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- **Refresher** on logs, odds, probability and linear regression
- Understand why linear regression not sensible for **binary data**
- Explain how **logit** and binomial model let us **extend linear regression**
- Be able to run a **simple logistic regression in R**
- Be able to explain basic R glm **output**
- Be able to explain **estimates** with categorical and continuous variables
- Explain **significance test results** on variables
- Things to watch out for!
- **Know where to go next!**

```
library(tidyverse)
library(boot)
library(broom)
library(skimr)
library(sjPlot)

dat ← read_csv("logreg_data_01_20190530.csv")

skim(dat)
```

Can we skip this bit?

$$\log_{10}(10) = 1$$

$$\log_{10}(1000) = 3$$

$$\log_{10}(0.01) = -2$$

We can have other bases e.g. e

$$\log_e(2.718) \simeq 1$$

And reversing this...

$$10^3 = 1000$$

$$e^2 \simeq 7.389$$

Probabilities have values from 0 ('never happens') to 1 ('always happens')

'events of interest' ÷ 'all events'

What is the probability that a fair coin lands on heads?

$$1/2 = 0.5$$

What is the probability that a 6 sided die 🎲 lands on 4?

$$1/6 \simeq 0.166$$

Odds have values from 0 ('never happen') to infinity ('always happens')

'events of interest' ÷ 'other events'

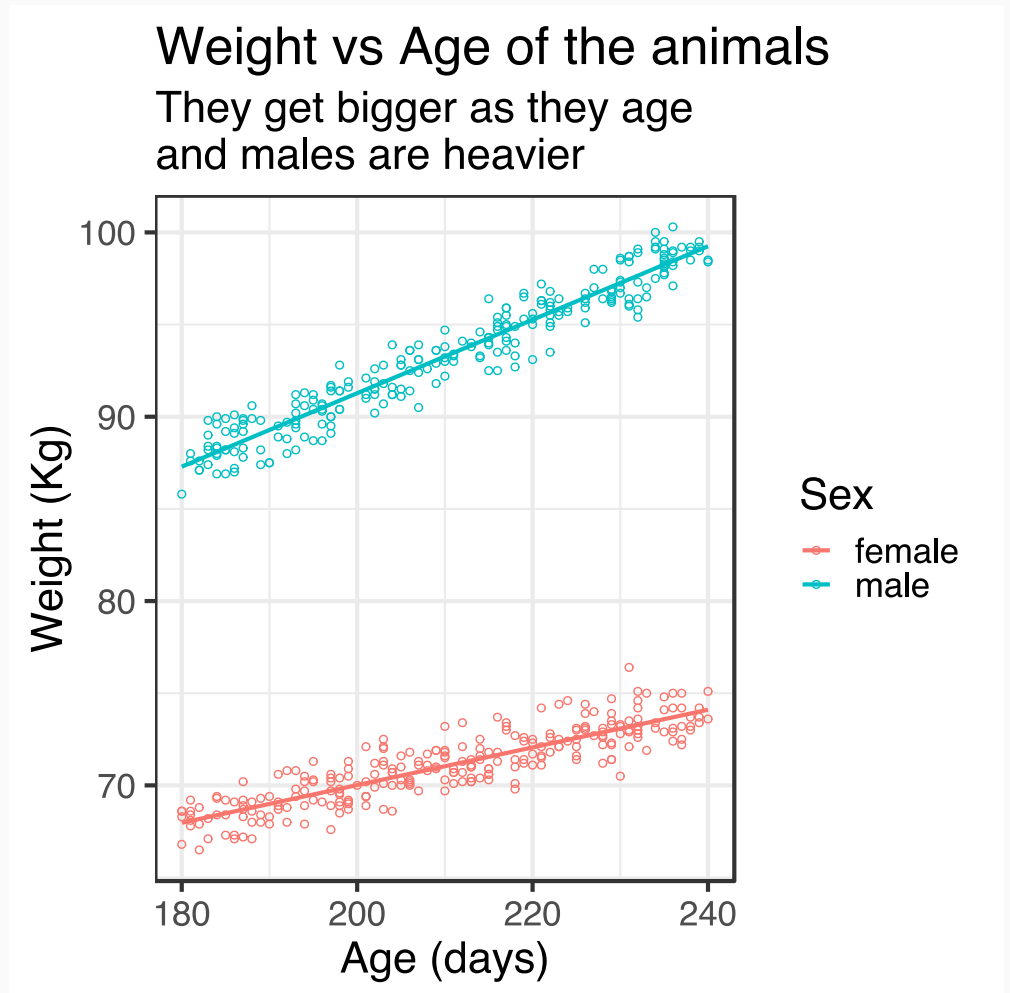
What is the odds that a fair coin lands on heads?

$$1/1 = 1$$

What is the odds that a 6 sided die 🎲 lands on 4?

$$1/5 = 0.2$$

- numerical outcome
- numerical / categorical predictors
- linear relationship



```
mod ← lm(weight ~ age + sex, data = dat)
```

Call:

```
lm(formula = weight ~ age + sex, data = dat)
```

Residuals:

Min	1Q	Median	3Q	Max
-3.54	-0.88	-0.02	0.89	3.07

Coefficients:

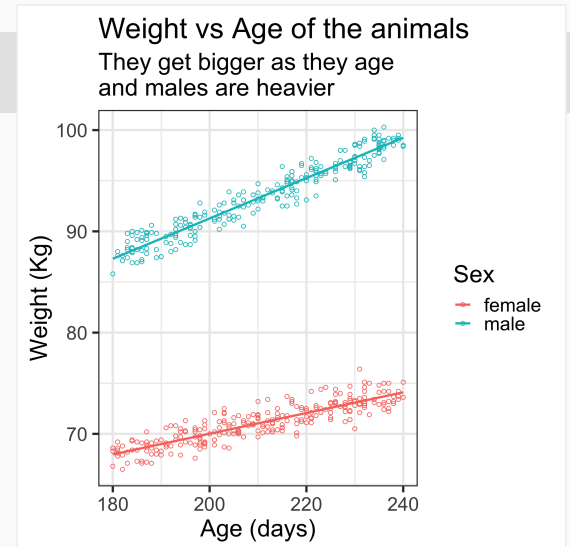
	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	39.1901	0.7026	56	<2e-16	***
age	0.1515	0.0033	46	<2e-16	***
sexmale	22.2796	0.1136	196	<2e-16	***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.3 on 497 degrees of freedom

Multiple R-squared: 0.99, Adjusted R-squared: 0.99

F-statistic: 2e+04 on 2 and 497 DF, p-value: <2e-16



- alive/dead
- healthy/diseased

Example data

ID	treatment	age	region	sex	weight	status
A0458	control	209	C	female	70.9	healthy
A0468	treated	190	C	female	68.3	healthy
A0143	control	239	B	female	73.7	diseased
A0413	control	235	D	male	97.8	healthy
A0319	control	197	B	male	89.1	healthy
A0257	control	194	B	female	69.7	healthy

Status vs treatment

treatment	diseased	healthy
control	43	168
treated	34	255

```
with(dat,  
      fisher.test(status, treatment))
```

Fisher's Exact Test for Count Data

data: status and treatment

p-value = 0.01175

alternative hypothesis: true odds ratio is not equal to 1

95 percent confidence interval:

1.142387 3.238178

sample estimates:

odds ratio

1.917107

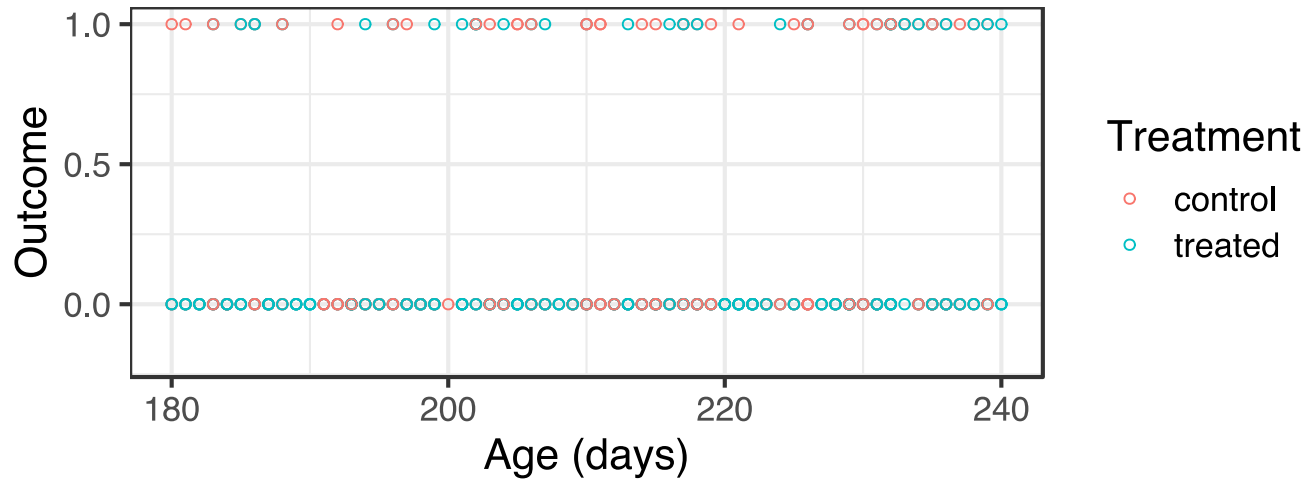
How about recoding the outcome as 0/1?

Example data

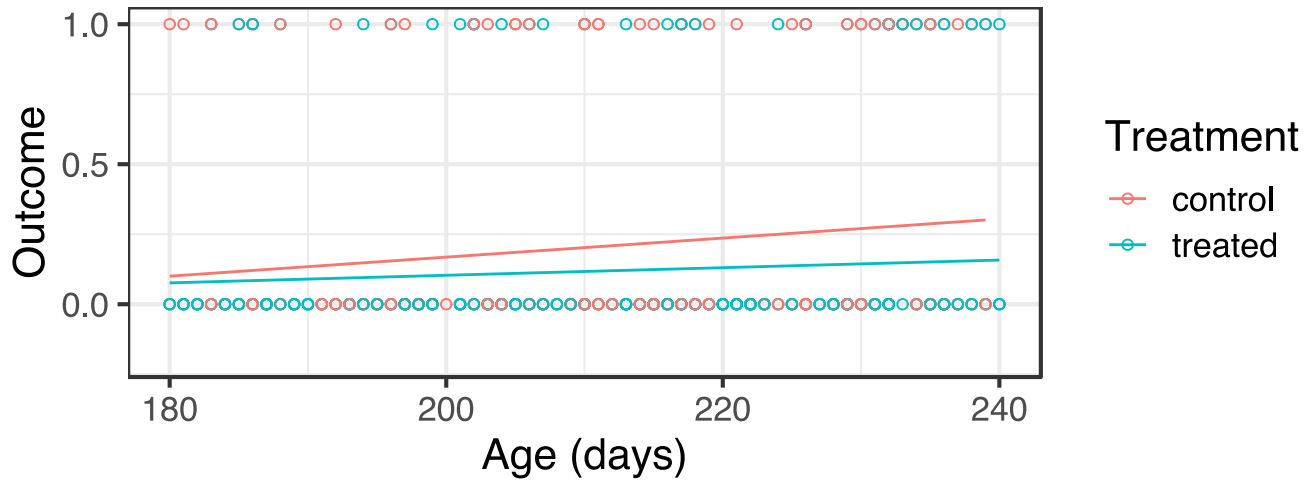
ID	treatment	age	region	sex	weight	status	status01
A0458	control	209	C	female	70.9	healthy	0
A0468	treated	190	C	female	68.3	healthy	0
A0143	control	239	B	female	73.7	diseased	1
A0413	control	235	D	male	97.8	healthy	0
A0319	control	197	B	male	89.1	healthy	0
A0257	control	194	B	female	69.7	healthy	0

Then use linear regression...

Status vs Age and treatment



Status vs Age and treatment



-predicts (impossible) intermediate values

-can predict <0 and >1

Linear regression does this...

$$weight \sim \beta_0 + \beta_1 age + \beta_2 sex + \epsilon$$

or in english...

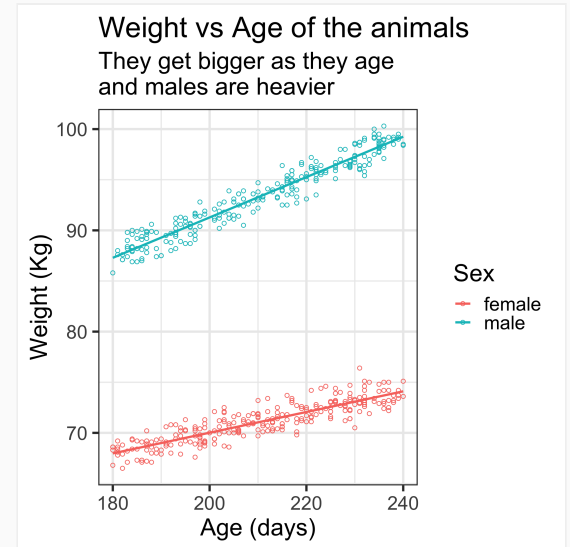
The outcome, *weight*, is related to the predictors by one or more straight lines.

For binary data we want

Our outcome to be 0 or 1

So rather than modelling the outcome.

We model the **probability** of something e.g. being diseased...

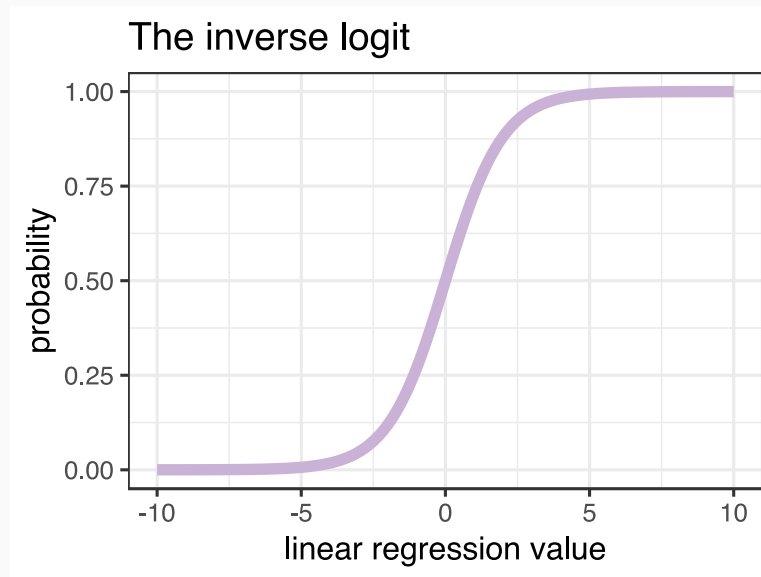


Linear regression models model numbers, any numbers!

Probabilities go from...

0 to 1

So we need to turn any number into 0 - 1



In fact the regression value is the log of the odds of the outcome.

So we have an outcome, e.g. being diseased vs healthy, that is coded 0 or 1

And our model is

$$\log_e\left(\frac{prob}{1 - prob}\right) \sim \beta_0 + \beta_1 age + \beta_2 treatment$$

or in english

The log of the odds of an animal being diseased are modelled by a linear combination of the predictor variables


```
head(dat)
```

ID	treatment	age	region	sex	weight	status	status01
A0001	control	219	A	female	71.4	diseased	1
A0002	control	218	A	female	70.1	healthy	0
A0003	treated	214	D	female	71.4	healthy	0
A0004	treated	194	D	female	68.9	healthy	0
A0005	control	185	D	female	67.3	healthy	0
A0006	treated	235	D	male	98.6	healthy	0

A linear model of weight

```
mod_weight <- lm(weight ~ age + sex, data = dat)
```

A logistic regression model of disease status

```
mod_disease <- glm(status01 ~ treatment + age, family = binomial, data = dat)
```

```
summary(mod_disease)
```

Call:

```
glm(formula = status01 ~ treatment + age, family = binomial,  
     data = dat)
```

Deviance Residuals:

	Min	1Q	Median	3Q	Max
	-0.8300	-0.6054	-0.5320	-0.4210	2.2841

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	-5.074379	1.613819	-3.144	0.00166	**
treatmenttreated	-0.662589	0.251643	-2.633	0.00846	**
age	0.017514	0.007521	2.329	0.01987	*

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 429.59 on 499 degrees of freedom
Residual deviance: 417.16 on 497 degrees of freedom
AIC: 423.16

Number of Fisher Scoring iterations: 4

```
print(summary(mod_disease), digits = 3)
```

```
Call:
glm(formula = status01 ~ treatment + age, family = binomial,
    data = dat)
```

```
Deviance Residuals:
```

Min	1Q	Median	3Q	Max
-0.830	-0.605	-0.532	-0.421	2.284

```
Coefficients:
```

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-5.07438	1.61382	-3.14	0.0017 **
treatmenttreated	-0.66259	0.25164	-2.63	0.0085 **
age	0.01751	0.00752	2.33	0.0199 *

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
(Dispersion parameter for binomial family taken to be 1)
```

```
Null deviance: 429.59  on 499  degrees of freedom
Residual deviance: 417.16  on 497  degrees of freedom
AIC: 423.2
```

```
Number of Fisher Scoring iterations: 4
```

Lets get 'tidy output...

```
tidy(mod_disease) #tidy from the broom package
```

```
# A tibble: 3 x 5
```

	term	estimate	std.error	statistic	p.value
	<chr>	<dbl>	<dbl>	<dbl>	<dbl>
1	(Intercept)	-5.07	1.61	-3.14	0.00166
2	treatmenttreated	-0.663	0.252	-2.63	0.00846
3	age	0.0175	0.00752	2.33	0.0199

The estimates = log(odds ratios)

i.e.

$$\frac{\text{odds of outcome if have factor}}{\text{odds of outcome if dont have factor}}$$

So we get odds ratios by 'inverse logging them'.

We can remove the intercept.

```
tidy(mod_disease) %>%  
  filter(term  $\neq$  "(Intercept)") %>%  
  mutate(OR = exp(estimate))
```

A tibble: 2 x 6

	term	estimate	std.error	statistic	p.value	OR
	<chr>	<dbl>	<dbl>	<dbl>	<dbl>	<dbl>
1	treatmenttreated	-0.663	0.252	-2.63	0.00846	0.516
2	age	0.0175	0.00752	2.33	0.0199	1.02

```

tidy(mod_disease) %>%
  mutate(OR = exp(estimate)) %>%
  bind_cols(exp(confint(mod_disease)) %>%
    as_tibble()
  ) %>%
  filter(term  $\neq$  "(Intercept)") %>%
  select(term, OR, `2.5 %`, `97.5 %`, p.value)

```

term	OR	2.5 %	97.5 %	p.value
treatmenttreated	0.516	0.313	0.842	0.008
age	1.018	1.003	1.033	0.020

But what does it mean?

term	OR	2.5 %	97.5 %	p.value
treatmenttreated	0.516	0.313	0.842	0.008
age	1.018	1.003	1.033	0.020

multiply

How many times greater the odds of outcome are **if** the risk factor (etc) is present.

So for the treatment variable (which can be control or treatment) the odds of disease if treated are 0.516 **times greater** than if untreated (control).

term	OR	2.5 %	97.5 %	p.value
treatmenttreated	0.516	0.313	0.842	0.008
age	1.018	1.003	1.033	0.020

multiply

How many times greater the odds of outcome are for **each unit change** in the variable

So for the age variable the odds of disease are 1.018 **times greater** for each day older.

So for 3 days it's $1.018 \times 1.018 \times 1.018 \simeq 1.055$.

How does R know if you are predicting 'healthy' or 'diseased'?

E.g. all the males are diseased and all the females are healthy

Disease risk might go up and then down

Dohoo book