Problem Set 4 *

Ian McGroarty

Course Number: 625.603

February 28, 2019

^{*}Problem list 3.10.6, 3.10.16, 3.12.6, 3.12.8

Problem 3.10.6. Let $Y_1, Y_2, ..., Y_n$ be a random sample from the exponential pdf $f_y(y) = e^{-y}, y \ge 0$. What is the smallest n for which $P(Y_{min} < 0.2) > 0.9$?

Solution The evaluated n > 11.513, since x must be an integer (I'm assuming since these are trials) $n \ge 12$.

$$P(Y_{min} < 0.2) = \int_{0}^{0.2} f_{Y_{min}}(y)$$

$$= \int_{0}^{0.2} n[1 - F_{Y}(y)]^{n-1} f_{Y}(y) \qquad \text{Theorem 3.10.1.b (pg193)}$$

$$= \int_{0}^{0.2} n[1 - (1 - e^{-y})]^{n-1} (e^{-y})$$

$$= \int_{0}^{0.2} n(e^{-y})^{n-1} (e^{-y})$$

$$= \int_{0}^{0.2} n(e^{-ny})$$

$$= -e^{-ny} \Big|_{0}^{0.2}$$

$$= -e^{-(0.2)n} - (-1) > 0.9$$

$$= \log(e^{-0.2n}) < \log(0.1)$$

$$n > \frac{\log(0.1)}{-0.2} \approx 11.513$$

Problem 3.10.16. Suppose a device has three independent components, all of whose lifetimes (in months) are modeled by the exponential pdf, $f_y(y) = e^{-y}$, y > 0. What is the probability that all three components will fail within two months of one another?

Solution¹ Range = $Y_{max} - Y_{min} = Y_3^{'} - Y_1^{'}$. P(R < r) = 0.646

The memoryless property of the exponential distribution:

$$P(X \ge s + t | x \ge s) = P(X \ge t)$$

This implies that the level of Y_1 is inconsequential. Thus, we can assume that $Y_1 = 0$. In which case we are really only interested in $P(Y_{max} < r)$. Thus, we can apply Theorem 3.10.1.a (pg 193) with: n=3, $f_y(y) = e^{-y}$, and $F_Y(y) = \int_0^y f_Y(y) dy = 1 - e^{-y}$.

$$P(Y_{max} < m) = \int_{-\infty}^{m} n[F_Y(y)]^{n-1} f_Y(y)$$

$$P(Y_3^{'} < 2) = \int_0^2 3[1 - e^{-y}]^2 e^{-y}$$
 Enter WolframAlpha
$$\approx 0.646.$$

Problem 3.12.6. Find $M_Y(t)$ if Y has the pdf:

$$f_Y(y) = \begin{cases} y, & 0 \le y \le 1 \\ 2 - y, & 1 \le y \le 2 \\ 0, & \text{elsewhere} \end{cases}$$

Solution
$$M_Y(t) = \frac{1-e^t}{t^2}$$

¹To start, I want to note that understanding of the "memoryless property of the exponential distribution" was critical to even approaching success in this problem. I studied the proof in this pdf, http://www.cs.cmu.edu/afs/cs/academic/class/15750-s19/OldScribeNotes/lecture11.pdf (pg 2). I also used wolframalpha to do some of the calculations that were a to complex for my patience level.

Since X is continuous, the second part of Definition 3.12.1 (pg 206) applies:

$$\begin{split} M_Y(t) &= E(e^{tW}) = \int_{-\infty}^{\infty} e^{tw} f_W(w) dw \\ &= \int_{-\infty}^{0} e^{ty} 0 + \int_{0}^{1} e^{ty} y + \int_{1}^{2} e^{ty} (2-y) + \int_{0}^{\infty} e^{ty} 0 \\ &= \int_{0}^{1} e^{ty} y + \int_{1}^{2} e^{ty} (2-y) & \text{I used WolframAlpha here.} \\ &= (\frac{1}{t}y - \frac{1}{t^2})e^{ty}\Big|_{0}^{1} + \frac{2}{t}e^{ty}\Big|_{1}^{2} - (\frac{1}{t}y - \frac{1}{t^2})e^{ty}\Big|_{1}^{2} \\ &= (\frac{1}{t} - \frac{1}{t^2})e^t + (\frac{1}{t^2}) + \frac{2}{t}e^{2t} - \frac{2}{t}e^t - (\frac{2}{t} - \frac{1}{t^2})e^{2t} + (\frac{1}{t} - \frac{1}{t^2})e^t \\ &= \frac{1}{t^2} + \frac{1}{t^2}e^{2t} - \frac{2}{t^2}e^t \\ &= \frac{1}{t^2}(e^t - 1) \end{split}$$

Problem 3.12.8. Let Y be a continuous random variable with $f_Y(y) = ye^{-y}$, $o \le y$. Show that $M_Y(t) = \frac{1}{(1-t)^2}$

Solution Since X is continuous, the second part of Definition 3.12.1 (pg 206) applies:

$$\begin{split} M_Y(t) &= E(e^{tW}) = \int_{-\infty}^{\infty} e^{tw} f_W(w) dw \\ &= \int_{0}^{\infty} e^{ty} y e^{-y} \\ &= \int_{0}^{\infty} e^{ty-y} y \\ &= \frac{e^{(t-1)y}((t-1)y-1)}{(t-1)^2} \Big|_{0}^{\infty} \\ &= \lim_{y \to \infty} \frac{e^{(t-1)y}((t-1)y-1)}{(t-1)^2} - \lim_{y \to 0} \frac{e^{(t-1)y}((t-1)y-1)}{(t-1)^2} \\ &= \lim_{y \to \infty} \frac{e^{(t-1)y}((t-1)y-1)}{(t-1)^2} - \frac{e^{(t-1)0}((t-1)0-1)}{(t-1)^2} \\ &= \lim_{y \to \infty} \frac{e^{U}(U-1)}{(t-1)^2} - \frac{(-1)}{(t-1)^2} & \text{Distribute, let } U = (yt-y) \\ &= \frac{(\lim_{y \to \infty} e^{yt-y})(\lim_{y \to \infty} (yt-y-1))}{(t-1)^2} - \frac{(-1)}{(t-1)^2} & \text{Theorem 20.4 (Ross pg 156)} \end{split}$$

It will suffice to show that if $\lim_{y\to\infty} e^{yt-y}$ converges to 0 then $M_Y(t) \to \frac{1}{(t-1)^2}$. Since $0 * (\lim_{y\to\infty} (yt-y-1)) = 0$ and $\frac{0}{(t-1)^2} - \frac{(-1)}{(t-1)^2} = \frac{1}{(t-1)^2}$.

Proof. Let
$$\epsilon > 0$$
 and $\delta = \frac{\log(\epsilon)}{(t-1)}$. Note that, $\forall t$ such that $0 < t < 1$, $\delta > 0$. So if $0 < |x| < \delta$ then $|x| < \frac{\log(\epsilon)}{(t-1)} \implies (t-1)|x| = \log(\epsilon) \implies e^{xt-x} < \epsilon$ Thus, by Theorem 20.6 (Ross pg 159²), $\lim_{x\to\infty} e^{xt-x} = 0$.

²Ross, Kenneth: Elementary Analysis the Theory of Calculus. Undergraduate Texts in Mathematics, 2nd edn. Springer, New York/Heidelberg/Berlin (2013)

Problem Assignment. The Inverse Transform Sampling procedure requires $F_Y^-(u)$ where $F_Y^-(u)$ is the inverse of the cumulative distribution function, $F_Y(y)$. We saw in problem 3.10.6 that $F_Y(y) = (1 - e^-y)$. Solving for x = -log(1 - u) is our inverse CDF function. Now to be honest with you, I couldn't really follow your R code so I wrote my own. Now I'm not really sure what you are looking for to by way of the solution. But I've included my code below, and I am getting very close to 0.9 so I believe it is correct.

```
# inverse transfrom sampling

# This is the number of times you run the process,

positions <- 1000

# This will be a vector of minimums

samples <- c()

# You add one number to sample each process

# and you want to run the process position times.

while(length(samples) < positions)

{

#This is the number computed in the problem.

#The n needed to produce the minimum

num.samples <- 12
```

```
\# The uniform distribution from 0 to 1
                                                                         \leftarrow runif(num.samples, 0, 1)
\# The pdf
Y < - \exp(-U)
\# The inverse cdf function
                                                                        <--log(1-U)
 \# take the minimum
 sortedx \leftarrow sort(X)
 minx = min(sortedx)
 samples <- append(samples, minx)</pre>
  }
 \# Get the ratio of samples <0.2 to total samples.
 \# This is your probability.
  less < sum(samples < 0.2)
  ratio <- less/positions
\# p lot
 \mathbf{hist} \, (\, \mathbf{samples} \, , \, \, \, \mathbf{breaks} = 30, \, \, \, \mathbf{freq} = F, \, \, \, \mathbf{xlab} = \mathrm{'X'} \, , \, \, \, \mathbf{main} = \mathrm{'Generating} \, \subseteq \, \mathbf{Exponential} \, , \, \, \mathbf{xlab} = \mathrm{'X'} \, , \, \, \mathbf{main} = \mathrm{'Allowed Matter M
```

 $\mathbf{curve}(\mathbf{dexp}(X, \ \text{rate=2}) \ , \ 0\,, \ 3\,, \ \mathrm{lwd=2}, \ \mathrm{xlab} \ = \ "\," \,, \ \mathbf{ylab} \ = \ "\," \,, \ \mathbf{add} \ = \ T)$