

Data-based Statistical Decision Model

Lecture 9 supplement - Predicting binary outcomes

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Logistic regression

- A simple-minded understanding of logistic regression is to predict $y \in \{0, 1\}$ using x_1, \dots, x_p by a formula

$$y \sim \phi(x_1, \dots, x_p) = \phi(\beta_0 + \beta_1 x_1 + \dots + \beta_p x_p) = \phi(\beta' X)$$

- Logistic regression starts with different model setup than linear regression: instead of modeling Y as a function of $X = (x_1, \dots, x_p)$ directly, we model the probability that Y is equal to class 1, given X . First, abbreviate $p(X) = P(Y = 1|X)$. Then the logistic model is

$$p(X) = \frac{\exp(\beta' X)}{1 + \exp(\beta' X)}.$$

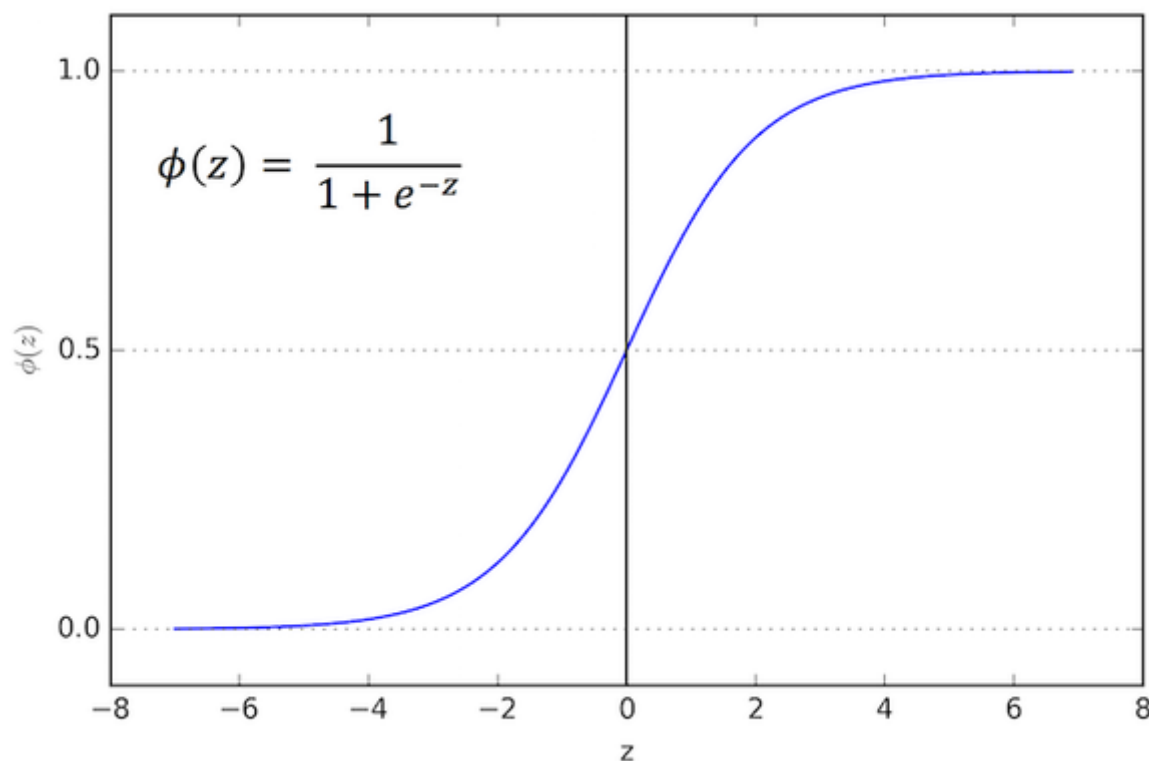
Logistic regression as a *generalized* linear regression

- “Linear”: Explanatory variables to an intermediate predictor z :

$$(x_1, \dots, x_p) \rightarrow z = \beta_0 + \beta_1 x_1 + \dots + \beta_p x_p$$

- “Generalized”: z to estimated probability p :

$$p = \phi(z)$$



Interpreting coefficients

- How can we interpret the role of the coefficients β ?
- The logistic model is rearranged:

$$\log\left(\frac{p(X)}{1 - p(X)}\right) = \beta'X$$

- The left-hand side $\log\left(\frac{p(X)}{1 - p(X)}\right)$ is called *log-odds* of Class 1.
 - $p(X) = P(Y = 1 \mid X)$ = Probability of Class 1 (given X).
 - $p(X)/(1 - p(X))$ = *Odds* of Class 1 (given X)

Interpreting odds

Odds are an alternative scale to probability for representing chance

- As a way to express the payoffs for bets
- *Evens* bet: Winner gets paid an equal amount to that staked [Odds = 1]
- 3-1 *against* bet: pay \$3 for every \$1 [Odds = 1/3]
- 3-1 *on* bet: pay \$1 for every \$3 [Odds = 3]
- if the games are fair, if you win with probability p , then you would make in the long run

$$E(\text{payout}) = (1 - p)(-1) + p\left(\frac{1}{\text{Odds}}\right) = 0$$

- That is, Odds = $\frac{p}{1-p}$.

A simple 1-d example

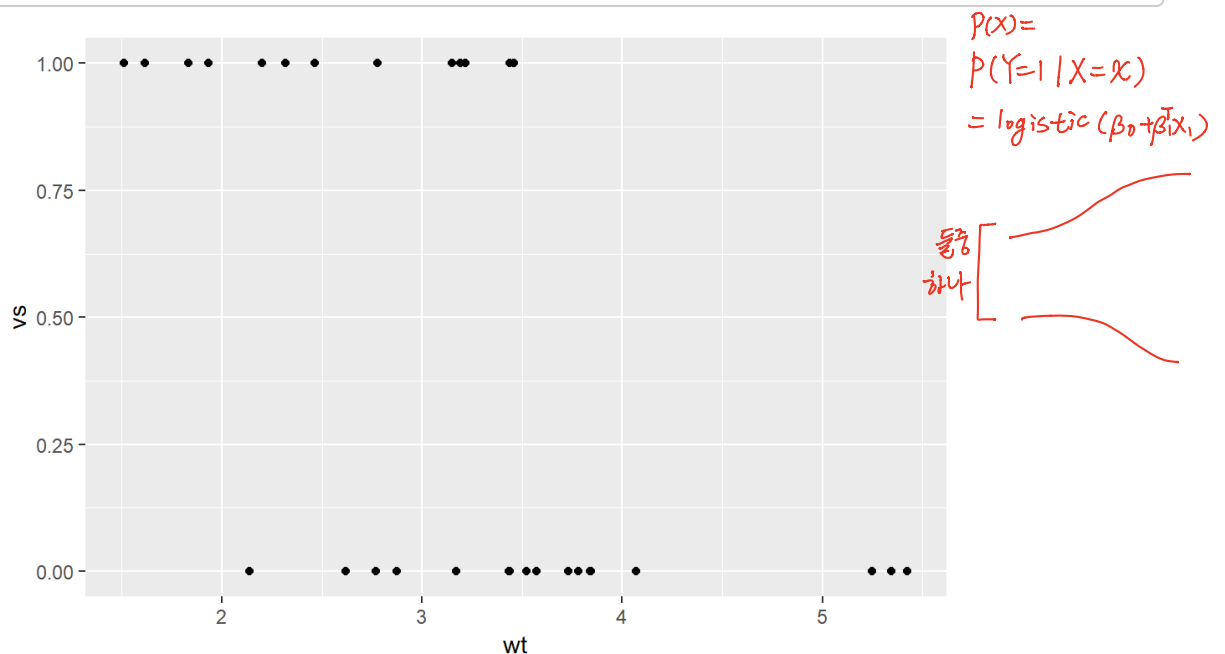
Motor Trend Car Road Tests

The data was extracted from the 1974 Motor Trend US magazine, and comprises fuel consumption and 10 aspects of automobile design and performance for 32 automobiles (1973-74 models). We will look at the following two variables:

- wt : Weight (1000 lbs)
- vs : Engine (0 = V-shaped, 1 = straight)

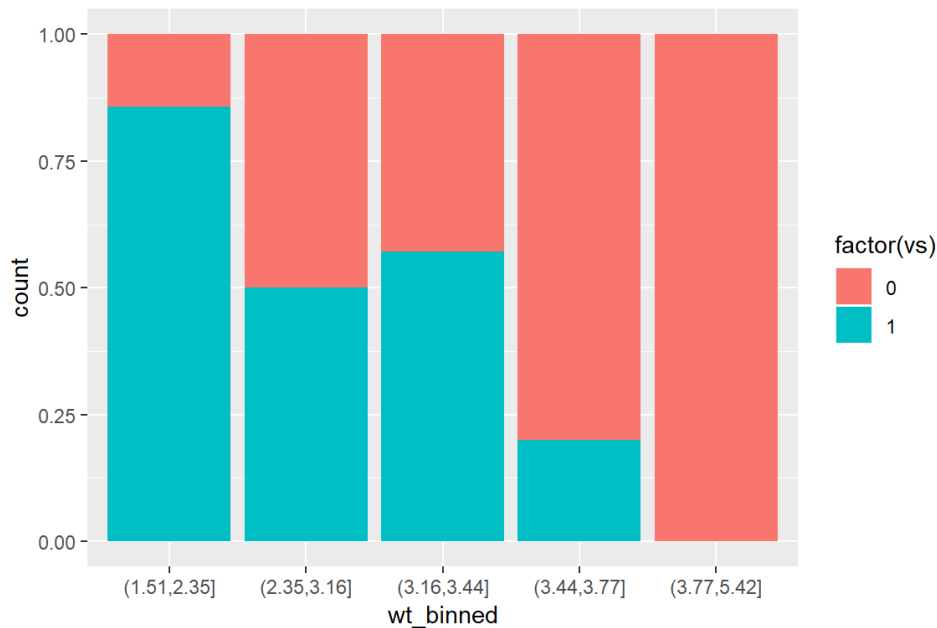
The data set is simple enough to visualize. What are we modeling here?

```
data(mtcars)
mtcars %>% ggplot(aes(x = wt, y = vs)) + geom_point()
```



The conditional probability is easier to see if we bin the observations.

```
wt_f <- unique(quantile(mtcars$wt, probs = seq(0,1, by = 0.2)))
wt_f[1] = wt_f - (1e-10)
mtcars %>% mutate(wt_binned = cut(wt, breaks = wt_f) ) %>% ggplot(aes(x = wt_binned, fill = factor(vs))) + geom_bar(position = "fill")
```

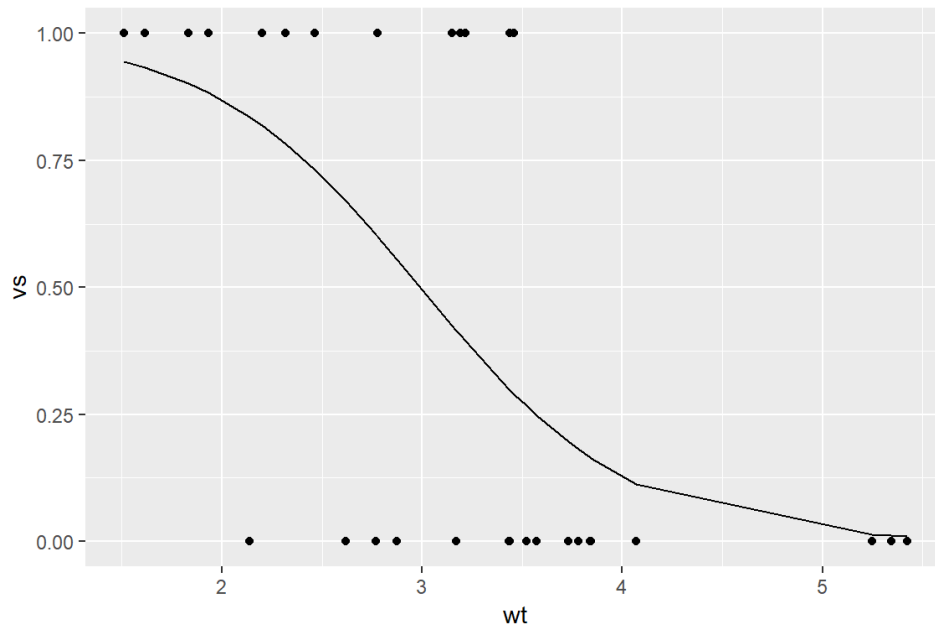


```
out <- glm(vs ~ wt, data = mtcars, family = "binomial")
summary(out)
```

```
##
## Call:
## glm(formula = vs ~ wt, family = "binomial", data = mtcars)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.9003  -0.7641  -0.1559   0.7223   1.5736
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  $\beta_0$  5.7147      2.3014   2.483  0.01302 *
## wt           $\beta_1$  -1.9105      0.7279  -2.625  0.00867 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 43.860  on 31  degrees of freedom
## Residual deviance: 31.367  on 30  degrees of freedom
## AIC: 35.367
##
## Number of Fisher Scoring iterations: 5
```

1. What is the model formula? *Conditional Mean을 Model화 하는 것. $E(Y | X=x) = P(Y=1 | X=x) = p(x) = \text{logistic}(\beta_0 + \beta_1 x)$*
 $\log(\text{Odds}(Y=1 | X=x)) = \beta_0 + \beta_1 x$
2. What are the odds of having a straight engine (vs = 1) if wt = 4000 lbs?
3. How does the odds change if the weight of the car increases by 1000 lbs? *$\log \text{odds} + \beta_1$*
 $\text{odds}(e\beta_1)$
4. Under which value of wt are the odds even? *$\log 1 = \beta_0 + \beta_1 x = 0$*
 $\rightarrow \text{odds}=1$
5. Is there a simple explanation for the relation between β_1 and the conditional probability?
*wt가 1만큼 증가하면 logodds가 변함. but 확률은 위에서 얘기한 두 값다.
 β_1 이 전방적인 slope을 얘기하지만, 구역별로 기울기가 달라서 명백히 설명X.
기원점을 잡고 그 구간 내에서 설명해야 한다.*

```
mtcars %>% ggplot(aes(x = wt, y = vs)) +
  geom_point() +
  geom_line(aes(x = wt, y = out$fitted.values)) +
  labs(main = "Data with fitted probability")
```



A bit more involved example

Western Collaborative Group Study

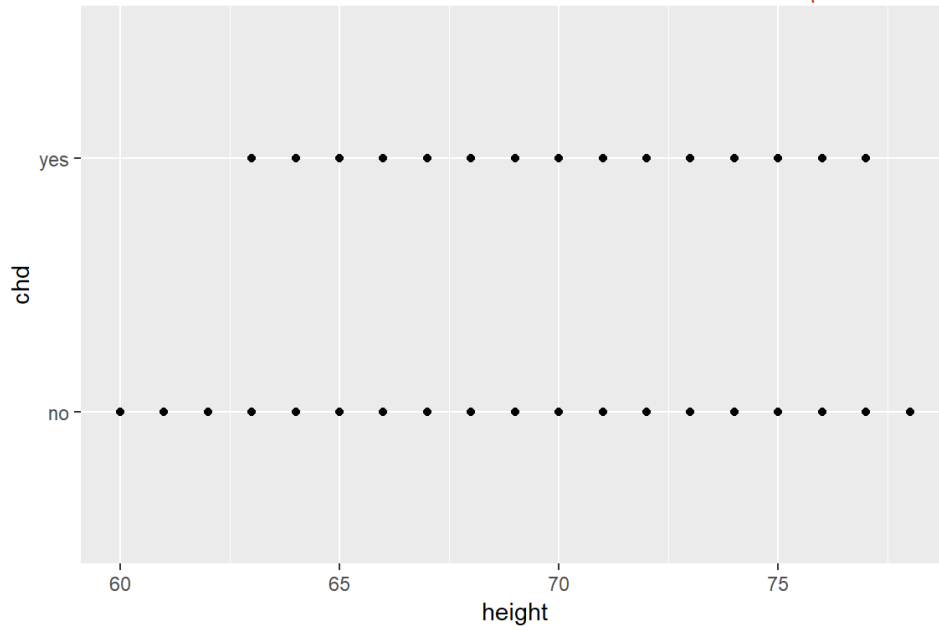
3154 healthy young men aged 39-59 from the San Francisco area were assessed for their personality type. All were free from coronary heart disease at the start of the research. Eight and a half years later change in this situation was recorded.

- chd : coronary heart disease developed is a factor with levels no yes
- dibep : behavior type a factor with levels A (Agressive) B (Passive)
- height : height in inches (→ - inches)

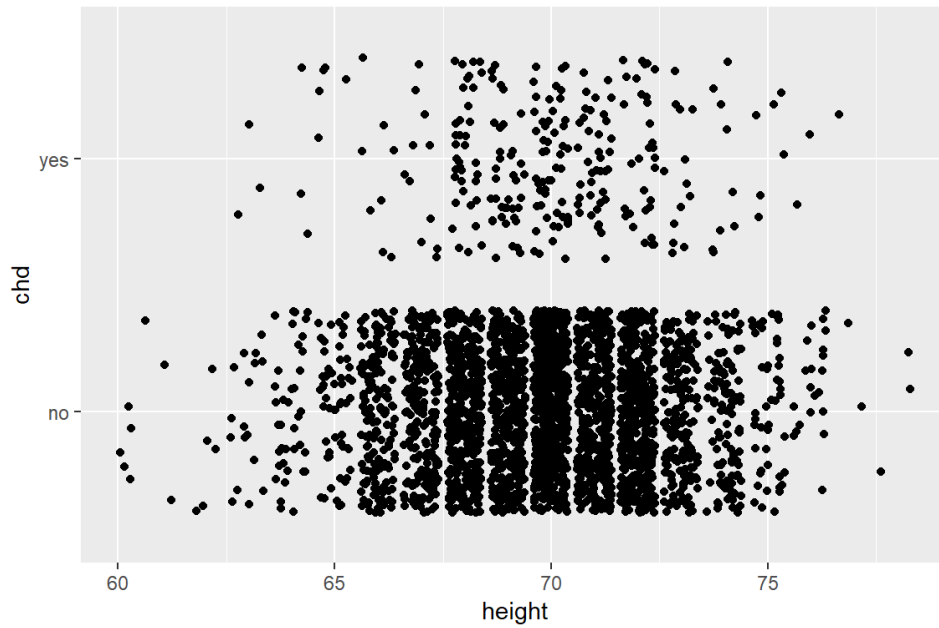
Model chd ~ height

```
library(faraway)
data(wcgs)
wcgs %>% ggplot(aes(x = height, y = chd)) + geom_point() # What's going on here?
```

→ 결구이 정음의 정려가 있다.

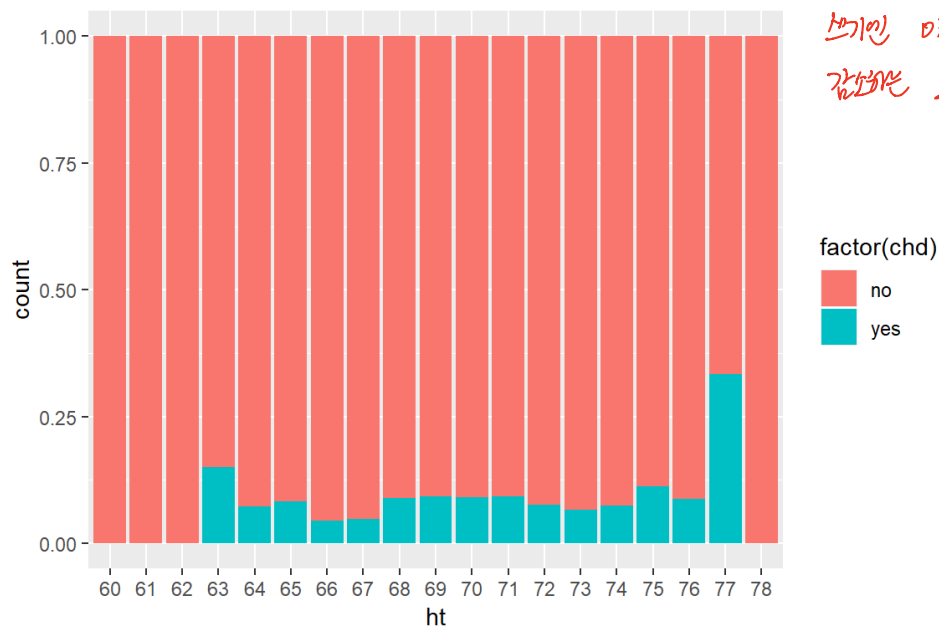


```
wcgs %>% ggplot(aes(x = height, y = chd)) + geom_jitter()
```



```
wcgs %>% mutate(ht = factor(height)) %>% ggplot(aes(x = ht, fill = factor(chd))) + geom_bar(position = "fill")
```

쓰기인 매체. 일정하게 증가하루
감소하는 모형을 안보임.



```
out <- glm(chd ~ height, data = wcgs, family = "binomial")
summary(out)
```

```
##
## Call:
## glm(formula = chd ~ height, family = "binomial", data = wcgs)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.4587  -0.4186  -0.4131  -0.4024   2.3157
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -4.33732    1.81231  -2.393   0.0167 *
## height       0.02742    0.02590   1.058   0.2899
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 1781.2  on 3153  degrees of freedom
## Residual deviance: 1780.1  on 3152  degrees of freedom
## AIC: 1784.1
##
## Number of Fisher Scoring iterations: 5
```

```
height_grd <- seq(min(wcgs$height), max(wcgs$height), 1)
height_pred <- predict(out, newdata = data.frame(height = height_grd), type = "response")
tibble(x = height_grd, y = height_pred)
```

```
## # A tibble: 19 x 2
##       x     y
##   <dbl> <dbl>
## 1    60 0.0634
## 2    61 0.0651
## 3    62 0.0668
## 4    63 0.0685
## 5    64 0.0703
## 6    65 0.0721
## 7    66 0.0739
## 8    67 0.0758
## 9    68 0.0778
## 10   69 0.0798
## 11   70 0.0818
## 12   71 0.0839
## 13   72 0.0860
## 14   73 0.0882
## 15   74 0.0904
## 16   75 0.0927
## 17   76 0.0950
## 18   77 0.0974
## 19   78 0.0998
```

$0.06 < y < 0.1$ 정도

\Rightarrow logistic Regression 이 잘 안맞는 data.

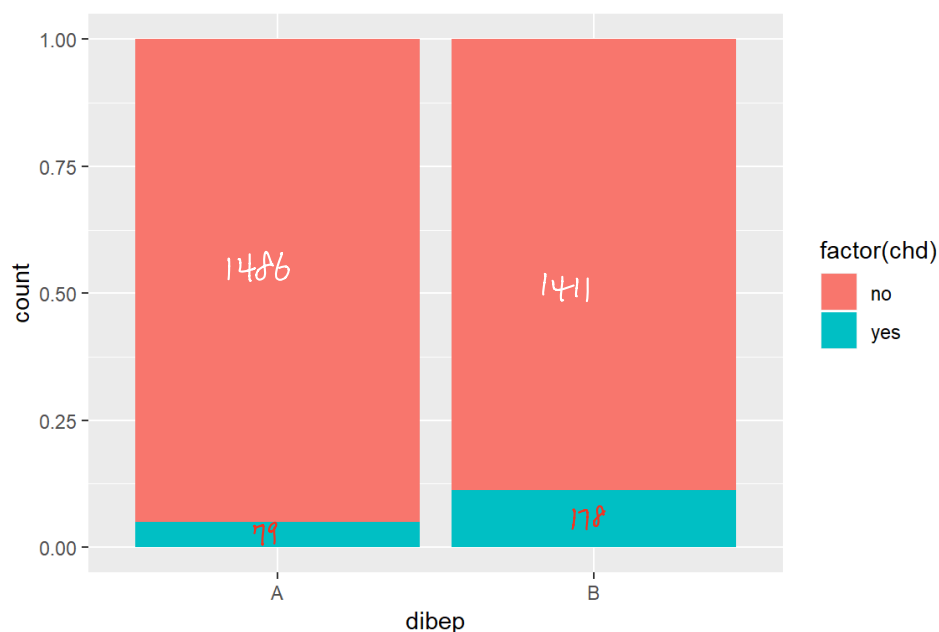
Model chd ~ dibep

Both variables are binary. Logistic regression becomes quite simple.

```
with(wcgs, table(chd,dibep))
```

```
##      dibep
## chd      A      B
## no  1486 1411
## yes   79  178
```

```
wcgs %>% ggplot(aes(x = dibep, fill = factor(chd))) + geom_bar(position = "fill")
```




```
out <- glm(chd ~ dibep, data = wogs, family = "binomial")
summary(out)
```

```
##
## Call:
## glm(formula = chd ~ dibep, family = "binomial", data = wogs)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.4875  -0.4875  -0.3219  -0.3219   2.4438
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  -2.9344     0.1155 -25.416  < 2e-16 ***
## dibepB (factor)  0.8641     0.1402   6.163  7.12e-10 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 1781.2  on 3153  degrees of freedom
## Residual deviance: 1740.3  on 3152  degrees of freedom
## AIC: 1744.3
##
## Number of Fisher Scoring iterations: 5
```

type A = 0, type B = 1.

```
unique(out$fitted.values)
```

```
## [1] 0.11202014 0.05047923
```

(A, B 두 점만 표시)

1. What is the model?
2. Can you interpret β_1 ?
3. What is the probability of coronary heart disease when the behavior type is passive? Aggressive?