

Natural selection in the Health and Retirement Study

David Hugh-Jones

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I investigate natural selection on polygenic scores in the contemporary US, using the Health and Retirement Study. Results support the economic theory of fertility as an explanation for natural selection: among both white and black respondents, scores which correlate negatively (positively) with education are selected for (against). However, selection effects are not larger among people with less education. For whites, selection effects are smaller among married people.

Hugh-Jones and Abdellaoui (2022) explain patterns of natural selection on polygenic scores in the UK, using an economic theory of fertility derived from Becker and Tomes (1976). The theory has two components.

1. There is a trade-off between time spent working and raising children. This leads people with more human capital and higher expected wages to have fewer children.
2. The trade-off is sharper for low-income people, people with low human capital, and single parents. As a result, natural selection is stronger among these groups.

The evidence for point 1 is that polygenic scores which correlate positively with human capital correlate negatively with number of children, i.e. they are being selected against. And vice versa, scores which correlate positively with human capital are being selected for. The evidence for point 2 is that correlations with number of children are larger and more significant among people with lower income, less education and single parents.

Here, I test the same theory in the US population, using the Health and Retirement Survey (HRS; XXX). This gives an independent test of the theory. The HRS is more representative of the population than UK Biobank, which addresses one potential

weakness of the previous paper. Also, the HRS provides precalculated polygenic scores (XXX Aysu) for both black and (non-hispanic) white participants, so I can check whether the human capital theory of natural selection works in both these ethnicities.

Economic theory treats all humans as the same. Once you have controlled appropriately for people's preferences, resources and beliefs, there should be no difference between ethnicities. So, the economic theory of fertility predicts that patterns of natural selection should be the same in both white and black ethnic groups. Of course, ethnic groups may have different economic characteristics. Also, most of the polygenic scores were created using people of European ancestry. This means they are less predictive for people with African ancestry. If natural selection works via scores' correlation with phenotypes, then we would expect these correlations to be smaller among black people.

Data

The HRS sample focuses on cohorts born between 1920 and 1960, but contains some younger and older participants. I include only male participants born before 1965 and female participants born before 1970, which guarantees that most will have completed their fertility by 2010.

TODO: I think for all these guys, the raevborn question will have been asked by the time they were indeed over 45/40. Double check.

The resulting sample contains 12046 genotyped white participants and 3067 genotyped black participants. Genotyping took place in 2006, 2008 and subsequent years. I use polygenic scores (PGS) precalculated by XXX. I discard obsolete PGS for which there is a newer, more accurate score targeting the same phenotype. I also discard PGS for number of children ever born (but keep scores for age at first birth).

PGS are rescaled to zero mean and unit variance within each ethnic group, so effect sizes are not directly comparable between ethnic groups, but are measured in standard deviations of the within-ethnic-group score. In all regressions using PGS, I control for ten principal components of the DNA array data.

I use realised fertility (number of children ever born) as the key dependent variable. I use this rather than relative lifetime reproductive success for ease of interpretation and because I don't have data for dates when children were born. My focus is on examining patterns of natural selection across PGS, rather than estimating their effects on PGS over time; raw fertility is good enough for this.

The HRS contains weights which match survey respondents to the US population. I use weights for the biomarker subsample (*BIOWGTR in the HRS tracker file). Since half the sample enters the extended interview including biomarker data in each biannual survey, I weight individuals by either their 2010 weight or their 2012 weight. This maximizes the available sample of both black and white respondents, and should approximately match the US population of the sample cohorts between 2010 and 2012.

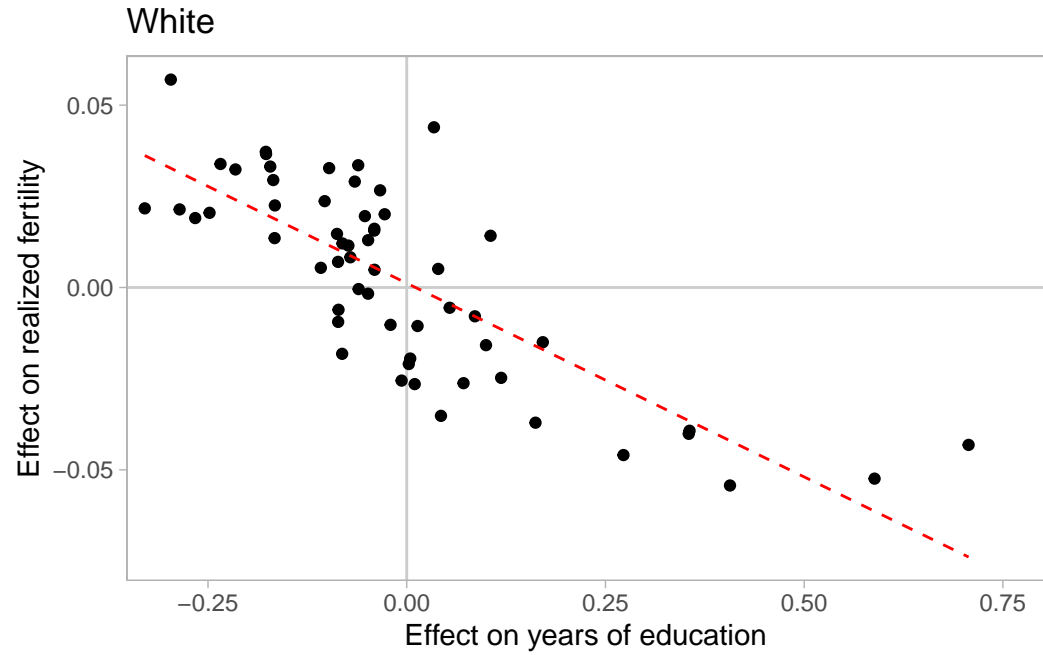
Results

Appendix Figure 7 shows effect sizes of PGS correlations with realised fertility among each ethnic group. Standard errors are large because of the relatively low sample sizes, and no scores are significant at Bonferroni-corrected $p < 0.05/116$. But I am most concerned with looking at patterns across scores rather than judging the significance of individual scores.

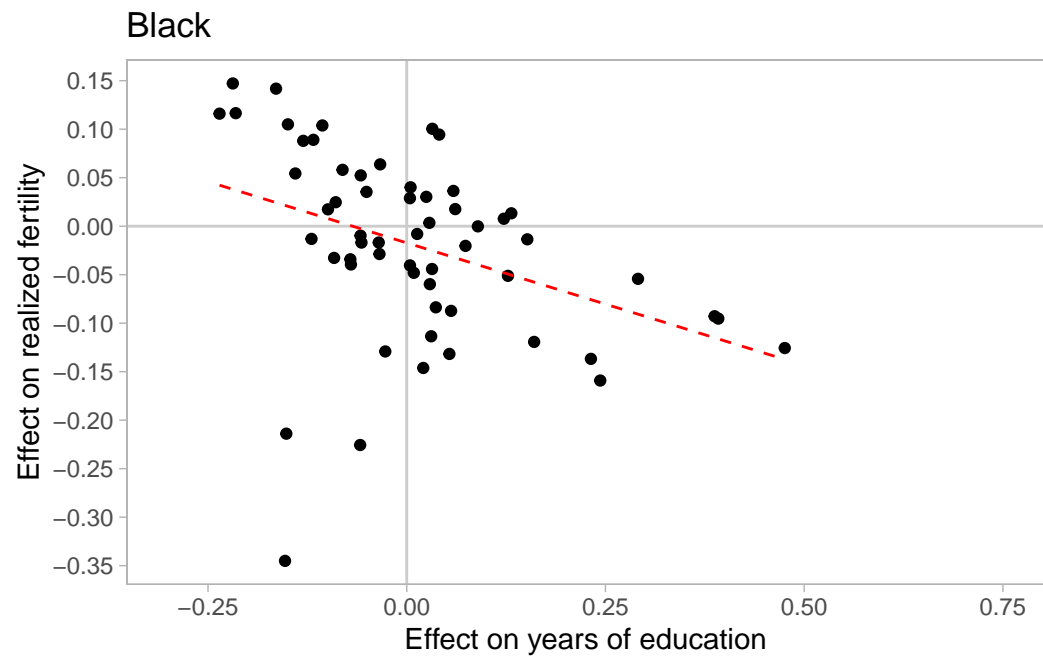
Figure 1 plots each PGS's correlation with fertility against its correlation with educational attainment. The relationship is negative among whites (correlation -0.786, bootstrapped 95% C.I. -0.989 to -0.583) and negative but insignificant among blacks (correlation -0.384, bootstrapped 95% C.I. -0.832 to 0.063). I use bootstraps so as to make inferences from the sample of individuals, rather than the "sample" of different PGS.

The relationship is imprecisely estimated in the smaller black sample, but it is also steeper. Put another way, if we measured PGS in units of "years of associated education", estimated selection effects would be larger among blacks than whites. If this holds in the population, one reason may be that the correlation between phenotypic years of education and fertility is larger for blacks than for whites. In bivariate regressions, an extra year of education reduces fertility by -0.1 for whites and -0.229 for blacks (cf. Goldscheider and Uhlenberg 1969; Johnson 1979; Yang and Morgan 2003). In turn, the simplest explanation for this fact would be point 2 of the theory. Because black people on average have lower incomes and fewer years of education, and are more likely to be single parents, the fertility-wage tradeoff is especially acute for them (cf. Maralani 2013).

I can also examine natural selection in the previous generation, by regressing PGS on respondents' number of siblings (including dead siblings) in 2010. I reweight respondents by the reciprocal of their number of siblings, to account for parents of many siblings being more likely to be a parent of a respondent. Parents of no siblings cannot be included, so effect sizes are not comparable across the generations. Appendix Figure 8 plots effects on number of siblings versus effects on years of



(a)



(b)

Figure 1: Scatterplot of PGS effect sizes on fertility and years of education. Controls include 10 principal components of genetic array data. Dashed lines show linear regressions.

education. The correlation is negative and significant for whites (correlation -0.161, bootstrapped 95% C.I. -0.581 to 0.26). It is positive for blacks (correlation 0.2), although this is driven by two outlier PGS, which are scores for Alzheimer’s with all SNPs included; without these, the correlation is negative.

I next test part 2 of the theory more directly by interacting PGS with measures of education, income and marital status. Education is years of education, split at 12 years. Income is respondent’s mean wage income over all surveys, residualized on a full set of birth year dummies, and median-split.

Figure 2 shows that for both whites and blacks, there is essentially no difference in the distribution of absolute effect sizes between respondents with more and less education. Figure 3, on the other hand, shows that absolute effect sizes appear larger among lower-income respondents, for both whites and blacks. However, in bootstrap tests, the median difference in absolute effect sizes is not significantly different from 0, for either ethnic group. Figure 4 shows absolute effect sizes for respondents who were married in 2010 against all other statuses. Among blacks, effect sizes look similar across marital status, but among whites they are larger for unmarried respondents. Figure 5 splits respondents by age at first birth (median within each ethnicity). There are no large differences between early and late parents. Note that the N is lower here due to missing values.

Discussion

The results here provide qualified support for the economic theory of fertility as an explanation for contemporary natural selection in humans. PGS which predict less education are being selected for, and PGS which predict more education are being selected against. However, there is no evidence that PGS effects on fertility are larger among people with less education, unlike the UK (Hugh-Jones and Abdellaoui 2022). And there’s not enough evidence to show that effects on fertility are larger for lower-income people. Also, the smaller black sample made most tests inconclusive for this population.

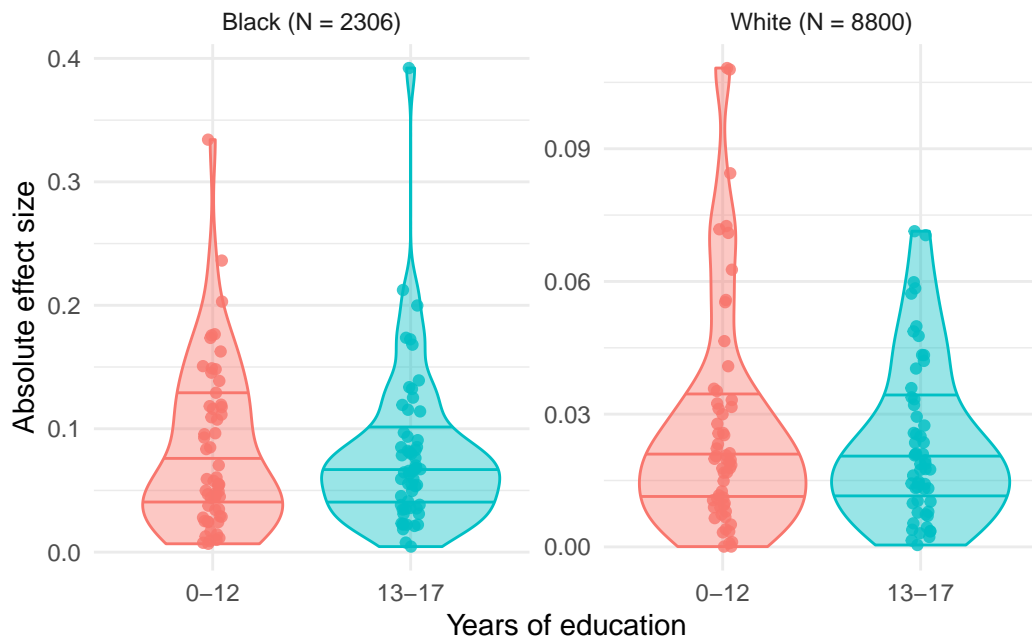


Figure 2: Violin plots of polygenic score effects on fertility by education. Effect sizes are absolute sizes of effects of a one standard deviation (within-ethnicity) change of the PGS on fertility, controlling for 10 principal components of genetic array data, estimated within the low/high education group. Lines show quartiles.

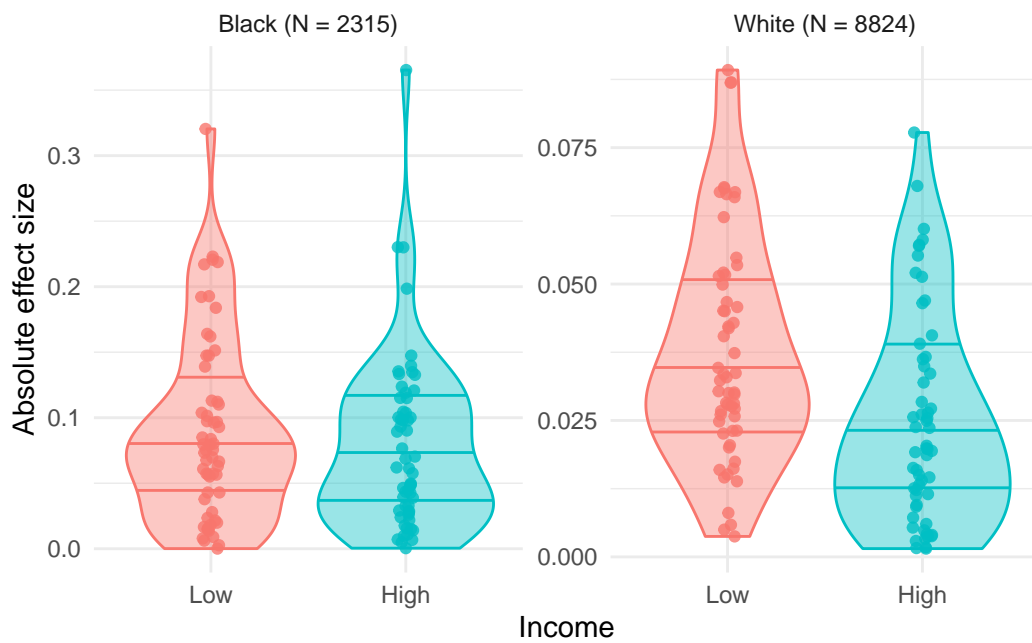


Figure 3: Violin plots of polygenic score effects on fertility by income. Effect sizes are absolute sizes of effects of a one standard deviation (within-ethnicity) change of the PGS on fertility, controlling for 10 principal components of genetic array data, estimated within the low/high income group. Lines show quartiles.



Figure 4: Violin plots of polygenic score effects on fertility by marital status. Effect sizes are absolute sizes of effects of a one standard deviation (within-ethnicity) change of the PGS on fertility, controlling for 10 principal components of genetic array data, estimated within the low/high income group. Lines show quartiles.

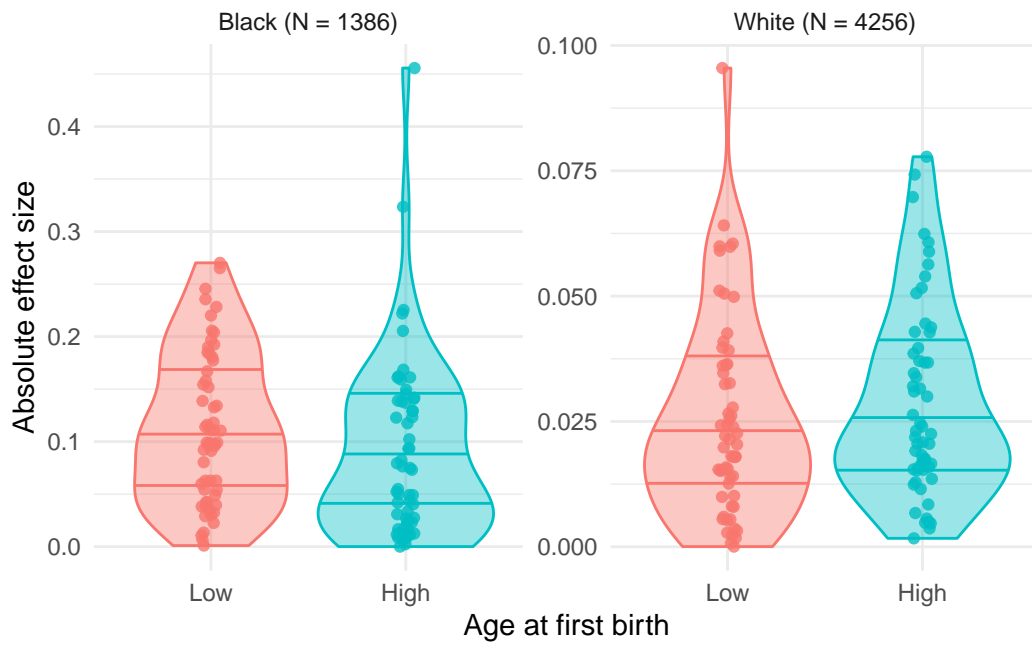


Figure 5: Violin plots of polygenic score effects on fertility by age at first birth (median split). Effect sizes are absolute sizes of effects of a one standard deviation (within-ethnicity) change of the PGS on fertility, controlling for 10 principal components of genetic array data, estimated within the low/high income group. Lines show quartiles.

Appendix

Figures

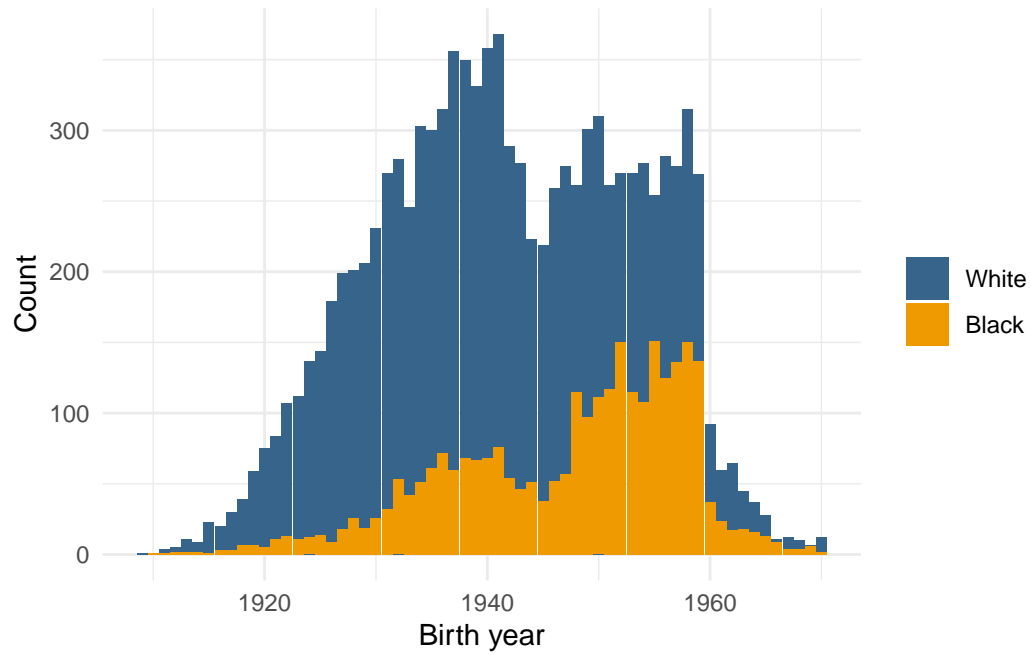


Figure 6: Distribution of birth years for the sample

Effects of polygenic scores on fertility

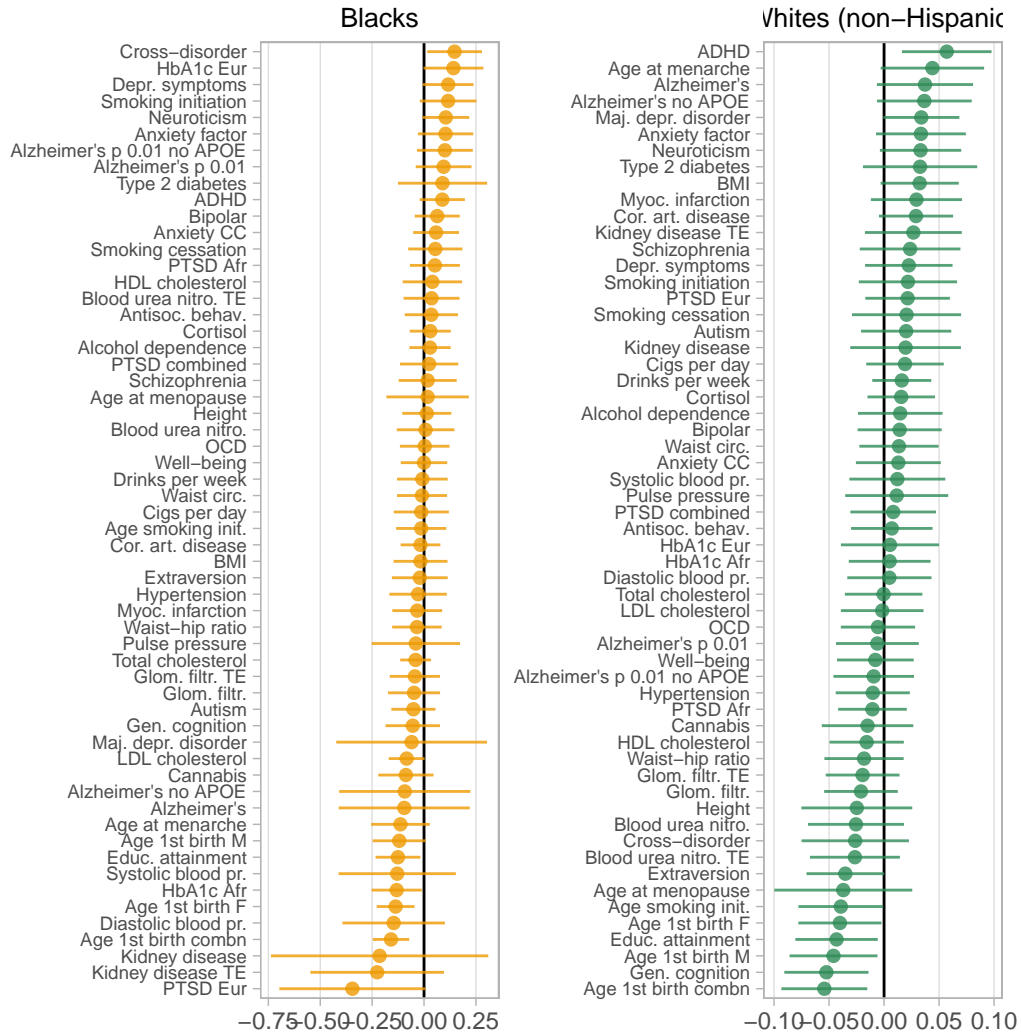
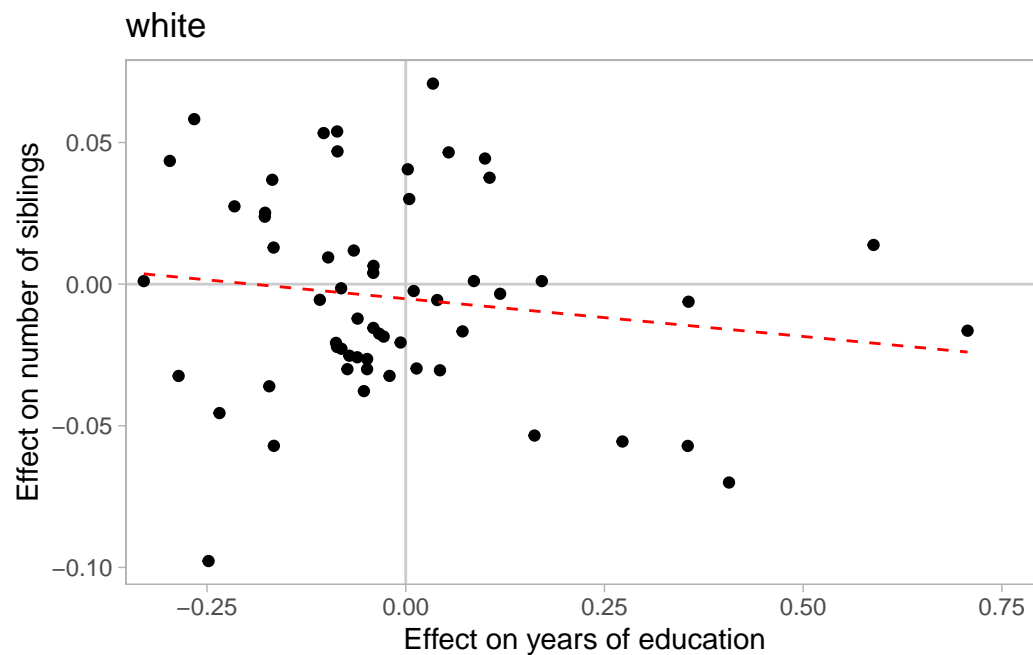
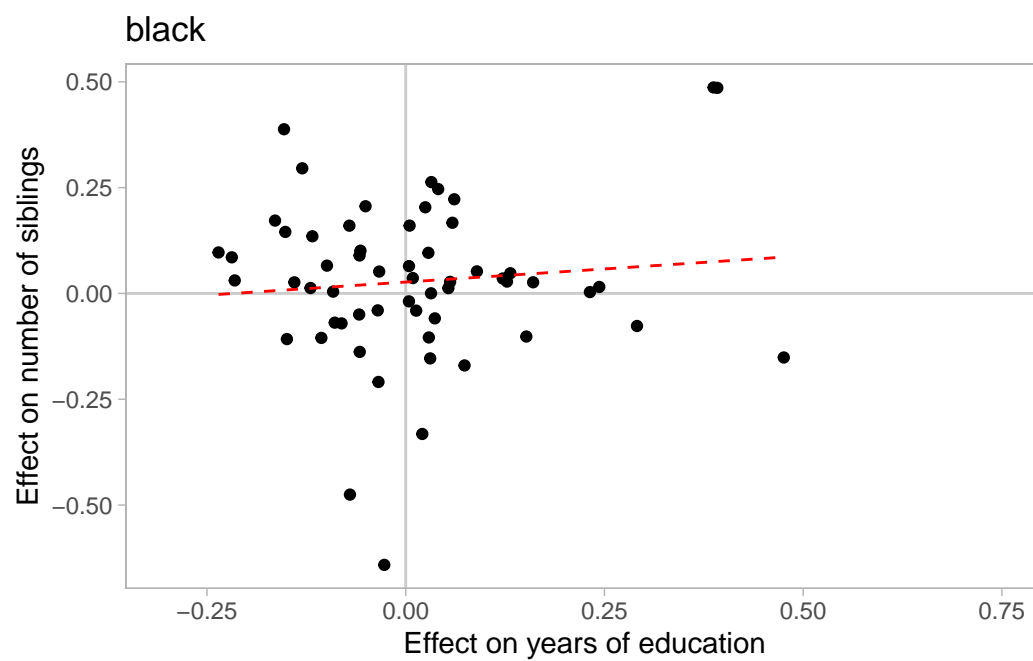


Figure 7: Effects of PGS on realized fertility. Values are partial correlations of a one standard deviation (within ethnicity) change of the PGS with fertility, controlling for 10 principal components of genomic array data. Scales differ between the two charts.



(a)



(b)

Figure 8: Scatterplot of PGS effect sizes on number of live siblings and years of education. Controls include 10 principal components of genetic array data. Dashed lines show linear regressions.

Bibliography

- Becker, Gary S, and Nigel Tomes. 1976. "Child Endowments and the Quantity and Quality of Children." *Journal of Political Economy* 84 (4, Part 2): S143–62.
- Goldscheider, Calvin, and Peter R Uhlenberg. 1969. "Minority Group Status and Fertility." *American Journal of Sociology* 74 (4): 361–72.
- Hugh-Jones, David, and Abdel Abdellaoui. 2022. "Human Capital Mediates Natural Selection in Contemporary Humans." *Behavior Genetics* 52 (4-5): 205–34.
- Johnson, Nan E. 1979. "Minority-Group Status and the Fertility of Black Americans, 1970: A New Look." *American Journal of Sociology* 84 (6): 1386–1400.
- Maralani, Vida. 2013. "The Demography of Social Mobility: Black-White Differences in the Process of Educational Reproduction." *American Journal of Sociology* 118 (6): 1509–58.
- Yang, Yang, and S Philip Morgan. 2003. "How Big Are Educational and Racial Fertility Differentials in the US?" *Social Biology* 50 (3-4): 167–87.