时间序列hw5

18300290007 加兴华

1

Generate

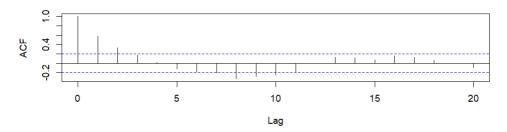
```
n <- 100
ar <- arima.sim(list(order=c(1,0,0), ar=c(0.6)), n=n)
ma <- arima.sim(list(order=c(0,0,1), ma=c(0.9)), n=n)
arma <- arima.sim(list(order=c(1,0,1), ar=c(0.6), ma=c(0.9)), n=n)</pre>
```

参数均小于1, 因此模型causal 和 invertible.

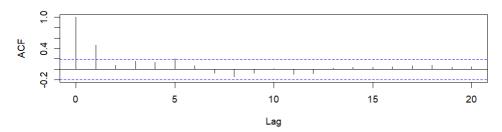
ACF

```
par(mfrow=c(3,1))
acf(ar, main='ACF for AR(1)')
acf(ma, main='ACF for MA(1)')
acf(arma, main='ACF for ARMA(1,1)')
```

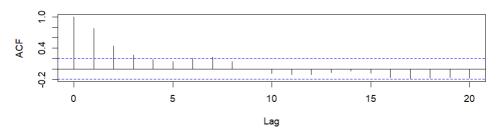
ACF for AR(1)



ACF for MA(1)



ACF for ARMA(1,1)



AR(1):
$$ho(h)=\phi^h,\quad h\geq 0$$

指数衰减趋于0,符合理论值.

MA(1):
$$ho(0)=1; \quad
ho(1)=rac{ heta}{1+ heta^2}; \quad
ho(h)=0, \quad h\geq 2$$

只在0和1处显著,符合理论值.

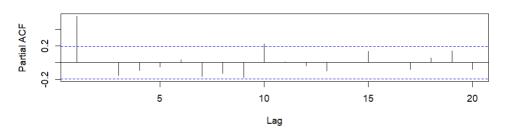
ARMA(1,1):
$$ho(0)=1;\quad
ho(h)=rac{(1+ heta\phi)(\phi+ heta)}{1+2 heta\phi+ heta^2}\phi^{h-1},\quad h\geq 1$$

指数衰减趋于0,符合理论值.

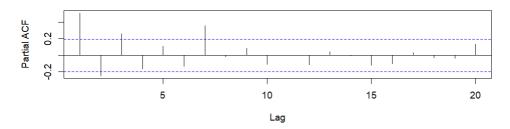
PACF

```
par(mfrow=c(3,1))
acf(ar, main='ACF for AR(1)')
acf(ma, main='ACF for MA(1)')
acf(arma, main='ACF for ARMA(1,1)')
```

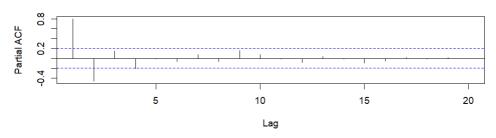
PACF for AR(1)



PACF for MA(1)



PACF for ARMA(1,1)



AR(1):
$$\phi_{11}=\phi; \quad \phi_{ii}=0, \quad i\geq 1$$

只在1处显著,符合理论值.

MA(1):
$$\phi_{hh}=-rac{(- heta)^h(1- heta^2)}{1- heta^{2h+2}},\quad h\geq 1$$

来回震荡幅度衰减到0,符合理论值.

ARMA(1,1):可逆ARMA模型有一个 $AR(\infty)$ 表示,从而理论上不会截尾.

来回震荡幅度衰减到0,符合理论值.

和表 3.1 对比

纵观所有图,只有AR(1)的样本PACF和MA(1)的样本ACF在lag=1发生截尾,其余图都算tails off, 和表结论一致.

2.
$$\hat{X}_{teh} = \beta_1 \hat{X}_{teh+1} + \beta_{h+1} \hat{X}_{teh}$$

$$L(\alpha) \stackrel{?}{=} E(\hat{X}_{teh} - \alpha_1 \hat{X}_{teh+1} - \beta_{h+1} \hat{X}_{teh})^{\frac{1}{2}}$$

$$\Rightarrow \begin{cases} \beta = \operatorname{argmin} L(\alpha). \\ \frac{\partial L}{\partial \alpha_1} = -E_2 \hat{X}_{teh+1} (\hat{X}_{teh} - \alpha_1 \hat{X}_{teh+1} - \beta_{h+1} \hat{X}_{teh}) \\ \Rightarrow E_1 \hat{X}_{teh+1} (\hat{X}_{teh} - \beta_1 \hat{X}_{teh+1} - \beta_{h+1} \hat{X}_{teh}) = 0. \quad i = 1, 2, \dots, h^{-1} \end{cases}$$

i.e. $\langle \hat{X}_{teh+1}, \hat{X}_{teh} - \hat{X}_{teh} \rangle = 0. \quad i = 1, 2, \dots, h^{-1}$

$$\therefore \hat{X}_{teh} - \hat{X}_{teh} + \hat{X}_{teh+1}, \dots, \hat{X}_{teh+1}, \dots, \hat{X}_{teh+1} \hat{I}.$$

$$|\hat{I}| = \frac{1}{2} \frac{1}{2} \frac{1}{2} \hat{X}_{teh+1} + \hat{X}_{teh+1}, \dots, \hat{X}_{teh+1} \hat{I}.$$

4

(a) | < x, y> | ≤ || x || · || y || . \ x, y ∈ 2

< tx+y, tx+y> >0. Yt

- => <x,x>t+2<x,y>t+<y,y>>0. yt.
- ⇒ \(\alpha = 4 < \x, y > 4 < \x, x > < y, y > \(\in 0 \).
- = 1<x,y>1 € 11×11·11/11.
- (b) ||x+y || € ||x|| + ||y||. Yx,y ∈ H

 $LHS = \langle x + y, x + y \rangle = \langle x, x \rangle + \langle y, y \rangle + 2 \langle x, y \rangle$

RHS = < x, x> + < y, y> + 2 11x111y11

由四声 <x,y> < 11x11-11y11.

· LHS & RHS 将证.

由投影唯一性

 $\Rightarrow \chi_{t+h} = \sum_{j=1}^{p} \phi_j \chi_{t+h-i}$

6.
$$d_{nh} = -\frac{(-\theta)^{3}(1-\theta)^{3}}{1-\theta^{2(h+1)}}$$
 $h \ge 1$, $MA(1) = X_{t} = W_{t} + \theta W_{24}$, $|\theta| < 1$

$$Q_{hh} = corc X_{t} + ch - \hat{X}_{t} + h$$
, $X_{t} - \hat{X}_{t} = \hat{X}_{t}$.

$$\hat{X}_{t} = \sum_{i=1}^{h-1} \beta_{i} X_{t} + h \cdot i$$
. $\hat{X}_{t} = \sum_{i=1}^{h-1} \beta_{i} X_{t} + i$.

$$\beta = \underset{i=1}{argmin} E(X_{t} + h - \sum_{i=1}^{h-1} d_{i} X_{t} + h \cdot i)^{2}$$

$$= \underset{i=1}{argmin} E(W_{2th} + \theta W_{t} + h \cdot i)^{2}$$

$$= \underset{i=1}{curgmin} (1 + (d_{1} - \theta)^{2} + \sum_{i=1}^{h-1} (d_{i} \theta + d_{i} + 1)^{2} + (\theta d_{h} \cdot i)^{2})^{2}$$

$$= \underset{i=1}{curgmin} (1 + (d_{1} - \theta)^{2} + \sum_{i=1}^{h-1} (d_{i} \theta + d_{i} + 1)^{2} + (\theta d_{h} \cdot i)^{2})^{2}$$

$$= \underset{i=1}{curgmin} (1 + (d_{1} - \theta)^{2} + \sum_{i=1}^{h-1} (d_{i} \theta + d_{i} + 1)^{2} + (\theta d_{h} \cdot i)^{2})^{2}$$

$$= \underset{i=1}{curgmin} (1 + (d_{1} - \theta)^{2} + d_{1} \cdot i)^{2} + (d_{1} \theta + d_{1} \cdot i)^{2} + (d_{1} \theta + d_{1} \cdot i)^{2} + (d_{1} \theta + d_{1} \cdot i)^{2}$$

$$= \underset{i=1}{curgmin} (1 + (d_{1} - \theta)^{2} + d_{1} \cdot i)^{2} + (d_{1} \theta + d_{1} \cdot i)^{2} +$$

$$\frac{2}{3}i = 1 : \beta_{1} = -0 \cdot (1 + 0^{2} + \cdots + 0^{2(h-2)}) \cdot (0\beta_{1} - 1)$$

$$= -0 \cdot \frac{1 - 0^{2(h-1)}}{1 - 0^{2}} \cdot (0\beta_{1} - 1)$$

$$= \beta_{1} \left[1 + 0^{2} + \frac{1 - 0^{2h-2}}{1 - 0^{2}} \right] = 0 \cdot \frac{1 - 0^{2h-2}}{1 - 0^{2}} \cdot \frac{\pi}{3} \beta_{1}$$

$$= \beta_{1} = 0 \cdot \frac{1 - 0^{2h-2}}{1 - 0^{2h}} ; \quad 0\beta_{1} - 1 = -\frac{1 - 0^{2}}{1 - 0^{2h}}$$

$$= \beta_{1} = -(-0)^{\frac{1}{1}} \frac{1 - 0^{2h-2}}{1 - 0^{2h}}$$

$$= \beta_{1} = -(-0)^{\frac{1}{1}} \frac{1 - 0^{2h-2}}{1 - 0^{2h}}$$

$$Cov \left(\begin{array}{c} X_{t}ph - \hat{X}_{t}h, & X_{t} - \hat{X}_{t} \end{array} \right)$$

$$= Cov \left(\begin{array}{c} W_{t}h + 0 & W_{t}h - \hat{X}_{t}h - \hat{X}_{t}h \\ & = Cov \left(\begin{array}{c} W_{t}h + 0 & W_{t}h - \hat{X}_{t}h - \hat{X}_{t}h \\ & = Cov \left(\begin{array}{c} W_{t}h + 0 & W_{t}h - \hat{X}_{t}h - \hat{X}_{t}h \\ & = Cov \left(\begin{array}{c} W_{t}h + 0 & W_{t}h - \hat{X}_{t}h - \hat{X}_{t}h \\ & = Cov \left(\begin{array}{c} W_{t}h + 0 & W_{t}h - \hat{X}_{t}h - \hat{X}_{t}h$$

$$= \sqrt{W^2 \frac{1-2^{\lambda_{12}}}{1-\theta^{\lambda_{11}}}} = Var(X_{12} - \hat{X}_{13})$$

$$\varphi_{hh} = \frac{-(-0)^{h}(1-0^{2})}{1-0^{h+2}}$$