Modeling the Cell Proliferation using Compound Non-homogeneous Poisson Distribution

The Author

1 Model

The evolution of the average cell counts of the active stem cell, $S(t)$, and differentiated cells, $F(t)$, is described by the following differential equations

$$\frac{dS(t)}{dt} = r[p_1(t) - p_3(t)]S(t), \quad S(0) = S_0,$$
$$\frac{dF(t)}{dt} = r[p_2(t) + 2p_3(t)]S(t), \quad F(0) = 0,$$

where $r > 0$ and $p_1(t), p_2(t), p_3(t) > 0$ and $0 < p_1(t) + p_2(t) + p_3(t) \leq 1$ for any $t > 0$. The solutions of the differential equations are given by

$$\begin{aligned} S(t) &= S_0 \exp\Big\{rP(t)\Big\}, \\ F(t) &= \int_0^t r[p_2(v) + 2p_3(v)]S(v)dv \\ &= S_0 \cdot r \cdot \int_0^t [p_2(v) + 2p_3(v)] \exp\Big\{rP(v)\Big\}dv, \end{aligned}$$

where $P(t) = \int_0^t [p_1(v) - p_3(v)]dv$.

We model the stem cell and differentiated cell counts using the compound nonhomogeneous Poisson process as follows. Let $N(t)$ be a nonhomogeneous Poisson process with rate $\lambda(t) = rS(t)$. Given a sequence of arrival times $S_1, S_2, S_3, \ldots$, the pairs of random variables $(X_1, Y_1), (X_2, Y_2), (X_3, Y_3), \ldots$ are independent with the following marginal distributions

$$(X_k, Y_k) = \begin{cases} (+1, 0), & p_1(S_k) \\ (0, +1), & p_2(S_k) \\ (-1, +2), & p_3(S_k) \\ (0, 0), & p_4(S_k) \end{cases},$$

where $p_4(t) = 1 - p_1(t) - p_2(t) - p_3(t)$. First, we define two compound nonhomogeneous Poisson processes, $X(t)$ and $Y(t)$, as follows

$$X(t) = \sum_{k=1}^{N(t)} X_k, \quad \text{and} \quad Y(t) = \sum_{k=1}^{N(t)} Y_k.$$

Next, we define the stopping time τ (when the stem cell count reaches 0) by

$$\tau = \inf\{t > 0 : X(t) = -S_0\}.$$

Under assumption $\mathbb{P}(\tau < \infty)$ the counts of stem cells and differentiated cells are modelled by processes $S_0 + X(t \wedge \tau)$ and $Y(t \wedge \tau)$, respectively.

2 Some Observations

From Theorem 2.1 and Corollary to Proposition 2.2 of Chen and Savits (1993) (cite), the characteristic functions, expected values, and variances of $X(t)$ and $Y(t)$ are given by

$$\begin{aligned}
\mathbb{E}[e^{iuX(t)}] &= \exp\Big\{\int_0^t \Big[e^{iu}p_2(v) - p_2(v) + e^{-iu}p_3(v) - p_3(v)\Big]rS(v)dv\Big\},\\
\mathbb{E}[X(t)] &= \int_0^t [p_1(v) - p_3(v)]rS(v)dv,\\
\mathbb{V}(X(t)) &= \int_0^t [p_1(v) + p_3(v)]rS(v)dv,\\
\mathbb{E}[e^{iuY(t)}] &= \exp\Big\{\int_0^t \Big[e^{iu}p_2(v) - p_2(v) + e^{2iu}p_3(v) - p_3(v)\Big]rS(v)dv\Big\},\\
\mathbb{E}[Y(t)] &= \int_0^t [p_2(v) + 2p_3(v)]rS(v)dv,\\
\mathbb{V}(Y(t)) &= \int_0^t [p_2(v) + 4p_3(v)]rS(v)dv.
\end{aligned}$$

Note that

$$\mathbb{E}[X(t)] = \int_0^t [p_1(v) - p_3(v)] \cdot r \cdot S_0 \exp\Big\{rP(v)\Big\}dv.$$

Therefore, using substitution with $q = P(v) = \int_0^v [p_1(u) - p_3(u)]du$ and $dq/dv = [p_1(v) - p_3(v)]$, we get

$$\mathbb{E}[X(t)] = S_0 r \int_0^{P(t)} \exp\{rq\}dq = S_0 \exp\{rq\}\Big|_0^{P(t)} = S_0 \exp\{rP(t)\} - S_0.$$

Thus, $\mathbb{E}[S_0 + X(t)] = S(t)$.

Proposition 1. *For a fixed t, $(S_0 + X(t))/S(t) \xrightarrow{P} 1$ and $Y(t)/F(t) \xrightarrow{P} 1$ as $S_0 \to \infty$.*

Proof. Using the first-order Taylor's expansion for the exponential function we have,

$$\begin{aligned}
\mathbb{E}\Big[e^{iu\frac{S_0+X(t)}{S(t)}}\Big] &= e^{iu\frac{S_0}{S(t)}} \exp\Big\{\int_0^t \Big[e^{i\frac{u}{S(t)}}p_1(v) - p_1(v) + e^{-i\frac{u}{S(t)}}p_3(v) - p_3(v)\Big]rS(v)dv\Big\}\\
&= e^{iu\frac{S_0}{S(t)}} \exp\Big\{\int_0^t \Big[\Big(\frac{iu}{S(t)} + o\left(S_0^{-1}\right)\Big)p_1(v) + \Big(-\frac{iu}{S(t)} + o\left(S_0^{-1}\right)\Big)p_3(v)\Big]rS(v)dv\Big\}\\
&= e^{iu\frac{S_0}{S(t)}} \exp\Big\{\frac{iu}{S(t)}\int_0^t [p_1(v) - p_3(v)]rS(v)dv + o(S_0^{-1})\int_0^t [p_1(v) + p_3(v)]rS(v)dv\Big\}\\
&= e^{iu\frac{S_0}{S(t)}} \exp\Big\{\frac{iu}{S(t)}\mathbb{E}(X(t) + o(1)\int_0^t [p_1(v) + p_3(v)]r\exp\{rP(v)\}dv\Big\}\\
&= \exp\Big\{\frac{iu}{S(t)}(\mathbb{E}(X(t) + S_0) + o(1)\int_0^t [p_1(v) + p_3(v)]r\exp\{rP(v)\}dv\Big\}\\
&\to \exp\Big\{iu\Big\},
\end{aligned}$$

as $S_0 \to \infty$. Thus, $(S_0 + X(t))/S(t) \xrightarrow{D} 1$, and, therefore, $(S_0 + X(t))/S(t) \xrightarrow{P} 1$. Similarly, we get the convergence for $Y(t)$.

□

Proposition 2. *For a fixed t, as $S_0 \to \infty$,*

$$\frac{S_0 + X(t) - S(t)}{\sqrt{Var(X(t))}} \xrightarrow{D} N(0,1),$$

and

$$\frac{Y(t)-F(t)}{\sqrt{\mathbb{V}(Y(t))}} \xrightarrow{D} N(0,1).$$

Proof. Let $\sigma_{x_t} = \sqrt{\mathbb{V}(X(t))}$. Using the second-order Taylor's expansion, we get

$$\begin{aligned}
\mathbb{E}\Big[e^{iu\frac{X(t)}{\sigma_{x_t}}}\Big] &= \exp\Big\{\int_0^t \Big[e^{i\frac{u}{\sigma_{x_t}}}p_1(v) - p_1(v) + e^{-i\frac{u}{\sigma_{x_t}}}p_3(v) - p_3(v)\Big]rS(v)dv\Big\} \\
&= \exp\Big\{\int_0^t \Big[\Big(\frac{iu}{\sigma_{x_t}} + \frac{i^2u^2}{2\sigma_{x_t}^2} + o(\sigma_{x_t}^{-2})\Big)p_1(v) + \\
&\Big(-\frac{iu}{\sigma_{x_t}} + \frac{i^2u^2}{2\sigma_{x_t}^2} + o(\sigma_{x_t}^{-2})\Big)p_3(v)\Big]rS(v)dv\Big\} \\
&= \exp\Big\{\frac{iu}{\sigma_{x_t}}\int_0^t [p_1(v) - p_3(v)]rS(v)dv + \frac{i^2u^2}{2\sigma_{x_t}^2}\int_0^t [p_1(v) + p_3(v)]rS(v)dv \\
&+ o(\sigma_{x_t}^{-2})\int_0^t [p_1(v) + p_3(v)]rS(v)dv\Big\} \\
&= \exp\Big\{\frac{iu}{\sigma_{x_t}}(S(t) - S_0) + \frac{i^2u^2}{2\sigma_{x_t}^2}\mathbb{V}(X(t)) \\
&+ o(S_0^{-1})\int_0^t [p_1(v) + p_3(v)]r\exp\{rP(v)\}dv\Big\}.
\end{aligned}$$

Therefore, as $S_o \to \infty$ we have

$$\mathbb{E}\Big[e^{iu\frac{S_0+X(t)-S(t)}{\sigma_{x_t}}}\Big] \to e^{-\frac{u^2}{2}}.$$

Again, the convergence for $Y(t)$ can be shown in a similar way. □

3 Covariance of $X(t)$ and $Y(t)$ when $p_4(t) = 0$

When $p_4(t) = 0 \forall t$, i.e all the stem cells can undergo further division and there is no inactive stem cell, the number of divisions equal to the sum of stem cells and differentiated cells $N(t) = X(t) + Y(t)$. For each division occurrence, the sum of stem cells and ependymal cells increases by 1 (Table 1). Then, $E[(N(t))^2] = E[(X(t))^2] + E[X(t)Y(t)] + E[(Y(t))^2]$.

	$p_1(t)$	$p_2(t)$	$p_3(t)$
$X(t)$	$+1$	$+0$	-1
$Y(t)$	$+0$	$+1$	$+2$
$X(t)+Y(t)$	$+1$	$+1$	$+1$

Table 1: Change of total cell counts with each division occurrence when $p_4(t) = 0$.

Since $N(t)$ is an nonhomogeneous Poisson process with rate $\lambda(t) = rS(t)$, $N(t)$ is a Poisson random variable with mean and variance

$$E[N(t)] = Var(N(t)) = \int_0^t rS(v)dv.$$

We obtain the covariance of $X(t)$ and $Y(t)$ when $p_4(t) = 0$

$$\begin{aligned}
Cov(X(t), Y(t)) &= E[X(t)Y(t)] - E[X(t)]E[Y(t)] \\
&= E[(N(t))^2] - E[(X(t))^2] - E[(Y(t))^2] - E[X(t)]E[Y(t)] \\
&= (E[N(t)])^2 + Var(N(t)) - (E[X(t)])^2 - Var(X(t)) - (E[Y(t)])^2 - Var(Y(t)) \\
&- E[X(t)]E[Y(t)] \\
&= \Big[\int_0^t rS(v)dv\Big]^2 - \Big[\int_0^t [p_1(v) - p_3(v)]rS(v)dv\Big]^2 - \Big[\int_0^t [p_2(v) + 2p_3(v)]rS(v)dv\Big]^2 \\
&- \int_0^t 4p_3(v)rS(v)dv - \int_0^t [p_1(v) - p_3(v)]rS(v)dv \cdot \int_0^t [p_2(v) + 2p_3(v)]rS(v)dv.
\end{aligned} \tag{1}$$

Limit of the $X(t)$ and $Y(t)$

Show the variance of $X(t)$ and of $Y(t)$ converge as t is large.

$$Var(X(t)) = \int_0^t [p_1(v) + p_3(v)]rS(v)dv,$$

$$Var(Y(t)) = \int_0^t [p_2(v) + 4p_3(v)]rS(v)dv.$$

Since $p_1(v) + p_3(v) \leq 1$ for any $v \in (0, \infty)$, then $[p_1(v) + p_3(v)]rS(v) \leq rS(v)$. Then using the comparison test if $\int_0^\infty rS(v)dv$ converges then $\int_0^\infty [p_1(v) + p_3(v)]rS(v)dv$ converges

(Note that in the compound nonhomogeneous Poisson process, $E[N(t)] = Var(N(t)) = \int_0^t rS(v)dv$).

$$\int_0^\infty rS(v)dv = r\int_0^\infty S_0 \exp\Big\{r\int_0^v [p_1(u) - p_3(u)]du\Big\}dv \tag{2}$$

Assume that $p_1(u) - p_3(u)$ is a continuous function on $(0, \infty)$ and there exists $u_0 \in (0, \infty)$ such that $\forall u > u_0, p_1(u) - p_3(u) \leq -q < 0$,

$$\begin{aligned}
\int_0^v [p_1(u) - p_3(u)]du &= \int_0^{u_0} [p_1(u) - p_3(u)]du + \int_{u_0}^v [p_1(u) - p_3(u)]du \\
&\leq \int_0^{u_0} 1du + \int_{u_0}^v (-q)du \quad \text{since } p_1(u) - p_3(u) \leq 1 \forall u \\
&= u_0 - rq(v - u_0).
\end{aligned} \tag{3}$$

Then,

$$\begin{aligned}
\int_0^\infty rS(v)dv &= r\int_0^\infty S_0 \exp\Big\{r\int_0^v [p_1(u) - p_3(u)]du\Big\}dv \\
&\leq r\int_0^\infty S_0 \exp\{r(u_0 - rq(v - u_0))\}dv \\
&= r\int_0^\infty S_0 \exp\{u_0(1 + rq)\}\exp\{-rqv\}dv \\
&= S_0 \cdot r \cdot \exp\{u_0(1 + rq)\}\int_0^\infty \exp\{-rqv\}dv \\
&= \frac{S_0 \exp\{u_0(1 + rq)\}}{q} < \infty.
\end{aligned} \tag{4}$$

Similarly for $Var(Y(t))$, since $[p_2(v) + 4p_3(v)] \leq 4$ for any $v \in (0, \infty)$,

$$\int_0^\infty [p_2(v) + 4p_3(v)]rS(v)dv \leq \int_0^\infty 4rS(v)dv \leq \frac{4S_0 \exp\{u_0(1 + rq)\}}{q} < \infty.$$