## STA3010 Regression Analysis

Feng YIN

The Chinese University of Hong Kong (Shenzhen) yinfeng@cuhk.edu.cn

March 16, 2020

### Overview

1 Empirical Variance-Stabilizing Transformations

2 Analytical Variance-Stabilizing Transformations

Transformations to Linearize A Model

2 / 22

Feng YIN (CUHKSZ) Lecture 7 March 16, 2020

## Variance-Stabilizing

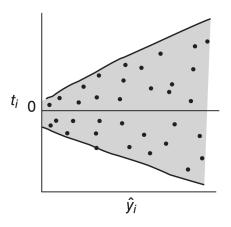
Thus far, we have assumed implicitly that:

- The model errors have zero mean and constant variance and they are uncorrelated.
- The model errors have a normal distribution; This assumption is made in order to conduct hypothesis tests and construct CI. Under this assumption, the errors are independent.

Plots of residuals can be used for detecting violations of these basic regression assumptions. But how to remedy?

Feng YIN (CUHKSZ)

A common reason for violating the "constant variance" assumption lies in that the variance of the output y is functionally related to its mean, E(y).



#### State-of-the-art variance-stabilizing transformations:

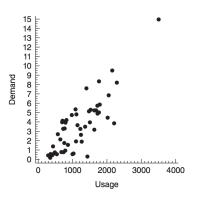
| Relationship of $\sigma^2$ to $E(y)$ | Transformation                                                               |  |
|--------------------------------------|------------------------------------------------------------------------------|--|
| $\sigma^2 \propto \text{constant}$   | y' = y (no transformation)                                                   |  |
| $\sigma^2 \propto E(y)$              | $y' = \sqrt{y}$ (square root; Poisson data)                                  |  |
| $\sigma^2 \propto E(y)[1 - E(y)]$    | $y' = \sin^{-1}(\sqrt{y})$ (arcsin; binomial proportions $0 \le y_i \le 1$ ) |  |
| $\sigma^2 \propto [E(y)]^2$          | $y' = \ln(y)(\log)$                                                          |  |
| $\sigma^2 \propto [E(y)]^3$          | $y' = y^{-1/2}$ (reciprocal square root)                                     |  |
| $\sigma^2 \propto [E(y)]^4$          | $y' = y^{-1}(reciprocal)$                                                    |  |

In practice, a transformation may be selected empirically according to the residual plots.

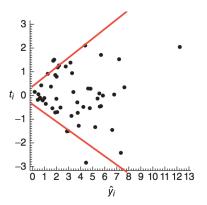
Feng YIN (CUHKSZ) Lecture 7 March 16, 2020 6 / 2:

Textbook Example: An electric utility is interested in developing a model to relate the peak-hour demand, y, to the total energy usage during the month, x.

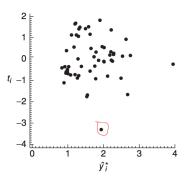
Data set for 53 residential customers collected for one month is shown below:



After applying the simple linear least-squares (LS) fit, we plot the the R-student residuals versus the fitted values  $\hat{y}_i$ , i = 1, 2, 3..., n.



- Apply  $y^* = \sqrt{y}$  as a variance stabilizing transformation and re-do LS fitting.
- The new R-student residuals are plotted against the fitted values.



### Analytical Box-Cox Transformation

An important class of transformations for stabilizing the variance is the power transformation/Box-Cox transformation  $y^{\lambda}$ , where  $\lambda$  is a parameter to be determined.

Due to the difficulties of using  $y^{\lambda}$  directly, it is modified to be:

$$y^{(\lambda)} = \begin{cases} \frac{y^{\lambda} - 1}{\lambda \dot{y}^{\lambda - 1}}, & \lambda \neq 0\\ \dot{y} \ln y, & \lambda = 0 \end{cases}$$
 (1)

where  $\dot{y}$  is defined to be the geometric mean of the outputs, i.e.,  $\dot{y} = \left(\prod_{i=1}^n y_i\right)^{1/n}$ , which is a scale factor that make the residual sum of squares for models with different values of  $\lambda$  comparable.

After the Box-Cox transformation, we have the new regression model:

$$\mathbf{y}^{(\lambda)} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}'. \tag{2}$$

Feng YIN (CUHKSZ) Lecture 7 March 16, 2020 10 / 22

### Analytical Box-Cox Transformation

#### Complete procedure:

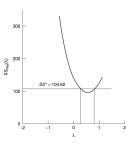
- Select a few (discrete)  $\lambda$  values, for instance 0:0.5:5
- 2 For each  $\lambda$  value, do:
  - Transform the original outputs  ${\bf y}$  according to Box-Cox transformation, and get transformed output  ${\bf y}^{(\lambda)}$
  - ② Fit the regression model  $\mathbf{y}^{(\lambda)} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}'$
  - **3** Compute the residual sum of squares  $SS_{Res}(\lambda)$ .
- **9** Plot  $SS_{Res}(\lambda)$  versus  $\lambda$  and find out the  $\lambda$  value that yields the smallest  $SS_{Res}$  value.
- **①** Optional: In the vicinity of the found  $\lambda$  value, repeat the above procedure for refined result.

### Analytical Box-Cox Transformation

Textbook Example: An electric utility is interested in developing a model to relate the peak-hour demand, y, to the total energy usage during the month, x.

The values of  $SS_{Res}(\lambda)$  versus  $\lambda$  are shown below:

| λ     | $SS_{Res}(\lambda)$ |  |
|-------|---------------------|--|
| -2    | 34,101.0381         |  |
| -1    | 986.0423            |  |
| -0.5  | 291.5834            |  |
| 0     | 134.0940            |  |
| 0.125 | 118,1982            |  |
| 0.25  | 107.2057            |  |
| 0.375 | 100.2561            |  |
| 0.5   | 96,9495             |  |
| 0.625 | 97.2889             |  |
| 0.75  | 101.6869            |  |
| 1     | 126.8660            |  |
| 2     | 1,275,5555          |  |



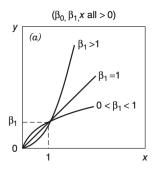
Square-root transformation ( $\lambda = 0.5$ ) is very close to the optimum choice.

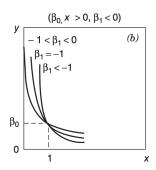
Sometimes, a nonlinear regression model can be linearized through appropriate transformation. Such nonlinear models are called intrinsically or transformably linear.

#### Some examples:

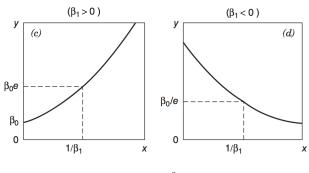
TABLE 5.4 Linearizable Functions and Corresponding Linear Form

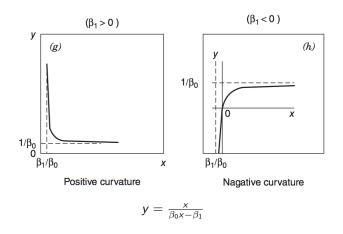
| Figure                                             | Linearizable Function                                  | Transformation                          | Linear Form                                                        |
|----------------------------------------------------|--------------------------------------------------------|-----------------------------------------|--------------------------------------------------------------------|
| 5.4 <i>a</i> , <i>b</i><br>5.4 <i>c</i> , <i>d</i> | $y = \beta_0 x^{\beta_1} $ $y = \beta_0 e^{\beta_1 x}$ | $y' = \log y, x' = \log x$ $y' = \ln y$ | $y' = \log \beta_0 + \beta_1 x'$<br>$y' = \ln \beta_0 + \beta_1 x$ |
| 5.4g, h                                            | $y = \frac{x}{\beta_0 x - \beta_1}$                    | $y' = \frac{1}{y}, x' = \frac{1}{x}$    | $y' = \beta_0 - \beta_1 x'$                                        |





$$y = \beta_0 x^{\beta_1}$$





#### One famous transformation:

$$y = \beta_0 e^{\beta_1 x} \varepsilon, \tag{3}$$

where we have multiplicative random error instead of additive random error.

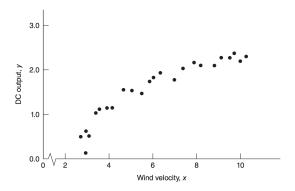
This function is intrinsically linear since it can be transformed to a straight line by logarithmic transformation as

$$\ln y = \ln \beta_0 + \beta_1 x + \ln \varepsilon. \tag{4}$$

If the transformed error terms  $\ln \varepsilon$  are normally and independently distributed with mean zero and variance  $\sigma^2$ , then the multiplicative error terms  $\varepsilon$  in the original model are i.i.d. log-normal distributed.

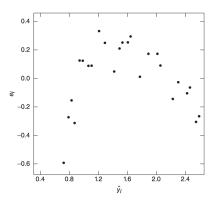
4 D > 4 B > 4 E > E + 9 Q (\*)

Textbook example: A research engineer is investigating the use of a windmill to generate electricity. He has collected data on the DC output from his windmill and the corresponding wind velocity.

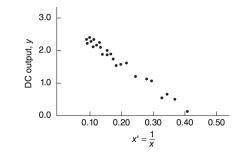


18 / 22

- Fitting a straight-line model to the data, yields  $\hat{y} = 0.13 + 0.24x$ , moreover  $R^2 = 0.8745$  and  $F_0 = 160.26$ .
- ② A plot of the residuals versus  $\hat{y}_i$  is shown below.

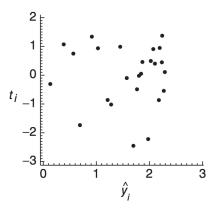


- **1** A more reasonable model would be  $y = \beta_0 + \beta_1(\frac{1}{x}) + \varepsilon$ .
- 2 Refit the model using LS yields  $\hat{y} = 2.98 6.93x'$ .
- The summary statistics for the new model are  $R^2 = 0.98$ ,  $MS_{Res} = 0.0089$ , and  $F_0 = 1128.43$ .



Feng YIN (CUHKSZ)

A plot of the R-student residuals versus the fitted values,  $\hat{y}_i$ , is



# Summary

- Variance-Stabilizing
- 2 Empirical transformations
- Box-Cox transformation
- Intrincically linear model
- Model linearization