# Calibration of Degrees of Freedom: Global Financial Cycle Project Part-time RA for Wenbin Wu

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### 1 Introduction

From the previous estimation (1. Load Data and Model Preparations, 2. Estimate the Gaussian Mode, 3. Get Posterior Draws, 4. Impulse Response Plots, 5. Produce Tables), we have already got the sample distribution of residuals for the Gaussian-errors model, say, e and  $e_{scaled}$  (residuals up to scaling by Lambda).

Noticing that  $e_{scaled}$  is the output generated directly from the MCMC process, and e is adjusted by variance scaling factors in different stages, say,  $lambda_scaling$ . This may cause confusion in subsequent variable names, but it is worth noting that the error without the scaled subscript is the scaled error. To respect the original code, the subscripts here remain the same, but readers need to distinguish them.

Next, assuming that e has an independent Student-t distribution with a degrees of freedom and unit scale, I am about to calibrate the degrees of freedom using the sample distribution of e and  $e_{scaled}$ .

## 2 Econometric Framework

The likelihood function of t-distribution is,

$$f(x\,|\,\theta) = \frac{\Gamma(\frac{\theta+1}{2})}{\sqrt{\theta\pi}\Gamma(\frac{\theta}{2})}(1+\frac{x^2}{\theta})^{\frac{\theta+1}{2}}$$

where  $\theta$  is the degrees of freedom (DoF).

The scaled log-likelihood function of t-distribution is,

$$ll(\theta) = \sum f(e_i \,|\, \theta) = \sum \left[ \log \Gamma\left(\frac{\theta+1}{2}\right) - \log \Gamma\left(\frac{\theta}{2}\right) - \frac{1}{2}\log(\theta\pi) - \frac{\theta+1}{2}\log\left(1 + \frac{e^2}{\theta}\right) \right]$$

The log-likelihood function of t-distribution is,

$$ll_{scaled}(\theta) = \sum f(e_{scaled,\,i} \,|\, \theta) = \sum \left\lceil \log \Gamma\left(\frac{\theta+1}{2}\right) - \log \Gamma\left(\frac{\theta}{2}\right) - \frac{1}{2}\log(\theta\pi) - \frac{\theta+1}{2}\log\left(1 + \frac{e_{scaled}^2}{\theta}\right) \right\rceil$$

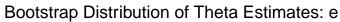
# 3 Calibration of DoF using R: Boostrapping & MLE

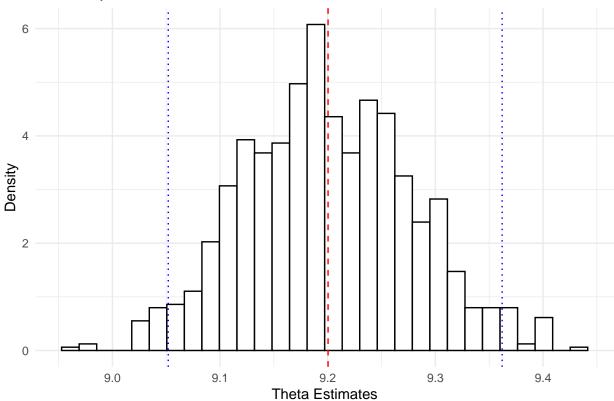
```
## 1.2 MLE
### 1.2.1 Log-likelihood function for unscaled residuals
loglike <- function(theta, e) {</pre>
    if (theta <= 0)</pre>
        return(-Inf) # Ensure theta is positive
    11 \leftarrow sum(log(gamma((theta + 1)/2)) - log(theta * pi)/2 - log(gamma(theta/2)) -
        (theta + 1)/2 * log(1 + e^2/theta))
    return(11)
}
neg.loglike <- function(theta, e) -loglike(theta, e)</pre>
### 1.2.2 Define Wrapper Function for Optimization
neg.loglike.wrap <- function(e) {</pre>
    function(theta) {
        neg.loglike(theta, e)
    }
}
## 1.3 Bootstrapping
set.seed(2045)
n_bootstrap <- 1000 # Number of bootstrap samples</pre>
bootstrap_results <- numeric(n_bootstrap) # Store bootstrap results</pre>
for (i in 1:n_bootstrap) {
    # Resample the data
    e bootstrap <- sample(e, length(e), replace = TRUE)</pre>
    # Optimize using resampled data
    result <- optim(par = df, fn = neg.loglike.wrap(e_bootstrap), method = "BFGS")
    # Store the estimated parameter
    bootstrap_results[i] <- result$par</pre>
}
## 1.4 Calculate the mean and confidence interval of theta
theta_median <- mean(bootstrap_results)</pre>
theta_ci <- quantile(bootstrap_results, c(0.025, 0.975))
# 2. Calibration: MLE 2.1 Setup Load the data generated from model_choices =
# 'qaussian'
df <- theta_median # Starting point of degrees of freedom (DoF) of t-distribution
## 2.2 MLE 2.2.1 Log-likelihood function for unscaled residuals
loglike <- function(theta) {</pre>
    if (theta <= 0)</pre>
        return(-Inf) # Ensure theta is positive
    11 \leftarrow sum(log(gamma(theta + 1)/2)) - log(theta * pi)/2 - log(gamma(theta/2)) -
        (theta + 1)/2 * log(1 + e^2/theta))
    return(11)
}
neg.loglike <- function(theta) -loglike(theta)</pre>
```

```
### 2.3.1 Calibration
set.seed(3045)
n_iterations <- 10  # Maximum number of MLE iterations</pre>
mle_results <- numeric(n_iterations) # Store MLE results</pre>
tolerance <- 1e-10 # Convergence threshold
for (i in 1:n iterations) {
    est <- bbmle::mle2(neg.loglike, start = list(theta = df))</pre>
   new_df <- est@coef # Use the result as the new starting point</pre>
   mle_results[i] <- new_df</pre>
    cat("Iteration", i, ": theta =", new df, "\n")
    # Check for convergence
   if (i > 1 && abs(new_df - mle_results[i - 1]) < tolerance) {</pre>
        cat("Convergence achieved after", i, "iterations.\n")
        break
   }
   df <- new_df # Update df for the next iteration</pre>
}
Iteration 1 : theta = 9.200688
Iteration 2: theta = 9.200888
Iteration 3: theta = 9.200996
Iteration 4: theta = 9.201055
Iteration 5: theta = 9.201086
Iteration 6: theta = 9.201103
Iteration 7: theta = 9.201105
Iteration 8 : theta = 9.201113
Iteration 9: theta = 9.201113
Iteration 10: theta = 9.201118
## 3.1 Output: Bootstrapping 3.1.1 Median and Confidence Intervals
cat("Median of theta:", theta_median, "\n")
Median of theta: 9.200318
cat("95% Confidence Interval of theta:", theta_ci, "\n")
95% Confidence Interval of theta: 9.051872 9.362034
### 3.1.2 Histogram
library(ggplot2)
bootstrap_df <- data.frame(theta = bootstrap_results)</pre>
ggplot(bootstrap_df, aes(x = theta)) + geom_histogram(aes(y = after_stat(density)),
   bins = 30, color = "black", fill = "white") + labs(title = "Bootstrap Distribution of
    → Theta Estimates: e",
   x = "Theta Estimates", y = "Density") + geom_vline(xintercept = theta_median,
    color = "red", linetype = "dashed") + geom_vline(xintercept = theta_ci[1], color =

    "blue",

   linetype = "dotted") + geom_vline(xintercept = theta_ci[2], color = "blue", linetype
    theme_minimal()
```





```
## 3.2 Output: MLE
cat("MLE Method:\n")
```

### MLE Method:

```
cat("Theta estimates across iterations:", mle_results, "\n")
```

Theta estimates across iterations:  $9.200688 \ 9.200888 \ 9.200996 \ 9.201055 \ 9.201086 \ 9.201103 \ 9.201105 \ 9.201$  cat("Final Theta estimate:", df, "\n")

Final Theta estimate: 9.201118