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#### **APPLICATION**



## POLYGENE: Population genetics analyses for autopolyploids based on allelic phenotypes

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#### **Abstract**

- 1. Polyploidy has appeared in almost every ancestral plant lineage, and in extant species, occurs frequently. When present, polyploidy presents problems for genetic data analysis, which are caused by both genotypic ambiguities and double-reduction.
- 2. To address these problems, we developed a new software package, POLYGENE, which enables the estimation of genotypic frequencies for a number of polysomic inheritance models. Specifically, POLYGENE obtains posterior probabilities for genotypes hidden within allelic phenotypes.
- 3. Comprehensive modes of genetic analyses are provided by POLYGENE, which include genetic diversity analysis, tests for allelic phenotypic or genotypic distributions, linkage disequilibrium and genetic differentiation, genetic distance analysis, principal coordinates analysis, hierarchical clustering analysis, individual inbreeding coefficient estimation, individual heterozygosity index estimation, population assignment, pairwise relatedness estimation, parentage analysis, analysis of molecular variance and Bayesian clustering.
- 4. POLYGENE enables easy and convenient allelic phenotype- or genotype-based analysis for both autopolyploids and diploids. POLYGENE will thus facilitate molecular ecology research involving autopolyploids.

allelic phenotype, AMOVA, double-reduction, parentage analysis, polysomic inheritance, population genetics

#### 1 | INTRODUCTION

Polyploids are cells or organisms having a genome with more than two sets of homologous chromosomes. Polyploidy has occurred in almost every ancestral plant lineage, and frequently occurs in extant species, especially in plants (Barker, Arrigo, Baniaga, Li, & Levin, 2016). Due to their often important role in plant speciation,

polyploids are regularly the subject of theoretical and experimental studies for evolutionary biology, molecular ecology and agriculture (Ling et al., 2018).

There are two distinct mechanisms of genome duplication that result in polyploidy: allopolyploidy and autopolyploidy. In allopolyploidy, chromosomes originate from two species; in autopolyploidy, all chromosomes originate within a single species, often due to

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Methods in Ecology and Evolution HUANG ET AL.

unreduced gametes. This paper mainly focuses on autopolyploids and allopolyploids that display polysomic inheritance.

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In autopolyploids, both bivalents and multivalents can be formed during meiosis, resulting in disomic and polysomic inheritance, respectively. These are two extremes and many autopolyploid taxa represent intermediate stages (Butruille & Boiteux, 2000). Allopolyploids generally display disomic inheritance (Luo et al., 2006), because chromosomes from different species are not completely homologous. However, multivalent pairing can also occur in allopolyploids, resulting in a mixed inheritance pattern across loci in the genome, termed segmental allopolyploidy (Stebbins, 1950).

Because of differences in data formats and inheritance models. computer software designed for diploid organisms, such as GENEPOP (Rousset, 2008) and ARLEQUIN (Excoffier & Lischer, 2010), cannot be used for autopolyploid species. In general, polyploid population genetics analyses present two major challenges: (a) genotyping ambiguities and (b) double-reduction.

For PCR-based markers, the dosage of alleles cannot be directly determined by electrophoresis. For example, for autotetraploids, the genotype AABB has the same electrophoresis band pattern as the genotype AAAB. Therefore, the true genotype of individuals cannot be determined by electrophoresis and only the allelic phenotype is available for these markers. Novel technologies such as genotyping-by-sequencing (GBS) also uses PCR and suffers from such problems. Some researchers have claimed that genotypes can be assigned from the band intensity (Esselink, Nybom, & Vosman, 2004) or allelic read depth (Voorrips, Gort, & Vosman, 2011) using such methods. However, due to varying amplification efficiency among alleles in PCR, the results of these methods are still unreliable. Hereafter, we denote the allelic phenotype as a set of alleles and the genotype as a multiset of alleles, with the allelic phenotype being abbreviated as the 'phenotype'.

A peculiarity of polysomic inheritance is double-reduction, which occurs from a combination of three major events during meiosis: (a) the crossing-over between non-sister chromatids, (b) an appropriated pattern of disjunction, and (c) the subsequent migration of the chromosomal segments carrying a pair of sister chromatids to the same gamete (Darlington, 1929). For example, assume a tetraploid individual ABCD produces a gamete AA. Double reduction will result in gametes carrying identical-bydescent (IBD) alleles and increased homozygosity. A brief description of polysomic inheritance models can be found in the Supporting Information S1.

To solve these two problems, we developed a new software package: POLYGENE. We note that there are several other current software packages for polysomic inheritance, such as POLYSAT (Clark & Jasieniuk, 2011), spagedi (Hardy & Vekemans, 2002), polyrelated-NESS (Huang, Ritland, Guo, Shattuck, & Li, 2014), GENODIVE (Meirmans & Tienderen, 2004) and STRUCTURE (Pritchard, Stephens, & Donnelly, 2000). POLYSAT converts the phenotype into a binary array and uses methods for dominant markers for analysis. Because the models of dominant markers and codominant markers are different, such a

method is biased. STRUCTURE cannot handle ambiguous genotypes. The remaining three software packages assume that genotypic frequencies are in accordance with the Hardy-Weinberg equilibrium (HWE). Because genotypic frequencies under double reduction or inbreeding will always deviate from those expected under the HWE, this will override any heterozygous genotypes and bias the results. Moreover, some applications are sensitive to double reduction. For example, in a parentage analysis, the true father of the genotype ABBB will be excluded when the offspring genotype is AACC.

### 2 | THE NEW SOFTWARE PACKAGE 'POLYGENE'

POLYGENE v1.0b is written in C++ and C# with a graphical user interface. The source code is available from the GitHub website (http:// github.com/huangkang1987/polygene). The binary executables for three platforms (Windows, Linux and Mac OS X) are provided. For the remaining Linux platforms, users can compile their own source code. Instructions on how to do this can be found in the user manual. To ensure free copying, distribution and modifications of the software and its source code, POLYGENE is distributed under the terms of a GNU General Public License, version 3.

POLYGENE supports a maximum ploidy level of 10, with different ploidy levels among populations also supported. The ploidy levels of all individuals within the same population are assumed to be 4 equal, so that tests for allelic phenotypic distribution, linkage disequilibrium and differentiations can all be conducted. POLYGENE also supports multi-level region definition (Huang, Li, Dunn, Zhang, & Li, 2019), so as to analyse the variance component in different hierarchies in analysis of molecular variance (AMOVA). Other analyses are also supported for each hierarchy (e.g. genetic diversity, genetic distance, principal coordinate analysis (PCoA), hierarchical clustering). The hardware requirements to run POLYGENE are not high; a computer with a 4 GiB memory is sufficient.

### | INPUT AND OUTPUT FORMAT

An example for the format of inputting the allelic phenotypic and genotypic data is shown in Table 1. Here, the first row is the header with each subsequent row representing an individual. The first three columns are: (a) the individual, (b) the population and

TABLE 1 An example of inputting allelic phenotypic/genotypic data

ID	Pop	Ploidy	Loc1	Loc2
Ind1	pop1	4	2,3,4	1,2,4
Ind2	pop1	4	4	2,3,4
Ind3	pop2	4	2,3,4,4	2,3,4,3
Ind4	pop2	4	1,3,4,1	3,4,3,3

HUANG ET AL. Methods in Ecology and Evolution

(c) the ploidy level. Each cell of the fourth column onwards represents either a phenotype (the first two individuals) or a genotype (the latter two individuals) at a specific locus, with these columns being separated by tabs to enable data to be pasted from a spread-sheet. The alleles within a phenotype are identified by positive integers, which are separated by commas. For missing values, the corresponding cells are left blank. POLYGENE supports multi-level region definition, where each region is defined as a collection of populations or regions (e.g. Level I region is a collection of populations, Level II region is a collection of level I regions). The region can be defined in the parameters of the *analysis of molecular variance* (AMOVA, Figure 2).

POLYGENE is able to handle both phenotypic and genotypic data. The function 'remove duplicated alleles' is provided so as to convert genotypes into phenotypes. If this function is not used, the candidate genotypes will be generated based on the dosage of the inputted alleles. For example, an incomplete tetraploid genotype ABB will generate two candidate genotypes: ABBA and ABBB. If null alleles are also considered, the genotype ABBY will also be added to the list, where Y is the null allele. The missing data are usually not analyzed, except for the dummy haplotype-based methods or when null alleles are considered.

Performing each analysis in POLYGENE is relatively straightforward. First, the phenotypic data are formatted and then entered into the box marked 'Phenotypes/Genotypes.' Second, the required analytical method(s) is selected and the corresponding parameters on the page marked 'Parameters' is configured (Figure 1). Finally, the

button marked 'Calc' in the toolbox menu at the top of the window is 'clicked'. All analyses will then be performed simultaneously, and the progress bar will show the progress of the current analysis. After a short delay, the results will be displayed in plain text on the corresponding pages (Figure 2).

#### 4 | MATERIALS AND METHODS

POLYGENE improves on existing packages and also includes four polysomic inheritance models, that is, random chromosome segregation (RCS) (Muller, 1914), pure random chromatid segregation (PRCS) (Haldane, 1930), complete equational segregation (Mather, 1935) and partial equational segregation (Huang, Wang, et al., 2019). A brief description of these models can be found in the Supporting Information S1 (section double reduction models). The disomic inheritance model can also be used, whose genotypic frequency is equivalent to the RCS model.

These inheritance models determine the a priori probabilities of the genotypic frequencies, based on the genotypic and phenotypic frequencies under each model (Huang, Wang, et al., 2019). We are able to calculate the posterior probabilities of the genotypes hidden behind the phenotypes and these genotypes are weighted by their posterior probabilities. The correct model can accurately obtain the posterior probabilities of hidden genotypes and improve the accuracy of the analyses. The optimal model can be evaluated by the *Bayesian information criterion* (BIC), and the BIC of an

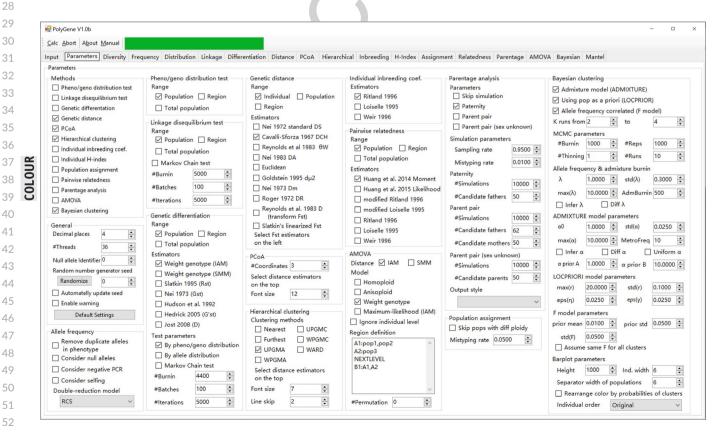


FIGURE 1 The graphical user interface of POLYGENE

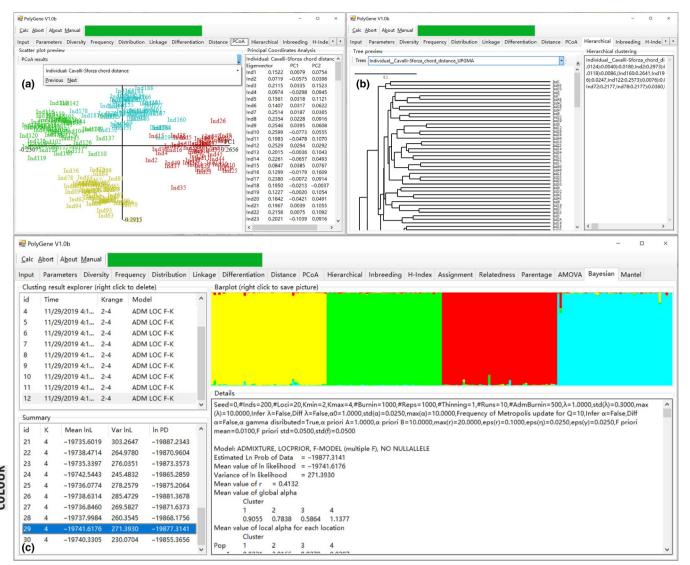


FIGURE 2 An example of the results output of PCoA, hierarchical clustering and Bayesian clustering. Simulated dataset D using the CES model is presented as an example. PCoA and hierarchical clustering are based on individual Cavalli-Sforza chord distance. The default settings are applied for Bayesian clustering

inheritance model is sum of all of the BIC values over all loci for all populations. The BIC is provided in the results of a phenotypic distribution test.

Multiple genetic analysis methods are incorporated into POLYGENE, and these can be classified into four categories: (a) weighted genotype, (b) allele frequency, (c) phenotype and (d) dummy allele/haplotype. The effects of double reduction, null alleles, negative amplification and self-fertilization can be freely taken into account in most methods provided by our new software. The analytical methods and their associated parameters are presented in Figure 1. An extensive description of the methods used in POLYGENE can be found in the Supporting Information S1. Because the genotypes are weighted at each locus independently, this assumes linkage equilibrium and some prior distribution of genotypes in all methods except for those that are phenotype-based. We briefly describe these methods below.

Weighted genotype-based methods are based on the weighting of extracted genotypes according to their posterior probabilities. These methods are allele frequency estimation based on an EM algorithm (De Silva, Hall, Rikkerink, McNeilage, & Fraser, 2005; Kalinowski & Taper, 2006), genetic diversity analysis, parentage analysis (Huang, Mi, Dunn, Wang, & Li, 2018; Kalinowski, Taper, & Marshall, 2007), kinship coefficient estimation (Loiselle, Sork, Nason, & Graham, 1995; Ritland, 1996; Weir, 1996), and pairwise relatedness coefficient estimation (Huang, Guo, et al., 2015; Huang, Ritland, et al., 2015; Huang et al., 2014). The effects of null alleles, negative amplification and self-fertilization can be freely taken into account in most methods, which use the genotypic and phenotypic frequencies. The likelihood equations in parentage analysis can incorporate self-fertilization.

Allele frequency-based methods are based on the estimated allele frequencies, which do not need specific modification for

1 polysomic inheritance. These are  $F_{\rm ST}$  and its analogues (Hedrick, 2 3 4 5 6 7 8 9 10 11 12 13 14 15

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2005; Hudson, Slatkin, & Maddison, 1992; Jost, 2008; Nei, 1973; Slatkin, 1995), genetic distance (Cavalli-Sforza & Edwards, 1967; Goldstein, Ruiz, Cavallisforza, & Feldman, 1995; Nei, 1972; Nei & Roychoudhury, 1974; Nei, Tajima, & Tateno, 1983; Reynolds, Weir, & Cockerham, 1983; Rogers, 1972; Slatkin, 1995), principal coordinate analysis, hierarchical clustering analysis and Mantel tests (Mantel, 1967; Smouse & Sokal, 1986). Genetic distance indices can be extended to that between two individuals. For instance, the frequencies of alleles within an individual can be obtained from the phenotypes by using the posterior probabilities of the candidate genotypes. The latter three methods are based on the estimated genetic distance matrix. The phenotype-based methods are based on the observed or

expected distribution of phenotypes. These are tests for linkage disequilibrium, genetic differentiation, phenotypic distribution and population assignment. The first two tests are respectively performed by Fisher's G test and the Markov Chain test (Raymond & Rousset, 1995) with the null hypotheses that two loci are under linkage equilibrium or the phenotypes in multiple populations are drawn from the same distribution. The phenotypic distribution test is performed by Fisher's G test, the null hypothesis of which is that the observed phenotypes accord with double-reduction. Population assignment assigns each individual to the population with the highest likelihood (Paetkau, Slade, Burden, & Estoup, 2004) to find the natal population of each individual.

Dummy allele or haplotype-based methods are based on the weighting of allele copies in the phenotypes. This technique is used in Weir and Cockerham's (1984)  $F_{ST}$  estimator, AMOVA (Cockerham, 1973; Excoffier, Smouse, & Quattro, 1992) and Bayesian clustering (Pritchard et al., 2000). AMOVA is generalized to any number of hierarchies, and the genetic distance between dummy alleles or haplotypes are calculated during variance decomposition. Bayesian clustering implements three models: the ADMIXTURE model

TABLE 2 RMSE values of estimated effective number of alleles, heterozygosity and F-statistics for our six simulated datasets

	Dataset	$A_{e}$	H <sub>I</sub>	H <sub>s</sub>	H <sub>T</sub>	F <sub>IS</sub>	F <sub>ST</sub>
SPAGEDI	A	0.2784	0.3467	0.0227	0.0170	0.4485	0.0041
	В	0.2504	0.3467	0.0205	0.0010	0.4486	0.0134
	С	0.2671	0.4211	0.0219	0.0146	0.5598	0.0036
	D	0.2364	0.4211	0.0196	0.0012	0.5583	0.0139
	E	0.3029	0.3000	0.0245	0.0224	0.3783	0.0040
	F	0.2727	0.3000	0.0223	0.0009	0.3783	0.0159
	G	0.2714	0.3677	0.0222	0.0195	0.4791	0.0041
	Н	0.2389	0.3677	0.0197	0.0009	0.4795	0.0129
GENODIVE	Α	0.2123	0.3151	0.0207	0.0094	0.4433	0.0028
	В	0.1950	0.3144	0.0164	0.0014	0.4434	0.0126
	С	0.2410	0.3858	0.0231	0.0099	0.5515	0.0033
	D	0.2098	0.3862	0.0182	0.0013	0.5497	0.0138
	Е	0.2139	0.2391	0.0194	0.0150	0.3270	0.0135
	F	0.1877	0.2364	0.0155	0.0010	0.3258	0.0128
	G	0.2354	0.2618	0.0206	0.0143	0.3659	0.0127
	Н	0.1917	0.2580	0.0156	0.0010	0.3613	0.0124
POLYGENE <sup>a</sup>	Α	0.2181	0.0373	0.0215	0.0086	0.0414	0.0066
	В	0.1769	0.0370	0.0168	0.0019	0.0454	0.0129
	С	0.2387	0.0378	0.0228	0.0093	0.0478	0.0079
	D	0.1879	0.0386	0.0177	0.0021	0.0517	0.0136
	Е	0.2120	0.0331	0.0202	0.0085	0.0351	0.0060
	F	0.1746	0.0325	0.0166	0.0016	0.0381	0.0131
	G	0.2150	0.0378	0.0210	0.0083	0.0455	0.0068
	Н	0.1879	0.0638	0.0177	0.0021	0.0853	0.0136

Note: Header row:  $A_a$ : effective number of alleles;  $H_i$ : observed heterozygosity;  $H_s$ : expected heterozygosity in a subpopulation;  $H_{T}$ : expected heterozygosity in the total population;  $H_{I}$ ,  $H_{S}$  and  $F_{\rm IS}$  is calculated for each subpopulation at each locus; while  $H_{\rm T}$  and  $F_{\rm ST}$  is calculated in the total population at each locus. The optimal RMSEs are marked in bold.

<sup>a</sup>The optimal model evaluated from BIC are: A: PRCS (ΔBIC = 10.36), B: PRCS (ΔBIC = 8.44), C: CES (ΔBIC = 81.25), D: CES (ΔBIC = 88.80), E: RCS (ΔBIC = 28.33), F: RCS (ΔBIC = 35.64), G: CES  $(\Delta BIC = 12.94), H: CES (\Delta BIC = 13.89).$ 

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(Pritchard et al., 2000), the LOCPRIORI model (Hubisz, Falush, Stephens, & Pritchard, 2009) and the F model (Falush, Stephens, & Pritchard, 2003) and calculates the probability that each dummy allele is drawn from each cluster.

#### SIMULATIONS AND COMPARISONS

Here, we use both simulated and empirical datasets to evaluate the accuracy of estimated genetic diversity indices (including the effective number of alleles  $A_e$ , the observed heterozygosity  $H_I$ , the expected heterozygosity in a subpopulation  $H_S$  and in the total population  $H_{\mathrm{T}}$ ) and F-statistics ( $F_{\mathrm{IS}}$  and  $F_{\mathrm{ST}}$ ) of the software packages GENODIVE V2.0b23 (Meirmans & Tienderen, 2004), SPAGEDI V1.5d (Hardy & Vekemans, 2002) and POLYGENE V1.0b.

Eight simulated datasets (denoted by Dataset A-H) were generated using the simulation function of POLYGENE. Specifically, Datasets A and E simulate no inbreeding and no differentiation; Datasets B and F simulate differentiation; Datasets C and G simulate inbreeding; and Datasets D and H simulate both inbreeding and differentiation. All datasets were generated with a random number generator seed equal to zero.

There were four populations in each dataset and each had 50 autotetraploids. Twenty tetra-allelic loci were simulated, whose allele frequencies were drawn from a triangular distribution (0.1, 0.2, 0.3 and 0.4), the allele frequencies being equal among populations in Datasets A, C, E and G. Frequencies in Datasets B, D, F and H were shifted accordingly to simulate differentiation. For Datasets C, D, G and H, we used a selfing rate of 0.3 to simulate inbreeding. The genotypes at 20 tetra-allelic loci for these individuals were generated under the PRCS model (Datasets A-D) or the RCS/disomic model (Dataset E-G) and then converted into phenotypes. The true values of the compared statistics were derived using the methods of Huang, Dunn, et al. (2019), Huang, Wang, et al. (2019), and are shown in Table S32.

For POLYGENE, we used the BIC in order to select the best inheritance model. Selfing was considered for Datasets C and D. We estimated statistics other than  $F_{\rm ST}$  using the 'Genetic Diversity' function, with the locus specific  $F_{\rm ST}$  estimated using Nei's (1973) estimator. In SPAGEDI, the option '2.3 global Gst and pairwise Gst' was used with all additional options disabled except 'Report allele freq and diversity coef'. In GENODIVE, all statistics were estimated using the 'Correct for unknown dosage of alleles' option, in which A<sub>e</sub> is estimated by the 'Genetic Diversity' function for each population,  $F_{IS}$  is estimated

Dataset	Species	#Individuals	#Populations	#Loci
Zwart et al. (2016)	Grey willow Salix cinereal	139	1	7
Yang et al. (2018)	Lotus sessilifolius	48	8	11
Hoeltgebaum, Londoño, Lando, and Reis (2017)	Varronia curassavica	681	4	8
Julier et al. (2010)	Lucerne Medicago sativa	463	16	7

TABLE 3 General information for the four empirical datasets used for comparisons

TABLE 4 RMSE values of estimated effective number of alleles, heterozygosity and F-statistics for the four empirical datasets

Dataset	$A_{\rm e}$	H <sub>I</sub>	H <sub>s</sub>	H <sub>T</sub>	F <sub>IS</sub>	F <sub>ST</sub>
Zwart et al. (2016)	2.6848	0.5608	0.3625		0.8837	
Yang et al. (2017)	0.7930	0.4156	0.0443	0.0112	0.7857	0.0468
Hoeltgebaum et al. (2017)	0.6031	0.5274	0.0517	0.0427	0.9028	0.0145
Julier et al. (2010)	0.8240	0.3499	0.0558	0.0469	0.3794	0.0112
Zwart et al. (2016)	2.4647	0.3447	0.3572		0.1900	
Yang et al. (2017)	0.1167	0.1550	0.0328	0.0083	0.2048	0.0582
Hoeltgebaum et al. (2017)	0.3499	0.2014	0.0181	0.0122	0.3040	0.0079
Julier et al. (2010)	0.3754	0.1919	0.0287	0.0520	0.2342	0.0104
Zwart et al. (2016)	0.1703	0.0073	0.0130		0.0221	
Yang et al. (2017)	0.3743	0.0290	0.0487	0.0221	0.0774	0.0303
Hoeltgebaum et al. (2017)	0.3247	0.0164	0.0220	0.0180	0.0424	0.0057
Julier et al. (2010)	0.2770	0.0121	0.0196	0.0164	0.0216	0.0024
	Zwart et al. (2016) Yang et al. (2017) Hoeltgebaum et al. (2017) Julier et al. (2010) Zwart et al. (2016) Yang et al. (2017) Hoeltgebaum et al. (2017) Julier et al. (2010) Zwart et al. (2016) Yang et al. (2017) Hoeltgebaum et al. (2017)	Zwart et al. (2016) 2.6848 Yang et al. (2017) 0.7930 Hoeltgebaum et al. (2017) 0.6031 Julier et al. (2010) 0.8240 Zwart et al. (2016) 2.4647 Yang et al. (2017) 0.1167 Hoeltgebaum et al. (2017) 0.3499 Julier et al. (2010) 0.3754 Zwart et al. (2016) 0.1703 Yang et al. (2017) 0.3743 Hoeltgebaum et al. (2017) 0.3247	Zwart et al. (2016)       2.6848       0.5608         Yang et al. (2017)       0.7930       0.4156         Hoeltgebaum et al. (2017)       0.6031       0.5274         Julier et al. (2010)       0.8240       0.3499         Zwart et al. (2016)       2.4647       0.3447         Yang et al. (2017)       0.1167       0.1550         Hoeltgebaum et al. (2017)       0.3499       0.2014         Julier et al. (2010)       0.3754       0.1919         Zwart et al. (2016)       0.1703       0.0073         Yang et al. (2017)       0.3743       0.0290         Hoeltgebaum et al. (2017)       0.3247       0.0164	Zwart et al. (2016)       2.6848       0.5608       0.3625         Yang et al. (2017)       0.7930       0.4156       0.0443         Hoeltgebaum et al. (2017)       0.6031       0.5274       0.0517         Julier et al. (2010)       0.8240       0.3499       0.0558         Zwart et al. (2016)       2.4647       0.3447       0.3572         Yang et al. (2017)       0.1167       0.1550       0.0328         Hoeltgebaum et al. (2017)       0.3499       0.2014       0.0181         Julier et al. (2010)       0.3754       0.1919       0.0287         Zwart et al. (2016)       0.1703       0.0073       0.0130         Yang et al. (2017)       0.3743       0.0290       0.0487         Hoeltgebaum et al. (2017)       0.3247       0.0164       0.0220	Zwart et al. (2016)       2.6848       0.5608       0.3625         Yang et al. (2017)       0.7930       0.4156       0.0443       0.0112         Hoeltgebaum et al. (2017)       0.6031       0.5274       0.0517       0.0427         Julier et al. (2010)       0.8240       0.3499       0.0558       0.0469         Zwart et al. (2016)       2.4647       0.3447       0.3572         Yang et al. (2017)       0.1167       0.1550       0.0328       0.0083         Hoeltgebaum et al. (2017)       0.3499       0.2014       0.0181       0.0122         Julier et al. (2010)       0.3754       0.1919       0.0287       0.0520         Zwart et al. (2016)       0.1703       0.0073       0.0130         Yang et al. (2017)       0.3743       0.0290       0.0487       0.0221         Hoeltgebaum et al. (2017)       0.3247       0.0164       0.0220       0.0180	Zwart et al. (2016)       2.6848       0.5608       0.3625       0.8837         Yang et al. (2017)       0.7930       0.4156       0.0443       0.0112       0.7857         Hoeltgebaum et al. (2017)       0.6031       0.5274       0.0517       0.0427       0.9028         Julier et al. (2010)       0.8240       0.3499       0.0558       0.0469       0.3794         Zwart et al. (2016)       2.4647       0.3447       0.3572       0.1900         Yang et al. (2017)       0.1167       0.1550       0.0328       0.0083       0.2048         Hoeltgebaum et al. (2017)       0.3499       0.2014       0.0181       0.0122       0.3040         Julier et al. (2010)       0.3754       0.1919       0.0287       0.0520       0.2342         Zwart et al. (2016)       0.1703       0.0073       0.0130       0.0221       0.0774         Hoeltgebaum et al. (2017)       0.3743       0.0290       0.0487       0.0221       0.0774         Hoeltgebaum et al. (2017)       0.3247       0.0164       0.0220       0.0180       0.0424

Note: The optimal RMSEs are marked in bold.

<sup>&</sup>lt;sup>a</sup>The optimal model evaluated from BIC are: Zwart et al. (2016): PES ( $r_s = 0.5$ ) (ΔBIC = 1.95); Yang et al. (2017): PES ( $r_s = 0.25$ ) (ΔBIC = 3.41); Hoeltgebaum et al.'s (2017): CES ( $\Delta$ BIC = 1.40); Julier et al. (2010): CES ( $\Delta$ BIC = 2.24).

by the 'Hardy-Weinberg' function, and  $H_{\rm S}$ ,  $H_{\rm T}$  and  $F_{\rm ST}$  are all estimated using the 'G-statistics' function.

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The results for all simulations are presented in Table 2. We found that POLYGENE is the most accurate package for estimating  $A_{\rm e}$ ,  $H_{\rm I}$ ,  $H_{\rm T}$  and  $F_{\rm IS}$ . Especially for  $H_{\rm I}$  and  $F_{\rm IS}$ , the RMSE using Polygene is greatly improved. Previously available software packages thus appear to be unable to accurately estimate  $H_{\rm I}$  and  $F_{\rm IS}$ , because the estimates for  $H_{\rm I}$  are close to one and those for  $F_{\rm IS}$  are negative. However, Polygene is marginally less accurate than genodive in estimating both  $H_{\rm S}$  and  $F_{\rm ST}$ . For  $H_{\rm T}$ , spaged is most accurate for high differentiation with Polygene being most accurate for low differentiation.

In natural populations, the true values of the compared statistics are usually unknown. To evaluate accuracy, we adopted the statistics calculated from the genotypic datasets for each of several software packages as the true value, and those calculated from the phenotypic datasets as the observed values. We used four empirical genotypic datasets of autotetraploid species (details presented in Table 3).

The software configurations used were the same as for the simulated datasets and the results for the empirical datasets are presented in Table 4. POLYGENE performs best with the RMSE of all statistics being relatively low. In particular, the RMSE of all statistics are optimal for the datasets of Zwart, Elliott, Hopley, Lovell, and Young (2016) and Julier, Semiani, and Laouar (2010). Similar to the results obtained from the simulated datasets, both spaged and Genodive produce inaccurate estimates for  $F_{\rm IS}$  and  $H_{\rm I}$ , with each of these two software packages producing values that deviate from the values calculated from the genotypic data. For Polygene, the results from both the genotypic and phenotypic datasets are similar, with the RMSE reduced by ~90%.

#### 6 | CONCLUSION

We show here that POLYGENE is a useful new tool for molecular ecology studies involving autopolyploids. This new software package utilizes genotypic frequencies under double reduction to perform subsequent analyses, and extends the genetic analyses methods that can be used for polysomic inheritance. POLYGENE also supports data from haploids, diploids and varying ploidy levels among populations, and can be run on Windows, Linux and Mac OS X. POLYGENE enables the analysis of autopolyploids based on allelic phenotypes or genotypes to be as easy and convenient as when using data from diploids. This new software will thus likely facilitate molecular ecology research involving autopolyploid species.

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The authors declare no conflict of interest.

#### **AUTHORS' CONTRIBUTIONS**

K.H. and B.L. conceived the ideas, K.R. designed methodology, D.W.D. checked the model, K.H. and D.W.D. wrote the draft and D.W.D. and K.R. edited the manuscript.

#### DATA AVAILABILITY STATEMENT

The binary executable, user manual and source code of POLYGENE, and the simulated and empirical results are available at GitHub (https://github.com/huangkang1987/polygene) and at Zenodo (https://doi.org/10.5281/zenodo.3541408).

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HUANG ET AL. Methods in Ecology and Evolution

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#### SUPPORTING INFORMATION

Additional supporting information may be found online in the Supporting Information section.

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