

CAS MA575: Assignment #3

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1.

Statement (d) is true, and following are my reasons:

Through Figure 2.8, we can easily tell that $|y_i - \hat{y}_i|$ in model 1 is smaller compare to that in model 2. Since $RSS = \sum_{i=1}^n (y_i - \hat{y}_i)^2$, RSS of model 1 should be less than RSS of model 2.

And because $RSS + SSreg = SST = \sum_{i=1}^n (y_i - \bar{y})^2$ is fixed, RSS of model 1 is less than RSS of model 2 means that SSreg of model 1 is greater than SSreg of model 2.

2.(a)

Because y_i, y_j is independent when $i \neq j$, which means $cov(y_i, y_j) = 0$ when $i \neq j$.

$$\begin{aligned} cov(y_{n+1}, y_{n+1} + n\hat{\mu}_n) &= cov(y_{n+1}, \sum_{i=1}^{n+1} y_i) \\ &= \sum_{i=1}^n cov(y_{n+1}, y_i) \\ &= cov(y_{n+1}, y_{n+1}) \\ &= \sigma^2 \end{aligned}$$

2.(b)

$$\begin{aligned}
\text{cov}(\hat{\mu}_n, y_{n+1} + n\hat{\mu}_n) &= \text{cov}\left(\frac{1}{n} \sum_{i=1}^n y_i, \sum_{i=1}^{n+1} y_i\right) \\
&= \frac{1}{n} \sum_{i=1}^n \sum_{j=1}^{n+1} \text{cov}(y_i, y_j) \\
&= \frac{1}{n} \sum_{i=1}^n \text{cov}(y_i, y_i) \\
&= \frac{1}{n} n \sigma^2 \\
&= \sigma^2
\end{aligned}$$

2.(c)

$$\begin{aligned}
\text{cov}(y_{n+1} - \hat{\mu}_n, y_{n+1} + n\hat{\mu}_n) &= \text{cov}\left(y_{n+1} - \frac{1}{n} \sum_{i=1}^n y_i, \sum_{i=1}^{n+1} y_i\right) \\
&= \text{cov}\left(y_{n+1}, \sum_{i=1}^{n+1} y_i\right) - \text{cov}\left(\frac{1}{n} \sum_{i=1}^n y_i, \sum_{i=1}^{n+1} y_i\right) \\
&= \sigma^2 - \sigma^2 \\
&= 0
\end{aligned}$$

Since $y_{n+1} - \hat{\mu}_n$ and $y_{n+1} + n\hat{\mu}_n$ are normal distribution, $\text{cov}(y_{n+1} - \hat{\mu}_n, y_{n+1} + n\hat{\mu}_n) = 0$ means that they are independent.

3.(a)

Because $y_i, i = 1, \dots, n$ is $N(0, \sigma^2)$, $\frac{y_i}{\sigma}$ is $N(0, 1)$. Therefore $\sigma^{-2} \sum_{i=1}^m y_i^2 = \sum_{i=1}^m \left(\frac{y_i}{\sigma}\right)^2$ is square sum of m independent $N(0, 1)$, which is χ_m^2 , for any $1 \leq m \leq n$.

3.(b)

First of all, $\frac{y_{m+1}}{\sigma}$ is $N(0, 1)$, for $1 \leq m \leq n-1$.

And $\sum_{i=1}^m \left(\frac{y_i}{\sigma}\right)^2$ is $\chi_m^2 \rightarrow \frac{\frac{y_{m+1}}{\sigma}}{\sqrt{\frac{\sum_{i=1}^m y_i^2}{m\sigma^2}}}$ is t_m .

Which equals $\frac{y_{m+1}}{\sqrt{m^{-1} \sum_{i=1}^m y_i^2}}$ is t_m , for any $1 \leq m \leq n-1$.

3.(c)

Because $\sum_{i=m+1}^n (\frac{y_i}{\sigma})^2$ is χ_{n-m}^2 , $\sum_{i=1}^m (\frac{y_i}{\sigma})^2$ is χ_m^2 .

And since $y_i, i = 1, \dots, n$ are independent, $\sum_{i=m+1}^n (\frac{y_i}{\sigma})^2$ and $\sum_{i=1}^m (\frac{y_i}{\sigma})^2$ are independent.

Therefore $\frac{\sum_{i=m+1}^n (\frac{y_i}{\sigma})^2 / (n-m)}{\sum_{i=1}^m (\frac{y_i}{\sigma})^2 / m} = \frac{(n-m)^{-1} \sum_{i=m+1}^n y_i^2}{m^{-1} \sum_{i=1}^m y_i^2}$ is $F_{n-m, m}$.