#### 09\_Repeated measures

### Repeated measures, simple methods

#### Anders Stockmarr

Course developers: Anders Stockmarr, Elisabeth Wreford Andersen, Per Bruun Brockhoff

> DTU Department of Applied Mathematics and Computer Science Section for Statistics and Data Analysis Technical University of Denmark anst@dtu.dk

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### Plan for this week

#### 09\_Repeated measures

Monday Statistical inference, and the t-test
Tuesday Simple and Multiple regression
Wednesday ANOVA, ANCOVA, and linear models
Thursday Categorical data, Writing statistical reports,
Logistic regression
Friday Repeated measurements, Principal Component

Analysis

## Overview of this module Repeated measures

- The repeated measurements setup
  - Aspirin Example
  - Activity of rats
- Separate analysis for each time-point
  - Example: rats data
- Analysis of a summary statistic
  - Example: rats data
- Random effects model simple version
  - Example: rats data
- Random effects model advanced version
- Pros and cons of simple approaches

### Aim of this module

09\_Repeated measures

- Present simple methods for dealing with repeated measurements and a few more advanced applications
- Easy to use, and a lot better than pretending to have independent observations
- Useful even after more advanced models are presented
- See how to specify these models in R

### Overview

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- Opening in the implement of the imple

## The repeated measurements setup

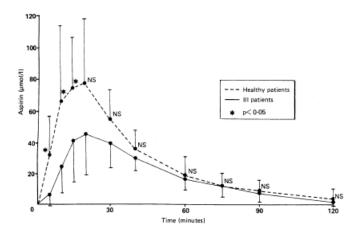
- Several "individuals"
- Several measurements on each individual
- Two measurements on the same individual might be correlated
- Might even be highly correlated if "close" and less correlated if "far apart"

## The repeated measurements setup

- Several "individuals"
- Several measurements on each individual
- Two measurements on the same individual might be correlated
- Might even be highly correlated if "close" and less correlated if "far apart"
- Typical example (Matthews et al 1990):
  - Two groups of patients (ill/healthy)
  - Treated with asprin
  - The asprin concentration is measured in the blood 10 times during 2 hours

To pretend all observations are independent can lead to wrong conclusions

## Traditional figure of means at each time point



## What is the aim of the structure measures

It is always important to consider the aim of the study!

- 1 To describe pattern over time
  - Do we have a change over time?
  - Linear or curve?
  - Is the pattern the same for all groups?
- 2 Can we see a difference between groups?
  - Is the difference the same at all time points?
  - Is the difference in the levels?
  - Is the difference in the trend?

## Problems using the traditional presentation of the data

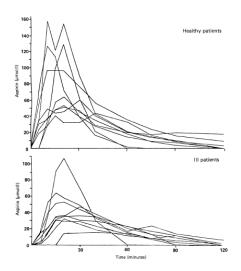
- Comparing the groups at each time point.
  - Not effecient.
  - Tests are not independent, carried out on the same individuals
  - Difficult to interpret
- The pattern over time
  - This can only be studied if we use that each individual has several observations

## Take care when using average christies

- Important structures can be overlooked if you start with the average curves.
- You cannot see the variation over time.
- Always make a plot of the individual observations.
- Only use the average if the patterns are similar, i.e. we only see shifts up and down.

### Individual curves

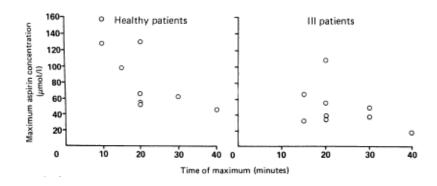
#### 09\_Repeated measures



## Analysis of summary statistics measures

- Choose a single measure to summarize the individual curves
- This again reduces the data set to independent observations
- Popular choices of summary measures "relevant" feature extraction:
  - Total increase (last point minus first point)
  - Area under curve (AUC)
  - Maximum or minimum point
  - Average over time
  - Slope in regression with time
  - (or higher order polynomial coefficients)
- Good method with few and easily checked assumptions
- Information may be lost
- Important to choose a good summary measure

### 09\_Repeated measures



## Analysis of aspirin study Repeated measures

TABLE I — Analysis of data from aspirin study

	Area under curve	Maximum concentration		
Health	patients (n = 9)			
Arithmetic mean (SD) (µmol/l)	26.5 (8.8)	86.0 (41.5)		
Geometric mean (µmol/l)	25-4	77.8		
Ill p	atients $(n = 9)$			
Arithmetic mean (SD) (µmol/l)	17-5 (5-0)	46.7 (26.3)		
Geometric mean (µmol/l)	16-8	41.2		
Ratio of geometric means	1.52	1.89		
95% Confidence interval	1·11 to 2·08	1·14 to 3·13		
t Test	2.83 (df = 16)	2.66 (df = 16)		
p Value	0.01	0.02		

## Example: Activity of rate Repeated measures

#### Summary of experiment:

- 3 treatments: 1, 2, 3 (concentrations)
- 10 cages per treatment
- 10 months
- ullet The response is activity ( $\log(\text{count})$  of intersections of light beam during 57 hours)

### The rats data

#### 09 Repeated measures

```
rats <- read.csv("rats.txt")
# make treatment and cage factors
rats$treatm <- factor(rats$treatm)
rats$cage <- factor(rats$cage)
# make two versions of the time variable
 # - one quantitative and one factor
rats$monthQ <- rats$month
rats$month <- factor(rats$month)
summary(rats)
str(rats)
```

In this example we have repeated measurements from each cage.

## Data in two formats, long Repeated measures

This is the format we want. One response (here Inc) in each line

## Data in two formats, wid Repeated measures

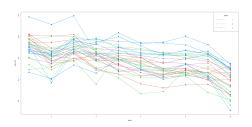
One line for each cage.

```
treatm cage
                  lnc.1
                          lnc.2 lnc.3 lnc.4 lnc.5 lnc.6
                                                              lnc. 7
                                                                      1nc.8
##
                   9.9323 9.6447 9.7628 9.6014 9.3227 9.2463 9.0739 9.2077
##
  11
                  10.0547 9.7384 9.6928 9.4676 9.5297 9.3837 9.2052 9.2582
                   9.7448 9.4278 9.5995 9.2614 9.3744 9.0857 9.1068 9.1732
  31
                   9.8660 9.8687 9.9329 9.7183 9.6861 9.5047 9.2622 9.2699
##
                   9.9552 9.5536 9.4930 9.4369 9.4520 9.4147 9.4107 9.3751
       Inc. 9 Inc. 10
##
      9.1670 8.8319
      9.0681 8.8753
     8.9907 8.9995
   31 9.1555 9.1584
  41 9.2475 9.0435
```

Not usually the format wanted for analysis but adapted to plots and group averages.

### The rats data - plotting the individual patterns

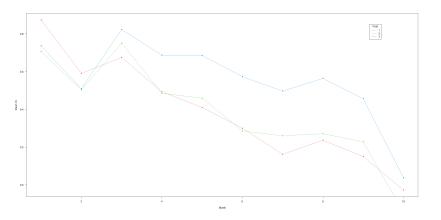
```
plot(rats$monthQ,rats$lnc,xlab="Months",
  ylab="log(count)",type ="n")
for(i in 1:dim(rats2)[1]){
  lines(1:10,rats2[i,-(1:2)],type="b",col=rats2$treatm[i]+1,
  pch=16, lwd=2, cex=2)
legend("topright",legend=1:3,col=2:4,pch=16,lty=2,title="treatm")
```



### The rats data - plotting the average patterns

```
Mean_data <- aggregate(rats$lnc, by = list(rats$month, rats$treatm)
                      mean)
names(Mean_data) <- c("month","treatm","Meanlnc" )</pre>
#Plot the means
Grp1<-subset(Mean_data,treatm==1)</pre>
Grp2<-subset(Mean_data,treatm==2)</pre>
Grp3<-subset(Mean_data, treatm==3)</pre>
plot(as.numeric(Grp1$month), Grp1$Meanlnc, type = "b", pch=16,
  xlab = "Month",ylab = "Mean lnc", las = 1, col= 2)
lines(as.numeric(Grp2$month), Grp2$Meanlnc, type = "b",
     col = 3, pch=16)
lines(as.numeric(Grp3$month), Grp3$Meanlnc, type = "b",
     col = 4, pch=16)
legend = c("1","2","3"),
      title = "Treat", lty = c(2,2,2), col = 2:4)
```

### The rats data - plotting the average patterns



### Overview

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- Sandom effects model advanced version
- 6 Pros and cons of simple approaches

## Separate analysis for each time point (continued)

- Select a fixed time point
- The observations at that time (one from each individual) are independent
- Do a separate analysis for the observations at that time
- This is not wrong, but (possibly) a lot of information is wasted
- This can be done for several time-points, but
  - Difficult to reach a coherent conclusion
  - Sub-tests are not independent
  - Tempting to select time-points supporting our preference
  - Mass significance: If many tests are carried out at 5% level some might be significant by chance. (Bonferroni correction: Use significance level 0.05/n instead of 0.05)

## Separate analysis of rats\_datad measures

## Separate analysis of rats data measures

• The model at each month is:

$$lnc_i = \mu + \alpha(treatm_i) + \varepsilon_i$$
,  $\varepsilon_i \sim i.i.d.$   $N(0, \sigma^2)$ ,  $i = 1...30$ 

• The result of the ten tests for no treatment effect:

Month	1	2	3	4	5	6	7	8	9	10
F-value	1.22	0.27	1.02	2.30	3.87	4.10	4.70	7.29	4.09	0.88

Compare with  $F_{95\%;2,27}=3.35$  or  $F_{99.5\%;2,27}=6.49$  if Bonferroni correction is used

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### Rats data analyzed via of unpracty and sure

- The log of the total activity in each cage is chosen as summary measure lnTot = log(Total count)
- The one way ANOVA model becomes:

$${\tt lnTot}_i = \mu + \alpha({\tt treatm}_i) + \varepsilon_i, \quad \varepsilon_i \sim {\sf i.i.d.} \ N(0,\sigma^2), \quad i=1\dots 30$$

• Notice the simplicity of the model and the relative few assumptions

# Rats data analyzed via of Repeated measure – log total activity

```
rats$count <- exp(rats$lnc)
DataTot <- aggregate(rats$count, by = list(rats$cage,
                                         rats$treatm), sum)
names(DataTot) <- c("cage", "treatm", "Tot")</pre>
head(DataTot)
    cage treatm Tot
##
##
              1 124864
## 2 3 1 131277
## 3 5 1 110166
## 4 7
         1 145418
## 5 9
          1 128819
## 6
      11
              1 136214
```

## log total activity (contdo)\_Repeated measures

```
reg1 <- lm(log(Tot) ~ treatm, data = DataTot); summary (reg1)
## Coefficients:
##
          Estimate Std. Error t value Pr(>|t|)
## (Intercept) 11.73354   0.05409   216.91   <2e-16 ***
## treatm2 -0.00567 0.07650 -0.07 0.941
## treatm3 0.16728 0.07650 2.19 0.038 *
##
## Residual standard error: 0.171 on 27 degrees of freedom
## Multiple R-squared: 0.196, Adjusted R-squared: 0.137
## F-statistic: 3.3 on 2 and 27 DF, p-value: 0.0522
anova (reg1)
## Analysis of Variance Table
##
## Response: log(Tot)
     Df Sum Sq Mean Sq F value Pr(>F)
##
## treatm 2 0.193 0.0965 3.3 0.052 .
## Residuals 27 0.790 0.0293
```

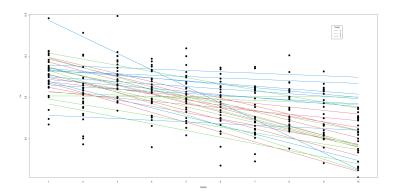
### Rats data analyzed via a supplement measure - slopes

```
# byCage is a regression with different slope and intercept
# for each cage. Saves the coefficients
by Cage \leftarrow coef (lm(lnc ~ -1 + cage + monthQ: cage, data = rats))
slope1 <- data.frame(matrix(byCage, nrow=30, byrow=F))</pre>
names(slope1) <- c("Intercept", "Slope")</pre>
#We also need info about cage and treat
SlopeData <- cbind(DataMean[ , 1:2], slope1)</pre>
head(SlopeData)
##
     cage treatm Intercept Slope
             1 9.9685 -0.107163
## 1
## 2 3 1 10.0411 -0.111593
## 3 5 1 9.6981 -0.076675
## 4 7
              1 10.0831 -0.098336
              1 9.8067 -0.066990
## 5 9
      11
              1 9.7310 -0.043147
## 6
```

### Rats data analyzed via of supermary measure - plotting slopes

```
# Individual slopes:
fit <- unlist(by(rats, rats$cage,
          function(x) fitted.values(lm(lnc ~ monthQ, data=x))))
names(fit) <- NULL
#plotting the linear fit by cage
interaction.plot(rats$monthQ, rats$cage, fit,
  xlab="Month", ylab="lnc",legend=F,
  col=as.numeric(SlopeData$treatm)+1,lty=1,lwd=2)
lines(rats$monthQ,rats$lnc, type="p",pch=16,cex=2)
legend(locator(1), # we will place it with a mouse click
       legend = c("1","2","3"), title = "Treat",
       lty = c(1,1,1), lwd=2, col= 2:4)
```

## Rats data analyzed via og\_supmenary measure - plotting slopes



## Rats data analyzed via of summary measure - slopes (contd)

## Rats data analyzed via of summary measure - slopes (contd)

- Treat 1 has been chosen as reference with estimated slope -0.09.
- ullet Treat 2 is very similar to Treat 1 with slope -0.09+0.01=-0.08
- Treat 3 has slope closer to 0: -0.09 + 0.04 = -0.05

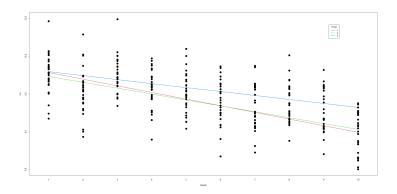
## Rats data analyzed via of summary measure - slopes (contd)

From the one-way analysis of variance we do not see a significant difference in the slopes between the three groups.

## Rats data analyzed via of support and support support and support support and support support and support support and support support and support support and support and support and support and support support

```
# Treatment average slopes:
fit2 <- unlist(by(rats, rats$treatm,</pre>
          function(x) fitted.values(lm(lnc ~ monthQ, data=x))))
names(fit2) <- NULL</pre>
#plotting the linear fit by treatment
interaction.plot(rats$monthQ, rats$treatm, fit2,
  xlab="Month", ylab="lnc",legend=F,
  col=2:4,lty=1,lwd=2,ylim=c(8.5,10.5))
lines(rats$monthQ,rats$lnc, type="p",pch=16,cex=2)
legend(locator(1),
       legend = c("1","2","3"), title = "Treat",
       lty = c(1,1,1), lwd=2, col= 2:4)
```

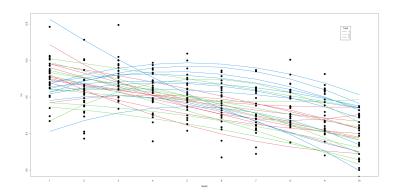
## Rats data analyzed via of supermary measure - plotting slopes



## Rats data analyzed via of summary measure - plotting curves

```
# Individual curves: Look similar apart from one of the blue
rats$monthQ2 <- rats$monthQ^2
fit3 <- unlist(by(rats, rats$cage,
   function(x) fitted.values(lm(lnc ~ monthQ+monthQ2, data=x))))
names(fit3) <- NULL</pre>
#plotting the fit by cage
interaction.plot(rats$monthQ, rats$cage, fit3,
  xlab="Month", ylab="lnc",legend=F,
  col=as.numeric(SlopeData$treatm)+1,lty=1,lwd=2,)
lines(rats$monthQ,rats$lnc, type="p",pch=16,cex=2)
legend(locator(1),
       legend = c("1","2","3"), title = "Treat",
       lty = c(1,1,1), lwd=2, col= 2:4)
```

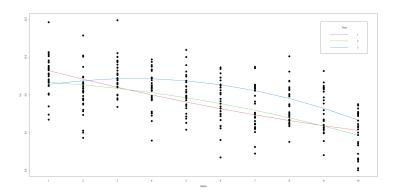
## Rats data analyzed via 09\_Repeated measure - plotting curves



## Rats data analyzed via of supplied measure - plotting curves

```
# Treatment average curves:
fit4 <- unlist(by(rats, rats$treatm,</pre>
  function(x) fitted.values(lm(lnc ~ monthQ+monthQ2, data=x))))
names(fit4) <- NULL</pre>
#plotting the fit by treatment
interaction.plot(rats$monthQ, rats$treatm, fit4,
  xlab="Month", ylab="lnc",legend=F,
  col=2:4,lty=1,lwd=2,ylim=c(8.5,10.5))
lines(rats$monthQ,rats$lnc, type="p",pch=16,cex=2)
legend(locator(1),
       legend = c("1","2","3"), title = "Treat",
       lty = c(1,1,1), lwd=2, col= 2:4)
```

## Rats data analyzed via of Repeated measure - plotting curves



## Rats data analyzed via of summary measure - curvatures

```
byCage2 <- coef(lm(lnc ~ -1 + cage + monthQ:cage +
                    monthQ2:cage, data = rats))
#I want a data frame, one row for each cage
curve <- data.frame(matrix(byCage2, nrow=30, byrow=F))</pre>
names(curve) <- c("Intercept", "monthQ", "monthQ2")</pre>
#We also need info about cage and treat
CurveData <- cbind(DataMean[ , 1:2], curve)</pre>
head(CurveData)
##
    cage treatm Intercept monthQ monthQ2
              1 10.0465 -0.146171 0.0035462
## 1
## 2 3
              1 10.0886 -0.135343 0.0021591
## 3 5
              1 9.8302 -0.142721 0.0060042
## 4
              1 10.0237 -0.068641 -0.0026996
              1 9.8664 -0.096853 0.0027148
## 5
              1 9.8445 -0.099881 0.0051576
## 6
      11
```

## Rats data analyzed via of summary measure - curvatures (contd)

### Each estimate of the curve is used as the response

```
reg4 <- lm(monthQ2 ~ treatm, data = CurveData)</pre>
summary (reg4)
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
##
  (Intercept) 0.00320 0.00298 1.08 0.29185
  treatm2 -0.00854 0.00421 -2.03 0.05237.
  treatm3 -0.01608 0.00421 -3.82 0.00071 ***
##
## Residual standard error: 0.00941 on 27 degrees of freedom
## Multiple R-squared: 0.351, Adjusted R-squared: 0.303
## F-statistic: 7.3 on 2 and 27 DF, p-value: 0.00291
```

- For Treat 1 the estimated curve is positive (nearly zero 0.003)
- For Treat 2 the estimated curve is negative (0.003-0.009=-0.006)
- For Treat 2 the estimated curve is also negative (0.003-0.016=-0.013)

# Rats data analyzed via of Repeated Areas measure - curvatures (contd)

When using the curvature as the response we get a significant difference between the three groups (p=0.003<0.05)

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## Random effects model of simple yersion

- This model uses all observations instead of reducing to one observation per individual
- We can test the time-by-treatment interaction
- Add "individual" as a random effect.
- Makes measurements on same individual correlated
- Unfortunately equally correlated no matter if they are "close" or "far apart"
- Can be considered first step in modeling the actual covariance structure
- Usually only good for short series

## Rats data analyzed via gandom effects approach

• The model can now be enhanced to:

$$\mathtt{lnc}_i = \mu + \alpha(\mathtt{treatm}_i) + \beta(\mathtt{month}_i) + \gamma(\mathtt{treatm}_i, \mathtt{month}_i) + d(\mathtt{cage}_i) + \varepsilon_i \;,$$

• The covariance structure of this model is:

$$\mathrm{cov}(y_{i_1},y_{i_2}) = \left\{ \begin{array}{ll} 0 & \text{, if } \mathrm{cage}_{i_1} \neq \mathrm{cage}_{i_2} \mathrm{\ and\ } i_1 \neq i_2 \\ \sigma_d^2 & \text{, if } \mathrm{cage}_{i_1} = \mathrm{cage}_{i_2} \mathrm{\ and\ } i_1 \neq i_2 \\ \sigma_d^2 + \sigma^2 & \text{, if } i_1 = i_2 \end{array} \right.$$

### Rats data analyzed via gandom effects approach

• The model can now be enhanced to:

$$\texttt{lnc}_i = \mu + \alpha(\texttt{treatm}_i) + \beta(\texttt{month}_i) + \gamma(\texttt{treatm}_i, \texttt{month}_i) + d(\texttt{cage}_i) + \varepsilon_i \;,$$

• The covariance structure of this model is:

$$\mathrm{cov}(y_{i_1},y_{i_2}) = \left\{ \begin{array}{ll} 0 & \text{, if } \mathrm{cage}_{i_1} \neq \mathrm{cage}_{i_2} \mathrm{\ and\ } i_1 \neq i_2 \\ \sigma_d^2 & \text{, if } \mathrm{cage}_{i_1} = \mathrm{cage}_{i_2} \mathrm{\ and\ } i_1 \neq i_2 \\ \sigma_d^2 + \sigma^2 & \text{, if } i_1 = i_2 \end{array} \right.$$

One can think of the model as:

log(count)="mean"+"between cage variation" + "within cage variation"

## Multilevel model structus Repeated measures

• This model is also called a two-level model.

Level	1	2
Unit	Each separate obs.	Cages
Variation	Within cage	Between cage
	$\sigma^2$	$\sigma_d^2$
Covariates	month	treatm
	month:treatm	

### Model synonyms

#### 09\_Repeated measures

- Two-level model
- Mixed model with random subject level
- Mixed model with random intercept
- Model with compound symmetry correlation structure
- Model with exchangable correlation structure

### Compound symmetry 09 Repeated measures

The model implies that all observations from the same cage are correlated with the intra-class correlation:

$$Corr(Y_{i_1}, Y_{i_2}) = \frac{cov(Y_{i_1}, Y_{i_2})}{\sqrt{var(Y_{i_1})}\sqrt{var(Y_{i_1})}} = \frac{\sigma_d^2}{\sigma_d^2 + \sigma^2}$$

Here  $cage_{i_1} = cage_{i_2}$  and  $i_1 \neq i_2$ .

We are not taking the distance between observations into account.

Maybe too simple? Perhaps observations close in time are more similar than observations far apart.

This correlation structure is called exchangable or compound symmetry.

## Potential mistakes when keaving the random effect out

Level	Unit	Covariates
1	Each separate obs.	month
		month:treatm
2	Cages	treatm

### If a random effect is present:

- Potential bias in the mean (ignoring that observations go together)
- ullet Estimates on Level 1 may have too much variation. Ignoring pairs o too high p-values. Effects can be overlooked
- Estimates on Level 2 may have too little variation → too small p-values. "Noise" can become an effect

## Rats data analyzed via gardom effects approach in R

So we have a significant interaction between treatment and month.

## Table of fixed effects 09 Repeated measures

```
#This time summary gives a lot of output!
#summary(model1)
#A table of fixed effects estimates
summary(model1)$tTable
##
                      Value Std.Error DF
                                            t-value
                                                        p-value
  (Intercept)
                             0.080856 243
                                          122.12192 4.1653e-220
                  9.874280
## month2
                  -0.282870 0.087062 243
                                           -3.24908 1.3213e-03
## month3
                  -0.199010 0.087062 243
                                           -2.28585 2.3124e-02
## month4
                   -0.381285
                             0.087062 243
                                           -4.37948
                                                     1.7692e-05
## month5
                   -0.464289
                             0.087062 243
                                           -5.33288
                                                     2.2101e-07
## month7:treatm3
                   0.472180
                             0.123124 243
                                            3.83500
                                                     1,6005e-04
  month8:treatm3
                   0.465124
                             0.123124 243
                                            3.77770
                                                     1.9913e-04
## month9:treatm3 0.443817
                             0.123124 243
                                            3.60464
                                                     3.7919e-04
                                                     1.0190e-01
## month10:treatm3
                   0.202159
                             0.123124 243
                                            1.64192
```

## Table of random effect 99\_Repeated measures

```
#The estimates random effects
VarCorr(model1)

## cage = pdLogChol(1)

## Variance StdDev

## (Intercept) 0.027478 0.16577

## Residual 0.037899 0.19468

#intra class correlation
0.027478/(0.027478+0.037899)

## [1] 0.4203
```

$$\hat{\sigma}_d^2 = 0.03$$
 and  $\hat{\sigma}^2 = 0.04$ 

### Overview

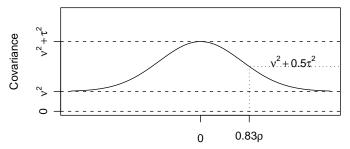
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## Advanced Covariance Structures Gaussian spatial correlation

- Spatial covariance structures, depending on "how far" observations are apart.
- Gaussian decline of dependency:

$$Cov(Y_{i_1},Y_{i_2}) \quad = \quad \left\{ \begin{array}{l} 0 \\ \nu^2 + \tau^2 \exp\left\{\frac{-(t_{i_1} - t_{i_2})^2}{\rho^2}\right\} \\ \nu^2 + \tau^2 + \sigma^2 \end{array} \right. \quad \text{, if } \mathrm{indiv}_{i_1} \neq \mathrm{indiv}_{i_2} \text{ and } i_1 \neq i_2 \\ \end{array}$$



Distance  $t_{i_1} - t_{i_2}$ 

## Rats data via spatial Goussian correlation model

• Model for the entire observational vector:

$$\begin{array}{rcl} \mathbf{Y} & \sim & N(\mu,\mathbf{V}), & \text{where} \\ \mu_i & = & \mu + \alpha(\mathtt{treatm}_i) + \beta(\mathtt{month}_i) + \gamma(\mathtt{treatm}_i,\mathtt{month}_i), & \text{and} \\ \\ V_{i_1,i_2} & = & \left\{ \begin{array}{l} 0 & , & \mathrm{if} \; \mathtt{cage}_{i_1} \neq \mathtt{cage}_{i_2} \; \mathtt{and} \; i_1 \neq i_2 \\ \\ \nu^2 + \tau^2 \exp \left\{ \frac{-(\mathtt{month}_{i_1} - \mathtt{month}_{i_2})^2}{\rho^2} \right\} & , & \mathrm{if} \; \mathtt{cage}_{i_1} = \mathtt{cage}_{i_2} \; \mathtt{and} \; i_1 \neq i_2 \\ \\ \nu^2 + \tau^2 + \sigma^2 & , & \mathrm{if} \; i_1 = i_2 \end{array} \right.$$

• R code for model implementation:

```
analysis<-lme(lnc~month+treatm+month:treatm,
  random=~1 | cage,
  correlation=corGaus(form=~as.numeric(month) | cage,nugget=T),
  data=rats)</pre>
```

#### 09\_Repeated measures

Partial R output:

```
Random effects: Formula: ~1 | cage (Intercept) Residual StdDev: 0.1404056 (= \hat{\nu}) 0.2171559 (= \sqrt{\hat{\sigma}^2 + \hat{\tau}^2}) Correlation Structure: Gaussian spatial correlation Formula: ~as.numeric(month) | cage Parameter estimate(s): range nugget 2.3863954 (= \hat{\rho}^2) 0.2186744 (= \hat{\sigma}^2/(\hat{\sigma}^2 + \hat{\tau}^2)) Number of Observations: 300 Number of Groups: 30
```

Notice the R parametrization of the variance parameters

## Other spatial correlation structures were

• R has a number of build-in correlation structures. A few examples:

correlation=	Name	Correlation term
sp(gau)(t)	Gaussian	$\tau^2 \exp\{\frac{-(t_{i_1}-t_{i_2})^2}{\rho^2}\}$
sp(exp)(t)	exponential	$\tau^2 \exp\{\frac{- t_{i_1} - t_{i_2} }{\rho}\}$
ar(1)	${\sf autoregressive}(1)$	$ au^2 ho^{ i_1-\dot{i}_2 }$
un	unstructured	$ au^2_{i_1,i_2}$

## How to select the correlation structure - Stationarity

- For all the listed covariance structures, the value of  $Cov(Y_t, Y_{t+u})$  does not depend on the time t; only on the time difference u.
- We say that the error proces  $\varepsilon_t$  is weakly stationary of order 2;
- weakly because the stationarity is defined from moments, and not distributions (for Gaussian processes this is the same though);
- of order 2 because the stationarity is defined through 2nd order moments (variance, covariance).

### The semi-variogram

### 09\_Repeated measures

• The semi-variogram plots  $\gamma(u)=\frac{1}{2}V(Y_t-Y_{t+u}),\,u>0$ :

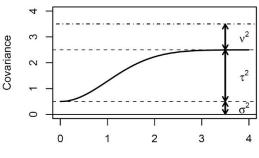
$$\begin{split} \gamma(u) &= \frac{1}{2}V(Y_t - Y_{t+u}) \\ &= \frac{1}{2}\left(V(Y_t) + V(Y_{t+u}) - 2Cov(Y_t, Y_{t+u})\right) \\ &= V(Y_t) - Cov(Y_t, Y_{t+u}) \\ &= \nu^2 + \tau^2 + \sigma^2 - \nu^2 - \tau^2 \lambda(u) \\ &= \sigma^2 + \tau^2(1 - \lambda(u)) \,, \end{split}$$

where  $\lambda(u) = exp(-u^2/\rho^2)$ ,  $exp(-u/\rho)$ ,  $\rho^u$  for spatial Gaussian, exponential and AR1 correlation structures, respectively.

### The semi-variogram

09\_Repeated measures

- Theoretical semi-variogram, Gaussian correlation structure:
- Plot of  $\sigma^2 + \tau^2(1 \lambda(u))$ , where  $\lambda(u) = exp(-u^2/\rho^2)$ :

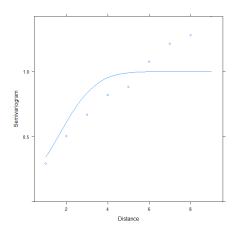


Time difference

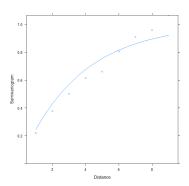
### The semi-variogram

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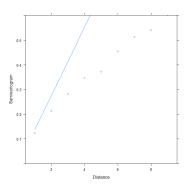
plot(Variogram(analysis,form=~monthQ|cage,data=rats))



## The semi-variogram, expended by decreasing correlation



## The semi-variogram, linearly decreasing correlation



### The semi-variogram, summing un Repeated measures

- The semi-variogram compares a theoretical and empircal correlation structure. Similarity indicates a good model.
- For the rats data, the semi-variogram indicates that the optimal correlation structure seems to be an exponetially decreasing spatial correlation structure, rather than the standard spatial Gaussian structure.

## The rats data, partial final model output

Random effects: Formula: ~1 | cage (Intercept) Residual StdDev:  $0.07828533 \ (=\hat{\nu}) \ 0.2510369 \ (=\sqrt{\hat{\sigma}^2+\hat{\tau}^2})$  Correlation Structure: Exponential spatial correlation Formula: ~as.numeric(month) | cage Parameter estimate(s): range nugget  $3.556503e+00 \ (=\hat{\rho}) \ 3.341370e-08 \ (=\hat{\sigma}^2/(\hat{\sigma}^2+\hat{\tau}^2))$  Number of Observations: 300 Number of Groups: 30

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## Pros and cons of approaches area measures

### Separate analysis for each time-point

- + Not wrong
- Can be confusing
- Difficult to reach coherent conclusion
- In general not very informative

### Analysis of summary statistic

- + Good method with few and easily checked assumptions
- Important to choose good summary measure(s)

### Random effects approach - simple version

- + Good method for short series
- + Uses all observations
- Usually not good for long series

### Random effects approach - qadvanced version

- + Works for short and long series
- + Uses all observations
- Requires appropriate choice of covariance decay function

## Overview of this module\_Repeated measures

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## Exercise: Histamin in degrepeated measures

In an experiment with 16 dogs the blood histamine concentration was measured 0, 1, 3, and 5 minutes after injection of morphine or trimethaphane. Before injection the dogs were classified into two groups according to their level of histamine (intact or depleted).

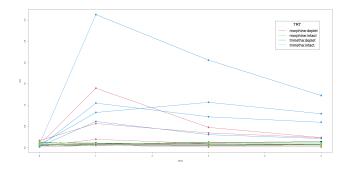
### Histamin in dogs

#### 09\_Repeated measures

```
plot(histamin$minQ,histamin$hist,xlab="minQ",ylab="hist",pch="")
for(i in 1:64){
   temp<-histamin[histamin$dog==i,]
lines(temp$minQ,temp$hist,col=(1:4)[temp$TRT[1]],lwd=2,type="b",
   pch=16,cex=2)
   }
legend(locator(1),
        legend = levels(histamin$TRT), title = "TRT",
        lty = c(1,1,1,1),lwd=2,col= 2:5,cex=2)</pre>
```

## Histamin in dogs

### 09\_Repeated measures



## Histamin in dogs - questings -

First of all use the TRT factor for the analysis (defining 4 groups of dogs).

- Make some plots of the data, for instance one line for each dog (maybe colored differently in each TRT group).
- Analyze these data using one or more of "the simple methods".
- How would you approach making a conclusion for the "real" treatment: morphine vs. trimetha?