

# STA 6326 Homework 1 Solutions

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- 1.1 (a) All 4 bit sequences of 0s and 1s, where 0 represents tails and 1 represents heads, i.e.  $\Omega = \{\omega_1, \omega_2, \omega_3, \omega_4\}$  where  $\omega_i \in \{0, 1\}$ .
- (b) All  $n$  bit sequences of 0s and 1s, where 0 represents undamaged and 1 represents damaged and  $n$  is the number of leaves on the plant. Then the number of damaged leaves is the number of 1 bits in the sequence (a RV).
- (c) All  $x \in (0, \infty)$ .
- (d) All  $x \in (0, k)$  where  $k$  is some reasonable maximum weight for a 10-day old rat (probably like 100 pounds).
- (e) All  $x \in [0, 1]$ .

- 1.2 (a)  $A \setminus (A \cap B) = A \cap (A \cap B)^c = A \cap (A^c \cup B^c) = (A \cap A^c) \cup (A \cap B^c) = \emptyset \cup (A \cap B^c) = A \cap B^c$ .
- (b)  $(B \cap A) \cup (B \cap A^c) = B \cap (A \cup A^c) = B \cap \Omega = B$ .
- (c)  $B \setminus A = B \cap A^c$  by definition.
- (d)  $A \cup (B \cap A^c) = (A \cup B) \cap (A \cup A^c) = (A \cup B) \cap \Omega = A \cup B$ .

- 1.4 (a) Probability of  $A \vee B \iff P(A \cup B) = P(A) + P(B) - P(A \cap B)$ .
- (b) Probability of  $(A \vee B) \wedge \neg(A \wedge B) \iff P(A \cup B) - P(A \cap B)$  and

$$\begin{aligned} P(A \cup B) - P(A \cap B) &= P(A) + P(B) - P(A \cap B) - P(A \cap B) \\ &= P(A) + P(B) - 2 \cdot P(A \cap B) \end{aligned}$$

- (c) Same as (a).
- (d) Same as (b).

1.13  $A \cap B = \emptyset \implies A \subset B^c \implies P(A) \leq P(B^c) \implies 1/3 \leq 1/4$  which is a contradiction.

- 1.14 Each subset involves making  $n$  binary choices, one choice per element in the set  $S$ , of whether to include the element in that subset or exclude that element from the subset. Therefore by **Thm 1.2.14** in Casella there are  $\underbrace{2 \times 2 \times \cdots \times 2}_{n \text{ times}} = 2^n$  different ways to pick which elements to include in a subset.

- 1.16 (a) Choose 3 from 26 with replacement hence  $26^3 = 17576$ .
- (b) Choose 3 from 26 with replacement or (therefore the  $\cup$ ) choose 2 from 26 with replacement hence  $26^3 + 26^2$ .
- (c) Choose 4 from 26 with replacement or (therefore the  $\cup$ ) choose 3 from 26 with replacement or choose 2 from 26 with replacement hence  $26^4 + 26^3 + 26^2$ .

1.18 Imagine laying out the  $n$  balls in a particular order then going through one by one and assigning a cell to each ball. This induces an implicit order on the cell assignments since the balls have a fixed ordering. The sample space is all such assignments, i.e. all  $n$ -tuples with each entry being an integer from 1 to  $n$ . There are  $n^n$  different ways to assign the  $n$  cells to the  $n$  balls. A successful outcome is then one where exactly one cell is assigned twice, i.e. exactly two repetitions in the assignment list. There are  $n$  ways to choose which cell will be assigned twice and  $\binom{n}{2}$  to choose which balls will bear the repetition. Then there are  $n - 1$  ways to choose a cell that will remain empty. Finally there are  $(n - 2)!$  way to assign the remaining  $n - 2$  cells to the  $n - 2$  balls. Hence the probability is  $n\binom{n}{2}(n - 1)(n - 2)!/n^n = \binom{n}{2}n!/n^n$ .

1.21 The “experiment” is choosing  $2r$  shoes from  $2n$  shoes without replacement. There are  $\binom{2n}{2r}$  ways to choose  $2r$  shoes from  $2n$  shoes, i.e.  $n$  pairs of shoes. There are  $\binom{n}{2r}$  different ways to choose  $2r$  different pairs of shoes from  $n$  pairs of shoes, i.e. since we’ll only be choosing one shoe from each pair we must choose which pairs exactly it is we’ll be taking a single shoe from. Then finally there are  $2^{2r}$  different ways to choose either the left shoe or right shoe from each pair of previously  $2r$  chosen shoes. Hence the fraction of choices of  $2r$  shoes from  $2n$  shoes is

$$\frac{\text{number of choices of } 2r \text{ pairs} \times \text{number of ways of choosing either left or right}}{\text{number of ways of choosing } 2r \text{ shoes}} = \frac{\binom{n}{2r} 2^{2r}}{\binom{2n}{2r}}$$

1.22 (a) The “experiment” is to draw 180 days from the year. The sample spaces is all 180-element subsets of the 366 days, so there are  $\binom{366}{180}$  ways to draw 180 days from anywhere in the year for the 180 lottery tickets. In order that the days be evenly distributed there must be  $180/12 = 15$  days chosen from each month. Then there are  $\underbrace{\binom{30}{15} \times \binom{31}{15} \times \cdots \times \binom{29}{15}}_{12} =$

$\binom{30}{15}^4 \times \binom{31}{15}^7 \times \binom{29}{15}$  ways to assign exactly 15 lottery tickets to each of the 12 months. Hence the probability of having all 180 lottery tickets evenly distributed is  $\frac{\binom{30}{15}^4 \times \binom{31}{15}^7 \times \binom{29}{15}}{\binom{366}{180}} = .167 \times 10^{-8}$

(b) The experiment is again to draws days from the year. There are  $\binom{336}{30}$  ways to choose 30 days from all the days in the year that are not in September ( $366 - 30 = 336$ ) and  $\binom{366}{30}$  ways to choose 30 days from any of the days of the year. Hence the probability of not selecting any days from September is  $\binom{336}{30}/\binom{366}{30} = 0.0686905$ .

1.23 Every “draw” has 4 possibilities: {HH, HT, TH, TT}. So there are  $4^n$  different outcomes of the  $n$  “draws”. Then there are  $\binom{n}{i}\binom{n}{i}$  different each person could flip  $i$  heads, for  $i = 0, 1, 2, \dots, n$ . Hence the probability of each person flipping the same number of heads is  $\sum_{i=0}^n \binom{n}{i}^2 / 4^n$ . To see that this equals  $\binom{2n}{n} / 4^n$  we prove Vandermonde’s identity, namely that

$$\binom{m+n}{r} = \sum_{k=0}^r \binom{m}{k} \binom{n}{r-k}$$

To see that this is the case take 2 urns, one with  $m$  red balls and one with  $n$  green balls and consider how many collections of size  $r$  can be drawn from both urns. Certainly it’s  $\binom{m+n}{r}$ . But it’s also the sum, for all  $k$ , of the number of collections with  $k$  red balls and  $n - k$  green balls. For fixed  $k$  this is  $\binom{m}{k}\binom{n}{r-k}$  and thence follows Vandermonde’s identity. Finally letting  $m = r = n$  and using the fact that  $\binom{n}{n-i} = \binom{n}{i}$  we have

$$\frac{1}{4^n} \sum_{i=0}^n \binom{n}{i} \binom{n}{i} = \frac{1}{4^n} \sum_{i=0}^n \binom{n}{i} \binom{n}{n-i} = \frac{1}{4^n} \binom{n+n}{n} = \frac{1}{4^n} \binom{2n}{n}$$

- 1.24 (a) Number of flips until heads is geometrically distributed, i.e. the probability that the  $k$ th flip is a head  $P(X = k) = (1 - p)^{k-1}p$  with  $p = 1/2$ . The probability that  $A$  wins is tantamount to the probability that  $k$  is odd (since  $A$  flips first). Therefore

$$P(A \text{ wins}) = P(X = k = 2i - 1 | i = 1, 2, 3, \dots) = \sum_{i=1}^{\infty} (1 - p)^{(2i-1)-1} p = -\frac{(p-1)^2}{(1-p)^2(p-2)}$$

For  $p = 1/2$  this is  $-\frac{(1/2-1)^2}{(1-1/2)^2(1/2-2)} = \frac{2}{3}$ .

- (b) Simplifying the result from (a) we get that

$$\begin{aligned} P(A \text{ wins}) &= -\frac{(p-1)^2}{(1-p)^2(p-2)} \\ &= -\frac{(1-p)^2}{(1-p)^2(p-2)} \\ &= \frac{1}{2-p} \\ &= \frac{p}{2p-p^2} \\ &= \frac{p}{1-(1-2p+p^2)} \\ &= \frac{p}{1-(1-p)^2} \end{aligned}$$

- (c) Since  $P(A \text{ wins}) = \frac{1}{2-p}$  we see that if  $0 < p < 1$  then  $\frac{1}{2} < P(A \text{ wins}) = \frac{1}{2-p} < 1$ .

- 1.27 (a) Using the binomial theorem

$$0 = (1 - 1)^n = \sum_{k=0}^n \binom{n}{k} 1^{n-k} (-1)^k = \sum_{k=0}^n (-1)^k \binom{n}{k}$$

- (b) Using the binomial theorem

$$\begin{aligned} \sum_{k=1}^n k \binom{n}{k} &= \sum_{k=1}^n k \frac{n!}{k!(n-k)!} \\ &= \sum_{k=1}^n \frac{n!}{(k-1)!(n-k)!} \\ &= \sum_{k=1}^n \frac{n \cdot (n-1)!}{(k-1)!(n-1-(k-1))!} \\ &= n \sum_{k=1}^n \binom{n-1}{k-1} && \text{let } j = k-1 \\ &= n \sum_{j=0}^{n-1} \binom{n-1}{j} \\ &= n(1+1)^{n-1} \\ &= n2^{n-1} \end{aligned}$$

- (c) Using essentially the same calculation (steps 1-4) from part (a) we have that

$$\sum_{k=1}^n (-1)^{k+1} k \binom{n}{k} = n \sum_{k=1}^n (-1)^{k+1} \binom{n-1}{k-1}$$

. Let  $j = k - 1$  then

$$\sum_{k=1}^n (-1)^{k+1} \binom{n-1}{k-1} = (-1)^2 n \sum_{j=0}^{n-1} (-1)^j \binom{n-1}{j} = 0$$

by part (a).

1.33 The probability is

$$\begin{aligned} P(\text{male}|\text{colorblind}) &= \frac{P(\text{colorblind}|\text{male})P(\text{male})}{\sum_{i \in \text{sex}} P(\text{colorblind}|i)P(i)} \\ &= \frac{P(\text{colorblind}|\text{male})P(\text{male})}{P(\text{colorblind}|\text{male})P(\text{male}) + P(\text{colorblind}|\text{female})P(\text{female})} \\ &= \frac{.05 \cdot .5}{.05 \cdot .5 + .25 \cdot .5} \\ &= \frac{1}{6} \end{aligned}$$