高级计量经济学

Assignment 2

1. Show that the OLS residuals of the following regression models are numerically identical.

$$(1) \mathbf{y} = \mathbf{X}_1 \beta_1 + \mathbf{X}_2 \beta_2 + \varepsilon,$$

(2)
$$\mathbf{y}^* = \mathbf{X}_2^* \beta_2^* + \varepsilon^*$$
,

where
$$\mathbf{X}_2^* = \mathbf{M}_1 \mathbf{X}_2$$
, $\mathbf{y}^* = \mathbf{M}_1 \mathbf{y}$, and $\mathbf{M}_1 = \mathbf{I} - \mathbf{X}_1 (\mathbf{X}_1' \mathbf{X}_1)^{-1} \mathbf{X}_1'$.

2. Consider random variables x and y whose joint density function is f(x, y). Prove the following properties.

(1)
$$E[xy] = E_x[xE[y|x]]$$

(2)
$$Cov[x, y] = Cov_x[x, E[y|x]] = \int_x (x - E[x]) E[y|x] f_x(x) dx.$$

Solution

1.

We start with regression (2). The residual maker matrix of this regression is

$$\mathbf{M}^* = \mathbf{I} - \mathbf{X}_2^* (\mathbf{X}_2^{*'} \mathbf{X}_2^*)^{-1} \mathbf{X}_2^{*'}.$$

The OLS residual is then

$$\mathbf{M}^* \mathbf{y}^* = \mathbf{y}^* - \mathbf{X}_2^* (\mathbf{X}_2^{*'} \mathbf{X}_2^*)^{-1} \mathbf{X}_2^{*'} \mathbf{y}^*$$

$$= \mathbf{M}_1 \mathbf{y} - \mathbf{M}_1 \mathbf{X}_2 (\mathbf{X}_2' \mathbf{M}_1' \mathbf{M}_1 \mathbf{X}_2)^{-1} (\mathbf{M}_1 \mathbf{X}_2)' \mathbf{M}_1 \mathbf{y}$$

$$= \mathbf{M}_1 \mathbf{y} - \mathbf{M}_1 \mathbf{X}_2 (\mathbf{X}_2' \mathbf{M}_1 \mathbf{X}_2)^{-1} \mathbf{X}_2' \mathbf{M}_1 \mathbf{y}$$

For regression (1), from the deviation of the FWL Theorem, one sees that

$$\mathbf{b}_2 = (\mathbf{X}_2^{*'}\mathbf{X}_2^{*})\mathbf{X}_2^{*'}\mathbf{y}^{*} = (\mathbf{X}_2'\mathbf{M}_1\mathbf{X}_2)^{-1}\mathbf{X}_2'\mathbf{M}_1\mathbf{y}$$

and

$$\begin{aligned} \mathbf{b}_1 &= (\mathbf{X}_1' \mathbf{X}_1)^{-1} \mathbf{X}_1' (\mathbf{y} - \mathbf{X}_2 \mathbf{b}_2) \\ &= (\mathbf{X}_1' \mathbf{X}_1)^{-1} \mathbf{X}_1' \left[\mathbf{y} - \mathbf{X}_2 (\mathbf{X}_2' \mathbf{M}_1 \mathbf{X}_2)^{-1} \mathbf{X}_2' \mathbf{M}_1 \mathbf{y} \right] \\ &= (\mathbf{X}_1' \mathbf{X}_1)^{-1} \mathbf{X}_1' \left[\mathbf{I} - \mathbf{X}_2 (\mathbf{X}_2' \mathbf{M}_1 \mathbf{X}_2)^{-1} \mathbf{X}_2' \mathbf{M}_1 \right] \mathbf{y} \end{aligned}$$

Therefore, the OLS residual is

$$\begin{split} \mathbf{y} - \mathbf{X}_1 \mathbf{b}_1 - \mathbf{X}_2 \mathbf{b}_2 &= \mathbf{y} - \mathbf{X}_1 (\mathbf{X}_1' \mathbf{X}_1)^{-1} \mathbf{X}_1' \big[\mathbf{I} - \mathbf{X}_2 (\mathbf{X}_2' \mathbf{M}_1 \mathbf{X}_2)^{-1} \mathbf{X}_2' \mathbf{M}_1 \big] \mathbf{y} \\ &- \mathbf{X}_2 (\mathbf{X}_2' \mathbf{M}_1 \mathbf{X}_2)^{-1} \mathbf{X}_2' \mathbf{M}_1 \mathbf{y} \\ &= \big[\mathbf{I} - \mathbf{X}_1 (\mathbf{X}_1' \mathbf{X}_1)^{-1} \mathbf{X}_1' + \mathbf{X}_1 (\mathbf{X}_1' \mathbf{X}_1)^{-1} \mathbf{X}_1' \mathbf{X}_2 (\mathbf{X}_2' \mathbf{M}_1 \mathbf{X}_2)^{-1} \mathbf{X}_2' \mathbf{M}_1 \big] \mathbf{y} \\ &- \mathbf{X}_2 (\mathbf{X}_2' \mathbf{M}_1 \mathbf{X}_2)^{-1} \mathbf{X}_2' \mathbf{M}_1 \big] \mathbf{y} \\ &= \big[\mathbf{M}_1 - \mathbf{M}_1 \mathbf{X}_2 (\mathbf{X}_2' \mathbf{M}_1 \mathbf{X}_2)^{-1} \mathbf{X}_2' \mathbf{M}_1 \big] \mathbf{y} \end{split}$$

which coincides with the OLS residual of regression (2).

2. (1)

By definition, one has

$$E[xy] = \int_{x} \int_{y} xy f(x, y) dy dx$$

$$= \int_{x} \int_{y} xy f(y|x) f_{x}(x) dy dx$$

$$= \int_{x} x \left[\int_{y} y f(y|x) dy \right] f_{x}(x) dx$$

$$= \int_{x} x E[y|x] f_{x}(x) dx$$

$$= E_{x} [x E[y|x]]$$

(2) The first equality follows from applying the property Cov(x, y) = E[xy] - E[x]E[y] by replacing y with E[y|x], i.e.,

$$\begin{aligned} \operatorname{Cov}_{\boldsymbol{x}}(\boldsymbol{x}, \operatorname{E}[\boldsymbol{y} \,|\, \boldsymbol{x}]) &= \operatorname{E}_{\boldsymbol{x}}\big[\boldsymbol{x} \operatorname{E}[\boldsymbol{y} \,|\, \boldsymbol{x}]\big] - \operatorname{E}[\boldsymbol{x}] \operatorname{E}_{\boldsymbol{x}}\big[\operatorname{E}[\boldsymbol{y} \,|\, \boldsymbol{x}]\big] \\ &= \operatorname{E}[\boldsymbol{x} \, \boldsymbol{y}] - \operatorname{E}[\boldsymbol{x}] \operatorname{E}[\boldsymbol{y}] \qquad \text{(usign (1) and Law of Iterated Expectation)} \\ &= \operatorname{Cov}[\boldsymbol{x}, \boldsymbol{y}] \end{aligned}$$

The second equality can then be derived from definition:

$$Cov_{x}(x, E[y|x]) = \int_{x} (x - E[x]) (E[y|x] - E_{x}[E[y|x]]) f_{x}(x) dx$$

$$= \int_{x} (x - E[x]) E[y|x] f_{x}(x) dx - \int_{x} (x - E[x]) E[y] f_{x}(x) dx$$

$$= \int_{x} (x - E[x]) E[y|x] f_{x}(x) dx - E[y] \int_{x} (x - E[x]) f_{x}(x) dx$$

$$= \int_{x} (x - E[x]) E[y|x] f_{x}(x) dx - E[y] \cdot 0$$

$$= \int_{x} (x - E[x]) E[y|x] f_{x}(x) dx$$