

3) Uma distribuição de Pareto
 X_0 e α desconhecido

$$f(x|X_0, \alpha) = \frac{\alpha X_0^\alpha}{x^{\alpha+1}} \mathbb{1}\{x \geq x_m\} \begin{cases} 1, & x \geq x_m \\ 0, & x < x_m \end{cases} \quad \begin{aligned} T_1 &= \min\{X_1, \dots, X_n\} \\ T_2 &= \prod_{i=1}^m X_i \end{aligned}$$

$$f(\bar{x}|X_0, \alpha) = \left(\prod_{i=1}^m x_i \right)^{-(\alpha+1)} \cdot \alpha X_0^\alpha$$

$$= t_2^{-(\alpha+1)} \cdot \alpha X_0^\alpha$$

$$v \left[\begin{matrix} T_1 & T_2 \\ r_1(x) & r_2(x) \end{matrix}; X_0, \alpha \right] = t_2^{-(\alpha+1)} \alpha X_0^\alpha g(x_0, t_1)$$

$$u(x) = 1$$

$$\begin{cases} 1, & x_0 \leq t_1 \\ 0, & x_0 > t_1 \end{cases}$$

8) X_1, X_2, \dots, X_n formam uma amostra de uma distrib
 exponencial β desconhecido.

EMV é uma estatística suficiente

$$\begin{aligned} f(x|\theta) &= \theta e^{-\theta x} \\ L(\theta) &= f_m(x|\theta) = \theta^m e^{-\theta y} \end{aligned}$$

$$\log L(\theta) = m \log \theta - \theta y$$

$$(\log L(\theta))' = 0$$

$$\frac{m}{\theta} - y = 0 \Leftrightarrow \theta = \frac{m}{y} = \frac{1}{\bar{x}_m}$$

$$T = 1/\bar{x}_m \quad f(x|\beta) = \beta e^{-\beta x} \rightarrow f_m(x|\beta) = \beta^m e^{-\beta \sum x_i}$$

$$f_m(x|\beta) = \beta^m e^{-\beta m/T}, \quad v(x(x), \beta) \sim u(x) = 1$$

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 X_1, X_2, \dots, X_m , com pdf: θ desconhecido

$$f(x|\theta) = \begin{cases} 2x/\theta^2, & 0 \leq x \leq \theta \\ 0, & \text{c.c.} \end{cases}$$

MLE da mediana é mínimo suficiente

$$l_m(x|\theta) = \prod_{i=1}^m \frac{2x_i}{\theta^2} \mathbb{1}\{x \in [0, \theta]\}$$

$$= 2^m \prod_{i=1}^m x_i \mathbb{1}\{x_1, \dots, x_m \in [0, \theta]\}$$

$$\theta^{2m} \rightarrow \hat{\theta} = \max\{x_1, \dots, x_m\}$$

$$\int_{-\infty}^m f(x|\theta) dx, \quad x_i > 0 \rightarrow \int_0^m \frac{2x}{\theta^2} dx = \frac{1}{2}$$

$$\frac{x^2}{\theta^2} \Big|_0^m = \frac{1}{2} \Leftrightarrow \frac{m^2}{\theta^2} = \frac{1}{2}$$

$$2m^2 = \theta^2$$

$$|m| = \frac{1}{\sqrt{2}} \theta$$

$$\hat{m} = \frac{1}{\sqrt{2}} \theta$$

$$l_m(x|\theta) = \underbrace{2^m \left(\prod_{i=1}^m x_i \right)}_{u(x)} \cdot \frac{1}{\theta^{2m}} \mathbb{1}\{\hat{\theta}_{MLE} < \theta\}$$

$$v(\hat{\theta}_{MLE}, \theta)$$



(16) X_1, X_2, \dots, X_m com distrib de Poisson média = λ
posteriori desconhecida
♥ $\alpha_1 = \alpha + \sum X_i$ $\beta_1 = \beta + m$

$$\frac{\theta^x e^{-\theta}}{x!}$$

$$E[\lambda | X] = \frac{\alpha + \sum X_i}{\beta + m} \quad \xrightarrow{T}$$

$$f(x|\lambda) = \frac{\lambda^x e^{-\lambda}}{x!} \rightarrow f_m(\bar{x}|\lambda) = \frac{\lambda^{\sum x_i} e^{-m\lambda}}{\prod_{i=1}^m x_i!} \quad \pi\{x_1, x_2, \dots, x_m > 0\}$$

$$= \underbrace{\frac{1}{\prod_{i=1}^m x_i!}}_{u(x)} \cdot \underbrace{\lambda^{\sum x_i} e^{-m\lambda}}_{v(T, \lambda)} \quad \text{é suficiente}$$

O estimador de Bayes (função injetiva de T) também é suficiente. ♥