

# TP5: Fisher Information

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Credits: Vincenzo Schimmenti

Recall that for a distribution  $p(x|\theta)$  with parameters  $\theta = \{\theta_1, \dots, \theta_n\}$  the Fisher information matrix reads:

$$I_{i,j}(\theta) = -\mathbb{E} \left[ \frac{\partial^2 \log p(x|\theta)}{\partial \theta_i \partial \theta_j} \right] \quad (1)$$

where the expected value is taken w.r.t. to  $p(x|\theta)$ , keeping  $\theta$  fixed. The Cramer-Rao bound for an unbiased estimator  $\hat{\theta}$  of a parameter  $\theta$  is:

$$\text{Var}(\hat{\theta}) \geq \frac{1}{NI(\theta)} \quad (2)$$

where  $N$  is the number of i.i.d. samples in the estimator. Recall that an unbiased estimator  $\hat{\theta}$  is such that:

$$\mathbb{E}[\hat{\theta} - \theta] = 0 \quad (3)$$

where  $\theta$  is the true parameter.

**Example:** Exponential distribution  $p(x) = \lambda e^{-\lambda x}$ . The log probability is:

$$-\log p(x|\lambda) = -\log \lambda + \lambda x$$

By taking two times the derivatives w.r.t.  $\lambda$  we get:

$$-\partial_\lambda^2 \log p(x|\lambda) = \frac{1}{\lambda^2}$$

Since the second derivative does not depend on  $x$ :

$$\mathbb{E} [-\partial_\lambda^2 \log p(x|\lambda)] = \frac{1}{\lambda^2}$$

Hence the Fisher information reads:

$$I(\lambda) = \frac{1}{\lambda^2}$$

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If we want rephrase the Fisher information using the average parameter of the exponential distribution  $\mu = 1/\lambda$  we need to use the change of parameter property of the Fisher information, namely:

$$I(\eta) = I(\theta(\eta)) \left( \frac{d\theta}{d\eta} \right)^2$$

By choosing  $\theta = \lambda$  and  $\eta = \mu$  we have:

$$\frac{d\lambda(\mu)}{d\mu} = -1/\mu^2$$

and

$$I(\mu) = I(\lambda(\mu)) \left( \frac{d\lambda(\mu)}{d\mu} \right)^2 = 1/\mu^2$$

When we estimate  $\mu$  from data using Maximum Likelihood approach we know that the unbiased estimator for  $\mu$  is:

$$\hat{\mu} = \frac{1}{N} \sum_{k=1}^N x_k$$

If the true parameter is  $\mu$ , the Variance of the estimator is:

$$\text{Var}(\hat{\mu}) = \frac{\mu^2}{N}$$

By substituting our results in the Cramer-Rao bound:

$$\frac{\mu^2}{N} \geq \frac{\mu^2}{N}$$

In this case the inequality is saturated.

**Problem 1.** Compute the Fisher information matrix for a Gaussian distribution with parameters  $\mu$  and  $\sigma^2$ :

$$p(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2} \quad (4)$$

The matrix is given by three elements:

$$I_{\mu, \mu} = -\mathbb{E} \left[ \frac{\partial^2 \log p(x|\mu, \sigma^2)}{\partial \mu^2} \right] \quad (5)$$

$$I_{\sigma^2, \sigma^2} = -\mathbb{E} \left[ \frac{\partial^2 \log p(x|\mu, \sigma^2)}{\partial (\sigma^2)^2} \right] \quad (6)$$

$$I_{\mu, \sigma^2} = -\mathbb{E} \left[ \frac{\partial^2 \log p(x|\mu, \sigma^2)}{\partial \mu \partial (\sigma^2)} \right] \quad (7)$$

We treat  $\sigma^2$  as the parameter so we take derivatives w.r.t.  $\sigma^2$  and not  $\sigma$  (since two parameters  $\sigma$  and  $-\sigma$  would characterize the same model).

- Repeat the discussion about the Cramer-Rao bound for the parameter  $\mu$ . Recall that an unbiased estimator for  $\mu$  is:

$$\hat{\mu} = \frac{1}{N} \sum_{k=1}^N x_k$$

- (NOT MANDATORY) Do the same for the parameter  $\sigma^2$  knowing that the unbiased estimator is:

$$\hat{\sigma}^2 = \frac{1}{N-1} \sum_{k=1}^N (x_k - \hat{\mu})^2$$

For solving the second point, it is necessary to compute the variance of  $\hat{\sigma}^2$  which might not be straightforward.

**Problem 2.** Repeat the same analysis for a Bernoulli random variable:

$$P(X = 0|\theta) = 1 - \theta \quad (8)$$

$$P(X = 1|\theta) = \theta \quad (9)$$

Remember that the unbiased estimator for  $\theta$  is:

$$\hat{\theta} = \frac{1}{N} \sum_{k=1}^N x_k$$

where  $x_k = 0, 1$ . The variance of  $X$  w.r.t. the Bernoulli distribution is:

$$\begin{aligned} \text{Var}(X) &= \sum_{x=0,1} x^2 p(X=x|\theta) - \left( \sum_{x=0,1} x p(X=x|\theta) \right)^2 = \\ &= \theta(1 - \theta) \end{aligned}$$

## Comments