

Laboratory Measurement Bias and Repeatability Model (LBM)

This section has four parts: (1) a description of the Fifth International Radiocarbon Intercomparison (VIRI) dataset to which the LBM was fitted, (2) model formulas and prior distributions for parameters, (3) a description of the posterior parameter values, and (4) a description of Hamiltonian Monte Carlo (HMC) diagnostics with a posterior predictive check. The goal of this model is to estimate parameters that describe inter-lab variation in the measurement of the ^{14}C content of a sample as well as intra-laboratory variation over repeated measurements of a sample. This involves estimating the distributions of mean ^{14}C values and ^{14}C standard deviations across laboratories.

1.1. The VIRI Dataset

To fit the model, we first aggregated data presented in the VIRI (1–3). These data included all ^{14}C measurements across all sample materials for which ^{14}C measurements were reported (samples B, D, F, G, H, and I). In total, this spans 420 measurements performed by 80 laboratories.

We transformed these data in three steps. First, we centered all ^{14}C measurements on the median value for each sample material. Second, we used these centered measurements to investigate the presence of outliers. Outliers were defined very conservatively as measurements that fall at least six times the distance of the interquartile range (IQR) either below the first or above the third quartile within each set of ^{14}C measurements for a sample material. The values identified as outliers may have resulted from unusually poor quality control for some laboratories or from other anomalies in measurement. After outliers were identified, all measurements associated with the laboratory that produced an outlier were removed from the dataset (Figure S3.1). This reduced the number of laboratories from 80 to 68, and the number of ^{14}C measurements from 420 to 361. Finally, these median-centered measurements were converted to z-scores (Figure S3.1).

1.2. Model Formula and Prior Distributions for Parameters

Each of the 361 median-centered ^{14}C z-scores is associated with four additional variables: a categorical sample material ID (B, D, F, G, H, or I), a categorical laboratory ID, a dummy variable indicating whether the laboratory performed an AMS (0) or GPC/LSC measurement (1), and the reported measurement errors for ^{14}C values. Although measurement errors should represent uncertainty in the reported means, laboratories calculate these errors in a variety of ways that may not be comparable (4). As such, we treat them as predictor variables for the dispersion of reported ^{14}C means, with the expectation that within-laboratory repeatability ^{14}C measurements decreases with larger reported errors.

First, we defined the likelihood for median-centered ^{14}C z-scores as

$$^{14}\text{C} \sim N(\mu, \sigma). \quad (\text{Equation S3.1})$$

Median-centered ^{14}C z-scores are distributed $N(\mu, \sigma)$. We then modelled μ as a linear outcome of the ^{14}C value of sample material i , $C_s[\text{sample } i]$, and an offset from that sample material value that depends on the laboratory ID, $C_o[\text{lab } j]$. The laboratory ID offset also varies based on the AMS vs GPC/LSC dummy variable, AMS , through parameter C_{AMS} :

$$\mu = C_s[\text{sample } i] + C_o[\text{lab } j] \times (1 + (C_{\text{AMS}} - 1) \times \text{AMS}). \quad (\text{Equation S3.2})$$

Since median-centered ^{14}C z-scores likely approximate a true sample-specific value near zero, we use a prior distribution of $N(0, 1)$ for each $C_s[\text{sample } i]$. We modelled $C_o[\text{lab } j]$ as distributed $N(0, C_o^\sigma)$, with the prior for C_o^σ set to $\exp(2)$. C_{AMS} may reduce or increase the effect of $C_o[\text{lab } j]$, but it does not affect the sign of $C_o[\text{lab } j]$, taking only positive values. Values less than 1 reduce the lab specific offset $C_o[\text{lab } j]$, while values greater than 1 increase the lab specific offset $C_o[\text{lab } j]$. As such, for C_{AMS} we use a gamma distribution with the mean centered on 1 as a prior: $\text{gamma}(1, 0.5)$. The parameterization of this prior does not follow the base *dgamma* R function, but instead uses the *dgamma2* parameterization included in the *rethinking* R package (5).

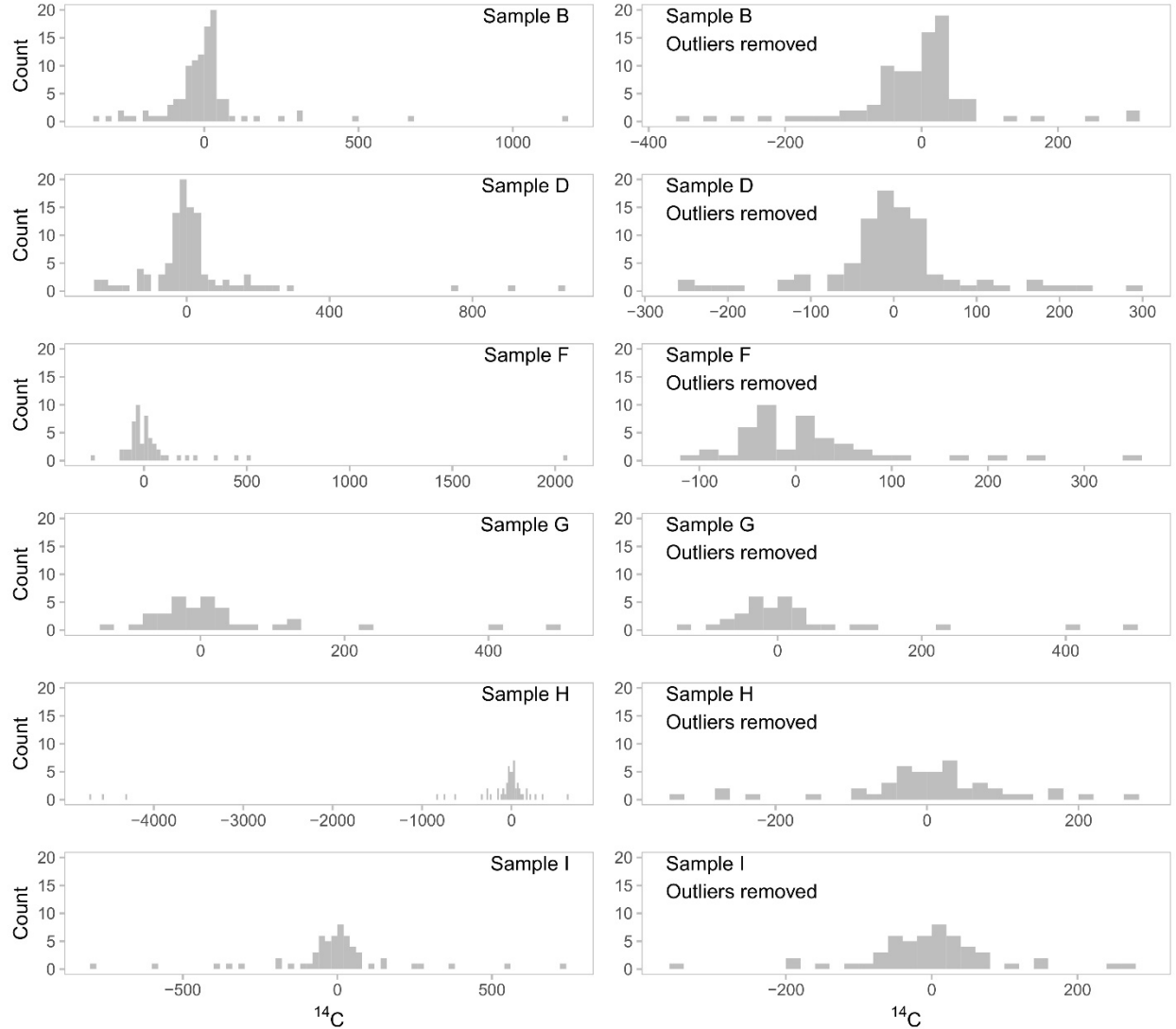


Figure S3.1. Histograms of median-centered ^{14}C z-scores for samples B, D, F, G, H, and I. Left panels show all available measurements ($n = 420$), and right panels show only those measurements that originate from laboratories that did not produce outlier values ($n = 361$). The values in the right panels were used to fit the LBM.

σ can then be interpreted as intra-laboratory variation in median-centered ^{14}C z-scores. We modelled σ as the linear outcome of a baseline parameter shared by all laboratories, σ_l , a laboratory ID specific offset parameter, $\sigma_{lab}[\text{lab } j]$, the log reported measurement error ($\log(\text{ME})$)

multiplied by parameter σ_{ME} , and a parameter that is expressed only for GPC/LSC laboratories (σ_{AMS}). As such, the dispersion of repeated intra-laboratory ^{14}C z-scores depends on laboratory specific variation in repeatability, the reported measurement error associated with those measurements, and whether the measurement was obtained via AMS or GPC/LSC. Laboratory specific variation depends on each reported ^{14}C measurement error, which can vary from value to value within a single laboratory. To constrain σ on the positive scale, we used a log link function:

$$\log(\sigma) = \sigma_{lab}[lab\ j] + \sigma_l + \sigma_{ME} \times \log(ME) + \sigma_{AMS} \times AMS. \quad (\text{Equation S3.3})$$

We assigned the same informative normal prior distribution to σ_l , σ_i , and σ_{AMS} : $N(0, 1)$. We modelled $\sigma_{lab}[lab\ j]$ as distributed $N(0, \sigma_{lab}^\sigma)$, with the prior for σ_{lab}^σ set to $\exp(2)$. Readers may note that we have not modelled covariance between the random laboratory parameters, $C_o[lab\ j]$ and $\sigma_{lab}[lab\ j]$. For the simulation, the practical implication of this decision is that these parameter values are sampled independently for simulated laboratories rather than from a multivariate distribution. Such covariance is often modelled as multivariate normal, which would be inappropriate here, as σ_{lab} should vary with *only the magnitude* of C_o rather than the *magnitude and sign* of C_o (i.e., a parabolic rather than monotonic relationship). In other words, we might expect within-laboratory dispersion to vary with absolute laboratory ID offset, regardless of whether the laboratory ID offset is above or below the target ^{14}C value (Figure S3.2). For the sake of model simplicity and interpretability, we did not attempt to model a parabolic relationship between these parameters.

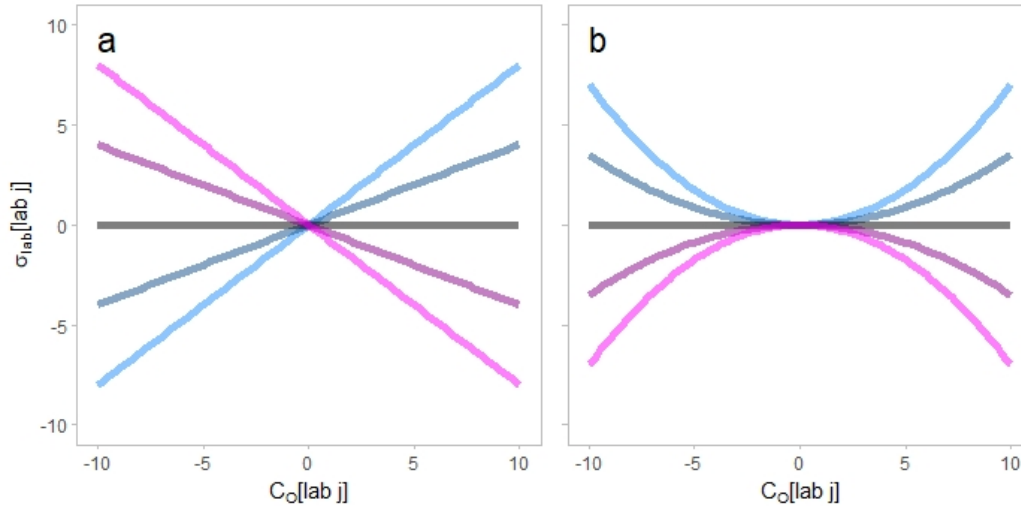


Figure S3.2. (a) Example relationships that might be expected if laboratory parameters covaried monotonically. (b) Example relationships that might be expected given a parabolic relationship between parameters. Given the lack of a covariance component in the model, laboratory parameters were estimated independently, conforming to the horizontal grey relationship in each panel (i.e., no covariance).

This model assumes that systematic offsets between mean laboratory measurements and target ^{14}C values are maintained across any sample materials that a laboratory might measure. For example, consider a laboratory faced with measuring three different sample materials: A, B, and C. If the laboratory takes measurements that are on-average -10 ^{14}C years from the target

value of sample A, this mean offset will be also be present when the same laboratory measures the ^{14}C values of sample materials B and C. If the systematic offset varies between sample materials A, B, and C, this cannot be captured by the model. This model also assumes that within-laboratory repeatability is uniform across sample materials A, B, and C. In reality, different sample material types (e.g., bone, wood, grass seeds) and variability in target ^{14}C values may affect both systematic offsets and within-laboratory repeatability. Unfortunately, the available VIRI dataset is insufficient to explore these issues in detail with this model.

The LBM was specified using the *map2stan* function in the *rethinking* R package (5) and fitted through Hamiltonian Monte Carlo (HMC) simulation in RStan (6). The model was fitted with four chains, each of which performed 5000 warmup and 2500 sampling iterations (10,000 total sampling iterations).

1.3. Posterior Parameter Values

Posterior estimates for each sample material ^{14}C value are close to each material's observed sample median (Figure S3.3a). Each posterior distribution for these sample materials includes its respective observed median value in a high-density region (since sample values are median centered, the observed median value is 0 for each distribution). The posterior distribution for C_{AMS} has a mean of 1.26, indicating that GPC/LSC laboratories generally have larger mean offsets from target ^{14}C values than do AMS laboratories (Figure S3.3b). However, this posterior is fairly dispersed, with 36.6% of the distribution falling below 1. Values below 1 correspond to a scenario where GPC/LSC laboratories have smaller mean offsets than those mean offsets associated with AMS laboratories.

The modelled laboratory offset parameters, $C_o[\text{lab } i]$, have mean posterior parameter values ranging from -29.0 to 26.3 ^{14}C years across the 68 laboratories (Table S3.1; Figure S3.4a). These posteriors show high overlap. At first glance, laboratory offset posteriors appear to show that between-laboratory variability is much higher than within-laboratory variability (Figure S3.4). However, this is only in the hypothetical scenario where a laboratory reports 0 measurement error. When measurement error is included, modelled within laboratory σ values increase rapidly and exceed the mean laboratory offsets (Figure S3.5; Figure S3.3e).

Posterior distributions for σ parameters are expressed on the log scale, which makes their combined effects on σ less intuitive to interpret than the previously discussed parameters (Figure S3.5c-e). The posterior for the global parameter for σ , σ_i , has a mean value of 2.8 when this distribution is exponentiated and transformed back into the scale of ^{14}C years (95% HPDI: 0.7—5.6). This represents within-laboratory measurement repeatability for the *average* AMS laboratory when reported ^{14}C measurement error is zero. When σ_{lab} distributions are added to this average value, within-laboratory repeatability varies between laboratories. For the 68 labs in this dataset, the mean posterior σ value ranges from 1.53 to 9.92 ^{14}C years (Table S3.1). However, these posterior distributions are dispersed and show considerable overlap. In general, the effect of σ_{AMS} causes GPC/LSC laboratories to have higher within-laboratory variability than AMS laboratories (Figure S3.3d; Figure S3.4b; Figure S3.5).

The generative aspect of this model allows one to simulate hypothetical pairs of laboratory parameters (Figure S3.6). As expected, GPC/LSC laboratories have generally larger mean offsets and within-laboratory σ values than AMS laboratories.

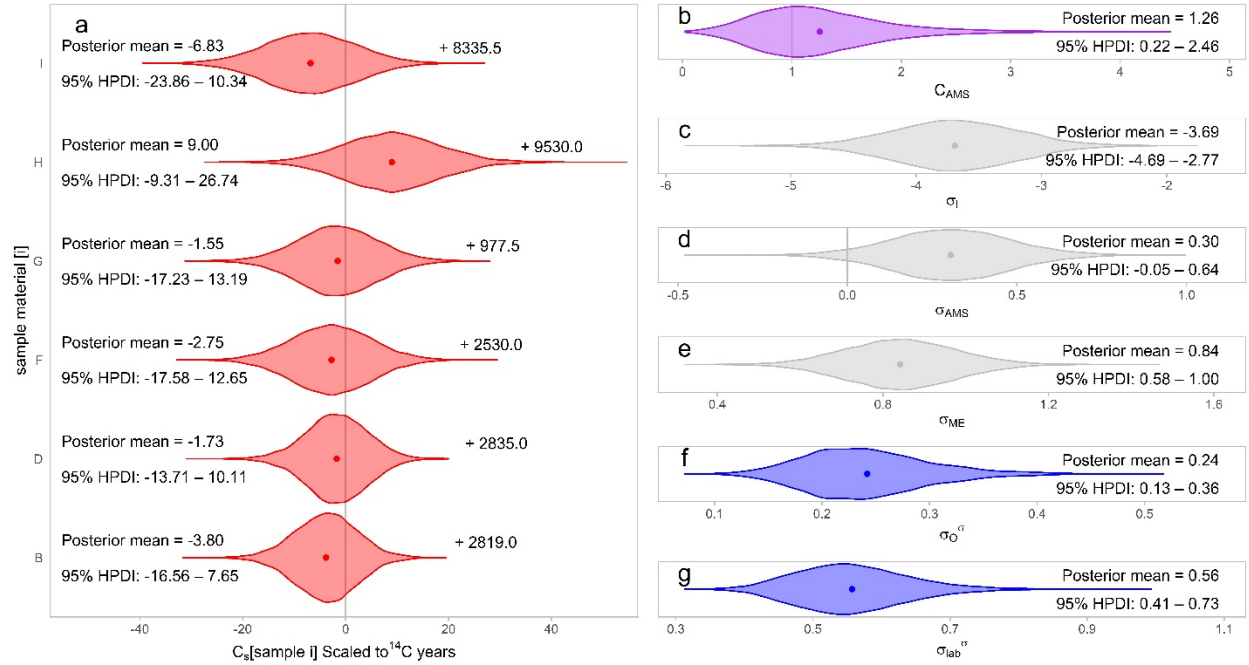


Figure S3.3. Posterior densities, posterior means (dots), and 95% highest posterior density intervals (HPDI) for model parameters. (a) Posteriors for each median-centered sample material. The sample medians are displayed to the right of each density to indicate over which ^{14}C years the non-centered distributions fall. (b) Posterior distribution for C_{AMS} , which adjusts mean laboratory offsets if they are GPC/LSC measurements. Note, most of the density sits above 1, indicating that GPC/LSC laboratory offsets are probably more dispersed than their AMS counterparts. (c-d) Posterior distributions for the standard deviation parameters (σ_I , σ_{AMS} , and σ_{ME}), which are displayed on the log scale. (f-g) Posterior distributions for the standard deviations of the distributions of each laboratory specific parameter, σ_O^σ and σ_{lab}^σ (i.e., mean laboratory-specific offsets and laboratory-specific standard deviations).

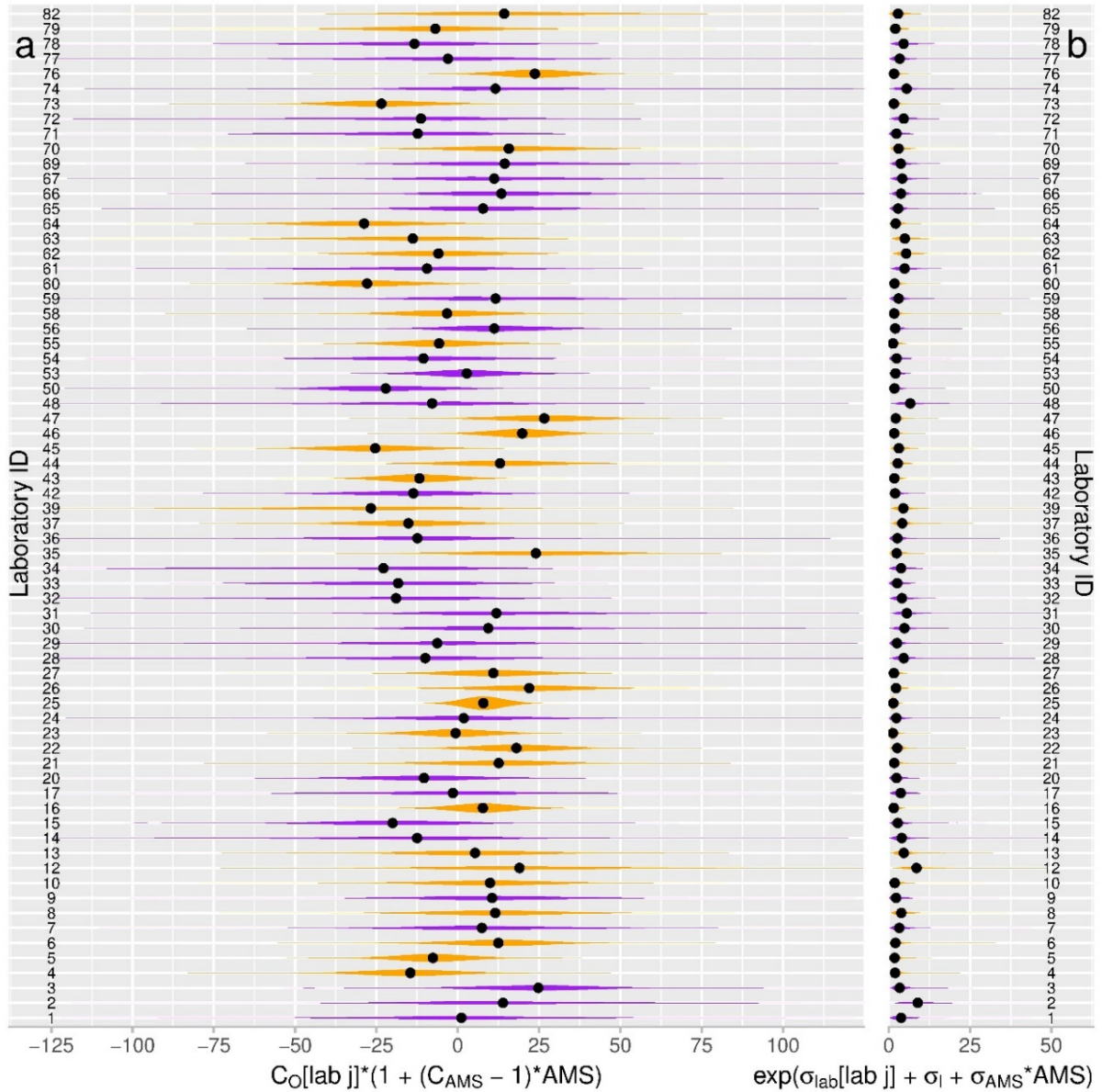


Figure S3.4. Posterior distributions for (a) mean laboratory offsets and (b) within-laboratory standard deviations. Gold densities show AMS laboratories and purple densities show GPC/LSC laboratories. GPC/LSC effects on laboratory offsets and within laboratory standard deviation values are included in these posterior distributions. Black dots mark the medians of each distribution. Note, within-laboratory standard deviations (b) assume 0 reported measurement error, and, in practice, these values become larger (see Figure S3.5). X-axes are on the ^{14}C year scale.

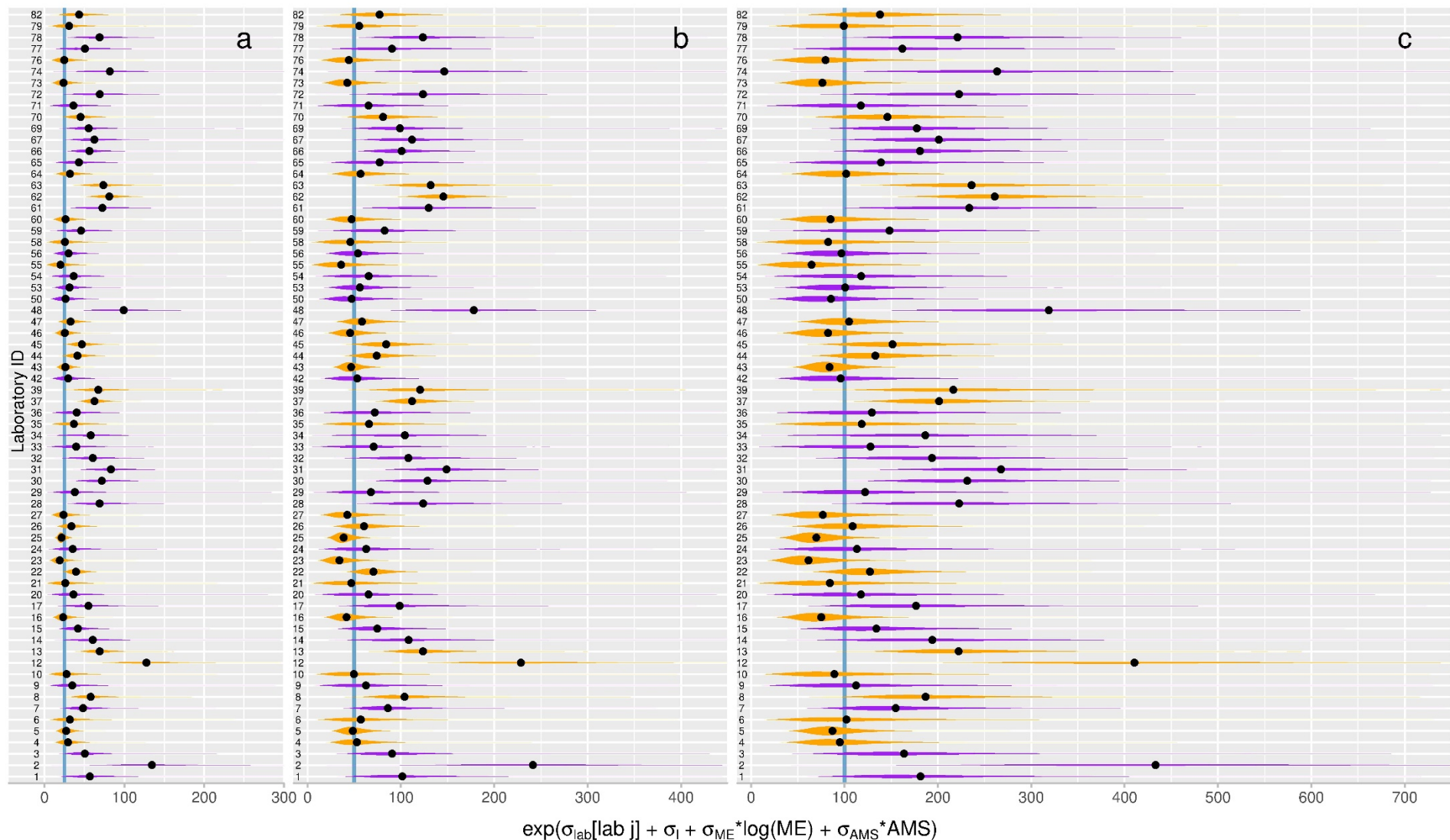


Figure S3.5. Three examples showing how reported measurement error (ME) affects within-laboratory standard deviations: (a) ME = 25 ^{14}C years, (b) ME = 50 ^{14}C years, and (c) ME = 100 ^{14}C years. Vertical blue lines mark the reported error values. Gold posterior densities are AMS laboratories and purple posterior densities are GPC/LSC laboratories. Black dots mark median values in each posterior density. GPC/LSC effects on within-laboratory standard deviations are included in these posterior distributions. X-axes are on the ^{14}C year scale.

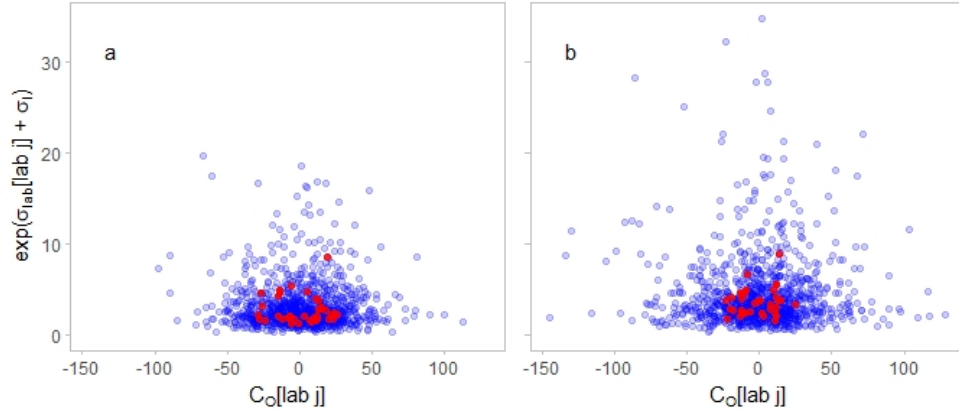


Figure S3.6. Laboratory parameters for (a) AMS and (b) GPC/LSC laboratories. Blue dots show 1000 simulated laboratories and red dots show mean posterior values for laboratories in the VIRI dataset. Both axes are displayed on the ^{14}C year scale.

Table S3.1. Posterior means and 95% highest posterior density intervals (HPDI) for VIRI laboratory parameters. The effects of GPC/LSC measurement methods are excluded here. In other words, all laboratories are treated here as AMS laboratories to express variability that is due to laboratory identity exclusive of ^{14}C measurement method.

Lab. ID [j]	Laboratory mean offset (^{14}C years) $C_0[\text{lab } j]$		Intra-laboratory standard deviation (^{14}C years) $\exp(\sigma_{\text{lab}}[\text{lab } j] + \sigma_l)$	
1	1.27	(-48.45 – 47.78)	3.54	(0.41 – 8.77)
2	14.87	(-38.29 – 71.87)	7.89	(1.44 – 17.99)
3	21.31	(-8.09 – 56.94)	2.88	(0.65 – 6.11)
4	-14.20	(-42.09 – 13.91)	2.32	(0.47 – 5.08)
5	-7.52	(-29.32 – 14.78)	2.09	(0.47 – 4.42)
6	11.97	(-18.54 – 41.52)	2.62	(0.41 – 6.40)
7	7.09	(-31.19 – 49.82)	2.85	(0.48 – 6.37)
8	11.62	(-25.75 – 47.68)	4.54	(0.94 – 9.87)
9	8.95	(-30.30 – 49.86)	2.18	(0.24 – 5.52)
10	9.47	(-27.11 – 45.79)	2.32	(0.25 – 5.59)
12	21.16	(-27.98 – 77.48)	9.92	(2.11 – 21.35)
13	5.61	(-31.05 – 41.74)	5.30	(1.25 – 11.18)
14	-12.61	(-65.43 – 37.58)	3.70	(0.53 – 8.91)
15	-16.92	(-52.35 – 17.87)	2.44	(0.42 – 5.30)
16	7.74	(-10.95 – 26.83)	1.82	(0.35 – 3.92)
17	-1.70	(-51.05 – 46.33)	3.49	(0.41 – 8.72)
20	-8.79	(-44.27 – 26.74)	2.24	(0.19 – 5.50)
21	11.63	(-24.49 – 43.44)	2.19	(0.18 – 5.57)
22	18.04	(-6.20 – 42.05)	3.03	(0.76 – 6.40)
23	-0.60	(-23.73 – 22.72)	1.53	(0.24 – 3.42)
24	2.00	(-38.73 – 45.11)	2.21	(0.20 – 5.35)
25	7.93	(-4.37 – 20.12)	1.60	(0.49 – 3.10)
26	21.80	(-8.74 – 50.63)	2.65	(0.58 – 5.81)
27	10.90	(-17.98 – 41.21)	1.92	(0.37 – 4.29)

28	-9.63	(-58.45 – 39.77)	4.14	(0.78 – 9.69)
29	-6.62	(-55.05 – 43.62)	2.42	(0.22 – 6.21)
30	9.20	(-35.61 – 55.32)	4.28	(0.69 – 9.66)
31	11.28	(-33.59 – 57.48)	4.85	(0.84 – 10.67)
32	-18.65	(-73.86 – 29.13)	3.62	(0.47 – 8.45)
33	-15.23	(-59.45 – 28.70)	2.42	(0.27 – 6.03)
34	-21.21	(-73.84 – 29.64)	3.47	(0.30 – 8.31)
35	23.47	(-17.02 – 63.13)	3.03	(0.33 – 7.25)
36	-10.87	(-54.42 – 29.74)	2.47	(0.33 – 6.01)
37	-15.38	(-45.41 – 15.47)	4.73	(1.16 – 9.54)
39	-29.02	(-84.32 – 20.78)	5.36	(0.81 – 11.78)
42	-11.70	(-46.12 – 24.47)	1.80	(0.26 – 4.24)
43	-11.68	(-32.10 – 9.40)	2.02	(0.44 – 4.22)
44	13.21	(-20.36 – 44.39)	3.24	(0.59 – 7.06)
45	-25.20	(-52.19 – 2.26)	3.59	(0.82 – 7.41)
46	19.72	(-0.37 – 39.30)	1.95	(0.45 – 4.03)
47	26.32	(-1.93 – 52.93)	2.53	(0.53 – 5.38)
48	-8.16	(-59.22 – 40.86)	5.93	(0.94 – 13.58)
50	-18.49	(-51.68 – 12.90)	1.58	(0.20 – 3.61)
53	2.49	(-21.70 – 24.94)	1.82	(0.34 – 3.93)
54	-8.95	(-48.71 – 30.48)	2.26	(0.25 – 5.58)
55	-5.29	(-34.24 – 22.08)	1.67	(0.14 – 4.15)
56	9.84	(-21.37 – 39.56)	1.79	(0.30 – 4.04)
58	-3.11	(-32.11 – 28.43)	2.14	(0.17 – 5.16)
59	11.95	(-35.21 – 61.90)	2.84	(0.29 – 6.99)
60	-26.96	(-53.13 – 0.31)	2.10	(0.37 – 4.65)
61	-9.55	(-58.01 – 39.37)	4.40	(0.62 – 10.24)
62	-6.03	(-38.43 – 26.03)	6.26	(1.31 – 13.27)
63	-14.42	(-60.64 – 27.33)	5.83	(1.07 – 12.76)
64	-28.35	(-60.61 – 6.38)	2.55	(0.44 – 5.71)
65	6.83	(-29.71 – 45.51)	2.62	(0.35 – 6.23)
66	13.15	(-21.49 – 52.30)	3.26	(0.54 – 7.00)
67	10.71	(-34.82 – 57.40)	3.77	(0.62 – 8.79)
69	13.29	(-30.11 – 60.55)	3.29	(0.50 – 7.46)
70	15.81	(-17.36 – 51.36)	3.53	(0.74 – 7.60)
71	-11.08	(-55.76 – 34.09)	2.24	(0.22 – 5.64)
72	-12.08	(-65.29 – 41.20)	4.21	(0.56 – 10.05)
73	-22.75	(-50.96 – 6.54)	1.92	(0.28 – 4.38)
74	12.59	(-38.32 – 67.89)	4.98	(0.68 – 11.78)
76	23.24	(0.83 – 44.15)	1.98	(0.34 – 4.50)
77	-3.51	(-51.84 – 47.22)	3.24	(0.34 – 8.18)
78	-13.89	(-60.91 – 29.29)	4.11	(0.60 – 9.37)
79	-6.77	(-42.76 – 29.40)	2.61	(0.32 – 6.37)
82	15.92	(-33.75 – 69.33)	3.67	(0.40 – 9.18)

1.4. HMC Diagnostics and Posterior Predictive Checks

To ensure convergence of HMC chains, we examined trace plots, R-hat values, and effective sample sizes. Trace plots suggest convergence of HMC chains for all parameters. All R-hat values are below 1.01. Effective sample sizes close to 10,000 indicate efficient sampling of posterior parameter spaces. Of the 148 parameters, 101 have effective sample sizes of 10,000. The median effective sample size is also 10,000, the mean is 8362, and the minimum is 1056.4 (Figure S3.7).

We also performed a posterior predictive check by simulating data from the model and plotting these simulated data against the VIRI ^{14}C measurements (Figure S3.8). For simplicity, where a laboratory contributed multiple ^{14}C measurements for a given sample material, we used the average reported measurement error for that laboratory ID and sample (for example, laboratories 5 and 22 in the panel for Sample G). In all but nine of the 361 VIRI ^{14}C measurements (97.5%), the 95% prediction intervals overlap the observed values. This indicates that the LBM does a reasonable job recovering the observed values.

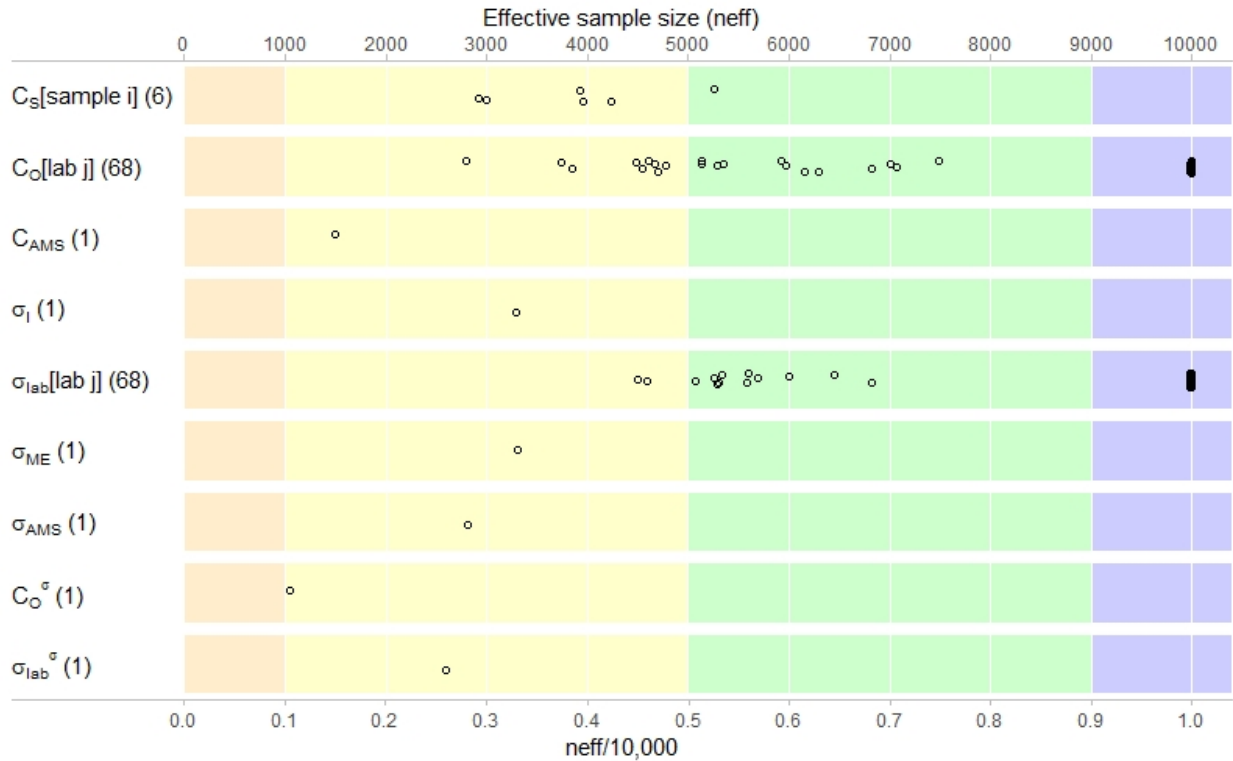


Figure S3.7. Distribution of effective sample sizes across parameters. Parenthetical values show the number of parameter distributions that were sampled for each parameter type.

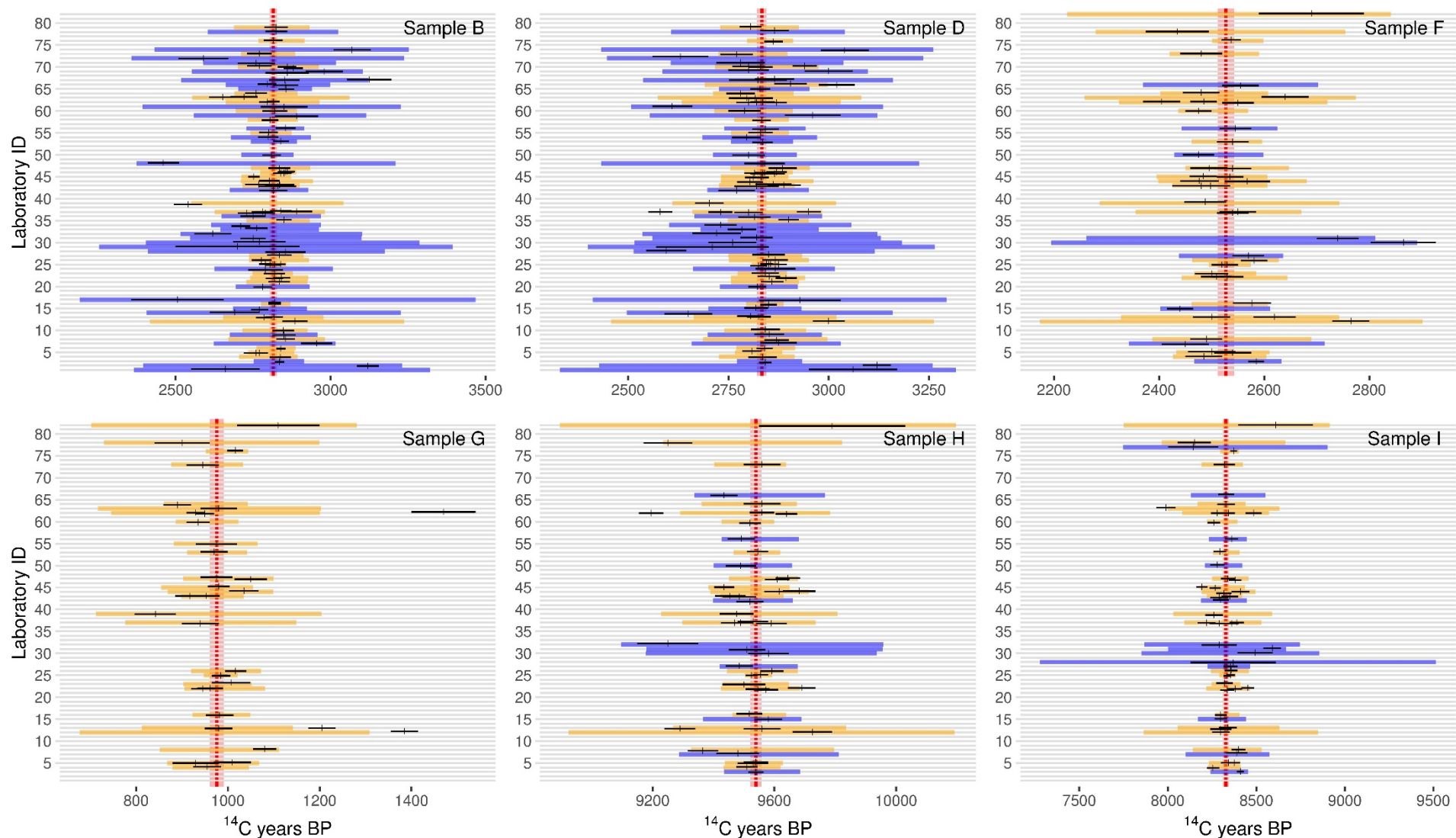


Figure S3.8. Posterior predictive check for the LBM. Each panel is a sample material reported in the VIRI study (Scott et al. 2007, 2010a, 2010b). Laboratory IDs are listed along the y-axes. Reported ^{14}C measurements and their associated errors (1σ) are indicated by vertical and horizontal black segments, respectively. Horizontal gold and purple bars show the 95% posterior prediction intervals for AMS and GPC/LSC laboratories, respectively. The dashed red lines and bands indicate the mean and 95% highest posterior density intervals for the ^{14}C value of each sample material.

Citations

1. Scott EM, Cook GT, Naysmith P, Bryant C, O'Donnell D (2007) A Report on Phase 1 of the 5th International Radiocarbon Intercomparison (VIRI). *Radiocarbon* 49(02):409–426.
2. Scott EM, Cook GT, Naysmith P (2010) The Fifth International Radiocarbon Intercomparison (VIRI): An Assessment of Laboratory Performance in Stage 3. *Radiocarbon* 52(03):859–865.
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