PROBLEH 15 12

Dendrancianon:

Show that 
$$\mathbb{E}[\hat{g}(\vec{x})] = \mathbb{E}[g(x)]$$

 $E[\hat{g}(\hat{x})] = E[\frac{1}{N} \xi g(x_i)] = \begin{cases} u \text{ wereing of expectation} \\ x_i \text{ are from} \end{cases}$ 

$$= \frac{1}{N} \sum_{i=1}^{N} g(x_i) = \frac{1}{N} \cdot M = [g(x_i)] = \frac{1}{N} \cdot M = \frac{1}{$$

Supu that  $|der[\hat{g}(\vec{x})] = \frac{1}{N} |der(g(x))|$ 

DEMONSRATION:  $Vous[\hat{g}(\vec{x})] = Vous[\frac{1}{N}, \frac{N}{N}g(x;)] =$ 

$$= \frac{1}{N^2} \operatorname{Vol}\left[\sum_{i=1}^{N} g(x_i)\right] = \frac{1}{N^2} \sum_{i=1}^{N} \sum_{i=1}^{N}$$

 $\Sigma$  - couple ance hotely of  $4g(x_i)$  - uplance of  $g(x_i)$  is on the  $i^{th}$  element of diagonal. Other Elements of the haten are  $(bu(g(x_i), g(x_i))$ . Since  $g(x_i)$  &  $g(x_i)$  are independent, they are edges.

(\*) CAN BE REWELTEN AS

$$= \frac{1}{N^2} \left( \sum_{i=1}^{N} \log \left[ g(x_i) \right] \right) = \frac{1}{N^2} \sum_{i=1}^{N} \log \left[ g(x_i) \right] = \frac{1}{N^2} \left[ \log \left[ g(x_i) \right] \right] = \frac{1}{N^2} \left[ \log \left[ g(x_i) \right] \right]$$

THOSE PROCURS ARE NOT APPLIED TO THE CORE WHEN X; ARE GENERATED FROM MCHC SAMPLER BECOMES X; DEPEND ON X;-1 (IF i > 0).

P(C=T | R=T, S=T, W=T) = 
$$\frac{P(R=T, S=T, W=T | C=T)}{P(R=T, S=T, W=T)}$$

$$= P(R=T|C=T)P(S=T|C=T)P(C=T) + P(R=T|C=F)P(S=T|C=F)P(C=F) = \frac{48}{10} \cdot \frac{1}{10} \cdot \frac{1}{10} \cdot \frac{1}{10} \cdot \frac{1}{10} \cdot \frac{1}{10} \cdot \frac{1}{10} \cdot \frac{1}{100} = \frac{9}{100}$$

• 
$$P(P=T, S=T, W=T|C=T) = P(P=T, S=T|C=T) =$$
  
=  $P(P=T|C=T) P(S=T|C=T) P(C=T) = \frac{8^4}{10} \cdot \frac{1}{10} \cdot \frac{1}{10} = \frac{4}{100}$ 

• P(C=T|P=T,S=T, ω=T) = 
$$\frac{u}{100}$$
 •  $\frac{100}{9}$  =  $\frac{u}{9}$  ≈ 0.44

$$\frac{P(C=T \mid R=F, S=T, W=T)}{P(R=F, S=T, W=T)} = \frac{P(R=F, S=T, W=T)}{P(R=F, S=T, W=T)}$$

• P(P=F, S=T, W=T) = P(P=F, S=T | C=T) =

$$= P(P=F|C=T)P(S=T|C=T)P(C=T) = \frac{\Delta}{1000}$$

$$\frac{P(R=T \mid C=T, S=T, \omega=T)}{P(C=T, S=T, \omega=T)} = \frac{P(R=T, C=T, S=T, \omega=T)}{P(C=T, S=T, \omega=T)}$$

$$= \frac{100}{40} \cdot \frac{100}{8} + \frac{100}{40} \cdot \frac{100}{8} = \frac{1000}{400} + \frac{1000}{180} = \frac{1000}{450}$$

$$= \frac{qq}{\omega_0} \cdot \frac{8}{\lambda_0} = \frac{4q2}{\lambda_000}$$
•  $P(P=T \mid C=T, S=T, \omega=T) = \frac{4q2}{1000} \cdot \frac{1000}{942} = \frac{4q2}{942} \approx 0.82$ 

• P(R=T, C=T, 8=T, W=T) = P(W=T | R=T, C=T, S=T) P(R=T | C=T, S=T)=

$$P(R=T \mid C=F, S=T, \omega=T) = \frac{P(R=T, C=F, S=T, \omega=T)}{P(C=F, S=T, \omega=T)}$$

$$= \frac{40}{100} \cdot \frac{5}{10} + \frac{100}{100} \cdot \frac{10}{10} = \frac{1000}{1000} + \frac{1000}{1000} = \frac{1000}{1000}$$

$$= \frac{100}{66} \cdot \frac{100}{5} = \frac{100}{100}$$

i) ANAWTICAL COMPUTATION OF MARGINAL PROBABILITY

$$P(R=T|S=T, \omega=T) = \frac{P(S=T, \omega=T, R=T)}{P(S=T, \omega=T)}$$

## Project 6

#### Contents

oblem 16	1
Gibbs sampler (b)	1
Marginal probability (c)	2
Auto-correlation and effective sample size (ESS) (d)	2
Gibbs sampling of 50000 samples (e)	4
Relative frequencies to detect burn-in phase (f) $\dots \dots \dots \dots \dots \dots \dots \dots \dots \dots \dots$	5
Potential scale reduction factor (g) $\dots$	6
Re-estimation of marginal probability (h) $\dots \dots \dots$	7
Analytical solution of marginal probability (i)	7

#### Problem 16

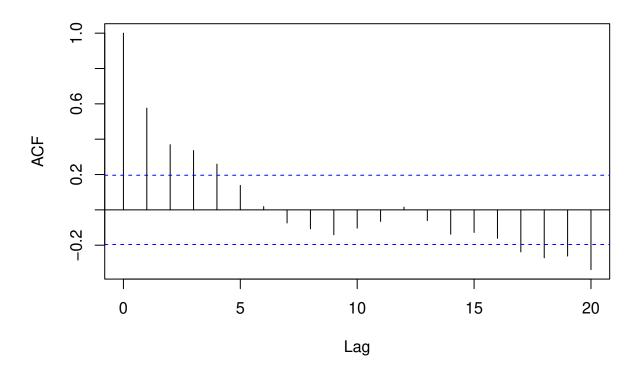
# <sup>5.1</sup> Gibbs sampler (b)

```
# Encoding conditional probabilities
p <- list(C=0.5, S.C1=0.1, S.C0=0.5, R.C1=0.8, R.C0=0.2,
          W.S1_R1=0.99, W.S1_R0=0.9, W.S0_R1=0.9, W.S0_R0=0.01)
p\$C.R1\_S1\_V1 <- p\$R.C1*p\$S.C1*p\$C/(p\$R.C1*p\$S.C1*p\$C+p\$R.C0*p\$S.C0*(1-p\$C))
p$C.R0_S1_W1 < (1-p$R.C1)*p$S.C1*p$C/((1-p$R.C1)*p$S.C1*p$C+(1-p$R.C0)*p$S.C0*(1-p$C))
p$R.C1_S1_W1 <- p$W.S1_R1*p$R.C1/(p$W.S1_R1*p$R.C1+p$W.S1_R0*(1-p$R.C1))
p$R.CO_S1_W1 <- p$W.S1_R1*p$R.CO/(p$W.S1_R1*p$R.CO+p$W.S1_R0*(1-p$R.CO))
Gibbs_sampler <- function(N=100) {</pre>
    samples <- matrix(NA, ncol=2, nrow=N, dimnames=list(NULL, c("R", "C")))</pre>
    # Initialisation (beginning state)
    samples[1,] <- c(TRUE, TRUE)</pre>
    for(n in 2:N) {
        # Random sampling to choose which variable (R or C) is updated
        indicator \leftarrow sample(c("R", "C"), size=1, prob=c(0.5, 0.5))
        if(indicator == "R") {
            \# Sampling from P(R_{n+1}/C_{n}, S=T, W=T)
            if(samples[(n-1), "C"]) sampling_prob <- p$R.C1_S1_W1 else sampling_prob <- p$R.C0_S1_W1
            sampled_R <- sample(c(TRUE, FALSE), size=1, prob=c(sampling_prob, (1-sampling_prob)))</pre>
            samples[n,] <- c(sampled_R, samples[(n-1), "C"])</pre>
        } else {
            # Sampling from P(C_{n+1}/R_{n}, S=T, W=T)
            if(samples[(n-1), "R"]) sampling_prob <- p$C.R1_S1_W1 else sampling_prob <- p$C.R0_S1_W1
            sampled_C <- sample(c(TRUE, FALSE), size=1, prob=c(sampling_prob, (1-sampling_prob)))</pre>
            samples[n,] <- c(samples[(n-1), "R"], sampled_C)</pre>
        }
```

```
}
    return(samples)
set.seed(42)
N <- 100
samples <- Gibbs_sampler(N)</pre>
result <- table(samples[, "R"], samples[, "C"])/N</pre>
rownames(result) <- c("R0", "R1")</pre>
colnames(result) <- c("CO", "C1")</pre>
result
##
##
          CO
                C1
     RO 0.59 0.05
##
     R1 0.18 0.18
##
Marginal probability (c)
Computing marginal probability P(R=T|S=T,W=T).
marginal.R.S1_W1 <- sum(result[2,])</pre>
marginal.R.S1_W1
## [1] 0.36
Auto-correlation and effective sample size (ESS) (d)
acf_output_R <- acf(as.numeric(samples[, "R"]), main="Rain")</pre>
```

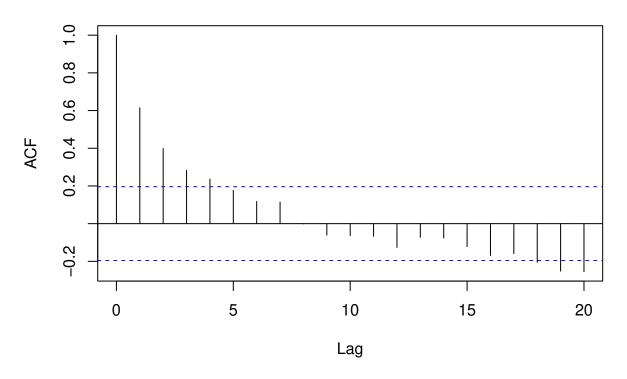
6.1

## Rain



acf\_output\_C <- acf(as.numeric(samples[, "C"]), main="Cloudy")</pre>

## Cloudy



```
ESS_R <- acf_output_R$n.used/(1+2*sum(acf_output_R$acf))
ESS_C <- acf_output_C$n.used/(1+2*sum(acf_output_C$acf))
print(paste("Effective sample size for 'rain':", ESS_R))

## [1] "Effective sample size for 'rain': 44.4718962322423"
print(paste("Effective sample size for 'cloudy':", ESS_C))
```

## [1] "Effective sample size for 'cloudy': 27.5547672392332"

#### Gibbs sampling of 50000 samples (e)

```
set.seed(42)
N <- 50000
samples1 <- Gibbs_sampler(N)

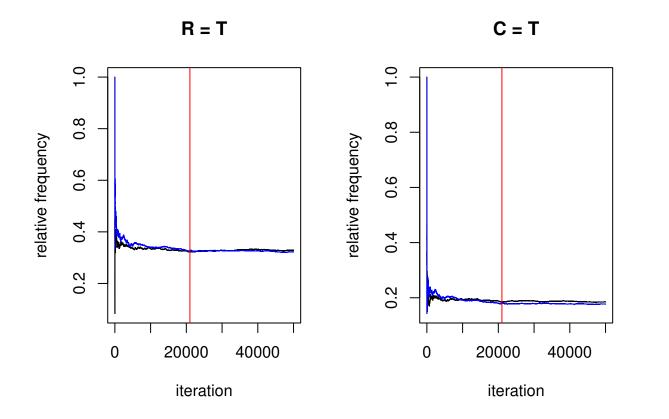
result1 <- table(samples1[, "R"], samples1[, "C"])/N
rownames(result1) <- c("R0", "R1")
colnames(result1) <- c("C0", "C1")

samples2 <- Gibbs_sampler(N)

result2 <- table(samples2[, "R"], samples2[, "C"])/N
rownames(result2) <- c("R0", "R1")
colnames(result2) <- c("C0", "C1")</pre>
```

```
result1
##
##
              CO
                      C1
     RO 0.63588 0.03570
##
##
     R1 0.17904 0.14938
result2
##
##
              CO
                      C1
     RO 0.64450 0.03334
##
     R1 0.17826 0.14390
##
```

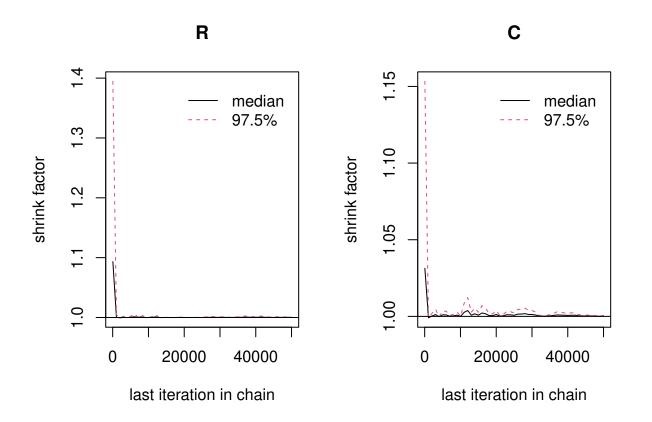
#### 9.1 Relative frequencies to detect burn-in phase (f)



Based on this plot, burn-in time could be suggested to be 21000. After this time point curves of relative frequency for both runs and both variables seem to be converged.

### 10.1

#### Potential scale reduction factor (g)



```
gelman.diag(mcmcs)
```

```
## Potential scale reduction factors:
##
## Point est. Upper C.I.
## R 1 1
## C 1 1
##
## Multivariate psrf
##
## 1
```

After around 20000 iterations, the fluctuations around the potential scale reduction factor are negligible for both variables, therefore a suggested burn-in time is 20000 iterations.

#### 11.1 Re-estimation of marginal probability (h)

Computing marginal probability P(R=T|S=T,W=T) excluding the samples from the burn-in phase. The length of the burn-in phase was chosen to be the average between the burn-in phase estimate from relative frequency plot and plot of potential scale reduction factor.

```
burn_in_phase <- 20500</pre>
result1 <- table(samples1[(burn_in_phase+1):N, "R"],</pre>
    samples1[(burn_in_phase+1):N, "C"])/(N-burn_in_phase)
rownames(result1) <- c("RO", "R1")</pre>
colnames(result1) <- c("CO", "C1")</pre>
result1
##
##
                 CO
                            C1
##
     RO 0.63315254 0.03471186
     R1 0.18230508 0.14983051
##
result2 <- table(samples2[(burn_in_phase+1):N, "R"],</pre>
    samples2[(burn_in_phase+1):N, "C"])/(N-burn_in_phase)
rownames(result2) <- c("R0", "R1")</pre>
colnames(result2) <- c("CO", "C1")</pre>
result2
##
##
                 CO
                            C1
##
     RO 0.64752542 0.03288136
##
     R1 0.17630508 0.14328814
marginal1.R.S1_W1 <- sum(result1[2,])</pre>
marginal2.R.S1_W1 <- sum(result2[2,])</pre>
print("Marginal probability using the first run of 50000 iterations and excluded burn-in phase:")
## [1] "Marginal probability using the first run of 50000 iterations and excluded burn-in phase:"
print(marginal1.R.S1 W1)
## [1] 0.3321356
print("Marginal probability using the second run of 50000 iterations and excluded burn-in phase:")
## [1] "Marginal probability using the second run of 50000 iterations and excluded burn-in phase:"
print(marginal2.R.S1 W1)
## [1] 0.3195932
```

#### <sup>11.2</sup> Analytical solution of marginal probability (i)

```
print("Analytical solution for the marginal probability:")
## [1] "Analytical solution for the marginal probability:"
print(analytical_marginal.R.S1_W1)
## [1] 0.3203883
```

The marginal probabilities of the (h) part are closer to the analytical solution, meanwhile the probability of (c) part differs more. It is intuitive, since the sample size in (c) was small (100).

```
library(rmarkdown)
render("Gibbs.Rmd", pdf_document(TRUE), "Indilewitsch_Toidze_Houhamdi_Pudziuvelyte_Project6.pdf")
```

# Index of comments

1.1	great!
1.2	2/2
2.1	1/1
2.2	0.048
5.1	1/1
6.1	1/1
8.1	0.75/1 should look at acf plots and sum until convergence, for large values of lag the sample correlation is too noisy
9.1	1/1
10.1	1/1
11.1	1/1
11.2	