

# Foundation of Cryptography, Lecture 2

## Pseudorandom Generators<sup>1</sup>

### Handout Mode

Iftach Haitner

Tel Aviv University.

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# Part I

## **Statistical Vs. Computational distance**

# Section 1

## **Distributions and Statistical Distance**

## Distributions and Statistical Distance

Let  $P$  and  $Q$  be two distributions over a finite set  $\mathcal{U}$ . Their **statistical distance** (also known as, variation distance) is defined as

$$\text{SD}(P, Q) := \frac{1}{2} \sum_{x \in \mathcal{U}} |P(x) - Q(x)| = \max_{S \subseteq \mathcal{U}} (P(S) - Q(S))$$

We will only consider **finite** distributions.

### Claim 1

For any pair of (finite) distribution  $P$  and  $Q$ , it holds that

$$\text{SD}(P, Q) = \max_D \{ \Pr_{x \leftarrow P}[D(x) = 1] - \Pr_{x \leftarrow Q}[D(x) = 1] \},$$

where  $D$  is **any** algorithm.

Interpretation?

## Some useful facts

Let  $P, Q, R$  be finite distributions, then

**Triangle inequality:**

$$\text{SD}(P, R) \leq \text{SD}(P, Q) + \text{SD}(Q, R)$$

**Repeated sampling:**

$$\text{SD}((P, P), (Q, Q)) \leq 2 \cdot \text{SD}(P, Q)$$

# Distribution ensembles and statistical indistinguishability

## Definition 2 (distribution ensembles)

$\mathcal{P} = \{P_n\}_{n \in \mathbb{N}}$  is a **distribution ensemble**, if  $P_n$  is a (finite) distribution for any  $n \in \mathbb{N}$ .

$\mathcal{P}$  is **efficiently samplable** (or just efficient), if  $\exists$  PPT  $Samp$  with  $Sam(1^n) \equiv P_n$ .

## Definition 3 (statistical indistinguishability)

Two distribution ensembles  $\mathcal{P}$  and  $\mathcal{Q}$  are **statistically indistinguishable**, if  $SD(P_n, Q_n) = \text{neg}(n)$ .

Alternatively, if  $\left| \Delta_{(\mathcal{P}, \mathcal{Q})}^D(n) \right| = \text{neg}(n)$ , for *any* algorithm  $D$ , where

$$\Delta_{(\mathcal{P}, \mathcal{Q})}^D(n) := \Pr_{x \leftarrow P_n} [D(1^n, x) = 1] - \Pr_{x \leftarrow Q_n} [D(1^n, x) = 1] \quad (1)$$

## Section 2

# Computational Indistinguishability

# Computational Indistinguishability

## Definition 4 (computational indistinguishability)

Two distribution ensembles  $\mathcal{P}$  and  $\mathcal{Q}$  are **computationally indistinguishable**, if  $|\Delta_{(\mathcal{P}, \mathcal{Q})}^D(n)| = \text{neg}(n)$ , for any **PPT**  $D$ .

- ▶ Can it be different from the statistical case?
- ▶ Non uniform variant
- ▶ Sometime behaves different then expected!



# Repeated sampling

## Question 5

Assume that  $\mathcal{P}$  and  $\mathcal{Q}$  are computationally indistinguishable, is it always true that  $\mathcal{P}^2 = (\mathcal{P}, \mathcal{P})$  and  $\mathcal{Q}^2 = (\mathcal{Q}, \mathcal{Q})$  are?

Let  $D$  be an algorithm and let  $\delta(n) = \left| \Delta_{(\mathcal{P}^2, \mathcal{Q}^2)}^D(n) \right|$

$$\begin{aligned}\delta(n) &= \left| \Pr_{x \leftarrow \mathcal{P}_n^2} [D(x) = 1] - \Pr_{x \leftarrow \mathcal{Q}_n^2} [D(x) = 1] \right| \\ &\leq \left| \Pr_{x \leftarrow \mathcal{P}_n^2} [D(x) = 1] - \Pr_{x \leftarrow (\mathcal{P}_n, \mathcal{Q}_n)} [D(x) = 1] \right| \\ &\quad + \left| \Pr_{x \leftarrow (\mathcal{P}_n, \mathcal{Q}_n)} [D(x) = 1] - \Pr_{x \leftarrow \mathcal{Q}_n^2} [D(x) = 1] \right| \\ &= \left| \Delta_{(\mathcal{P}^2, (\mathcal{P}, \mathcal{Q}))}^D(n) \right| + \left| \Delta_{((\mathcal{P}, \mathcal{Q}), \mathcal{Q}^2)}^D(n) \right|\end{aligned}$$

So either  $\left| \Delta_{(\mathcal{P}^2, (\mathcal{P}, \mathcal{Q}))}^D(n) \right| \geq \delta(n)/2$ , or  $\left| \Delta_{((\mathcal{P}, \mathcal{Q}), \mathcal{Q}^2)}^D(n) \right| \geq \delta(n)/2$

- ▶ Assume  $D$  is a PPT and that  $\left| \Delta_{(\mathcal{P}^2, \mathcal{Q}^2)}^D(n) \right| \geq 1/p(n)$  for some  $p \in \text{poly}$  and infinitely many  $n$ 's, and assume wlg. that  $\left| \Delta_{\mathcal{P}^2, (\mathcal{P}, \mathcal{Q})}^D(n) \right| \geq 1/2p(n)$  for infinitely many  $n$ 's.
- ▶ Can we use  $D$  to contradict the fact that  $\mathcal{P}$  and  $\mathcal{Q}$  are computationally close?
- ▶ Assuming that  $\mathcal{P}$  and  $\mathcal{Q}$  are efficiently samplable
- ▶ Non-uniform settings

## Repeated sampling cont.

Given  $t = t(n) \in \mathbb{N}$  and a distribution ensemble  $\mathcal{P} = \{P_n\}_{n \in \mathbb{N}}$ , let  $\mathcal{P}^t = \{P_n^{t(n)}\}_{n \in \mathbb{N}}$ .

### Question 6

Let  $t = t(n) \leq \text{poly}(n)$  be an eff. computable integer function. Assume that  $\mathcal{P}$  and  $\mathcal{Q}$  are eff. samplable and computationally indistinguishable, does it mean that  $\mathcal{P}^t$  and  $\mathcal{Q}^t$  are?

Proof:

- ▶ Induction?
- ▶ Hybrid

## Hybrid argument

Let  $D$  be an algorithm and let  $\delta(n) = \left| \Delta_{(\mathcal{P}^t, \mathcal{Q}^t)}^D(n) \right|$ .

- ▶ Fix  $n \in \mathbb{N}$ , and for  $i \in \{0, \dots, t = t(n)\}$ , let  $H^i = (p_1, \dots, p_i, q_{i+1}, \dots, q_t)$ , where the  $p$ 's [resp.,  $q$ 's] are uniformly (and independently) chosen from  $\mathcal{P}_n$  [resp., from  $\mathcal{Q}_n$ ].
- ▶ Since  $\delta(n) = \left| \Delta_{H^t, H^0}^D(t) \right| = \left| \sum_{i \in [t]} \Delta_{H^i, H^{i-1}}^D(t) \right|$ , there exists  $i \in [t]$  with  $\left| \Delta_{H^i, H^{i-1}}^D(t) \right| \geq \delta(n)/t(n)$ .
- ▶ How do we use it?

## Using hybrid argument via estimation

### Algorithm 7 ( $D'$ )

Input:  $1^n$  and  $x \in \{0, 1\}^*$

1. Find  $i \in [t]$  with  $\left| \Delta_{H^i, H^{i-1}}^D(t) \right| \geq \delta(n)/2t(n)$
2. Let  $(p_1, \dots, p_i, q_{i+1}, \dots, q_t) \leftarrow H^i$
3. Return  $D(1^t, p_1, \dots, p_{i-1}, x, q_{i+1}, \dots, q_t), .$

1. how do we find  $i$ ? why  $\delta(n)/2t(n)$
2. Easy in the non-uniform case

## Using hybrid argument via **sampling**

### Algorithm 8 ( $D'$ )

Input:  $1^n$  and  $x \in \{0, 1\}^*$

1. Sample  $i \leftarrow [t = t(n)]$
2. Let  $(p_1, \dots, p_i, q_{i+1}, \dots, q_t) \leftarrow H^i$
3. Return  $D(1^t, p_1, \dots, p_{i-1}, x, q_{i+1}, \dots, q_t)$ .

$$\begin{aligned} \left| \Delta_{(\mathcal{P}, \mathcal{Q})}^{D'}(n) \right| &= \left| \Pr_{p \leftarrow P_n} [D'(p) = 1] - \Pr_{q \leftarrow Q_n} [D'(q) = 1] \right| \\ &= \left| \frac{1}{t} \sum_{i \in [t]} \Pr_{x \leftarrow H_i} [D(x) = 1] - \frac{1}{t} \sum_{i \in [t]} \Pr_{x \leftarrow H_{i-1}} [D(x) = 1] \right| \\ &= \left| \frac{1}{t} \left( \Pr_{x \leftarrow H_t} [D(x) = 1] - \Pr_{x \leftarrow H_0} [D(x) = 1] \right) \right| \\ &= \delta(n)/t(n) \end{aligned}$$

# Part II

## Pseudorandom Generators

# Pseudorandom generator

## Definition 9 (pseudorandom distributions)

A distribution ensemble  $\mathcal{P}$  over  $\{\{0, 1\}^{\ell(n)}\}_{n \in \mathbb{N}}$  is **pseudorandom**, if it is computationally indistinguishable from  $\{U_{\ell(n)}\}_{n \in \mathbb{N}}$ .

- ▶ Do such distributions exist?

## Definition 10 (pseudorandom generators (PRGs))

An efficiently computable function  $g : \{0, 1\}^n \mapsto \{0, 1\}^{\ell(n)}$  is a **pseudorandom generator**, if

- ▶  $g$  is length extending (i.e.,  $\ell(n) > n$  for any  $n$ )
- ▶  $g(U_n)$  is pseudorandom

- ▶ Do such generators exist?
- ▶ Imply one-way functions (homework)
- ▶ Do they have any use?



## Section 3

# Hardcore Predicates

# Hardcore predicates

## Definition 11 (hardcore predicates)

An efficiently computable function  $b : \{0, 1\}^n \mapsto \{0, 1\}$  is a **hardcore predicate** of  $f : \{0, 1\}^n \mapsto \{0, 1\}^n$ , if

$$\Pr_{x \leftarrow \{0, 1\}^n} [P(f(x)) = b(x)] \leq \frac{1}{2} + \text{neg}(n),$$

for any PPT  $P$ .

- ▶ Does the existence of a hardcore predicate for  $f$ , implies that  $f$  is one way? If  $f$  is injective?
- ▶ Fact: any OWF has a hardcore predicate (next class)
- ▶ Building blocks in constructions of PRGS from OWF

## Section 4

# **PRGs from OWPs**

## Claim 12

Let  $f : \{0, 1\}^n \mapsto \{0, 1\}^n$  be an eff. permutation and let  $b : \{0, 1\}^n \mapsto \{0, 1\}$  be a hardcore predicate for  $f$ , then  $g(x) = (f(x), b(x))$  is a PRG.

Proof: Assume  $\exists$  a PPT  $D$ , and infinite set  $\mathcal{I} \subseteq \mathbb{N}$  and  $p \in \text{poly}$  with

$$\left| \Delta_{g(U_n), U_{n+1}}^D \right| > \varepsilon(n) = 1/p(n)$$

for any  $n \in \mathcal{I}$ . We use  $D$  for breaking the hardness of  $b$ .

- We assume wlg. that  $\Pr[D(g(U_n)) = 1] - \Pr[D(U_{n+1}) = 1] \geq \varepsilon(n)$  for any  $n \in \mathcal{I}$  (?), and fix  $n \in \mathcal{I}$ .

## OWP to PRG cont.

- ▶ Let  $\delta(n) = \Pr[D(U_{n+1}) = 1]$  (note that  $\Pr[D(g(U_n)) = 1] = \delta + \varepsilon$ ).
- ▶ Compute

$$\begin{aligned}\delta &= \Pr[D(f(U_n), U_1) = 1] && (?) \\ &= \Pr[U_1 = b(U_n)] \cdot \Pr[D(f(U_n), U_1) = 1 \mid U_1 = b(U_n)] \\ &\quad + \Pr[U_1 = \overline{b(U_n)}] \cdot \Pr[D(f(U_n), U_1) = 1 \mid U_1 = \overline{b(U_n)}] \\ &= \frac{1}{2}(\delta + \varepsilon) + \frac{1}{2} \cdot \Pr[D(f(U_n), U_1) = 1 \mid U_1 = \overline{b(U_n)}].\end{aligned}$$

Hence,

$$\Pr[D(f(U_n), \overline{b(U_n)}) = 1] = \delta - \varepsilon \quad (2)$$

## OWP to PRG cont.

- ▶  $\Pr[D(f(U_n), b(U_n)) = 1] = \delta + \varepsilon$
- ▶  $\Pr[D(f(U_n), \overline{b(U_n)}) = 1] = \delta - \varepsilon$
- ▶ Consider the following algorithm for predicting  $b$ :

### Algorithm 13 (P)

Input:  $y \in \{0, 1\}^n$

1. Flip a random coin  $c \leftarrow \{0, 1\}$ .
2. If  $D(y, c) = 1$  output  $c$ , otherwise, output  $\bar{c}$ .

- ▶ It follows that

$$\begin{aligned} & \Pr[P(f(U_n)) = b(U_n)] \\ &= \Pr[c = b(U_n)] \cdot \Pr[D(f(U_n), c) = 1 \mid c = b(U_n)] \\ & \quad + \Pr[c = \overline{b(U_n)}] \cdot \Pr[D(f(U_n), c) = 0 \mid c = \overline{b(U_n)}] \\ &= \frac{1}{2} \cdot (\delta + \varepsilon) + \frac{1}{2}(1 - \delta + \varepsilon) = \frac{1}{2} + \varepsilon. \end{aligned}$$

## OWP to PRG cont.

### Remark 14

- ▶ Prediction to distinguishing (homework)
- ▶ PRG from any OWF: (1) Regular OWFs, first use pairwise hashing to convert into “almost” permutation. (2) Any OWF, harder

# PRG Length Extension

## Construction 15 (iterated function)

Given  $g: \{0, 1\}^n \mapsto \{0, 1\}^{n+1}$  and  $i \in \mathbb{N}$ , define  $g^i: \{0, 1\}^n \mapsto \{0, 1\}^{n+i}$  as

$$g^i(x) = g(x)_1, g^{i-1}(g(x)_{2,\dots,n+1}),$$

where  $g^0(x) = x$ .

## Claim 16

Let  $g: \{0, 1\}^n \mapsto \{0, 1\}^{n+1}$  be a PRG, then  $g^{t(n)}: \{0, 1\}^n \mapsto \{0, 1\}^{n+t(n)}$  is a PRG, for any  $t \in \text{poly}$ .

Proof: Assume  $\exists$  a PPT  $D$ , an infinite set  $\mathcal{I} \subseteq \mathbb{N}$  and  $p \in \text{poly}$  with

$$\left| \Delta_{g^t(U_n), U_{n+t(n)}}^D \right| > \varepsilon(n) = 1/p(n),$$

for any  $n \in \mathcal{I}$ . We use  $D$  for breaking the hardness of  $g$ .



## PRG Length Extension cont.

- ▶ Fix  $n \in \mathbb{N}$ , for  $i \in \{0, \dots, t = t(n)\}$ , let  $H^i = U_{t-i}, g^i(U_n)$  (i.e., the distribution of  $H^i$  is  $(x, g^i(x'))_{x \leftarrow \{0,1\}^{t-i}, x' \leftarrow \{0,1\}^n}$ )
- ▶ Note that  $H^0 \equiv U_{n+t}$  and  $H^t \equiv g^t(U_n)$ .

### Algorithm 17 ( $D'$ )

Input:  $1^n$  and  $y \in \{0, 1\}^{n+1}$

1. Sample  $i \leftarrow [t]$
2. Return  $D(1^n, U_{t-i}, y_1, g^{i-1}(y_2, \dots, y_{n+1}))$ .

### Claim 18

It holds that  $\left| \Delta_{g(U_n), U_{n+1}}^{D'} \right| > \varepsilon(n)/t(n)$

Proof: ...