# CSC 4020 Fundamental of Machine Learning: Linear Regression

Baoyuan Wu School of Data Science, CUHK-SZ

January 25/27, 2021

#### Outline

- Some illustrations and review of last week
- 2 Linear Regression: A Deterministic Perspective
- 3 Linear Regression: A Probabilistic Perspective
  - Probabilistic modeling
  - Robust linear regression
  - Ridge regression
  - Lasso regression
- 4 Generalized Linear Regression

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- $\bullet$  Welcome to the office hour at Wednesday 10:30–11:30am in DY 411.

#### Review of last week

- Probability theory:
  - Discrete probability distributions: Bernoulli, Binomial, Beta
  - Continuous probability distributions: Gaussian, Student t, Laplace

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#### • Information theory:

• Information

Entropy, marginal/conditional/joint entropy, relative entropy (KL divergence, mutual information)

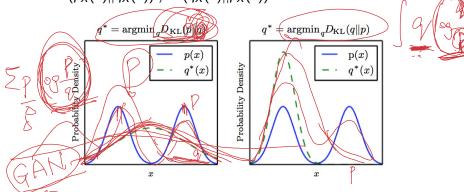
11/1/2) H(1/2) H(1/2)

P(X,Y) leg P(Y|X)

## Properties of KL divergence

•  $D(p_X(x)||q_X(x)) \ge 0$  with equality if and only if  $p_X(x) = q_X(x)$ .

•  $D(p_X(x)||q_X(x)) \neq D(q_X(x)||p_X(x))$ 



One constraint with respect to q is missing at last time, *i.e.*, it is the single mode distribution! More detailed derivations could be found at https://dibyaghosh.com/blog/probability/kldivergence.html.

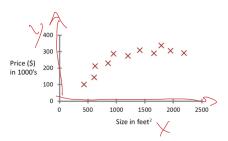
# Linear regression

• Here we start from a simple example of one dimensional input variable, and the training dataset  $D = \{(x_i, y_i)\}_{i=1}^m$  can be plotted on the x - y plane.

## Linear regression

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- ullet m indicates the number of training samples; x denotes the input variable/feature; y denotes the output variable.

Size in feet $(x)$	Price in 1000's (y)
7: 2104	460 y
1416	232
1514	315
852	178
·	



 Our goal is to find a <u>linear hypothesis</u> function to well fit the training data D, i.e.,

$$h_{\theta}(x) = \underline{\theta_0} + \underline{\theta_1} \phi(x) = [\theta_0, \theta_1][1; \phi(x)] = \hat{\phi}(x)$$

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where  $\phi(x)$  is called **basis expansion**, which is specified as different forms, such as  $\phi(x) = x$  or  $\phi(x) = [x^3, x^2; x]$ . In the following, we will use  $\phi(x) = x$  as example, while other expansions will be introduced later.



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- Given  $\theta_0, \theta_1, h_{\theta}(x)$  is the function of x.
- Given x,  $h_{\theta}(x)$  is a linear function of  $\theta = [\theta_0; \theta_1]$ . This is why it is called linear regression.

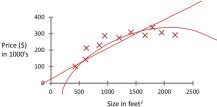


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- Given x,  $h_{\theta}(x)$  is a **linear function** of  $\theta = [\theta_0; \theta_1]$ . This is why it is called **linear regression**.
- Then, given D, how to learn  $\theta$ ?



#### Cost function

 $((K), \emptyset)$ 

• We design the following **cost function** to minimize the difference between the prediction  $h_{\theta}(x_i)$  and the ground-truth value  $y_i$ , *i.e.*,

$$\mathcal{M}(\boldsymbol{\theta}) = \frac{1}{2} \sum_{i=1}^{m} (h_{\boldsymbol{\theta}}(x_i) - y_i)^2 \qquad (2)$$

$$= \frac{1}{2} \sum_{i=1}^{m} (\underline{\theta_0} + \underline{\theta_1} x_i - y_i)^2, \qquad (3)$$

$$= \frac{1}{2} \sum_{i=1}^{m} (\bar{x}_i^{\top} \underline{\theta} - y_i)^2 \qquad (4)$$

which is called **residual sum of squares** (RSS) or sum of squared errors (SSE).

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 (2)

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which is called **residual sum of squares** (RSS) or sum of squared errors (SSE).

•  $J(\theta)$  is a convex or non-convex function? What is the shape of I(t)?

• The linear regression is formulated to the following optimization problem

$$\boldsymbol{\theta}^* = \arg\min_{\boldsymbol{\theta}} J(\boldsymbol{\theta}) = \frac{1}{2} \sum_{i=1}^{m} (\bar{\boldsymbol{x}}_i^{\top} \boldsymbol{\theta} - y_i)^2.$$
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•  $\theta$  can be updated by gradient descent algorithm,

$$\boldsymbol{\theta} \leftarrow \boldsymbol{\theta} - \alpha \frac{\partial J(\boldsymbol{\theta})}{\partial \boldsymbol{\theta}}, \quad \underline{\partial \boldsymbol{\theta}} = \sum_{i=1}^{m} (\bar{\boldsymbol{x}}_{i}^{\top} \boldsymbol{\theta} - y_{i}) \bar{\boldsymbol{x}}_{i}$$
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where  $\alpha$  is called step-size or learning rate.



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• Does gradient descent always converge to the optimal solution?

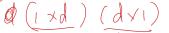
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• Does gradient descent always converge to the optimal solution? (Plot the trajectory of gradient descent on curve or contours)

## Analytical solution

• If we set the gradient to 0, then we can get the following solution

$$\frac{\partial J(\boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \sum_{i=1}^{m} (\bar{\boldsymbol{x}}_{i}^{\top} \boldsymbol{\theta} - y_{i}) \bar{\boldsymbol{x}}_{i} = \boldsymbol{X}^{\top} \boldsymbol{X} \boldsymbol{\theta} = \boldsymbol{X}^{\top} \boldsymbol{Y} \boldsymbol{\theta} = \boldsymbol{X}^{\top} \boldsymbol{Y} \boldsymbol{\theta}$$

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 (7)

$$\Rightarrow \underline{\boldsymbol{\theta}^* = (\boldsymbol{X}^\top \boldsymbol{X})^{-1} \boldsymbol{X}^\top \boldsymbol{y}},\tag{8}$$

which are called **normal equation** and **ordinary least squares** (OLS) solution, respectively.  $\boldsymbol{X} = [\bar{\boldsymbol{x}}_1^\top; \bar{\boldsymbol{x}}_2^\top; \dots; \bar{\boldsymbol{x}}_m^\top] \in \mathbb{R}^{m \times d}$ .

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• Since there is a closed-form solution, why do we need gradient descent algorithm?

## Geometric interpretation

ullet Since  $oldsymbol{ heta}^* = (oldsymbol{X}^ op oldsymbol{X})^{-1} oldsymbol{X}^ op oldsymbol{y},$  then the predictions of  $oldsymbol{X}$  can be obtained by

$$\hat{\boldsymbol{y}} = X\boldsymbol{\theta}^* = X(X^\top X)^{-1} X^\top \boldsymbol{y}, \tag{9}$$

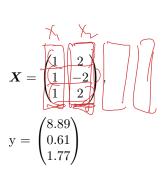
which corresponds to the <u>orthogonal projection</u> of y onto the <u>column</u> space of X.

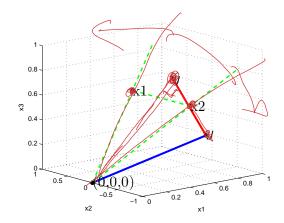
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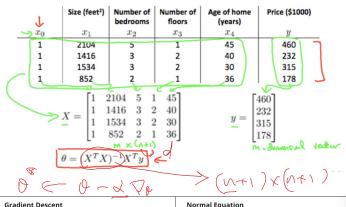




# Normal equation vs. gradient descent

J	Size (feet²)	Number of bedrooms	Number of floors	Age of home (years)	Price (\$1000	)
$\rightarrow x_0$	$x_1$	$x_2$	$x_3$	$x_4$	y	
((1)	2104	5	1,)	45	460	7
1	1416	3	2	40	232	1
1	1534	3	2	30	315	- 1
(1	852	2 (	1	36	178	7
	$X = \begin{bmatrix} 1 & 1 \end{bmatrix}$	2104 5 1 1416 3 2 1534 3 2		<i>u</i> –	460 232	
ſ	4[1	852 2 A	36	_	315 178	testor
L	$\theta = (X^T X$	$)^{-1}X^{T}y$	<b>&lt;</b>			

# Normal equation vs. gradient descent



Gradient Descent	Normal Equation
Need to choose alpha	No need to choose alpha
Needs many iterations	No need to iterate
0 (k)2)	$\bigcirc$ ( $n^3$ ), need to calculate inverse of $X^TX$
Works well when n is large	Slow if n is very large

# Probabilistic modeling

• We assume that the relationship between the input variable/feature x and the output variable y is O(x)

$$y = \underbrace{\boldsymbol{\theta}^{\top} \boldsymbol{x} + \boldsymbol{e}, \text{ where } \boldsymbol{e} \sim \mathcal{N}(0, \sigma^2),}_{\text{2}}$$
(10)

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where e is called **observation noise** or **residual error**, and it is independent with any specific input x.

• Thus, the output y can also be seen as a random variable, and its conditional probability is formulated as y



## Maximum log-likelihood estimation

• The parameter  $\theta$  can be learned by maximum log-likelihood estimation (MLE), given the training dataset  $D = \{(x_i, y_i)\}_{i=1}^m$ , as follows

$$\boldsymbol{\theta}_{MLE} = \arg\max_{\boldsymbol{\theta}} \underbrace{\log \mathcal{L}(\boldsymbol{\theta}|D)}_{(12)}$$

$$= \sum_{i}^{m} \log p(y|\mathbf{x}, \boldsymbol{\theta}) = \sum_{i}^{m} \log \mathcal{N}(\boldsymbol{\theta}^{\top} \mathbf{x}, \sigma^{2})$$
 (13)

$$= -\log(\sigma^{m}(2\pi)^{\frac{m}{2}}) - \frac{1}{2\sigma^{2}} \sum_{i}^{m} (y_{i} - \boldsymbol{\theta}^{\top} \boldsymbol{x}_{i})$$
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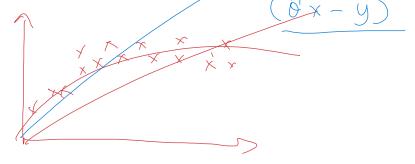
• Removing the constants w.r.t.  $\theta$ ,

$$\boldsymbol{\theta}_{MLE} = \arg\min_{\boldsymbol{\theta}} \frac{1}{2} \sum_{i}^{m} (y_i - \boldsymbol{\theta}^{\top} \boldsymbol{x}_i)^2, \tag{15}$$

which is exactly same with the cost function from the deterministic perspective.

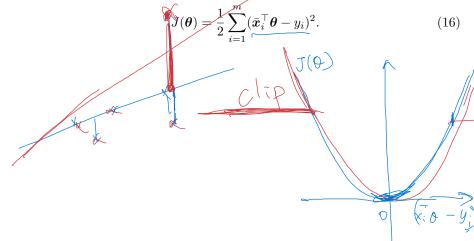
## Robust linear regression

• When there is a few outliers in the training data D, which are far from most other points, then learned parameters  $\theta_{MLF}$  will be significantly influenced, leading to very poor fit.



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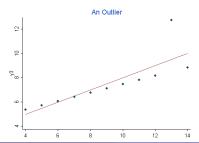
$$J(\boldsymbol{\theta}) = \frac{1}{2} \sum_{i=1}^{m} (\bar{\boldsymbol{x}}_i^{\top} \boldsymbol{\theta} - y_i)^2.$$
 (16)

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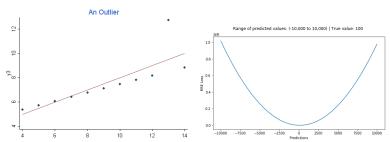
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- How to alleviate the significant influence of outliers?



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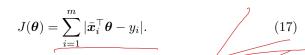
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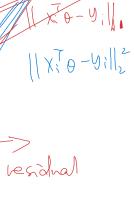
• We adopt the  $\ell_1$  loss to replace the  $\ell_2$  loss, as follows

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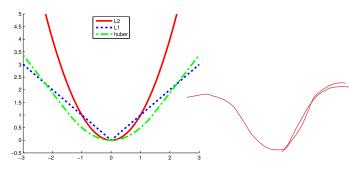
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- The curves of  $\ell_1$  and  $\ell_2$  losses are shown ad follows.
- When the residual is large, the  $\ell_1$  loss is much smaller than the  $\ell_2$  loss, such that the influence of outliers could be alleviated.



• Actually, the above  $\ell_1$  loss can also be derived from the probabilistic perspective, by assuming that

$$p(y|\boldsymbol{x},\boldsymbol{\theta},\underline{b}) \neq \text{Lap}(y|\boldsymbol{x},\boldsymbol{\theta},b) \propto \exp(-\frac{1}{b}|y-\boldsymbol{\theta}^{\top}\boldsymbol{x}|)$$
 (18)

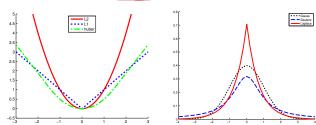
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 (18)

• Applying the maximum log-likelihood estimation (MLE), we will obtain

$$\boldsymbol{\theta}_{MLE} = \arg \max_{\boldsymbol{\theta}} \log \mathcal{L}(\boldsymbol{\theta}|D) = \sum_{i}^{m} \log p(y|\boldsymbol{x}, \boldsymbol{\theta})$$
 (19)

$$\equiv \arg\min_{\boldsymbol{\theta}} \sum_{i=1}^{m} |\bar{\boldsymbol{x}}_{i}^{\top} \boldsymbol{\theta} - y_{i}| \tag{20}$$



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• However, the  $\ell_1$  loss function is non-differentiable and non-linear. The gradient descent algorithm cannot be adopted.

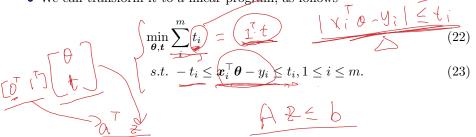
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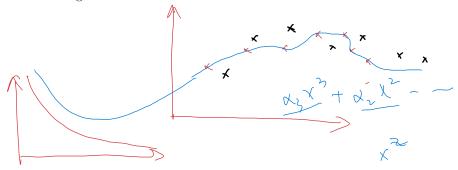
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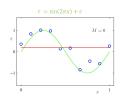


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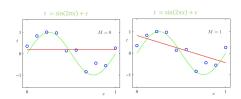


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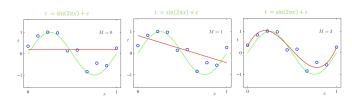
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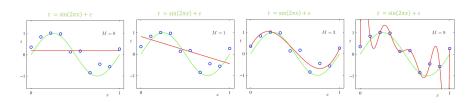
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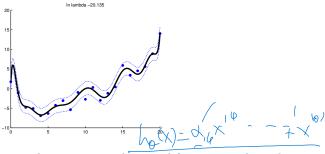


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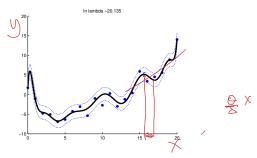


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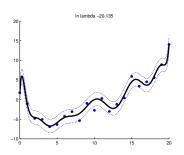


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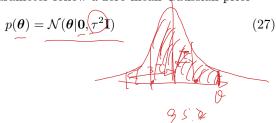
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• There are many large positive/negative values, such that a small change of features could lead to significant change of output.

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• Utilizing this prior, we obtain the maximum a posteriori (MAP) estimation

$$\theta_{MAP} = \arg\max_{\boldsymbol{\theta}} \sum_{i}^{m} \log p(y|\boldsymbol{x}, \boldsymbol{\theta}) + \log p(\boldsymbol{\theta})$$

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$$\equiv \arg\min_{\boldsymbol{\theta}} \sum_{i=1}^{m} (\bar{\boldsymbol{x}}_{i}^{\top} \boldsymbol{\theta} - y_{i})^{2} + \lambda \|\boldsymbol{\theta}\|_{2}^{2}. \qquad (30)$$

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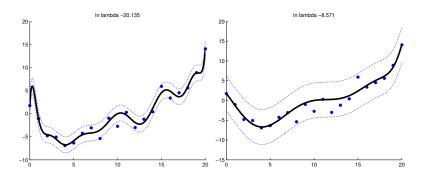
The corresponding closed-form solution is given by

$$\boldsymbol{\theta}_{MAP} = (\boldsymbol{I} + \boldsymbol{X}^{\top} \boldsymbol{X})^{-1} \boldsymbol{X}^{\top} y. \tag{31}$$

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- As shown below, when we set a larger  $\lambda$ , i.e., more weight on the prior, the resulting curve will be smoother.



• We can replace the Gaussian prior by a Laplacian prior, i.e.,

$$p(\boldsymbol{\theta}) = \text{Lap}(\boldsymbol{\theta}|\mathbf{0}, b) = \frac{1}{2b} \exp\left(-\frac{|\boldsymbol{\theta}|}{b}\right),$$
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• The combination of the Gaussian distribution of  $p(y|x, \theta)$  and the Laplacian prior, leading to

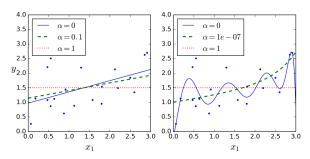
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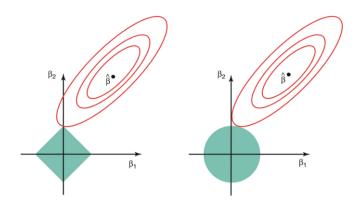
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# Geometry of Ridge and Lasso regression

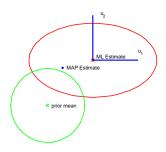
• Geometry of Ridge and Lasso regression. Which one is Ridge?



# Summary of different linear regressions

Note that the uniform distribution will not change the mode of the likelihood. Thus, MAP estimation with a uniform prior corresponds to MLE.

			r
p(y)	$y oldsymbol{x},oldsymbol{ heta})$	$p(\boldsymbol{\theta})$	regression method
Ga	ussian	Uniform	Least squares
Ga	ussian	Gaussian	Ridge regression
Ga	ussian	Laplace	Lasso regression
La	aplace	Uniform	Robust regression
St	udent	Uniform	Robust regression



# Generalized linear regression

#### • Linear model:

$$\mu(\boldsymbol{x}|\boldsymbol{\theta}) = \boldsymbol{\theta}^{\top} \phi(\boldsymbol{x}), \tag{36}$$

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• The standard linear model is a special case of GLM with g(a) = a.

# Why we need generalized linear regression

• Why we need generalized linear model? Let's see one example.

In the early stages of a disease epidemic, the rate at which new cases occur can often increase exponentially through time. Hence, if  $\mu_i$  is the expected number of new cases on day  $t_i$ , a model of the form

$$\mu_i = \gamma \exp(\delta t_i)$$

seems appropriate.

Such a model can be turned into GLM form, by using a log link so that

$$\log(\mu_i) = \log(\gamma) + \delta t_i = \beta_0 + \beta_1 t_i.$$

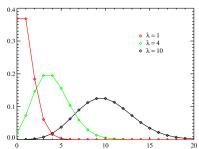
Since this is a count, the Poisson distribution (with expected value  $\mu_i$ ) is probably a reasonable distribution to try.

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- A discrete random variable X is said to have a Poisson distribution with parameter  $\lambda > 0$  if for k = 0, 1, 2, ..., the probability mass function of X is given by

$$f(k;\lambda) = P(X = k|\lambda) = \frac{\lambda^k e^{-\lambda}}{k!},$$
 (40)

where e is Euler's number (e = 2.71828...), we k is the number of occurrences, k! is the factorial of k.



• We assume that the conditional probability follows

$$P(y_i|\mathbf{x}_i, \boldsymbol{\theta}) = Poisson(\lambda_i) = \frac{\lambda^{y_i} e^{-\lambda}}{y_i!}, \quad \ln \lambda_i = \boldsymbol{\theta}^{\top} \mathbf{x}_i$$
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• Plot the log-linear regression as below.

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- Since  $\frac{y_i}{N} \in [0, 1]$ , it can be seen as the posterior probability. Thus, logistic regression is a classification model, rather than regression.

# Summary

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- Linear model is a special case of generalized linear model, while generalized linear model is not always linear
- Choosing different linear models is equivalent to choosing different distributions of  $p(y|\mathbf{x}, \boldsymbol{\theta})$  and  $p(\boldsymbol{\theta})$ , according to the task and the data

#### Reading material

• https://www.stat.cmu.edu/~ryantibs/advmethods/notes/glm.pdf