Application of Information Theory, Lecture 9 Parallel Repetition of Interactive Arguments

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Part I

Interactive Proofs and Arguments

\mathcal{NP} as a Non-interactive Proofs

Definition 1 (\mathcal{NP})

 $\mathcal{L} \in \mathcal{NP}$ iff \exists and poly-time algorithm \lor such that:

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- ▶ V(x, w) = 0 for every $x \notin \mathcal{L}$ and $w \in \{0, 1\}^*$

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- Efficient verifier, efficient prover (given the witness)
- Soundness holds unconditionally

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- Games no-input protocols.

Section 1

Interactive Proof for Graph Non-Isomorphism

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- We will show a simple interactive proof for GNT Idea: Beer tasting...

Interactive proof for \mathcal{GNI}

Protocol 4 ((P, V)(G₀ = ([
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- 1. V chooses $b \leftarrow \{0,1\}$ and $\pi \leftarrow \Pi_m$, and sends $\pi(E_b)$ to P.^a
- **2.** P send b' to V (tries to set b' = b).
- **3.** V accepts iff b' = b.
 - ${}^{a}\pi(E) = \{(\pi(u), \pi(v) \colon (u, v) \in E\}.$

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Claim 5

The above protocol is IP for \mathcal{GNI} , with perfect completeness and soundness error $\frac{1}{2}$.

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Hence,

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: $Pr[b' = b] \le \frac{1}{2}$.
 $G_0 \not\equiv G_1$: $Pr[b' = b] = 1$ (i.e., P can, possibly inefficiently, extracted from $\pi(E_i)$)



Part II

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- ▶ Public-coin interactive proof/argument in each round the verifier flips coins and sends them to the prover. To compute its output, the verifier applies some (fixed) function to the protocol's transcript.

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- ▶ In the following we focus on games (no input protocols)

Section 2

Parallel repetition of public-coin interactive argument



Theorem 6

Let $\pi = (P, V)$ be m-round, public-coin protocol with $\Pr\left[(\widetilde{P}, V) = 1\right] \le \varepsilon$ for any t-time \widetilde{P} , then $\Pr\left[(\widetilde{P^{(k)}}, V^{(k)}) = 1^k\right] \le \varepsilon^{k/4}$ for any $t \cdot \frac{\varepsilon^{k/4}}{mk^3t_V}$ -time $\widetilde{P^{(k)}}$, where t_V is V's running time.

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Proof plan: Let $\widetilde{\mathsf{P}^{(k)}}$ be $t^{(k)}$ -time algorithm with $\Pr\left[\left(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}\right)=1^k\right]=\varepsilon^{(k)}$, we construct $t^{(k)}\cdot\frac{mk^3t_V}{\varepsilon^{(k)}}$ -time $\widetilde{\mathsf{P}}$ with $\Pr\left[\left(\widetilde{\mathsf{P}},\mathsf{V}\right)=1\right]\geq (\varepsilon^{(k)})^{4/k}$.

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- Assume for simplicity that $P^{(k)}$ is deterministic
- Assume wlg. that V sends the first message in π and that in each round it samples and sends ℓ coins.
- ▶ We view the coins of $V^{(k)}$ as a matrix $R \in \{0,1\}^{m \times (k\ell)}$, letting R_j denote the coins of the j'th round, and $R_{1,...,j}$ the coins of the first j rounds.

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- ▶ Let $\mathbf{R} \sim \{0,1\}^{m \times (k\ell)}$

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- 1. Let $i^* \leftarrow [k]$.
- **2.** Upon getting the j'th message r from V, do:
 - **2.1** Let $R \leftarrow \{0,1\}^{m \times (k\ell)}$, conditioned that $R_{1,\dots,j-1} = \widetilde{R}_{1,\dots,j-1}$ and $R_{j,j^*} = r$.
 - **2.2** If $(P^{(k)}, V^{(k)}(R)) = 1^k$:
 - **2.2.1** Set $\widetilde{R}_j = R_j$
 - **2.2.2** Send a_{j,i^*} back to V, for a_j being the j'th message $P^{(k)}$ send to $V^{(k)}$ in $(P^{(k)}, V^{(k)}(R))$.

Else, GOTO Line 2.1

2.3 Abort if the overall number of sampling exceeds $\lceil qm/\varepsilon^{(k)} \rceil$.

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- **2.3** Abort if the overall number of sampling exceeds $\lceil qm/\varepsilon^{(k)} \rceil$.
- Let \widetilde{P}' be the non aborting variant of \widetilde{P}' , let \widetilde{R} and \widetilde{N} be the value of \widetilde{R} and # of samples done in a random execution of $(\widetilde{P}', V^{(k)})$.
- $\qquad \qquad \Pr\left[(\widetilde{\mathsf{P}},\mathsf{V})=1\right] \geq \Pr\left[\mathsf{win}(\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) := (\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\widetilde{\mathbf{R}})) = 1^k \wedge \widetilde{\mathbf{N}} \leq qm/\varepsilon^{(k)} \right].$

Experiment 8 (P)

- 1. Let $R \leftarrow \{0,1\}^{m \times \ell}$, conditioned that $R_{1,\dots,j-1} = \hat{R}_{1,\dots,j-1}$.
- **2.** If $(\widetilde{P^{(k)}}, V^{(k)}(R)) = 1^k$, set $\hat{R}_j = R_j$. Else, GOTO Line 1.

Experiment 8 (P)

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- $\blacktriangleright \ \hat{\boldsymbol{R}} \sim \boldsymbol{R}|_{\widetilde{(P^{(k)},V^{(k)}(\boldsymbol{R}))}=1^k}$

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- ► In particular, $\Pr\left[(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\hat{\mathbf{R}})=1^k\right]=1$

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- ▶ Let N̂ be # of samples done in R̂.

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- ► In particular, $\Pr\left[(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\hat{\mathbf{R}})=1^k\right]=1$
- ▶ Let \hat{N} be # of samples done in \hat{R} .

Experiment 8 (P)

For j = 1 to m:

- 1. Let $R \leftarrow \{0,1\}^{m \times \ell}$, conditioned that $R_{1,\dots,j-1} = \hat{R}_{1,\dots,j-1}$.
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- $\blacktriangleright |\hat{\mathbf{R}} \sim \mathbf{R}|_{(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\mathbf{R}))=1^k}$
- ► In particular, $Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(\hat{\mathbf{R}}) = 1^k\right] = 1$
- ▶ Let N̂ be # of samples done in R̂.

Lemma 9

$$\Pr\left[\hat{\mathbf{N}} \leq qm/arepsilon^{(k)}
ight] \geq 1 - rac{1}{q}$$

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▶ Let $(X_1, ..., X_m) = \mathbf{R}$ and $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$

- ▶ Let $(X_1, ..., X_m) = \mathbf{R}$ and $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$
- ▶ $v(\mathbf{y} = (y_1, ..., y_j)) := \Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(X^m)) = 1^k \mid X^j = \mathbf{y}\right]$ (letting $X^j = (X_1, ..., X_j)$)

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- Conditioned on $Y^j = \mathbf{y} = (y_1, \dots, y_j)$, the expected # of samples done in (j+1)'th round of \widehat{P} is $\frac{1}{\nu(\mathbf{y})}$.
- ▶ We prove Lemma 9 showing that $E\left[\frac{1}{\nu(Y')}\right] \leq \frac{1}{\varepsilon^{(k)}}$ for every $j \in \{0, \dots, m-1\}$

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Claim 10

For $j \in \{0, \dots, m-1\}$ and $\mathbf{y} \in \operatorname{Supp}(Y^j)$ it holds that $\Pr_{Y^j}[\mathbf{y}] = \Pr_{X^j}[\mathbf{y}] \cdot \frac{\nu(\mathbf{y})}{\varepsilon^{(k)}}$

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Hence,
$$\mathsf{E}_{Y^j}\left[\frac{1}{v(Y^j)}\right] = \sum_{\mathbf{y} \in \mathsf{Supp}(Y^j)} \mathsf{Pr}[Y^j = \mathbf{y}] \cdot \frac{1}{v(\mathbf{y})}$$

- ► Let $(X_1, ..., X_m) = \mathbf{R}$ and $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$
- ► $v(\mathbf{y} = (y_1, ..., y_j)) := \Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(X^m)) = 1^k \mid X^j = \mathbf{y}\right]$ (letting $X^j = (X_1, ..., X_j)$)
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Hence,
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$$= \sum_{\mathbf{y}} \mathsf{Pr}[\mathsf{X}^j = \mathbf{y}] \cdot \frac{\nu(\mathbf{y})}{\varepsilon^{(k)}} \cdot \frac{1}{\nu(\mathbf{y})} = \frac{1}{\varepsilon^{(k)}} \cdot \sum_{\mathbf{y} \in \mathsf{Supp}(\mathsf{Y}^j)} \mathsf{Pr}[\mathsf{X}^j = \mathbf{y}] \leq \frac{1}{\varepsilon^{(k)}}. \ \Box$$

Note that

$$\Pr_{Y_{j}|Y^{j-1}=\mathbf{y}_{1...,j-1}}[y_{j}] = \sum_{\ell=1}^{\infty} (1 - v(\mathbf{y}_{1...,j-1}))^{\ell-1} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1...,j-1}}[y_{j}] \cdot v(\mathbf{y}) \qquad (1)$$

$$= \frac{1}{v(\mathbf{y}_{1...,j-1})} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1...,j-1}}[y_{j}] \cdot v(\mathbf{y})$$

Note that

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$$\Pr_{Y_j}[\mathbf{y}] = \Pr_{Y_{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \Pr_{Y_j|Y^{j-1} = \mathbf{y}_{1...,j-1}}[y_j]$$

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$$= \Pr_{X_{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \frac{v(\mathbf{y}_{1...,j-1})}{\varepsilon^{(k)}} \cdot \Pr_{Y_{j}|Y^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}]$$
(i.h.)

Note that

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$$\begin{aligned}
&\Pr_{\mathbf{y}_{j}}[\mathbf{y}] = \Pr_{\mathbf{y}_{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \Pr_{\mathbf{y}_{j}|\mathbf{y}_{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \\
&= \Pr_{\mathbf{x}_{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \frac{\mathbf{v}(\mathbf{y}_{1...,j-1})}{\varepsilon^{(k)}} \cdot \Pr_{\mathbf{y}_{j}|\mathbf{y}_{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \\
&= \Pr_{\mathbf{x}_{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \frac{\mathbf{v}(\mathbf{y}_{1...,j-1})}{\varepsilon^{(k)}} \cdot \frac{\mathbf{v}(\mathbf{y})}{\mathbf{v}(\mathbf{y}_{1...,j-1})} \cdot \Pr_{\mathbf{x}_{j}|\mathbf{x}_{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \quad \text{(Eq. (1))}
\end{aligned}$$

Note that

$$\Pr_{Y_{j}|Y^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] = \sum_{\ell=1}^{\infty} (1 - v(\mathbf{y}_{1,...,j-1}))^{\ell-1} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] \cdot v(\mathbf{y}) \qquad (1)$$

$$= \frac{1}{v(\mathbf{y}_{1,...,j-1})} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] \cdot v(\mathbf{y})$$

$$\begin{aligned}
&\Pr_{\mathbf{Y}^{j}}[\mathbf{y}] = \Pr_{\mathbf{Y}^{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \Pr_{\mathbf{Y}_{j}|\mathbf{Y}^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \\
&= \Pr_{\mathbf{X}^{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \frac{v(\mathbf{y}_{1...,j-1})}{\varepsilon^{(k)}} \cdot \Pr_{\mathbf{Y}_{j}|\mathbf{Y}^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \\
&= \Pr_{\mathbf{X}^{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \frac{v(\mathbf{y}_{1...,j-1})}{\varepsilon^{(k)}} \cdot \frac{v(\mathbf{y})}{v(\mathbf{y}_{1...,j-1})} \cdot \Pr_{\mathbf{X}_{j}|\mathbf{X}^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \quad (\text{Eq. (1)}) \\
&= \Pr_{\mathbf{Y}^{j}}[\mathbf{y}] \cdot \frac{v(\mathbf{y})}{\varepsilon^{(k)}}.
\end{aligned}$$

Experiment 11 (P)

- 1. Let $i^* \leftarrow [k]$.
- **2.** For for j = 1 to m:
 - **2.1** Let $R \leftarrow \{0,1\}^{m \times \ell}$, conditioned on $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$.
 - **2.2** If $(P^{(k)}, V^{(k)}(R)) = 1^k$, set $\widehat{R}_{j,j^*} = R_{j,j^*}$. Else, GOTO Line 2.1.
 - **2.3** Let $R \leftarrow \{0,1\}^{m \times \ell}$, conditioned on $R_{1,...,j-1} = \widehat{R}_{1,...,j-1}$ and $R_{j,i^*} = \widehat{R}_{j,i^*}$.
 - **2.4** If $(P^{(k)}, V^{(k)}(R)) = 1^k$, set $\hat{R}_i = R_i$. Else, GOTO Line 2.3.

Experiment 11 (P)

- **1.** Let $i^* \leftarrow [k]$.
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 - **2.4** If $(P^{(k)}, V^{(k)}(R)) = 1^k$, set $\hat{R}_i = R_i$. Else, GOTO Line 2.3.
- Let $\widehat{\mathbf{R}}$ be the final value of $\widehat{\mathbf{R}}$ in $\widehat{\mathbf{P}}$.

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- $\blacktriangleright \ \widehat{\boldsymbol{\mathsf{R}}} \sim \boldsymbol{\mathsf{R}}|_{(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\boldsymbol{\mathsf{R}}))=1^k}$

Let $\hat{\mathbf{R}}$ be the final value of $\hat{\mathbf{R}}$ in $\hat{\mathbf{P}}$.

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- 1. Let $i^* \leftarrow [k]$.
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- ▶ Let $\hat{\mathbf{R}}$ be the final value of $\hat{\mathbf{R}}$ in $\hat{\mathbf{P}}$.
- $ho \ \widehat{f R} \sim {f R}|_{\widetilde{(P^{(k)},V^{(k)}({f R}))}=1^k}$
- ▶ Let \hat{N} be the # of Step-2.3-samples done in \hat{P} .

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 - **2.4** If $(P^{(k)}, V^{(k)}(R)) = 1^k$, set $\hat{R}_j = R_j$. Else, GOTO Line 2.3.
 - Let $\hat{\mathbf{R}}$ be the final value of $\hat{\mathbf{R}}$ in $\hat{\mathbf{P}}$.
- $\blacktriangleright \ \widehat{\boldsymbol{\mathsf{R}}} \sim \boldsymbol{\mathsf{R}}|_{(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\boldsymbol{\mathsf{R}}))=1^k}$
- ▶ Let \hat{N} be the # of Step-2.3-samples done in \hat{P} .

Lemma 12

$$\Pr\left[\text{win}(\widehat{\pmb{R}}, \widehat{\pmb{N}}) \right] \ge 1 - \frac{1}{q}$$

Let
$$\widetilde{\mathbf{R}}_i = \widetilde{\mathbf{R}}|_{i^*=i}$$
 and $\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$ $(= \widehat{\mathbf{R}})$.

Let
$$\widetilde{\mathbf{R}}_i = \widetilde{\mathbf{R}}|_{j^*=i}$$
 and $\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{j^*=i}$ (= $\widehat{\mathbf{R}}$).

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}||\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i||\widetilde{\mathbf{R}}_i).$$

Let
$$\widetilde{\mathbf{R}}_i = \widetilde{\mathbf{R}}|_{i^*=i}$$
 and $\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$ $(= \widehat{\mathbf{R}})$.

Claim 13

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

$$\sum_{i\in[k]}D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i)\leq D(\widehat{\mathbf{R}}||\mathbf{R}).$$

Let
$$\widetilde{\mathbf{R}}_i = \widetilde{\mathbf{R}}|_{i^*=i}$$
 and $\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$ $(= \widehat{\mathbf{R}})$.

Claim 13

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

Claim 14

$$\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i) \leq D(\widehat{\mathbf{R}}||\mathbf{R}).$$

► Thm. 7 in Lecture 7 $\implies D(\widehat{\mathbf{R}}||\mathbf{R}) \leq \log \frac{1}{\Pr[(\widehat{\mathbf{P}^{(k)}}, \mathbf{V}^{(k)}(\mathbf{R})) = 1^k]} = \log \frac{1}{\varepsilon^{(k)}}$

Let
$$\widetilde{\mathbf{R}}_i = \widetilde{\mathbf{R}}|_{i^*=i}$$
 and $\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$ $(= \widehat{\mathbf{R}})$.

Claim 13

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

$$\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i) \leq D(\widehat{\mathbf{R}}||\mathbf{R}).$$

- ▶ Thm. 7 in Lecture 7 $\implies D(\widehat{\mathbf{R}}||\mathbf{R}) \le \log \frac{1}{\Pr[(\widetilde{P^{(k)}}, \mathbf{V}^{(k)}(\mathbf{R})) = 1^k]} = \log \frac{1}{\varepsilon^{(k)}}$
- ► Hence, $D(\text{win}(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})||\text{win}(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})) \leq D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}||\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq -\frac{1}{k} \cdot \log \varepsilon^{(k)}$

Let
$$\widetilde{\mathbf{R}}_i = \widetilde{\mathbf{R}}|_{i^*=i}$$
 and $\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$ (= $\widehat{\mathbf{R}}$).

Claim 13

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

$$\sum_{i\in[k]}D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i)\leq D(\widehat{\mathbf{R}}||\mathbf{R}).$$

- ► Thm. 7 in Lecture 7 $\implies D(\widehat{\mathbf{R}}||\mathbf{R}) \le \log \frac{1}{\Pr[(\widehat{P^{(k)}}, V^{(k)}(\mathbf{R})) = 1^k]} = \log \frac{1}{\varepsilon^{(k)}}$
- ► Hence, $D(\text{win}(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})||\text{win}(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})) \leq D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}||\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq -\frac{1}{k} \cdot \log \varepsilon^{(k)}$
- ► Claim 12 $\implies \alpha := \Pr[\text{win}(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})] \ge 1 \frac{1}{q}$, and let $\beta := \Pr[\text{win}(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})]$.

Let
$$\widetilde{\mathbf{R}}_i = \widetilde{\mathbf{R}}|_{i^*=i}$$
 and $\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$ (= $\widehat{\mathbf{R}}$).

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$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

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- ► Hence, $D(\text{win}(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})||\text{win}(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})) \leq D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}||\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq -\frac{1}{k} \cdot \log \varepsilon^{(k)}$
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- ▶ Thus, $\alpha \cdot \log \frac{\alpha}{\beta} + (1 \alpha) \log(1 \alpha) \le -\frac{1}{k} \cdot \log \varepsilon^{(k)}$

Let
$$\widetilde{\mathbf{R}}_i = \widetilde{\mathbf{R}}|_{i^*=i}$$
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$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}||\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i||\widetilde{\mathbf{R}}_i).$$

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- ► Thm. 7 in Lecture 7 $\implies D(\widehat{\mathbf{R}}||\mathbf{R}) \leq \log \frac{1}{\Pr[(\widetilde{P^{(k)}}, V^{(k)}(\mathbf{R})) = 1^k]} = \log \frac{1}{\varepsilon^{(k)}}$
- ► Hence, $D(\text{win}(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})||\text{win}(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})) \le D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}||\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \le -\frac{1}{k} \cdot \log \varepsilon^{(k)}$
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- ► Thus, $\alpha \cdot \log \frac{\alpha}{\beta} + (1 \alpha) \log(1 \alpha) \le -\frac{1}{k} \cdot \log \varepsilon^{(k)}$

$$\implies \beta \geq 2^{\log \alpha + \frac{1-\alpha}{\alpha} \log(1-\alpha) + \frac{1}{\alpha k} \log \varepsilon^{(k)}}$$

Let
$$\widetilde{\mathbf{R}}_i = \widetilde{\mathbf{R}}|_{i^*=i}$$
 and $\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$ (= $\widehat{\mathbf{R}}$).

Claim 13

$$D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i||\widetilde{\mathbf{R}}_i).$$

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- ► Hence, $D(\text{win}(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})||\text{win}(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})) \leq D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}||\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq -\frac{1}{k} \cdot \log \varepsilon^{(k)}$
- ► Claim 12 $\implies \alpha := \Pr[\text{win}(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})] \ge 1 \frac{1}{q}$, and let $\beta := \Pr[\text{win}(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})]$.
- ► Thus, $\alpha \cdot \log \frac{\alpha}{\beta} + (1 \alpha) \log (1 \alpha) \le -\frac{1}{k} \cdot \log \varepsilon^{(k)}$ $\implies \beta \ge 2^{\log \alpha + \frac{1 \alpha}{\alpha} \log (1 \alpha) + \frac{1}{\alpha k} \log \varepsilon^{(k)}}$
- ► Recalling $q = k^2$, $\alpha \ge 2^{-\frac{2}{q}} \ge 2^{-\frac{1}{k}}$ and $\frac{1-\alpha}{\alpha} \log(1-\alpha) \ge -\frac{4 \log k}{k^2} \ge -\frac{1}{k}$

Let
$$\widetilde{\mathbf{R}}_i = \widetilde{\mathbf{R}}|_{i^*=i}$$
 and $\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$ $(= \widehat{\mathbf{R}})$.

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$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

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- ► Hence, $D(\text{win}(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})||\text{win}(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})) \le D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}||\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \le -\frac{1}{k} \cdot \log \varepsilon^{(k)}$
- ► Claim 12 $\implies \alpha := \Pr[\text{win}(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})] \ge 1 \frac{1}{q}$, and let $\beta := \Pr[\text{win}(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})]$.
- ► Thus, $\alpha \cdot \log \frac{\alpha}{\beta} + (1 \alpha) \log(1 \alpha) \le -\frac{1}{k} \cdot \log \varepsilon^{(k)}$ ⇒ $\beta \ge 2^{\log \alpha + \frac{1 - \alpha}{\alpha} \log(1 - \alpha) + \frac{1}{\alpha k} \log \varepsilon^{(k)}}$
- ▶ Recalling $q = k^2$, $\alpha \ge 2^{-\frac{2}{q}} \ge 2^{-\frac{1}{k}}$ and $\frac{1-\alpha}{\alpha} \log(1-\alpha) \ge -\frac{4\log k}{k^2} \ge -\frac{1}{k}$
- We conclude that $\beta \geq 2^{\frac{4}{k}\log \varepsilon^{(k)}} = \sqrt[k/4]{\varepsilon^{(k)}}$.

HW...

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Lemma 15

Let $Z = \{Z_{ij}\}_{(i,j)\in[k]\times[m]}$ be iids, let W be an event, and let

$$D_i(z) := \textstyle \prod_{j=1}^m \text{Pr}\left[Z_{j,i} = z_{i,j}\right] \cdot \text{Pr}\left[Z_{j,-i} = z_{i,j-1} | Z_{1,\dots,j-1} = z_{1,\dots,j-1} \wedge Z_{j,i} = z_{i,j} \wedge W\right].$$

Then $\sum_{i=1}^{k} D(Z_{W}||D_{i}) \leq D(Z_{W}||Z)$.

Lemma 15

Let $Z = \{Z_{ij}\}_{(i,j) \in [k] \times [m]}$ be iids, let W be an event, and let

$$D_i(z) := \prod_{j=1}^m \Pr\left[Z_{j,i} = z_{i,j} \right] \cdot \Pr\left[Z_{j,-i} = z_{i,j-1} | Z_{1,...,j-1} = z_{1,...,j-1} \wedge Z_{j,i} = z_{i,j} \wedge W \right].$$

Then $\sum_{i=1}^{k} D(Z_{W}||D_{i}) \leq D(Z_{W}||Z)$.

Letting $Z = \mathbf{R}$ and W be the event $(\widetilde{\mathsf{P}^{(k)}}, \mathsf{V}^{(k)}(\mathbf{R})) = 1^k$, Lemma 15 yields that $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widetilde{\mathbf{R}}_i) \le D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$. \square

Lemma 15

Let $Z = \{Z_{ij}\}_{(i,j) \in [k] \times [m]}$ be iids, let W be an event, and let

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Then
$$\sum_{i=1}^{k} D(Z_{W}||D_{i}) \leq D(Z_{W}||Z)$$
.

Letting $Z = \mathbf{R}$ and W be the event $(\widehat{\mathsf{P}^{(k)}}, \mathsf{V}^{(k)}(\mathbf{R})) = 1^k$, Lemma 15 yields that $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widehat{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widehat{\mathbf{R}}_i) \leq D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$. \square

Lemma 15

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Letting $Z = \mathbf{R}$ and W be the event $(\widehat{\mathsf{P}^{(k)}}, \mathsf{V}^{(k)}(\mathbf{R})) = 1^k$, Lemma 15 yields that $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widehat{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widehat{\mathbf{R}}_i) \leq D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$. \square

Proof: (of Lemma 15) We prove for m = k = 2.

▶ Let $X = Z_1$ and $Y = Z_2$

Lemma 15

Let $Z = \{Z_{ij}\}_{(i,j) \in [k] \times [m]}$ be iids, let W be an event, and let

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Then
$$\sum_{i=1}^{k} D(Z_{W}||D_{i}) \leq D(Z_{W}||Z)$$
.

Letting $Z = \mathbf{R}$ and W be the event $(\widehat{P^{(k)}}, V^{(k)}(\mathbf{R})) = 1^k$, Lemma 15 yields that $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widehat{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widehat{\mathbf{R}}_i) \le D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$. \square

- ► Let $X = Z_1$ and $Y = Z_2$
- $U(x_1, x_2, y_1, y_2) := \Pr_{(X,Y)} [(x_1, x_2, y_1, y_2)]$

Lemma 15

Let $Z = \{Z_{ij}\}_{(i,j)\in[k]\times[m]}$ be iids, let W be an event, and let

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Then
$$\sum_{i=1}^{k} D(Z_{W}||D_{i}) \leq D(Z_{W}||Z)$$
.

Letting $Z = \mathbf{R}$ and W be the event $(\widehat{P^{(k)}}, V^{(k)}(\mathbf{R})) = 1^k$, Lemma 15 yields that $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widehat{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widehat{\mathbf{R}}_i) \le D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$. \square

- ► Let $X = Z_1$ and $Y = Z_2$
- $U(x_1, x_2, y_1, y_2) := \Pr_{(X,Y)} [(x_1, x_2, y_1, y_2)]$
- $ightharpoonup C(x_1, x_2, y_1, y_1) := (X|_W)(x_1, x_2, y_1, y_1)$

Lemma 15

Let $Z = \{Z_{ij}\}_{(i,j)\in[k]\times[m]}$ be iids, let W be an event, and let

$$D_i(z) := \prod_{j=1}^m \Pr\left[Z_{j,i} = z_{i,j}\right] \cdot \Pr\left[Z_{j,-i} = z_{i,j-1} | Z_{1,...,j-1} = z_{1,...,j-1} \wedge Z_{j,i} = z_{i,j} \wedge W\right].$$

Then
$$\sum_{i=1}^{k} D(Z_{W}||D_{i}) \leq D(Z_{W}||Z)$$
.

Letting
$$Z = \mathbf{R}$$
 and W be the event $(\widehat{\mathsf{P}^{(k)}}, \mathsf{V}^{(k)}(\mathbf{R})) = 1^k$, Lemma 15 yields that $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widehat{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widehat{\mathbf{R}}_i) \leq D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$. \square

- ▶ Let $X = Z_1$ and $Y = Z_2$
- $U(x_1, x_2, y_1, y_2) := \Pr_{(X,Y)} [(x_1, x_2, y_1, y_2)]$
- $C(x_1,x_2,y_1,y_1) := (X|_W)(x_1,x_2,y_1,y_1)$
- $Pr[X_1, x_2, y_1, y_1) := Pr[X_1 = x_1 | W] \cdot Pr[X_2 = x_2 | W] \cdot Pr[Y_1 = y_1 | W, X = (x_1, x_2)] \cdot Pr[Y_2 = y_2 | W, X = (x_1, x_2)]$

Lemma 15

Let
$$Z = \{Z_{ij}\}_{(i,j)\in[k]\times[m]}$$
 be iids, let W be an event, and let

$$D_i(z) := \prod_{j=1}^m \Pr\left[Z_{j,i} = z_{i,j}\right] \cdot \Pr\left[Z_{j,-i} = z_{i,j-1} | Z_{1,...,j-1} = z_{1,...,j-1} \wedge Z_{j,i} = z_{i,j} \wedge W\right].$$

Then
$$\sum_{i=1}^{k} D(Z_{W}||D_{i}) \leq D(Z_{W}||Z)$$
.

Letting
$$Z = \mathbf{R}$$
 and W be the event $(\widehat{\mathsf{P}^{(k)}}, \mathsf{V}^{(k)}(\mathbf{R})) = 1^k$, Lemma 15 yields that $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widehat{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widehat{\mathbf{R}}_i) \le D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$. \square

- ▶ Let $X = Z_1$ and $Y = Z_2$
- $U(x_1, x_2, y_1, y_2) := \Pr_{(X,Y)} [(x_1, x_2, y_1, y_2)]$
- $C(x_1, x_2, y_1, y_1) := (X|_W)(x_1, x_2, y_1, y_1)$
- $P(X_1, X_2, y_1, y_1) := Pr[X_1 = x_1 | W] \cdot Pr[X_2 = x_2 | W] \cdot Pr[Y_1 = y_1 | W, X = (x_1, x_2)] \cdot Pr[Y_2 = y_2 | W, X = (x_1, x_2)]$
- ► We write $\frac{C(x_1, x_2, y_1, y_1)}{U(x_1, x_2, y_1, y_1)} = \frac{\Pr[X_1 = x_1 | W] \cdot \Pr[Y_1 = y_1 | W, X = (x_1, x_2)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \cdot \frac{\Pr[X_2 = x_2 | W] \cdot \Pr[Y_2 = y_2 | W, X = (x_1, x_2)]}{\Pr[X_2 = x_2] \cdot \Pr[Y_2 = y_2]} \cdot \frac{C(x_1, x_2, y_1, y_1)}{Q(x_1, x_2, y_1, y_1)}$

$$\begin{split} D(C||U) &= \mathop{\mathbb{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[\log \frac{\Pr[X_1 = x_1 | W] \cdot \Pr[Y_1 = y_1 | W, X = (x_1, x_2)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \right] \\ &+ \mathop{\mathbb{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[\log \frac{\Pr[X_2 = x_2 | W] \cdot \Pr[Y_2 = y_2 | W, X = (x_1, x_2)]}{\Pr[X_2 = x_2] \cdot \Pr[Y_2 = y_2]} \right] \\ &+ \mathop{\mathbb{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[\log \frac{C(x_1, x_2, y_1, y_2)}{Q(x_1, x_2, y_1, y_2)} \right]. \end{split}$$

$$\begin{split} D(C||U) &= \mathop{\mathsf{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[\log \frac{\Pr[X_1 = x_1 | W] \cdot \Pr[Y_1 = y_1 | W, X = (x_1, x_2)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \right] \\ &+ \mathop{\mathsf{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[\log \frac{\Pr[X_2 = x_2 | W] \cdot \Pr[Y_2 = y_2 | W, X = (x_1, x_2)]}{\Pr[X_2 = x_2] \cdot \Pr[Y_2 = y_2]} \right] \\ &+ \mathop{\mathsf{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[\log \frac{C(x_1, x_2, y_1, y_2)}{Q(x_1, x_2, y_1, y_2)} \right]. \end{split}$$

It follows that

$$D(C||U) = D(X_1|_W, X_2|_{W,X_1}, Y_1|_{W,X}, Y_2|_{W,X,Y_1}||X_1, X_2|_{W,X_1}, Y_1, Y_2|_{W,X,Y_1})$$

$$+ D(X_2|_W, X_1|_{W,X_2}, Y_2|_{W,X}, Y_1|_{W,X,Y_2}||X_2, X_1|_{W,X_2}, Y_2, Y_1|_{W,X,Y_2})$$

$$+ D(C||Q),$$

$$\begin{split} D(C||U) &= \mathop{\mathsf{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[\log \frac{\Pr[X_1 = x_1 | W] \cdot \Pr[Y_1 = y_1 | W, X = (x_1, x_2)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \right] \\ &+ \mathop{\mathsf{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[\log \frac{\Pr[X_2 = x_2 | W] \cdot \Pr[Y_2 = y_2 | W, X = (x_1, x_2)]}{\Pr[X_2 = x_2] \cdot \Pr[Y_2 = y_2]} \right] \\ &+ \mathop{\mathsf{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[\log \frac{C(x_1, x_2, y_1, y_2)}{Q(x_1, x_2, y_1, y_2)} \right]. \end{split}$$

It follows that

$$D(C||U) = D(X_1|_W, X_2|_{W,X_1}, Y_1|_{W,X}, Y_2|_{W,X,Y_1}||X_1, X_2|_{W,X_1}, Y_1, Y_2|_{W,X,Y_1})$$

$$+ D(X_2|_W, X_1|_{W,X_2}, Y_2|_{W,X}, Y_1|_{W,X,Y_2}||X_2, X_1|_{W,X_2}, Y_2, Y_1|_{W,X,Y_2})$$

$$+ D(C||Q),$$

and the proof follows since $D(\cdot||\cdot) \geq 0$. \square

Similar proof to the public-coin proof we gave above.

- Similar proof to the public-coin proof we gave above.
- ▶ In each round, the attacker \widetilde{P} samples random continuations of $(\widetilde{P^{(k)}}, V^{(k)})$, till he gets an accepting execution.

- Similar proof to the public-coin proof we gave above.
- ▶ In each round, the attacker \widetilde{P} samples random continuations of $(\widetilde{P^{(k)}}, V^{(k)})$, till he gets an accepting execution.
- Why fails us to extend this approach for non-public-coin interactive arguments?

Section 3

Parallel amplification for any interactive argument



Parallel amplification theorem for any protocol

Can we amplify the security of any interactive argument "in parallel"?

Parallel amplification theorem for any protocol

- Can we amplify the security of any interactive argument "in parallel"?
- Yes we can!