## **Problem Set 11**

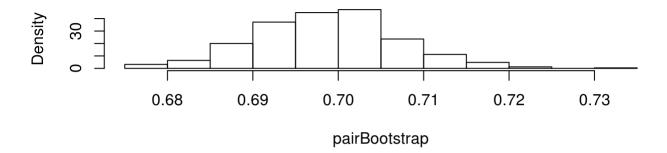
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## Problem 1:

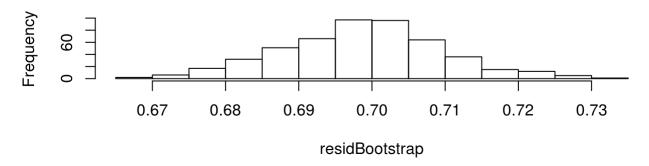
We can see from the histograms below that the pair bootstrap has a longer tail. In gerneral it sees to be a bit more skewed. Though both are centerted around the same mean.

```
#### Set up data ####
R = c(68, 77, 299, 220, 142, 287, 276, 115, 64, 206, 222, 205, 233, 228, 188, 132, 285, 188, 224
, 121, 311, 166, 248, 161, 226, 67, 201, 267, 121, 301, 244, 222, 195, 203, 210, 275, 286, 275,
304, 214)
S = c(56, 62, 445, 279, 138, 428, 319, 102, 51, 289, 351, 282, 310, 266, 256, 144, 447, 186, 389)
, 113, 412, 176, 313, 162, 368, 54, 214, 429, 115, 407, 265, 301, 234, 229, 270, 478, 419, 490,
430, 235)
salmon <- as.data.frame(R)</pre>
salmon$S <- S
salmon$R2 <- 1/R
salmon$S2 <- 1/S
#### First regression #####
origRgrsn <- lm(formula = R2 ~ S2, data=salmon)
#### Bootstrap Bill Turner ####
  # Set Parameters
    B <- 500
    pairBootstrap = rep(0,B)
      residBootstrap = rep(0,B)
  for (j in 1:B) {
       bootstrap covariates and response
      ## Get Sample Data - sample by row number
          resampledObs <- base::sample(seq(1:length(salmon[,1])),length(salmon[,1]), replace = T
RUE)
          newDataSet = salmon[resampledObs,]
        ## Run regression and store coefficient
            newRgrsn = lm(formula = R2 ~ S2, data=newDataSet)
            pairBootstrap[j] = newRgrsn$coefficients[2]
        bootstrap residuals
          ## The the residual for each observation.
            resampledResids <- origRgrsn$residuals[resampledObs]</pre>
          ## Get prediction
            salmon$alteredRspns = resampledResids + predict(origRgrsn)
        newResidRgrsn = lm(alteredRspns ~ S2,data=salmon)
        residBootstrap[j] = newResidRgrsn$coefficients[2]
  }
#plot(S,R)
par(mfrow = c(2,1))
hist(pairBootstrap, freq=FALSE)
hist(residBootstrap)
```

#### Histogram of pairBootstrap



#### Histogram of residBootstrap



## Problem 2

Generate 100 samples  $X_1, \cdots, X_{20}$  from a normal population  $N(\theta,1)$  with  $\theta=1$ .

(a) For each sample compute the bootstrap and jackknife estimate for variance for  $\hat{\theta}=\bar{X}$  and compute the mean and standard deviation of these variance estimates over the 100

samples.

```
B <- 1000
theta <- c()
## Build the functions
  bootstrap.var <- function(DATA,B){</pre>
      for (i in 1:B){
           # Make sure the data is in vector form
             vctrzdData <- as.vector(DATA)</pre>
           # Sample of size n with replacement from data
             btstrpSmpl = base::sample(vctrzdData,length(vctrzdData), replace = TRUE)
           # Sort
             btstrpSmpl.ordered <- sort(btstrpSmpl)</pre>
           # Take the mean of the middle 4 numbers
           # theta[i] <- mean(btstrpSmpl.ordered[3:6])</pre>
             theta[i] <- mean(btstrpSmpl.ordered)</pre>
        }
    # Take the mean of the B thetas
      theta.mean <- mean(theta)</pre>
    # Take the variance of the B thetas (var(theta) also works)
      theta.var \langle (1/(B-1))*(sum((theta-theta.mean)^2))
      return(c(theta.mean,theta.var))
  }
  jackknife.var <- function(DATA){</pre>
    # Clear theta
      theta <- c()
    # Get the Length of DATA
      n <- length(DATA)</pre>
    # Get the theta j values for the jackknife samples.
      for (j in 1:n){
        # Generate the sample excluding the ith value of X1
           btstrpSmpl <- DATA[-j]</pre>
        # Take the mean of the sample.
           theta[j] <- mean(btstrpSmpl)</pre>
      }
    # Compute theta(.)
      theta.dot <- mean(theta)</pre>
    # Jackknife variance
      dif <- (theta - theta.dot)^2</pre>
      var.jackknife <- sqrt(((n-1)/n)*sum(dif))</pre>
    # Return
        return(c(theta.dot,var.jackknife))
    # print(paste0("Jackknife: Mean = ",theta.dot," Variance = ", var.jackknife))
  }
## Run
jack.mean <- c()</pre>
jack.var <- c()</pre>
boot.mean <- c()</pre>
boot.var <- c()</pre>
for ( i in 1:100){
  # Generate X1
```

```
X1 <- rnorm(n=20,mean=1,sd=1)
# Bootstrap
boot <- bootstrap.var(X1,B)
boot.mean[i] <- boot[1]
boot.var[i] <- boot[2]
# Jackknife
jack <- jackknife.var(X1)
jack.mean[i] <- jack[1]
jack.var[i] <- jack[2]
}

print(paste0("Bootstrap: B = ",B, " Mean = ",mean(boot.mean), " mean(Variance) = ", mean(boot.var), " St. Dev.(variance) = ",sd(boot.var)))</pre>
```

```
## [1] "Bootstrap: B = 1000 \text{ Mean} = 0.994970100498525 \text{ mean(Variance)} = 0.047877760262376 \text{ St. Dev.} (variance) = 0.0170191611810896"
```

```
print(paste0("Jackknife: Mean = ",mean(jack.mean)," mean(Variance) = ", mean(jack.var)," St. De
v.(variance) = ",sd(jack.var)))
```

```
## [1] "Jackknife: Mean = 0.994536074257312 mean(Variance) = 0.222451568199411 St. Dev.(variance) = 0.0409148554100715"
```

# (b) Repeat (a) for $\hat{ heta}=ar{X}^2$

```
B <- 1000
theta <- c()
bootstrap.var <- function(DATA,B){</pre>
    for (i in 1:B){
        # Make sure the data is in vector form
           vctrzdData <- as.vector(DATA)</pre>
        # Sample of size n with replacement from data
           btstrpSmpl = base::sample(vctrzdData,length(vctrzdData), replace = TRUE)
        # Sort
           btstrpSmpl.ordered <- sort(btstrpSmpl)</pre>
        # Take the mean of the middle 4 numbers
        # theta[i] <- mean(btstrpSmpl.ordered[3:6])</pre>
           theta[i] <- mean(btstrpSmpl.ordered)</pre>
      }
  # Take the mean of the B thetas
    theta.mean <- (mean(theta))^2
  # Take the variance of the B thetas (var(theta) also works)
    theta.var \langle (1/(B-1))*(sum((theta-theta.mean)^2))
  # Return
    return(c(theta.mean,theta.var))
}
jackknife.var <- function(DATA){</pre>
  # Clear theta
    theta <- c()
  # Get the length of DATA
    n <- length(DATA)</pre>
  # Get the theta_j values for the jackknife samples.
    for (j in 1:n){
      # Generate the sample excluding the ith value of X1
        btstrpSmpl <- DATA[-j]</pre>
      # Take the mean of the sample.
        theta[j] <- (mean(btstrpSmpl))^2</pre>
  # Compute theta(.)
    theta.dot <- mean(theta)</pre>
  # Jackknife variance
    dif <- (theta - theta.dot)^2</pre>
    var.jackknife <- sqrt(((n-1)/n)*sum(dif))</pre>
  # Return
      return(c(theta.dot,var.jackknife))
  # print(paste0("Jackknife: Mean = ",theta.dot," Variance = ", var.jackknife))
}
jack.mean <- c()</pre>
jack.var <- c()</pre>
boot.mean <- c()</pre>
boot.var <- c()</pre>
for ( i in 1:100){
  # Generate X1
    X1 <- rnorm(n=20,mean=1,sd=1)</pre>
  # Bootstrap
```

```
boot <- bootstrap.var(X1,B)
        boot.mean[i] <- boot[1]
        boot.var[i] <- boot[2]

# Jackknife
    jack <- jackknife.var(X1)
        jack.mean[i] <- jack[1]
        jack.var[i] <- jack[2]
}

print(paste0("Bootstrap: B = ",B, " Mean = ",mean(boot.mean), " mean(Variance) = ", mean(boot.var), " St. Dev.(variance) = ",sd(boot.var)))</pre>
```

```
## [1] "Bootstrap: B = 1000 Mean = 1.1038158182034 mean(Variance) = 0.110581802674557 St. Dev.(v ariance) = 0.126440650553676"
```

```
print(paste0("Jackknife: Mean = ",mean(jack.mean)," mean(Variance) = ", mean(jack.var)," St. De
v.(variance) = ",sd(jack.var)))
```

```
## [1] "Jackknife: Mean = 1.10764142696516 mean(Variance) = 0.449888186545517 St. Dev.(variance) = 0.114055888362321"
```

### Problem 3

Find the number of bootstraps necessary to accurately calculate a 95% confidence interval in the difference of two population means. That is, simulate n values of X, where \$X\_i N(\_1, \_1^2) and m values of Y , where \$Y\_j N(\_2, \_2^2) and calculate a 95% confidence interval in  $\mu_1 - \mu_2$ . Now calculate a 95% confidence interval for  $\mu_1 - \mu_2$  using the bootstrap techniques learned in this module. How large does B have to be until you accurately and dependably reproduce the confidence interval you originally calculated?

```
# Set parameters
  B <- 200
  n <- 200
  mu1 <- 1
  sig1 <- 1
  m <- 20
  mu2 <- 1
  sig2 <- 2
# Generate the true samples
  X <- rnorm(n,mean=mu1,sd=sig1)</pre>
  Y <- rnorm(m,mean=mu2,sd=sig2)
# Theoretical Confidence interval
  # Sample variance (Larsen & Marx pg 475)
    sample.var <- function(DATA){</pre>
      (length(DATA)*(sum(DATA^2)) - (sum(DATA))^2)/(length(DATA)*(length(DATA)-1))
    }
  # T statistic
    t \leftarrow qt(0.025, df=(n+m-2), lower.tail = FALSE)
    term <- t*(sqrt((sample.var(X)/n)+(sample.var(Y)/m)))</pre>
  # Confidence interval
    print(paste0("The 95% Confidence interval for X-Y is (",
                  mean(Y)-mean(Y)-term,",", mean(Y)-mean(Y)+term,")"))
```

```
## [1] "The 95% Confidence interval for X-Y is (-0.665609515885879,0.665609515885879)"
```

```
# Bootstrap
    mean.X <- c()
    mean.Y <- c()
  for (i in 1:B){
      # Sample of size n with replacement from data
        btstrpSmpl.X = base::sample(X,length(X), replace = TRUE)
        btstrpSmpl.Y = base::sample(Y,length(Y), replace = TRUE)
      # Mean
        mean.X[i] <- mean(btstrpSmpl.X)</pre>
        mean.Y[i] <- mean(btstrpSmpl.Y)</pre>
# Calculate theta
  theta <- mean.X-mean.Y
# Order Theta
  theta.ordered <- sort(theta)</pre>
# Where do the intervales lie
  min < -(B-(0.95*B))/2
# Interval
  print(paste0("The 95% Confidence interval for X-Y is (",
                theta.ordered[min],",", theta.ordered[(B-min)],")"))
```

```
## [1] "The 95% Confidence interval for X-Y is (-1.54235337620315,-0.116471898835519)"
```

```
# Set parameters
  B <- 200
  n <- 200
  mu1 <- 1
  sig1 <- 1
  m <- 20
  mu2 <- 1
  sig2 <- 2
# Generate the true samples
  X <- rnorm(n,mean=mu1,sd=sig1)</pre>
  Y <- rnorm(m, mean=mu2, sd=sig2)
# Theoretical Confidence interval
  # Sample variance (Larsen & Marx pg 475)
    sample.var <- function(DATA){</pre>
      (length(DATA)*(sum(DATA^2)) - (sum(DATA))^2)/(length(DATA)*(length(DATA)-1))
    }
  # T statistic
    t \leftarrow qt(0.025,df=(n+m-2), lower.tail = FALSE)
    term <- t*(sqrt((sample.var(X)/n)+(sample.var(Y)/m)))</pre>
  # Confidence interval
    print(paste0("The 95% Confidence interval for X-Y is (",
                  mean(Y)-mean(Y)-term,",", mean(Y)-mean(Y)+term,")"))
```

```
## [1] "The 95% Confidence interval for X-Y is (-0.955542866058812,0.955542866058812)"
```

```
# Bootstrap
    bootstrap.twosamp.confint <- function(DATA1,DATA2,B){</pre>
         mean.X <- c()
        mean.Y \leftarrow c()
        theta <- c()
         min < - c()
        max <- c()
      for (i in 1:B){
           # Sample of size n with replacement from data
             btstrpSmpl.X = base::sample(DATA1,length(DATA1), replace = TRUE)
             btstrpSmpl.Y = base::sample(DATA2,length(DATA2), replace = TRUE)
           # Mean
             mean.X[i] <- mean(btstrpSmpl.X)</pre>
             mean.Y[i] <- mean(btstrpSmpl.Y)</pre>
           # Calculate theta
             theta[i] <- mean.X[i]-mean.Y[i]</pre>
         }
    # Order Theta
      theta.ordered <- sort(theta)</pre>
    # Where do the intervals lie
      int \langle -(B-(0.95*B))/2 \rangle
    # Interval
      min <- theta.ordered[int]</pre>
      max <- theta.ordered[(B-int)]</pre>
      return(c(min,max))
    }
# Lets do this for B
  bootstrap.conf <- bootstrap.twosamp.confint(X,Y,B)</pre>
    print(paste0("The 95% Confidence interval for X-Y is (",
                 bootstrap.conf[1],",", bootstrap.conf[2],")"))
```

```
## [1] "The 95% Confidence interval for X-Y is (-0.84968537443988,0.863594086427095)"
```

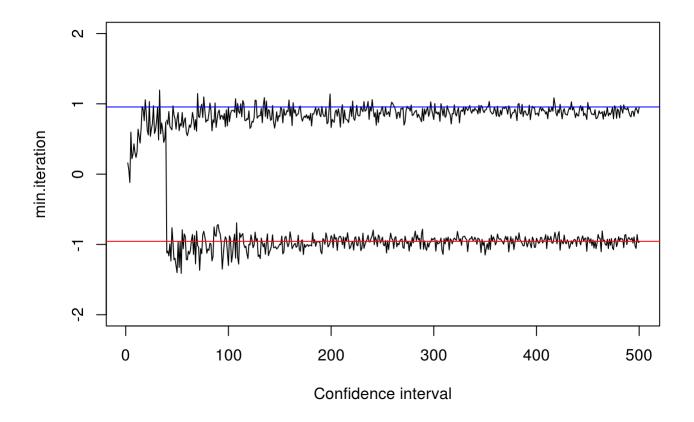
```
# For the plot
min.iteration <- c()
max.iteration <- c()
for (j in 1:500){
   iteration <- bootstrap.twosamp.confint(X,Y,j)
   min.iteration[j] <- iteration[1]
   max.iteration[j] <- iteration[2]
}

plot(min.iteration, type='l',
        ylim=c(-2,2),xlab="Confidence interval")
lines(max.iteration, type='l')
abline(h=(mean(Y)-mean(Y)-term), type='l', col='red')</pre>
```

```
## Warning in int_abline(a = a, b = b, h = h, v = v, untf = untf, ...): graphical
## parameter "type" is obsolete
```

```
abline(h=(mean(Y)-mean(Y)+term), type='l', col='blue')
```

```
## Warning in int_abline(a = a, b = b, h = h, v = v, untf = untf, ...): graphical ## parameter "type" is obsolete
```



From the plot above, it looks like the confidence interval is pretty stable once it reaches (approximately) B= 40.