Lecture 15: Markov and Chebyshev's Inequalities

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We talk about Chebyshev's inequality and compute the variance of the geometric distribution.

The variance gives a powerful way to measure the probability that a random variable deviates from its expectation by a lot. As we have seen, $\mathbb{E}[X]$ does not tell us anything about how far X can be from its expectation. However, we do have the following simple inequality (we proved it in the last lecture):

Fact 1 (Markov's inequality). *If* X *is a non-negative random variable, then*

$$p(X \ge \alpha) \le \frac{\mathbb{E}[X]}{\alpha}.$$

Applying Markov's inequality to the variance gives us Chebyshev's inequality:

Fact 2 (Chebyshev's inequality). *If* $\alpha \geq 0$,

$$p(|X - \mathbb{E}[X]| \ge \alpha) \le \frac{\operatorname{Var}[X]}{\alpha^2}.$$

Proof.

$$p(|X - \mathbb{E}[X]| \ge \alpha) = p((X - \mathbb{E}[X])^2 \ge \alpha^2)$$

$$\le \frac{\mathbb{E}[(X - \mathbb{E}[X])^2]}{\alpha^2}$$

$$= \frac{\text{Var}[X]}{\alpha^2}.$$

by Markov's inequality applied to the non-negative random variable $(X - \mathbb{E}[X])^2$.

Let us apply Markov and Chebyshev's inequality to the geometric distribution.

Example: Geometric Distribution

Suppose we repeatedly toss a coin until we see heads. Suppose the probability of heads in each coin toss is p. Let X be the number of coin tosses.

We saw in the last lecture that

$$\mathbb{E}[X] = \sum_{i=1}^{\infty} p \cdot (1-p)^{i-1} i = 1/p.$$

To calculate the variance,

$$\mathbb{E}\left[X^{2}\right] = \sum_{i=1}^{\infty} p \cdot (1-p)^{i-1} i^{2} = p \sum_{i=1}^{\infty} (1-p)^{i-1} i^{2}.$$

To calculate this, we use Taylor series. We have the identity

$$\frac{1}{1-x} = 1 + x + x^2 + \dots$$

This already shows that

$$\mathbb{E}[X] = p + p \cdot 2(1-p) + p \cdot 3(1-p)^2 + \dots = \frac{p}{1 - (1-p)}.$$

Taking the derivative gives

$$\frac{1}{(1-x)^2} = 1 + 2x + 3x^2 + 4x^3 + \dots$$

This already shows that

$$\mathbb{E}[X] = p + p \cdot 2(1-p) + p \cdot 3(1-p)^2 + \dots = \frac{p}{(1-(1-p))^2} = \frac{1}{p}.$$

Multiplying the Taylor series by *x* gives

$$\frac{x}{(1-x)^2} = x + 2x^2 + 3x^3 + 4x^4 + \dots$$

Then taking the derivative again, we get

$$\frac{1}{(1-x)^2} + \frac{2x}{(1-x)^3} = 1 + 2^2 \cdot x + 3^2 \cdot x^2 + 4^2 \cdot x^3 + \dots,$$

which gives

$$\frac{1+x}{(1-x)^3} = 1 + 2^2 \cdot x + 3^2 \cdot x^2 + 4^2 \cdot x^3 + \dots,$$

so:

$$\mathbb{E}\left[X^2\right] = p \sum_{i=1}^{\infty} (1-p)^{i-1} i^2 = p \cdot \frac{1+1-p}{(1-(1-p))^3} = \frac{2-p}{p^2}.$$

So the variance is:

Var
$$[X] = \mathbb{E}[X^2] - \mathbb{E}[X]^2 = \frac{2-p}{p^2} - \frac{1}{p^2} = \frac{1-p}{p^2}.$$

Now, let us estimate the probability that we have to wait twice as long to see the first heads. Markov's inequality gives:

$$p(X \ge 2/p) \le \frac{\mathbb{E}[X]}{2/p} \le 1/2.$$

Chebyshev's inequality gives

$$p(X\geq 2/p)=p(|X-\mathbb{E}\left[X\right]|\geq 1/p)\leq \frac{\mathrm{Var}\left[X\right]}{1/p^2}=\frac{(1-p)/p^2}{1/p^2}=1-p.$$

This is better than Markov's inequality when p > 1/2, but worse otherwise.

What if we toss the coin until we see *n* heads? Let *X* denote the number of coin tosses. Let X_1 be the number of coin tosses to see the first heads, let X_2 denote the number of additional coin tosses to see the second heads and so on. Then we see that

$$X = X_1 + X_2 + \ldots + X_n.$$

Moreover, $X_1, X_2, ..., X_n$ are mutually independent. We have

$$\mathbb{E}[X] = \mathbb{E}[X_1] + \mathbb{E}[X_2] + \ldots + \mathbb{E}[X_n] = n/p,$$

by linearity of expectation. Since the variables are independent, we have

$$Var[X] = Var[X_1] + Var[X_2] + ... + Var[X_n] = \frac{1-p}{p^2} \cdot n.$$

Let us again estimate the probability that the number of tosses is twice as many as we expect. Markov's inequality gives:

$$p(X \ge 2n/p) \le \frac{n/p}{2n/p} = \frac{1}{2}.$$

Chebyshev's inequality gives

$$p(X \ge 2n/p) = p(|X - \mathbb{E}[X]| \ge n/p) \le \frac{\text{Var}[X]}{(n/p)^2} = \frac{(1-p)/p^2 \cdot n}{n^2/p^2} = \frac{1-p}{n}.$$

As *n* gets larger, Chebyshev's inequality gives a much stronger bound.

You can also calculate directly that $p(X \ge 2/p) \le (1-p)^{2/p-1}$.