

MAS Ü11

1) (Ω, \mathcal{F}, P) ... Wahrscheinlichkeitsraum, $\mathcal{A} \subseteq \mathcal{F}$... σ -Algebra, $1 \leq p$, $q := \frac{p}{p-1}$

$$X \in L_p(\Omega, \mathcal{F}, P), Y \in L_p(\Omega, \mathcal{A}, P)$$

$$z.z.: Y = E(X|\mathcal{A}) \Leftrightarrow \forall Z \in L_q(\Omega, \mathcal{A}, P): \int Y Z dP = \int X Z dP$$

$$\Leftrightarrow z.z.: Y \text{ ist } \mathcal{A} \text{ messbar und } \forall A \in \mathcal{A}: E[X \mathbb{1}_A] = \int_A X dP = \int_A Y dP = E[Y \mathbb{1}_A]$$

Da $Y \in L_p(\Omega, \mathcal{A}, P)$ ist Y \mathcal{A} -messbar.

$$\text{Sei } A \in \mathcal{A} \text{ bel.} \Rightarrow \mathbb{1}_A \in L_q(\Omega, \mathcal{A}, P), \text{ da } \int \mathbb{1}_A^q dP = \int \mathbb{1}_A dP = P(A) \leq 1 < \infty \quad z = \mathbb{1}_A$$

$$\Rightarrow \int Y \mathbb{1}_A dP = \int X \mathbb{1}_A dP \Rightarrow E[X \mathbb{1}_A] = E[Y \mathbb{1}_A] \Rightarrow Y = E(X|\mathcal{A})$$

\Rightarrow

$$g: L_q(\Omega, \mathcal{A}, P) \rightarrow \mathbb{R} \quad \text{ist linear und stetig}$$

$$z \mapsto \int z \cdot X dP$$

$$\Rightarrow \exists! h \in L_q(\Omega, \mathcal{A}, P) : \forall Z \in L_q(\Omega, \mathcal{A}, P): g(Z) = \langle h, Z \rangle \text{ nach Riesz}$$

$$\text{also } \forall Z \in L_q(\Omega, \mathcal{A}, P): \int Z \cdot X dP = \int Z \cdot h dP$$

$$\Rightarrow \forall A \in \mathcal{A}: \int \mathbb{1}_A \cdot X dP = \int \mathbb{1}_A \cdot h dP \Rightarrow h = E(X|\mathcal{A}) = Y$$

$$\Rightarrow \forall Z \in L_q(\Omega, \mathcal{A}, P): \int Z \cdot X dP = \int Z \cdot Y dP$$



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3) (Ω, \mathcal{F}) ... Messraum, \mathcal{F} ... Familie von Wahrscheinlichkeitsmaßen

$\mathcal{A} \in \mathcal{S}$... Sub- σ -Algebra heißt suffizient für \mathcal{F} $\Leftrightarrow \forall A \in \mathcal{F} \exists h_A \dots \mathcal{A}$ -messbare Fkt.
mit $h_A = \mathbb{E}_P[1_A | \mathcal{A}] \quad \forall P \in \mathcal{F}$

Sei $\Omega = \{1, 2, 3, 4\}$, $\mathcal{F} = 2^\Omega$, $\mathcal{F} = \{P_1, P_2\}$ wobei $P_1(A) = 1_A(1)$, $P_2(A) = 1_A(4)$

$\mathcal{A}_1 = \{\emptyset, \Omega, \{1, 3\}, \{2, 4\}\}$, $\mathcal{A}_2 = \{\emptyset, \Omega, \{1, 2\}, \{3, 4\}\}$

zz: $\mathcal{A}_1, \mathcal{A}_2$ sind suffizient für \mathcal{F}

①

Sei $h_A := P_1(\{1, 3\} \cap A) 1_{\{1, 3\}} + P_2(\{2, 4\} \cap A) 1_{\{2, 4\}}$ ist \mathcal{A}_1 -messbar

$$= 1_{\{1, 3\} \cap A}(1) 1_{\{1, 3\}} + 1_{\{2, 4\} \cap A}(4) 1_{\{2, 4\}} = 1_A(1) 1_{\{1, 3\}} + 1_A(4) 1_{\{2, 4\}}$$

$$\Rightarrow h_A(\omega) = \begin{cases} 1, & (\omega \in A \wedge \omega \in \{1, 3\}) \vee (\omega \in A \wedge \omega \in \{2, 4\}) \\ 0, & \text{sonst} \end{cases}$$

$$\text{zz: } \forall A \in \mathcal{P}(\Omega): h_A = \mathbb{E}_{P_1}[1_A | \mathcal{A}_1] \wedge h_A = \mathbb{E}_{P_2}[1_A | \mathcal{A}_1]$$

Sei $B \in \mathcal{A}_1$ bel.

$$\mathbb{E}_{P_1}[1_A 1_B] = \int_{A \cap B} dP_1 = P_1(A \cap B) = 1_{A \cap B}(1)$$

$$\begin{aligned} \mathbb{E}_{P_1}[h_A 1_B] &= \int_B h_A dP_1 = 1_A(1) \int_B 1_{\{1, 3\}}(\omega) dP_1(\omega) + 1_A(4) \int_B 1_{\{2, 4\}}(\omega) dP_1(\omega) \\ &= 1_A(1) P_1(B) + 0 = 1_A(1) 1_B(1) = 1_{A \cap B}(1) \end{aligned}$$

$$\mathbb{E}_{P_2}[1_A 1_B] = \int_{A \cap B} dP_2 = 1_{A \cap B}(4)$$

$$\mathbb{E}_{P_2}[h_A 1_B] = 1_A(1) \int_B 1_{\{1, 3\}}(\omega) dP_2(\omega) + 1_A(4) \int_B 1_{\{2, 4\}}(\omega) dP_2(\omega) = 1_A(4) P_2(B) = 1_{A \cap B}(4)$$

② $h_A = P_1(\{1, 2\} \cap A) 1_{\{1, 2\}} + P_2(\{3, 4\} \cap A) 1_{\{3, 4\}} = 1_A(1) 1_{\{1, 2\}} + 1_A(4) 1_{\{3, 4\}}$ ist \mathcal{A}_2 -messbar

$$\mathbb{E}_{P_1}[1_A 1_B] = \int_{A \cap B} dP_1 = 1_{A \cap B}(1)$$

$$\mathbb{E}_{P_1}[h_A 1_B] = 1_A(1) \int_B 1_{\{1, 2\}}(\omega) dP_1(\omega) + 1_A(4) \int_B 1_{\{3, 4\}}(\omega) dP_1(\omega) = 1_A(1) P_1(B) = 1_{A \cap B}(1)$$

$$\mathbb{E}_{P_2}[1_A 1_B] = 1_{A \cap B}(4)$$

$$\mathbb{E}_{P_2}[h_A 1_B] = 1_A(1) \int_B 1_{\{1, 2\}}(\omega) dP_2(\omega) + 1_A(4) \int_B 1_{\{3, 4\}}(\omega) dP_2(\omega) = 1_A(4) P_2(B) = 1_{A \cap B}(4)$$

$\mathcal{A}_1 \cap \mathcal{A}_2 = \{\emptyset, \Omega\}$ suffizient für \mathcal{F} ? $h_A \dots \mathcal{A}_1 \cap \mathcal{A}_2$ -messbar $\Rightarrow h_A = c \dots$ konstant

Für $B = \Omega$ gilt: $\int_{\Omega} 1_{\mathcal{F}} dP_1 = \int_{\Omega} c dP_1 \wedge \int_{\Omega} 1_{\mathcal{F}} dP_2 = \int_{\Omega} c dP_2 \Rightarrow 1 = P_1(\{1\}) = c = P_2(\{1\}) = 0 \quad \downarrow$

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4) $N \dots$ SG auf \mathbb{N} verteilt mit $P[N=k] = p_k$, $(X_i)_{i \in \mathbb{N}}$ Folge, u.a. von N , mit $\forall i \in \mathbb{N}$ F_i :
 $E(X_i) = \mu_i$, $V(X_i) = \sigma_i^2$; $Z = X_N \dots$ Auswahlgröße, zufällig X_0, X_1, X_2, \dots an der N -ten Stelle
 wird beobachtet.

zz: VF von Z $F_Z(x) = \sum_{k \geq 0} p_k F_k(x) \wedge$ falls $\exists M_{X_k}$ auf $(-d, d) \forall k=0,1,2,\dots \Rightarrow M_Z(t) = \sum_{k \geq 0} p_k M_{X_k}(t)$
 existiert für, falls Summe endlich

$$F_Z(x) = P(Z \leq x) = P(X_N \leq x) = \sum_{k=0}^{\infty} P(X_N \leq x | N=k) \cdot P(N=k)$$

$$= \sum_{k \geq 0} P(X_k \leq x) \cdot p_k = \sum_{k \geq 0} p_k \cdot F_k(x)$$

$$M_Z(t) = E(e^{tZ}) = E(e^{tX_N}) = \sum_{k \geq 0} E(e^{tX_k} | N=k) P(N=k)$$

$$= \sum_{k \geq 0} p_k \cdot E(e^{tX_k}) = \sum_{k \geq 0} p_k \cdot M_{X_k}(t) \quad \text{falls endliche Summe} \Rightarrow M_Z(t) < \infty$$

$N \sim \text{Pois}(\theta)$ mit Rate $\theta > 0$, $X_i \sim N(0, i)$, $X_0 = 0$

ges: $M_Z(t)$

$$p_k = P(N=k) = \frac{\theta^k}{k!} e^{-\theta}$$

$$M_{X_0}(t) = E(e^{tX_0}) = E(e^0) = E(1) = 1 (= e^{\frac{1}{2} \cdot 0 \cdot t^2})$$

$$M_{X_k}(t) = e^{\frac{1}{2} k t^2}$$

$$\Rightarrow M_Z(t) = \sum_{k \geq 0} p_k M_{X_k}(t) = \sum_{k \geq 0} \frac{\theta^k}{k!} e^{-\theta} \cdot e^{\frac{1}{2} k t^2} = e^{-\theta} \sum_{k \geq 0} \frac{1}{k!} (\theta \cdot e^{\frac{1}{2} t^2})^k$$

$$= e^{-\theta} \exp(\theta \cdot e^{\frac{1}{2} t^2}) = \exp(\theta (e^{\frac{1}{2} t^2} - 1))$$

□

(5) $\frac{d}{dt} M_{X_k}(t) = e^{\frac{1}{2} k t^2} \cdot \frac{1}{2} k \cdot 2t = k \cdot t \cdot e^{\frac{1}{2} k t^2}$

$$\frac{d}{dt} M_Z(t) = \exp(\theta (e^{\frac{1}{2} t^2} - 1)) \cdot \theta \cdot e^{\frac{1}{2} t^2 - 1} \cdot \frac{1}{2} 2t = \theta \cdot t \cdot e^{\frac{1}{2} t^2 - 1} \cdot \exp(\theta (e^{\frac{1}{2} t^2} - 1)) = \theta \cdot t \cdot \exp(\frac{1}{2} t^2 - 1 + \theta (e^{\frac{1}{2} t^2} - 1))$$

$$\sum_{k=0}^{\infty} p_k \frac{d}{dt} M_{X_k}(t) = \sum_{k=1}^{\infty} \frac{\theta^k}{k!} e^{-\theta} \cdot k \cdot t \cdot e^{\frac{1}{2} k t^2} = e^{-\theta} \cdot t \cdot \sum_{k=1}^{\infty} \frac{\theta^k}{(k-1)!} e^{\frac{1}{2} k t^2}$$

$$= \theta \cdot e^{\frac{1}{2} t^2} \cdot e^{-\theta} \cdot t \cdot \sum_{k=1}^{\infty} \frac{1}{(k-1)!} (\theta \cdot e^{\frac{1}{2} t^2})^{k-1} = \theta \cdot t \cdot e^{\frac{1}{2} t^2 - \theta} \cdot \sum_{k=0}^{\infty} \frac{1}{k!} (\theta \cdot e^{\frac{1}{2} t^2})^k$$

$$= \theta \cdot t \cdot e^{\frac{1}{2} t^2 - \theta} \cdot \exp(\theta \cdot e^{\frac{1}{2} t^2}) = \theta \cdot t \cdot \exp(\frac{1}{2} t^2 - \theta + \theta e^{\frac{1}{2} t^2}) = \frac{d}{dt} M_Z(t)$$

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5) N...SG auf \mathbb{N} , $P[N=k]=p_k$ X_i ... Folge, u.a. von \mathbb{N} $E X_i = \mu_i$, $V X_i = \sigma_i^2$

X_i hat VF F_i , $Z = X_N$... Auswahlgröße

beschränkter Fall: $\exists n \forall k > n: p_k = 0$

$$z.z.: E(Z) = \sum_{k=0}^{\infty} p_k \mu_k \quad V(Z) = E(\sigma_N^2) + V(\mu_N)$$

$$\begin{aligned} E(Z) &= \left. \frac{d}{dt} M_Z(t) \right|_{t=0} = \left. \frac{d}{dt} \sum_{k=0}^n p_k M_{X_k}(t) \right|_{t=0} \stackrel{\Delta}{=} \sum_{k=0}^n p_k \left. \frac{d}{dt} M_{X_k}(t) \right|_{t=0} \\ &= \sum_{k=0}^n p_k E(X_k) = \sum_{k=0}^n p_k \mu_k \end{aligned}$$

$$E(Z^2) = \left. \frac{d^2}{dt^2} M_Z(t) \right|_{t=0} = \left. \frac{d^2}{dt^2} \sum_{k=0}^n p_k M_{X_k}(t) \right|_{t=0} \stackrel{\Delta}{=} \sum_{k=0}^n p_k \left. \frac{d^2}{dt^2} M_{X_k}(t) \right|_{t=0} = \sum_{k=0}^n p_k E(X_k^2)$$

$$V(Z) = E(Z^2) - E(Z)^2 = \sum_{k=0}^n p_k E(X_k^2) - \left(\sum_{k=0}^n p_k \mu_k \right)^2$$

$$\stackrel{*}{=} \sum_{k=0}^n p_k \sigma_k^2 + \sum_{k=0}^n p_k \mu_k^2 - \left(\sum_{k=0}^n p_k \mu_k \right)^2$$

$$= E(\sigma_N^2) + E(\mu_N^2) - (E(\mu_N))^2 = E(\sigma_N^2) + V(\mu_N)$$

$$\begin{aligned} * \sigma_k^2 &= V X_k = E(X_k^2) - E(X_k)^2 = E(X_k^2) - \mu_k^2 \\ &\Rightarrow E(X_k^2) = \sigma_k^2 + \mu_k^2 \end{aligned}$$

Δ Im unbeschränkter Fall möglicherweise problematisch.

$N \sim \text{Pois}(\theta)$, $X_i \sim N(0, 1)$, $X_0 \sim 0$

ges: $V(Z)$

Mit obiger Formel (eigentlich nicht beschränkt): $V(Z) = E(\sigma_N^2) + V(\mu_N)$

$$E(\sigma_N^2) = \sum_{k=0}^{\infty} p_k \sigma_k^2 = \sum_{k=1}^{\infty} \frac{\theta^k}{k!} e^{-\theta} \cdot k = \theta e^{-\theta} \sum_{k=1}^{\infty} \frac{\theta^{k-1}}{(k-1)!} = \theta e^{-\theta} \sum_{k=0}^{\infty} \frac{\theta^k}{k!} = \theta e^{-\theta} e^{\theta} = \theta$$

$V(\mu_N) = 0$, da $\forall k \in \mathbb{N}: \mu_k = 0$ also konstant

$$\Rightarrow V(Z) = \theta$$

$$\begin{aligned} \text{Mit moment erzeugende Funktion: } V(Z) &= E(Z^2) - E(Z)^2 = \left. \frac{d^2}{dt^2} M_Z(t) \right|_{t=0} - \left(\left. \frac{d}{dt} M_Z(t) \right|_{t=0} \right)^2 \\ &= \left. \frac{d^2}{dt^2} \exp(\theta \cdot e^{\frac{1}{2}t^2 - 1}) \right|_{t=0} - \left(\left. \frac{d}{dt} \exp(\theta \cdot e^{\frac{1}{2}t^2 - 1}) \right|_{t=0} \right)^2 \\ &= \theta \exp(\theta \cdot e^{-1} - 2)(e) - 0 = \theta e^{\frac{\theta}{2} - 2 + 1} = \theta e^{\frac{\theta}{2} - 1} \neq \theta, \text{ da siehe unten (4)} \end{aligned}$$

$$\begin{aligned} \frac{d}{dt} \exp(\theta \cdot e^{\frac{1}{2}t^2 - 1}) &= \exp(\theta \cdot e^{\frac{1}{2}t^2 - 1}) \cdot \theta \cdot \exp(\frac{1}{2}t^2 - 1) \cdot \frac{1}{2} \cdot 2 \cdot t = \theta t \cdot \exp(\theta \cdot e^{\frac{1}{2}t^2 - 1}) \cdot e^{\frac{1}{2}t^2 - 1} \\ \frac{d^2}{dt^2} \exp(\theta \cdot e^{\frac{1}{2}t^2 - 1}) &= \theta \exp(\theta \cdot e^{\frac{1}{2}t^2 - 1}) + \frac{1}{2} \cdot 2 \cdot t \cdot \theta t \cdot \exp(\theta \cdot e^{\frac{1}{2}t^2 - 1}) \cdot e^{\frac{1}{2}t^2 - 1} \end{aligned}$$

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6) X_i : Folge identisch verteilter SG, $X_i \sim X$, $\mathbb{E} X < \infty$

X_i : paarweise unkorreliert

$$\text{z.z.: } p\text{-}\lim_{n \rightarrow \infty} \bar{X}_n = p\text{-}\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X_i = \mathbb{E} X$$

[Tschebyscheff]-Ungleichung $P(|X - \mu_X| \geq k) \leq \frac{\sigma_X^2}{k^2}$ falls $\mu_X = \mathbb{E}(X)$, $\sigma_X^2 = V(X) < \infty$, $k > 0$

$$\mathbb{E}(\bar{X}_n) = \frac{1}{n} \sum_{i=1}^n \mathbb{E}(X_i) = \frac{1}{n} \sum_{i=1}^n \mathbb{E}(X) = \mathbb{E}(X) = \mu_X < \infty$$

$$\sigma_{\bar{X}_n}^2 = V(\bar{X}_n) = \frac{1}{n^2} V\left(\sum_{i=1}^n X_i\right) = \frac{1}{n^2} \sum_{i=1}^n V(X_i) = \frac{1}{n^2} (n \cdot V(X)) = \frac{V(X)}{n} = \frac{\sigma_X^2}{n}$$

unkorreliert $\Rightarrow V(X+Y) = V(X) + V(Y) + 2 \underbrace{\text{Cov}(X,Y)}_{=0} = V(X) + V(Y)$

$$P(|\bar{X}_n - \mu_X| \geq k) \leq \frac{\sigma_X^2}{n} \cdot \frac{1}{k^2} \xrightarrow{n \rightarrow \infty} 0 \quad \forall k > 0$$

$$\Rightarrow p\text{-}\lim_{n \rightarrow \infty} \bar{X}_n = \mu_X = \mathbb{E} X$$

□

[Def] $p\text{-}\lim_{n \rightarrow \infty} X_n = X \Leftrightarrow \forall \varepsilon > 0: \lim_{n \rightarrow \infty} P(|X_n - X| \geq \varepsilon) = 0$

$X_i, X_j \dots$ unkorreliert $\Rightarrow 0 = C(X_i, X_j) = \mathbb{E}(X_i \cdot X_j) - \mathbb{E}(X_i) \cdot \mathbb{E}(X_j) = \mathbb{E}(X^2) - \mu_X^2$
 $\Rightarrow \mathbb{E}(X^2) = \mu_X^2 < \infty$
 $\Rightarrow V(X) < \infty$ (eigentlich sogar $V(X) = 0$)

\hookrightarrow bei $V(X) = 0 \Rightarrow V(\bar{X}_n) = 0 \Rightarrow \bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i = \mathbb{E} X$ also auch p-lim