## STAT3006 Assignment 2

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#### 1 Question 1

For n = 500 i.i.d. r.v.s which follow Poisson( $\lambda$ ), unobserved variables are  $\lambda, y_1, y_2, ..., y_{68}$ , in which y's denote the r.v.s which are larger than or equal to 4.

Prior: 
$$\pi(\lambda) \propto \frac{1}{\lambda}$$

$$P(X,Y|\lambda) = \prod_{i=1}^{432} \frac{e^{-\lambda} \lambda^{x_i}}{x_i!} \prod_{j=1}^{68} \frac{e^{-\lambda} \lambda^{y_j}}{y_j!} I(y_j \ge 4)$$

$$P(\lambda|X,Y) \propto P(X,Y|\lambda) \pi(\lambda) \propto e^{-n\lambda} \lambda^{\sum_i x_i + \sum_j y_j - 1}$$

$$P(y_j|X,\lambda) = \frac{e^{-\lambda} \lambda^{y_j}}{y_j!} I(y_j \ge 4) \propto \frac{\lambda^{y_j}}{y_j!} I(y_j \ge 4)$$

$$\lambda|X,Y \propto Gamma(\sum_i x_i + \sum_j y_j, n)$$
The MH-Step to sample 68 unobserved  $y'$  is  $y_j^* = \begin{cases} y_j^{(t)} - 1, & \text{with probability } \frac{1}{3} \\ y_j^{(t)}, & \text{with probability } \frac{1}{3} \end{cases}$ 

$$r = \min \left\{ \frac{\left[\lambda^{(t+1)}\right]^{y_j^*} / y_j^*!}{\left[\lambda^{(t+1)}\right]^{y_j^{(t)}} / y_j^{(t)}!} I(y_j \ge 4), 1 \right\}, \text{ which is the accept-reject ratio.}$$

The estimated  $\lambda$  is 1.45744.

#### 2 Question 2

The complete-data likelihood function is:

$$f(X, Z|\Pi, \Theta) = L(\Pi, \Theta|X, Z) = \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[ P(Z_i = k|\Pi, \Theta) P(X_{ij}, j = 1, 2|Z_j, \Pi, \Theta) \right]^{I(Z_j = k)}$$

$$= \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[ \pi_k \prod_{j=1}^{2} P(X_{ij}|Z_j, \Theta) \right]^{I(Z_j = k)}$$

$$= \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[ \pi_k \prod_{j=1}^{2} \binom{10 \times j}{X_{ij}} \theta_{jk}^{X_{ij}} (1 - \theta_{jk})^{10 \times j - X_{ij}} \right]^{I(Z_j = k)}$$

Given  $Z_i = k$ , we have each sample i,  $X_{ij} \sim Bino(10 \times j, \theta_{jk})$ ,  $P(\theta_{jk}) \propto Beta(a, b)$  with prior  $P(\theta_{jk}) \propto Beta(1, 1) \propto 1$  and  $(\pi_1, \pi_2, \pi_3) \sim Dirichlet(\alpha_1, \alpha_2, \alpha_3)$ .

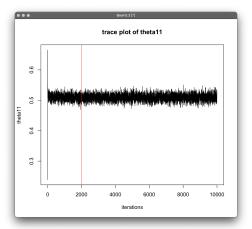
$$\begin{split} P(\Pi,\Theta|X,Z) &\propto P(\Pi)P(\Theta)P(X,Z|\Pi,\Theta) \\ &\propto \prod_{k=1}^{3} \pi_{k}^{\alpha_{k}-1} \prod_{j=1}^{2} \prod_{k=1}^{3} \theta_{jk}^{\alpha_{j}-1} (1-\theta_{jk})^{b-1} \prod_{i=1}^{1000} \prod_{k=1}^{3} \left[\pi_{k} \prod_{j=1}^{2} \binom{10 \times j}{X_{ij}} \theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10 \times j-X_{ij}}\right]^{I(Z_{j}=k)} \\ &\approx \prod_{k=1}^{3} \pi_{k}^{\alpha_{k}-1} \prod_{k=1}^{3} \prod_{k=1}^{3} \pi_{k}^{I(Z_{i}=k)} \\ &\propto \prod_{k=1}^{3} \pi_{k}^{\alpha_{k}-1} \prod_{k=1}^{3} \pi_{k}^{\sum_{i=1}^{1000} I(Z_{i}=k)} \\ &\propto \prod_{k=1}^{3} \pi_{k}^{\alpha_{k}-1} \prod_{k=1}^{2} \prod_{i=1}^{1000} I(Z_{i}=k) \\ &\propto \prod_{k=1}^{3} \pi_{k}^{\alpha_{k}-1} \prod_{k=1}^{2} \prod_{i=1}^{1000} I(Z_{i}=k) \\ &\propto \prod_{k=1}^{3} \pi_{k}^{\alpha_{k}-1} \prod_{k=1}^{2} \prod_{i=1}^{1000} I(Z_{i}=1), \alpha_{2} + \sum_{i=1}^{1000} I(Z_{i}=2), \alpha_{3} + \sum_{i=1}^{1000} I(Z_{i}=3)) \\ &\text{For } \theta \colon \theta_{jk} |\sim \alpha \prod_{j=1}^{2} \prod_{k=1}^{3} \theta_{jk}^{\alpha_{j}-1} (1-\theta_{jk})^{b-1} \prod_{i=1}^{1000} \prod_{k=1}^{3} \prod_{j=1}^{2} \binom{10 \times j}{X_{ij}} \theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10 \times j-X_{ij}} \prod_{i=1}^{I(Z_{j}=k)} \\ &\propto \theta_{jk}^{a-1} (1-\theta_{jk})^{b-1} \prod_{i=1}^{1000} \left[\theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10 \times j-X_{ij}} \right]^{I(Z_{j}=k)} \\ &\propto \theta_{jk}^{a-1} (1-\theta_{jk})^{b-1} \prod_{i=1}^{1000} \left[\theta_{jk}^{X_{ij}} (1-\theta_{jk})^{10 \times j-X_{ij}} \right]^{I(Z_{j}=k)} \\ &\propto \theta_{jk}^{a-1} \prod_{i=1}^{3} \prod_{k=1}^{2} \prod_{j=1}^{1000} \prod_{i=1}^{3} \prod_{i=1}^{2} \prod_{k=1}^{1000} \prod_{j=1}^{3} \prod_{i=1}^{2} \prod_{k=1}^{1000} \prod_{j=1}^{3} \prod_{i=1}^{2} \prod_{i=1}^{1000} \prod_{i=1}^{3} \prod_{i=1}^{2} \prod_{i=1}^{1000} \prod_{i=1}^{1000} \prod_{i=1}^{2} \prod$$

For Gibbs Sampler algorithm, given  $\Pi^{(t)}, \Theta^{(t)}, Z^{(t)}$ , the updates are as follows.

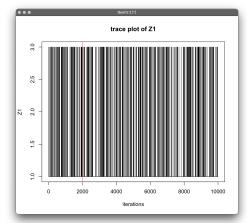
$$\begin{split} \Pi|\Theta, Z \sim & Dirichlet(\alpha_1 + \sum_{i=1}^{1000} I(Z_i^{(t)} = 1), \alpha_2 + \sum_{i=1}^{1000} I(Z_i^{(t)} = 2), \alpha_3 + \sum_{i=1}^{1000} I(Z_i^{(t)} = 3)) \\ \theta_{jk}|-\sim & Beta\left(a + \sum_{i=1}^{1000} X_{ij}I(Z_i^{(t)} = k), b + \sum_{i=1}^{1000} (10 \times j - X_{ij})I(Z_i^{(t)} = k)\right) \\ P(Z_i^{(t+1)} = k|-) = & \frac{\pi_k^{(t+1)} \prod_{j=1}^2 \binom{10 \times j}{X_{ij}} \theta_{jk}^{X_{ij}} (1 - \theta_{jk})^{10 \times j - X_{ij}}}{\sum_{l=1}^3 \pi_l^{(t+1)} \prod_{j=1}^2 \binom{10 \times j}{X_{ij}} \theta_{jl}^{X_{ik}} (1 - \theta_{jl})^{10 \times j - X_{ij}}} \end{split}$$

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The estimated \Theta = (\hat{\theta}_{11}, \hat{\theta}_{12}, \hat{\theta}_{13}, \hat{\theta}_{21}, \hat{\theta}_{22}, \hat{\theta}_{23}) = (0.5093235, 0.1909221, 0.8023880, 0.4920218, 0.7818400, 0.1984175).
The estimated Z are as follows.
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The estimated  $\Pi = (\hat{\pi}_1, \hat{\pi}_2, \hat{\pi}_3) = (0.3047501, 0.1008389, 0.5944110).$ 



[589] 985 988 989 990 992 994 995 996 997 998 999



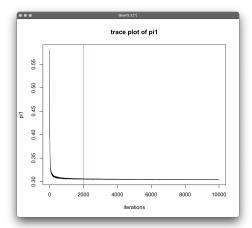


Figure 1: Trace Plot of  $\theta_{11}$  (left),  $Z_1$  (Middle) and  $\pi_1$  (Right)

#### 3 Question 3

For  $i = 1, 2, Y_i \sim multinomial(100, p_1, p_2, p_3, p_4)$ .

Prior: 
$$\pi(\mathbf{p}) \propto Dirchlet(\alpha_{1}, \alpha_{2}, \alpha_{3}, \alpha_{4}) \propto Dirchlet(2, 2, 2, 2) \propto p_{1}p_{2}p_{3}p_{4}$$

$$P(y_{i1}, y_{i2}, y_{i3}, y_{i4}|p_{1}, p_{2}, p_{3}, p_{4}) = \frac{100!}{y_{i1}!y_{i2}!y_{i3}!y_{i4}!} p_{1}^{y_{1}}p_{2}^{y_{2}}p_{3}^{y_{3}}y_{4}^{y_{4}}$$

$$P(p_{1}, p_{2}, p_{3}, p_{4}|y_{i1}, y_{i2}, y_{i3}, y_{i4}) \propto P(y_{i1}, y_{i2}, y_{i3}, y_{i4}|p_{1}, p_{2}, p_{3}, p_{4}) \cdot f(p_{1}, p_{2}, p_{3}, p_{4}|y_{i1}, y_{i2}, y_{i3}, y_{i4})$$

$$\propto \frac{100!}{y_{i1}!y_{i2}!y_{i3}!y_{i4}!} p_{1}^{y_{i1}+\alpha_{1}-1} p_{2}^{y_{i2}+\alpha_{2}-1} p_{3}^{y_{i3}+\alpha_{3}-1} p_{4}^{y_{i4}+\alpha_{4}-1}$$

$$p_{1}, p_{2}, p_{3}, p_{4}|y_{i1}, y_{i2}, y_{i3}, y_{i4} \sim Dirchlet(y_{i1} + \alpha_{1}, y_{i2} + \alpha_{2}, y_{i3} + \alpha_{3}, y_{i4} + \alpha_{4})$$

$$P(y_{12}|\mathbf{Y}, \mathbf{P}) \propto \frac{p_{2}^{y_{12}}}{y_{12}!} \cdot \frac{p_{3}^{y_{13}}}{y_{13}!} = \frac{p_{2}^{y_{12}}}{y_{12}!} \cdot \frac{p_{3}^{44-y_{12}}}{(44-y_{12})!}$$

$$P(y_{22}|\mathbf{Y}, \mathbf{P}) \propto \frac{p_{2}^{y_{22}}}{y_{22}!} \cdot \frac{p_{4}^{y_{24}}}{y_{24}!} = \frac{p_{2}^{y_{22}}}{y_{22}!} \cdot \frac{p_{4}^{48-y_{22}}}{(48-y_{22})!}$$

Denote the upper bound of  $y_{i2}$  by c,  $c = \begin{cases} 35, \text{ for } i = 1\\ 30, \text{ for } i = 2 \end{cases}$ 

Then MH-Step to update  $y_{i2}$  for i = 1, 2 is as follows.

$$\begin{split} y_{i2}^{(t+1)} &= \begin{cases} y_{i2}^{(t)} + 1 \text{ with probability } 0.5 \text{ or if } y_{i2}^{(t)} = 15 \\ y_{i2}^{(t)} - 1 \text{ with probability } 0.5 \text{ or if } y_{i2}^{(t)} = c \end{cases} \\ r &= \begin{cases} \min \{ 2 \times \frac{P(y_{i2}^{(t+1)}|Y,P)}{P(y_{i2}^{(t)}|Y,P)}, 1 \} & \text{,if } y_{i2}^{(t+1)} = c \text{ or } 15 \\ \min \{ \frac{1}{2} \times \frac{P(y_{i2}^{(t+1)}|Y,P)}{P(y_{i2}^{(t)}|Y,P)}, 1 \} & \text{,if } y_{i2}^{(t)} = c \text{ or } 15 \\ \min \{ \frac{P(y_{i2}^{(t+1)}|Y,P)}{P(y_{i2}^{(t)}|Y,P)}, 1 \} & \text{, otherwise} \end{cases} \end{split}$$

where r is the accept-reject ratio. For  $y_{13}$  and  $y_{24}$ ,

$$y_{13}^{(t+1)} = 100 - y_{11} - y_{14} - y_{12}^{(t+1)} = 44 - y_{12}^{(t+1)}$$
$$y_{24}^{(t+1)} = 100 - y_{21} - y_{23} - y_{22}^{(t+1)} = 48 - y_{22}^{(t+1)}$$

The estimated  $\mathbf{P} = (\hat{p}_1, \hat{p}_2, \hat{p}_3, \hat{p}_4) = (0.4561099, 0.2653649, 0.1546046, 0.1239205)$