

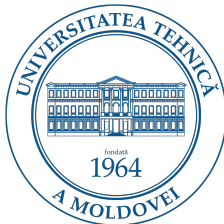
# Mathematics for Computer Science

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## Lecture 8



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- "How do you feel?" asked after a while terrorists.
- Professor answered: ...
- On the average just fine!



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A die is rolled. If an odd number turns up, we win an amount equal to this number; if an even number turns up, we lose an amount equal to this number.

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- Both of these quantities apply only to numerically-valued random variables.
- Consider the following game:  
A die is rolled. If an odd number turns up, we win an amount equal to this number; if an even number turns up, we lose an amount equal to this number.
- We want to decide if this is a **reasonable** game to play.  
You can try a computer simulation.

Winning	n = 100		n = 10000	
	Frequency	Relative Frequency	Frequency	Relative Frequency
1	17	.17	1681	.1681
-2	17	.17	1678	.1678
3	16	.16	1626	.1626
-4	18	.18	1696	.1696
5	16	.16	1686	.1686
-6	16	.16	1633	.1633

In the first run we have played the game 100 times.

In this run our average gain is  $-0.57$ .

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To get a better idea, we have played the game 10,000 times. In this case our average gain is  $-0.4949$ .

$$\mu = 1 \cdot \frac{1}{6} - 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{6} - 4 \cdot \frac{1}{6} + 5 \cdot \frac{1}{6} - 6 \cdot \frac{1}{6} = -0.5.$$

## Definition

Let  $X$  be a numerically-valued discrete random variable with sample space  $\Omega$  and distribution function  $m(x)$ .

The **expected value**  $E(X)$  is defined by

$$E(X) = \sum_{x \in \Omega} x \cdot m(x),$$

provided this sum converges absolutely.

Expected value  $E(X)$  is referred as the mean, and denoted also by  $\mu$ .

If the above sum does not converge absolutely, then we say that  $X$  does not have an expected value.

**Example 1.**

Let an experiment consist of tossing a fair coin 3 times.

Let  $X$  denote the number of heads which appear.

Then, the possible values of  $X$  are 0, 1, 2 and 3.

The corresponding probabilities are  $1/8$ ,  $3/8$ ,  $3/8$ , and  $1/8$ .



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Thus, the expected value of  $X$  equals:

$$E(X) = 0 \cdot \frac{1}{8} + 1 \cdot \frac{3}{8} + 2 \cdot \frac{3}{8} + 3 \cdot \frac{1}{8} = \frac{3}{2}.$$

Later we will see a quicker way to compute this expected value, based on the fact that  $X$  can be written as a sum of simpler random variables.

**Example 2.** Suppose that we toss a fair coin until a head first comes up, and let  $X$  represent the number of tosses which were made.

Then the possible values of  $X$  are  $1, 2, \dots$ , and the distribution function of  $X$  is defined by

$$m(i) = \frac{1}{2^i}.$$

Then,

$$E(X) = \sum_{i=1}^{\infty} i \cdot \frac{1}{2^i} = 2.$$

**Example 3.** Suppose that we flip a coin until a head first appears, and if the number of tosses equals  $n$ , then we are paid  $2^n$  dollars.

What is the expected value of the payment?

**Solution.** Let  $Y$  represent the payment. Then,

$$P(Y = 2^n) = \frac{1}{2^n},$$

$$E(Y) = \sum_{n=1}^{\infty} 2^n \frac{1}{2^n} = \sum_{n=1}^{\infty} 1,$$

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How much you would be willing to pay for this privilege?

It is unlikely that your answer is more than 10 dollars. Paradox!

**Example 4.**

Let  $T$  be the time for the first success in a Bernoulli trials process.

Take as sample space  $\mathbb{N}$  and assign geometric distribution

$$m(j) = P(T = j) = q^{j-1}p.$$

$$\begin{aligned} E(T) &= 1 \cdot p + 2qp + 3q^2p + \dots \\ &= p(1 + 2q + 3q^2 + \dots) \\ &= \frac{p}{(1 - q)^2} = \frac{1}{p}. \end{aligned}$$



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In particular, we see that if we toss a fair coin a sequence of times, the expected time until the first heads is  $\frac{1}{1/2} = 2$ . If we roll a die a sequence of times, the expected number of rolls until the first six is  $\frac{1}{1/6} = 6$ .

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- Using these properties, and definition of the variance to be introduced later, we shall be able to prove the  
**Law of Large Numbers.**
- This theorem will justify mathematically both the frequency concept of probability and the interpretation of expected value as the average value to be expected in a large number of experiments.

## Theorem

*If  $X$  is a discrete random variable with sample space  $\Omega$  and distribution function  $m(x)$ , and if  $\phi$  is a function, then*

$$E(\phi(X)) = \sum_{x \in \Omega} \phi(x) \cdot m(x),$$

*provided the series converges absolutely.*

## Theorem

*Let  $X$  and  $Y$  be random variables with finite expected values. Then*

$$\begin{aligned} E(X + Y) &= E(X) + E(Y), \\ E(cX) &= cE(X), \end{aligned}$$

*where  $c$  is any constant.*



### Theorem

*Let  $S_n$  be the number of successes in  $n$  Bernoulli trials with probability  $p$  for success on each trial. Then the expected number of successes is  $np$ . That is,*

$$E(S_n) = np.$$

### Proof.

Let  $X_j$  be a random variable which has the value 1 if the  $j$ -th outcome is a success and 0 if it is a failure. Then, for each  $X_j$ ,

$$\begin{aligned} E(X_j) &= 0 \cdot (1 - p) + 1 \cdot p = p, \\ S_n &= X_1 + X_2 + \dots + X_n. \end{aligned}$$

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The expected value of the sum is the sum of the expected values:

$$E(S_n) = np.$$

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If  $X$  and  $Y$  are two random variables, then in general

$$E(X \cdot Y) \neq E(X) \cdot E(Y).$$

However, this is true if  $X$  and  $Y$  are independent.

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### Theorem

*If  $X$  and  $Y$  are independent random variables, then*

$$E(X \cdot Y) = E(X) \cdot E(Y).$$

**Example.** A coin is tossed twice.  $X_i = 1$  if the  $i$ -th toss is heads and 0 otherwise. We know that  $X_1$  and  $X_2$  are independent. They each have expected value  $1/2$ . Thus

$$E(X_1 \cdot X_2) = E(X_1) \cdot E(X_2) = \frac{1}{4}.$$

**Example.** Consider a single toss of a coin. We define the random variable  $X$  to be 1 if heads turns up and 0 if tails turns up, and we set  $Y = 1 - X$ . Then

$$E(X) = E(Y) = \frac{1}{2}.$$

But  $XY = 0$  for either outcome. Hence,  $E(XY) = 0 \neq E(X)E(Y)$ .

We start keeping snowfall records this year and want to find the expected number of records that will occur in the next  $n$  years.

The first year is necessarily a record.

The second year will be a record if the snowfall in the second year is greater than that in the first year.

By symmetry, this probability is  $1/2$ .

More generally, let  $X_j$  be 1 if the  $j$ -th year is a record and 0 otherwise.

To find  $E(X_j)$ , we need only find the probability that the  $j$ -th year is a record.



But the record snowfall for the first  $j$  years is equally likely to fall in any one of these years, so  $E(X_j) = 1/j$ .

Therefore, if  $S_n$  is the total number of records observed in the first  $n$  years,

$$E(S_n) = 1 + \frac{1}{2} + \frac{1}{3} + \frac{1}{4} + \dots$$

This is the famous divergent harmonic series. It can be shown that

$$E(S_n) \sim \ln n \quad \text{as } n \rightarrow \infty$$

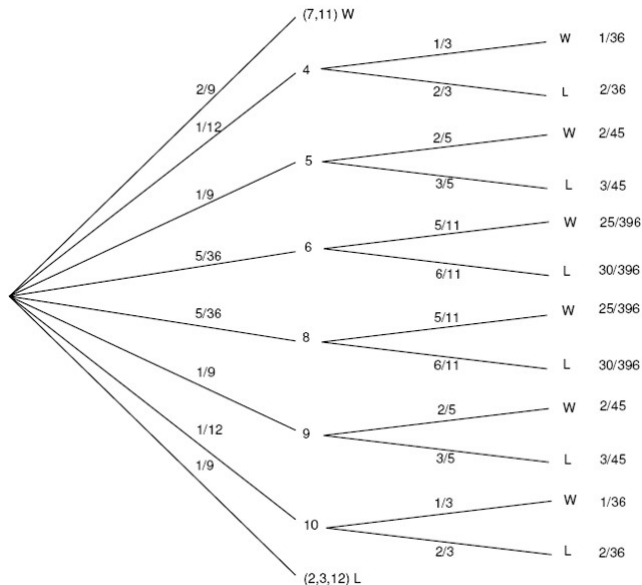
Therefore, in ten years the expected number of records is approximately  $\ln 10 = 2.3$ .

The exact value is the sum of the first ten terms of the harmonic series which is 2.9.



- In the game of craps, the player makes a bet and rolls a pair of dice.
- If the sum of the numbers is 7 or 11 the player wins, if it is 2, 3, or 12 the player loses.
- If any other number results, say  $r$ , then  $r$  becomes the player's point and he continues to roll until either  $r$  or 7 occurs. If  $r$  comes up first he wins, and if 7 comes up first he loses.
- Let us calculate the expected winnings on a single play and see if this game is favourable or not.
- We construct a two-stage tree measure as shown on the next page.

- The first stage represents the possible sums for his first roll.
- The second stage represents the possible outcomes for the game if it has not ended on the first roll. In this stage we are representing the possible outcomes of a sequence of rolls required to determine the final outcome. The branch probabilities for the first stage are computed in the usual way assuming all 36 possibilities for outcomes for the pair of dice are equally likely.
- For the second stage we assume that the game will eventually end, and we compute the conditional probabilities for obtaining either the point or a 7.
- For example, assume that the player's point is 6. Then the game will end when one of the eleven pairs, (1, 5), (2, 4), (3, 3), (4, 2), (5, 1), (1, 6), (2, 5), (3, 4), (4, 3), (5, 2), (6, 1), occurs.
- Assume that each of these possible pairs has the same probability. Then the player wins in the first 5 cases and loses in the last 6.



- Thus the probability of winning is  $5/11$  and the probability of losing is  $6/11$ .
- From the path probabilities, we can find the probability that the player wins 1 dollar; it is  $244/495$ .
- The probability of losing is then  $251/495$ . Thus if  $X$  is his winning for a dollar bet,

$$E(X) = 1 \cdot \frac{244}{495} + (-1) \cdot \frac{251}{495} = -0.0141.$$

- The game is unfavorable, but only slightly. The player's expected gain in  $n$  plays is  $-n(.0141)$ . If  $n$  is not large, this is a small expected loss for the player.
- The casino makes a large number of plays and so can afford a small average gain per play and still expect a large profit.

In Las Vegas, a roulette wheel has 38 slots numbered  $0, 00, 1, 2, \dots, 36$ . The 0 and 00 slots are green, and half of the remaining 36 slots are red and half are black. A croupier spins the wheel and throws an ivory ball. If you bet 1 dollar on red, you win 1 dollar if the ball stops in a red slot, and otherwise you lose a dollar. We wish to calculate the expected value of your winnings, if you bet 1 dollar on red.

Let  $X$  be the random variable which denotes your winnings in a 1 dollar bet on red in Las Vegas roulette. Then the distribution of  $X$  is given by

$$m_X = \begin{pmatrix} -1 & 1 \\ \frac{20}{38} & \frac{18}{38} \end{pmatrix}$$

Then,

$$E(X) = (-1) \cdot \frac{20}{38} + 1 \cdot \frac{18}{38} = -0.0526.$$

The usefulness of the expected value as a prediction for the outcome of an experiment is increased when the outcome is not likely to deviate too much from the expected value.

Let introduce a measure of this deviation, called the **variance**.

## Definition

Let  $X$  be a numerically valued random variable with expected value  $\mu = E(X)$ . Then, the **variance** of  $X$ , denoted by  $V(X)$ , is

$$V(X) = E((X - \mu)^2).$$

Note that,  $V(X)$  is given by

$$V(X) = \sum (x - \mu)^2 m(x),$$

where  $m$  is the distribution function of random variable  $X$ .



## Definition

The **standard deviation** of  $X$ , denoted by  $D(X)$ , is

$$D(X) = \sqrt{V(X)}.$$

We often write  $\sigma$  for  $D(X)$  and  $\sigma^2$  for  $V(X)$ .

**Example.** Consider one roll of a die. Let  $X$  be the number that turns up. To find  $V(X)$ , we must first find the expected value of  $X$ . This is

$$\mu = E(X) = 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + \dots + 6 \cdot \frac{1}{6} = \frac{7}{2}.$$

To find the variance of  $X$ , we form the new random variable  $(X - \mu)^2$  and compute its expectation.

$x$	$m(x)$	$(x - 7/2)^2$
1	1/6	25/4
2	1/6	9/4
3	1/6	1/4
4	1/6	1/4
5	1/6	9/4
6	1/6	25/4

From this table we find  $E((X - \mu)^2)$  is

$$\begin{aligned} V(X) &= \frac{25}{4} \cdot \frac{1}{6} + \frac{9}{4} \cdot \frac{1}{6} + \frac{1}{4} \cdot \frac{1}{6} + \frac{1}{4} \cdot \frac{1}{6} + \frac{9}{4} \cdot \frac{1}{6} + \frac{25}{4} \cdot \frac{1}{6} \\ &= \frac{35}{12} \end{aligned}$$

and the standard deviation

$$D(X) = \sqrt{35/12} = 1.707.$$

## Theorem

*If  $X$  is any random variable with  $E(X) = \mu$ , then*

$$V(X) = E(X^2) - \mu^2.$$

## Proof.

We have:

$$\begin{aligned} V(X) &= E((X - \mu)^2) \\ &= E(X^2 - 2\mu X + \mu^2) \\ &= E(X^2) - 2\mu E(X) + E(\mu^2) \\ &= E(X^2) - 2\mu\mu + \mu^2. \end{aligned}$$



Consider the example of rolling a die.

$$E(X^2) = 1^2 \cdot \frac{1}{6} + 2^2 \cdot \frac{1}{6} + 3^2 \cdot \frac{1}{6} + \dots + 6^2 \cdot \frac{1}{6} = \frac{91}{6}$$

and,

$$\begin{aligned} V(X) &= E(X^2) - \mu^2 \\ &= \frac{91}{6} - \left(\frac{7}{2}\right)^2 \\ &= \frac{35}{12}, \end{aligned}$$

which is in agreement with the value obtained directly from the definition of  $V(X)$ .

If  $c$  is any constant,  $E(cX) = cE(X)$  and  $E(X + c) = E(X) + c$ . These two statements imply that the expectation is a linear function. The variance is not linear.

## Theorem

*If  $X$  is any random variable and  $c$  is any constant, then*

$$V(cX) = c^2 V(X), \quad V(X + c) = V(X).$$

## Proof.

Let  $\mu = E(X)$ . Then  $E(cX) = c\mu$ , and

$$\begin{aligned} V(cX) &= E((cX - c\mu)^2) = E(c^2(X - \mu)^2) \\ &= c^2 E((X - \mu)^2) = c^2 V(X). \end{aligned}$$

The second assertion is proved similarly. □

## Theorem

*Let  $X$  and  $Y$  be two independent random variables. Then*

$$V(X + Y) = V(X) + V(Y).$$

## Proof.

Let  $E(X) = a$  and  $E(Y) = b$ . Then

$$\begin{aligned} V(X + Y) &= E((X + Y)^2) - (a + b)^2 \\ &= E(X^2) + 2E(XY) + E(Y^2) - a^2 - 2ab - b^2. \end{aligned}$$

Since  $X$  and  $Y$  are independent,  $E(XY) = E(X)E(Y) = ab$ . Thus,

$$\begin{aligned} V(X + Y) &= E(X^2) - a^2 + E(Y^2) - b^2 \\ &= V(X) + V(Y). \end{aligned}$$

## Theorem

Let  $X_1, X_2, \dots, X_n$  be an independent trials process with  $E(X_j) = \mu$  and  $V(X_j) = \sigma^2$ . Let

$$S_n = X_1 + X_2 + \dots + X_n$$

be the sum, and

$$A_n = \frac{S_n}{n}$$

be the average. Then

$$E(S_n) = n\mu, \quad V(S_n) = n\sigma^2,$$

$$E(A_n) = \mu, \quad V(A_n) = \frac{\sigma^2}{n}.$$

## Proof.

Since all the random variables  $X_j$  have the same expected value, we have

$$\begin{aligned}E(S_n) &= E(X_1) + \dots + E(X_n) = n\mu, \\V(S_n) &= V(X_1) + \dots + V(X_n) = n\sigma^2.\end{aligned}$$

We have seen that, if we multiply a random variable  $X$  with mean  $\mu$  and variance  $\sigma^2$  by a constant  $c$ , the new random variable has expected value  $c\mu$  and variance  $c^2\sigma^2$ . Thus,

$$\begin{aligned}E(A_n) &= E\left(\frac{S_n}{n}\right) = \frac{n\mu}{n} = \mu, \\V(A_n) &= V\left(\frac{S_n}{n}\right) = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}.\end{aligned}$$





Note that the standard deviation of  $A_n$  is given by

$$\sigma(A_n) = \frac{\sigma}{\sqrt{n}}.$$

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This statement implies that in an independent trials process, if the individual terms have finite variance, then the standard deviation of the average goes to 0 as  $n \rightarrow \infty$ .

Since the standard deviation tells us something about the spread of the distribution around the mean, we see that for large values of  $n$ , the value of  $A_n$  is usually very close to the mean of  $A_n$ , which equals  $\mu$ , as shown above.

This statement is called the **Law of Large Numbers**.

**Example.** Consider  $n$  rolls of a die.

We have seen that, if  $X_j$  is the outcome if the  $j$ -th roll, then  $E(X_j) = 7/2$  and  $V(X_j) = 35/12$ .

Thus, if  $S_n$  is the sum of the outcomes, and  $A_n = S_n/n$  is the average of the outcomes, we have  $E(A_n) = 7/2$  and  $V(A_n) = (35/12)/n$ .

Therefore, as  $n$  increases, the expected value of the average remains constant, but the variance tends to 0.

If the variance is a measure of the expected deviation from the mean this would indicate that, for large  $n$ , we can expect the average to be very near the expected value.

Consider next the general Bernoulli trials process. As usual, we let  $X_j = 1$  if the  $j$ -th outcome is a success and 0 if it is a failure. If  $p$  is the probability of a success, and  $q = 1 - p$ , then

$$E(X_j) = 0 \cdot q + 1 \cdot p = p,$$

$$E(X_j^2) = 0^2 \cdot q + 1^2 \cdot p = p,$$

$$V(X_j) = E(X_j^2) - (E(X_j))^2 = p - p^2 = pq.$$

Thus, for Bernoulli trials, if  $S_n = X_1 + X_2 + \dots + X_n$  is the number of successes, then

$$E(S_n) = np, \quad V(S_n) = npq, \quad D(S_n) = \sqrt{npq}.$$

If  $A_n = S_n/n$  is the average number of successes, then

$$E(A_n) = p, \quad V(A_n) = pq/n, \quad D(A_n) = \sqrt{pq/n}.$$

See that expected proportion of successes remains  $p$  and the variance tends to 0. This suggests that the frequency interpretation of probability is a correct one.

Let  $T$  denote the number of trials until the first success in a Bernoulli trials process.

Then  $T$  is geometrically distributed. What is the variance of  $T$ ?

$$m_T = \begin{pmatrix} 1 & 2 & 3 & \dots, \\ p & pq & pq^2 & \dots \end{pmatrix}$$

We showed that  $E(T) = 1/p$ . Thus,  $V(T) = E(T^2) - (1/p)^2$  so we need only find

$$\begin{aligned} E(T^2) &= 1 \cdot p + 4qp + 9q^2p + \dots \\ &= p(1 + 4q + 9q^2 + \dots) \\ &= \frac{1+q}{p^2}. \end{aligned}$$

Therefore,

$$\begin{aligned} V(T) &= E(T^2) - (E(T))^2 \\ &= \frac{1+q}{p^2} - \frac{1}{p^2} = \frac{q}{p^2}. \end{aligned}$$

For example, the variance for the number of tosses of a coin until the first head turns up is  $(1/2)/(1/2)^2 = 2$ .

The variance for the number of rolls of a die until the first six turns up is  $(5/6)/(1/6)^2 = 30$ .

Note that, as  $p$  decreases, the variance increases rapidly. This corresponds to the increased spread of the geometric distribution as  $p$  decreases.

# The End of Lecture