

$$E[y_i] = \mu_i(\beta) = \tilde{x}_i^T \beta \quad \begin{array}{l} \tilde{x}_i: i\text{-ésima linha de } \tilde{X} \\ \beta: p+1 \times 1 \end{array}$$

1. Escreva a log-verossimilhança e deduza seu gradiente e a sua derivada segunda (hessiana);

$$f(y_i | \beta, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2} (\tilde{x}_i \beta - y_i)^2\right)$$

$$f_X(Y | \beta, \sigma^2) = \prod \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2} (\tilde{x}_i \beta - y_i)^2\right)$$

$$L(\beta, \sigma^2 | Y) = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (\tilde{x}_i \beta - y_i)^2\right) = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left(-\frac{1}{2\sigma^2} \|\tilde{X}\beta - Y\|^2\right)$$

$$\log(f_X(Y | \beta, \sigma^2)) = -\frac{\|\tilde{X}\beta - Y\|^2}{2\sigma^2} - \frac{n}{2} \log(2\pi\sigma^2) \quad \xrightarrow{\quad} = (\tilde{X}\beta - Y)^T (\tilde{X}\beta - Y)$$

Gradiente:

$$\frac{\partial L}{\partial \beta} = -2 \tilde{X}^T \cdot \frac{(\tilde{X}\beta - Y)}{2\sigma^2} = \frac{-\tilde{X}^T (\tilde{X}\beta - Y)}{\sigma^2}$$

Hessiana:

$$H = \frac{\partial^2 L}{\partial \beta^2} = \frac{-\tilde{X}^T X}{\sigma^2}$$

2. Com base nos cálculos do item anterior, mostre a forma do estimador de máxima verossimilhança para β , $\hat{\beta}$;

$$\frac{\partial L}{\partial \beta} = 0$$

$$\frac{-\tilde{X}^T (\tilde{X}\beta - Y)}{\sigma^2} = 0$$

$$-\tilde{X}^T \tilde{X} \beta + \tilde{X}^T Y = 0$$

$$\tilde{X}^T X \beta = \tilde{X}^T Y$$

$$\hat{\beta} = (\tilde{X}^T X)^{-1} \tilde{X}^T Y$$

3. Mostre que $\hat{\beta}$ é não-viesado;

$$\text{queremos } E[\hat{\beta}] = \beta, \quad E[Y] = E[\tilde{X}\beta + \epsilon] = \tilde{X}\beta$$

$$E[(\tilde{X}^T \tilde{X})^{-1} \tilde{X}^T Y] = \underbrace{(\tilde{X}^T \tilde{X})^{-1} \tilde{X}^T \tilde{X}}_I \beta = \beta$$

4. Considere um outro estimador não-viesado de β : $\tilde{\beta} = M y$, onde

$$M = (X^T X)^{-1} X^T + D,$$

e D é uma matriz $P \times n$ cujas entradas são não-zero.

Mostre que $R := \text{Var}(\tilde{\beta}) - \text{Var}(\hat{\beta})$ é positiva-definida.

Dica: Compute $E[\tilde{\beta}]$ e considere o que deve valer para D sob a premissa de que $\tilde{\beta}$ é não-viesado.

$$\tilde{\beta} \text{ não-viesado: } E[\tilde{\beta}] = \beta, \quad E[y] = E[\tilde{X}\beta + \epsilon] = \tilde{X}\beta$$

$$E\left[\left((X^T X)^{-1} X^T + D\right) y\right] = E\left[\left((X^T X)^{-1} X^T y + D y\right)\right] = \beta$$

$$\underbrace{(X^T X)^{-1} X^T}_{I} \tilde{X} \beta + D \tilde{X} \beta = \beta$$

$$\beta + D \tilde{X} \beta = \beta$$

$$D \tilde{X} \beta = 0$$

$$D \tilde{X} = 0$$

$$\text{Var}(Ay) = A \text{Var}(y) A^T$$

$$\begin{aligned} \text{Var}(\tilde{\beta}) &= \text{Var}\left(\left((X^T X)^{-1} X^T + D\right) y\right) = \left((X^T X)^{-1} X^T + D\right) \underbrace{\text{Var}(y)}_{\sigma^2} \left((X^T X)^{-1} X^T + D\right)^T \\ &= \sigma^2 \left((X^T X)^{-1} X^T + D\right) \left((X^T X)^{-1} X^T + D\right)^T \\ &= \sigma^2 \left((X^T X)^{-1} X^T + D\right) \left(\left((X^T X)^{-1} X^T\right)^T + D^T\right) \\ &= \sigma^2 \left((X^T X)^{-1} X^T + D\right) \left(X (X^T X)^{-1} + D^T\right) \\ &= \sigma^2 \left(\underbrace{(X^T X)^{-1} X^T X (X^T X)^{-1}}_I + \cancel{(X^T X)^{-1} X^T D^T} + \cancel{D X (X^T X)^{-1}} + D D^T \right) \\ &= \sigma^2 (X^T X)^{-1} + \sigma^2 (D D^T) \end{aligned}$$

$$\begin{aligned} \text{Var}(\hat{\beta}) &= \text{Var}\left((\tilde{X}^T \tilde{X})^{-1} \tilde{X}^T y\right) = \left((\tilde{X}^T \tilde{X})^{-1} \tilde{X}^T\right) \text{Var}(y) \left((\tilde{X}^T \tilde{X})^{-1} \tilde{X}^T\right)^T \\ &= \sigma^2 \underbrace{\left((\tilde{X}^T \tilde{X})^{-1} \tilde{X}^T\right) \tilde{X} (\tilde{X}^T \tilde{X})^{-1}}_I \\ &= \sigma^2 (X^T \tilde{X})^{-1} \end{aligned}$$

$$\begin{aligned} R = \text{Var}(\tilde{\beta}) - \text{Var}(\hat{\beta}) &= \cancel{\sigma^2 (X^T X)^{-1}} + \sigma^2 (D D^T) - \cancel{\sigma^2 (X^T \tilde{X})^{-1}} \\ &= \underbrace{\sigma^2 (D D^T)}_{\text{positiva definida}} \quad X X^T \text{ positiva definida} \\ &\quad \text{positiva definida} \end{aligned}$$