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## A PROOFS

**Lemma 1.** Let  $X$  be a non-negative random variable and  $\Phi$  be a continuous function on  $[0, \infty)$ . If  $\Phi'$  is integrable on all closed intervals in  $[0, \infty)$ ,

$$\mathbb{E}[\Phi(X)] = \Phi(0) + \int_0^\infty \Phi'(t)\Pr(X \geq t)dt$$

*Proof.*

$$\begin{aligned} \Phi(0) + \int_0^\infty \Phi'(t)\Pr(X \geq t)dt &= \Phi(0) + \int_0^\infty \int_t^\infty \Phi'(t)p(x)dxdt \\ &= \Phi(0) + \int_{\{x \geq t, t \geq 0\}} \Phi'(t)p(x)d\left(\frac{x}{t}\right) \\ &= \Phi(0) + \int_{\{t \leq x, t \geq 0\}} \Phi'(t)p(x)d\left(\frac{x}{t}\right) \\ &= \Phi(0) + \int_0^\infty \int_0^x \Phi'(t)p(x)dtdx \\ &= \Phi(0) + \int_0^\infty \left( \int_0^x \Phi'(t)dt \right) p(x)dx \\ &= \Phi(0) + \int_0^\infty (\Phi(x) - \Phi(0))p(x)dx \quad (\text{2nd FTC}) \\ &= \Phi(0) + \int_0^\infty \Phi(x)p(x)dx - \int_0^\infty \Phi(0)p(x)dx \\ &= \Phi(0) + \int_0^\infty \Phi(x)p(x)dx - \Phi(0) \int_0^\infty p(x)dx \\ &= \Phi(0) + \mathbb{E}[\Phi(X)] - \Phi(0) \\ &= \mathbb{E}[\Phi(X)] \end{aligned}$$

□

**Lemma 2.** Under the choice of  $\Phi_\tau(\cdot)$  above and its associated  $\Phi'_\tau(\cdot)$ ,

$$\mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} [\Phi_{\tau_i}(d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)))] = \tau_i - \int_{\delta\tau_i}^{\tau_i} \Pr(d(\mathbf{x}_i, T_\theta(\mathbf{z})) < t)dt.$$

*Proof.* By definition,  $\Phi_{\tau_i}(0) = \delta\tau_i$ .

$$\begin{aligned} \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} [\Phi_{\tau_i}(d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)))] &= \Phi_{\tau_i}(0) + \int_0^\infty \Phi'_{\tau_i}(t)\Pr(d(\mathbf{x}_i, T_\theta(\mathbf{z})) \geq t)dt \quad (\text{Lemma 1}) \\ &= \delta\tau_i + \int_{\delta\tau_i}^{\tau_i} \Pr(d(\mathbf{x}_i, T_\theta(\mathbf{z})) \geq t)dt \\ &= \delta\tau_i + \int_{\delta\tau_i}^{\tau_i} (1 - \Pr(d(\mathbf{x}_i, T_\theta(\mathbf{z})) < t)) dt \\ &= \delta\tau_i + (\tau_i - \delta\tau_i) - \int_{\delta\tau_i}^{\tau_i} \Pr(d(\mathbf{x}_i, T_\theta(\mathbf{z})) < t)dt \\ &= \tau_i - \int_{\delta\tau_i}^{\tau_i} \Pr(d(\mathbf{x}_i, T_\theta(\mathbf{z})) < t)dt \end{aligned}$$

□

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**Lemma 3.** Under the choice of  $\Phi_\tau(\cdot)$  above and its associated  $\Phi'_\tau(\cdot)$ ,

$$\mathcal{L}_{\{\tau_i\}_i}(\theta) = \frac{1}{n} \sum_{i=1}^n \frac{1}{mw_i} \sum_{j=1}^m \int_{\delta\tau_i}^{\tau_i} \Pr(d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)) < t) dt.$$

*Proof.*

$$\begin{aligned} \mathcal{L}_{\{\tau_i\}_i}(\theta) &= \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ \frac{1}{n} \sum_{i=1}^n \frac{1}{w_i} \left( \tau_i - \frac{1}{m} \sum_{j=1}^m \Phi_{\tau_i}(d(\mathbf{x}_i, T_\theta(\mathbf{z}_j))) \right) \right] \\ &= \frac{1}{n} \sum_{i=1}^n \frac{1}{w_i} \left( \tau_i - \frac{1}{m} \sum_{j=1}^m \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} [\Phi_{\tau_i}(d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)))] \right) \\ &= \frac{1}{n} \sum_{i=1}^n \frac{1}{w_i} \left( \tau_i - \frac{1}{m} \sum_{j=1}^m \left( \tau_i - \int_{\delta\tau_i}^{\tau_i} \Pr(d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)) < t) dt \right) \right) \quad (\text{Lemma 2}) \\ &= \frac{1}{n} \sum_{i=1}^n \frac{1}{mw_i} \sum_{j=1}^m \int_{\delta\tau_i}^{\tau_i} \Pr(d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)) < t) dt \end{aligned}$$

□

### Equation 3.

*Proof.*

$$\begin{aligned} \mathcal{L}_{\{\tau_i\}_i}(\theta) &= \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ \frac{1}{n} \sum_{i=1}^n \frac{1}{w_i} \left( \tau_i - \frac{m-1}{m} \tau_i - \frac{1}{m} \max(\min_{j \in [m]} d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)), \delta\tau_i) \right) \right] \\ &= \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ \frac{1}{n} \sum_{i=1}^n \frac{1}{w_i} \left( \frac{1}{m} \tau_i - \frac{1}{m} \max(\min_{j \in [m]} d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)), \delta\tau_i) \right) \right] \\ &= \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ \frac{1}{nm} \sum_{i=1}^n \frac{1}{w_i} \left( \tau_i - \max(\min_{j \in [m]} d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)), \delta\tau_i) \right) \right] \end{aligned}$$

□

### Equation 4.

*Proof.*

$$\begin{aligned} \arg \max_{\theta} \mathcal{L}_{\{\tau_i\}_i}(\theta) &= \arg \max_{\theta} \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ \frac{1}{nm} \sum_{i=1}^n \frac{1}{w_i} \left( \tau_i - \max(\min_{j \in [m]} d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)), \delta\tau_i) \right) \right] \\ &= \arg \max_{\theta} \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ \sum_{i=1}^n \frac{1}{w_i} \left( \tau_i - \max(\min_{j \in [m]} d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)), \delta\tau_i) \right) \right] \\ &= \arg \max_{\theta} \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ \sum_{i=1}^n \frac{\tau_i}{w_i} - \sum_{i=1}^n \frac{1}{w_i} \max(\min_{j \in [m]} d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)), \delta\tau_i) \right] \\ &= \arg \max_{\theta} \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ - \sum_{i=1}^n \frac{1}{w_i} \max(\min_{j \in [m]} d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)), \delta\tau_i) \right] \\ &= \arg \min_{\theta} \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ \sum_{i=1}^n \frac{1}{w_i} \max(\min_{j \in [m]} d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)), \delta\tau_i) \right] \end{aligned}$$

□

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**Lemma 4.** Suppose  $p_\theta$  is continuous at all data points  $\mathbf{x}_1, \dots, \mathbf{x}_n$ , under the choice of  $w_i = \int_{\delta\tau_i}^{\tau_i} \text{vol}(B_t(\mathbf{x}_i))dt := \int_{\delta\tau_i}^{\tau_i} \int_{B_t(\mathbf{x}_i)} d\mathbf{x} dt$ , where  $B_r(\mathbf{x}) = \{\mathbf{y} | d(\mathbf{y}, \mathbf{x}) < r\}$  is an open ball of radius  $r$  centred at  $\mathbf{x}$ ,

$$\lim_{\{\tau_i \rightarrow 0^+\}_i} \mathcal{L}_{\{\tau_i\}_i}(\theta) = \frac{1}{n} \sum_{i=1}^n p_\theta(\mathbf{x}_i)$$

*Proof.*

$$\begin{aligned} \mathcal{L}_{\{\tau_i\}_i}(\theta) &= \frac{1}{n} \sum_{i=1}^n \frac{1}{mw_i} \sum_{j=1}^m \int_{\delta\tau_i}^{\tau_i} \Pr(d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)) < t) dt \quad (\text{Lemma 3}) \\ &= \frac{1}{n} \sum_{i=1}^n \frac{1}{mw_i} \sum_{j=1}^m \int_{\delta\tau_i}^{\tau_i} \int_{B_t(\mathbf{x}_i)} p_\theta(\mathbf{x}) d\mathbf{x} dt \\ &= \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m \frac{1}{w_i} \int_{\delta\tau_i}^{\tau_i} \int_{B_t(\mathbf{x}_i)} p_\theta(\mathbf{x}) d\mathbf{x} dt \\ &= \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m \frac{\int_{\delta\tau_i}^{\tau_i} \int_{B_t(\mathbf{x}_i)} p_\theta(\mathbf{x}) d\mathbf{x} dt}{\int_{\delta\tau_i}^{\tau_i} \int_{B_t(\mathbf{x}_i)} d\mathbf{x} dt} \end{aligned}$$

$$\begin{aligned} \lim_{\{\tau_i \rightarrow 0^+\}_i} \mathcal{L}_{\{\tau_i\}_i}(\theta) &= \frac{1}{nm} \sum_{i=1}^n \left( \lim_{\tau_i \rightarrow 0^+} \left( \sum_{j=1}^m \frac{\int_{\delta\tau_i}^{\tau_i} \int_{B_t(\mathbf{x}_i)} p_\theta(\mathbf{x}) d\mathbf{x} dt}{\int_{\delta\tau_i}^{\tau_i} \int_{B_t(\mathbf{x}_i)} d\mathbf{x} dt} \right) \right) \\ &= \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m \left( \lim_{\tau_i \rightarrow 0^+} \frac{\int_{\delta\tau_i}^{\tau_i} \int_{B_t(\mathbf{x}_i)} p_\theta(\mathbf{x}) d\mathbf{x} dt}{\int_{\delta\tau_i}^{\tau_i} \int_{B_t(\mathbf{x}_i)} d\mathbf{x} dt} \right) \\ &= \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m \left( \lim_{\tau_i \rightarrow 0^+} \frac{\int_{B_{\tau_i}(\mathbf{x}_i)} p_\theta(\mathbf{x}) d\mathbf{x} - \delta \int_{B_{\delta\tau_i}(\mathbf{x}_i)} p_\theta(\mathbf{x}) d\mathbf{x}}{\int_{B_{\tau_i}(\mathbf{x}_i)} d\mathbf{x} - \delta \int_{B_{\delta\tau_i}(\mathbf{x}_i)} d\mathbf{x}} \right) \quad (\text{L'Hôpital and 2nd FTC}) \\ &= \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m \left( \lim_{\tau_i \rightarrow 0^+} \frac{\int_{B_{\tau_i}(\mathbf{x}_i)} p_\theta(\mathbf{x})(1 - \delta \mathbf{1}_{B_{\delta\tau_i}(\mathbf{x}_i)}(\mathbf{x})) d\mathbf{x}}{\int_{B_{\tau_i}(\mathbf{x}_i)} 1 - \delta \mathbf{1}_{B_{\delta\tau_i}(\mathbf{x}_i)}(\mathbf{x}) d\mathbf{x}} \right) \\ &= \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m \left( \lim_{\tau_i \rightarrow 0^+} \frac{\int_0^{\tau_i} (1 - \delta \mathbf{1}_{\{r < \delta\tau_i\}}(r)) \int_{\{\mathbf{x} | d(\mathbf{x}, \mathbf{x}_i) = r\}} p_\theta(\mathbf{x}) d\mathbf{x} dr}{\int_0^{\tau_i} (1 - \delta \mathbf{1}_{\{r < \delta\tau_i\}}(r)) \int_{\{\mathbf{x} | d(\mathbf{x}, \mathbf{x}_i) = r\}} d\mathbf{x} dr} \right) \\ &= \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m \left( \lim_{\tau_i \rightarrow 0^+} \frac{\int_{\{\mathbf{x} | d(\mathbf{x}, \mathbf{x}_i) = \tau_i\}} p_\theta(\mathbf{x}) d\mathbf{x}}{\int_{\{\mathbf{x} | d(\mathbf{x}, \mathbf{x}_i) = \tau_i\}} d\mathbf{x}} \right) \quad (\text{L'Hôpital and 2nd FTC}) \\ &= \frac{1}{nm} \sum_{i=1}^n \sum_{j=1}^m p_\theta(\mathbf{x}_i) \quad (\text{Continuity of } p_\theta) \\ &= \frac{1}{n} \sum_{i=1}^n p_\theta(\mathbf{x}_i) \end{aligned}$$

□

Note that under common metrics like  $\ell_p$  distances,  $w_i$  can be found in closed form, i.e.,  $\text{vol}(B_t(\mathbf{x}_i)) = (2t)^d \frac{\Gamma(1+1/p)^d}{\Gamma(1+d/p)}$ , and so  $w_i = \int_{\delta\tau_i}^{\tau_i} \text{vol}(B_t(\mathbf{x}_i))dt = \int_{\delta\tau_i}^{\tau_i} (2t)^d \frac{\Gamma(1+1/p)^d}{\Gamma(1+d/p)} dt = \frac{(2(1-\delta)\tau_i)^{d+1}}{2(d+1)} \cdot \frac{\Gamma(1+1/p)^d}{\Gamma(1+d/p)}$ , where  $\Gamma(\cdot)$  denotes the gamma function.

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## B PSEUDO CODE FOR IMLE

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**Algorithm 1** Implicit maximum likelihood estimation (IMLE) procedure

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**Require:** The set of inputs  $\{\mathbf{x}_i\}_{i=1}^n$

- 1: Initialize the parameters  $\theta$  of the generator  $T_\theta$
- 2: **for**  $k = 1$  to  $K$  **do**
- 3:     Pick a random batch  $S \subseteq [n]$
- 4:     Draw latent codes  $Z \leftarrow \mathbf{z}_1, \dots, \mathbf{z}_m$  from  $\mathcal{N}(0, \mathbf{I})$
- 5:      $\sigma(i) \leftarrow \arg \min_{j \in [m]} d(\mathbf{x}_i, T_\theta(\mathbf{z}_j)) \forall i \in S$
- 6:     **for**  $l = 1$  to  $L$  **do**
- 7:         Pick a random mini batch  $\tilde{S} \subseteq S$
- 8:          $\theta \leftarrow \theta - \eta \nabla_\theta (\sum_{i \in \tilde{S}} d(\mathbf{x}_i, T_\theta(\mathbf{z}_{\sigma(i)}))) / |\tilde{S}|$
- 9:     **end for**
- 10: **end for**
- 11: **return**  $\theta$

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## C ADDITIONAL RESULTS

We show more interpolation results for Adaptive IMLE on FFHQ subset and Obama in Fig. 1. We also show randomly generated samples for FFHQ subset and Obama in Fig. 3 and 2.

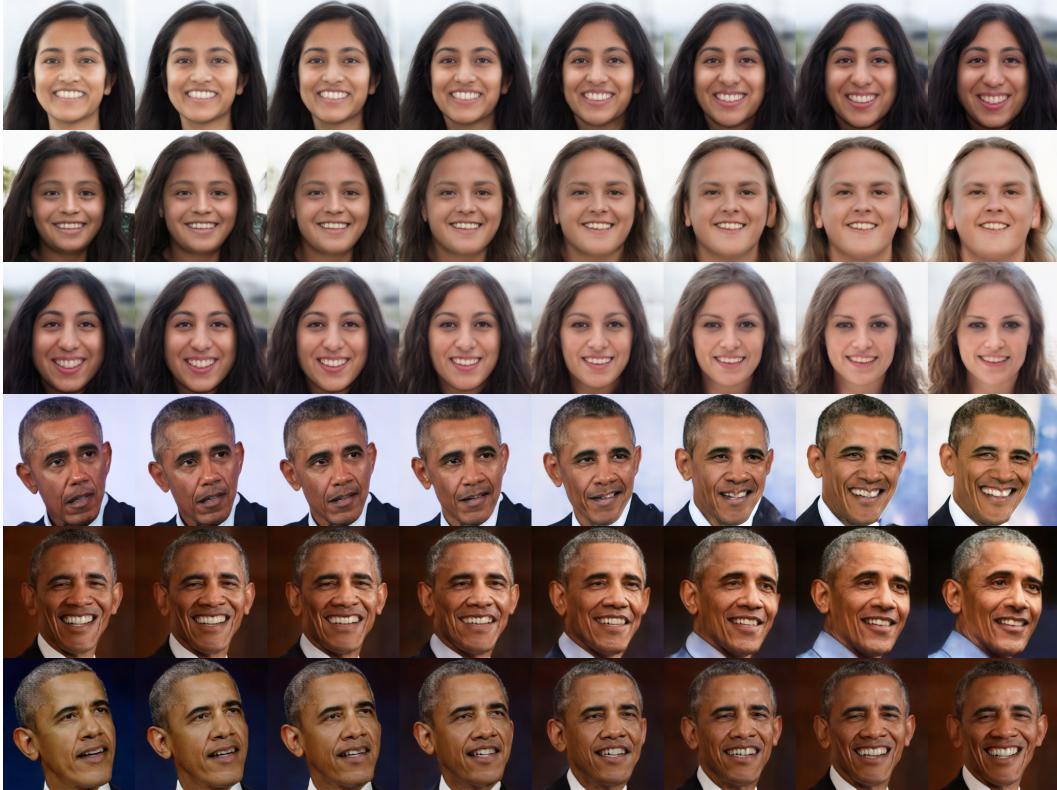


Figure 1: Interpolation results for Adaptive IMLE (Ours). Each row shows a different interpolation.



Figure 2: Adaptive IMLE (Ours): *FFHQ subset* randomly generated samples.

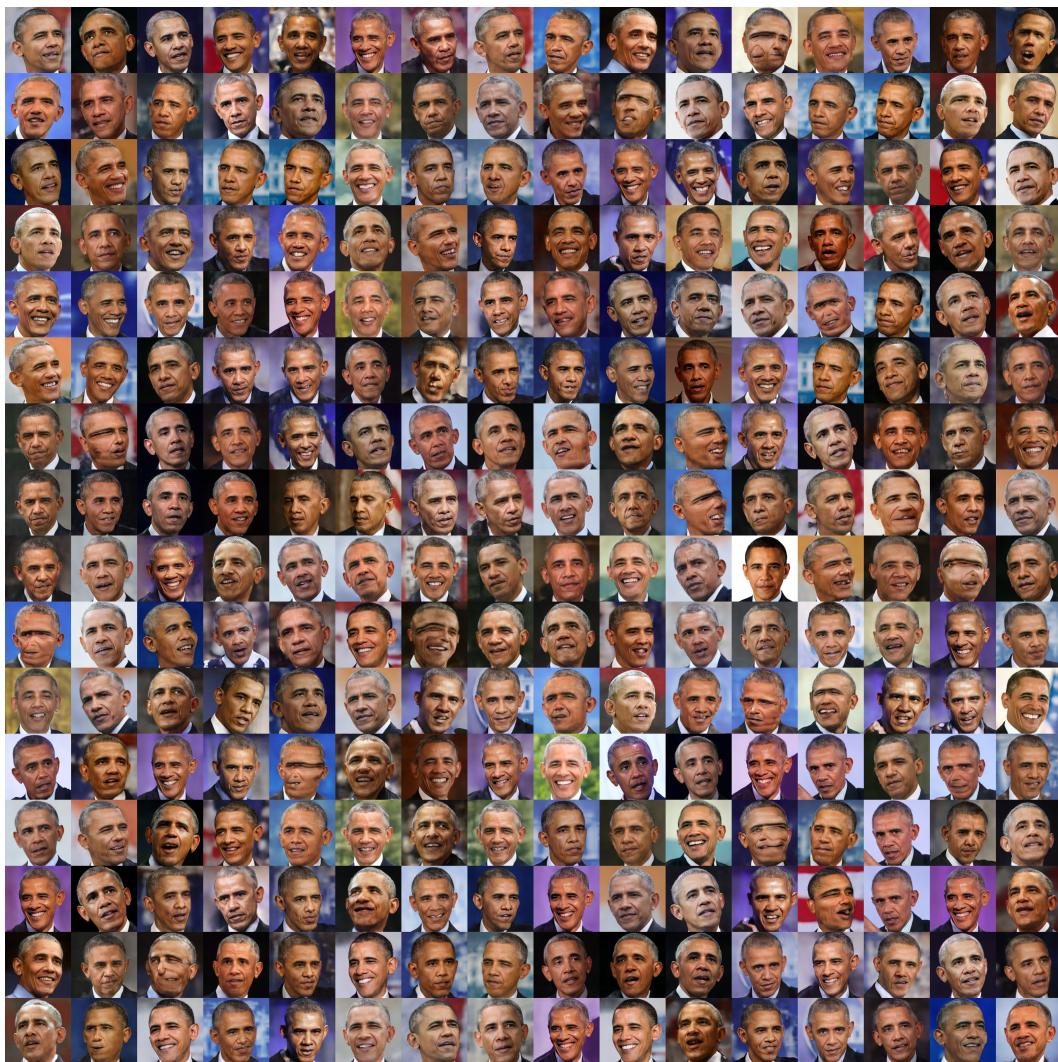


Figure 3: Adaptive IMLE (Ours): *Obama* randomly generated samples.

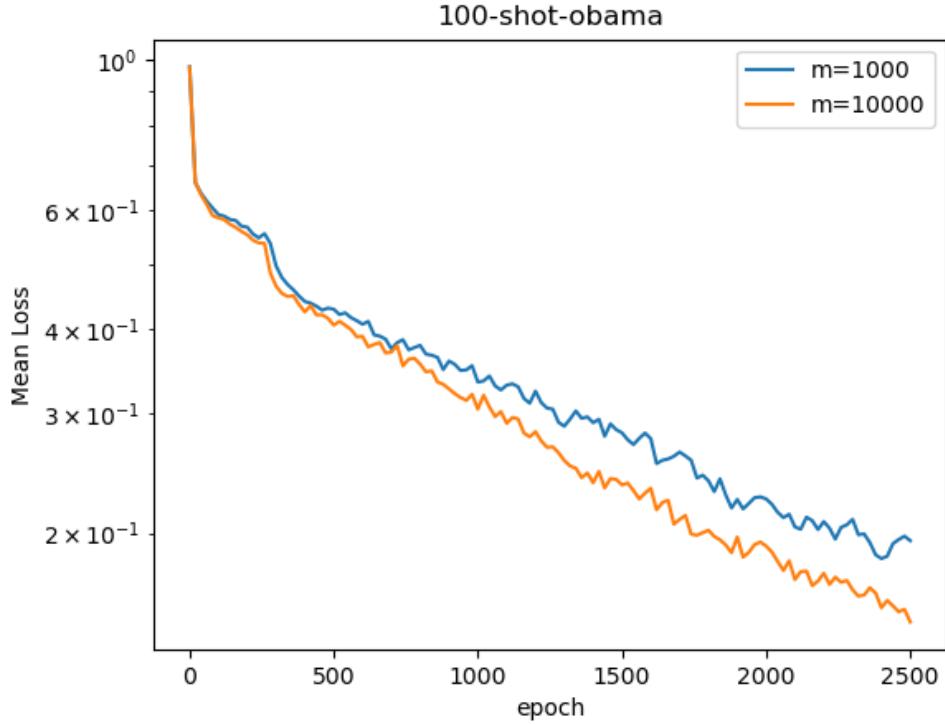


Figure 4: Average loss across different values of  $m$ , i.e., sample size. We observe that larger  $m$  results in faster convergence but slower nearest neighbour search.

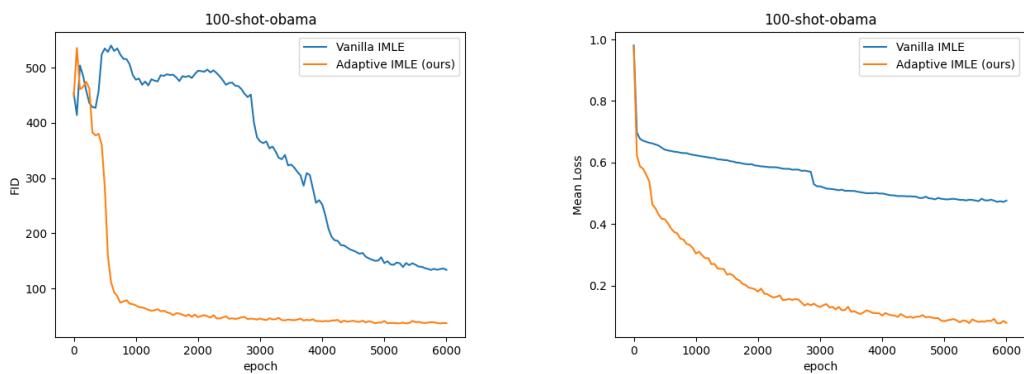


Figure 5: Comparison of Adaptive IMLE and Vanilla IMLE during training for the Obama dataset. On left, FID during training is shown. On the right, average loss during the training is shown.

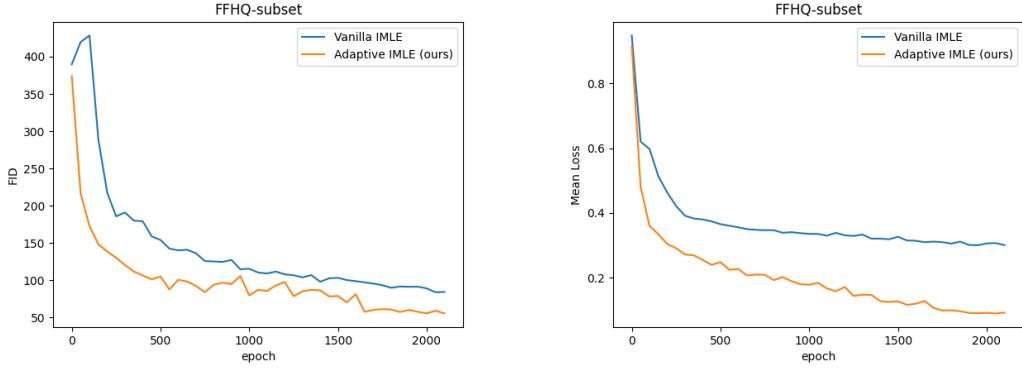


Figure 6: Comparison of Adaptive IMLE and Vanilla IMLE during training for the FFHQ subset dataset. On left, FID during training is shown. On the right, average loss during the training is shown.

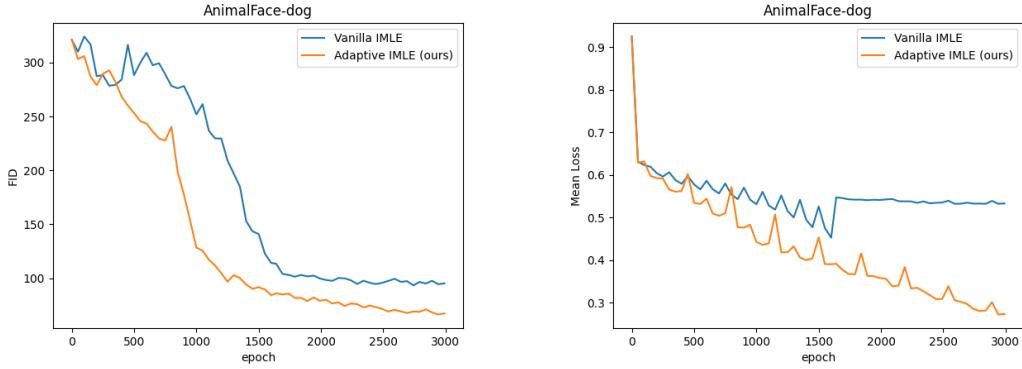


Figure 7: Comparison of Adaptive IMLE and Vanilla IMLE during training for the Dog dataset. On left, FID during training is shown. On the right, average loss during the training is shown.

## D EQUIVALENCE OF OPTIMIZATION FOR WEIGHTED AND UNWEIGHTED OBJECTIVES

Consider the weighted objective:

$$\arg \max_{\theta} \mathcal{L}_{\{\tau_i^k\}_i}(\theta) = \arg \min_{\theta} \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ \sum_{i=1}^n \frac{1}{w_i} \max \left( \min_{j \in [m]} d(\mathbf{x}_i, T_{\theta}(\mathbf{z}_j)), \delta \tau_i^k \right) \right] \quad (1)$$

Let's consider the sequence of optimization problems in the curriculum. Each optimization problem uses different values for  $\{\tau_i\}_i$ , which denote the radii of the neighbourhoods around each  $\mathbf{x}_i$ . We denote  $k$ th optimization problem as  $\tilde{\mathcal{L}}_{\{\tau_i^k\}_i}$  and the neighbourhood radii values it uses as  $\{\tau_i^k\}_i$ . Because  $0 < \delta < 1$  and  $\tau_i^k = \delta \tau_i^{k-1}$ ,  $\lim_{k \rightarrow \infty} \tau_i^k = 0 \forall i \in [n]$ . Therefore,  $\forall i \in [n] \forall \epsilon_i > 0, \exists N_i$  s.t.  $\forall N'_i > N_i$  we have  $\tau_i^{N'_i} < \epsilon$ . Define  $N = \max_i N_i$ . Then  $\forall \epsilon_i > 0, \forall N' > N$  we have  $\tau_i^{N'} < \epsilon \forall i \in [n]$ .

Recall that  $\tau_i^{k-1}$  is the minimum distance between  $\mathbf{x}_i$  and its nearest sample at the start of optimizing  $\tilde{\mathcal{L}}_{\{\tau_i^k\}_i}$ . Consider the setting of the model parameters at the start of optimizing  $\tilde{\mathcal{L}}_{\{\tau_i^k\}_i}$ , which we will call  $\epsilon_0 > 0$ . Then,  $\tau_i^{k-1} = \min_{j \in [m]} d(\mathbf{x}_i, T_{\theta_0}(\mathbf{z}_j)) < \epsilon_0 \forall i \in [n]$ , i.e., the distance between each data example  $x_i$  and its nearest sample is smaller than  $\epsilon_0$ .

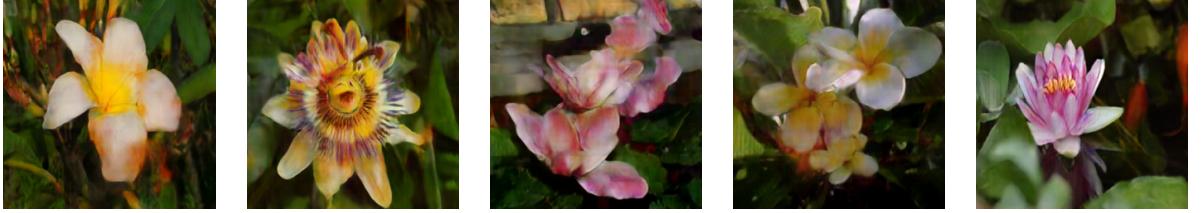


Figure 8: Generated samples for 10-shot Oxford Flowers dataset. The model is trained on 10 random images of Oxford Flowers dataset. This both illustrates qualitative results of our method for a non-facial dataset and also with fewer data examples, i.e., 10.

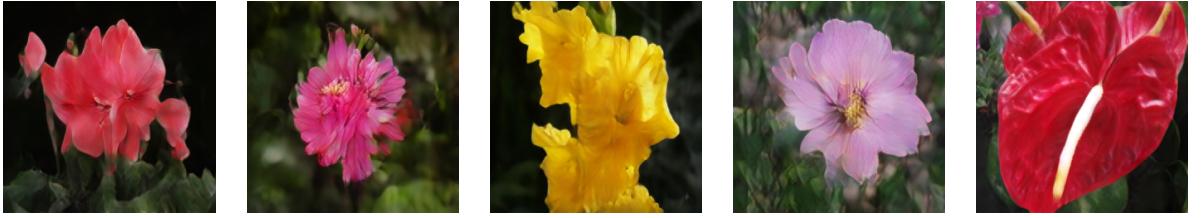


Figure 9: Generated samples for 100-shot Oxford Flowers dataset. The model is trained on 100 random images of Oxford Flowers dataset. This illustrates qualitative results of our method for a non-facial dataset.

Now let's consider an unweighted optimization problem  $k'$  such that  $\delta\tau_i^{k'} > \epsilon_0 \forall i \in [n]$ :

$$\tilde{\mathcal{L}}_{\{\tau_i^k\}_i} = \arg \min_{\theta} \mathbb{E}_{z_1, \dots, z_m \sim \mathcal{N}(0, I)} \left[ \sum_{i=1}^n \max(d(\mathbf{x}_i, T_\theta(\mathbf{z}_{s_i})), \delta\tau_i^{k'}) \right] \quad (2)$$

where  $s_i = \arg \min_{j \in [m]} d(\mathbf{x}_i, T_\theta(\mathbf{z}_j))$ , i.e., the index of the closest generated sample to  $\mathbf{x}_i$ . As established above, we know for  $\theta_0$  we have  $d(\mathbf{x}_i, T_{\theta_0}(\mathbf{z}_{s_i})) < \epsilon_0$ , by transitivity,  $d(\mathbf{x}_i, T_{\theta_0}(\mathbf{z}_{s_i})) < \delta\tau_i^k$  and therefore the optimal objective value for  $\mathcal{L}_{\{\tau_i^k\}_i}$  is  $\sum_{i=1}^n \frac{1}{w_i} \delta\tau_i^{k'}$ . Now if we consider the set of parameters that achieve this optimal objective value, because  $w_i$ 's are strictly positive, this set is also a solution to the unweighted objective, i.e.,  $\sum_{i=1}^n \delta\tau_i^k$ . Because this is true for any optimization problem  $\tilde{\mathcal{L}}_{\{\tau_i^{k'}\}_i}$ , it means that the set of optimizers for the weighted and unweighted objectives are equivalent and hence optimizing the unweighted objective is equivalent to optimizing the weighted objective.