


BMJ Open Revisiting systematic geographical variations in tonsils surgery in children in the Spanish National Health System: spatiotemporal ecological study on hospital administrative data

Eusebio Castaño-Riera,¹ Manuel Ridao,^{2,3} Julian Librero,^{3,4} Natalia Martínez-Lizaga,^{2,3} Micaela Comendeiro-Maaløe,^{2,3} Ester Angulo-Pueyo,^{2,3} Salvador Peiro,^{3,5} Enrique Bernal-Delgado ^{2,3}

To cite: Castaño-Riera E, Ridao M, Librero J, *et al*. Revisiting systematic geographical variations in tonsils surgery in children in the Spanish National Health System: spatiotemporal ecological study on hospital administrative data. *BMJ Open* 2022;**12**:e064009. doi:10.1136/bmjopen-2022-064009

► Prepublication history and additional supplemental material for this paper are available online. To view these files, please visit the journal online (<http://dx.doi.org/10.1136/bmjopen-2022-064009>).

Received 21 April 2022
Accepted 09 November 2022



© Author(s) (or their employer(s)) 2022. Re-use permitted under CC BY-NC. No commercial re-use. See rights and permissions. Published by BMJ.

For numbered affiliations see end of article.

Correspondence to

Dr Enrique Bernal-Delgado; ebernal.ias@aragon.es

ABSTRACT

Objective To provide new evidence on how tonsils surgery in children has geographically varied over time in the context of the Spanish National Health System.

Design Observational ecological spatiotemporal study on geographical variations in medical practice, using linked administrative datasets, including virtually all surgeries performed from 2003 to 2015.

Setting The Spanish National Health System, a quasi-federal structure with 17 autonomous communities (ACs), and 203 healthcare areas (HCAs).

Participants Patients aged 19 and younger residing in the HCAs and ACs.

Interventions Tonsillectomy with adenoidectomy (T&A); and tonsillectomies alone (T).

Main endpoints (1) Evolution of T&A and T rates; (2) spatiotemporal variation in the risk of receiving T&A or T surgery at regional level (ACs) and HCAs; and (3) the fraction of the variation (FV) attributed to each of the components of variation—ACs, HCAs, year and interaction ACs year.

Results T&A age-sex standardised rates increased over the period of analysis from 15.2 to 20.9 (5.7 points per 10 000 inhabitants). T alone remained relatively lower than T&A rates, evolving from 3.6 in 2003 to 3.9 in 2015 (0.3 points per 10 000 inhabitants). Most of the risk variation was captured at the HCAs level in both procedures (FV: 55.3% in T&A and 72.5% in T). The ACs level explained 27.6% of the FV in the risk in T&A versus 8% in T. The interaction ACs year was similar in both procedures (FV: 15.5% in T&A and 17.5% in T). The average trend hardly explained 1.46% and 1.83% of the variation, respectively.

Conclusion Our study showed wide persistent variations with a steady increase in rates and risk of T&A and a stagnation of T alone, where most of the variation risk was explained at HCA level.

INTRODUCTION

'In some of the early years of the School Medical Service it had seemed necessary to recommend further provision for the

STRENGTHS AND LIMITATIONS OF THIS STUDY

- ⇒ This study covers virtually all the surgery provided to the population aged 0–19 in public settings in the whole country along years of observation; the cases have been effectively allocated to the place of residence so that the risks are unambiguously attributed to healthcare areas (HCAs), autonomous communities (ACs) and year.
- ⇒ A robust model spatiotemporal was used to describe the variation on different levels (ACs—locus of health policy decisions), HCAs (locus of provision) and over a long period of time (13 years).
- ⇒ A limitation of the study was the non-inclusion of information on surgical interventions in private hospitals with no public financing, although private provision in Spain is limited.
- ⇒ Variation in the risks of getting a tonsils surgery could be explained by factors both, at HCAs and ACs level, but also on how unevenly ACs have adopted an evidence-based approach to tackle with unexpectedly high rates of tonsils surgery.

operative treatment of tonsils, but the tide rose so fast that in his Annual Report for 1923, Sir George Newman issued the first of his many warnings against premature resort to operation'. This statement is part of the contextual arguments given to justify a dedicated session of the Section of Epidemiology and State Medicine at the Royal Society of Medicine on 28 May 1938. Sir Alison Glover, acting as speaker, provided evidence on differences in the rates of tonsillectomies that could not be warranted by differences in epidemiology.¹ These proceedings, fairly deemed the first formal approach to the study of unwarranted variations in medical practice, paved the way for other works on the use of tonsillectomies in the paediatric population.^{2 3}

Numerous studies have shown wide geographical variations in tonsillectomy rates.^{4–10} As in Glover's work, these large variations are suggested to be a proof of differences in care intensity and therapeutic choices across territories rather than differences in need. In Spain, a 2006 study based on 2002 data showed extreme variation in tonsillectomy rates, as large as 13-fold difference across 147 healthcare areas (HCAs), very unlikely to be explained by differences in children's needs across HCAs.¹¹

Since the 2006 study, no recent information on unwarranted variations in tonsillectomy rates in the Spanish National Health System has been released or published while guidelines,^{12–18} systematic reviews,^{19–23} randomised clinical trials^{24–28} and a cohort study²⁹ on the indications for tonsillectomy and adenoidectomy multiplied. In the Spanish context, the societies of ENL and Paediatrics provided clear recommendations on the indications of tonsillectomy.³⁰ There is a need to understand if this wealth of new evidence has translated into a change in the prescription patterns of tonsils surgery in children and consequently, into a change in the variation of population rates.

This paper provides new evidence on how tonsils surgery has geographically varied over time in the context of the Spanish National Health System.

METHODS

Design

We conducted an observational spatiotemporal ecological study on geographical variations in medical practice. Epistemically, unlike utilisation studies at patient level or studies on practice appropriateness, studies on geographical variations in medical practice aim at analysing and interpreting whether similar populations are differently exposed to health services because of the place of residence, traditionally measured as systematic variations in age-sex standardised utilisation rates or age-sex standardised utilisation ratios.

As in small area analyses and disease mapping in epidemiology, the methods seek differentiating what part of the variation is on the features of the population (typically, age and sex as proxies of burden of disease) and what part is in other factors, so named supply-side factors (institutional features, available resources, incentives, etc), once chance as well as information biases are discarded as potential explanations.

In this study, we followed two complementary approaches; the first one, focused on the evolution of surgical rates of tonsillectomy with adenoidectomy (T&A) and tonsillectomy alone (T); and, the second one, meant to elicit spatiotemporal variation in the risk of receiving tonsils surgery. In both cases, the population exposure was analysed at the HCAs and Region (namely, autonomous communities (ACs)) of residence, over time.

Population

The population aged 19 and younger residing in all the 203 HCAs (the locus where hospital and primary care services are provided) of the 17 ACs (the locus where financing and policy-making decisions are taken) composing the Spanish National Health System was included.

Virtually all tonsillectomies with or without adenoidectomy performed on this population between 2003 and 2015 in all public hospitals in the country were included in the study. Tonsillectomy and/or adenoidectomy were defined as per ICD9-MC taxonomy. Codes 28.2 (tonsillectomies without adenoidectomy) and 28.3 (tonsillectomies with adenoidectomy) were considered for the analysis. All the tonsil surgeries identified were effectively allocated to the HCA and AC of residence.

Setting

The Spanish National Health System, composed of 17 AC health systems, covers virtually all the population residing in Spain for a wide and deep basket of benefits. Hospital care, outpatient specialised care, primary care and emergency are fully covered for free at the point of care, with no patients' copayments. As a strongly regulated system, all the population is mandated to register to a primary care doctor (general medicine or paediatrics), and is allocated to an HCA where residents are served for specialised care. The private sector in Spain is essentially subsidiary; in 2015, the last year of our study, represented less than 5% of the total expenditure in health,³¹ with 85% of the hospital and outpatient specialised care provided in public settings.³²

A usual pathway of care in the Spanish system would start from a visit to the primary care doctor who will decide on whether there is need for the patient to be transferred out to specialised care. In the case of tonsils surgery, the usual path for children with a tonsil condition would be as follows. The child will first visit the paediatrician at primary care level. Then, the paediatrician will decide whether the case requires transfer out to the ENL specialist. Usually, those ENL doctors specialised in children will provide the services to children transferred for all the HCAs in the region, and will decide on the treatment, whether watchful waiting, T&A or T alone. In the case of surgery, the intervention will be usually performed as a day-case care in dedicated hospital facilities or outpatient specialised care premises.

Main endpoints

Three endpoints were considered: (1) evolution of T&A and T standardised rates; (2) spatiotemporal variation in the risk of receiving T&A or T surgery at regional (ACs) and HCAs levels; and (3) the fraction of the variation attributed to each of the components of variation; so, ACs, HCAs, the year when the surgery was provided, and the interaction between the ACs and the year when the surgery was performed.

Sources of data

The study is based on administrative data from the so named, Minimum Basic Hospital Discharge Dataset (CMBD in Spanish), that collects information from both in-hospital surgeries and day-case surgeries. Specifically, this data source provides administrative data (place of residence, provider where the population is allocated, and hospital of treatment); clinical information (main diagnosis, secondary diagnoses and procedures performed during the stay); and demographic information (age and sex) on all hospital discharges in the Spanish National Health System (public hospitals and private hospitals with public funding). The completeness of this data source is guaranteed since ACs have the legal obligation of duly reporting to the Ministry of Health, in a yearly basis, on the hospital activity (see the corresponding Royal Decree at <https://www.boe.es/eli/es/rd/2015/02/06/69/com>).

In addition, the denominator for the population rates was obtained from the annual census of the Spanish National Institute of Statistics. Atlas VPM allocates each of the episodes of treatment to the place of residence, then to the HCA and region of residence, both the units of analysis in this study.³³ Notably, as part of its data quality assurance methodology, Atlas VPM conducts regular (once a year) quality checks to, for example, avoid over time coding variations, correct semantic and syntactic inconsistencies in the variables, or reallocate admission episodes to the place of residence if there are changes in geolocation of administrative areas or hospital providers.

Analysis

All 2003 to 2015 T&A and T interventions of patients aged 19 and younger were selected, linked to their area of residence and aggregated into four age-sex specific subgroups (0–4, 5–9, 10–14, 15–19). For each year in the series, age-standardised and sex-standardised rates per 10000 inhabitants were calculated. As a first approach to variation the extremal quotient between standardised rates (95th to 5th percentiles ratio) and the interquartile interval (75th to 25th percentiles ratio) were calculated. Then, the systematic component of variation (SCV) provided the part of the variation that was not random (SCV assumes the expected cases for each HCA follow a Poisson distribution).

To understand the systematic differences in the risk of receiving T&A and T, risk ratios at HCAs and ACs level were estimated; thus, the variation across HCAs and ACs was decomposed in four components: spatial at HCA level, spatial at AC level, spatiotemporal at AC level, and overall temporal effect. Risks estimation and estimates for the fraction of variance followed a two-step approach: (1) under the hypothesis that risk remained constant across areas over time, the expected number of cases per HCA was estimated applying the overall rate over the period of study to the population at risk in each HCAs in the respective year. Standardised HCAs ratios (RR) were estimated using the ratio of observed-to-expected cases, interpretable as the maximum estimate

of the risk ratio of admissions for T or T&A in each area in each period. (2) A Bayesian hierarchical approach was then used to model the different risk structures of dependence in both space and time.^{34–37} In the *first level* of this hierarchical modelling, we assumed that, conditional on the underlying relative risk, the number of counts y_{jt} in the j th area at the t th time period followed a Poisson distribution with mean $u_{jt} = e_{jt} r_{jt}$, where e_{jt} is the number of expected counts and r_{jt} the unknown relative risk. In the *second level*, the $\log(r_{jt})$ was expressed as the sum of the components representing the individual and independent contributions to the risk in a specific HCAs and period [$\log(r_{jt}) = \text{intercept} + s_i + T_j + ST_{ij}$], where the intercept term gives the initial level of risk that is shared by all regions and periods. The main effects S_i and T_j represented the additional risk of living in ACs i and period j and the second order interaction term ST_{ij} represented the risk contribution due to a combination of the effects that cannot be explained additively by the main effects. In the *third level*, a hyperparameter-prior distribution was assumed where the spatial effect was modelled following a convolution CAR prior.³⁴ The temporal main effect was a combination of a time-unstructured (exchangeable) and a time-structured effects (first-order random walk), and the interaction term, as the independent effect of unobserved covariates for each combination of ACs and period (i, j), thus without any structure.³⁵

The use of a Bayesian approach in this study lays on a number of arguments. First, unlike in traditional studies on geographical variations, this study cannot assume that the exposure to tonsils surgery of the population residing in one area is independent of the exposure in neighboured HCAs; on the contrary, children from different HCAs get tonsils surgery at a single tertiary care setting that serves them all. On the other hand, traditional methods provide ecological associations (and interpretations) based on observed variables while latent unobserved factors that occur within the units of analysis are simply omitted; this Bayesian model allows inference on those latent unobserved variables that occur in units of analysis relevant to the study; so, we are able to know what part of the variation is in the regions (so, common to all areas), what part of the variation is in the HCA, and within the HCA, what part is structured (dependent on the interaction with other areas) or what is not (basically attributed to phenomena occurring exclusively within the area); and, what part is in the time trends (in the utilisation of certain procedures) and whether trends differ across regions. Finally, this Bayesian approach has been observed to be more robust in multiple testing conditions and more precise in the case of high heterogeneity across HCAs.

Patient and public involvement

Patients or the public were not involved in the setting of the research question, nor in developing plans for design/writing of our draft. Patients or the public were

Table 1 Cases, rates and variation in tonsils surgery in people aged 19 and younger (2003–2015)

	2003	2004	2005	2006	2007	2008	2009	2010	2011	2012	2013	2014	2015
Population aged 0–19*	8.41	8.46	8.57	8.68	8.78	8.96	9.11	9.19	9.23	9.24	9.21	9.14	9.12
T&A cases	12218	13084	13980	14763	16200	17060	17765	17380	17610	17696	18319	19149	18110
Crude rate	14.51	15.46	16.30	17	18.43	19.05	19.50	18.90	19.06	19.14	19.89	20.93	19.84
Stand. rate	15.24	16.24	16.87	17.36	18.71	19.19	20.04	19.92	20.21	20.01	20.51	21.54	20.96
EQ5-95	12.35	7.96	9.58	6.49	6.84	6.11	7.19	9.80	5.68	6.86	5.46	5.88	8.21
IQ25-75	2.49	2.25	2.24	2.40	2.33	2.02	2.02	1.99	1.91	1.92	2.05	1.95	2.05
SCV	0.27	0.21	0.22	0.28	0.29	0.24	0.29	0.29	0.28	0.28	0.24	0.31	0.25
EB	0.42	0.3	0.28	0.3	0.27	0.23	0.22	0.26	0.2	0.21	0.22	0.23	0.22
T alone cases	2685	2900	3278	3351	3523	3598	3772	3597	3362	3246	3398	3545	3289
Crude rate	3.19	3.43	3.82	3.86	4	4.01	4.14	3.91	3.64	3.51	3.69	3.87	3.60
Stand. rate	3.57	3.65	4.02	3.97	4.04	4.03	4.34	4.20	3.85	3.73	3.98	4.06	3.88
EQ5-95	8.01	6.86	7.23	7.16	6.13	5.22	4.92	6.23	5.26	5.16	5.92	5.88	6.48
IQ25-75	2.04	2.10	2.16	2.04	2.06	1.96	2.09	2.05	2.10	2.18	2.12	2.12	2.08
SCV	0.37	0.2	0.25	0.16	0.26	0.2	0.39	0.36	0.19	0.21	0.48	0.23	0.4
EB	0.34	0.22	0.22	0.16	0.19	0.16	0.25	0.24	0.16	0.18	0.22	0.17	0.2

*Figures are expressed in millions.
EB, empirical bayes; EQ5-95, 95th to 5th percentiles ratio; IQ25-75, 75th to 25th percentiles ratio; SCV, systematic component of variation ; T, tonsillectomy alone; T&A, tonsillectomy and adenoidectomy.

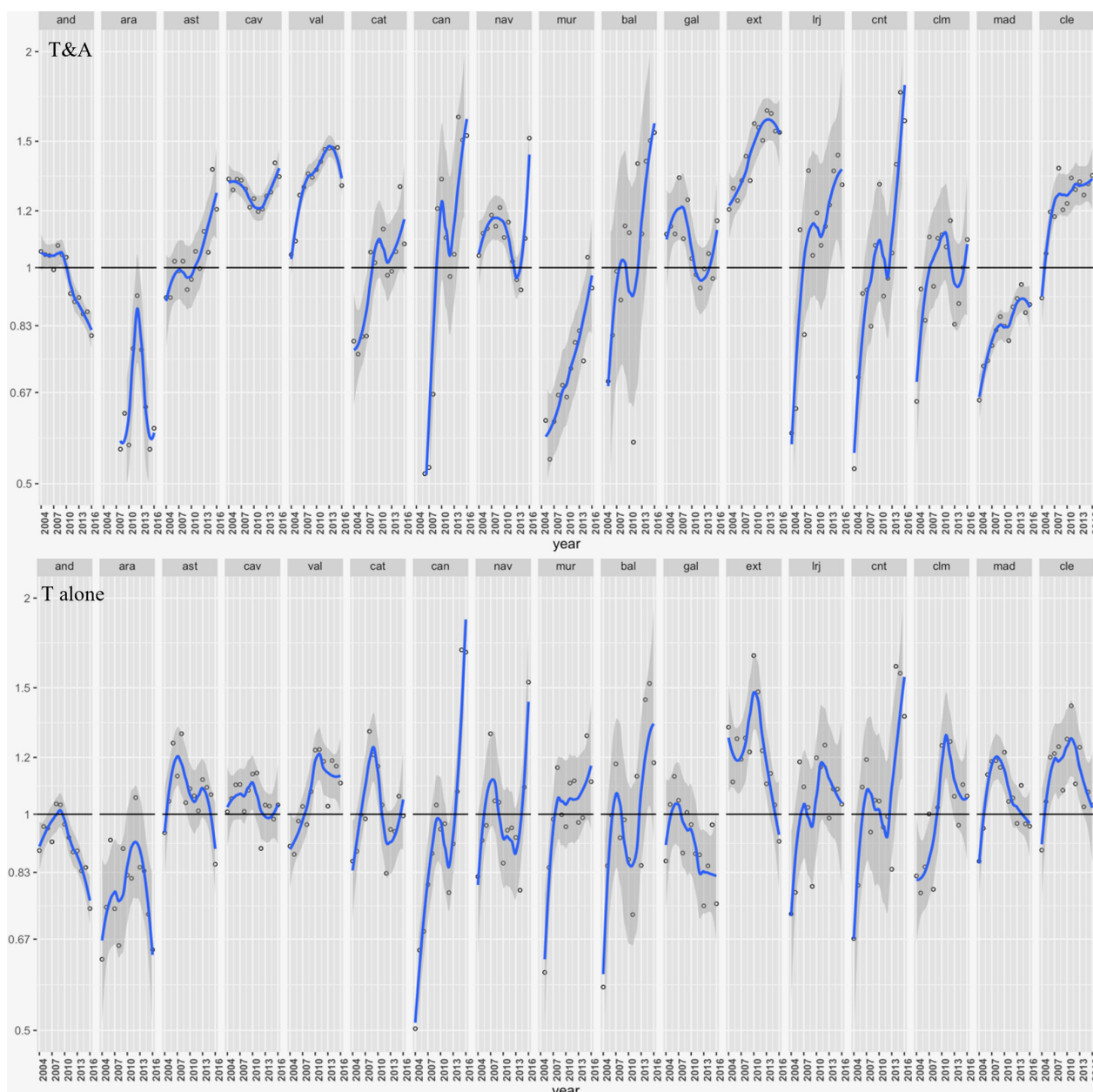


Figure 1 Surgery risk ratios across regions along the study period.

not asked to advise on the interpretation or writing up of findings.

RESULTS

While the population 19 and younger increased by 8% until 2009, from 8.4M to 9.1M lives, remaining constant until 2015, tonsillectomies with or without adenoidectomies procedures grew at a faster pace: from 12218 in 2003 to 18110 in 2015 (48% increase).

T&A age-sex standardised rates steadily increased over the period from 15.24 to 20.96 (5.42 points per 10000 inhabitants). T rates in turn, lower than T&A rates, slightly increased over the period—from 3.57 to 2003 to 3.88 in 2015 (0.31 points per 10000 inhabitants) (table 1). By age groups (tables in online supplemental

appendix 1), T&A were mainly performed in children 5 to 9 years old (46.8%) and 0 to 4 years old (35.6%), with an increase in T&A standardised rates of 8.11 and 7.63 points per 10000 inhabitants, respectively. T were more prevalent in the group between 15 and 19 (39.2%) and in children between 5 and 9 (33.5%), with an increase in T standardised rates of 1.42 and 0.59 points per 10000 inhabitants, respectively.

Variation statistics showed high variation figures; the systematic component of the variation was above 1.2 in all the period, slightly higher in the case of tonsillectomies alone (table 1). By age groups, the systematic variation in T&A was extreme; in the adolescents (15–19 years old), the SCV was between 0.68 and 1.2; in the group between 5 and 9, the SCV ranged between 0.5 and 0.63. In the case

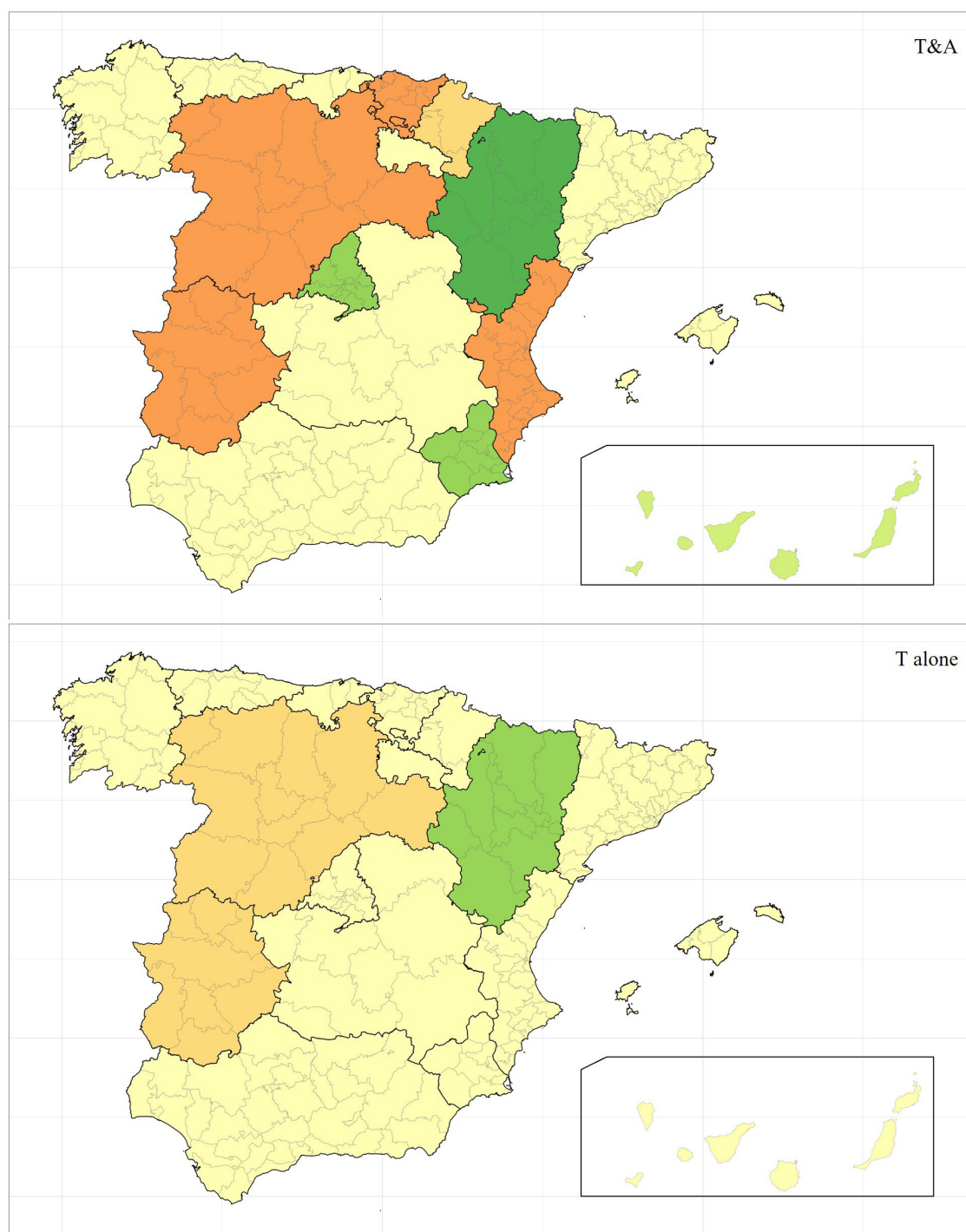


Figure 2 Spatial component surgery risk at regional level.

of tonsillectomies alone, the systematic variation in these age subgroups was even larger; in the adolescents (15–19 years old), the SCV was between 0.75 and 1, while in the group between 5 and 9, the SCV ranged between 0.85 and 1.36 (tables in online supplemental appendix 1).

Figure 1 exhibits very large differences of tonsils surgery across ACs and over the study period, with larger variation in T&A.

Figures 2–4 show the spatial and temporal components of the variation in the observed to expected cases at AC level, and as HCA and AC jointly; in addition, it is

also shown the average time evolution of the two prior models, for T&A and T alone (figure 2). As observed, T&A showed higher variation (9 out of 17 ACs showed risk different to expected) while T showed a more homogeneous pattern of risks (only 3 out of 17 ACs had risk ratios slightly over or below the expected).

Figure 3 represents the joint model for both ACs and HCAs; the images in both T&A and T alone suggest that the higher contribution in the variation would correspond to the HCA. Finally, in figure 4, the time component shows a different pattern in T&A and T alone. While

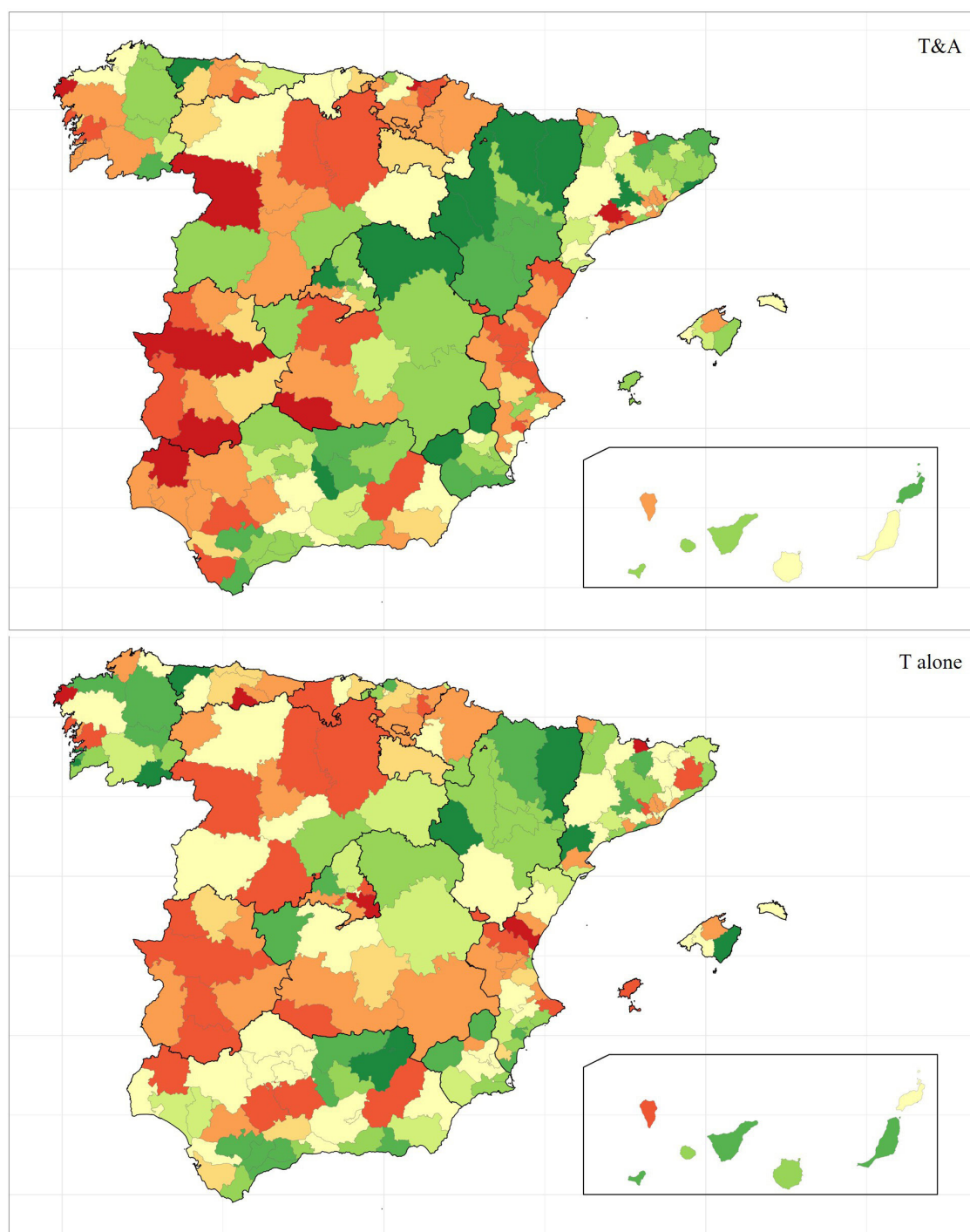


Figure 3 Spatial component surgery risk at healthcare area level.

T&A keeps growing since the beginning of the study period, T alone has remained almost flat since 2006 (risk slightly above 1 with a slight increase in the last period).

Table 2 exhibits the relative contribution of each of these components of variation (ie, variance decomposition). Consistent with the observations in figures 2–4, most of the variance in T&A was captured by the spatial phenomenon, 55.4% explained by HCAs and 27.6% by ACs; the interaction between spatial and temporal components explained 15.6% of the variance while time just explained 1.5%.

In turn, the HCAs component explained 72.6% of variance in T (figures 2–4) with an additional 8.1% of spatial component of variance attributed to ACs level. The spatiotemporal interaction explained 17.5% of the variance while the temporal component explained a small 1.8% of the variation.

DISCUSSION

Our study showed a continuous increase in the rates of T&A procedures and a stagnation in T. The variability

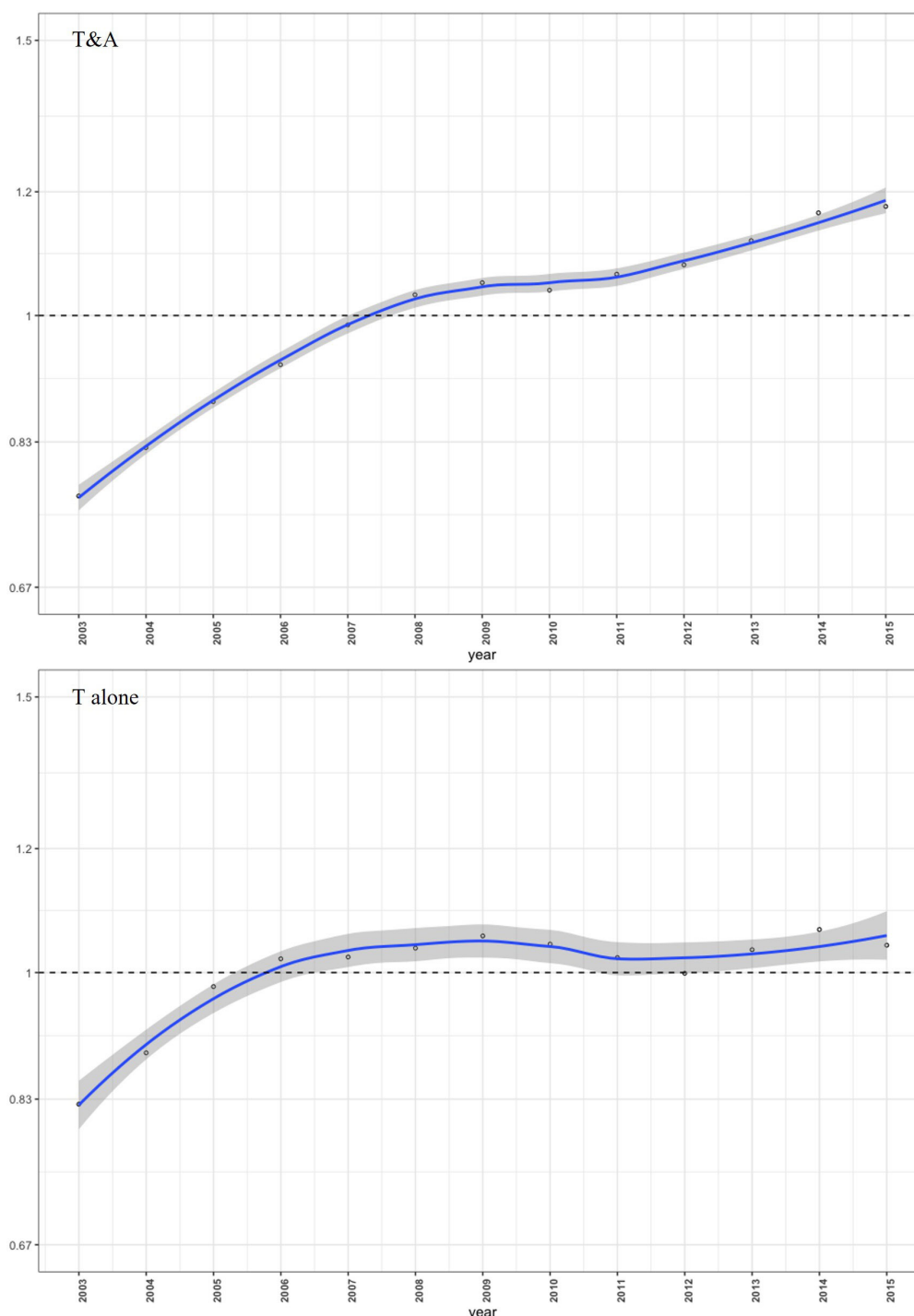


Figure 4 Time component surgery risk.

was very high over the study period in both procedures, extreme in some age groups. The spatial component of the risk of surgery was the most relevant in the explanation of the variance, with HCAs capturing most of the variation (55% in T&A and 73% in T alone). In T&A, the variation attributable to the AC of residence was substantial, explaining a 27.6% of the variance, and notably different to T. The component of variance attributable to the spatiotemporal interaction was substantially relevant in T alone, explaining a 17.5% of the variation.

Variations in tonsils surgery are ubiquitous and wide. In our study, the extremal quotient was as high as 6.5 times in the case of tonsillectomies alone, and as 8 times in the case of T&A. Other studies exhibited a sevenfold variation in tonsillectomy rates across local authority areas in England in people under 15,⁶ 4.3-fold difference in USA⁷ and 3.4-fold in local areas in Veneto.¹⁰

Likewise, our large variation in risk ratios and increasing intervention rates is also consistent with previous international and national evidence.^{4–10} Interestingly, the

Table 2 Variance decomposition to the model per each surgery

	T&A (%)	T alone (%)
Spatial HCAs	55.37	72.56
Spatial ACs	27.61	8.09
Temporal	1.46	1.83
Spatiotemporal	15.56	17.52
ACs, autonomous communities; HCAs, healthcare areas; T&A, tonsillectomy with adenoidectomy.		

behaviour is different between T&A and T. While the growing utilisation of tonsillectomies alone reached a plateau that has remained over the years, in the case of the combination with adenoidectomy has steadily grown. This relatively lower growth in T may be explained by the fact that there has been increasing evidence on the actual indications of this surgery as compared with the ambiguity of evidence regarding T&A indications.^{38–41} Further exploring our data to test this hypothesis, we developed a sensitive definition for conditions that a priori could be deemed as appropriate indications of tonsils surgery. We found that 56% of tonsillectomies alone could be considered as an evidence-based indication, while in the case of tonsillectomies with adenoidectomy this figure dropped to 26%.

As in previous research on geographical variations in elective surgery in Spain, the variance component at the HCA level has been observed more important than the one at ACs level when explaining variation.³⁷ This is consistent with the institutional design of the Spanish Health System and how decisions on provision are made; in this particular case, ENT services are based and provided at HCA level. However, this study has also found that the ACs of residence explained a substantial part of the variation when it comes to T&A (27.6% of the variation). As per the analytical approach this finding would suggest that some latent unobserved factors common to all the HCAs within an ACs would drive the decision of prescribing T&A instead of conserving the tonsils and adenoids. A potential latent variable (consistent with the observation of greater rates of A&T in children aged 9 or less) that would explain the high fraction of variation at ACs level is that ENT specialised in younger children care, unlike the rest of ENT doctors, provide *de facto* services at ACs level.

However, this explanation does not appear sufficient to explain that, for the overall period, some ACs are well below the expected while others are well above (figure 2). We may find a second potential explanation for these differences across ACs in light of the importance of the spatiotemporal component observed in our results (ie, 15.6% of the variance is explained by the interaction between the ACs and the evolution of rates over time in T&A). This interaction is coherent with the changes in the indication profile over the years, uneven across ACs. So, while in the first years of our time-series infections were

the underlying reason for surgery, currently, obstructive conditions are those which drive the decision.

Limitations

Although the strengths of this work (i.e., the study covers all the Spanish population aged 19 or younger, virtually all the surgery provided in public settings in the whole country along 13 years of observation and the cases have been effectively allocated to the place of residence so that the risks are unambiguously attributed to HCAs, ACs and year) the study is not exempt from some potential bias.

As our study builds on administrative data, the specificity of codes and the coding practices are potential sources of bias, if recording differs across the units of analysis. In the case of the Spanish National Health System coding practices are fairly homogeneous, as the Ministry of Health provides guidance since the late 1990s on when and how to code the conditions and procedures included in the International Classification of Diseases (ICD). However, in the case of tonsils surgery ICD-MC 9th, the classification used until 2015, very specific in the coding of tonsillectomy and/or adenoidectomy, does not include any specific code for tonsillotomy, leading coders to register this procedures as a tonsillectomy. Fortunately, this fact, which would have implied a major limitation in appropriateness studies, has not resulted in any bias when it comes to the estimates of variation, since all tonsillotomies were also included in the study.

As a second caveat, our analyses refer to the activity carried out in public hospitals and private hospitals with public financing. No information was included on those private hospitals with no public financing. As private hospitals in Spain do not distribute homogeneously across regions, there is a potential for bias in the estimates of variation and in the evolution of rates. Therefore, our estimates are expected to be conservative in those regions with larger provision of private services. If we were accounting for those cases performed in private hospitals in those regions, and those cases represented a substantial amount, variation will likely increase at ACs level although we cannot confirm the same for variation at HCAs level.

As a third consideration, the strength of the results relies on the large variation that is irrespective of age and sex and the robustness of the estimates over time. But, when it comes to the interpretation of the results, high (low) rates or high (low) risk ratios cannot be strictly interpreted in terms of adequacy—that is, more is necessarily worse, or less is necessarily better. For that purpose, we would have needed that the registry of the underlying conditions and the encoding system (ICD 9th) would have unequivocally flagged those cases deemed appropriate for treatment. As this was not the case, we cannot make any strong judgement about the right rate on tonsils surgery. However, as this study builds on the whole population in mid to large-size areas, we should not expect that the potential differences in the prevalence of sleep-apnoea disorders in children or severe infectious tonsillitis were as large as

seven times, a variation observed in our series. This fact provides a fair indication that tonsils surgery rates are not necessarily driven by need.

As a fourth question, it would be ideal to continue the study until the most recent available data—2019. However, we prevented ourselves from continuing because of the discontinuation of the coding system in 2016 (ICD 10th started to be implemented) which entailed uneven coverage of cases across HCAs and ACs and severe data quality issues (eg, coding of secondary diagnoses). This is a limitation for the generalisation of our results nowadays, particularly if they were to inform any policy aiming the reduction of rates or the variation. Nonetheless, in an exploratory work carried out with a sample of HCAs from 2016 to 2018, the rates were observed to decrease but the variation in terms of extremal quotient remained very high, up to 7.

As a final comment, these findings have to be interpreted in the context of the Spanish National Health System—institutional design and care pathways. While the methodology is directly applicable to other contexts, the results (ie, the magnitude of the observed difference across providers, and the fraction of variation attributable to each level of analysis) may likely differ from one system to another. Thus, for a proper generalisation of the results to a different context, a judgement on the institutional design and usual care pathways will be required.

Implications

Spanish rates are low compared with OECD countries; among 30 OECD countries in the period 1994–2014, Spain was the 4th country with the lowest rates of tonsillectomies.⁹ However, variation still remains very high and a question on whether surgery is provided to eligible patients comes up when in our sample half of the patients receiving tonsillectomy alone and three in four patients getting both T&A did not have any condition deemed as an indication of surgery.

With the corresponding caveats out of these figures, while in the case of tonsillectomy alone, clear guidelines, proper antibiotic policies and audit and feedback mechanisms could reduce unnecessary variation in surgery, the lack of evidence about the use of T&A would indicate that policies should seek the reduction of the rates, for example, using, second opinion or disinvestment programmes. As a note for the potential impact of these policies, if those HCAs had reduced their T&A intervention at the level of the HCAs in the first decile of rates (0 rate could be sensibly assumed as inappropriate), the Spanish National Health System would have saved, over the study period, 121 535 interventions with a dubious indication.

As a final note of caveat, some evidence suggests that reductions in T&A rates may translate into an increase in hospital admissions for retropharyngeal and parapharyngeal abscesses or for acute tonsillitis.⁴² Before making any recommendation on substantial reductions in T&A rates, ENT doctors should propose guidance on those

indications that would avoid those adverse events to appear.

CONCLUSIONS

Despite the rather low Spanish tonsils surgery rates, our study showed wide persistent variations with a stagnation of tonsillectomies alone and a steady increase in T&A rates, across HCAs and ACs, and over the years. Variation in the exposure to tonsils surgery could be explained by factors both, at HCAs and ACs levels, but also on how unevenly ACs have adopted an evidence-based approach to tackle with unexpectedly high rates of tonsils surgery.

Author affiliations

¹Servicio de Planificación, Govern de les Illes Balears Conselleria de Salut, Palma de Mallorca, Spain

²Data Science for Health Services and Policy Research, Instituto Aragonés de Ciencias de la Salud, Zaragoza, Spain

³Network for Research on Chronicity, Primary Care, and Health Promotion (RICAPS), Instituto de Salud Carlos III, Madrid, Spain

⁴Navarrabiomed—Unidad de metodología, Navarre Institute of Health Research, Pamplona, Spain

⁵Fundación para el Fomento de la Investigación Sanitaria y Biomédica de la Comunidad Valenciana, FISABIO, Valencia, Spain

Twitter Salvador Peiro @VoroPeiro

Contributors All authors contributed to conceptualising and designing the study. NM-L prepare the final dataset. JL and MC-M run the statistical analysis. MR, JL, MC-M, EA-P and SP commented on important intellectual content and made revisions. EC-R, MR and EB-D drafted the manuscript. All authors read and approved the final version of the manuscript. EC-R, MR and EB-D accept full responsibility for the work and/or the conduct of the study, had access to the data, and controlled the decision to publish.

Funding MR and EAP are partially funded by the Spanish Health Services Research on Chronic Patients Network (REDISSEC) - Institute of Health Carlos III. Grant number: RD16/0001/0007. There is also a non-conditional additional public contribution from the Institute of Health Carlos III. Grant number RD21/0016/0023.

Map disclaimer The inclusion of any map (including the depiction of any boundaries therein), or of any geographic or locational reference, does not imply the expression of any opinion whatsoever on the part of BMJ concerning the legal status of any country, territory, jurisdiction or area or of its authorities. Any such expression remains solely that of the relevant source and is not endorsed by BMJ. Maps are provided without any warranty of any kind, either express or implied.

Competing interests None declared.

Patient and public involvement Patients and/or the public were not involved in the design, or conduct, or reporting, or dissemination plans of this research.

Patient consent for publication Not applicable.

Ethics approval This study, observational by design, uses retrospective anonymised non-identifiable and non-traceable data, and was conducted in accordance with the amended Helsinki Declaration, the International Guidelines for Ethical Review of Epidemiological Studies, and Spanish laws on data protection and patients' rights.

Provenance and peer review Not commissioned; externally peer reviewed.

Data availability statement Data may be obtained from a third party and are not publicly available. Data may be obtained from a third party under request and are not publicly available.

Supplemental material This content has been supplied by the author(s). It has not been vetted by BMJ Publishing Group Limited (BMJ) and may not have been peer-reviewed. Any opinions or recommendations discussed are solely those of the author(s) and are not endorsed by BMJ. BMJ disclaims all liability and responsibility arising from any reliance placed on the content. Where the content includes any translated material, BMJ does not warrant the accuracy and reliability of the translations (including but not limited to local regulations, clinical guidelines,

terminology, drug names and drug dosages), and is not responsible for any error and/or omissions arising from translation and adaptation or otherwise.

Open access This is an open access article distributed in accordance with the Creative Commons Attribution Non Commercial (CC BY-NC 4.0) license, which permits others to distribute, remix, adapt, build upon this work non-commercially, and license their derivative works on different terms, provided the original work is properly cited, appropriate credit is given, any changes made indicated, and the use is non-commercial. See: <http://creativecommons.org/licenses/by-nc/4.0/>.

ORCID iD

Enrique Bernal-Delgado <http://orcid.org/0000-0002-0961-3298>

REFERENCES

- Glover JA. The incidence of tonsillectomy in school children. *Proc R Soc Med* 1938;31:1219–36.
- Bakwin H. Pseudodoxia pediatrica. *N Engl J Med Overseas Ed* 1945;232:691–7.
- Wennberg JE, Gittelsohn AM. Variations in medical care among small areas. *Sci Am* 1982;264:100–11.
- Glover JA. The paediatric approach to tonsillectomy. *Arch Dis Child* 1948;23:1–6.
- Van Den Akker EH, Hoes AW, Burton MJ, et al. Large international differences in (adeno)tonsillectomy rates. *Clin Otolaryngol Allied Sci* 2004;29:161–4.
- Suleman M, Clark MPA, Goldacre M, et al. Exploring the variation in paediatric tonsillectomy rates between English regions: a 5-year NHS and independent sector data analysis. *Clin Otolaryngol* 2010;35:111–7.
- Boss EF, Marsteller JA, Simon AE. Outpatient tonsillectomy in children: demographic and geographic variation in the United States, 2006. *J Pediatr* 2012;160:814–9.
- McPherson K, Wennberg JE, Hovind OB, et al. Small-area variations in the use of common surgical procedures: an international comparison of new England, England, and Norway. *N Engl J Med* 1982;307:1310–4.
- Crowson MG, Ryan MA, Rocke DJ, et al. Variation in tonsillectomy rates by health care system type. *Int J Pediatr Otorhinolaryngol* 2017;94:40–4.
- Fedeli U, Marchesan M, Avossa F, et al. Variability of adenoidectomy/tonsillectomy rates among children of the veneto region, Italy. *BMC Health Serv Res* 2009;9:25.
- Oterino de la Fuente D, Castaño E, Libero J, et al. Y grupo VPM-SNS variaciones en hospitalizaciones pediátricas POR procedimientos quirúrgicos Y diagnósticos seleccionados. Variaciones en pediatría en El sistema nacional de Salud. *Atlas Var Pract Med Sist Nac Salud* 2006;3:99–114.
- Scottish Intercollegiate Guidelines Network. Management of sore throat and indications for tonsillectomy: a national clinical guideline. N° 34; 1999.
- Melchor Díaz MA. Amigdalitis aguda. Criterios de amigdalectomía. *An Pediatr, Monogr* 2003;1:62–8.
- Scottish Intercollegiate Guidelines Network. Management of sore throat and indications for tonsillectomy: a national clinical guideline. N° 117, 2010. Available: <http://www.show.scot.nhs.uk/sign/home.htm>
- Baugh RF, Archer SM, Mitchell RB, et al. Clinical practice guideline: tonsillectomy in children. *Otolaryngol Head Neck Surg* 2011;144:S1–30.
- Patel HH, Straight CE, Lehman EB, et al. Indications for tonsillectomy: a 10 year retrospective review. *Int J Pediatr Otorhinolaryngol* 2014;78:2151–5.
- Mitchell RB, Archer SM, Ishman SL, et al. Clinical practice guideline: tonsillectomy in children (Update)-executive summary. *Otolaryngol Head Neck Surg* 2019;160:187–205.
- Mandavia R, Schilder AGM, Dimitriadis PA, et al. Addressing the challenges in tonsillectomy research to inform health care policy: a review. *JAMA Otolaryngol Head Neck Surg* 2017;143:943–7.
- Shaikh W, Vayda E, Feldman W. A systematic review of the literature on evaluative studies of tonsillectomy and adenoidectomy. *Pediatrics* 1976;57:401–7.
- Windfuhr JP. Indications for tonsillectomy stratified by the level of evidence. *GMS Curr Top Otorhinolaryngol Head Neck Surg* 2016;15:Doc09.
- Burton MJ, Glasziou PP, Chong LY, et al. Tonsillectomy or adenotonsillectomy versus non-surgical treatment for chronic/recurrent acute tonsillitis. *Cochrane Database Syst Rev* 2014;334.
- Lim J, McKean MC. Adenotonsillectomy for obstructive sleep apnoea in children. *Cochrane Database Syst Rev* 2009;2:CD003136.
- Venekamp RP, Hearne BJ, Chandrasekharan D. Tonsillectomy or adenotonsillectomy versus non-surgical management for obstructive sleep-disordered breathing in children. *Cochrane Database Syst Rev* 2015;10:CD011165.
- Lock C, Wilson J, Steen N, et al. North of England and Scotland study of tonsillectomy and adeno-tonsillectomy in children (NESSTAC): a pragmatic randomised controlled trial with a parallel non-randomised preference study. *Health Technol Assess* 2010;14:1–164.
- Wilson JA, Steen IN, Lock CA, et al. Tonsillectomy: a cost-effective option for childhood sore throat? further analysis of a randomised controlled trial. *Otolaryngology Head and Neck Surgery* 2010;146:122–8.
- Paradise JL, Bluestone CD, Colborn DK, et al. Tonsillectomy and adenotonsillectomy for recurrent throat infection in moderately affected children. *Pediatrics* 2002;110:7–15.
- van Staa BK, van den Akker EH, Rovers MM, et al. Effectiveness of adenotonsillectomy in children with mild symptoms of throat infections or adenotonsillar hypertrophy: open, randomised controlled trial. *Clin Otolaryngol* 2005;30:60–3.
- van Staa BK, van den Akker EH, Rovers MM, et al. Effectiveness of adenotonsillectomy in children with mild symptoms of throat infections or adenotonsillar hypertrophy: open, randomised controlled trial. *BMJ* 2004;329:651.
- Šumilo D, Nichols L, Ryan R, et al. Incidence of indications for tonsillectomy and frequency of evidence-based surgery: a 12-year retrospective cohort study of primary care electronic records. *Br J Gen Pract* 2019;69:e33–41.
- Cervera Escario J, Del Castillo Martín F, Gómez Campderá JA, et al. Indicaciones de adenoidectomía Y amigdalectomía: documento de consenso entre la sociedad española de otorrinolaringología Y patología cervicofacial Y la asociación española de pediatría. *Acta Otorrinolaringológica Española* 2006;57:59–65.
- Bernal-Delgado E, García-Armesto S, Oliva J, et al. Spain: health system review. *Health Syst Transit* 2018;20:1–179.
- Ministerio de Sanidad, Servicios Sociales e Igualdad. Informe anual del Sistema Nacional de Salud, 2015. Available: <http://www.msssi.gob.es/estadEstudios/estadisticas/sisInfSanSNS/tablasEstadisticas/InfAnSNS.htm>
- Angulo-Pueyo E, Comendeiro-Maaløe M, Estupiñán-Romero F, et al. Atlas VPM: two decades informing on unwarranted variations in health care in Spain. *Res Health Serv Reg* 2022;1:5.
- Besag J, York J, Molli A. Bayesian image restoration, with two applications in spatial statistics. *Ann Inst Stat Math* 1991;43:1–20.
- Knorr-Held L. Bayesian modelling of inseparable space-time variation in disease risk. *Stat Med* 2000;19:2555–67.
- Vaughan AS, Kramer MR, Waller LA, et al. Comparing methods of measuring geographic patterns in temporal trends: an application to county-level heart disease mortality in the United States, 1973 to 2010. *Ann Epidemiol* 2015;25:329–35.
- Librero J, Ibañez B, Martínez-Lizaga N, et al. Applying spatio-temporal models to assess variations across health care areas and regions: lessons from the decentralized Spanish National health system. *PLoS One* 2017;12:e0170480.
- NHS. Value based clinical commissioning policy. version 7.0 review, 2019. Available: <http://www.northcumbriaccg.nhs.uk/about-us/key-policies-and-documents/VBCCP/vbccp-v7-april-2019.PDF>
- Audit Commission. Croydon list: reducing spending on low clinical value treatments, health briefin. Available: <http://archive.auditcommission.gov.uk/auditcommission/sitecollectiondocuments/Downloads/20110414reducingexpenditure.pdf>
- Goodman DC, Challener GJ. Tonsillectomy: a procedure in search of evidence. *J Pediatr* 2012;160:716–8.
- Doganer YC, Rohrer JE, Aydoğan U, et al. Tonsillectomy, adenoidectomy and adenotonsillectomy rates in school-aged children: relative contributions of socio-demographic and clinical features. *Int J Pediatr Otorhinolaryngol* 2015;79:969–74.
- Douglas CM, Altmeyer U, Cottom L, et al. A 20-year observational cohort of a 5 million patient population-tonsillectomy rates in the context of two national policy changes. *Clin Otolaryngol* 2019;44:7–13.