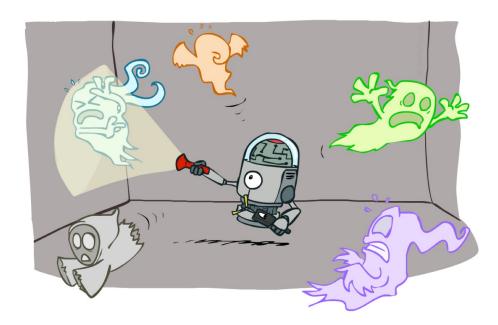
Artificial Intelligence - INFOF311

HMM, particle filters and DBN

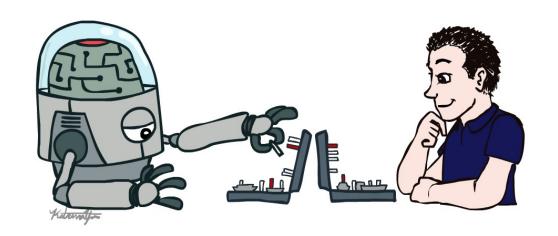


Instructor: Tom Lenaerts

Acknowledgement

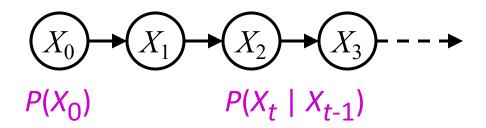
We thank Stuart Russell for his generosity in allowing us to use the slide set of the UC Berkeley Course CS188, Introduction to Artificial Intelligence. These slides were created by Dan Klein, Pieter Abbeel and Anca Dragan for CS188 Intro to AI at UC Berkeley. All CS188 materials are available at http://ai.berkeley.edu.





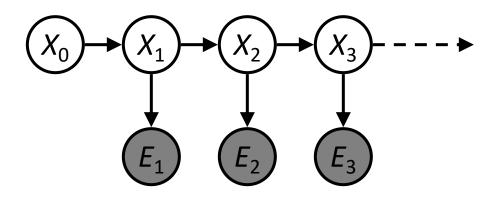
The slides for INFOF311 are slightly modified versions of the slides of the spring and summer CS188 sessions in 2021 and 2022

Markov Chains



- Stationarity assumption: transition probabilities are the same at all times
- Markov assumption: "future is independent of the past given the present"
- Mini-Forward Algorithm: $P(X_t) = \sum_{X_{t-1}} P(X_{t-1} = X_{t-1}) P(X_t | X_{t-1} = X_{t-1})$
- Equivalently: $P_{t+1} = T^T P_t$
- Stationary Distribution: $P_{\infty} = T^{T} P_{\infty}$
- Stationary distribution does not depend on the starting distribution

Hidden Markov Models



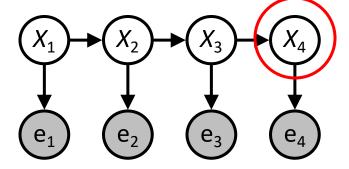
- Sensor models are the same at all times
- Current evidence is independent of everything else given the current state
- Inference tasks; Filtering, prediction, ...

Inference tasks

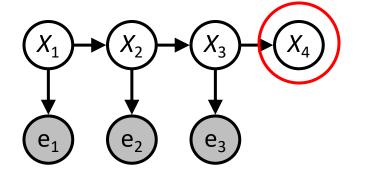
- Filtering: $P(X_t | e_{1:t})$
 - belief state—input to the decision process of a rational agent
- **Prediction**: $P(X_{t+k}|e_{1:t})$ for k > 0
 - evaluation of possible action sequences; like filtering without the evidence
- Smoothing: $P(X_k | e_{1:t})$ for $0 \le k < t$
 - better estimate of past states, essential for learning
- Most likely explanation: $arg max_{x_{1:t}} P(x_{1:t} | e_{1:t})$
 - speech recognition, decoding with a noisy channel

Inference tasks

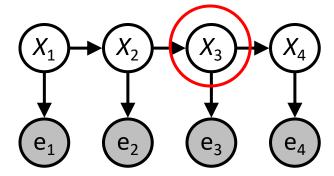
Filtering: $P(X_t | e_{1:t})$



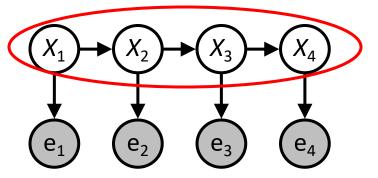
Prediction: $P(X_{t+k} | e_{1:t})$



Smoothing: $P(X_k | e_{1:t})$, k<t



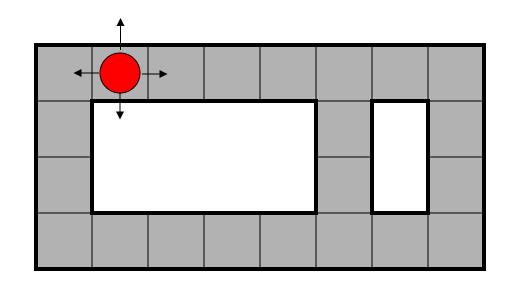
Explanation: $P(X_{1:t} | e_{1:t})$

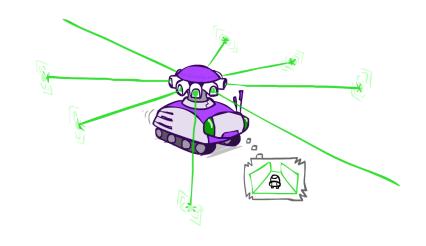


Filtering / Monitoring

- Filtering, or monitoring, or state estimation, is the task of maintaining the distribution $f_{1:t} = P(X_t | e_{1:t})$ over time
- We start with f_0 in an initial setting, usually uniform
- Filtering is a fundamental task in engineering and science
- The Kalman filter (continuous variables, linear dynamics, Gaussian noise) was invented in 1960 and used for trajectory estimation in the Apollo program; core ideas used by Gauss for planetary observations; 1110000 papers on Google Scholar (spring 2023).

Example from Michael Pfeiffer

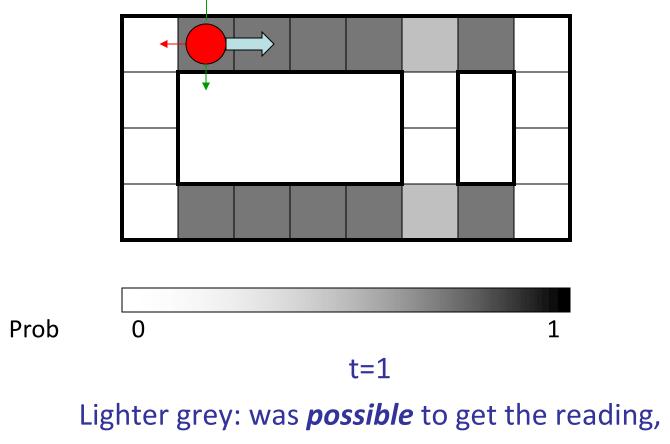


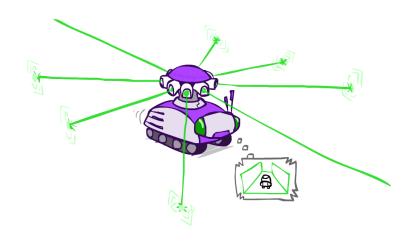




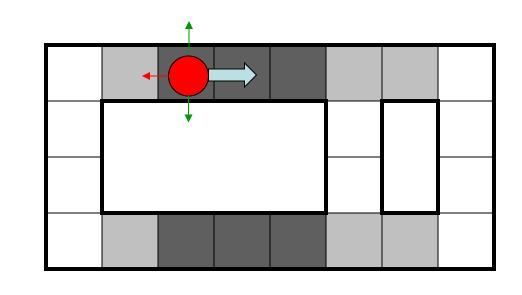
Sensor model: four bits for wall/no-wall in each direction, never more than 1 mistake

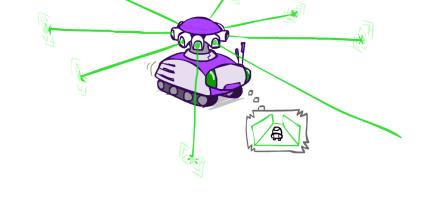
Transition model: action may fail with small prob.



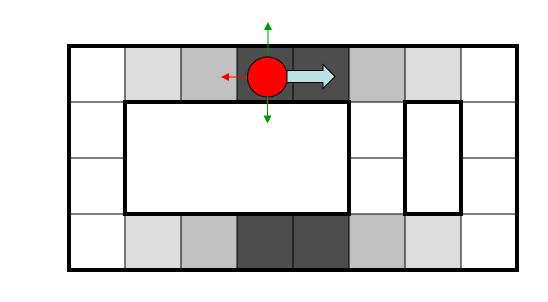


but *less likely* (required 1 mistake)

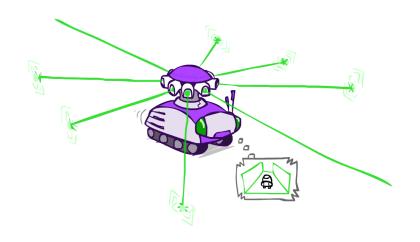


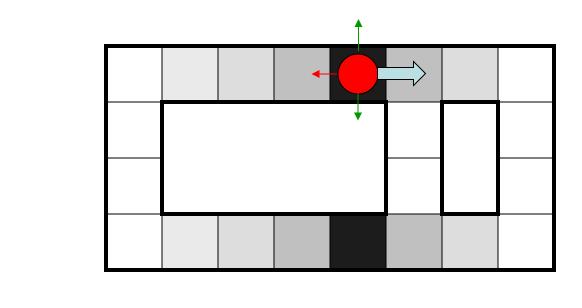


Prob 0 1





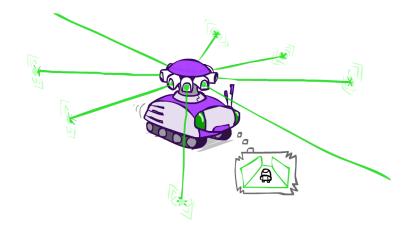


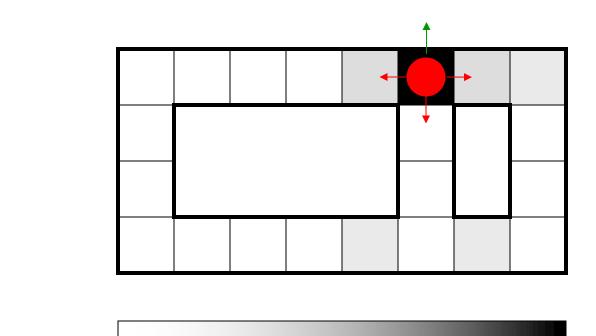




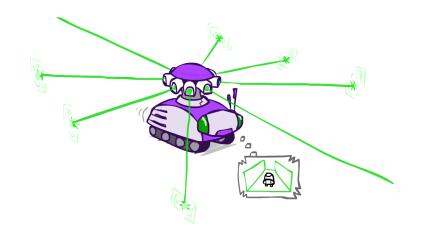


Prob





Prob



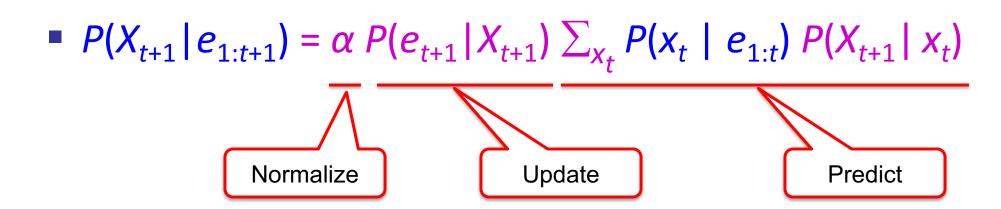
Filtering algorithm

• Aim: devise a recursive filtering algorithm of the form

$$P(X_{t+1}|e_{1:t+1}) = g(e_{t+1}, P(X_t|e_{1:t}))$$

Apply Bayes' rule Apply conditional independence $P(X_{t+1} | e_{1:t+1}) = P(X_{t+1} | e_{1:t}, e_{t+1})$ $= \alpha P(e_{t+1} | X_{t+1}, e_{1:t}) P(e_{t+1} | X_{t+1}, e_{1:t})$ Condition on X_t $= \alpha \, P(e_{t+1} \, | \, X_{t+1}) \, P(X_{t+1} \, | \, e_{1:t})$ Apply conditional independence $\sum_{X_t} P(X_{t+1} | e_{1:t}) P(X_{t+1} | X_t, e_{1:t})$ Normalize $|e_{t+1}|$ Update Predict $(X_{t+1} | X_t)$

Filtering algorithm



- $f_{1:t+1} = FORWARD(f_{1:t}, e_{t+1})$
- Cost per time step: $O(|X|^2)$ where |X| is the number of states
- Time and space costs are constant, independent of t
- $O(|X|^2)$ is infeasible for models with many state variables
- We get to invent really cool approximate filtering algorithms

And the same thing in linear algebra

- Transition matrix T, observation matrix O_t
 - Observation matrix has state likelihoods for E_t along diagonal

■ E.g., for
$$U_1$$
 = true, $O_1 = \begin{pmatrix} 0.2 & 0 \\ 0 & 0.9 \end{pmatrix}$

- Filtering algorithm becomes

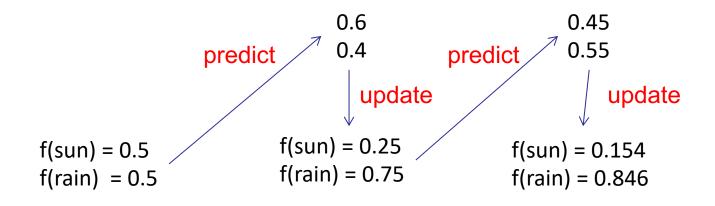
X _{t-1}	P(X _t X _{t-1})	
	sun	rain
sun	0.9	0.1
rain	0.3	0.7

W_{t}	$P(U_t W_t)$	
	true	false
sun	0.2	0.8
rain	0.9	0.1

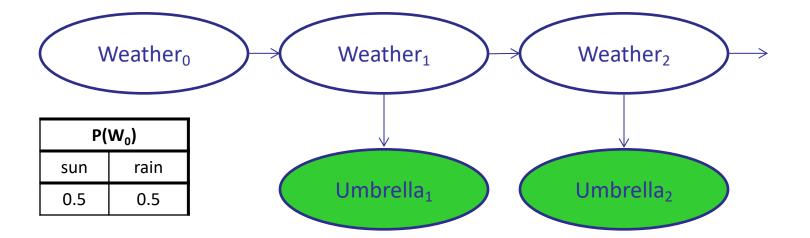
Example: Weather HMM







\mathbf{W}_{t-1}	$P(W_t W_{t-1})$	
	sun	rain
sun	0.9	0.1
rain	0.3	0.7



\mathbf{W}_{t}	P(U _t W _t)	
	true	false
sun	0.2	0.8
rain	0.9	0.1

Pacman – Hunting Invisible Ghosts with Sonar



Video of Demo Pacman – Sonar (with beliefs)



Most Likely Explanation

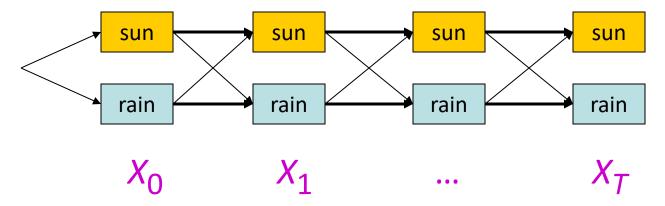


Inference tasks

- Filtering: $P(X_t|e_{1:t})$
 - belief state—input to the decision process of a rational agent
- **Prediction**: $P(X_{t+k}|e_{1:t})$ for k > 0
 - evaluation of possible action sequences; like filtering without the evidence
- Smoothing: $P(X_k | e_{1:t})$ for $0 \le k < t$
 - better estimate of past states, essential for learning
- Most likely explanation: $arg max_{x_{1:t}} P(x_{1:t} | e_{1:t})$
 - speech recognition, decoding with a noisy channel

Most likely explanation = most probable path

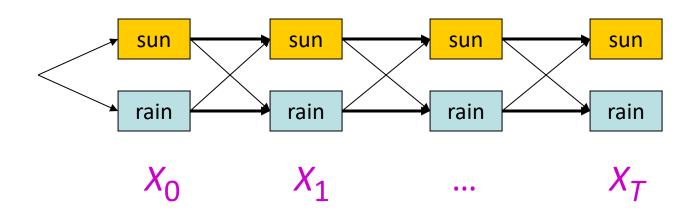
State trellis: graph of states and transitions over time



 $arg \max_{x_{1:t}} P(x_{1:t} | e_{1:t})$ $= arg \max_{x_{1:t}} \alpha P(x_{1:t}, e_{1:t})$ $= arg \max_{x_{1:t}} P(x_{1:t}, e_{1:t})$ $= arg \max_{x_{1:t}} P(x_0) \prod_{t} P(x_t | x_{t-1}) P(e_t | x_t)$

- Each arc represents some transition $x_{t-1} \rightarrow x_t$
- Each arc has weight $P(x_t \mid x_{t-1}) P(e_t \mid x_t)$ (arcs to initial states have weight $P(x_0)$)
- The product of weights on a path is proportional to that state sequence's probability
- Forward algorithm computes sums of paths, Viterbi algorithm computes best paths

Forward / Viterbi algorithms



Forward Algorithm (sum)

For each state at time *t*, keep track of the *total probability of all paths* to it

$$f_{1:t+1} = FORWARD(f_{1:t}, e_{t+1})$$

= $\alpha P(e_{t+1}|X_{t+1}) \sum_{X_t} P(X_{t+1}|X_t) f_{1:t}$

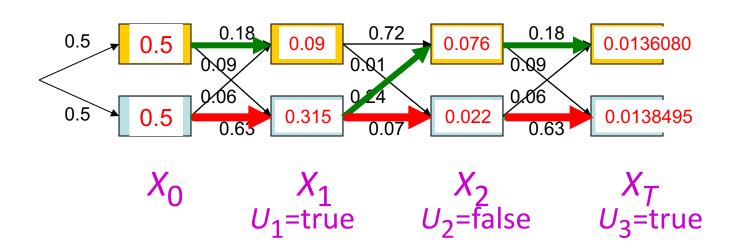
Viterbi Algorithm (max)

For each state at time *t*, keep track of the *maximum probability of any path* to it

$$m_{1:t+1} = VITERBI(m_{1:t}, e_{t+1})$$

= $P(e_{t+1}|X_{t+1}) \max_{X_t} P(X_{t+1}|X_t) m_{1:t}$

Viterbi algorithm contd.



W_{t-1}	$P(W_t W_{t-1})$	
	sun	rain
sun	0.9	0.1
rain	0.3	0.7

\mathbf{W}_{t}	$P(U_t W_t)$	
	true	false
sun	0.2	0.8
rain	0.9	0.1

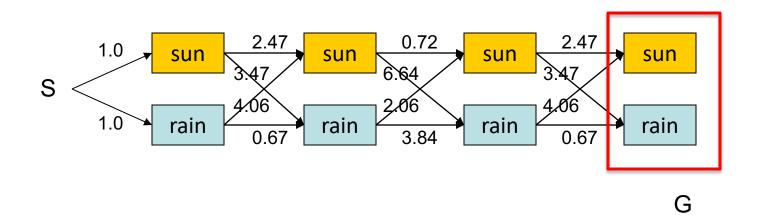
Time complexity?
O(|X|² T)

Space complexity?

O(|X|T)

Number of paths? O(|X|^T)

Viterbi in negative log space



W _t	P(U)	false
W_t	P(U	$_{t} W_{t})$
	$P(U_t W_t)$	
rain	0.3	0.7
sun	0.9	0.1

sun

 $P(W_t|W_{t-1})$

rain

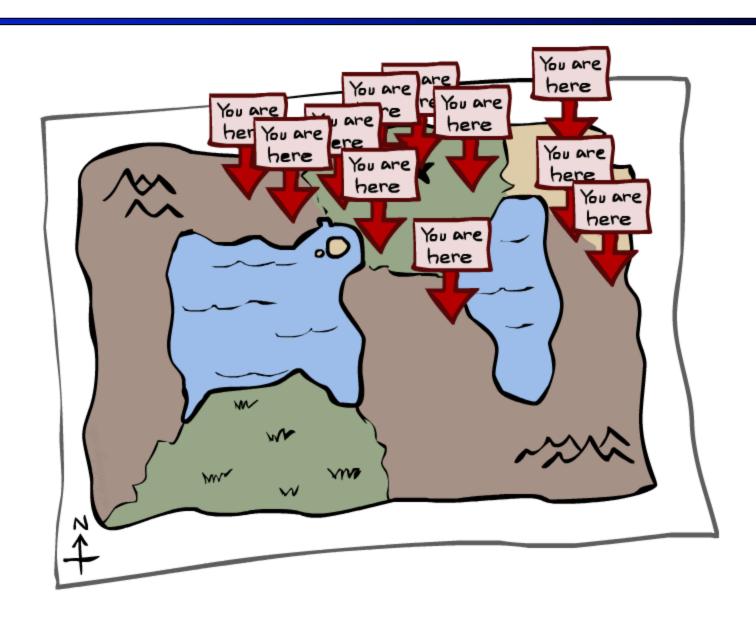
argmax of product of probabilities

- = argmin of sum of negative log probabilities
- = minimum-cost path

Viterbi is essentially breadth-first graph search What about A*?

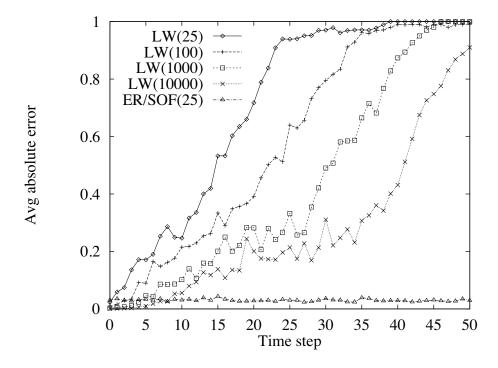
_		
W _t	P(U	_t W _t)
	true	false
sun	0.2	0.8
rain	0.9	0.1

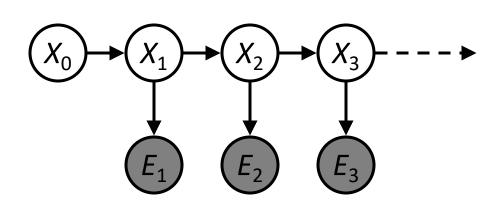
Particle Filtering



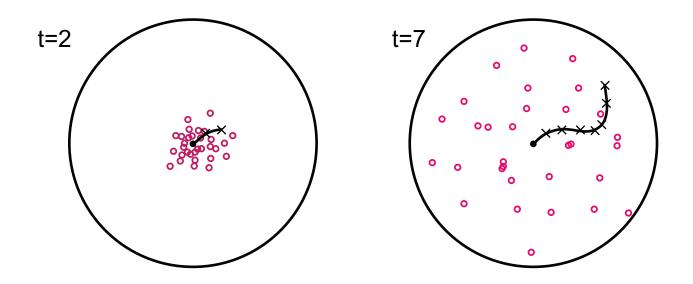
We need a new algorithm!

- When size of the state space is large, exact inference becomes infeasible
- Likelihood weighting also fails completely number of samples needed grows exponentially with T





Particle Filtering: Resample



- The problem of likelihood weighting: sample state trajectories go off into low-probability regions; too few "reasonable" samples
- Solution: kill the bad ones, make more of the good ones
- This way the population of samples stays in the high-probability region

Particle Filtering

- Solution: approximate inference
 - Track samples of X, not all values
 - Samples are called particles
 - Time per step is linear in the number of samples
 - But: number needed may be large
 - In memory: list of particles, not states

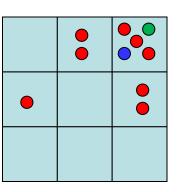
0.0	0.1	0.0
0.0	0.0	0.2
0.0	0.2	0.5



	•
• •	

Representation: Particles

- Our representation of P(X) is now a list of $N \ll |X|$ particles
- P(x) approximated by number of particles with value x
 - So, many x may have P(x) = 0!
 - More particles => more accuracy
- For now, all particles have a weight of 1



Particles:

(3,3)

(2,3)

(3,3)

(3,2)

(3,3)

(3,2)

(1,2)

(3,3)

(3,3)

(2,3)

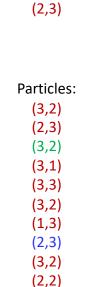
Particle Filtering: Prediction Step

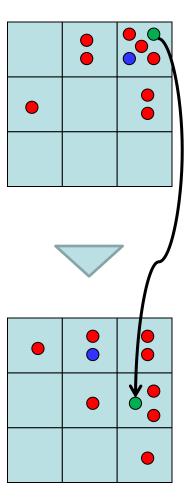
 Each particle is moved by sampling its next position from the transition model

$$x' = \text{sample}(P(X'|x))$$

- This is like prior sampling samples' frequencies reflect the transition probabilities
- Here, most samples move clockwise, but some move in another direction or stay in place
- This captures the passage of time
 - If enough samples, close to exact values before and after (consistent)

Particles: (3,3) (2,3) (3,3) (3,2) (3,3) (3,2) (1,2) (3,3) (3,3)





Particle Filtering: Update step

- After observing e_{t+1} :
 - As in likelihood weighting, weight each sample based on the evidence
 - $w^{(j)} = P(e_{t+1} | X_{t+1}^{(j)})$
 - Normalize the weights: particles that fit the data better get higher weights, others get lower weights

Particles:

(3,2)

(2,3)

(3,2)

(3,1)

(3,3)

(3,2)

(1,3)

(2,3)

(3,2)

(2,2)

Particles:

(3,2) w=.9

(2,3) w=.2

(3,2) w=.9

(3,1) w=.4

(3,3) w=.4

(3,2) w=.9

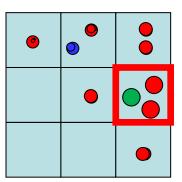
(1,3) w=.1

(2,3) w=.2

(3,2) w=.9

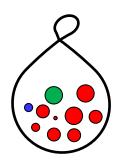
(2,2) w=.4





Particle Filtering: Resample

- Rather than tracking weighted samples, we *resample*
- N times, we choose from our weighted sample distribution (i.e., draw with replacement)
- Now the update is complete for this time step, continue with the next one (with weights reset to 1)

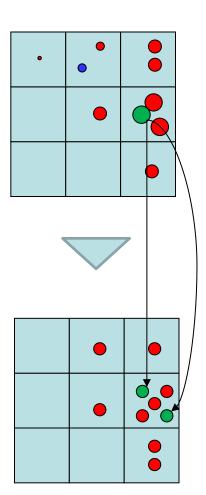


Particles:

- (3,2) w=.9
- (2,3) w=.2
- (3,2) w=.9
- (3,1) w=.4
- (3,3) w=.4
- (3,2) w=.9
- (1,3) w=.1
- (2,3) w=.2
- (3,2) w=.9
- (2,2) w=.4

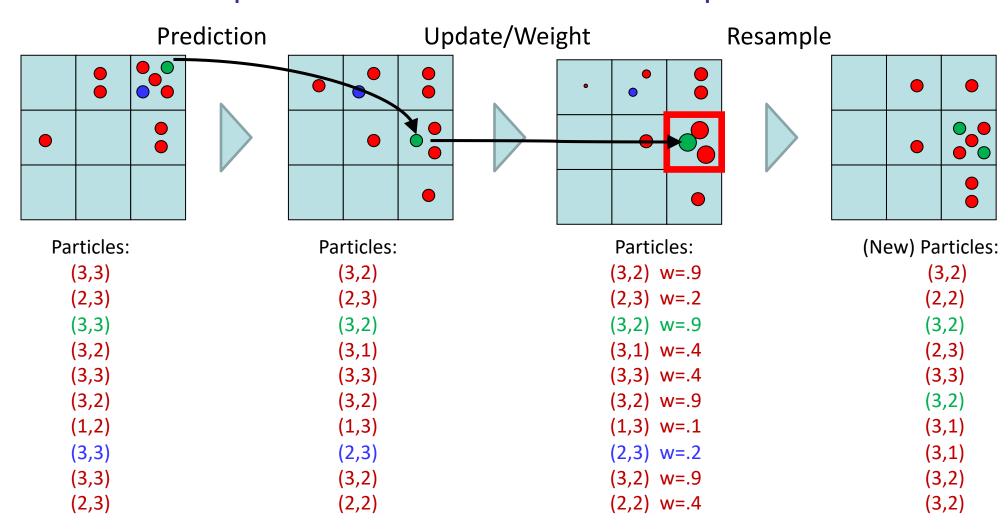
(New) Particles:

- (3,2)
- (2,2)
- (3,2)
- (2,3)
- (3,3)
- (3,2)
- (3,1)
- (3,1)
- (3,2)
- (3,2)



Summary: Particle Filtering

Particles: track samples of states rather than an explicit distribution



Consistency: see proof in AIMA Ch. 14

[Demos: ghostbusters particle filtering (L15D3,4,5)]

Particle Filter Localization (Sonar)

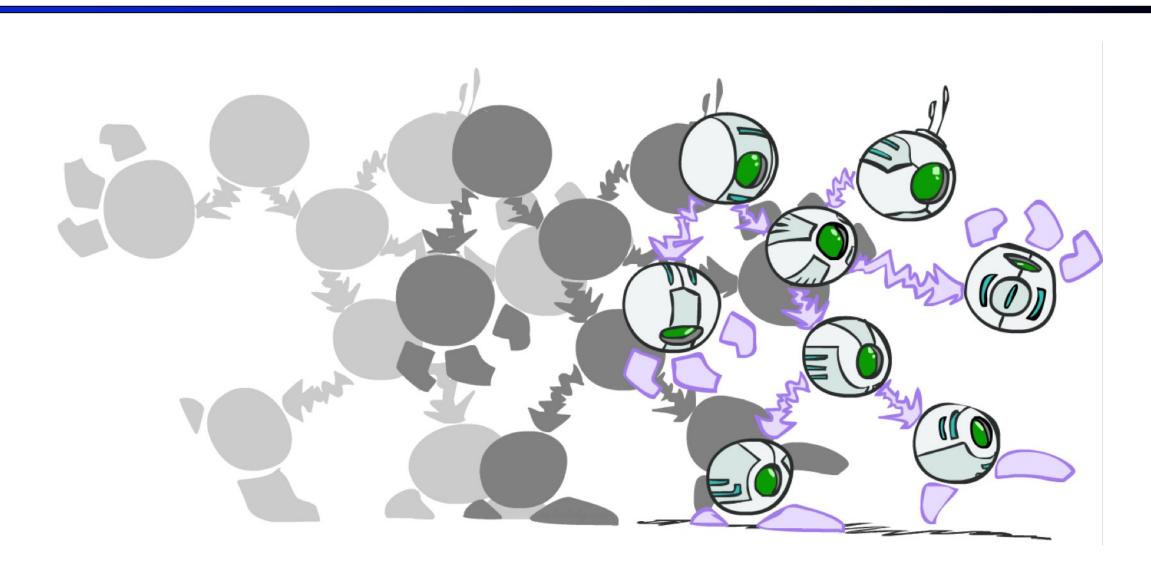


Particle Filter SLAM

Simultaneous localization and mapping

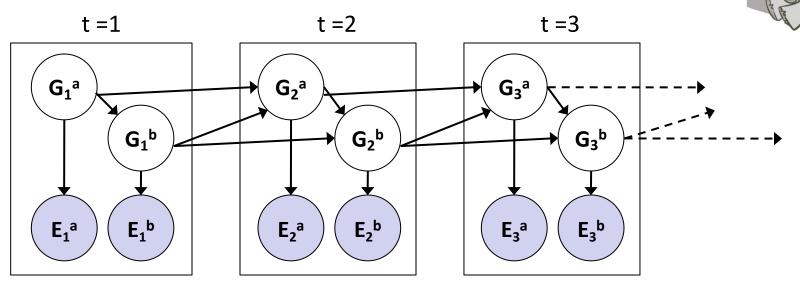


Dynamic Bayes Nets



Dynamic Bayes Nets (DBNs)

- We want to track multiple variables over time, using multiple sources of evidence
- Idea: Repeat a fixed Bayes net structure at each time
- Variables at time t can have parents at time t-1

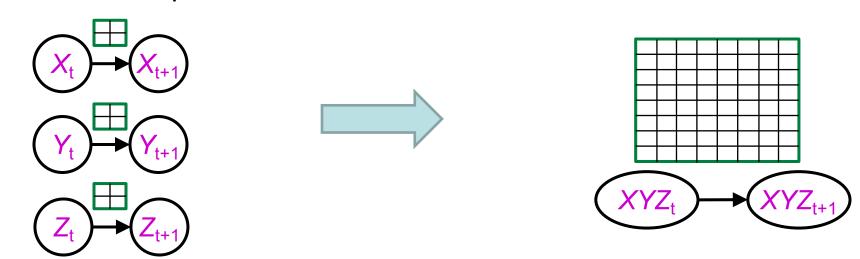






DBNs and **HMMs**

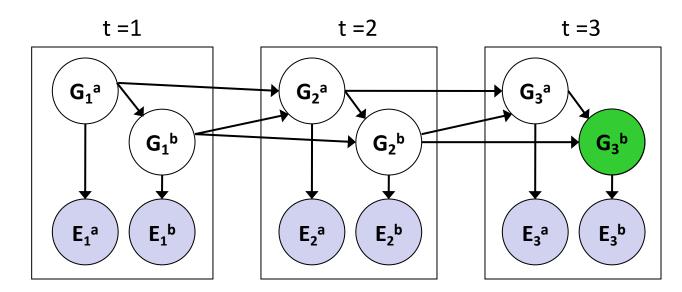
- Every HMM is a single-variable DBN
- Every discrete DBN is an HMM
 - HMM state is Cartesian product of DBN state variables



- Sparse dependencies => exponentially fewer parameters in DBN
 - E.g., 20 state variables, 3 parents each; DBN has $20 \times 2^3 = 160$ parameters, HMM has $2^{20} \times 2^{20} = 10^{12}$ parameters

Exact Inference in DBNs

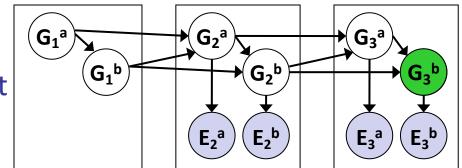
- Variable elimination applies to dynamic Bayes nets
- Offline: "unroll" the network for T time steps, then eliminate variables to find $P(X_T | e_{1:T})$



- Online: eliminate all variables from the previous time step; store factors for current time only
- Problem: largest factor contains all variables for current time (plus a few more)

DBN Particle Filters

- A particle is a complete sample for a time step
- Initialize: Generate prior samples for the t=1 Bayes net
 - Example particle: $G_1^a = (3,3) G_1^b = (5,3)$



- Elapse time: Sample a successor for each particle
 - Example successor: $G_2^a = (2,3) G_2^b = (6,3)$
- Observe: Weight each <u>entire</u> sample by the likelihood of the evidence conditioned on the sample
 - Likelihood: $P(E_2^a | G_2^a) * P(E_2^b | G_2^b)$
- Resample: Select prior samples (tuples of values) in proportion to their likelihood