Statistical Computing HW 4

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Problem 4.3)

Give an efficient algorithm to simulate the value of a random variable X such that

$$P{X = 1} = 0.3$$

 $P{X = 2} = 0.2$
 $P{X = 3} = 0.35$
 $P{X = 4} = 0.15$

Answer

Derivation

The CDF of the distribution is:

$$F(x) = \begin{cases} 0 & x < 1 \\ 0.3 & 1 \le x < 2 \\ 0.5 & 2 \le x < 3 \\ 0.85 & 3 \le x < 4 \\ 1 & 4 \le x \end{cases}$$

The can be used along with a standard uniform random variable to generate values of the random variable X.

Algorithm

This is slightly different than the inverse transform method in the book but the outcome is the same. 1. Generate a number u from the standard uniform distribution 2. Set $x = F^{-1}(u) + 1$ 3. Repeat n times

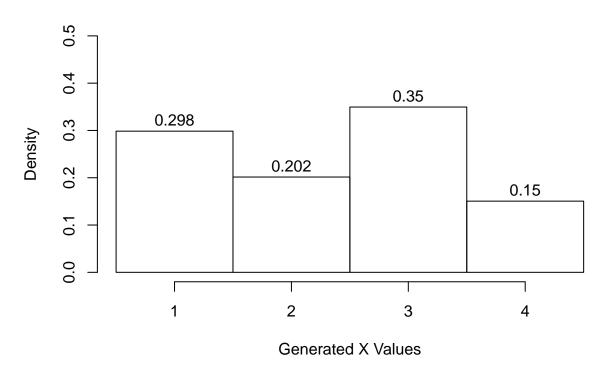
Program

Below is a program and histogram using the derivation and algorithm. 10,000 values were generated. The distribution is quite good.

```
# F inverse function
F_inv <- function(u){
    x <- rep(0, length(u))
    if(u < 0.30)
        x <- 0 + 1
    else if((u >= 0.30) && (u < 0.50))
        x <- 1 + 1
    else if((u >= 0.50) && (u < 0.85))
        x <- 2 + 1
    else if((u >= 0.85) && (u < 1))</pre>
```

```
x <- 3 + 1
return(x)
}
# This line is generates the x values
x <- sapply(array(runif(10^4)), FUN = F_inv)</pre>
```

Distribution of 10,000 generated X values



Problem 4.14a)

Let X be a binomial random variable with parameters n and p. Suppose that we want to generate a random variable Y whose probability mass function is the same as the conditional mass function of X given that $X \ge k$, for some $k \le n$. Let $\alpha = P\{X \ge k\}$ and suppose that the value of α has been computed. Give the inverse transform method for generating Y.

Answer

Derivation

We are given

$$Y = \frac{P(X=x)}{P(X \ge k)} = \frac{P(X=x)}{\alpha} \text{ for } x = k, k+1, ..., n$$

Then:

$$P(Y = y) = \frac{\binom{n}{y}p^{y}(1-p)^{n-y}}{\alpha}$$
 for $y = k, k+1, ..., n$

Then we can let $y_0 = k$, $y_1 = k + 1$, $y_2 = k + 2$, and so on until $y_{n-k} = n$. Then finally we use the rule for the inverse transform method:

$$Y = y_j \text{ if } F(y_{j-1}) \le U < F(y_j)$$

where

$$F(y_j) = \sum_{i=k}^{k+j} P(Y=i)$$

Algorithm

- 1. Generate u from U(0,1)
- 2. Set $y = y_j$ according to $F(y_{j-1}) \le u < F(y_j)$
- 3. Repeat n times

Problem 4.15)

Give a method for simulating X, having the probability mass function p_j , j = 5, 6, 7, ..., 14, where

$$p_j = \begin{cases} 0.11 & \text{when } j \text{ is odd and } 5 \le j \le 13\\ 0.09 & \text{when } j \text{ is even and } 6 \le j \le 14 \end{cases}$$

Use the text's random number sequence to generate X.

Answer

(What is the text's random number sequence?)

The method used will the composition approach.

Derivation

The composition approach works by finding α , $p_i^{(1)}$, and $p_i^{(2)}$ so that

$$P\{X = j\} = \alpha p_j^{(1)} + (1 - \alpha)p_j^{(2)}$$

Let $\alpha = 0.55$ and let

$$p_j^{(1)} = P\{X = j\} = \begin{cases} 0 & \text{for } x \text{ even} \\ 0.2 & \text{for } x \text{ odd} \end{cases}$$

and

$$p_j^{(2)} = P\{X = j\} = \begin{cases} 0 & \text{for } x \text{ odd} \\ 0.2 & \text{for } x \text{ even} \end{cases}$$

In other words, if j_0 is odd, then $P(X = j_0) = (0.55)(0.20) + (0.45)(0) = 0.11$ and if j_0 is even then $P(X = j_0) = (0.55)(0) + (0.45)(0.20) = 0.09$. This is the desired outcome.

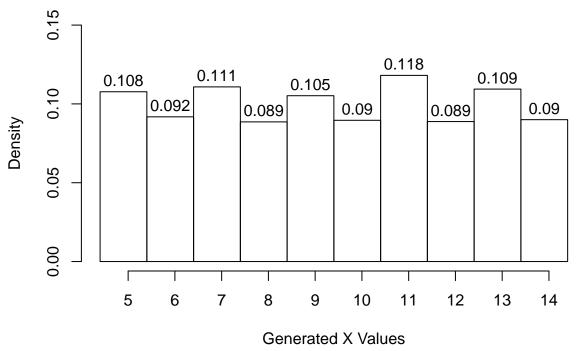
Algorithm

- 1. Generate u_1 from standard uniform distribution
- 2. Generate u_2 from standard uniform distribution
- 3. If $u_1 < \alpha$, (odd case) set $x = 2(floor(5u_2)) + 5$
- 4. Otherwise (even case) set $x = 2(floor(5u_2)) + 6$
- 5. Go back to step 1, repeat n times

Program

```
# Initialize variables, generate uniform random numbers
N <- 10^4
alpha <- 0.55
uniform1 <- runif(N)
uniform2 <- runif(N)
x_dist <- c(rep(0, N))
# These two lines generate the x values
x_dist[uniform1 < 0.55] <- 2*floor(5*uniform2[uniform1 < 0.55])+5
x_dist[x_dist == 0] <- 2*floor(5*uniform2[uniform1 >= 0.55])+6
```

Distribution of 10,000 Generated X Values



This plot of the 10,000 generated x values appears reasonable given what we would expect from the distribution defined in the question.

Problem 5.10)

A casualty insurance company has 1000 policyholders, each of whom will independently present a claim in the next month with probability 0.05. Assuming that the amounts of the claims made are independent exponential random variables with mean \$800, use simulation to estimate the probability that the sum of these claims exceeds \$50,000.

Answer

Derivation

In order to generate random values from an exponential distribution with mean 800, we'll use the fact that the CDF of this distribution is

$$F(x) = 1 - e^{-x/800}$$

Given that F(x) has a standard uniform distribution, then:

$$u = 1 - e^{-x/800}$$

$$e^{-x/800} = 1 - u$$

$$\frac{-x}{800} = \ln(1 - u)$$

$$x = (-800)\ln(u)$$

Note that the random variables U and 1-U have the same distribution if U is standard uniform.

Also, since each policyholder has a 5% chance of making a claim, then the number of claims made out of 1000 has a binomial distribution with n = 1000 and p = 0.05. The random variable we are interested in is then

$$S = X_1 + X_2 + X_3 + \dots + X_N$$

where the X_i are *iid* exponential with mean 800 and N is binomial with n = 1000 and p = 0.05. For simplicity in the program, values from the binomial distribution will be done using an R function, but the exponential values will be generated using values from a uniform distribution according to the derivation above.

Algorithm

- For each trial:
 - Generate a value n from Binomial (1000, 0.05)
 - Generate n values $u_1, u_2, ..., u_n$ from U(0,1)
 - Calculate each of the n claim amounts x by evaluating $x_i = (-800)ln(u_i)$
 - Find the sum of the n claims for that trial
- After the sum of each trial is calculated, then determine how many are greater than \$50,000

The R code appears different than this algorithm, but in practice this is what is happening. The R language benefits from avoiding explicitly writing for loops, and instead using other methods like apply, lapply, and sapply.

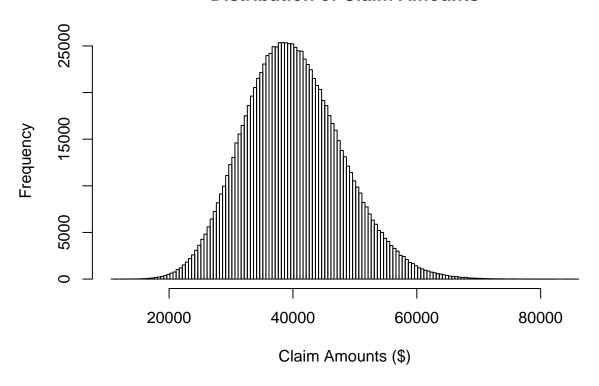
Program

The program uses the $generate_claims()$ function to generate n claims from an exponential distribution, where n has a binomial distribution. 1,000,000 trials are run, and then the sum of the claims from each trial is calculated.

```
\# Set seed, set N = number of simulations
set.seed(1234)
N = 10^6
# Returns variable number claims from exponential distribution
generate_claims <- function(vector, holders = 1000, chance = 0.05, beta = 800){</pre>
  # Number of claims that will occur is binomial(holders, chance)
  num_of_claims <- rbinom(1, holders, chance)</pre>
  # Generate claim amounts from exponential distribution for each claimant
  claims <- -beta * log(runif(num_of_claims))</pre>
  # Return vector of claims
  return(claims)
}
# Initialize claims list
claims \leftarrow \text{rep}(\text{list}(\text{rep}(0,100)), N)
# Simulate N times
claims <- lapply(claims, generate_claims)</pre>
# Find the sum for each trial
distribution <- sapply(claims, sum)</pre>
```

Below is a histogram of the results. The distribution is a sum of exponential random variables and thus is a continuous distribution. It appears to be slightly right tailed with a mean around \$40,000.

Distribution of Claim Amounts



The distribution seems to be centered at around \$40,000 which we would expect since this is the product of the mean of the two distributions (800 and 50). The following R command returns the estimated probability that the sum of the claims is above \$50,000:

mean(distribution>50000)

```
## [1] 0.107028
```

Out of 1,000,000 trials, the approximate proportion of trials where the total value of the claims was above \$50,000 was 0.107028.

Analytical Result

The best idea I can come up with is that this is a mixture distribution. The different distributions are Erlang distributions (sum of exponential distributions) and the weights are binomial weights for each. If we are interested in the distribution of the sum of the claims, denoted s(x), then

$$s(x) = \sum_{i=0}^{1000} w_i p_i(x)$$

$$= \sum_{i=0}^{1000} \left[\binom{1000}{i} (0.05)^i (0.95)^{1000-i} \right] \left[\frac{(1/800)^i x^{i-1} e^{-x/800}}{(i-1)!} \right]$$

$$= \sum_{i=1}^{1000} \left[\binom{1000}{i} (0.05)^i (0.95)^{1000-i} \right] \left[\frac{(1/800)^i x^{i-1} e^{-x/800}}{(i-1)!} \right]$$

where the left term is the binomial weight and the right term is the Erlang PDF. The first term in the sum is 0 (the sum of zero exponential random variables is zero). The CDF of the Erlang distribution is

$$F(y) = 1 - \sum_{n=0}^{k-1} \frac{1}{n!} e^{-\lambda y} (\lambda y)^n$$

which then means

$$P(Y>y) = \sum_{n=0}^{k-1} \frac{1}{n!} e^{-\lambda y} (\lambda y)^n$$

Then if we place this term into the mixture distribution formula above and let x = 50,000 and $\lambda = 1/800$ we get:

$$P(X > 50000) = \sum_{i=1}^{1000} \left[{1000 \choose i} (0.05)^i (0.95)^{1000-i} \right] \left[\sum_{n=0}^{i-1} \frac{1}{n!} e^{-62.5} (62.5)^n \right]$$

Evaluating this sum should give us the chance that the sum of the claims is above \$50,000. That's as far as an analytical result that I could reach. To carry out the calculations I used R. The only problem is that there are very large factorials involved. The largest factorial that R can handle is 170!. Though there are factorials up to 1000 in the equation, notice that since the binomial probabilities are very small for i > 170 and the factorial in the denominator of the Erlang term also make those terms very small, calculating from i = 1 to i = 170 will give us an approximation to the analytical result.

```
probability <- c(rep(0,170))
for (i in 1:length(probability)){
   probability[i] <- dbinom(i, 1000, 0.05) * sum(1/factorial(0:(i-1)) * 62.5^(0:(i-1)))
}
probability <- exp(-62.5) * sum(probability)</pre>
```

probability

[1] 0.1070977

This method says that the probability that the sum of the claims is above \$50,000 is approximately 0.1070977 which is a little different than the simulation's result. The absolute difference is 0.0000697. The estimation of the analytical result is a very slight underestimation because the terms after i = 170 were excluded from the summation, and even though they are extremely small, they are still positive.