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S E M I N A R A R B E I T

Using PCA on EEG Data to Distinguish Sleep Stages

ausgeführt am

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1 Introduction

[TODO general introduction]

1.1 EEG Data and Sleep Stages

Ganong [Gan97] describes typical patterns observed in electroencephalogram (EEG) data of a sleeping person. He describes the EEG patterns associated with rapid eye movement (REM) sleep and non-REM (NREM) sleep.

NREM sleep is further partitioned into three (although some use four) stages, termed Stage 1 (S1) to Stage 3 (S3). Example EEG data of these different sleep stages can be seen in Figure ?? [TODO image]. The EEG data of these stages is characterized by

S1: low amplitude, high frequency

S2: appearance of sleep spindles (bursts of higher amplitude, lower frequency waves)

S3: high amplitude, low frequency

In REM sleep the EEG data is that of high frequency and low amplitude patterns, resembling the data observed in alert humans.

2 Study of Literature

A substantial body of scientific research has been devoted to exploring Principal Component Analysis (PCA). The foundation of this method was laid by Pearson [Pea01] and Hotelling [Hot33].

An introduction to PCA, as well as a good overview on how to derive the formula used to compute the Principal Components (PC) is given by Shlens [Shl14]. Recent applications and variants of PCA are explored by Jolliffe et. al. [JC16].

Shlens discusses the limitations of PCA, as well as examples in which PCA fails, such as the requirement of linearly dependent data. Tenenbaum proposes a non-linear method to combat this problem [TDSL00].

Generally speaking the variables must not have third or higher order dependencies¹ between them. In some cases it is possible to reduce a problem with higher order dependencies to a second order one by applying a non-linear transformation beforehand. This method is called kernel PCA [SSM97].

Another method for dealing with this problem is Independent Component Analysis (ICA) which is discussed by Naik et. al. [NK11].

The given problem of distinguishing sleep stages given some EEG data has been investigated by use of PCA, as well as neural networks. Some of these works are summarized below.

A review of different methods in the preprocessing, feature extraction and classification is given by Boostani et. al. [BKN17]. They find that using a random forest classifier [Bre01] and entropy of wavelet coefficients [CMP94] as feature gives the best results.

Tăuțan et. al. [TRdFI21] compare different methods of dimensionality reduction on EEG data, such as PCA, factor analysis and autoencoders. They conclude that the use of PCA and factor analysis improves the accuracy of the model.

Putilov [Put15] used PCA to find boundaries between Stage 1, Stage 2 and Stage 3. Changes in the first two PC were related to changes between the Stage 1 and Stage 2, while changes in the fourth PC exhibited a change in sign at the boundary of Stage 2 and Stage 3. This suggests that changes between Stage 1 and Stage 2 are easier to detect than ones between Stage 2 and Stage 3.

Metzner et. al. [MST⁺23] try to rediscover the different human-defined sleep stages. They find that using PCA on the results makes clusters apparent. These clusters could then be used as a basis for a redefinition of sleep stages.

The PhysioNet/Computing in Cardiology Challenge 2018 [GMwHL⁺18] was a competition using a similar dataset. The goal was to identify arousal during sleep from EEG, EOG, EMG, ECG and SaO2 data given. The winning paper [HPPB18] of this competition describes the use of a dense recurrent convolutional neural network (DRCNN) comprised

¹e.g. $\mathbb{E}[x_i x_j x_k] \neq 0$ for some i, j, k assuming mean-free variables

of multiple dense convolutional layers, a bidirectional long-short term memory layer and a softmax output layer.

As shown in this section, the utilization of PCA to analyze EEG data has been used with success.

3 Mathematical Basics

We define mathematical notation, which will be used in Section 4 to define PCA.

3.1 Covariance

Assume we have two sets of n observations of variables with mean 0. Let us call the first list of observations $\mathbf{a} = (a_1, \dots, a_n)$ and the second $\mathbf{b} = (b_1, \dots, b_n)$.

Definition 1 (Covariance). *Let us define the covariance of $\mathbf{a} \in \mathbb{R}^n$ and $\mathbf{b} \in \mathbb{R}^n$ as*

$$\sigma_{\mathbf{ab}} := \frac{1}{n} \sum_{i=1}^n a_i b_i = \frac{1}{n} \mathbf{a} \cdot \mathbf{b}^T.$$

From the definition it is obvious that the covariance is symmetric, $\sigma_{\mathbf{ab}} = \sigma_{\mathbf{ba}}$. In the special case $\mathbf{a} = \mathbf{b}$ the covariance $\sigma_{\mathbf{aa}}$ is called *variance* $\sigma_{\mathbf{a}}^2$.

Definition 2 (Covariance Matrix). *Generalizing to m variables $\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_m)$, each having been observed n times, gives us the covariance matrix.*

$$\mathbf{C}_{\mathbf{X}} := \begin{pmatrix} \sigma_{\mathbf{x}_1 \mathbf{x}_1} & \cdots & \sigma_{\mathbf{x}_1 \mathbf{x}_m} \\ \vdots & \ddots & \vdots \\ \sigma_{\mathbf{x}_m \mathbf{x}_1} & \cdots & \sigma_{\mathbf{x}_m \mathbf{x}_m} \end{pmatrix} = \frac{1}{n} \mathbf{X} \mathbf{X}^T$$

The covariance matrix is a symmetric $m \times m$ matrix.

3.2 Diagonalizable Matrix

Definition 3 (Diagonalizable Matrix). *A square matrix \mathbf{A} is called diagonalizable, if there exists an invertible matrix \mathbf{P} and a diagonal matrix \mathbf{D} such that $\mathbf{A} = \mathbf{P} \mathbf{D} \mathbf{P}^{-1}$.*

Definition 4 (Symmetric matrix). *A square matrix \mathbf{A} is called symmetric, if $\mathbf{A}^T = \mathbf{A}$.*

Theorem 1. *Every symmetric matrix is diagonalizable.*

This is the main theorem we need in order to derive PCA. The proof of this theorem requires some preparation, which we will do now.

Definition 5 (Eigenvalues and Eigenvectors). *Let \mathbf{A} be a real $m \times m$ matrix. $\lambda \in \mathbb{C}$ is called a eigenvalue with eigenvector $\mathbf{v} \in \mathbb{C}^m \setminus \{\mathbf{0}\}$ if*

$$\mathbf{A} \mathbf{v} = \lambda \mathbf{v}. \tag{3.1}$$

Lemma 1. *Every square $m \times m$ matrix has m (not necessarily unique) eigenvalues.*

Proof. We can rewrite equation 3.1 as

$$(\mathbf{A} - \lambda \mathbf{I})\mathbf{v} = \mathbf{0}$$

This allows us to interpret $(\mathbf{A} - \lambda \mathbf{I})$ as a function, which takes vectors $\mathbf{v} \in \mathbb{C}^m$. For λ to be an eigenvalue of \mathbf{A} with eigenvector \mathbf{v} it has to satisfy $\mathbf{v} \in \ker(\mathbf{A} - \lambda \mathbf{I})$ and $\mathbf{v} \neq \mathbf{0}$. From this we gather that all λ with $\ker(\mathbf{A} - \lambda \mathbf{I}) \neq \{\mathbf{0}\}$ are eigenvalues. We know this holds if and only if $\det(\mathbf{A} - \lambda \mathbf{I}) = 0$. The determinant is a polynomial of degree m which can be expressed in the form $(\lambda - \lambda_1) \dots (\lambda - \lambda_m)$ with $\lambda_1, \dots, \lambda_m \in \mathbb{C}$. These $\lambda_1, \dots, \lambda_m$ are the m eigenvalues we wanted to find. \square

Lemma 2. *A symmetric matrix has real eigenvalues.*

Proof. Let $\bar{\cdot}$ denote the complex conjugate. Define a complex dot product

$$(\mathbf{u}, \mathbf{v}) := \sum_{i=1}^m u_i \bar{v}_i$$

This dot product has the following properties for all $\mathbf{A} \in \mathbb{C}^{m \times m}$, $\mathbf{u}, \mathbf{v} \in \mathbb{C}^m$, $\lambda \in \mathbb{C}$

- $(\mathbf{A}\mathbf{u}, \mathbf{v}) = (\mathbf{u}, \mathbf{A}^T \mathbf{v})$,
- $(\lambda \mathbf{u}, \mathbf{v}) = \lambda(\mathbf{u}, \mathbf{v})$,
- $(\mathbf{u}, \lambda \mathbf{v}) = \bar{\lambda}(\mathbf{u}, \mathbf{v})$
- $(\mathbf{u}, \mathbf{u}) = 0 \iff \mathbf{u} = \mathbf{0}$

Let \mathbf{A} be a symmetric matrix with eigenvalue $\lambda \in \mathbb{C}$.

For all $\mathbf{u} \in \mathbb{C}^m$ we have

$$\begin{aligned} \lambda(\mathbf{u}, \mathbf{u}) &= (\lambda \mathbf{u}, \mathbf{u}) = (\mathbf{A}\mathbf{u}, \mathbf{u}) = (\mathbf{u}, \mathbf{A}^T \mathbf{u}) = \\ &= (\mathbf{u}, \mathbf{A}\mathbf{u}) = (\mathbf{u}, \lambda \mathbf{u}) = \bar{\lambda}(\mathbf{u}, \mathbf{u}). \end{aligned}$$

For $\mathbf{u} \neq \mathbf{0}$ we get $\lambda = \bar{\lambda}$ and thus $\lambda \in \mathbb{R}$. \square

Are the corresponding eigenvectors real? From the proof of lemma 1 we know that the eigenvector \mathbf{v} of eigenvalue λ is in $\ker(\mathbf{A} - \lambda \mathbf{I})$. Both the matrix \mathbf{A} and λ are real, so \mathbf{v} must be in \mathbb{R}^m as well.

Lemma 3. *The eigenvectors of a symmetric matrix with distinct eigenvalues are orthogonal.*

Proof. Let λ_1, λ_2 be two distinct eigenvalues with eigenvectors $\mathbf{v}_1, \mathbf{v}_2$ of the matrix \mathbf{A} .

$$\begin{aligned} \lambda_1 \mathbf{v}_1 \cdot \mathbf{v}_2 &= (\lambda_1 \mathbf{v}_1)^T \mathbf{v}_2 = (\mathbf{A} \mathbf{v}_1)^T \mathbf{v}_2 = \mathbf{v}_1^T \mathbf{A}^T \mathbf{v}_2 = \\ &= \mathbf{v}_1^T \mathbf{A} \mathbf{v}_2 = \mathbf{v}_1^T (\lambda_2 \mathbf{v}_2) = \lambda_2 \mathbf{v}_1 \cdot \mathbf{v}_2 \end{aligned}$$

This shows $(\lambda_1 - \lambda_2) \mathbf{v}_1 \cdot \mathbf{v}_2 = 0$ and as λ_1 and λ_2 are distinct, \mathbf{v}_1 and \mathbf{v}_2 must be orthogonal. \square

What if the eigenvalues of the matrix are not distinct? In the proof of lemma 1 we showed that every $\mathbf{v} \in \ker(\mathbf{A} - \lambda\mathbf{I}) \setminus \{\mathbf{0}\}$ is a eigenvector. If and only if $(\lambda - \lambda_i)$ appears $k \geq 2$ times in the determinant of $(\mathbf{A} - \lambda\mathbf{I})$ then \mathbf{A} has a non unique eigenvalue λ_i . As $\dim(\ker(\mathbf{A} - \lambda_i\mathbf{I})) = k$ we can choose orthogonal eigenvectors.

Now we have everything we need to prove theorem 1.

Proof of Theorem 1. Let $\mathbf{A} \in \mathbb{R}^{m \times m}$ be a symmetric matrix. From lemma 1 we know that eigenvalues $\lambda_1, \dots, \lambda_m$ with corresponding eigenvectors $\mathbf{v}_1, \dots, \mathbf{v}_m$ exist.

Define the following matrices

$$\mathbf{D} := \begin{pmatrix} \lambda_1 & 0 & \cdots & 0 \\ 0 & \lambda_2 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \lambda_m \end{pmatrix} \quad \mathbf{V} := \begin{pmatrix} \mathbf{v}_1 & \mathbf{v}_2 & \cdots & \mathbf{v}_m \end{pmatrix}$$

The definition of eigenvalues and eigenvectors gives us

$$\mathbf{A}\mathbf{V} = (\mathbf{A}\mathbf{v}_1 \quad \cdots \quad \mathbf{A}\mathbf{v}_m) = (\lambda_1\mathbf{v}_1 \quad \cdots \quad \lambda_m\mathbf{v}_m) = \mathbf{V}\mathbf{D}. \quad (3.2)$$

From lemma 3 we know that the eigenvectors, and therefore the columns of \mathbf{V} , are orthogonal. It follows that $\text{rank}(\mathbf{V}) = m$ which gives us the existence of \mathbf{V}^{-1} .

Rearranging equation 3.2 now gives us $\mathbf{A} = \mathbf{V}\mathbf{D}\mathbf{V}^{-1}$ which is what we wanted to show.

This shows that \mathbf{A} is diagonalizable. \square

Lemma 4. *If the columns of matrix \mathbf{A} are orthonormal, then $\mathbf{A}^{-1} = \mathbf{A}^T$.*

Proof. Let $(\mathbf{a}_i)_{i=1, \dots, m}$ be the columns of the matrix. The columns are orthogonal and normed, therefore

$$\forall i, j : \mathbf{a}_i^T \mathbf{a}_j = \begin{cases} 1 & \text{if } i = j \\ 0 & \text{otherwise} \end{cases} \implies \mathbf{A}^T \mathbf{A} = \mathbf{I}$$

This shows $\mathbf{A}^{-1} = \mathbf{A}^T$. \square

4 Principal Component Analysis

Combining the concepts in section 3 we derive the ideas and implementation of PCA.

Assume we have gathered observations of different variables as part of an experiment. If we have n variables, each having been observed m times, we can create a $m \times n$ matrix of this data. The goal is to get more insight and find underlying patterns in the collected data. For $n = 2$ we could try to plot the data, with the first variable as the x -axis and the second as the y axis. An exemplary plot of some data can be seen in figure 4.1.

For larger values of n this gets increasingly difficult¹. PCA tries to solve this problem by transforming the data in such a way that the most interesting features are in the first few axis of the transformed m dimensional space. This makes it easy to look at a low dimension representation of the data, without losing much information.

An example of PCA being applied to the data from figure 4.1 can be seen in figure 4.2. In the top figure the normed data and the direction of the new axis (called Principal Components (PC)) in relation to the two original axis are shown. The bottom figure depicts the transformed data. One can see that the variance is maximal in the PC 1 axis.

Now we derive how to compute PCA. First we formulate a goal and define some assumptions.

We assume that the most interesting features are those that have a large variance². Our goal is to find a transformation into new coordinates such that:

- the variance in the each axis is as large as possible.
- the axis are all orthogonal to each other.
- the axis are sorted (descending) by the variance in the axis.

From this we gather that another assumption is, that the axis are orthogonal. Lastly we are only concerned with linear dependent features in the data. Some example cases in which PCA fails are shown in figure 4.3.

One way to achieve the goal is as follows:

1. Find the direction which maximizes the variance.
2. Save this direction as the next axis.
3. Determine the subspace that is orthogonal to all axis we found so far.
4. If the subspace is non-trivial start at the first step again.

¹For higher dimensionality we have to use some projection. Depending on the chosen projection the interpretation changes making it difficult to interpret the resulting image.

²This assumption can be false. For data where the noise has a larger variance than the feature we are trying to observe, PCA fails because this assumption is not met.

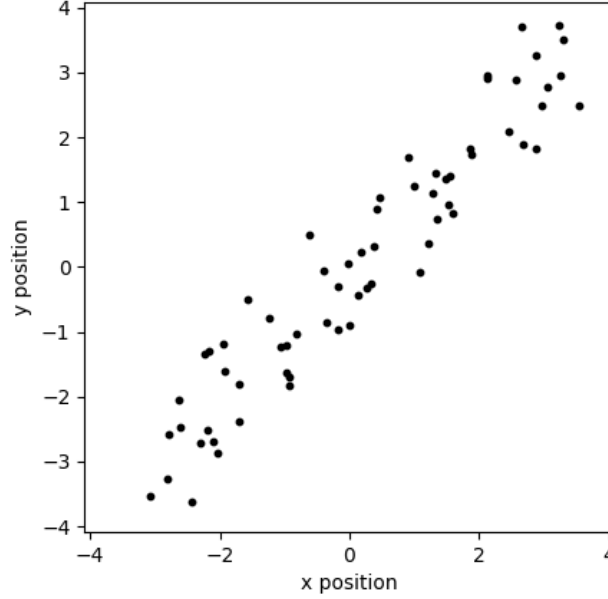


Figure 4.1: Randomly generated sample data. The data lies along a line with slope 1 and has mean $\mathbf{0}$.

5. If the subspace is trivial we have found all axis.

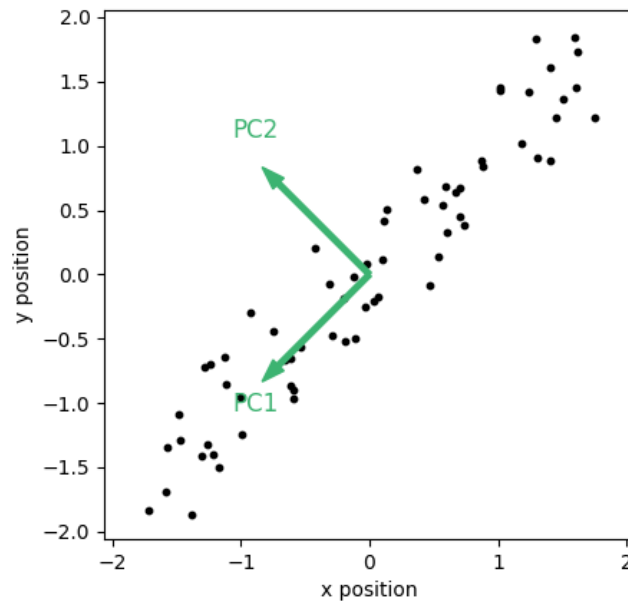
While this algorithm shows us what conceptually has to be done, we do not know how to compute the axis yet. We will now investigate this problem using the mathematical concepts from section 3. This will lead us to an algorithm in which all axis can be computed simultaneously.

Let $\mathbf{X} \in \mathbb{R}^{m \times n}$ be the data matrix. We want to find some orthonormal matrix \mathbf{P} such that $\mathbf{Y} := \mathbf{P}\mathbf{X}$ has a diagonal covariance matrix $\mathbf{C}_\mathbf{Y}$.

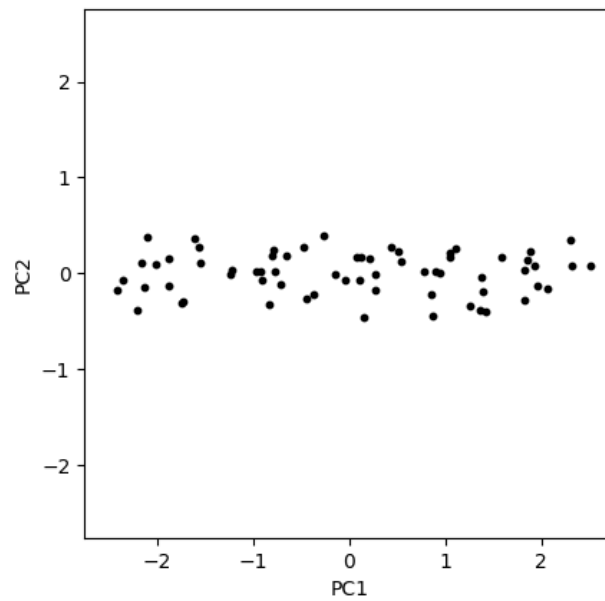
$$\begin{aligned} \mathbf{C}_\mathbf{Y} &= \frac{1}{n} \mathbf{Y}\mathbf{Y}^T = \frac{1}{n} (\mathbf{P}\mathbf{X})(\mathbf{P}\mathbf{X})^T = \frac{1}{n} \mathbf{P}\mathbf{X}\mathbf{X}^T\mathbf{P}^T = \\ &= \mathbf{P} \left(\frac{1}{n} \mathbf{X}\mathbf{X}^T \right) \mathbf{P}^T = \mathbf{P}\mathbf{C}_\mathbf{X}\mathbf{P}^T \end{aligned}$$

The covariance matrix $\mathbf{C}_\mathbf{X}$ is symmetric and therefore has a decomposition into an orthogonal matrix of eigenvectors \mathbf{V} and a diagonal matrix of eigenvalues \mathbf{D} . We choose $\mathbf{P} = \mathbf{V}^T$. From lemma 4 it follows that $\mathbf{V}^{-1} = \mathbf{V}^T$.

$$\begin{aligned} \mathbf{P}\mathbf{C}_\mathbf{X}\mathbf{P}^T &= \mathbf{P}(\mathbf{V}\mathbf{D}\mathbf{V}^{-1})\mathbf{P}^T = \mathbf{P}(\mathbf{V}\mathbf{D}\mathbf{V}^T)\mathbf{P}^T = \\ &= \mathbf{P}(\mathbf{P}^T\mathbf{D}\mathbf{P})\mathbf{P}^T = (\mathbf{P}\mathbf{P}^T)\mathbf{D}(\mathbf{P}\mathbf{P}^T) = \\ &= (\mathbf{P}\mathbf{P}^{-1})\mathbf{D}(\mathbf{P}\mathbf{P}^{-1}) = \mathbf{D} \end{aligned}$$

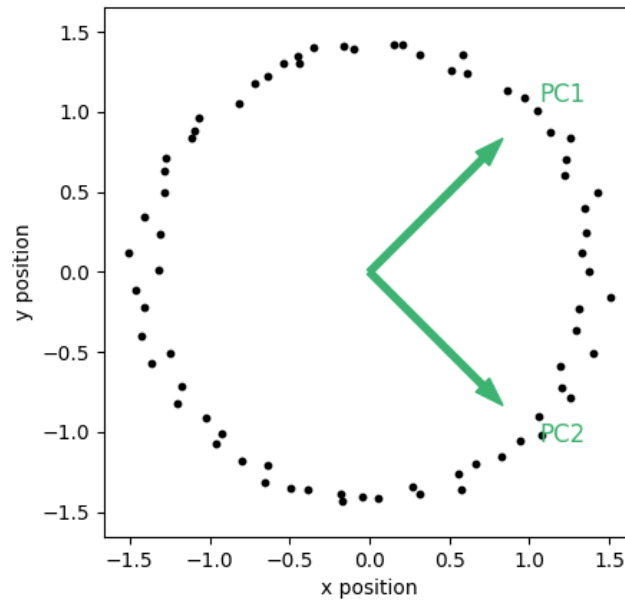


(a) Normed data (mean is zero and variance is one) and direction of the two PCs in relation to the x and y position.

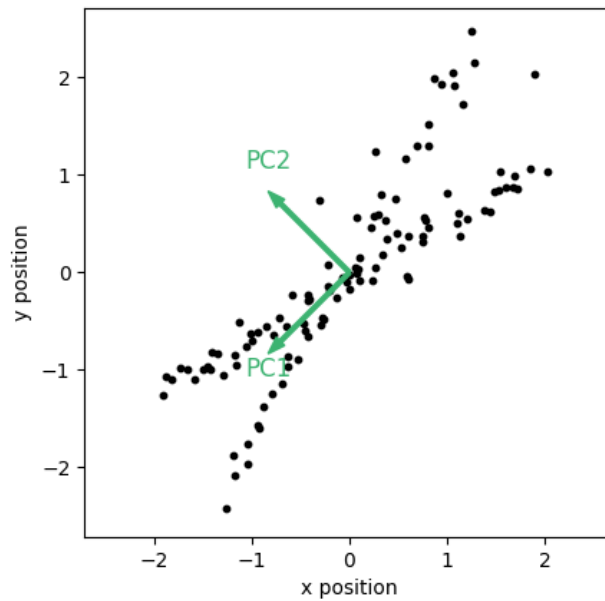


(b) The data after being transformed by PCA. The variance along the PC1 axis is maximal, therefore the data is spread out most along this axis.

Figure 4.2: Example application of PCA.



(a) Clearly the relationship in this figure is non-linear. PCA can not describe circular dependencies, as shown in this data.



(b) The two main axis along which the data is aligned are not orthogonal to each other. PCA always outputs orthogonal principal components, therefore it fails in this example.

Figure 4.3: Examples in which some of the assumptions of PCA are not valid. The results are sub-optimal.

In summary \mathbf{Y} has a diagonal covariance matrix if we choose $\mathbf{Y} = \mathbf{V}^T \mathbf{X}$, where \mathbf{V} is the matrix of eigenvectors of \mathbf{C}_X . The eigenvectors are the PCs and the eigenvalues are the variance in each new axis.

As pseudo code we get the program from algorithm 1 for calculating the PCA.

Algorithm 1 Principal Component Analysis

Require: matrix $X \in \mathbb{R}^{m \times n}$

Normalize each row in the matrix X

Calculate the covariance matrix C_X

Calculate the eigenvalues and eigenvectors of C_X

Sort the eigenvalues

Return sorted eigenvalues and corresponding eigenvectors

What happens if we skip the step in which we normalize each row in the matrix? A big variance is interpreted by the PCA algorithm as much information, thus the variance of the variables have an impact on how "important" the variable is deemed. As we do not want to prioritize certain variables we avoid this behavior by normalizing the data beforehand.

5 Methodology

For the algorithm in section 6 a few other methods will be used. As they are not the focus of this work we only give a short overview.

5.1 Fourier Transformation

TODO

5.2 Classification

In classification the objective is to find assignments between data points and categories. In our context we are interested in finding an assignment which closely matches some already categorized data. One simple approach to this problem is the k-nearest-neighbors algorithm.

For data that can be represented in \mathbb{R}^m and l categories the pseudo code is shown in algorithm 2.

Algorithm 2 k Nearest Neighbors

Require: data points $(p_i)_{i \in \mathbb{N}} \in \mathbb{R}^m$, categories $(c_i)_{i \in \mathbb{N}} \in \{0, 1, \dots, l\}$, point $x \in \mathbb{R}^m$, $k \in \mathbb{N}$
Calculate the distance between each point p_i and x
Take the k data points with the smallest distance to x
Return the category that most of the k points are assigned

Figure 5.1 shows a graphical representation of the algorithm for points in \mathbb{R}^2 and $k = 10$.

This algorithm is slow for big datasets as for each point the distance to x has to be calculated. One possibility to reduce calculation time is to partition space into smaller chunks. Then the loop only has to be over data points lying in the same or close chunks as the point we are interested in.

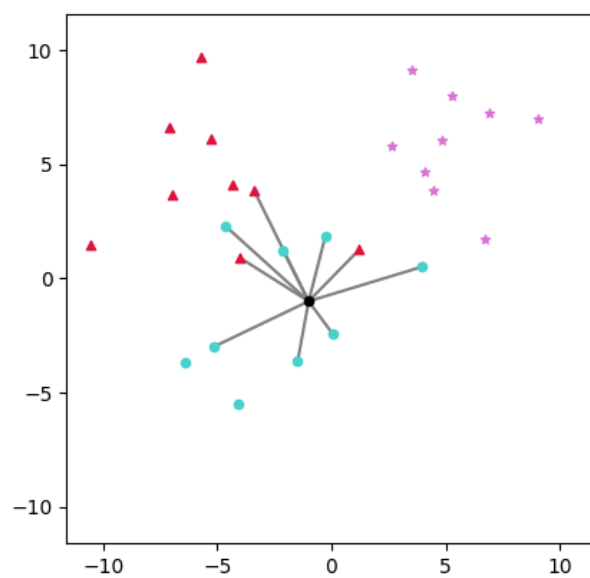


Figure 5.1: Data points in three categories are given. The black point is classified as a blue, circular point by the k-nearest-neighbors algorithm for $k = 10$.

6 Data and Algorithm

6.1 Data

The given dataset describes multiple channels of EEG data recorded during sleep. There are 14 sleep recordings which are between 490 and 735 minutes long. We are interested in the channels Fp2-F4 and Fp1-F3 as they record activity in the frontal area, where different sleep patterns are shown. [TODO cite something showing this.] The figure 6.1 shows where the sensors where placed.

[TODO describe dataset in more detail]

The data was recorded in 512 Hz for channels Fp2-F4 and Fp1-F3. [TODO other channels in other frequencies?] We resample to 200 Hz to reduce data size and standardize all channels.

6.2 Algorithm

We now describe the general structure of the algorithm from reading the data files to the clustering.

First we split the recorded data into 30 second sections. For each of these sections we get the data as a function of amplitude over time. An example segment is shown in the upper figure of 6.2. We apply a Fourier transformation as described in section 5.1 and get a function of amplitude over frequency. This gives a frequency based representation of the data. As the sleep stages are characterized by different frequencies and amplitudes, as described in section 1.1, this will help us distinguishing them. The output of such an FFT can be seen in the lower figure of 6.2. The pseudo code for this processing is

Algorithm 3 EEG data processing

```
for each patient do
  Read EEG data from file
  Apply a lowpass and highpass filter
  Resample with frequency 200 Hz
  Get one of the channels Fp2-F4, Fp1-F3, C4-A1, F3A2 or C3-A2 (in this order)
  Section out the part of data where the patient is asleep
  for each 30 second segment of the data do
    Do a Fourier transformation (FFT)
  end for
  Save all the outputs from the FFT to the patients file
end for
```

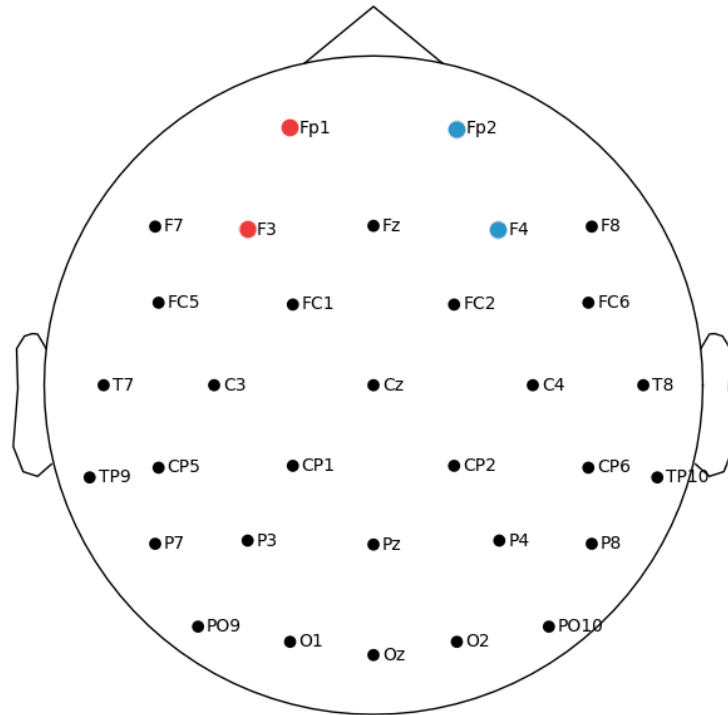


Figure 6.1: Scheme of sensor placement in recordings of EEG data. Fp1-F3 measure the difference in TODO between Fp1 and F3, both shown in red. Fp2-F4 is analogous for Fp2 and F4, shown in blue. All four points are located in the frontal area.

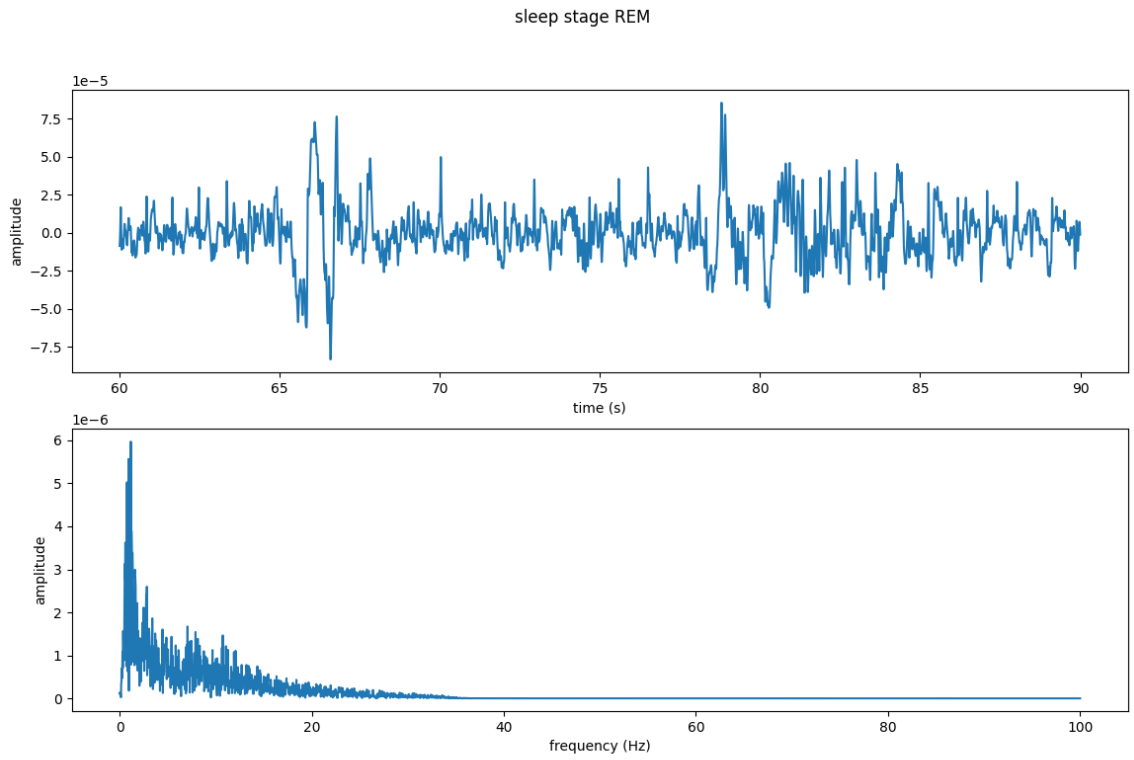


Figure 6.2: A 30 second segment of the data. The data was filtered with a 0.5 Hz highpass and a 30 Hz lowpass filter. The top shown the recorded signal of the EEG. The bottom shows the output of the Fourier transformation. This segment was assigned the REM sleep stage.

We now have high dimensional data points, as each 30 second segment is represented by the amplitudes for each of the distinct frequencies output by the FFT. Before we can start searching for cluster in the data we want to find a lower dimensional representation. This is where PCA can be used. From the output of algorithm 4 we get PCs, which reveal the frequencies where the most variance is shown. Under our assumptions this variation is caused by the different sleep stages each recording goes through. We show a visual representation of the data in the first three PC axis in figure 6.3.

Algorithm 4 Apply PCA to the EEG data

```
for each patient do
    Read FFT output from patients file
    Read the sleep stages from another file
end for
Concatenate FFT data of all patients
Concatenate sleep stages of all patients
Do PCA on the combined FFT data
Transform data according to the PC from the PCA
Visualize the data in the first three axis, with the color corresponding to the sleep stage
```

If we have a new recording and want to know the sleep stages we have to follow similar steps. First the recording has to be split into 30 second segments. These have to be transformed by a FFT. The output can be converted to the PCA basis by multiplying with the matrix of PCs. Lastly we can get a guess of the sleep stage by using the k nearest neighbor algorithm. The pseudo code is

Algorithm 5 Get guess for sleep stage

Require: EEG recording of sleep

```
for each 30s segment do
    Do FFT on the segment
    Multiply FFT output by PCs matrix
    Do k nearest neighbor
    Save the result of k nearest neighbor
end for
```

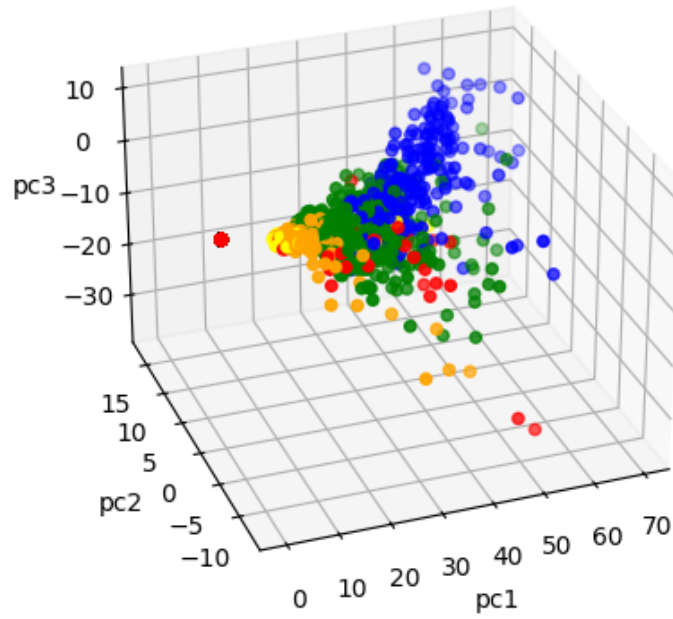


Figure 6.3: Three dimensional representation of the data. Each 30 second segment is represented by a dot in the color corresponding to the sleep stage this segment is assigned to. Awake is represented by red, orange is REM, yellow is S1, green S2 and blue S3.

7 Results

8 Conclusion

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