# Representing Uncertainty for Probabilistic Inference

Chapter 13

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#### Probabilistic Inference

How do we use probabilities in AI?

- You wake up with a headache
- Do you have the flu?
- H = headache, F = flu

Logical Inference: if *H* then *F* 

(but the world is usually not this simple)

Statistical Inference: compute the probability of a query/diagnosis/decision given (i.e., conditioned on) evidence/symptom/observation, i.e.,  $P(F \mid H)$ 

[Example from Andrew Moore]

# Uncertainty in the World

- An agent can often be uncertain about the state of the domain since there is often ambiguity and uncertainty
- Plausible/probabilistic inference
  - I've got this evidence; what's the chance that this conclusion is true?
    - I've got a sore neck; how likely am I to have meningitis?
    - A mammogram test is positive; what's the probability that the patient has breast cancer?

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# Uncertainty

• Say we have a rule:

**if** toothache **then** problem is cavity

• But not all patients have toothaches due to cavities, so we could set up rules like:

if toothache and ¬gum-disease and ¬filling and ...
then problem = cavity

• This gets complicated; better method:

if toothache then problem is cavity with 0.8 probability or  $P(cavity \mid toothache) = 0.8$ 

the probability of cavity is 0.8 given toothache is observed

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# Uncertainty in the World and our Models

- True uncertainty: rules are probabilistic in nature
  - quantum mechanics
  - rolling dice, flipping a coin
- Laziness: too hard to determine exception-less rules
  - takes too much work to determine all of the relevant factors
  - too hard to use the enormous rules that result
- Theoretical ignorance: don't know all the rules
  - problem domain has no complete, consistent theory (e.g., medical diagnosis)
- Practical ignorance: do know all the rules BUT
  - haven't collected all relevant information for a particular case

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# **Probability Theory**

- Probability theory serves as a formal means for
  - Representing and reasoning with uncertain knowledge
  - Modeling degrees of belief in a proposition (event, conclusion, diagnosis, etc.)
- Probability is the "language" of uncertainty
  - A key modeling tool in modern Al

# Logics

Logics are characterized by what they use as "primitives"

Logic	What Exists in World	Knowledge States
Propositional	facts	true/false/unknown
First-Order	facts, objects, relations	true/false/unknown
Temporal	facts, objects, relations, times	true/false/unknown
Probability Theory	facts	degree of belief 01
Fuzzy	degree of truth	degree of belief 01

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# Sample Space

- A space of events in which we assign probabilities
- Events can be binary, multi-valued, or continuous
- Events are mutually exclusive
- Examples

– Coin flip: {head, tail}

- Die roll: {1, 2, 3, 4, 5, 6}

– English words: a dictionary

- High temperature tomorrow: {-100, ..., 100}

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#### Random Variable

- A variable, X, whose domain is a sample space, and whose value is (somewhat) uncertain
- Examples:

X = coin flip outcome

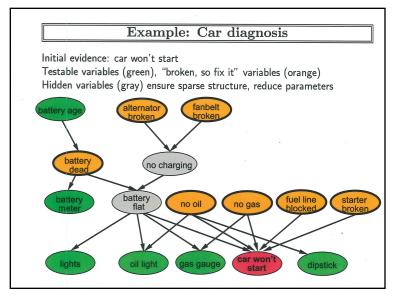
X = tomorrow's high temperature

- For a given task, the user defines a set of random variables for describing the world
- Each variable has a set of mutually exclusive and exhaustive possible values

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# **Probability for Discrete Events**

- An agent's uncertainty is represented by P(A=a) or simply P(a)
  - the agent's degree of belief that variable A takes on value a given no other information related to A
  - a single probability called an unconditional or prior probability



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# **Probability for Discrete Events**

- Examples
  - -P(head) = P(tail) = 0.5 fair coin
  - -P(head) = 0.51, P(tail) = 0.49 slightly biased coin
  - P(first word = "the" when flipping to a random
    page in R&N) = ?

• Book: The Book of Odds

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#### Source of Probabilities

- Frequentists
  - probabilities come from data
  - if 10 of 100 people tested have a cavity, P(cavity) = 0.1
  - probability means the fraction that would be observed in the limit of infinitely many samples
- Objectivists
  - probabilities are real aspects of the world
  - objects have a propensity to behave in certain ways
  - coin has propensity to come up heads with probability 0.5
- Subjectivists
  - probabilities characterize an agent's belief
  - have no external physical significance

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# **Probability Table**

Weather

sunny	cloudy	rainy
200/365	100/365	65/365

- P(Weather = sunny) = P(sunny) = 200/365
- **P**(*Weather*) = \( 200/365, 100/365, 65/365 \)
- We'll obtain the probabilities by counting frequencies from data

# **Probability Distributions**

Given A is a RV taking values in  $\langle a_1, a_2, \dots, a_k \rangle$ 

e.g., if *A* is *Sky*, then value is one of *<clear*, *partly\_cloudy*, *overcast>* 

- P(a) represents a single probability where A=a
  - e.g., if A is Sky, then P(a) means any one of P(clear),  $P(partly\_cloudy)$ , P(overcast)
- **P**(*A*) represents a probability distribution
  - the set of values:  $\langle P(a_1), P(a_2), ..., P(a_k) \rangle$
  - If A takes n values, then P(A) is a set of n probabilities
     e.g., if A is Sky, then P(Sky) is the set of probabilities:
     \(\setminus P(clear), P(partly \) cloudy), P(overcast)\(\rightarrow\)
  - Property:  $\sum P(a_i) = P(a_1) + P(a_2) + ... + P(a_k) = 1$ 
    - sum over all values in the domain of variable A is 1 because the domain is mutually exclusive and exhaustive

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# The Axioms of Probability

- 1.  $0 \le P(A) \le 1$
- 2. P(true) = 1, P(false) = 0
- 3.  $P(A \lor B) = P(A) + P(B) P(A \land B)$

**Note**: Here P(A) means P(A=a) for some value a and  $P(A \lor B)$  means  $P(A=a \lor B=b)$ 

# The Axioms of Probability • $0 \le P(A) \le 1$ • P(true) = 1, P(false) = 0• $P(A \lor B) = P(A) + P(B) - P(A \land B)$ Sample The fraction of A can't space be smaller than 0

The Axioms of Probability •  $0 \le P(A) \le 1$ The fraction of A can't be bigger than 1 P(true) = 1, P(false) = 0  $P(A \lor B) = P(A) + P(B) - P(A \land B)$ Sample space

The Axioms of Probability

•  $0 \le P(A) \le 1$ 

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- P(true) = 1, P(false) = 0
- $P(A \lor B) = P(A) + P(B) P(A \land B)$

Sample space Valid sentence: e.g., "X=head or X=tail"

The Axioms of Probability •  $0 \le P(A) \le 1$ • P(true) = 1, P(false) = 0 $P(A \lor B) = P(A) + P(B) - P(A \land B)$ 

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Invalid sentence:

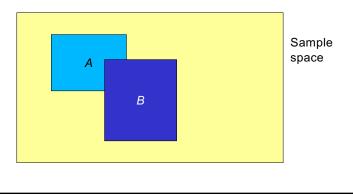
e.g., "X=head AND X=tail"

Sample

space

# The Axioms of Probability

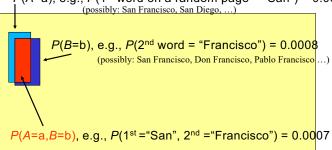
- $0 \le P(A) \le 1$
- P(true) = 1, P(false) = 0
- $P(A \lor B) = P(A) + P(B) P(A \land B)$



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# Joint Probability

The joint probability P(A=a, B=b) is shorthand for P(A=a ∧ B=b), i.e., the probability of both A=a and B=b happening P(A=a), e.g., P(1st word on a random page = "San") = 0.001 (possibly: San Francisco, San Diego, ...)



Some Theorems
Derived from the Axioms

- $P(\neg A) = 1 P(A)$
- If A can take k different values  $a_1, ..., a_k$ :  $P(A=a_1) + ... + P(A=a_k) = 1$
- $P(B) = P(B \land \neg A) + P(B \land A)$ , if A is a binary event
- $P(B) = \sum_{i=1...k} P(B \land A=a_i)$ , if A can take k values

Called Addition or Conditioning rule

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# Full Joint Probability Distribution (FJPD)

#### Weather

		sunny	cloudy	rainy
Temp	hot	150/365	40/365	5/365
	cold	50/365	60/365	60/365

- P(Temp=hot, Weather=rainy) = P(hot, rainy) = 5/365 = 0.014
- The full joint probability distribution table for n random variables, each taking k values, has k<sup>n</sup> entries

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## Full Joint Probability Distribution (FJPD)

Bird	Flier	Young	Probability
Т	Т	Т	0.0
Т	Т	F	0.2
Т	F	Т	0.04
Т	F	F	0.01
F	Т	Т	0.01
F	Т	F	0.01
F	F	Т	0.23
F	F	F	0.5

3 Boolean random variables  $\Rightarrow 2^3 - 1 = 7$  "degrees of freedom" (DOF) or "independent values"

Sums to 1

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# **Unconditional / Prior Probability**

- One's uncertainty or original assumption about an event prior to having any data about it or anything else in the domain
- P(Coin = heads) = 0.5
- P(Bird = T) = 0.0 + 0.2 + 0.04 + 0.01 = 0.22
- Compute from the FJPD by marginalization

## Computing from the FJPD

- Marginal Probabilities
  - -P(Bird=T) = P(bird) = 0.0 + 0.2 + 0.04 + 0.01 = 0.25
  - $-P(bird, \neg flier) = 0.04 + 0.01 = 0.05$
  - $-P(bird \lor flier) = 0.0 + 0.2 + 0.04 + 0.01 + 0.01 + 0.01 = 0.27$
- Sum over all other variables
- "Summing Out"
- "Marginalization"

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# Marginal Probability

#### Weather

		sunny	cloudy	rainy
Тетр	hot	150/365	40/365	5/365
TOTTIP	cold	50/365	60/365	60/365
	Σ	200/365	100/365	65/365

**P**(Weather) = (200/365, 100/365, 65/365)

Probability *distribution* for r.v. *Weather* 

The name comes from the old days when the sums were written in the margin of a page

# Marginal Probability

#### Weather

		sunny	cloudy	rainy	$\sum_{i}$
Temp	hot	150/365	40/365	5/365	195/365
remp	cold	50/365	60/365	60/365	170/365

**P**(*Temp*) = \langle 195/365, 170/365 \rangle

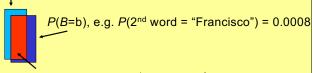
This is nothing but  $P(B) = \sum_{i=1...k} P(B \land A = a_i)$ , where A can take k values

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# **Conditional Probability**

The conditional probability  $P(A=a \mid B=b)$  is the fraction of time A=a, within the region where B=b

P(A=a), e.g.  $P(1^{st}$  word on a random page = "San") = 0.001



 $P(A=a \mid B=b)$ , e.g.  $P(1^{st}="San" \mid 2^{nd}="Francisco") = ?$  (possibly: San, Don, Pablo ...)

# **Conditional Probability**

- Conditional probabilities
  - formalizes the process of accumulating evidence and updating probabilities based on new evidence
  - specifies the belief in a proposition (event, conclusion, diagnosis, etc.) that is conditioned on a proposition (evidence, feature, symptom, etc.) being true
- $P(a \mid e)$ : conditional probability of A=a given E=e evidence is all that is known true

$$P(a \mid e) = P(a \land e) / P(e) = P(a, e) / P(e)$$

– conditional probability can viewed as the joint probability P(a, e) normalized by the prior probability, P(e)

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# **Conditional Probability**

P(s)=0.001

P(f)=0.0008

P(s,f)=0.0007

- P(san | francisco)
  - $= \#(1^{st} = s \text{ and } 2^{nd} = f) / \#(2^{nd} = f)$
  - = P(san ∧ francisco) / P(francisco)
  - = 0.0007 / 0.0008
  - = 0.875

P(B=b), e.g.  $P(2^{nd} \text{ word} = \text{"Francisco"}) = 0.0008$ 

 $P(A=a \mid B=b)$ , e.g.  $P(1^{st}="San" \mid 2^{nd}="Francisco") = 0.875$  (possibly: San, Don, Pablo ...)

Although "San" is rare and "Francisco" is rare, given "Francisco" then "San" is quite likely!

# **Conditional Probability**

Conditional probabilities behave exactly like standard probabilities; for example:

$$0 \le P(a \mid e) \le 1$$

conditional probabilities are between 0 and 1 inclusive

$$P(a_1 \mid e) + P(a_2 \mid e) + ... + P(a_k \mid e) = 1$$

conditional probabilities sum to 1 where  $a_1$ , ...,  $a_k$  are all values in the domain of random variable A

$$P(\neg a \mid e) = 1 - P(a \mid e)$$

negation for Boolean random variable A

**Computing Conditional Probability** 

$$P(\neg B \mid F) = ?$$

$$P(F) = ?$$

Note:  $P(\neg B \mid F)$  means  $P(B=\text{false} \mid F=\text{true})$  and P(F) means P(F=true)

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# **Full Joint Probability Distribution**

Bird (B)	Flier (F)	Young (Y)	Probability
Т	Т	Т	0.0
Т	Т	F	0.2
Т	F	Т	0.04
Т	F	F	0.01
F	Т	Т	0.01
F	Т	F	0.01
F	F	Т	0.23
F	F	F	0.5

3 Boolean random variables  $\Rightarrow 2^3 - 1 = 7$  "degrees of freedom" or "independent values"

Sums to 1

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# **Computing Conditional Probability**

$$P(\neg B \mid F)$$
 =  $P(\neg B, F)/P(F)$   
=  $(P(\neg B, F, Y) + P(\neg B, F, \neg Y))/P(F)$   
=  $(0.01 + 0.01)/P(F)$ 

$$P(F) = P(F, B, Y) + P(F, B, \neg Y) + P(F, \neg B, Y) + P(F, \neg B, \neg Y)$$

$$= 0.0 + 0.2 + 0.01 + 0.01$$

$$= 0.22$$
Marginalization

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# **Computing Conditional Probability**

- Instead of using Marginalization to compute P(F), can alternatively use Normalization:
- $P(\neg B \mid F) = .02/P(F)$  from previous slide
- $P(\neg B | F) + P(B | F) = 1$  by definition
- P(B|F) = P(B,F)/P(F) = (0.0 + 0.2)/P(F)
- So, 0.02/P(F) + 0.2/P(F) = 1
- Hence, P(F) = 0.22

Normalization

• In general,  $P(A \mid B) = \alpha P(A, B)$ where  $\alpha = 1/P(B) = 1/(P(A, B) + P(\neg A, B))$ 

•  $P(Q \mid E_1, ..., E_k) = \alpha P(Q, E_1, ..., E_k)$ =  $\alpha \sum_{Y} P(Q, E_1, ..., E_k, Y)$ 

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# Conditional Probability with Multiple Evidence

$$P(\neg B \mid F, \neg Y) = P(\neg B, F, \neg Y) / P(F, \neg Y)$$
  
=  $P(\neg B, F, \neg Y) / (P(\neg B, F, \neg Y) + P(B, F, \neg Y))$   
= .01 /(.01 + .2)  
= 0.048

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# **Conditional Probability**

- $P(X_1=x_1, ..., X_k=x_k \mid X_{k+1}=x_{k+1}, ..., X_n=x_n) =$ sum of all entries in FJPD where  $X_1=x_1, ..., X_n=x_n$  divided by sum of all entries where  $X_{k+1}=x_{k+1}, ..., X_n=x_n$
- But this means in general we need the entire FJPD table, requiring an exponential number of values to do probabilistic inference (i.e., compute conditional probabilities)

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#### The Chain Rule

From the definition of conditional probability we have

$$P(A, B) = P(B) * P(A \mid B) = P(A \mid B) * P(B)$$

• It also works the other way around:

$$P(A, B) = P(A) * P(B | A) = P(B | A) P(A)$$

• It works with more than 2 events too:

$$P(A_1, A_2, ..., A_n) =$$
  
 $P(A_1) * P(A_2 | A_1) * P(A_3 | A_1, A_2) * ...$   
 $* P(A_n | A_1, A_2, ..., A_{n-1})$ 

Called "Chain Rule"

Called

"Product

Rule"

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### Example

Statistical Inference: Compute the probability of a diagnosis, *F*, given symptom, *H*, where *H* = "has a headache" and *F* = "has flu"

That is, compute  $P(F \mid H)$ 

You know that

• P(H) = 0.1 "one in ten people has a headache"

• P(F) = 0.01 "one in 100 people has flu"

•  $P(H \mid F) = 0.9$  "90% of people who have flu have a headache"

[Example from Andrew Moore]

#### **Probabilistic Reasoning**

How do we use probabilities in AI?

- · You wake up with a headache
- Do you have the flu?
- H = headache, F = flu



Logical Inference: if H then F

(but the world is usually not this simple)

Statistical Inference: compute the probability of a query/diagnosis/decision given (i.e., conditioned on) evidence/symptom/observation, i.e.,  $P(F \mid H)$ 

[Example from Andrew Moore]

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# Inference with Bayes's Rule

Thomas Bayes, "Essay Towards Solving a Problem in the Doctrine of Chances." 1764

$$P(F \mid H) = P(F,H) = \frac{P(H \mid F)P(F)}{P(H)}$$
Def of cond. prob.

• P(H) = 0.1 "one in ten people has a headache"

- P(F) = 0.01 "one in 100 people has flu"
- P(H|F) = 0.9 "90% of people who have flu have a headache"
- P(F|H) = (0.9 \* 0.01) / 0.1 = 0.09
- So, there's a 9% chance you have flu much less than 90%
- But it's higher than P(F) = 1% since you have a headache

# Bayes's Rule

- Bayes's Rule is the basis for probabilistic reasoning given a prior model of the world, P(Q), and a new piece of evidence, E, Bayes's rule says how this piece of evidence decreases our ignorance about the world
- Initially, know P(Q) ("prior")
- Update after knowing E ("posterior"):

$$P(Q|E) = P(Q) \frac{P(E|Q)}{P(E)}$$

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# Bayes's Rule in Practice $\begin{array}{c} P_{d_1}(s_1,\overline{s}_8...s_1) = P_{d_1}P_{s_1}|d_1^{(1-P_{s_8}|d_1)...P_{s_1}|d_1} \\ P_{d_1}(s_1,\overline{s}_8...s_1) = P_{d_1}P_{s_1}|d_1^{(1-P_{s_8}|d_1)...P_{s_1}|d_1^{(1-P_{s_8}|d_1)...P_{s_1}|d_1^{(1-P_{s_8}|d_1)} \\ P_{d_1}(s_1,\overline{s}_8..$

# Inference with Bayes's Rule

#### $P(A \mid B) = P(B \mid A)P(A) / P(B)$

Bayes's rule

- Why do we make things this complicated?
  - Often P(B|A), P(A), P(B) are easier to get
- Some terms:
  - Prior: P(A): probability of A before any evidence
  - Likelihood: P(B|A): assuming A, how likely is the evidence B
  - Posterior: P(A|B): probability of A after knowing evidence B
  - (Deductive) Inference: deriving an unknown probability from known ones



# **Summary of Important Rules**

• Conditional Probability: P(A | B) = P(A,B)/P(B)

• Product rule:  $P(A,B) = P(A \mid B)P(B)$ 

• Chain rule: P(A,B,C,D) = P(A|B,C,D)P(B|C,D)P(C|D)P(D)

• Conditionalized version of Chain rule:

P(A,B|C) = P(A|B,C)P(B|C)

• Bayes's rule: P(A|B) = P(B|A)P(A)/P(B)

• Conditionalized version of Bayes's rule:

 $P(A \mid B, C) = P(B \mid A, C)P(A \mid C)/P(B \mid C)$ 

• Addition / Conditioning rule:  $P(A) = P(A,B) + P(A,\neg B)$ 

 $P(A) = P(A \mid B)P(B) + P(A \mid \neg B)P(\neg B)$ 

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#### Quiz

- A doctor performs a test that has 99% reliability, i.e., 99% of people who are sick test positive, and 99% of people who are healthy test negative. The doctor estimates that 1% of the population is sick.
- Question: A patient tests positive. What is the chance that the patient is sick?
- 0-25%, 25-75%, 75-95%, or 95-100%?

#### Common Mistake

• P(A) = 0.3 so  $P(\neg A) = 1 - P(A) = 0.7$ 

• P(A|B) = 0.4 so  $P(\neg A|B) = 1 - P(A|B) = 0.6$ because  $P(A|B) + P(\neg A|B) = 1$ 

**but**  $P(A|\neg B) \neq 0.6$  (in general) because  $P(A|B) + P(A|\neg B) \neq 1$  in general

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#### Quiz

- A doctor performs a test that has 99% reliability, i.e., 99% of people who are sick test positive, and 99% of people who are healthy test negative. The doctor estimates that 1% of the population is sick.
- Question: A patient tests positive. What is the chance that the patient is sick?
- 0-25%, 25-75%, 75-95%, or 95-100%?
- Common answer: 99%; Correct answer: 50%

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Given:

$$P(TP \mid S) = 0.99$$

TP = "tests positive" S = "is sick"

$$P(\neg TP \mid \neg S) = 0.99$$

$$P(S) = 0.01$$

Query:

$$P(S \mid TP) = ?$$

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# Inference with Bayes's Rule

- In a bag there are two envelopes
  - one has a red ball (worth \$100) and a black ball
  - one has two black balls. Black balls are worth nothing





- You randomly grab an envelope, and randomly take out one ball – it's black
- At this point you're given the option to switch envelopes. Should you switch or not?

Similar to the "Monty Hall Problem"

 $P(TP \mid S) = 0.99$   $P(\neg TP \mid \neg S) = 0.99$  P(S) = 0.01  $P(S \mid TP) =$   $P(TP \mid S) P(S) / P(TP)$  = (0.99)(0.01) / P(TP) = 0.0099/P(TP)  $P(\neg S \mid TP) = P(TP \mid \neg S)P(\neg S) / P(TP)$  = (1 - 0.99)(1 - 0.01) / P(TP) = 0.0099/P(TP) 0.0099/P(TP) + 0.0099/P(TP) = 1, so P(TP) = 0.0198So,  $P(S \mid TP) = 0.0099 / 0.0198 = 0.5$ 

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# Inference with Bayes's Rule

E: envelope, 1 = (R,B), 2 = (B,B)

B: the event of drawing a black ball

Given: P(B|E=1) = 0.5, P(B|E=2) = 1, P(E=1) = P(E=2) = 0.5

Query: Is P(E=1 | B) > P(E=2 | B)?

Use Bayes's rule: P(E|B) = P(B|E)\*P(E) / P(B)

P(B) = P(B|E=1)P(E=1) + P(B|E=2)P(E=2) = (.5)(.5) + (1)(.5) = .75

Conditioning rule

P(E=1|B) = P(B|E=1)P(E=1)/P(B) = (.5)(.5)/(.75) = 0.33

 $P(E=2 \mid B) = P(B \mid E=2)P(E=2)/P(B) = (1)(.5)/(.75) = 0.67$ 

After seeing a black ball, the posterior probability of this envelope being #1 (thus worth \$100) is *smaller* than it being #2

Thus you should switch!

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# Another Example

- 1% of women over 40 who are tested have breast cancer. 85% of women who really do have breast cancer have a positive mammography test (true positive rate). 8% who do not have cancer will have a positive mammography (false positive rate).
- Question: A patient gets a positive mammography test. What is the chance she has breast cancer?

 Let Boolean random variable M mean "positive mammography test"

• Let Boolean random variable *C* mean "has breast cancer"

• Given:

$$P(C) = 0.01$$

$$P(M|C) = 0.85$$

$$P(M|\neg C) = 0.08$$

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Compute the posterior probability: P(C|M)

- P(C|M) = P(M|C)P(C)/P(M) by Bayes's rule = (.85)(.01)/P(M)
- $P(M) = P(M|C)P(C) + P(M|\neg C)P(\neg C)$  by the Conditioning rule
- So, P(C|M) = .0085/[(.85)(.01) + (.08)(1-.01)]= 0.097
- So, there is only a 9.7% chance that if you have a positive test you really have cancer!

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# Independence

Two events *A*, *B* are **independent** if the following hold:

• 
$$P(A, B) = P(A) * P(B)$$

• 
$$P(A, \neg B) = P(A) * P(\neg B)$$

...

• 
$$P(A \mid B) = P(A)$$

• 
$$P(B \mid A) = P(B)$$

• 
$$P(A \mid \neg B) = P(A)$$

•••

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# Independence

- Given: P(B) = 0.001, P(E) = 0.002, P(B|E) = P(B)
- The full joint probability distribution table (FJPD) is:

Burglary	Earthquake	Prob.
В	Е	= P(B)P(E)
В	¬E	
¬В	Е	
¬В	¬E	

- Need only 2 numbers to fill in entire table
- Now we can do anything, since we have the FJPD

# Independence

- Independence is a kind of domain knowledge
  - Needs an understanding of *causation*
  - Very strong assumption
- Example: P(burglary) = 0.001 and P(earthquake) = 0.002
  - Let's say they are independent
  - The full joint probability table = ?

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# Independence

- Given n independent, Boolean random variables, the FJPD has 2<sup>n</sup> entries, but we only need n numbers (degrees of freedom) to fill in entire table
- Given *n* independent random variables, where each can take *k* values, the FJPD table has:
  - $-k^n$  entries
  - -Only n(k-1) numbers needed (DOFs)

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# Conditional Independence

- Random variables can be dependent, but conditionally independent
- Example: Your house has an alarm
  - Neighbor John calls when he hears the alarm
  - Neighbor Mary calls when she hears the alarm
  - Assume John and Mary don't talk to each other
- Is JohnCall independent of MaryCall?
  - No If John called, it is likely the alarm went off, which increases the probability of Mary calling
  - $-P(MaryCall \mid JohnCall) \neq P(MaryCall)$

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# Independence vs. Conditional Independence

- Say Alice and Bob each toss separate coins. A represents "Alice's coin toss is heads" and B represents "Bob's coin toss is heads"
- A and B are independent
- Now suppose Alice and Bob toss the same coin. Are A and B independent?
  - No. Say the coin may be biased towards heads. If A is heads, it will lead us to increase our belief in B beings heads. That is, P(B|A) > P(A)

# Conditional Independence

 But, if we know the status of the Alarm, JohnCall will not affect whether or not Mary calls

 $P(MaryCall \mid Alarm, JohnCall) = P(MaryCall \mid Alarm)$ 

- We say JohnCall and MaryCall are conditionally independent given Alarm
- In general, "A and B are conditionally independent given C" means:

$$P(A \mid B, C) = P(A \mid C)$$

$$P(B \mid A, C) = P(B \mid C)$$

$$P(A, B \mid C) = P(A \mid C) P(B \mid C)$$

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- Say we add a new variable, *C*: "the coin is biased towards heads"
- The values of A and B are dependent on C
- But if we know for certain the value of C (true or false), then any evidence about A cannot change our belief about B
- That is, P(B|C) = P(B|A, C)
- A and B are conditionally independent given C

# **Revisiting Earlier Example**

- Let Boolean random variable M mean "positive mammography test"
- Let Boolean random variable *C* mean "has breast cancer"
- Given:

P(C) = 0.01

P(M|C) = 0.85

 $P(M|\neg C) = 0.08$ 

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# Bayes's Rule with Multiple Evidence

• P(C|M1, M2) = P(M1, M2|C)P(C)/P(M1, M2)by Bayes's rule

=  $P(M1 \mid M2, C)P(M2 \mid C)P(C)/P(M1, M2)$ Conditionalized Chain rule

• P(M1, M2) = P(M1, M2 | C)P(C) +  $P(M1, M2 | \neg C)P(\neg C)$  by Conditioning rule = P(M1 | M2, C)P(M2 | C)P(C) + $P(M1 | M2, \neg C)P(M2 | \neg C)P(\neg C)$ 

by Conditionalized Chain rule

# Bayes's Rule with Multiple Evidence

- Say the same patient goes back and gets a second mammography and it too is positive.
   Now, what is the chance she has Cancer?
- Let M1, M2 be the 2 positive tests
- M1 and M2 are **not** independent
- Compute posterior: P(C|M1, M2)

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Cancer "causes" a positive test, so **M1** and **M2** are conditionally independent given **C**, so

```
• P(M1 | M2, C) = P(M1 | C) = 0.85
```

• P(M1, M2) = P(M1|M2, C)P(M2|C)P(C) + P(M1|M2, -C)P(M2|-C)P(-C)

= P(M1|C)P(M2|C)P(C) +

 $P(M1 | \neg C)P(M2 | \neg C)P(\neg C)$  by cond. indep.

= (.85)(.85)(.01) + (.08)(.08)(1-.01)

= 0.01356

So, P(C|M1, M2) = (.85)(.85)(.01)/.01356= 0.533 or 53.3%

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# Example

- Prior probability of having breast cancer:
   P(C) = 0.01
- Posterior probability of having breast cancer after 1 positive mammography:

P(C|M1) = 0.097

 Posterior probability of having breast cancer after 2 positive mammographies (and cond. independence assumption):

P(C|M1, M2) = 0.533

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# Bayes's Rule with Multiple Evidence

```
• P(C|M1, \neg M2) = P(M1, \neg M2 | C)P(C) / P(M1, \neg M2)
by Bayes's rule
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 $=P(M1\mid C)P(\neg M2\mid C)P(C)/P(M1,\neg M2)$ 

 $= (.85)(1-.85)(.01)/P(M1, \neg M2)$ 

•  $P(M1, \neg M2) = P(M1, \neg M2 \mid C)P(C) +$   $P(M1, \neg M2 \mid \neg C)P(\neg C)$  by Conditioning rule  $= P(M1 \mid \neg M2, C)P(\neg M2 \mid C)P(C) +$  $P(M1 \mid \neg M2, \neg C)P(\neg M2 \mid \neg C)P(\neg C)$ 

by Conditionalized Chain rule

Bayes's Rule with Multiple Evidence

- Say the same patient goes back and gets a second mammography and it is negative.
   Now, what is the chance she has cancer?
- Let M1 be the positive test and ¬M2 be the negative test
- Compute posterior:  $P(C|M1, \neg M2)$

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Cancer "causes" a positive test, so M1 and  $\neg M2$  are conditionally independent given C, so

 $P(M1|\neg M2, C)P(\neg M2|C)P(C) + P(M1|\neg M2, \neg C)P(\neg M2|\neg C)P(\neg C)$ 

 $= P(M1 \mid C)P(\neg M2 \mid C)P(C) + P(M1 \mid \neg C)P(\neg M2 \mid \neg C)P(\neg C) \text{ by cond. indep.}$ 

= (.85)(1 - .85)(.01) + (1 - .08)(.08)(1 - .01)

 $= 0.074139 \quad (= P(M1, \neg M2))$ 

So,  $P(C|M1, \neg M2) = (.85)(1 - .85)(.01)/.074139$ = 0.017 or 1.7%

# Bayes's Rule with Multiple Evidence and Conditional Independence

- Assume all evidence variables, B, C and D, are conditionally independent given the diagnosis variable, A
- P(A|B,C,D) = P(B,C,D|A)P(A)/P(B,C,D)= P(B|A)P(C|A)P(D|A)P(A)/P(D|B,C)P(C|B)P(B)

Conditionalized Chain rule + conditional independence

Chain rule

$$= P(A) \frac{P(B|A)}{P(B)} \frac{P(C|A)}{P(C|B)} \frac{P(D|A)}{P(D|B,C)}$$

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# Naïve Bayes Classifier

- Classification problem: Find the value of class/decision/diagnosis variable Y that is most likely given evidence/measurements/attributes X<sub>i</sub> = v<sub>i</sub>
- Use Bayes's rule and conditional independence:

$$P(Y = c \mid X_1 = v_1, X_2 = v_2, ..., X_n = v_n)$$

$$= P(Y = c)P(X_1 = v_1 \mid Y = c)...P(X_n = v_n \mid Y = c) / P(X_1 = v_1, ..., X_n = v_n)$$

- Try all possible values of Y and pick the value that gives the maximum probability
- But denominator,  $P(X_1=v_1, ..., X_n=v_n)$ , is a constant for all values of Y, so it won't affect which value of Y is best

# Inference Ignorance

- "Inferences about Testosterone Abuse Among Athletes," 2004
  - Mary Decker Slaney doping case
- "Justice Flunks Math," 2013
  - Amanda Knox trial in Italy

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# Naive Bayes Classifier Testing Phase

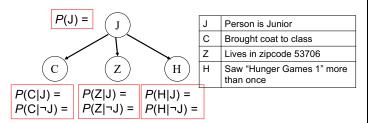
For a given test instance defined by X<sub>1</sub>=v<sub>1</sub>, ..., X<sub>n</sub>=v<sub>n</sub>, compute

$$argmax_c P(Y=c) \prod_{i=1}^n P(X_i=v_i \mid Y=c)$$
Class variable Evidence variable

- Assumes all evidence variables are conditionally independent of each other given the class variable
- Robust because it gives the right answer as long as the correct class is more likely than all others

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# **Naïve Bayes Classifier Training Phase**

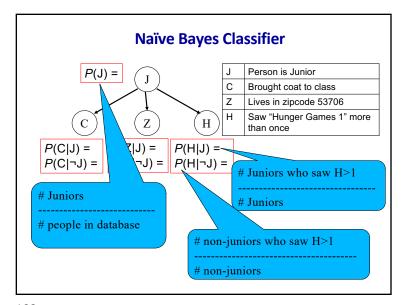


Compute from the Training set all the necessary Prior and Conditional probabilities

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# Naïve Bayes Classifier

- Assume k classes and n evidence (i.e., attribute) variables, each with m possible values
- k-1 values needed for computing P(Y=c)
- (m-1)k values needed for computing  $P(X_i=v_i \mid Y=c)$  for each evidence variable  $X_i$
- So, (k-1) + n(m-1)k values needed instead of exponential size FJPD table



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# Naïve Bayes Classifier

 Conditional probabilities can be very, very small, so instead use logarithms to avoid underflow:

$$\arg\max_{c} \log P(Y = c) + \sum_{i=1}^{n} \log P(X_{i} = v_{i} | Y = c)$$

# Add-1 Smoothing

- Unseen event problem: Training data may not include some cases
  - flip a coin 3 times, all heads → one-sided coin?
  - Conditional probability = 0
  - Just because a value doesn't occur in the training set doesn't mean it will never occur
- "Add-1 Smoothing" ensures that every conditional probability > 0 by pretending that you've seen each attribute's value 1 extra time

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# Add-1 Smoothing

• Compute Prior probabilities as

$$P(Y = c) = \frac{count(Y = c) + 1}{N + k}$$

where N = size of the training set, k = number of classes

# Add-1 Smoothing

• Compute Conditional probabilities as

number of times attribute X has value  $v_i$  in all training instances with class  $c_i$ 

$$P(X = vi \mid Y = c) = \frac{count(X = vi, Y = c) + 1}{count(Y = c) + m}$$

number of training instances with class  $\alpha$ 

• Note:  $\sum_{i=1}^{m} P(X = vi | Y = c) = 1$ 

where m = number of possible values for attribute X

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# **Laplace Smoothing**

- aka Add-δ Smoothing
- Instead of adding 1, add  $\delta$  (a positive real number)
- Compute conditional probabilities as

$$P(X=v_i \mid Y=c) = \frac{count(X=vi,Y=c) + \delta}{count(Y=c) + \delta m}$$

where m = number of possible values for attribute X

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