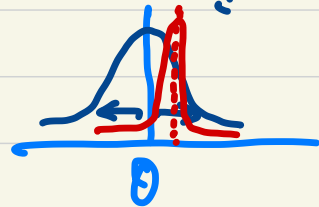




The Gauss-Markov Theorem.

When Assumptions A.1 - A.4 hold, the least squares estimator $b = (X'X)^{-1}X'y$ is the minimum variance linear unbiased estimator (or BLUE) of β .

- A.1 : $y = X\beta + \varepsilon$. The model is correctly specified.
- A.4 : homoskedasticity & no autocorrelation.
- linear estimator.
- trade-off of bias and variance.



Asymptotic (large sample) properties. $\|x - c\| > \varepsilon$

• $n \rightarrow \infty$

$|x_i - c_i| > \varepsilon.$

• Convergence in probability (Def. D1)

A random variable x_n converges in probability to a constant C (a random variable z) if

$$\lim_{n \rightarrow \infty} \text{Prob}(|x_n - C| > \varepsilon) = 0$$

for any positive ε .

$\Rightarrow \text{plim } x_n = C$

Let $x_n = \begin{cases} 0 & \text{with prob. } 1 - \frac{1}{n} \\ 1 & \text{" } \frac{1}{n} \end{cases}$ $\text{plim } x_n = 0$.

Consistent estimator (Def D.2).

An estimator $\hat{\theta}_n$ of a parameter θ is a consistent estimator of θ iff $\text{plim } \hat{\theta}_n = \theta$.

Theorem D.4

The mean of a random sample $\bar{x} = \frac{\sum x_i}{n}$ from any population with finite μ and σ^2 is a consistent estimator of μ .

The OLS estimator is consistent. $\Leftrightarrow \text{plim } \mathbf{b} = \boldsymbol{\beta}$

$$\mathbf{b} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{y} \quad (\text{A.1, A.2})$$

$$= (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'(\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon})$$

$$= \boldsymbol{\beta} + \underbrace{(\mathbf{X}'\mathbf{X})^{-1}}_{\cdot} \underbrace{(\mathbf{X}'\boldsymbol{\varepsilon})}_{\cdot}$$

$$\begin{aligned} \text{plim } \mathbf{b} &= \text{plim } \boldsymbol{\beta} + \text{plim } \underline{\hspace{2cm}} \\ &= \boldsymbol{\beta} + \text{plim } \underline{\hspace{2cm}} \end{aligned}$$

Theorem D.14

If $\text{plim } \mathbf{X}_n = \mathbf{A}$, $\text{plim } \mathbf{Y}_n = \mathbf{B}$, then
 $\text{plim } \mathbf{X}_n \mathbf{Y}_n = \mathbf{A} \mathbf{B}$.

$$\cdot (X'X)^{-1}.$$

$$X'X = \begin{matrix} (k \times n) & (n \times k) \\ x_s' & \begin{bmatrix} \text{---} \\ \text{---} \\ \text{---} \\ \text{---} \\ \text{---} \end{bmatrix} \end{matrix} \begin{matrix} x_t \\ \left\{ \begin{matrix} \vdots \\ x_{i,t} \\ \vdots \end{matrix} \right\} \end{matrix}$$

$\underbrace{\hspace{10em}}_{n \text{ obs.}} \quad \quad \quad \underbrace{\hspace{10em}}_{n \text{ obs.}}$

The (s, t) element of $X'X$ is $\sum_{i=1}^n x_{is} \cdot x_{it}$.

$$\frac{1}{n} X'X$$

A.S.a $(x_i, \varepsilon_i), i=1, \dots, n$ is a sequence of i.i.d. observations.

Theorem D.5. Law of Large numbers.

If $x_i, i=1, \dots, n$ is i.i.d. random sample with $E[x_i] = \mu$. then

$$\lim_{n \rightarrow \infty} \bar{x}_n = \lim_{n \rightarrow \infty} \left(\frac{1}{n} \sum_{i=1}^n x_i \right) = \mu$$

$$X'X = \begin{bmatrix} 1 & \dots & 1 & \dots & 1 \end{bmatrix} \begin{bmatrix} x_1 \\ \vdots \\ x_i \\ \vdots \\ x_n \end{bmatrix} = \sum_{i=1}^n x_i x_i'$$

\uparrow x_i
 $\leftarrow x_i'$

$$\begin{bmatrix} 0 & \dots & 0 & \dots & 0 \end{bmatrix} \begin{bmatrix} 0 \\ \vdots \\ 0 \\ \vdots \\ 0 \end{bmatrix} = \sum 0 \cdot 0$$

$$\frac{1}{n} X'X = \frac{1}{n} \sum_{i=1}^n x_i x_i' \xrightarrow{P} E[x_i x_i'] .$$

Let $E[x_i x_i'] = Q \leftarrow \text{positive definite, invertible.}$

We have $\text{plim} \left(\frac{1}{n} X'X \right) = Q.$

$$\Rightarrow \underline{\text{plim} \left(\frac{1}{n} X'X \right)^{-1} = Q^{-1} .}$$

$X'\varepsilon$. (A.3)

$$\begin{aligned} E[X'\varepsilon] &= E_x[E[X'\varepsilon|X]] = E_x[X' \underbrace{E[\varepsilon|X]}_0] \\ &= E_x[X'0] = 0, \end{aligned}$$

$$\Rightarrow E[x_i \varepsilon_i] = 0.$$

$$X'\varepsilon = \sum_{i=1}^n x_i \varepsilon_i$$

$$\underbrace{\text{plim} \left(\frac{1}{n} X'\varepsilon \right)} = \text{plim} \left(\frac{1}{n} \sum_{i=1}^n x_i \varepsilon_i \right) = \underline{E[x_i \varepsilon_i] = 0}.$$

$$\begin{aligned}\text{plim}(\hat{\beta}) &= \beta + \text{plim}\left(\frac{1}{n}X'X\right)^{-1} \cdot \text{plim}\left(\frac{1}{n}X'u\right) \\ &= \beta + Q^{-1} \cdot 0 \\ &= \beta.\end{aligned}$$

Under assumptions A.1 ~ A.3 and A.5.a,
 $\hat{\beta}$ is a consistent estimator of β .