Application of Information Theory, Lecture 2 Joint & Conditional Entropy, Mutual Information

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Part I

Joint and Conditional Entropy

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$$= \frac{1}{2} \cdot 1 + \frac{1}{2} \cdot 2 = 1\frac{1}{2}$$

Joint entropy, cont.

▶ The joint entropy of $(X_1, ..., X_n) \sim p$, is

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$$= -\sum_{z \in p_{Y|X}(Y|X)} \log z$$

Example

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$$= \frac{1}{2} H(Y|X = 0) + \frac{1}{2} H(Y|X = 1)$$

$$= \frac{1}{2} H(\frac{1}{2}, \frac{1}{2}) + \frac{1}{2} H(1, 0) = \frac{1}{2}.$$

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$$H(X|Y) = \mathop{\mathsf{E}}_{y \leftarrow Y} H(X|Y = y)$$

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$$= \frac{3}{4} H(\frac{1}{3}, \frac{2}{3}) + \frac{1}{4} H(1, 0) = 0.6887 \neq H(Y|X).$$

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- ► H(Y|X) = H(X, Y) H(X) is as an alternative definition for H(Y|X).

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 $H(P_{1,1}, \dots, P_{n,n})$
 $= H(q_1, \dots, q_n) + \sum_i q_i H(\frac{P_{i,1}}{q_i}, \dots, \frac{P_{i,n}}{q_i})$
 $= H(X) + H(Y|X).$

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For rvs X, Y, it holds that H(X, Y) = H(X) + H(Y|X).

Proof immediately follow by the grouping axiom:

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Another proof.

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For rvs X, Y, it holds that H(X, Y) = H(X) + H(Y|X).

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For rvs X_1, \ldots, X_k , it holds that

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Proof: ?

Extremely useful property!

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Examples

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 - Via mapping?

Let X_1, \ldots, X_n be Boolean iids with $X_i \sim (p, 1-p)$ and let $X = X_1, \ldots, X_n$. Let f be such that $\Pr[f(X) = z] = \Pr[f(X) = z']$, for every $k \in \mathbb{N}$ and $z, z' \in \{0, 1\}^k$. Let K = |f(X)|. Prove that $E K \le n \cdot h(p)$.

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Let $X_1, ..., X_n$ be Boolean iids with $X_i \sim (p, 1-p)$ and let $X = X_1, ..., X_n$. Let f be such that $\Pr[f(X) = z] = \Pr[f(X) = z']$, for every $k \in \mathbb{N}$ and $z, z' \in \{0, 1\}^k$. Let K = |f(X)|. Prove that $E K < n \cdot h(p)$.

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Interpretation

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- Interpretation
- Upper bounds

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 $\implies t \ge \log n! = \Theta(n \log n)$

Let $p=(p_1,\ldots,p_n)$ and $q=(q_1,\ldots,q_n)$ be two distributions, and for $\lambda\in[0,1]$ consider the distribution $\tau_\lambda=\lambda p+(1-\lambda)q$. (i.e., $\tau_\lambda=(\lambda p_1+(1-\lambda)q_1,\ldots,\lambda p_n+(1-\lambda)q_n)$.

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We are now certain that we drew the graph of the (two-dimensional) entropy function right...

Part II

Mutual Information

Mutual information

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▶ I(X; Y) — the "information" that X gives on Y

$$I(X; Y) := H(Y) - H(Y|X)$$

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$$\blacktriangleright$$

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- ► $I(X; Y) \ge 0$.

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- ▶ I(X; f(X)) = H(f(X)) (and smaller then H(X) if f is non-injective)

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- I(X; Y|Z) := H(Y|Z) H(Y|X,Z) ≥ 0
- ► I(X; Y|Z) = I(Y; X|Z) (since I(X'; Y') = I(Y'; X'))

X	0	1
0	$\frac{1}{4}$	$\frac{1}{4}$
1	1 2	0

X	0	1
0	1 4	1 4
1	1 2	0

$$I(X; Y) = H(X) - H(X|Y)$$

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$$= H(Y) - H(Y|X)$$

$$= h(\frac{1}{4}) - \frac{1}{2}h(\frac{1}{2})$$

Claim 4 (Chain rule for mutual information)

For rvs $X_1, ..., X_k, Y$, it holds that $I(X_1, ..., X_k; Y) = I(X; Y) + I(X_2; Y|X_1) + ... + I(X_k; Y|X_1, ..., X_{k-1})$.

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Proof: ?

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Proof: ? HW

Let X_1, \ldots, X_n be iid with $X_i \sim (p, 1 - p)$, under the condition that $\bigoplus_i x_i = 0$. Compute $I(X_1, \ldots, X_{n-1}; X_n)$.

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Compute $I(X_1, ..., X_{n-1}; X_n)$.

By chain rule

$$I(X_1, ..., X_{n-1}; X_n)$$

= $H(X_1; X_n) + H(X_2; X_n | X_1) + ... + H(X_{n-1}; X_n | X_1, ..., X_{n-2})$

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$$I(T; F) = H(T) - H(T|F)$$
$$= \log 6 - \log 4$$

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= log 3 - 1.

Part III

Data processing

Definition 5 (Markov Chain)

Rvs $(X, Y, Z) \sim p$ form a Markov chain, denoted $X \to Y \to Z$, if $p(x, y, z) = p_X(x) \cdot p_{Y|X}(y|x) \cdot p_{Z|Y}(z|y)$, for all x, y, z.

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- ▶ Since $I(X; Y|Z) \ge 0$, we conclude $I(X; Y) \ge I(X; Z)$.

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For any rvs X and Y, and any (even random) g, it holds that

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- ▶ We call \hat{X} an estimator for X (from Y).

► Let
$$E = \begin{cases} 1, & \hat{X} \neq X \\ 0, & \hat{X} = X. \end{cases}$$

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Let X and Y be rvs, let $\hat{X} = g(Y)$ and $P_e = \Pr \left[\hat{X} \neq X \right]$.

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