Proportion of individuals living with no topographic relation
 - $x_{sb} = 2.92\%$: Proportion of individuals living in the same sub-basin as a mine

Our analysis leads to an estimation of $e = 2.18\%$ the increased mortality rate because of industrial mining-induced water pollution. In our sample, 9,258 individuals live downstream of a mine and we count 822 deaths among them. The total number of additional deaths due to mining-induced water pollution is $9,258 \times 2.18\% = 202$ deaths. We now look at the 880 million children who were aged 0-2 years over 1981-2020 in our 26 countries ^{²⁰}. As we assume the representativity of the DHS surveys and the stable proportion of the population living in the vicinity of mines, this would mean that $1.94\% \times 880$ million = 17 million children lived within 45 km downstream of a mine. This leads to $2.18\% \times 17$ million = 370,600 deaths due to mining-induced water pollution over 1981-2020 in our 26 countries, i.e. 9,265 deaths per year, or 16 deaths per mine per year.^{²¹} To grasp a better sense of the magnitude of this figure, there are on average 840,000 births per year and per country (average within the 26 countries over 1981-2020), which means that the number of deaths caused by mining-induced water pollution over 26 countries represents 1.1% of the number of births per country^{²²}.

12.2 Extractive Industries Transparency Initiative members

We look at whether there is a significant difference across countries that have signed the Extractive Industries Transparency Initiative, launched in 2002 and which currently gathers 55 countries. Member countries commit to disclose information along the production value chain of oil, gas, and mining extraction and respect a common set of governing standards. We want to see if there is an effect of the EITI Rules signed by the member countries on the effect of industrial mining on child mortality. 18 out of the 26 countries included in our sample signed the EITI^{²³} which gather 76 percent of our sample of children.

and downstream up to the third sub-basin.

¹⁹up to the third sub-basin

²⁰Source: World Bank data.

²¹There are 604 mines in total in our main results' regressions.

²²As sampling weights are not considered in the calculation, we do not give a number per country.

²³Burkina Faso, Côte d'Ivoire, Democratic Republic of the Congo, Ethiopia, Ghana, Guinea, Liberia, Madagascar, Malawi, Mali, Niger, Nigeria, Senegal, Sierra Leone, Tanzania, Togo, Uganda, and Zambia are EITI members. Benin, Burundi, Kenya, Namibia, Rwanda, and Zimbabwe have not joined the EITI.

Table 20: Effects of industrial mining opening, across EITI membership.

Sample	Outcome	24-month mortality								
		All		Rural						
		Not an EITI member	EITI member	Not an EITI member	EITI member	(1)	(2)	(3)	(4)	(5)
Downstream × Open		0.0179 [0.0210]	0.0259** [0.0127]	0.0133 [0.0190]	-0.0428 [0.0518]	-0.0221 [0.0416]				-0.0557 [0.0669]
Surveyed after joining EITI				0.0237 [0.0303]						-0.0259 [0.0315]
D × O × Surv. after joining EITI				0.0243 [0.0236]						0.0221 [0.0705]
Downstream		-0.0552*** [0.0140]	-0.00934 [0.00876]	-0.00446 [0.0107]	0.00399 [0.0500]	0.0356 [0.0381]				0.0118 [0.0641]
Open		-0.00576 [0.0246]	-0.00588 [0.0112]	-0.00656 [0.0163]	-0.0456 [0.0651]	-0.00345 [0.0268]				-0.0114 [0.0443]
Controls		Yes	Yes	Yes	Yes	Yes				Yes
Birthmonth FE		Yes	Yes	Yes	Yes	Yes				Yes
Country-birthyear FE		Yes	Yes	Yes	Yes	Yes				Yes
Mine SB FE		Yes	Yes	Yes	Yes	Yes				Yes
Mine SB-birthyear trend		Yes	Yes	Yes	Yes	Yes				Yes
Commodity FE		Yes	Yes	Yes	Yes	Yes				Yes
N		8,434	26,810	26,810	2,251	8,685				8,685
R2		0.0373	0.0548	0.0548	0.0838	0.0677				0.0679
Outcome mean		0.0716	0.0920	0.0920	0.0604	0.0738				0.0738

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The same sample and controls as Table 2 Column 2 apply.

We estimate our main specification across the sample of countries that are members of the EITI or not. We find that our results hold even for countries who committed to improved governance of their extractive industries (Table 20), but which are also countries heavily relying on this activity in their national economy. We find no significant effect of our results when looking at whether surveys were conducted before or after their country signed the EITI Standards (triple interaction Downstream \times Open \times Surveyed after joining EITI in columns (3) and (6)).

13 Conclusion

This paper identifies a negative externality of industrial mining on local population living standards, as we show that industrial mining sites increase infant mortality in surrounding villages, indirectly through the contamination of water resources. We match geocoded repeated-cross sectional household surveys to geocoded data on industrial mine openings obtained through intensive handwork. We propose a staggered Difference-in-Difference strategy and isolate the mechanism of water pollution by building the treatment and control groups using an upstream-downstream comparison. We compare the effect on the health of villages located upstream and downstream of a mine deposit, before and after its opening. We are the first, to the best of our knowledge, to take into account the topography of mining areas using an upstream-downstream comparison and to empirically quantify this effect at the scale of 604 mines in 26 countries of Sub-Saharan Africa over 1981-2020.

We find that the opening of industrial mines increases by 25% the 24-month mortality rate among villages located downstream compared to villages located upstream, and thus indirectly isolates the channel of water pollution. We find almost no effects on other children's health outcomes, such as anthropometric measures, cough, fever, diarrhea, or anemia. We exploit the variation of the opening of a mine and show that our results are not driven by a change in women's fertility behavior, differential access to piped water, electricity, or health facilities but mainly by mining-induced water pollution, as children who were given plain water show increased mortality. The heterogeneity in the consumption of plain water seems to explain the null result on the 12-month mortality rates, as we observe a significant increase in the 12-month mortality rates exclusively for those who consume plain water. This can be interpreted as a proxy for having non-exclusive breastfeeding.

In an additional heterogeneity analysis, we show that our results are mainly driven by the pollution occurring during the time of mining activity. We find that the effects are even more harmful in rural areas, for open-pit and foreign-owned mines, and in places with a high density of mines. We also find that the effects increase with productivity intensity (proxied by international commodity prices). We run manyfold robustness checks and find that our results hold when controlling for in-migration, and when restricting to a balanced sample which deals with the issue of repeated cross-section surveys. Our results are also robust to the heterogeneous treatment effects estimator of de Chaisemartin and d'Haultfoeuille [2020], to measurement error tests, and a battery of placebo tests such as spatial and temporal randomization inference tests.

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A Appendix

B Descriptive Statistics

B.1 Data

Table 21 displays for each country the number and years of DHS waves, and the total number of DHS clusters and children under 5 years old. Overall, DHS sample gathers 36 countries overall Africa, from 1986 to 2018. In our main empirical analysis, we decided to only keep DHS countries that had at least two survey rounds, in order to have comparable temporal variation across countries. Our final sample accounts for the following countries (cf. Table 22): Tanzania, Burkina-Faso, Ghana, Zimbabwe, Mali, Democratic Republic of Congo, Guinea, Namibia, Madagascar, Cote d'Ivoire, Sierra Leone, Liberia, Nigeria, Senegal, Ethiopia, Uganda, Botswana, Malawi, Cameroon, Morocco, Niger, Kenya, Mauritania, Rwanda, Burundi, Lesotho, Togo, Eswatini, Algeria, Benin, Eritrea, Republic of the Congo, Guinea-Bissau, Somalia, Sudan, Tunisia, Djibouti, Equatorial Guinea (by order of importance in terms of mining activity according to Figure 23).

Table 21: DHS surveys overall across countries

Countries	Survey Years	Number of clusters	Number of children under 5
AO	2015	625	14,177
BF	1993, 1999, 2003, 2010	1,413	36,744
BJ	1996, 2001, 2012, 2017	1,752	31,884
BU	2010, 2016	930	20,824
CD	2007, 2013	836	27,307
CF	1994	230	2,639
CI	1994, 1998, 2012	674	12,227
CM	1991, 2004, 2011, 2018	1,619	31,279
EG	1992, 1995, 2000, 2003, 2005, 2008, 2014	7,741	75,394
ET	2000, 2005, 2010, 2016	2,313	42,173
GA	2012	334	5,911
GH	1993, 1998, 2003, 2008, 2014	2,037	17,931
GN	1999, 2005, 2012, 2018	1,289	26,588
KE	2003, 2008, 2014	2,391	32,235
KM	2012	252	3,134
LB	1986, 2007, 2013	776	16,224
LS	2004, 2009, 2014	1,199	10,269
MA	2003	480	6,030
MD	1997, 2008	860	15,932
ML	1996, 2001, 2006, 2012, 2018	1,867	52,996
MW	2000, 2004, 2010, 2015	2,655	56,688
MZ	2011	610	10,950
NG	1990, 2003, 2008, 2013, 2018	3,830	106,848
NI	1992, 1998	503	11,332
NM	2000, 2006, 2013	1,290	13,630
RW	2005, 2008, 2010, 2014	1176	21,927
SL	2008, 2013	787	17,483
SN	1993, 1997, 2005, 2010, 2012, 2014, 2015, 2016, 2017	2,572	73,084
SZ	2006	274	2,706
TD	2014	624	18,441
TG	1988, 1998, 2013	768	13,869
TZ	1999, 2010, 2015	1,259	20,520
UG	2000, 2006, 2011, 2016	1,765	37,603
ZA	2017	671	3,397
ZM	2007, 2013, 2018	1,585	29,105
ZW	1999, 2005, 2010, 2015	1,431	19,847

Notes: This table gives the sample size of children under five years old overall DHS surveys.

Table 22: DHS surveys in regression sample across countries

Countries	Survey Years	Number of clusters	Number of children under 5
BF	1993, 1999, 2003, 2010	694	23,846
BJ	2001, 2012, 2017	62	1,911
BU	2010, 2016	317	8,280
CD	2007, 2013	82	5,092
CI	1994, 1998, 2012	196	4,838
CM	1991, 2004, 2011, 2018	90	2,513
ET	2000, 2005, 2010, 2016	100	2,956
GH	1993, 1998, 2003, 2008, 2014	1,217	12,074
GN	1999, 2005, 2012, 2018	360	11,775
KE	2003, 2008, 2014	233	4,130
LB	1986, 2007, 2013	190	7,537
LS	2004, 2009, 2014	336	2,810
MD	1997, 2008	131	3,301
ML	1996, 2001, 2006, 2012, 2018	570	19,147
MW	2000, 2004, 2010, 2015	207	6,651
NG	1990, 2003, 2008, 2013, 2018	105	3,993
NI	1992, 1998	40	1,105
NM	2000, 2006, 2013	138	2,175
RW	2005, 2008, 2010, 2014	713	14,615
SL	2008, 2013	377	13,717
SN	1993, 1997, 2005, 2010, 2012, 2014, 2015, 2016, 2017	363	10,111
TG	1988, 1998, 2013	104	2,187
TZ	1999, 2010, 2015	325	6,866
UG	2000, 2006, 2011, 2016	305	9,031
ZM	2007, 2013, 2018	364	10,966
ZW	1999, 2005, 2010, 2015	468	8,307

Notes: This table gives the sample size of children under five years old that are in our main analysis, meaning within 100 km of an industrial mine.

Tables 23, 24, and 25 display the descriptive statistics of all our outcome and control variables for the sample of all individuals living within 45 km of an industrial mine, regardless of their topographic position, in the 26 countries of Sub-Saharan Africa with at least 2 waves of DHS and for heavy metals and coal mines. These descriptive figures are important to show that our analysis does not suffer from selection biases across the samples we use for our different regressions.

Table 23: Descriptive statistics of children's outcomes

	Mean	SD	Med	Min	Max	N
Mortality rates						
12-month mortality	0.064	.244	0	0	1	189,181
24-month mortality	0.083	.275	0	0	1	139,683
Control variables						
Birth order number	3.655	2.421	3	1	18	240,431
Male	0.508	0.500	1	0	1	240,431
Anthropometric measures						
Stunting	0.319	0.466	0	0	1	137,834
Underweight	0.234	0.423	0	0	1	136,043
Wasting	0.077	0.267	0	0	1	138,222
Weight and size at birth						
Less than 2.5 kg	0.164	0.370	0	0	1	117,651
Small or very small size	0.161	0.367	0	0	1	226,796
Measured anemia level						
Any anemia	0.633	0.482	1	0	1	67,567
Illness in the last 2 weeks						
Diarrhea	0.168	0.374	0	0	1	216,097
Cough	0.260	0.439	0	0	1	214,940
Fever	0.265	0.441	0	0	1	214,913
Nutrition						
Given plain water	0.187	0.390	0	0	1	122,915
Ever breastfed	0.980	0.140	1	0	1	223,039
Months breastfed	14.788	8.917	15	0	59	156,011
Health access						
No prenatal care	0.101	0.301	0	0	1	169,268
Ever vaccinated	0.788	0.409	1	0	1	82,082
Characteristic of paired mine						
Domestic mine	0.177	0.381	0	0	1	240,431
Open-pit mine	0.676	0.468	1	0	1	103,667

Notes: We present the mortality rates at n months, conditionnally on having reached n months, for the whole sample of children living within 45 km of an industrial mine and regardless of their topographic position. The sample is restricted to the 26 Sub-Saharan countries with at least two waves of DHS and to heavy metals and coal mines.

Table 24: Descriptive statistics of mothers' outcomes

	Mean	SD	Med	Min	Max	N
Mother's characteristics						
Mother's age	28.918	6.979	28	15	49	240,431
Years of education	3.985	4.226	3	0	22	240,332
Urban	0.287	0.452	0	0	1	236,966
Migrant	0.594	0.491	1	0	1	161,292
Access to sanitation and health facilities						
Piped water as main drinking water source	0.261	0.439	0	0	1	240,431
Has flushed toilet	0.086	0.280	0	0	1	239,773
Has electricity	0.218	0.413	0	0	1	236,692
Visited health facility in the last 12 months	0.623	0.485	1	0	1	218,053

Notes: The sample is restricted to all mothers of 0-5 years old children living within 45 km of an industrial mine in the 26 Sub-Saharan countries with at least two waves of DHS and to heavy metals and coal mines.

Table 25: Descriptive statistics of women's outcomes

	Mean	SD	Med	Min	Max	N
Fertility behavior and health						
Ever had a child	0.736	0.441	1	0	1	330,889
Total lifetime fertility	2.890	2.785	2	0	18	330,889
Currently pregnant	0.091	0.288	0	0	1	330,744
Ever had a miscarriage	0.127	0.333	0	0	1	296,235
Any anemia	0.378	0.485	0	0	1	115,481
Placebo disease						
Any STD	0.049	0.216	0	0	1	276,924
Heard of tuberculosis	0.935	0.246	1	0	1	88,438

Notes: The sample is restricted to all women aged 15-49 living within 45 km of an industrial mine in the 26 Sub-Saharan countries with at least two waves of DHS and to heavy metals and coal mines.

B.2 Handwork

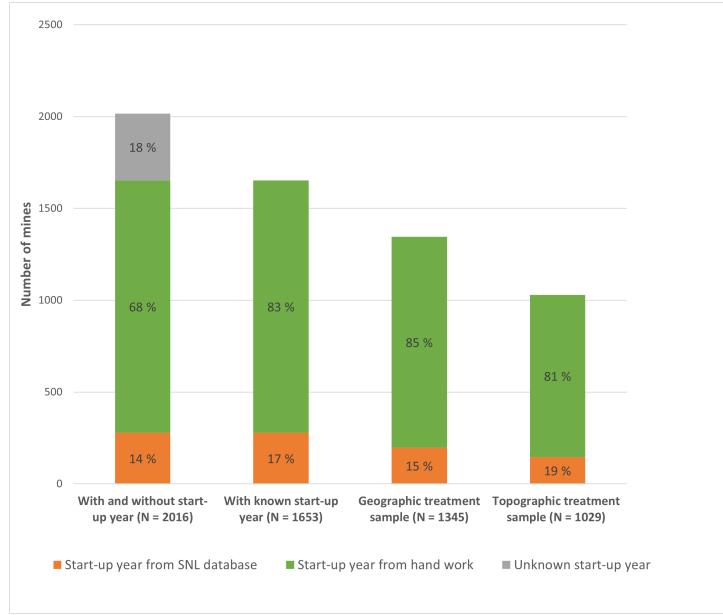
Out of the 3815 industrial mines recorded by the SNL database in Africa, 2016 were located within 100 km of a DHS cluster (with at least 2 waves of DHS). 278 had information on the opening and closing years within the database, and for the 1738 remaining mines, we searched for their years of opening ²⁴. The handwork consisted in reading the reports (comments and work history) available in the database and browsing through the aerial images available on the SNL platform which provided the exact GPS coordinates and main location labels. This information was corroborated with online research (press releases, mining companies' websites, specialized websites on global mining activities, etc.) as well as Google maps and Google timelapse satellite imagery. A mine opening corresponds to the beginning of the production.

The exact startup year could not be determined for 18 % of our sample (Figure 20 Bar (1)), and these mines are dropped in our regressions. In total we hand-checked 83% of the mines located within 100 km of a DHS cluster, and for which we know their year of opening (Figure 20 Bar (2)). Among the sample of mines with startup year, 83.2 % opened after 1981 (first year of birth within the DHS child surveys). For each of the following graphs, we study the whole sample of 2016 mines and plot the percentage of mines that were hand-checked and the percentage of mines that ends up having a startup year and are thus included in our study. We conduct this analysis on all the available mines within 100 km of a DHS cluster to be transparent on the creation of our sample compared to the original one.

The distribution across each mining site's primary commodity of production can be found in Figure 22. Half of our sample consists of gold mining sites. Figure 23 represents the distribution across country of location. Ownership information is available for 65 percent of our sample and the main owners are from the USA, UK, Canada, Australia, and China (Figure 25).

²⁴We also looked at their closure date as well as their current activity status, i.e. whether the mining site looked active or inactive. However, this was an information harder to retrieve and finally we focused on the date of opening.

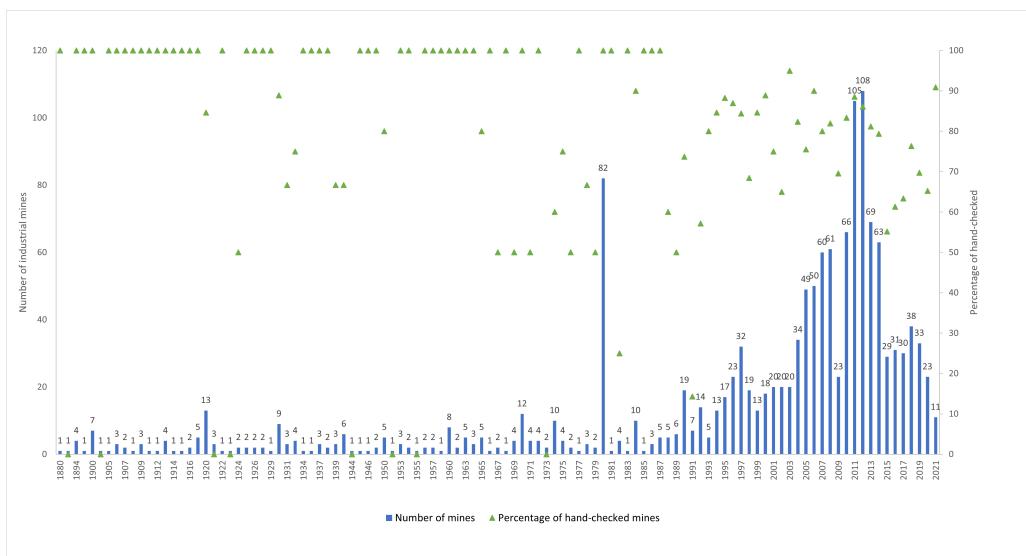
Figure 20: Description of hand work and industrial mines samples



Notes: This Figure gives the number and percentage of mines for which we have retrieved the year of opening by hand. Bars (1) and (2) gives it for all the mines located within 100 km of a DHS cluster, Bar (3) for the sample associated to the replication of Benshaul-Tolonen [2018] Section G. Bar (4) corresponds to the main analysis, i.e to mines that have at least one DHS cluster upstream within 100km, and one DHS cluster downstream within the three closest sub-basins (cf pairing strategy Section4.1.1) .

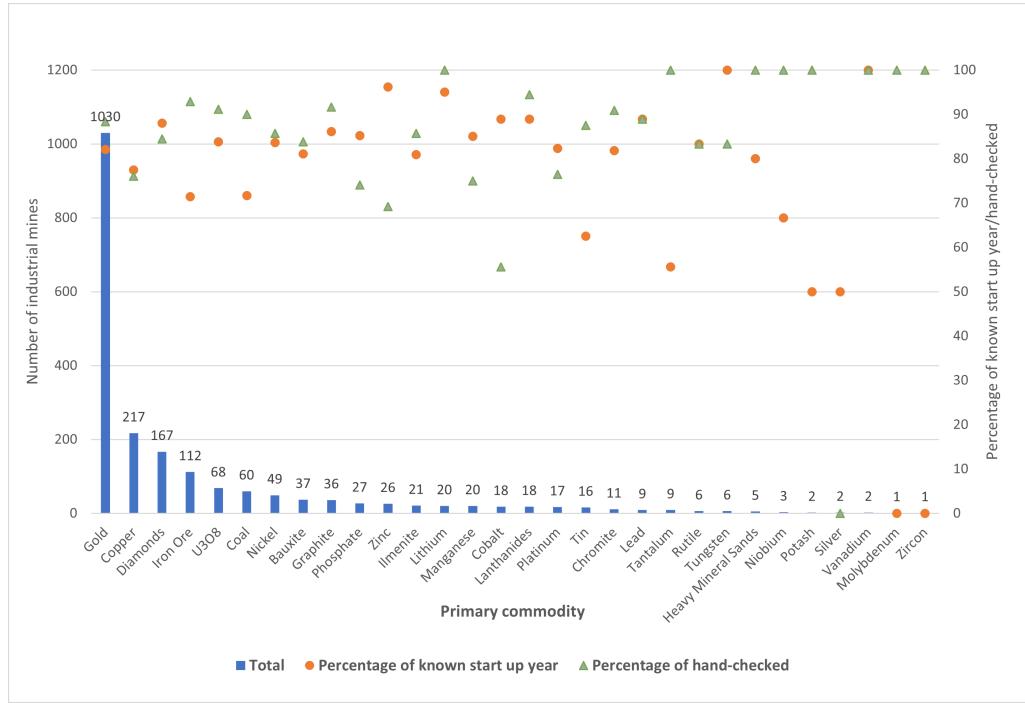
Sources: Authors' elaboration on DHS and SNL data.

Figure 21: Mines and percentage of hand-checked across start-up years



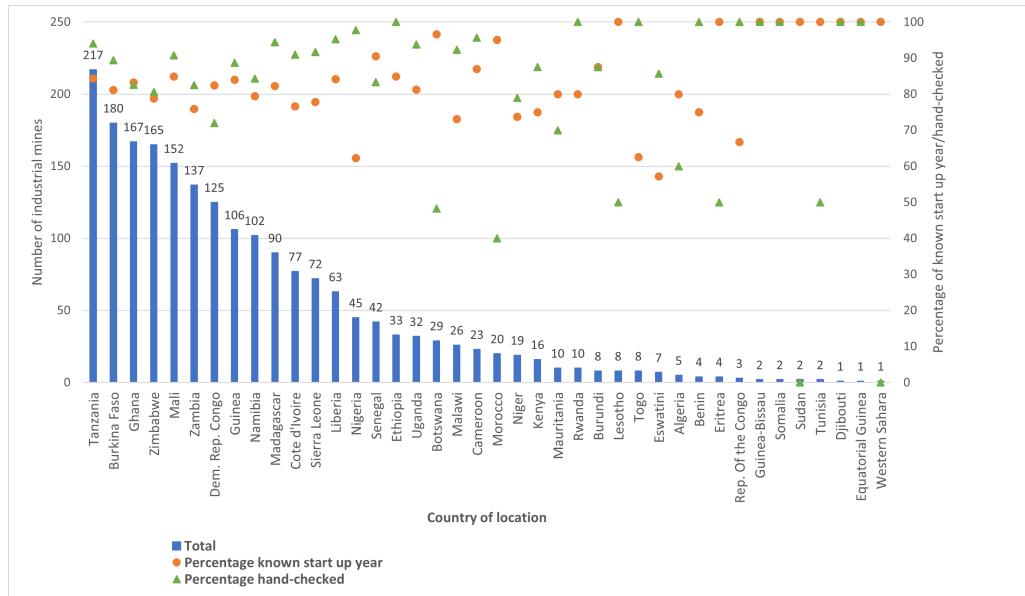
Notes: This graph displays the number of mines that opened during a specific year and the percentage of hand checked for the 2016 mines located within 100km of a DHS cluster.

Figure 22: Mines and percentage of hand-checked across primary commodities



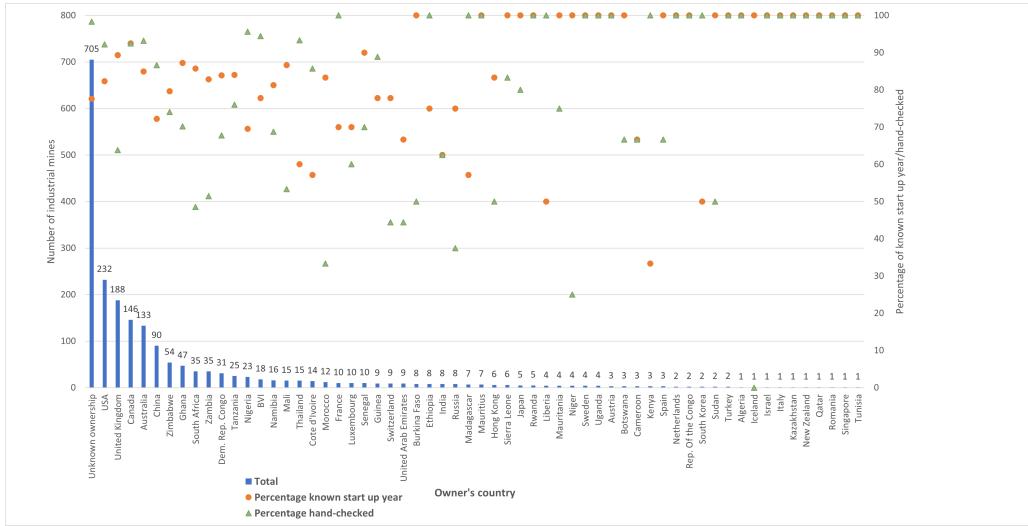
Notes: This graph gives the number of mines for each primary commodity, and the percentage of hand-checked, for the 2016 mines located within 100 km of a DHS cluster.

Figure 23: Mines and percentage of hand-checked across country of location



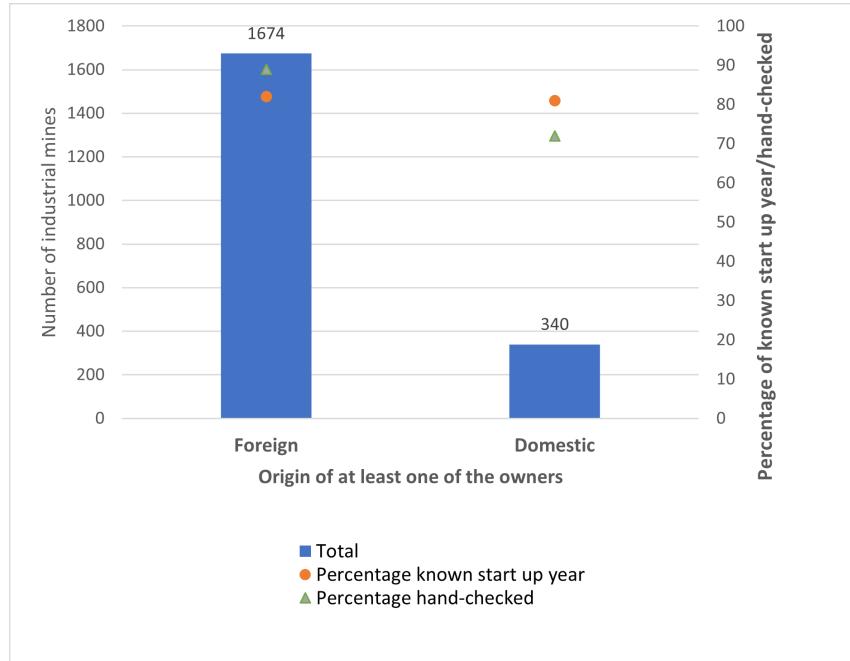
Notes: This graph gives the number of mines for each country of location, and the percentage of hand-checked, for the 2016 mines located within 100 km of a DHS cluster.

Figure 24: Mines and percentage of hand-checked across owner's country



Notes: This graph gives the number of mines by owning company's registration country, and the percentage of hand-checked, for the 2016 mines located within 100 km of a DHS cluster.

Figure 25: Mines across foreign and domestic ownership



Notes: This graph gives the number of mines across domestic and foreign ownership, and the percentage of hand-checked, for the 2016 mines located within 100 km of a DHS cluster.

C Context

C.1 Case study: the Essakane mine

Figure 26: Satellite image of Essakane Mine in 2019



Notes: Satellite image of the Essakane Mine in 2019. Retention dams can be seen.

Sources: Google Earth.

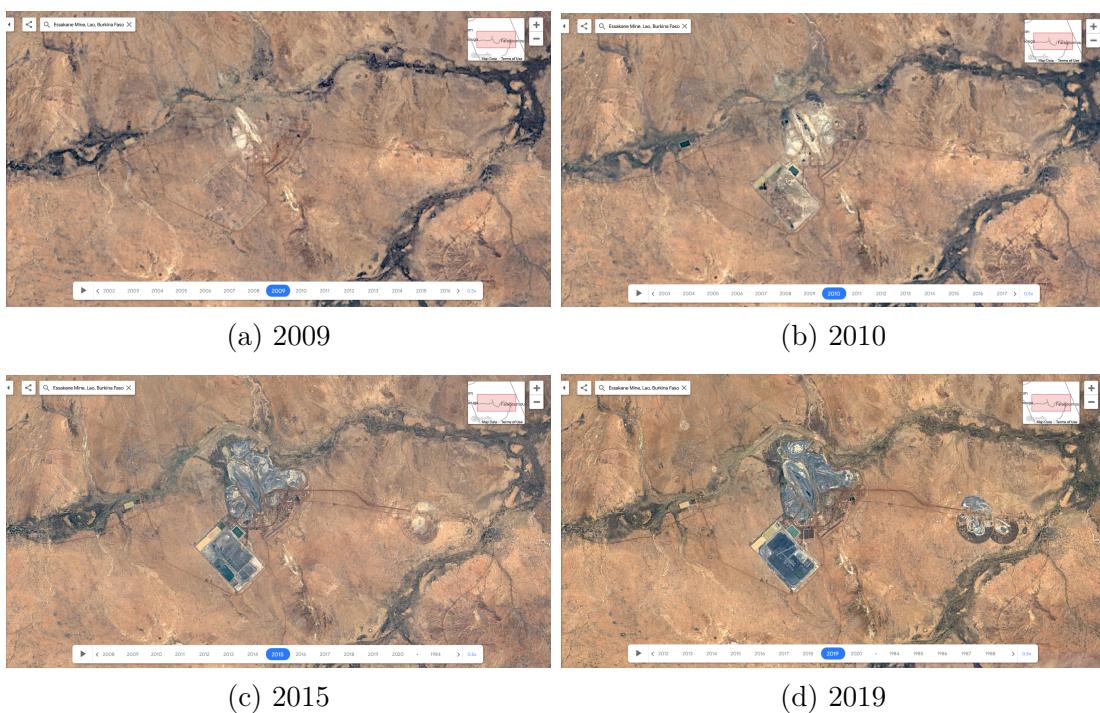
Figure 26 shows the satellite image of a mine from our sample, the Essakane mine in 2019, and Figure 27 shows the different stages of expansion and construction of the mine. Essakane is the most productive gold mine and the second largest in Burkina-Faso, still in activity. It is an open-pit gold mine that extends over a 100 km^2 area. It is located in the North-East of Burkina-Faso in the Oudalan province, near the Nigerian and Malian borders, and is hydrologically found in the sub-basin of Gorouol and Feildegasse rivers. It is exploited by the Burkinabé society Iamgold Essakane and belongs to the Canadian investor IamgoldInc (International African Mining Gold Corporation), who obtained the project in 2007. The installation in 2009 of the mine has forcibly displaced five villages, and 16,000 people with no choice, and the promised compensation to the communities for the displacement cost, loss of pastures and common forests have not been fulfilled [Atlas des Conflits pour la Justice Environnementale, 2022]. Mining at Essakane has been shown to have negative impacts on the environment and the health of the local population, both indirectly and accidentally.

In November 2015 Drechsel et al. [2018] has run qualitative interviews among the inhabitants of local communities of six active mining zones in Burkina-Faso, including

the Essakane zone. If the local population admits the benefits of the construction of a primary school and educational establishment, of a health center, of roads and electricity, the interviewed people do not find it sufficient to outweigh the negative aspects. The mine does not display formal employment to the local population, not educated enough to undertake the required skilled-work. On the contrary, due to the loss of agricultural land and the prohibition to practice gold panning, the local population fell into unemployment and poverty, as a farmer from the Essakane area explains: “*Before the mine arrival, we had better lives, we had animals, we were rich*” [Drechsel et al., 2018]. In 2010, the tailing storage facility of the mine collapsed, which caused the death of the surrounding livestock poisoned by chemicals used in the mine, and created tensions between the local population and the operator. In 2011 a truck carrying two containers of cyanide fell into the water source of Djibo’s dam and led to the death of all the fish in the dam. Tensions regarding water scarcity exist as well, the mine being water-consuming and reinforcing the vulnerability of the local population to droughts. Even if the regional government had prohibited the mine to use the village water, the national government overruled the decision, and the operator directly uses the water originally intended for the village. This led to the protestation of the local population around the mine in 2011, with no success. Finally, miners had major impacts on soil degradation, due to the construction of mining infrastructures, the multiplication of satellite pits and abandoned sterile holes devoid of gold [Porgo and Gokyay, 2016].

Environmental pollution has also degraded living conditions around the Essakane mine. Porgo and Gokyay [2016] use water sampling and digital calipers to capture water pollution and particle measurement in the Essakane zone and survey the surrounding population to understand the related health conditions. They find high levels of particles at the Essakane site center due to transportation and mining activity, such as the work of perforation, blasting, loading, transportation of ore, crushing, grinding, and energy production based on hydrocarbons. This air pollution mainly concerns mine workers, who develop acute respiratory infections (ARI), uncurable lung diseases caused by prolonged and severe inhalation of fine particles. The drilled well water samples display abnormally high concentrations of arsenic (higher than the WHO standard), which comes from the intensive use of acids (low pH) and the liberation of trace metals. The surrounding population presents diarrhea (13%) and affections of the skin and wounds (11%), reported to be caused by lack of hygiene, and use of drugs and chemicals. The main important health impact associated with the mine is the increase of malaria (20%), as stagnant water from mining dams attracts infected mosquitoes.

Figure 27: Expansion of the Essakane Mine, 2005-2019.



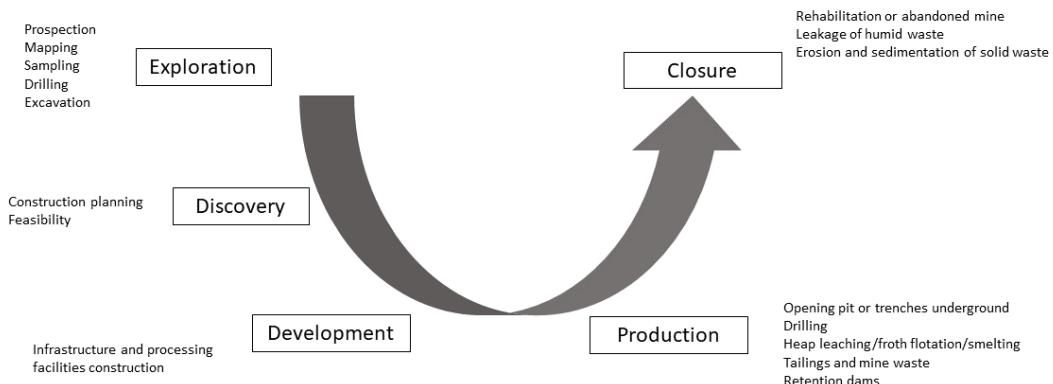
Notes: The six satellite images represent the expansion of the Essakane plant, in Burkina Faso. Retention dams can be seen.

Sources: Google Earth engine Timelapse.

C.2 Mine life cycles and types

Figure 28 gives the main stages of an industrial mining project, from the exploration phase to the start of the production and the closure of the mine. If it is hard to give the average length of each phase, yet the average mine lifetime is 16 years from the start of production to its closure (Figure 11). Figure 29 gives the time evolution of international prices for all commodities used in our main sample.

Figure 28: Industrial mine's life cycle

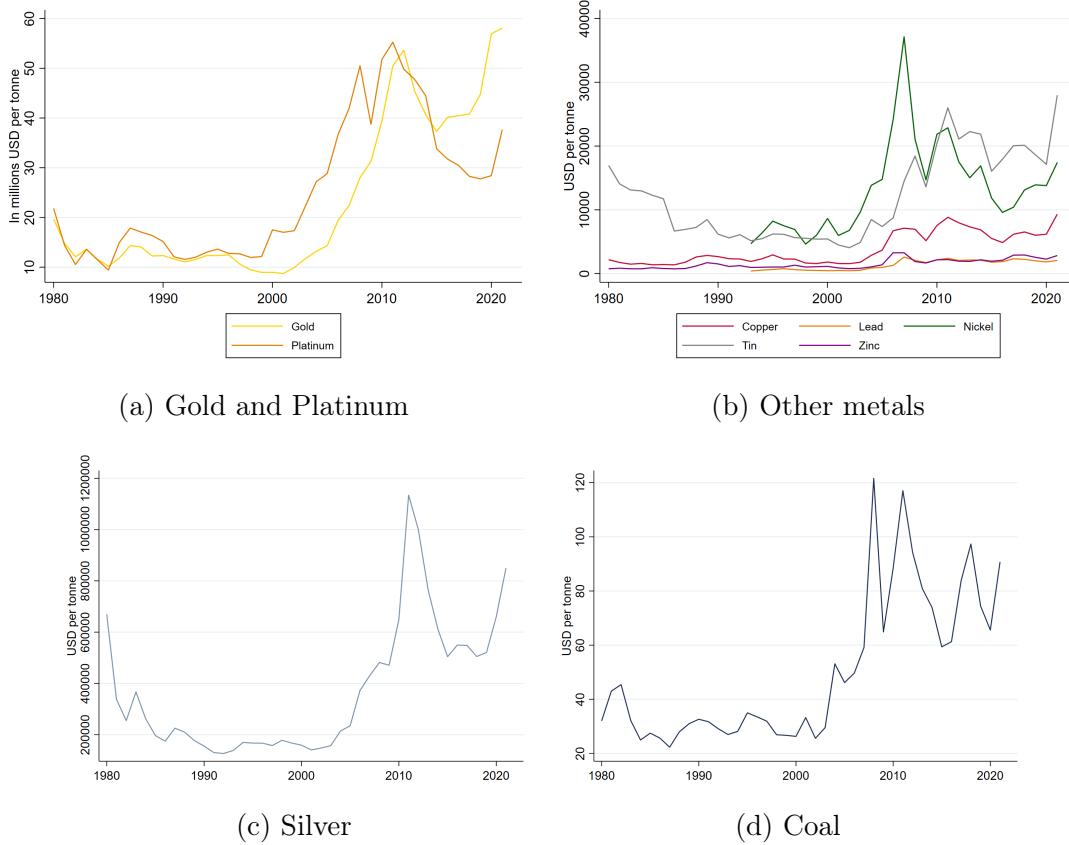


Notes: The figure schematizes the main stages of an industrial mining project .

Sources: Authors' elaboration, largely inspired by Coelho, Teixeira and Goncalves (2011)

Table 26 gives the chemical properties of each metal, including their main chemical compounds (Column 1), their density (Column 2), and displays their share in the main estimation sample (in terms of the number of mines Column (3) and Total Individual Sample (Column (4)). Heavy metals are defined according to their density as being greater than $5gcm^{-3}$ [Briffa et al., 2020]. If small amounts of heavy metals can be mandatory, a high and abnormal concentration of heavy metals may cause health issues due to chronic toxicity. Heavy metals released in mining activity are toxic elements that degrade the environment and human biology. This is the case as well for heavy metals released during the mining and burning of coal, which is linked to toxic heavy metals such as lead, mercury, arsenic, and nickel [Global Energy Monitor Wiki, 2021]. This is the reason why the main regression analysis includes heavy metals and coal mines, to capture the negative

Figure 29: Time evolution of international commodity prices



Notes: These Figures plot the evolution of metal prices from 1980 to 2020.

Sources: Authors' elaboration from SNL data and World Bank pink sheet data.

externalities linked to the most toxic mines.

Table 26: Metals, chemical properties and sample distribution

Metals	Main chemical compounds (1)	density (gcm^{-3}) (2)	Nb. Mines (3)	Total Individual Sample (%) (4)
Heavy Metals				
Gold	Gold	19.3	581	41.88
Copper	Copper	8.96	89	5.03
Iron ore	Iron	7.87	54	8.72
U308	Uranium	8.39	36	1.60
Nickel	Nickel	8.9	25	5.06
Platinum	Platinum	21.45	21	0.43
Zinc	Zinc	7.14	19	2.46
Chromite	Iron Chromium	[4.5,5.09]	16	0.57
Ilmenite	titanium	4.6	14	3.67
Lanthanides	Lanthane(57) Lutecium(71)	[6.1,9.8]	13	1.95
Manganese	Manganese	7.21	12	0.62
Tin	Tin	[5.7;7.26]	10	4.87
Cobalt	Cobalt	8.9	7	0.56
Tungsten	Tugsten	19.25	6	1.06
Tantalum	Tantalum	16.69	5	0.15
Vanadium	Vanadium	6.12	4	0.04
Niobium	Niobium	8.57	3	0.39
Heavy Mineral Sands	Zirconium Titanium Tungsten Thorium	[4.5,17.6]	3	0.16
Silver	Silver	10.49	1	0.00
Lead	Lead	11.29	1	0.06
Non-Heavy Metals				
Diamonds	Carbon	3.5	115	11.73
Coal	Carbon Mercury? Arsenic?	1.35	55	2.19
Bauxite	Aluminium	2.79	23	1.94
Graphite	Carbon	2.26	21	0.82
Phosphate	Phosphate	1.83	14	2.78
Lithium	Lithium	0.53	14	0.80
Rutile	titanium	4.23	2	0.29
Potash(Salt)	Potassium	0.89	1	0.17

Notes: This table gives for each metals the main chemical compounds (Column (1)) and their density (Column (2)). Columns (3) gives the number of mines within 100 km of a DHS cluster for which the metal is the main primary commodity, and Column (4) the percentage of children under 5 associated to these mines.

D Empirical Strategy

D.1 Descriptive Statistics

Table 27 gives the balance table for some household and mother characteristics. This is a descriptive table that accounts neither for controls nor fixed effects. Table 28 gives the effect of being downstream of a mine for the same variables. We observe no statistical difference in terms of access to piped water, and electricity, the age, and years of education of the mother. However, Table 28 shows that the proportion of urban households increases by 13 p.p once a mine has opened in downstream areas, compared to upstream areas. The proportion of mothers that are migrants also increases by 8 p.p.. These results suggest that the in-migrants coming after the mine opening, seeking jobs, for instance, settle down in downstream areas that become more urban. It suggests that the miner villages are located downstream of the mine. This gives the necessity to control for migration and verify that this is not driving our results (cf Section 6.2).

We plot in Figure 30 the distribution of mines opened within 100 km upstream or within the 3 closest sub-basins downstream during a child's birthyear, so as to see which countries gather the highest number of industrial mining activity in the vicinity of surveyed households over 1986-2018. Ghana, Zimbabwe, Tanzania, Zambia, Guinea and Sierra Leone have the highest density of open mines nearby DHS clusters, while Benin, Burundi, Cameroon, Lesotho and Niger have the lowest number of open mining sites. This figure also represents the variation in the number of mines that opened between the first and last year of surveys for each country. We can thus grasp the context of change in industrial mining activity over our period of interest. Ghana, Tanzania, Guinea, Mali, and Burkina-Faso witnessed the highest number of mine openings between 1986 and 2018.

Figures 31, 32 and 33 give the spatial variation of the infant mortality outcomes and the mine openings for the sample restricted to our main analysis.

Table 27: Balance Table - Double Difference with Topographic Treatment - Descriptive Statistics

Before Mine Opening					After Mine Opening					Within Up.	Within Dwn.	Within	
Upstream		Downstream		Diff	Upstream		Downstream		Diff	(7-2) (p.v)	(9-4) (p.v)	(12-11) (p.v)	
N	Mean /(SD)	N	Mean /(SD)	(4-2) (p.v)	N	Mean /(SD)	N	Mean /(SD)	(9-7) (p.v)	(11)	(12)	(13)	
(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	
Household Characteristics													
% Urban Household													
All	29,399	0.33	9,835	0.175	-0.155	16,285	0.385	6,174	0.291	-0.094	0.055	0.116	0.061
		(0.47)		(0.38)	(0)		(0.487)		(0.454)	(0)	(0)	(0)	(0)
Mines	244		237			190		193					
Has piped water													
All	29,399	0.307	9,835	0.193	-0.115	16,285	0.365	6,174	0.27	-0.095	0.057	0.077	0.019
		(0.461)		(0.395)	(0)		(0.481)		(0.444)	(0)	(0)	(0)	(0.002)
Has electricity													
All	29,399	0.211	9,835	0.14	-0.071	16,285	0.356	6,174	0.226	-0.13	0.145	0.086	-0.059
		(0.408)		(0.347)	(0)		(0.479)		(0.418)	(0)	(0)	(0)	(0)
Mother Characteristics													
Age													
All	29,399	29.106	9,835	29.187	0.081	16,285	28.779	6,174	28.818	0.039	-0.328	-0.369	-0.042
		(7.065)		(7.039)	(0.325)		(6.847)		(6.986)	(0.707)	(0)	(0.001)	(0.764)
Years of Education													
All	29,399	2.406	9,835	2.91	0.504	16,285	4.297	6,174	4.851	0.554	1.891	1.941	0.05
		(3.6)		(3.741)	(0)		(4.417)		(4.2)	(0)	(0)	(0)	(0.001)
% Migrant													
All	18,509	0.615	6,593	0.578	-0.037	9,773	0.597	3,962	0.589	-0.007	-0.019	0.011	0.029
		(0.487)		(0.494)	(0)		(0.491)		(0.492)	(0.421)	(0.002)	(0.278)	(0.094)

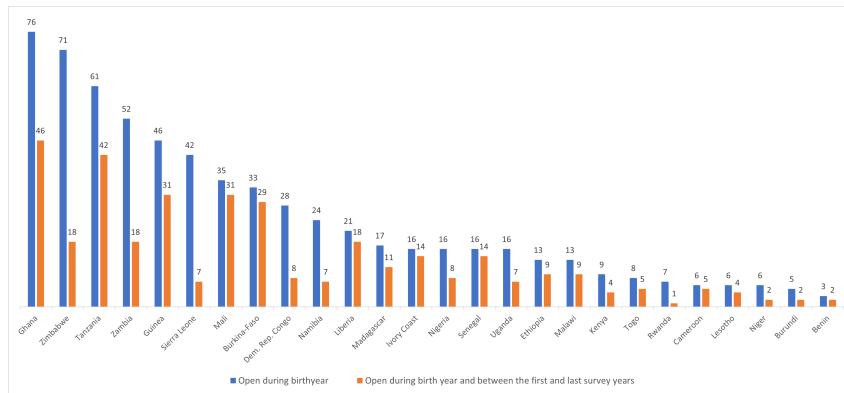
Notes: Standard errors and p-values in parentheses.

Table 28: Average effects of mine opening on control variables

	Household's characteristics			Mother characteristics		
	% urban households	Has piped water	Has electricity	Age	Yers of education	% migrant
	(1)	(2)	(3)	(4)	(5)	(6)
Downstream×Open	0.131*** [0.0422]	0.0205 [0.0277]	-0.00100 [0.0200]	0.0426 [0.162]	-0.0121 [0.144]	0.0881*** [0.0314]
Downstream	-0.0142 [0.0307]	-0.0411* [0.0239]	0.00433 [0.0143]	-0.0835 [0.132]	-0.119 [0.108]	-0.0327 [0.0272]
Open	-0.0455 [0.0293]	-0.00637 [0.0209]	-0.0104 [0.0181]	-0.136 [0.136]	-0.126 [0.112]	-0.0432* [0.0233]
Controls	Yes	Yes	Yes	Yes	Yes	Yes
Nb open mines	Yes	Yes	Yes	Yes	Yes	Yes
Birthmonth FE	Yes	Yes	Yes	Yes	Yes	Yes
Country-birthyear FE	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes	Yes	Yes
Mine SB-birthyear trend	Yes	Yes	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes	Yes	Yes
N	61690	61690	61179	61690	61690	38834
R2	0.608	0.489	0.547	0.681	0.463	0.185

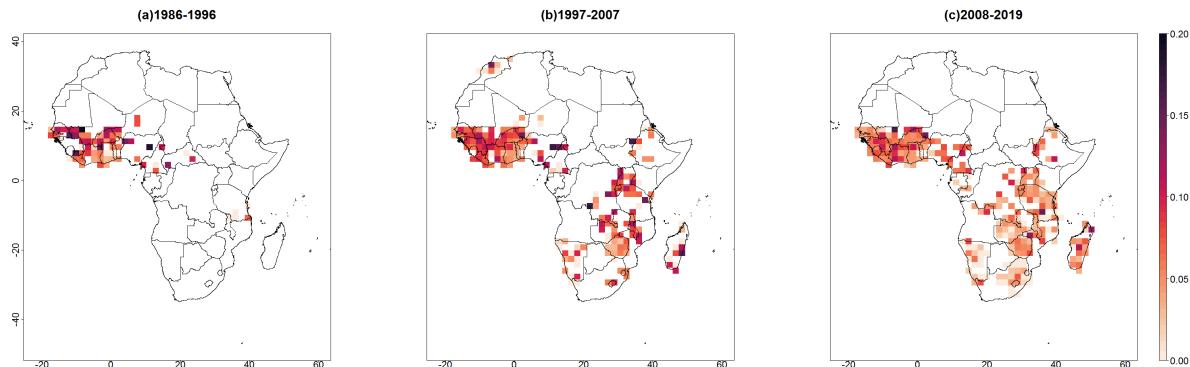
Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The variables Downstream and Opened are dummies which indicate whether the individual lives in a village downstream of at least one mining site, and whether the site opened before the birth year of the child. Each village DHS is paired to only one mining site, so that each individual appears only once in the regression. Other variables are control variables. The sample focuses on heavy metal mines.

Figure 30: Number of open mines during the birth year and between first and last wave



Notes: The figure represents the number of mines that were opened during the birth year of children located within our topographic treatment sample by country, and the number of mines that were opened during the birth year of children located within our topographic treatment sample and which opened between the first and last year of survey for each country.
Sources: Authors' elaboration on SNL and DHS data.

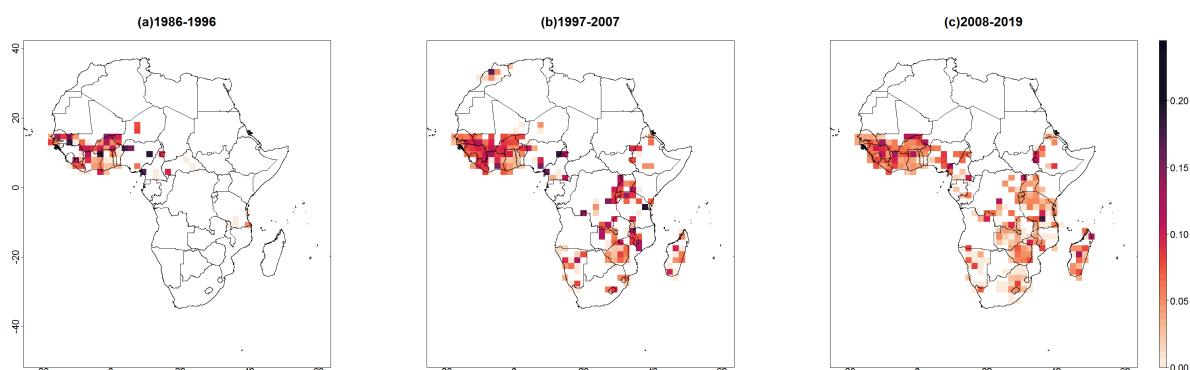
Figure 31: Spatial variation of 12-month mortality rates per period - Restricted Sample



Notes: The figures represent the means of 12-month mortality rates averaged at the grid level over (a) 1986-1996, (b) 1997-2008, and (c) 2008-2019, for the sample of the main analysis. The mortality rates are estimated without the children that did not reach 12 months at the time of the survey.

Sources: Authors' elaboration on DHS data.

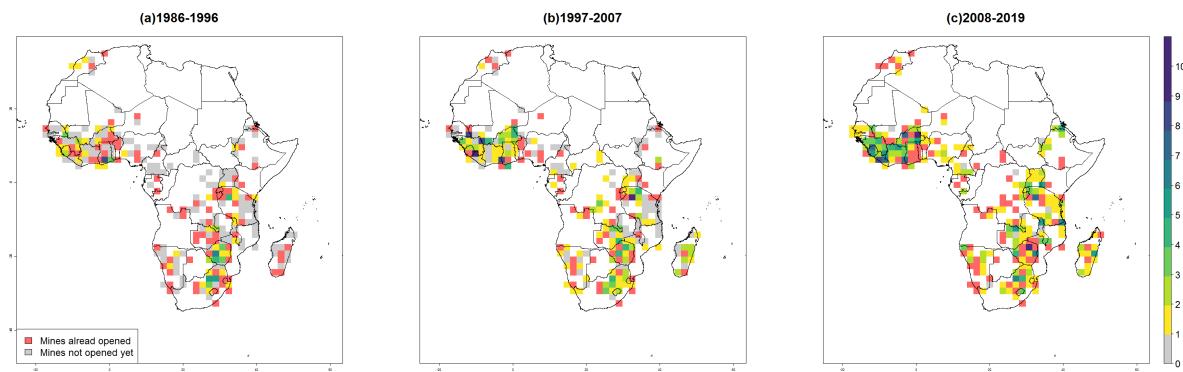
Figure 32: Spatial variation of 24 months mortality rates per period - Restricted Sample



Notes: The figures represent the means of 24-month mortality rates averaged at the grid level over (a) 1986-1996, (b) 1997-2008, and (c) 2008-2019, for the sample from the main analysis. The mortality rates are estimated without the children that did not reach 24 months at the time of the survey.

Sources: Authors' elaboration on DHS data.

Figure 33: Spatial variation of mine opening per period - Restricted Sample



Notes: The figures represent the number of mines that opened during the periods over the grid area (160 km on average). A red grid cell represents an area where no mine opened over the period, but where at least one mine has opened before the period. A grey cell represents an area where no mine opened over the period, but where at least one mine will open in the future.

Sources: Authors' elaboration on SNL data.

E Heterogeneity

Table 29: Effects of industrial mining activity, across sub-regions

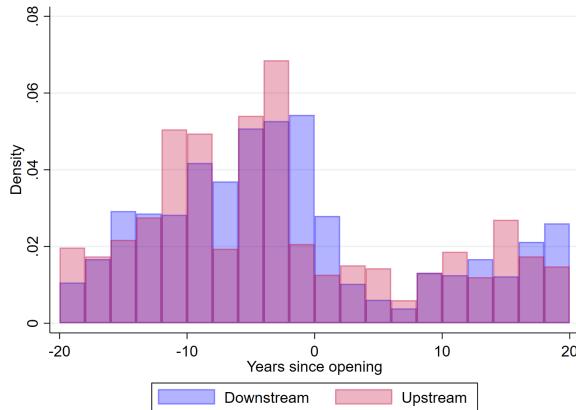
	Mortality under 24 months		
	Western Africa	Eastern Africa	Central and Southern Africa
	(1)	(2)	(3)
Downstream×Open	0.0443** [0.0176]	0.0292* [0.0154]	-0.0304 [0.0225]
Downstream	-0.0153 [0.00946]	-0.0369*** [0.0116]	0.0538** [0.0227]
Open	0.00523 [0.0133]	-0.0137 [0.0157]	0.0120 [0.00857]
Controls	Yes	Yes	Yes
Birthmonth FE	Yes	Yes	Yes
Country-birthyear FE	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes
Mine SB-birthyear trend	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes
N	21,006	13,484	20,014
R2	0.0521	0.0457	0.0238
Outcome Mean	0.0981	0.0712	0.0836

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

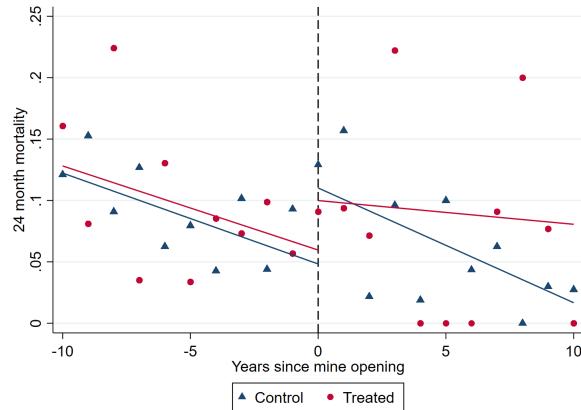
F Dynamic effects - pre trends and event study

In this section, we display the appendix figures associated with Section 8. It gives the same analysis as in the main Section restricted to the balanced sample. Figure 34 plots the parallel trends, while Figure 35 plots the event study for the balanced sample.

Figure 34: Linear trends of 24 month mortality - Balanced Sample



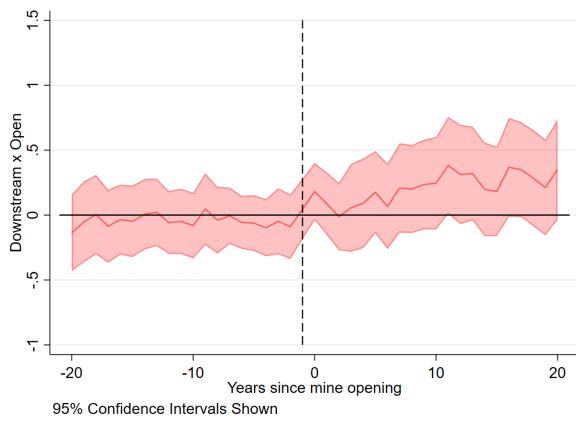
(a) Balanced sample



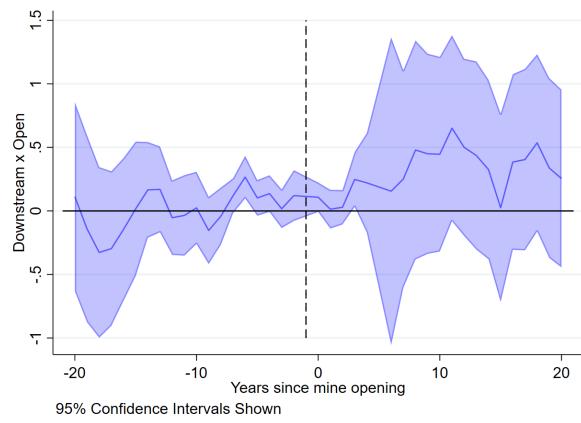
(b) Balanced sample

Notes: Figure (a) gives the distribution of the number of observations per opening year. Figure (b) plots the trends of the 24-month mortality rates according to the year of opening. The figures are made for the balanced sample and include neither control variables nor fixed effects.

Figure 35: Event study - dynamic effect of mine opening on under 24 months mortality - Balanced Sample



(a) Balanced sample - upstream



(b) Balanced sample - downstream

Notes: Figure (a) plots the event study for the upstream villages, while Figure (b) plots the event study for the downstream villages for the balanced sample. Controls and fixed effects are the same as in the main analysis (column (4) Table 2)

F.1 Sensitivity analysis

Table 30 shows that our result is stable when controlling for a dummy indicating whether the mine opening year has been found by hand or was given directly in the SNL database (column (2)).

Table 30: Effects of industrial mining opening, controlling for handwork.

Specification	Outcome Main result (1)	Mortality under 24 months		
		Adding control (2)	SNL database (3)	Handwork (4)
Downstream × Open	0.0218** [0.0108]	0.0218** [0.0108]	0.0222 [0.0351]	0.0344** [0.0138]
Dummy handwork		0.0254 [0.0335]		
Downstream	-0.0211*** [0.00739]	-0.0212*** [0.00739]	-0.0167 [0.0246]	-0.0316*** [0.00844]
Open	-0.00496 [0.0101]	-0.00489 [0.0101]	-0.0194 [0.0476]	-0.00973 [0.0126]
Controls	Yes	Yes	Yes	Yes
Birthmonth FE	Yes	Yes	Yes	Yes
Country-birthyear FE	Yes	Yes	Yes	Yes
Mine SB FE	Yes	Yes	Yes	Yes
Mine SB-birthyear trend	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes
N	35,638	35,638	6,702	22,017
R2	0.0511	0.0511	0.0615	0.0641
Outcome mean	0.0873	0.0873	0.0727	0.0954

Notes: Standard errors clustered at the DHS village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Columns (1) and (2) rely on the same sample and controls as Table 2 Column 2. Column (2) controls for the hand-work, while Columns (3) and (4) split the samples.

Figure 36 displays the DiD estimators for different regression with restricted samples, meaning while dropping each metal one by one, using the sample for the 24-month mortality rates, and the heavy metals and coal mine sample. This suggests that our main results are not driven by a specific metal. Accordingly, Figure 37, plots the DiD estimators while dropping countries one by one and show that our analysis is not driven by a particular country.

Figure 36: Regression results when dropping commodities one by one

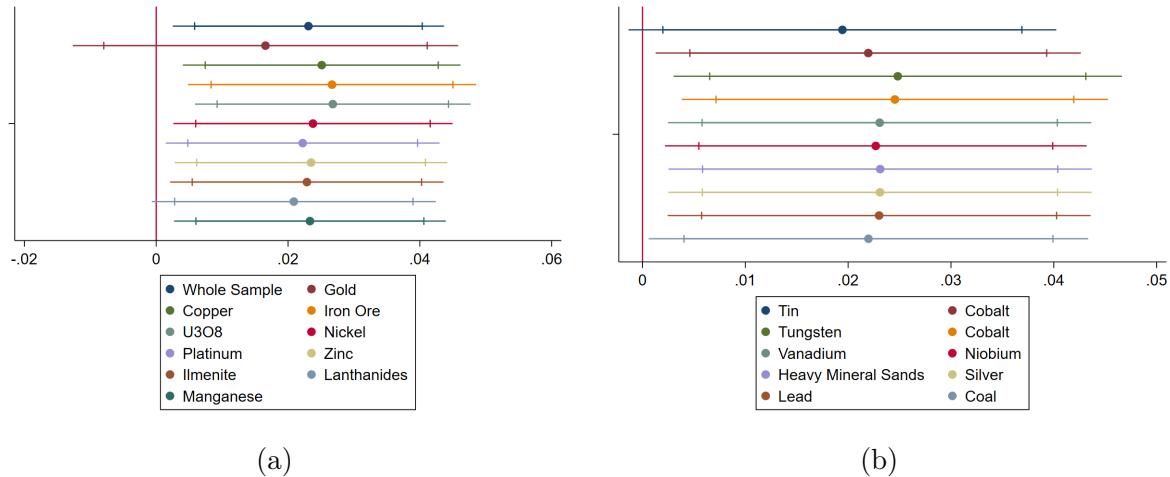
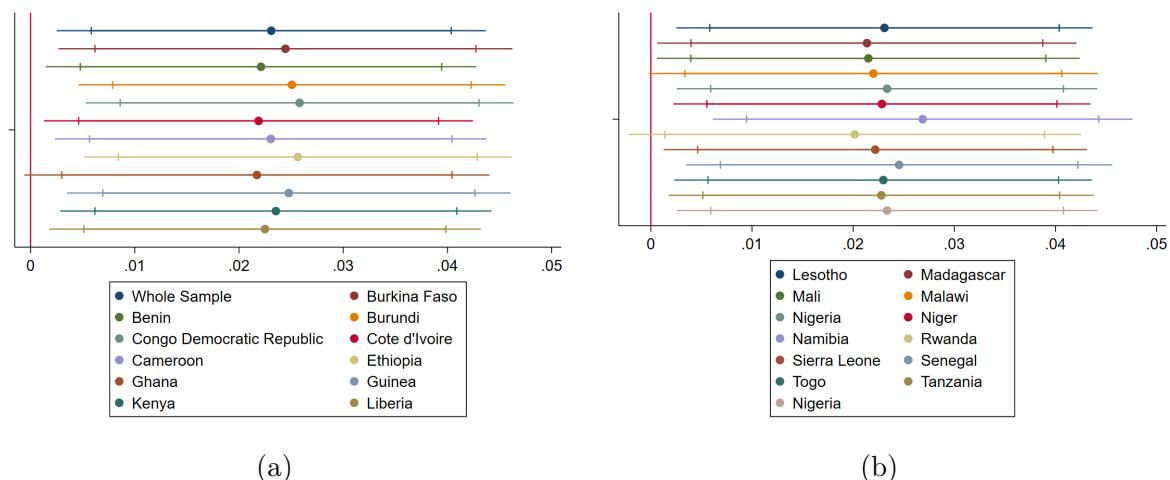


Figure 37: Regression results when dropping countries one by one



Sources: Authors' elaboration on DHS and SNL data.

G Geographic Treatment

In this section, we propose to replicate the empirical strategy of Benshaul-Tolonen [2018], who finds that a mine opening is associated with a 5.5 p.p decrease in the 12-month mortality. The identification strategy relies on a treatment based on proximity, comparing individuals living nearby to those living further from an industrial mine. In this estimation, geographical proximity is used as a proxy exposure to industrial mining activity, including both positive and negative externalities, such as exposure to mining pollution. The identification strategy relies on a DiD strategy. It compares within each district, the infant mortality in areas within 10 km of a mine deposit (treatment group) to infant mortality in DHS clusters further away from a mine deposit (10-100km, control group), before and after the opening of the mine deposit. As the strategy is a two-way fixed-effects, including a district-fixed effect, the comparison is made within each district. The identification can be formally written as:

$$\begin{aligned} Death_{i,v,c,m,SB} = & \alpha_0 + \alpha_1 Opened_{birthyear,i,v} + \alpha_2 MineDeposit_{[0;10km]v} \\ & + \alpha_3 Opened_{birthyear,i,v} \times MineDeposit_{[0;10km]v} + \alpha_4 X_i \quad (2) \\ & \gamma_d + \gamma_{d-bthtrend} + \gamma_{c,birthyear} + \epsilon_v \end{aligned}$$

With $Death_{i,v,c,district}$ a dummy equals to one if child i from DHS village v (within district d) of country c , has reached the n^{th} month and has died (n being 12 for the 12-month mortality, 24 and so on). $Opened_{birthyear,i,v}$ is a dummy equal to 1 if at least one mine located within 10 km for the treatment group, or within 100 km for the control group, has opened before child i 's year of birth (this cohort comparison can be considered here as a source of triple difference). $MineDeposit_{[0;10km]v}$ is a dummy of proximity (1 if village DHS v is within 10 km of a mine deposit, 0 if it is within 10-100km), X_i a vector of child/mother level controls (mother's age and age square, years of education, urban status). Finally, γ_d is a district fixed effect, $\gamma_{d-bthtrend}$ a district birthyear linear trend, and $\gamma_{c,birthyear}$ a country-birthyear fixed effect. Please note that the matching of DHS clusters to mines relies on the same strategy as in Benshaul-Tolonen [2018], and assigns a DHS cluster to the closest mine (without consideration of its opening status). Thanks to this pairing, if a DHS cluster is both in the treatment and control groups of two different mines (i.e within 10km of mine A and within 10-100km from Mine B), we assign it mechanically to the treatment group (so linked to mine A). This creates bias explained in Section 2.3, which explains the choice for a district fixed effects and reduces the noise linked to DHS random displacements.

Firstly, we give our estimators from the exact replication of Benshaul-Tolonen [2018] results, using our own calculation, and find similar impacts (Tables 31 and 32). Second, we propose the replication of the results using our extended sample, including more countries, DHS waves, types of mines, and mines hand-checked, and show the results from Benshaul-Tolonen [2018] are mainly determined by the choice of countries.

G.1 Exact replication of Benshaul-Tolonen [2018]

The geographic treatment proxies exposure to mining activity using the distance to the site and follows partly the analysis from Benshaul-Tolonen [2018], and finds contradictory impacts on infant mortality. To understand better how our results can be compared to the literature, we propose in this section a replication exercise of the main result from Benshaul-Tolonen [2018]²⁵.

For this replication analysis, we used the same mines and DHS survey rounds as Benshaul-Tolonen [2018]. Please note that we have few differences in terms of the whole sample, as Benshaul-Tolonen [2018] counts 37,365 children *vs* 41,902 for us, that might be explained by the way we calculated the 100km buffer distance ²⁷. A main difference between our paper and Benshaul-Tolonen [2018] is the independent variable, as we use as a shock the opening of the industrial mine whereas Benshaul-Tolonen [2018] uses the activity status based on production data given by the SNL product. This accounts for interim years, between the opening and final closing of the mine, where the production has been on hold. In this section, we replicate this exact same variable.

Table 31 displays the replication of the main results from Benshaul-Tolonen [2018] Table 2. We find that a mine opening within 10 kilometers is associated with a 4.7 percentage point decrease in infant mortality rates, while Benshaul-Tolonen [2018] found 5.5 p.p. Our results is slightly less significant than from Benshaul-Tolonen [2018], and we identify a different impact according to gender, with a significant reduction of girl mortality rates of 7 p.p *vs* a non-significant reduction for boys, which differs from the

²⁵Please note that a first difference between the two analyses is the sample, as Benshaul-Tolonen [2018] uses 43 gold mines that match with 31 DHS surveys from nine countries (Burkina Faso, Cote D'Ivoire, Ethiopia, Ghana, Guinea, Mali, Senegal, Tanzania, and DRC ²⁶). However, when pairing the DHS cluster to the same industrial mining sites from Benshaul-Tolonen [2018], no DHS from DRC remained. In the end, the analysis is only on the 8 first countries, in accordance with Figure A6 from Appendix of Benshaul-Tolonen [2018]), for a whole sample of 1-year-old children of 48,151.

²⁷In the replication codes of Benshaul-Tolonen [2018], one can observe that the distance has been determined using the Stata command *nearstat [...] dband(0,25)* which relies on different projections (not specified) as ours from *R* libraries, explaining the small sample differences

Table 31: Replication Benshaul-Tolonen [2018] Main Results

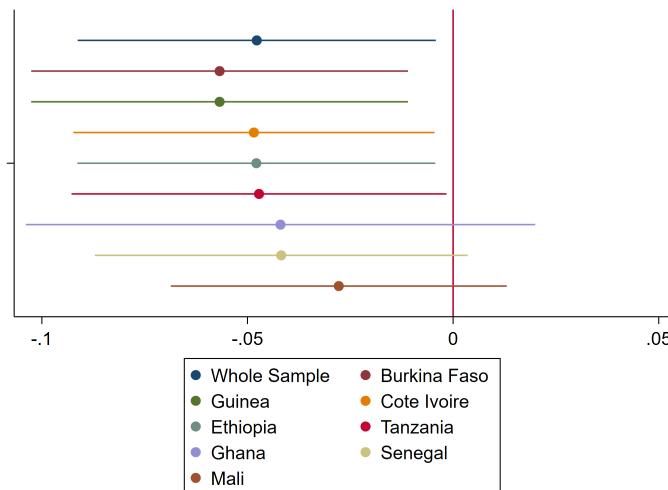
Dependent variable	Infant mortality first 12 months			
	Sample :	Children (1)	Children drop spillover (2)	Boys (3)
Industrial \times mine deposit (at birth)	-0.0472** [0.0230]	-0.0474* [0.0260]	-0.0289 [0.0320]	-0.0781*** [0.0301]
Mine deposit [0;10km]	0.0392** [0.0169]	0.0546*** [0.0195]	0.0517** [0.0229]	0.0561** [0.0231]
Mother's age	-0.0145*** [0.00190]	-0.0154*** [0.00210]	-0.0155*** [0.00274]	-0.0152*** [0.00297]
Mothers's age \times Mother's age	0.000222*** [0.0000302]	0.000236*** [0.0000335]	0.000223*** [0.0000435]	0.000245*** [0.0000475]
Years edu.	-0.00214*** [0.000489]	-0.00230*** [0.000547]	-0.00272*** [0.000827]	-0.00184** [0.000760]
Urban _h	-0.0125*** [0.00428]	-0.0120** [0.00480]	-0.00710 [0.00687]	-0.0183*** [0.00659]
Birth-month FE	Yes	Yes	Yes	Yes
Country birth year FE	Yes	Yes	Yes	Yes
District FE	Yes	Yes	Yes	Yes
District BirthYear trend	Yes	Yes	Yes	Yes
Drop10-30 km away	No	Yes	Yes	Yes
Drop investment phase	No	Yes	Yes	Yes
Mean of outcome	0.102	0.104	0.110	0.099
Mean(treatment, pre-treatment)	0.154	0.163	0.173	0.153
Observations	41902	34228	17534	16694

Notes: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Standard errors clustered at a DHS cluster level. The variables Mine deposit [0;10km] and Industrial \times mine deposit (at birth) are a replication from Benshaul-Tolonen [2018] and indicate whether the child is born within 10km of at least one industrial mining site and whether this site was active at the time of the birth. All regressions control for mother's age, age square, mother's education and whether the household is urban, for district, birth month and country-birth year. The main outcome is infant mortality in the 12 months since birth. Columns 2-5 drop the two years preceding the opening year, defined as investment phase in Benshaul-Tolonen [2018] and the individuals living within 10-30km of the closest industrial mine. Mean (treatment, pre-treatment) is the sample for the treatment group before the mine were active. dummies which indicate whether the individual lives in a village within at least one mining site, and whether the site opened before the birth year of the child. Each village DHS is paired to the closest opened mining site, so that each individual appears only once in the regression. Other variables are control variables.

previous study. To follow Benshaul-Tolonen [2018] example, we excluded in Columns 2-5 from Table 31 individuals born within 10-30 kilometers of the closest industrial mining site and those born the two years before the opening of a mine, which is a proxy for the investment phase according to Benshaul-Tolonen [2018].

Please note that in accordance with the descriptive statistics from Benshaul-Tolonen [2018] we have in the sample a very high mean of 12-month mortality rates (from 10 to 17 % according to the groups). These are relatively high numbers, that do not match with World Bank data. This is because Benshaul-Tolonen [2018] drops all the individuals who are still alive but did not reach the age of 12 months yet to measure the mortality, in order to avoid growing mechanically the mortality rates of these cohorts²⁸. For replication purposes, we propose to keep this variable and correct this in Table 32, where we observe average mortality rates around 7%. Figure 38 replicates the Figure A6 from Benshaul-Tolonen [2018], which shows the coefficient estimates of the main regression for *industrial* \times *mine deposit* on infant mortality, each regression excluding the sample from one country as indicated by the country name. The Figure 38 shows that results are highly sensitive to the presence of Mali, Senegal, and Ghana in the sample (whereas they do not consist for the majority of the sample (5847, 1098 and 5595 respectively).

Figure 38: Regression results when dropping one country at a time



Sources: Authors' elaboration on DHS and SNL data.

²⁸We can read in the codes that if the living individuals were dropped, the children that died before their 12 months from these specific cohorts were not dropped: mechanically, the mortality rates for all the years preceding the survey rounds are 100 %, which explain the high mean of outcomes.

Table 32: Replication Benshaul-Tolonen [2018] Main Results

Dependent variable	Infant mortality first 12 months corrected			
	Sample :	Children (1)	Children drop spillover (2)	Boys (3)
Industrial \times mine deposit (at birth)	-0.0494** [0.0229]	-0.0471* [0.0244]	-0.0439 [0.0317]	-0.0631** [0.0298]
Mine deposit [0;10km]	0.0394** [0.0179]	0.0587*** [0.0198]	0.0682*** [0.0255]	0.0513** [0.0235]
Mother's age	-0.0118*** [0.00175]	-0.0123*** [0.00196]	-0.0120*** [0.00256]	-0.0124*** [0.00283]
Mothers's age \times Mother's age	0.000182*** [0.0000279]	0.000189*** [0.0000312]	0.000172*** [0.0000405]	0.000203*** [0.0000452]
Years edu.	-0.00143*** [0.000455]	-0.00152*** [0.000510]	-0.00204*** [0.000772]	-0.000803 [0.000715]
Urban _h	-0.0106*** [0.00384]	-0.0113*** [0.00436]	-0.00501 [0.00661]	-0.0196*** [0.00600]
Birth-month FE	Yes	Yes	Yes	Yes
Country birth year FE	Yes	Yes	Yes	Yes
District FE	Yes	Yes	Yes	Yes
District BirthYear trend	Yes	Yes	Yes	Yes
Drop10-30 km away	No	Yes	Yes	Yes
Drop investment phase	No	Yes	Yes	Yes
Mean of outcome	0.079	0.080	0.083	0.077
Mean(treatment, pre-treatment)	0.109	0.118	0.120	0.115
Observations	40386	32873	16823	16050

Notes: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Standard errors clustered at a DHS cluster level. The variables Mine deposit [0;10km] and Industrial \times mine deposit (at birth) are a replication from Benshaul-Tolonen [2018] and indicate whether the child is born within 10km of at least one industrial mining site and whether this site was active at the time of the birth. All regressions control for mother's age, age square, mother's education and whether the household is urban, for district, birth month and country-birth year. The main outcome is infant mortality in the 12 months since birth. Columns 2-5 drop the two years preceding the opening year, defined as investment phase in Benshaul-Tolonen [2018] and the individuals living within 10-30km of the closest industrial mine. Mean (treatment, pre-treatment) is the sample for the treatment group before the mine were active. dummies which indicate whether the individual lives in a village within at least one mining site, and whether the site opened before the birth year of the child. Each village DHS is paired to the closest opened mining site, so that each individual appears only once in the regression. Other variables are control variables.

G.1.1 Replication using an extended sample

Table 33 and Table 34 display the results, replicating Benshaul-Tolonen [2018] estimation strategy, with our overall sample of mines and DHS surveys. Table 33 focuses on the 12-month mortality rates and shows that we find a significant reduction of infant mortality by 0.8 p.p only when controlling for migrants (column (2)). Columns (1) and (2) display the results for the whole sample, while columns (3) and (4) while dropping the spillovers effects (areas between [10-30]km and the two years before the mine opening, which represents the investment phase in Benshaul-Tolonen [2018]). Columns (5) and (6) replicate the analysis for the male sample while columns (7) and (8) for the girls.

Table 34 displays the result for the 12-month mortality rates (Columns (1)-(4)) and 24 months mortality rates (Columns (5)-(8)) and compares the estimators when not including the migrant control variable (Columns (1), (3) (5) and (7)), and when including it (Columns ((2),(4),(6) and (8))). We also display the estimators for the restricted sample of rural areas (Columns (3),(4), (7), and (8)). Again, we observe a significant reduction of 12 months mortality rates in Column (2), i.e for the overall sample while controlling for migrants, and find no results otherwise. This absence of results suggests that using proximity as a proxy for exposure to mining activity averages contradictory effects, including both positive and negative externalities, and shows the importance of our main estimation strategy which relies on topographic position.

Figure 39 plots the linear trends of the 12 and 24 months mortality rates for the geographic treatment, including our overall mine and DHS sample. We see that the linear trends assumption seems to be validated for the 24-month mortality, but not for the 12-month mortality rates.

Table 33: Geographic Treatment

	Infant mortality first 12 months							
	All	Drop	spillover	Boys		Girls		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Indus. × deposit	-0.00259 [0.00329]	-0.00823** [0.00418]	-0.00189 [0.00407]	-0.00575 [0.00537]	0.00250 [0.00570]	-0.00302 [0.00764]	-0.00513 [0.00522]	-0.00807 [0.00674]
Deposit	0.00130 [0.00252]	0.00374 [0.00317]	0.00103 [0.00392]	-0.000128 [0.00500]	0.00632 [0.00546]	0.0113 [0.00708]	-0.00366 [0.00513]	-0.0109* [0.00628]
Birth order	0.00389*** [0.000345]	0.00315*** [0.000428]	0.00360*** [0.000423]	0.00320*** [0.000518]	0.00349*** [0.000606]	0.00304*** [0.000742]	0.00382*** [0.000549]	0.00356*** [0.000671]
Mother's age	-0.0105*** [0.000541]	-0.0107*** [0.000668]	-0.0102*** [0.000669]	-0.0110*** [0.000824]	-0.0116*** [0.000953]	-0.0128*** [0.00119]	-0.00884*** [0.000903]	-0.00924*** [0.00111]
agesquare	0.000147*** [0.00000853]	0.000151*** [0.0000106]	0.000142*** [0.0000106]	0.000156*** [0.0000131]	0.000163*** [0.0000150]	0.000183*** [0.0000187]	0.000121*** [0.0000142]	0.000127*** [0.0000175]
Years edu.	-0.000877*** [0.000135]	-0.00103*** [0.000167]	-0.000874*** [0.000164]	-0.00101*** [0.000200]	-0.000881*** [0.000238]	-0.00103*** [0.000290]	-0.000873*** [0.000216]	-0.000968*** [0.000265]
Urban	-0.00610*** [0.00135]	-0.00725*** [0.00172]	-0.00708*** [0.00169]	-0.00906*** [0.00214]	-0.00825*** [0.00235]	-0.0111*** [0.00297]	-0.00563** [0.00227]	-0.00622** [0.00289]
migrant		0.00543*** [0.00120]		0.00509*** [0.00145]		0.00255 [0.00208]		0.00754*** [0.00196]
Constant	0.229*** [0.00826]	0.232*** [0.0101]	0.226*** [0.0103]	0.240*** [0.0126]	0.251*** [0.0146]	0.273*** [0.0181]	0.201*** [0.0138]	0.206*** [0.0169]
Birth-month FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Ctry-bthy FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
District FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Dist-bthy trend	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Drop10-30 km	No	No	Yes	Yes	No	No	No	No
Drop t-2	No	No	Yes	Yes	No	No	No	No
N	359219	243645	236573	165202	119860	83570	116696	81601

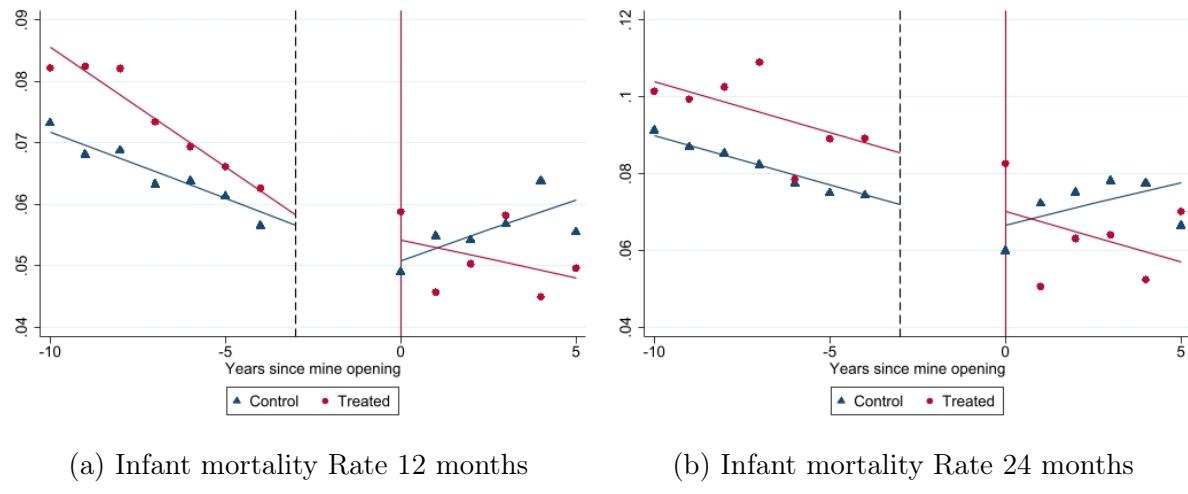
Notes: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. Standard errors clustered at a DHS cluster level. The variables Mine deposit [0;10km] and Industrial × mine deposit (at birth) are a replication from Benshaul-Tolonen [2018] and indicate whether the child is born within 10km of at least one industrial mining site and whether this site was active at the time of the birth. All regressions control for mother's age, age square, mother's education and whether the household is urban, for district, birth month and country-birth year. The main outcome is infant mortality in the 12 months since birth. dummies which indicate whether the individual lives in a village within at least one mining site, and whether the site opened before the birth year of the child. Each village DHS is paired to the closest opened mining site, so that each individual appears only once in the regression. Other variables are control variables.

Table 34: Effects of industrial mining activity on under 12, 24 mortality - Geographic Treatment - All Households

	All (1)	Death <12m			Death < 24m			
	All (1)	All (2)	Rural (3)	Rural (4)	All (5)	All (6)	Rural (7)	Rural (8)
indus. × deposit	-0.00259 [0.00329]	-0.00823** [0.00418]	-0.00259 [0.00329]	-0.00627 [0.00509]	0.000248 [0.00431]	-0.00264 [0.00535]	0.000248 [0.00431]	-0.00248 [0.00657]
Deposit	0.00130 [0.00252]	0.00374 [0.00317]	0.00130 [0.00252]	0.00313 [0.00368]	0.000627 [0.00321]	0.00121 [0.00411]	0.000627 [0.00321]	0.000859 [0.00477]
Indus.	0.00131 [0.00155]	0.00222 [0.00200]	0.00131 [0.00155]	0.00340 [0.00230]	0.00116 [0.00201]	0.00122 [0.00259]	0.00116 [0.00201]	0.00190 [0.00297]
Birth order	0.00389*** [0.000345]	0.00315*** [0.000428]	0.00389*** [0.000345]	0.00353*** [0.000500]	0.00512*** [0.000440]	0.00401*** [0.000549]	0.00512*** [0.000440]	0.00447*** [0.000642]
Mother's age	-0.0105*** [0.000541]	-0.0107*** [0.000668]	-0.0105*** [0.000541]	-0.0116*** [0.000787]	-0.0115*** [0.000704]	-0.0124*** [0.000873]	-0.0115*** [0.000704]	-0.0140*** [0.00103]
Age square	0.000147*** [0.00000853]	0.000151*** [0.0000106]	0.000147*** [0.00000853]	0.000161*** [0.0000122]	0.000151*** [0.0000110]	0.000167*** [0.0000136]	0.000151*** [0.0000110]	0.000187*** [0.0000159]
Years edu.	-0.000877*** [0.000135]	-0.00103*** [0.000167]	-0.000877*** [0.000135]	-0.000792*** [0.000219]	-0.00145*** [0.000173]	-0.00157*** [0.000215]	-0.00145*** [0.000173]	-0.00132*** [0.000283]
Urban	-0.00610*** [0.00135]	-0.00725*** [0.00172]	-0.00610*** [0.00135]		-0.00940*** [0.00175]	-0.00995*** [0.00222]	-0.00940*** [0.00175]	
migrant		0.00543*** [0.00120]		0.00514*** [0.00144]		0.00727*** [0.00155]		0.00630*** [0.00186]
Constant	0.229*** [0.00826]	0.232*** [0.0101]	0.229*** [0.00826]	0.247*** [0.0120]	0.273*** [0.0109]	0.286*** [0.0134]	0.273*** [0.0109]	0.315*** [0.0159]
Birthmonth FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Cty-Bthy FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mine FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Mine Bthy trend	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Commodity FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
N	359,219	243,645	359,219	179,155	265,735	179,729	265,735	132,398
R2	0.0195	0.0235	0.0195	0.0281	0.0289	0.0337	0.0289	0.0393
Mean	0.0630	0.0653	0.0630	0.0688	0.0816	0.0851	0.0816	0.0903

Notes: Standard errors clustered at the village level, * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The variables Proximity and Opened are dummies which indicate whether the individual lives in a DHS village within 10 km of at least one mining site, and whether the site opened before the birth year of the child. Each village DHS is paired to only one mining site, so that each individual appears only once in the regression. Other variables are control variables. The sample focuses on heavy metal mines.

Figure 39: Linear Trends dropping investment phase - Geographic Treatment



Sources: Authors' elaboration on DHS and SNL data.