STAT40810 — Stochastic Models

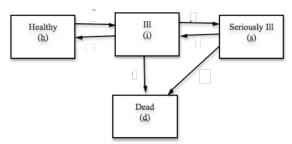
Brendan Murphy

Week 10

Multistate Markov Models

Multistate Markov Model

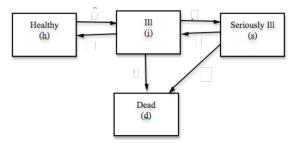
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Multistate Markov Model

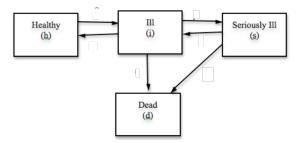
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- The process has four states:
 - healthy (h)
 - ill (i)
 - seriously ill (s)
 - dead (d)

Multistate Markov Model

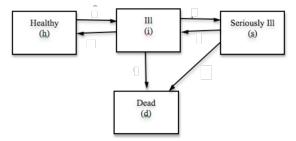
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 - healthy (h)
 - ill (i)
 - seriously ill (s)
 - dead (d)
- And people can move between states in continuous time.

Multistate Markov Model

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- The process has four states:
 - healthy (h)
 - ill (i)
 - seriously ill (s)
 - dead (d)
- And people can move between states in continuous time.
- The flow between states is controlled by unknown parameters.

Assumptions

- Consider the process at time t and what happens before time t + dt.
- \bullet We make the following assumptions:
 - The probabilities of being in any state at any subsequent time depends only on the time involved and on the state currently occupied.
 - 2 We assume a constant transition intensity λ_{gh} .

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Assumptions

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- We make the following assumptions:
 - 1 The probabilities of being in any state at any subsequent time depends only on the time involved and on the state currently occupied.
 - 2 We assume a constant transition intensity λ_{gh} .
 - **3** The probability of transitioning from state g to state h between time t and t+dt is $\lambda_{gh}dt+o(dt)$ as $dt\to 0$.
 - The probability of transitioning between two or more states is o(dt) as $dt \to 0$.

Data

• Suppose for a cohort of people we observed the following data:

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Data

• Suppose for a cohort of people we observed the following data:

Transi	Count		
I	h to i	n _{hi}	
i	to h	n _{ih}	
i	to <i>s</i>	n _{is}	
9	s to i	n _{si}	
i	to d	n _{id}	
	to d	n _{sd}	

Data

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Transitions	Count
h to i	n _{hi}
i to h	nih
i to s	n _{is}
<i>s</i> to <i>i</i>	nsi
<i>i</i> to <i>d</i>	n _{ia}
<i>s</i> to <i>d</i>	n _{sa}

• And...

State	Time
h	t _h
i	ti
S	t_s
d	t _d

Data

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s to i	n _{si}
i to d	n _{id}
s to d	n _{sd}

• And...

State	Time	
h	t _h	
i	ti	
S	t_s	
d	t_d	

• We have enough information to form the likelihood function.

Example

- The following data were collected from 100 individuals who were monitored:
 - total waiting time in state h: 4,100 months
 - total waiting time in state i: 1,000 months
 - total waiting time in state s: 500 months
 - total number of transfers form state h to state i: 110
 - total number of transfers from state i to state h: 90
 - total number of transfers from state i to state s: 40
 - total number of transfers from state s to state i: 20
 - total number of transfers from state i to state d: 7
 - total number of transfers from state s to state d: 10

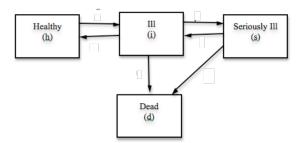
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Likelihood

• The likelihood is of the form

$$L = \exp(-\lambda_{hi}t_h) \exp(-(\lambda_{ih} + \lambda_{is} + \lambda_{id})t_i) \exp(-(\lambda_{si} + \lambda_{sd})t_s)$$
$$\lambda_{hi}^{n_{hi}} \lambda_{ih}^{n_{ih}} \lambda_{is}^{n_{id}} \lambda_{si}^{n_{id}} \lambda_{sd}^{n_{sd}}$$



• Note that t_d doesn't enter the expression.

Parameter Estimates

- The model parameters can be easily found using maximum likelihood.
- It turns out (check) that

$$\hat{\lambda}_{gh} = rac{n_{gh}}{t_g}$$

for any pair of states g and h.

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• Furthermore,

$$SE(\hat{\lambda}_{gh}) = rac{\sqrt{n_{gh}}}{t_g}.$$

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Confidence Intervals

• Hence, approximate 95% confidence intervals for the intensities are

$$\lambda_{hi} \in (0.022, 0.032)$$

$$\lambda_{ih} \in (0.071, 0.109)$$

$$\lambda_{is} \in (0.028, 0.052)$$

$$\lambda_{id} \in (0.002, 0.012)$$

$$\lambda_{si} \in (0.022, 0.058)$$

$$\lambda_{sd} \in (0.008, 0.032)$$

Estimates

• Hence, the MLEs are

$$\hat{\lambda}_{hi} = \frac{n_{hi}}{t_h} = \frac{110}{4100}$$

$$\hat{\lambda}_{ih} = \frac{n_{ih}}{t_i} = \frac{90}{1000}$$

$$\hat{\lambda}_{is} = \frac{n_{is}}{t_i} = \frac{40}{1000}$$

$$\hat{\lambda}_{id} = \frac{n_{id}}{t_i} = \frac{7}{1000}$$

$$\hat{\lambda}_{si} = \frac{n_{si}}{t_s} = \frac{20}{500}$$

$$\hat{\lambda}_{sd} = \frac{n_{sd}}{t_s} = \frac{10}{500}$$

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STAT40810 — Stochastic Models

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Week 10

Finite Mixture Models

Gender Data

 We have data from a study in Saxony, Germany, which seeks to identify the number of male children in 53680 families of size 8.

×	frequency
0	215
1	1485
2	5331
3	10649
4	14959
5	11929
6	6678
7	2092
8	342

- We could model the number of males as binomial with probability of a male, p, and number of trials n = 8.
- This yields, $\hat{p} = 0.515$, as a maximum likelihood estimate.

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Alzheimer Dataset

- Data were collected on early onset Alzheimer patient symptoms in St. James' Hospital, Dublin.
- Two hundred and forty patients had six behavioural and psychological symptoms (Hallucination, Activity, Aggression, Agitation, Diurnal and Affective) recorded.
- It is believed that the patients cannot be modeled using a single model because there are different subtypes within the disease.
- The number of distinct groups of patients gives an idea of the number of subclasses or syndromes.
- It is believed that two or three groups are more suitable to describe data.

Gender Data

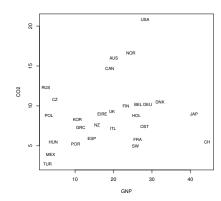
• We can see that the model fit is inadequate.

×	Frequency	Expected		
0	215.00	165.22		
1	1485.00	1401.69		
2	5331.00	5202.65		
3	10649.00	11034.65 14627.60		
4	14959.00			
5	11929.00	12409.87		
6	6678.00	6580.24		
7	2092.00	1993.78		
8	342.00	264.30		

- Also, the sample variance is bigger than expected.
- So, we need an alternative model.

CO₂ Emmissions

• Data were collected on CO₂ emmissions and per capita GNP for a number of countries.



• A close inspection of the scatter plot suggests that there could be two different linear relationships at play here.

Model-Based Clustering/Mixture Models

- Suppose we have data $x = (x_1, x_2, \dots, x_n)$ sampled from some population.
- We want to build a model that can account for substructure in the population.
- A finite mixture model has the following structure.
 - ullet Assume there are G groups (classes, components, subpopulations, clusters).
 - The probability of an observation coming from group g is τ_g .
 - Each observation within group is modeled using a standard statistical model $p(x|\theta_g)$
- This gives,

$$p(x_n|\tau,\theta,G) = \sum_{g=1}^{G} \tau_g p(x_n|\theta_g).$$

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Model-Based Clustering Approach

...when clustering samples from a population, no cluster method is, a priori believable without a statistical model.

(Aitkin et al, 1981)

...With the underlying probability model, the problems of determining the number of clusters and of choosing an appropriate clustering method become statistical model choice problems.

(Yeung, et al, 2001)

Mixture Models and Clustering

- Finite mixture models can be seen as a way of developing clustering methods which are based on statistical models.
- This is why many mixture model papers refer to *model-based* clustering.
- However, finite mixture models can also be used as a way of extending the flexibility of a standard model, without focussing on clustering.

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Gender Data

- Let's think about the gender data.
- Suppose we use a finite mixture model (with *G* components) for modeling this data.
- Each component will have probability τ_g and data within each component will be modeled as binomial with probability p_g .
- An interpretation of this is that:
 - ullet We have G different family types.
 - Each family type has its own propensity to have male children.
 - Each family type occurs with some probability.

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Gender Data: Code

• Let's fit some mixture models to the data and see what happens.

```
# Saxony data
n <- 0:8
f <- c(215,1485,5331,10649,14959,11929,6678,2092,342)
x <- rep(n,f)
y<-8-x
dat<-cbind(x,y)
colnames(dat)<-c("M","F")

# Load relevant package (flexmix)
library(mixtools)

# Fit a G component mixture model
G<-1
fit<-multmixEM(dat,k=G)

# Examine the fit
summary(fit)
```

Gender Data: Code

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```

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Gender Data: Model Fit (1)

• The model fit for G = 1 is:

```
summary of multmixEM object:
comp 1
lambda 1.000000
theta1 0.514677
theta2 0.485323
loglik at estimate: -95587.84
```

Gender Data: Model Fit (2)

• The model fit for G = 2 is:

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Gender Data: Model Fit (3)

• The model fit for G = 3 is:

Latent Class Analysis

- Latent Class Analysis (LCA) is a model for clustering categorical data.
- Let $x_n = (x_{n1}, x_{n2}, \dots, x_{nM})$ where x_{nm} takes a value from $\{1, 2, \dots, C_m\}$.
- In LCA we assume that there is local independence between variables, so that if we knew x_n was in class g we could write it's probability as

$$p(x_n|\theta_g) = \prod_{m=1}^M p(x_{nm}|\theta_{gm}) = \prod_{m=1}^M \prod_{c=1}^{C_m} \theta_{gmc}^{\mathcal{I}(x_{nm}=c)},$$

where $\{\theta_{gm1}, \dots, \theta_{gmC_m}\}$ give the probabilities of observing the categories $\{1, \dots, C_m\}$ in variable m.

ullet The $heta_g$ values will characterize and embody the differences between groups.

Model Choice

- The model with the highest BIC is the G = 2 model.
- It has been shown that BIC is consistent when choosing *G* in finite mixture models.

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- AIC is not consistent and it tends to overestimate G.
- What does this mean?
- Suppose we have data from a G component mixture model.
- The estimate \hat{G} is consistent if $\hat{G} \to G$ as $n \to \infty$.

Alzheimer Code

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```
#Load the BayesLCA package
library(BayesLCA)

# Load the data
data(Alzheimer)

#Fit the G=2 model
fit2 <- blca.em(Alzheimer, 2)
fit2

#Fit the G=3 model
fit3<-- blca.em(Alzheimer, 3, restarts=25)
fit3
```

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Alzheimer Results

Item Probabilities:

```
Hallucination Activity Aggression Agitation Diurnal Affective Group 1 0.069 0.540 0.108 0.126 0.140 0.598 Group 2 0.093 0.811 0.396 0.669 0.381 0.970
```

Membership Probabilities:

```
Group 1 Group 2
0.58 0.42
```

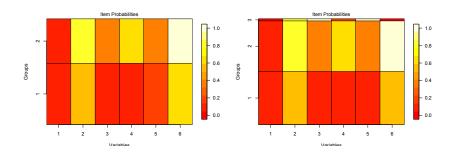
Item Probabilities:

		Hallucination	Activity	Aggression	Agitation	Diurnal	Affective
Group	1	0.062	0.518	0.063	0.132	0.096	0.549
Group	2	0.100	0.790	0.372	0.594	0.364	1.000
Group	3	0.000	0.821	0.998	0.208	1.000	0.000

Membership Probabilities:

```
Group 1 Group 2 Group 3
0.502 0.479 0.020
```

Alzhemier Results (2)



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CO₂ Code: flexmix

```
# Load the flexmix package
library(flexmix)
# Load the CO 2 data
data(CO2data)
\mbox{\tt\#} Fit a mixture of experts model with 50 random starting values for the EM algorithm.
# The highest BIC value is stored as bicval and the best fitting model as bestfit
bicval <- Inf
itermax <- 50
for (iter in 1:itermax)
  fit<-flexmix(CO2~GNP,data=CO2data,k=2)
   if (bicval>BIC(fit))
   bicval<-BIC(fit)
    bestfit<-fit
    print(c(iter,bicval))
# Explore the fitted model
summary(bestfit)
parameters(bestfit)
```

CO₂ Output: flexmix

```
Call:
flexmix(formula = CO2 ~ GNP, data = CO2data, k = 2)
```

prior size post>0 ratio Comp.1 0.244 6 10 0.600 Comp.2 0.756 22 27 0.815

'log Lik.' -66.98375 (df=7) AIC: 147.9675 BIC: 157.2929

 Comp.1
 Comp.2

 coef.(Intercept)
 1.4047073
 8.65077326

 coef.GNP
 0.6768985
 -0.02224341

 sigma
 0.8456090
 2.13942674

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CO₂ Code: mixtools

```
set.seed(1)
# Load the CO_2 data
library(flexmix)
data(CO2data)
# Load the mixtools package
library(mixtools)
# Fit a mixture of experts model with 50 random starting values for the EM algorithm. # The highest BIC value is stored as bicval and the best fitting model as bestfit
bicval <- -Inf
itermax <- 50
for (iter in 1:itermax)
  G<-2
  fit<-regmixEM(CO2data$CO2,CO2data$GNP,k=G)
  n<-nrow(CO2data)
  p<-nrow(fit$beta)*G+G+(G-1)
  fitbic <- 2*fit$loglik - log(n)*p
   if (bicval<fitbic)
    bicval<-fitbic
    bestfit<-fit
    print(c(iter,bicval))
# Explore the fitted model
summary(bestfit)
plot(bestfit, which=2)
```

CO₂ Output: mixtools

 summary
 of regmixEM object:

 comp 1
 comp 2

 lambda
 0.7549234
 0.245077

 sigma
 2.0493214
 0.809387

 beta1
 8.6789541
 1.415133

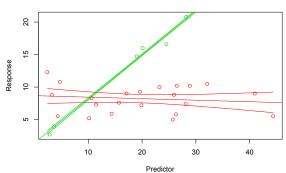
 beta2
 -0.0233429
 0.676597

 loglik
 at estimate:
 -66.93977

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CO₂ Output: mixtools

Most Probable Component Membership



CO₂ Output: mixtools

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With country names we get:

