

1    **TITLE**

2    Multidimensional scaling informed by *F*-ratio for hypothesis testing in microbial community analysis

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11   **SUMMARY**

12   Multidimensional scaling (MDS) is a common dimension reduction method to capture a global pattern  
13   and structure in microbial community. While a confirmatory analysis with the classical MDS is a  
14   promising approach in broadening an interpretation of the multivariate data (e.g., building a classifier,  
15   hypothesis testing), little attention is gained because the configuration can be biased towards the external  
16   information. Here we propose a confirmatory MDS informed by an *F*-ratio addressing a community-  
17   structural hypothesis testing result under a binary setting. Using a simulated or a 16S rRNA microbial  
18   community example, we present how the proposed configuration incorporates the testing result. We  
19   evaluate a performance by comparing to the previous confirmatory MDS, demonstrating that our method  
20   is less dependent on a selection of a model hyperparameter and minimally alters the classical MDS  
21   configuration. Our method proposes a broader applicability of MDS in modern biological and  
22   multivariate data analysis.

23   **INTRODUCTION**

24 Biotechnological advances in the past several decades have expanded a size and features of the  
25 multivariate data, necessitating a dimensionality reduction as a tool for the interpretation. By extracting an  
26 essential information from the biological data, the dimensionality reduction seeks a visual representation  
27 of the multivariate in a lower dimensional space. To retain a consistent data structure while performing  
28 the dimension reduction, a configuration is sought in a way that preserves dispersion or dissimilarity  
29 between samples, a process called as the multidimensional scaling (MDS). Compared to other nonlinear  
30 methods in dimensionality reduction [1-3], MDS is known to retain a global structure and represent a  
31 long-range interaction between samples, allowing its popularity for a long time since its inception.

32 In a classical MDS, the configuration is determined in a way that minimizes the difference between the  
33 dissimilarity in the original and the low dimensional space, a measure termed the stress. In microbial  
34 ecology community where the input data are compositional (i.e., a sample is comprised of different  
35 species expressing a level of abundance), the dissimilarity is measured by a difference in expression level  
36 of species within samples. For example, when interpreting a microbial community dataset such as 16S  
37 rRNA gene expression, a distance metric so-called the Unifrac [4] is used, and it allows to incorporate  
38 phylogenetic diversity as well as the compositional differences. The choice of an appropriate dissimilarity  
39 metric remains an important criterion for processing the input data and represent them in the low  
40 dimensional space.

41 In addition to visualizing the community structure with the dimensionality reduction, a quantitative  
42 analysis is carried out using the statistical inference such as hypothesis testing. Because the compositional  
43 abundances do not assume a probabilistic distribution *a priori* (highly skewed, zero-inflated [5]), a  
44 nonparametric model is preferred over parametric models for constructing a statistic. When the response  
45 is assumed to be independent between the samples, the statistic can be readily obtained by permuting the  
46 data labels [6]. For testing a difference in the ecological groups based on the sample dispersion, a  
47 (pseudo)  $F$  statistic has widely been used for each permutation [5, 7].

48 While hypothesis testing provides a quantitative perspective to understand the multivariate data, it should  
49 be noted that the testing result does not account for the MDS configuration. This is because most MDS  
50 (e.g., PCA, PCoA) do not consider data labels, whereas the hypothesis testing aims to infer whether  
51 dependent variables (e.g., class, label) are influenced by distribution of each sample group. For example,  
52 a configuration from the classical MDS is not able to explain a small but statistically meaningful  
53 difference between groups of different treatment.

54 The insufficient explanation by the classical MDS encourages to revise the approach and to address a  
55 structural hypothesis by including an external information, conferred by the responses or class [8] labels.  
56 Because the stress function is non-convex and an optimization algorithm can produce several local  
57 configurations, there is a notion allowing an altered configuration up to a point where its stress does not  
58 deviate too much from the classical configuration [9]. Broadly termed as the confirmatory MDS, it  
59 imposes an external constraint to the classical MDS carrying over an additional task to minimizing the  
60 stress.

61 In the recent confirmatory MDS methods, an objective function is constructed by adding a confirmatory  
62 term to the stress, former of which can be quantified by the labels then multiplied by a hyperparameter.  
63 While these confirmatory MDS methods have successfully visualized the multivariate structure in a way  
64 that differentiates each sample group with a discriminative purpose [10, 11], choosing a proper  
65 hyperparameter remains as a bottleneck towards a broader application of these inventive methods. For  
66 example, setting a high hyperparameter results in an undesirable stress and a misleading configuration  
67 distorted from the original.

68 In this study we propose an alternative MDS informed by a hypothesis testing inferences under a binary  
69 class setting. Our approach is motivated by a purpose to explain a statistical difference between groups, if  
70 any, using the MDS combined with a confirmatory analysis. Because the method does not target to  
71 directly discriminate between groups, the motivation distinguishes itself from previously proposed

72 confirmatory MDS. Furthermore, by characterizing the proposed framework we show that the revised  
73 configuration is less dependent to the choice of the model hyperparameter, mediating the previous issue  
74 with the distortion.

75 **METHODS**

76 **Problem formulation**

77 Consider a balanced design where the number of total observations is  $N$ , and each observation  $x_i$  is  $S$ -  
78 dimensional, pertaining to a set of labels  $y_i \in \{0,1\}$  for every  $i = 1 \dots N$ . Based on the observations  $(x_1, \dots$   
79  $x_N)$ , a distance matrix is obtained as  $\mathbf{d} = [d_{ij}] \in \mathbb{R}^{N \times N}$ . Now given the distance  $\mathbf{d}$ , we seek a two-  
80 dimensional configuration  $\mathbf{z} = (z_1, \dots z_N) \in \mathbb{R}^{N \times 2}$ , that best represents the original dimension by the  
81 following criteria.

82 **Classical MDS**

83 In classical MDS, a configuration is realized by minimizing the following objective function.

$$O(\mathbf{z}) = \frac{1}{2} \sum_{i,j} (d_{ij} - \|\mathbf{z}_i - \mathbf{z}_j\|_2)^2 \quad (1)$$

84  
85 In other words, the configuration  $\mathbf{z}$  is obtained in a way that tries to preserve a distance between a pair of  
86 observations  $(x_i, x_j)$  for each  $i, j \in N$ . Note that Equation (1) does not contain any terms related to  $y_i$ ,  
87 meaning that classical MDS does not consider the class labels.

88 **Confirmatory MDS for F-informed hypothesis testing**

89 *Hypothesis testing for non-parametric multivariate analysis of variance*

90 When testing a statistical difference between groups in multivariate analysis, the group variance is the  
91 measure of interest. It is represented by the  $F$ -statistic, a ratio between two group variances, each

92 respectively derived from across- and within-group. While the conventional  $F$ -test requires an assumption  
 93 that each observation follows a normal distribution, the convention is generalized by introducing an  
 94 analogous statistic (pseudo  $F$ -ratio) which is combined with label permutation for quantifying a statistical  
 95 significance. In this non-parametric approach [7], the pseudo  $F$ -ratio is defined as

$$F = \frac{\sum_{i,j} d_{ij}^2 - 2 \sum_{i,j} \mathbb{I}\{y_i = y_j\} d_{ij}^2}{2 \sum_{i,j} \mathbb{I}\{y_i = y_j\} d_{ij}^2} \cdot (N - 2), \quad (2)$$

96

97 where  $\mathbb{1}\{\cdot\}$  denotes an indicator function. Since the pseudo  $F$ -ratio does not follow an  $F$ -distribution  
 98 under the relaxed model assumption, it is instead evaluated by an empirical distribution that is created by  
 99 ‘permuting’ the labels. That is, in every permutation a new  $F$ -ratio,  $F^\pi$ , is obtained from the data structure,  
 100 and by repeating this, we have a  $P$ -value written as

$$P = \frac{\text{Number of case where } (F^\pi \geq F)}{\text{Number of total repeat}}. \quad (3)$$

101

102 Known as the permutational multivariate analysis of variance (PERMANOVA), the hypothesis testing  
 103 method has a broad application in microbial community analysis.

104 *Proposed MDS*

105 Now we propose a new multidimensional scaling that incorporates a hypothesis testing result in the  
 106 multivariate setting. This is enabled by adding a confirmatory term to the classical MDS (Eq. (1)), giving  
 107 an objective function as

$$O(\mathbf{z}) = \underbrace{\frac{1}{2} \sum_{i,j} (d_{ij} - \|\mathbf{z}_i - \mathbf{z}_j\|_2)^2}_{\text{MDS term}} + \lambda \cdot \underbrace{\frac{1}{2} \left| \sum_{i,j} [1 - (f_{\mathbf{z}}(\Phi_o) + 1) \mathbb{I}\{y_i = y_j\}] \|\mathbf{z}_i - \mathbf{z}_j\|_2^2 \right|}_{\text{Confirmatory term}} \quad (4)$$

108

109 where  $\Phi_o$  is a ratio constant expressed in terms of the distance  $\mathbf{d}$  and labels  $y_i$  and is defined as

110

$$\Phi_o(\mathbf{d}, \mathbf{y}) := \frac{\sum_{i,j} \mathbb{I}\{y_i \neq y_j\} d_{ij}^2}{\sum_{i,j} \mathbb{I}\{y_i = y_j\} d_{ij}^2},$$

111 and  $f_{\mathbf{z}}(\Phi_o) : \mathbf{R} \rightarrow \mathbf{R}$  is a mapping function which is determined by the configuration  $\mathbf{z}$ . An exact derivation  
 112 of  $f_{\mathbf{z}}$  and a detailed description on the confirmatory term in Equation (4) is described in Appendix B.  
 113 Given Equation (4), we want to find an optimal configuration  $\mathbf{z}^*$  such that  $\mathbf{z}^* = \operatorname{argmin}_{(z_1, \dots, z_N)} O(\mathbf{z})$ .

114 *Algorithm for the proposed MDS*

115 Because the confirmatory term in (4) is strictly convex in terms of  $\mathbf{z}$ , we are able to minimize  $O(\mathbf{z})$  by  
 116 using the Majorize-Minimization (MM) algorithm, a typical approach in the MDS optimization task [12].  
 117 While the implementation of MM algorithm is described in detail in Appendix C, we provide its update  
 118 rule as below.

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**Algorithm 1** MM algorithm for pseudo  $F$ -informed MDS
 

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For epoch  $t$  and every  $i = 1, \dots, N$ ,

$$\begin{aligned} \mathbf{z}_i^{[t+1]} &\leftarrow \frac{2}{2(N-1) + \lambda \delta(\mathbf{z}_i^{[t]})(N - (N-2)f_{\mathbf{z}}(\Phi))} \\ &\times \left[ (1 + \lambda \delta(\mathbf{z}_i^{[t]})) \sum_{\substack{j=1 \\ \epsilon_{ij}=0}}^N \mathbf{z}_j^{[t]} + (1 - \lambda f_{\mathbf{z}}(\Phi) \delta(\mathbf{z}^{[t]})) \sum_{\substack{j=1 \\ \epsilon_{ij}=1}}^N \mathbf{z}_j^{[t]} + \sum_{j=1}^N d_{ij} \frac{\mathbf{z}_i^{[t]} - \mathbf{z}_j^{[t]}}{\|\mathbf{z}_i^{[t]} - \mathbf{z}_j^{[t]}\|_2} \right], \end{aligned}$$

where  $\epsilon_{ij} = \mathbb{I}\{y_i = y_j\}$ ,  $\delta_i(\mathbf{z}) = \operatorname{sign} \sum_{j=1}^N [1 - (f_{\mathbf{z}}(\Phi) + 1)\epsilon_{ij}] \|\mathbf{z}_i - \mathbf{z}_j\|_2^2$ , with an initial value  
 119 obtained from a classical MDS.

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120 **RESULTS AND DISCUSSION**

121 Using Algorithm 1, we sought to determine how well our proposed MDS approach can produce a two-  
 122 dimensional configuration in simulated and experimental data. We then assess its performance by  
 123 comparing ours to a recent method in confirmatory MDS [13].

124 **Simulated data**

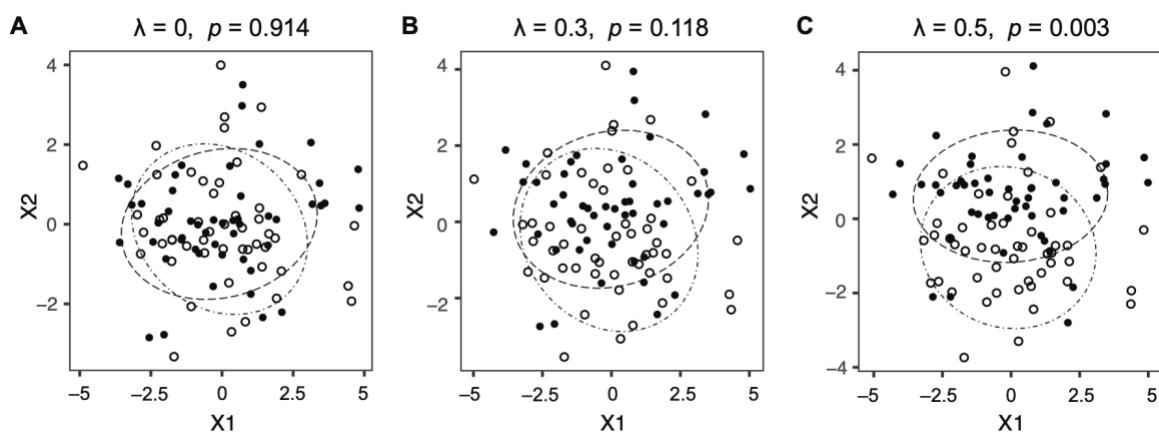
125 We first provide a representing case where our method can be useful in visualizing multidimensional data.

126 To do this we consider a binary labeled dataset where each group originates in a different multivariate  
 127 Gaussian distribution. In a three-dimensional setting, for example, consider a balanced design where an  
 128 observation expressed as

$$x_i \sim \begin{cases} \mathcal{N} \left( [0, 0, 0]^T, \begin{bmatrix} 3 & 0 & 0 \\ 0 & 3 & 0 \\ 0 & 0 & 1 \end{bmatrix} \right), & i = 1, 2, \dots, 50 \\ \mathcal{N} \left( [0, 0, 1]^T, \begin{bmatrix} 3 & 0 & 0 \\ 0 & 3 & 0 \\ 0 & 0 & 1 \end{bmatrix} \right), & i = 51, 52, \dots, 100. \end{cases} \quad (5)$$

129

130 As expected, PERMANOVA testing result indicates there is a statistically significant difference between  
 131 the groups with pseudo  $F = 5.402$  and  $p = 0.005$ . However, a classical MDS does not distinguish groups  
 132 in two-dimensional configuration (Figure 1a) with  $p = 0.914$ , because the difference is in the third  
 133 dimension with the lowest variance among the principal diagonals (Figure 6 in Appendix C).  
  
 134 On the other hand, our MDS configuration is able to display the difference between the groups, and it  
 135 becomes clearer when a hyperparameter  $\lambda$  is large (Figure 1b,c). The distinctions are also verified by a  
 136 low p-value resulting from PERMANOVA test using the two-dimensional configurations.



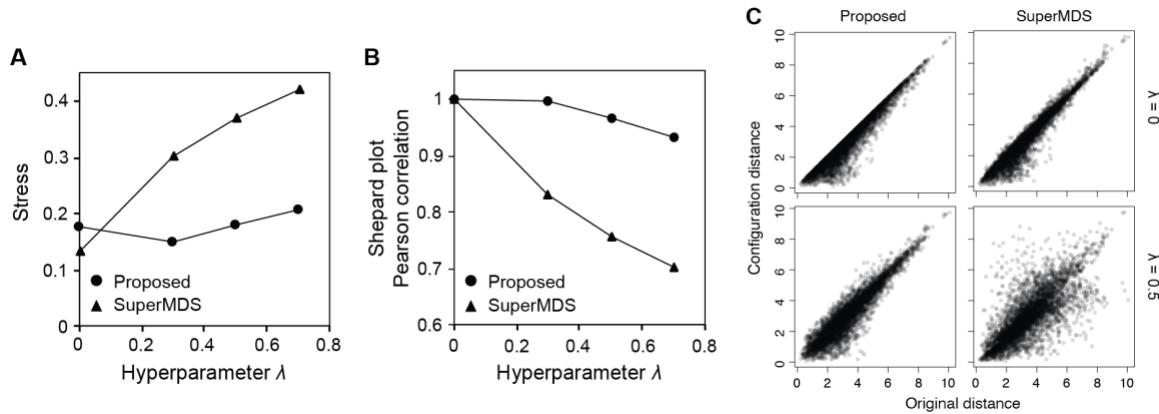
137

138 Figure 1: Two-dimensional visualization of proposed MDS with a hyperparameter (a)  $\lambda = 0$  (classical  
 139 MDS), (b)  $\lambda = 0.3$ , and (c)  $\lambda = 0.5$ . For each configuration, a p-value is given based on PERMANOVA  
 140 test. An ellipse is drawn for each group with a confidence interval of 80%.

141 We next evaluate the performance of our method to existing confirmatory MDS by calculating a stress,  
 142 which is defined as

$$143 \text{ Stress} = \frac{\sum_{i,j} (d_{ij} - \|z_i - z_j\|_2)^2}{\sum_{i,j} d_{ij}^2},$$

144 or by calculating a Spearman correlation from Shepard diagram, as measures of evaluating the  
 145 performance of MDS [14]. As a result, our proposed MDS visualization presents a Stress  $\sim 0.2$  regardless  
 146 of a choice of a hyperparameter  $\lambda$  (Figure 2a), suggesting either configuration can be used for visualizing  
 147 the simulated data [15]. This is in contrast to the previous method [13] where the stress increases as  $\lambda$   
 148 becomes large, implying their approach distinguishes groups at the expense of the original distance  
 149 structure. Similarly, in Shepard plot with a choice of  $\lambda$ , our method presents a more consistent correlation  
 150 of the sample pair distance in between three- and two-dimension (Figure 2b,2c, and 7 in Appendix A).



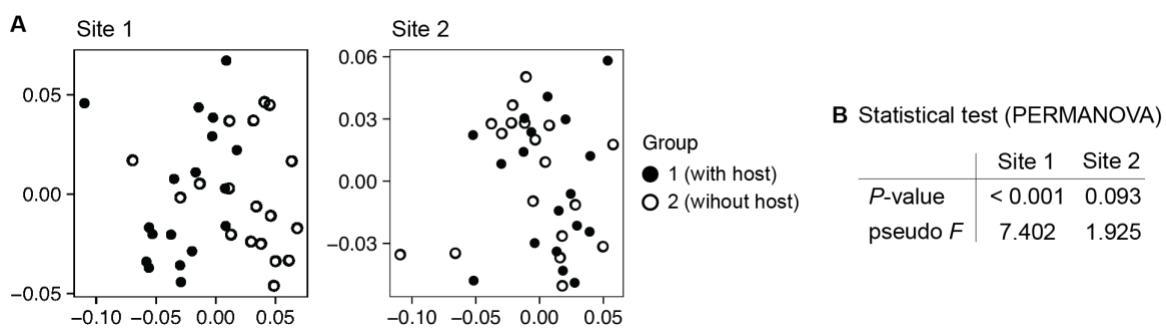
151  
 152 Figure 2: Performance of proposed MDS compared to an existing label-informed confirmatory MDS  
 153 (SuperMDS [13]) by using (a) Stress and (b) Spearman correlation from the simulated data. (c) Shepard

154 plot of the proposed MDS comparing to SuperMDS for a hyperparameter  $\lambda = 0.5$ . More results are  
155 displayed in Figure 7.

156 It is worth noting that proposed MDS is invariant by the choice of the hyperparameter, which has not  
157 been observed in the existing confirmatory MDS methods.

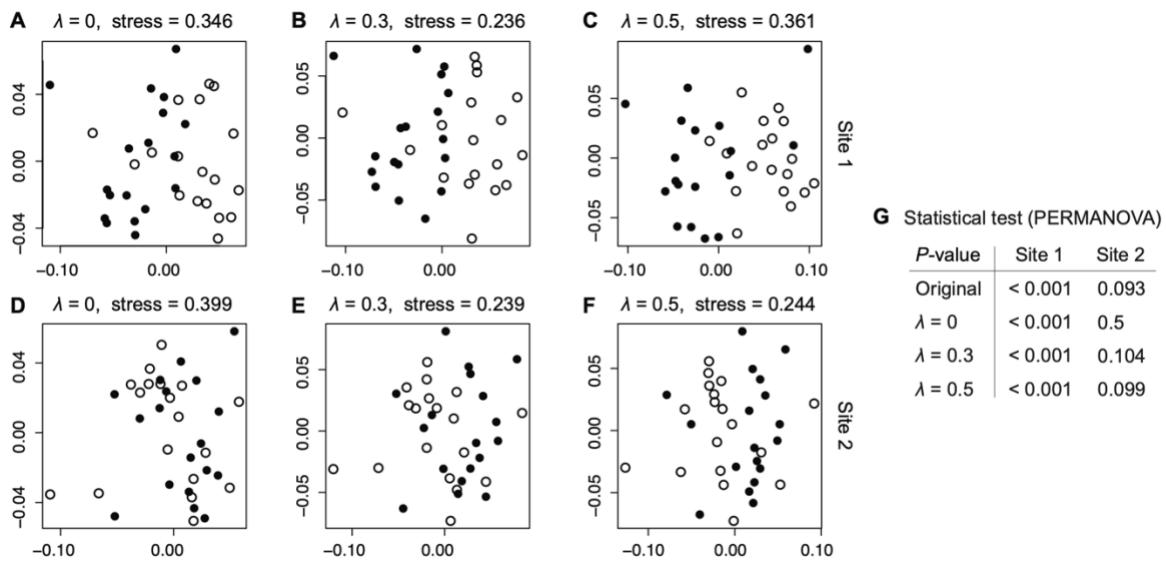
158 **Microbial community dataset**

159 Next, we provide an example where biological hypothesis testing result is conveyed to the MDS  
160 configuration. We take microbial community dataset containing thirty-six, balanced samples of a binary  
161 label (e.g., with or without a presence of microbial host) [16]. In detail, each data represents expression  
162 levels 16S rRNA gene of 72 bacterial taxa, and the distance between samples is measured using the  
163 weighted Unifrac [17]. Particular attention is made on these datasets, that the classical two-dimensional  
164 MDS configuration does not explain PERMANOVA test results on a group differences. As shown in  
165 Figure 3a, groups in site 1 are dispersed in a different location whereas site 2 groups are not, when  
166 visualized using the classical MDS. In both sites, however, moderately small P-values are obtained ( $< 0.1$ ,  
167 Figure 3b), indicating the group difference in the community structure is, in fact, statistically significant.



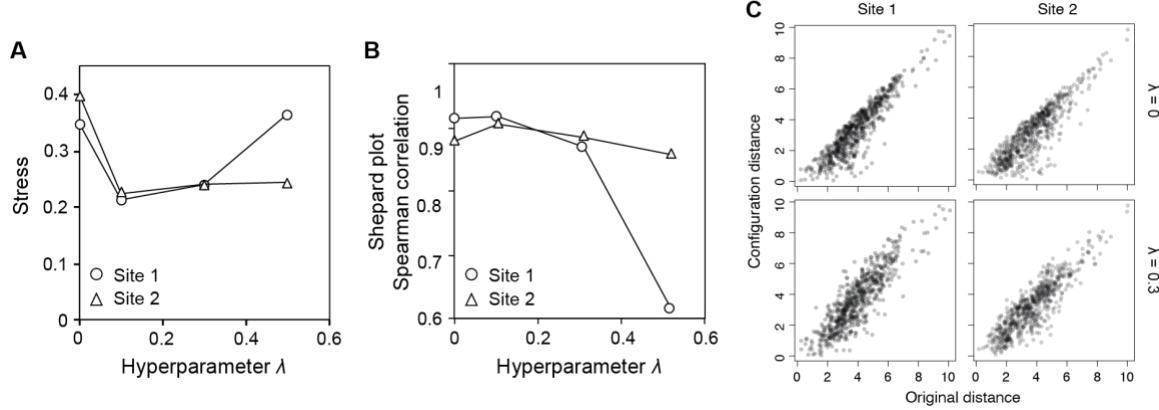
169 Figure 3: (a) Multidimensional scaling and (b) statistical test result on a group difference between two  
170 sample groups for each site. The sample presents a microbial community mea- sured by 16S rRNA gene  
171 expression.

172 Using the community dataset, we present a configuration with the proposed MDS visualization. As  
 173 expected, for site 1 community samples the configuration retains its distinction between the class labels  
 174 regardless of the choice of the hyperparameter  $\lambda$  (Figure 4a-c). Moreover, for site 2 samples we observe a  
 175 higher distinction between the groups with increasing  $\lambda$  (Figure 4d-f). The observation with the  
 176 visualizations is justified by a quantitative measure using P -value calculated on the 2D configurations  
 177 (Figure 4g).



178  
 179 Figure 4: Two-dimensional configuration of microbial community samples using the proposed MDS  
 180 method, where samples are collected from (a-c) site 1 and (d-f) site 2. (g) Statistical significance on the  
 181 group difference between two treatments using PERMANOVA test.

182 We then evaluate the performance of the proposed MDS using stress measurement and Shepard plot.  
 183 Again, we observe stress does not strictly depend on the hyperparameter  $\lambda$  or even show a decreased  
 184 value when  $\lambda$  is nonzero (e.g., 0.1, 0.3) compared to the classical MDS ( $\lambda = 0$ , Figure 5a). Shepard plot  
 185 and Spearman correlation also show that the configurations nicely preserve the original distance in the  
 186 microbial community data, except for a case when the largest  $\lambda$  is set to site 1 community.



187

188 Figure 5: Evaluation of the proposed MDS using microbial community data, measured by (a) stress, (b)  
189 Spearman correlation, and (c) Shepard plot from each sample site. More results are displayed in Figure 8  
190 (Appendix A).

191 **CONCLUSION**

- 192 • A new multidimensional scaling method which incorporates an F-statistic-informed hypothesis  
193 testing is proposed.
- 194 • We find that the performance of the proposed MDS excels existing MDS methods for addressing  
195 class labels, as evaluated by its stress and Shepard plot, validated using both simulated and real  
196 datasets.
- 197 • The proposed MDS can be useful when analyzing a biological dataset with F -informed hypothesis  
198 testing, providing informative and precise dimension reduction, especially for visualization.
- 199 • The method is less dependent on the choice of hyperparameter when producing the configuration.  
200 This lessens the risk of overreliance on the class labels in that the data are automatically grouped to a  
201 suitable degree. Also, users may avoid the hassle of hyperparameter selection using such as cross-  
202 validation.

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