# Application of Information Theory, Lecture 9 Parallel Repetition of Interactive Arguments

Iftach Haitner

Tel Aviv University.

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## Part I

# **Interactive Proofs and Arguments**

#### $\mathcal{NP}$ as a Non-interactive Proofs

#### **Definition 1** ( $\mathcal{NP}$ )

 $\mathcal{L} \in \mathcal{NP}$  iff  $\exists$  and poly-time algorithm  $\lor$  such that:

- ▶  $\forall x \in \mathcal{L}$  there exists  $w \in \{0, 1\}^*$  s.t. V(x, w) = 1
- ▶ V(x, w) = 0 for every  $x \notin \mathcal{L}$  and  $w \in \{0, 1\}^*$

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- Efficient verifier, efficient prover (given the witness)
- Soundness holds unconditionally

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**Completeness**  $\forall x \in \mathcal{L}$ :  $Pr[(P, V)(x) = 1] \ge 2/3$ .

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▶ IP = PSPACE!

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- Games no-input protocols.

#### Section 1

## **Interactive Proof for Graph Non-Isomorphism**

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- We will show a simple interactive proof for GNT Idea: Beer tasting...

#### Interactive proof for $\mathcal{GNI}$

#### **Protocol 4 ((P, V)(G**<sub>0</sub> = ([m], E<sub>0</sub>), G<sub>1</sub> = ([m], E<sub>1</sub>)))

- 1. V chooses  $b \leftarrow \{0,1\}$  and  $\pi \leftarrow \Pi_m$ , and sends  $\pi(E_b)$  to P.<sup>a</sup>
- **2.** P send b' to V (tries to set b' = b).
- **3.** V accepts iff b' = b.
  - ${}^{a}\pi(E) = \{(\pi(u), \pi(v) \colon (u, v) \in E\}.$

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#### Claim 5

The above protocol is IP for  $\mathcal{GNI}$ , with perfect completeness and soundness error  $\frac{1}{2}$ .

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#### Hence,

$$G_0 \equiv G_1$$
:  $\Pr[b' = b] \le \frac{1}{2}$ .  $G_0 \not\equiv G_1$ :  $\Pr[b' = b] = 1$  (i.e., P can, possibly inefficiently, extracted from  $\pi(E_i)$ )



## Part II

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- ▶ Public-coin interactive proof/argument in each round the verifier flips coins and sends them to the prover. To compute its output, the verifier applies some (fixed) function to the protocol's transcript.

▶ Give a protocol  $\pi = (P, V)$  and  $k \in \mathbb{N}$ , let  $\pi^{(k)} = (P^{(k)}, V^{(k)})$  be the k-fold parallel repetition of  $\pi$ : i.e., k parallel independent copies of  $\pi$ 

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- Why time?
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- ▶ In the following we focus on games (no input protocols)

## Section 2

# Parallel repetition of public-coin interactive argument



#### **Theorem 6**

Let 
$$\pi = (P, V)$$
 be m-round,  $t_{\pi}$ -time, public-coin protocol with  $\Pr\left[(\widetilde{P}, V) = 1\right] \leq \varepsilon$  for any  $t$ -time  $\widetilde{P}$ , then  $\Pr\left[(\widetilde{P^{(k)}}, V^{(k)}) = 1^k\right] \leq \varepsilon^{k/4}$  for any  $t \cdot \frac{\varepsilon^{k/4}}{mk^3t_{\pi}}$ -time  $\widetilde{P^{(k)}}$ .

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Proof plan: Let  $P^{(k)}$  be t'-time algorithm with  $\Pr\left[(P^{(k)}, V^{(k)}) = 1^k\right] = \varepsilon^{(k)}$ , we construct  $t' \cdot \frac{2mk^2t_\pi\log\frac{1}{\varepsilon^{(k)}}}{\varepsilon^{(k)}}$ -time  $\widetilde{P}$  with  $\Pr\left[(\widetilde{P}, V) = 1\right] \geq (\varepsilon^{(k)})^{4/k}$ .

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▶ The k/4 in the exponent can be pushed to be almost k.

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- ▶ We view the coins of  $V^{(k)}$  as a matrix  $R \in \{0, 1\}^{m \times \ell}$ , letting  $R_j$  denote the coins of the j'th round, and  $R_{1,...,j}$  the coins of the first j rounds.

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- ▶ Let  $\mathbb{R} \sim \{0,1\}^{m \times \ell}$

# Algorithm $\widetilde{P}$

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# Algorithm 7 ( $\widetilde{P}$ )

- 1. Let  $i^* \leftarrow [k]$ .
- **2.** Upon getting the j'th message r from V, do:
  - **2.1** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned that  $R_{1,...,j-1} = \widetilde{R}_{1,...,j-1}$  and  $R_{j,i^*} = r$ .
  - **2.2** If  $(\widetilde{P^{(k)}}, V^{(k)}(R)) = 1^k$ :
    - **2.2.1** Set  $\widetilde{R}_j = R_j$
    - **2.2.2** Send  $a_{j,i^*}$  back to V, for  $a_j$  being the j'th message  $\widetilde{P^{(k)}}$  send to  $V^{(k)}$  in  $(\widetilde{P^{(k)}}, V^{(k)}(R))$ .

Else, GOTO Line 2.1

**2.3** Abort if the overall number of sampling exceeds  $\lceil qm/\varepsilon^{(k)} \rceil$ .

# Algorithm P

Let 
$$q = 2k \cdot \log \frac{1}{\varepsilon^{(k)}}$$
.

# Algorithm 7 ( $\widetilde{P}$ )

- 1. Let  $i^* \leftarrow [k]$ .
- **2.** Upon getting the j'th message r from V, do:
  - **2.1** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned that  $R_{1,...,j-1} = \widetilde{R}_{1,...,j-1}$  and  $R_{j,i^*} = r$ .
  - **2.2** If  $(P^{(k)}, V^{(k)}(R)) = 1^k$ :
    - **2.2.1** Set  $\widetilde{R}_j = R_j$
    - **2.2.2** Send  $a_{j,i^*}$  back to V, for  $a_j$  being the j'th message  $\widetilde{P^{(k)}}$  send to  $V^{(k)}$  in  $(\widetilde{P^{(k)}}, V^{(k)}(R))$ .

Else, GOTO Line 2.1

- **2.3** Abort if the overall number of sampling exceeds  $\lceil qm/\varepsilon^{(k)} \rceil$ .
- Let  $\widetilde{P}'$  be the non aborting variant of  $\widetilde{P}'$ , let  $\widetilde{R}$  and  $\widetilde{N}$  be the value of  $\widetilde{R}$  and # of samples done in a random execution of  $(\widetilde{P}', V^{(k)})$ .

# **Algorithm** P

Let 
$$q = 2k \cdot \log \frac{1}{\varepsilon^{(k)}}$$
.

# Algorithm 7 (P)

- 1. Let  $i^* \leftarrow [k]$ .
- **2.** Upon getting the j'th message r from V, do:
  - **2.1** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned that  $R_{1,...,j-1} = \widetilde{R}_{1,...,j-1}$  and  $R_{j,i^*} = r$ .
  - **2.2** If  $(P^{(k)}, V^{(k)}(R)) = 1^k$ :
    - **2.2.1** Set  $\widetilde{R}_j = R_j$
    - **2.2.2** Send  $a_{j,i^*}$  back to V, for  $a_j$  being the j'th message  $\widetilde{P^{(k)}}$  send to  $V^{(k)}$  in  $(\widetilde{P^{(k)}}, V^{(k)}(R))$ .

Else, GOTO Line 2.1

- **2.3** Abort if the overall number of sampling exceeds  $\lceil qm/\varepsilon^{(k)} \rceil$ .
- Let  $\widetilde{P}'$  be the non aborting variant of  $\widetilde{P}'$ , let  $\widetilde{R}$  and  $\widetilde{N}$  be the value of  $\widetilde{R}$  and # of samples done in a random execution of  $(\widetilde{P}', V^{(k)})$ .
- $\qquad \qquad \Pr\left[(\widetilde{\mathsf{P}},\mathsf{V})=1\right] \geq \Pr\left[\text{win}(\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) := (\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\widetilde{\mathbf{R}})) = 1^k \wedge \widetilde{\mathbf{N}} \leq qm/\varepsilon^{(k)}\right].$

# Experiment 8 (P)

- 1. Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned that  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
- **2.** If  $(\widetilde{P^{(k)}}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 1.

# Experiment 8 (P)

- 1. Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned that  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
- **2.** If  $(\widetilde{P^{(k)}}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 1.
- Let  $\hat{\mathbf{R}}$  be the value of  $\hat{\mathbf{R}}$  in the end of a random execution of  $\hat{\mathbf{P}}$ .

# Experiment 8 (P)

- 1. Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned that  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
- **2.** If  $(\widetilde{P^{(k)}}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 1.
- Let  $\hat{\mathbf{R}}$  be the value of  $\hat{\mathbf{R}}$  in the end of a random execution of  $\hat{\mathbf{P}}$ .
- $\blacktriangleright \ \widehat{\boldsymbol{\mathsf{R}}} \sim \boldsymbol{\mathsf{R}}|_{(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\boldsymbol{\mathsf{R}}))=1^k}$

# Experiment 8 ( $\widehat{P}$ )

- 1. Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned that  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
- **2.** If  $(\widetilde{P^{(k)}}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 1.
- Let  $\hat{\mathbf{R}}$  be the value of  $\hat{\mathbf{R}}$  in the end of a random execution of  $\hat{\mathbf{P}}$ .
- $\blacktriangleright \ \widehat{\boldsymbol{R}} \sim \boldsymbol{R}|_{(\widetilde{P^{(k)}},V^{(k)}(\boldsymbol{R}))=1^k}$
- ► In particular,  $\Pr\left[(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\widehat{\mathbf{R}})=1^k\right]=1$

# Experiment 8 (P)

- 1. Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned that  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
- **2.** If  $(\widetilde{P^{(k)}}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 1.
  - Let  $\hat{\mathbf{R}}$  be the value of  $\hat{\mathbf{R}}$  in the end of a random execution of  $\hat{\mathbf{P}}$ .
- $\blacktriangleright \ \widehat{\boldsymbol{\mathsf{R}}} \sim \boldsymbol{\mathsf{R}}|_{(\widetilde{\boldsymbol{\mathsf{P}}^{(k)}},\boldsymbol{\mathsf{V}}^{(k)}(\boldsymbol{\mathsf{R}}))=1^k}$
- ► In particular,  $\Pr\left[(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\widehat{\mathbf{R}})=1^k\right]=1$
- Let  $\hat{N}$  be # of samples done in  $\hat{R}$ .

# Experiment 8 (P)

- 1. Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned that  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
- **2.** If  $(\widetilde{P^{(k)}}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 1.
  - Let  $\hat{\mathbf{R}}$  be the value of  $\hat{\mathbf{R}}$  in the end of a random execution of  $\hat{\mathbf{P}}$ .
- $\blacktriangleright \ \widehat{\boldsymbol{\mathsf{R}}} \sim \boldsymbol{\mathsf{R}}|_{(\widetilde{\boldsymbol{\mathsf{P}}^{(k)}},\boldsymbol{\mathsf{V}}^{(k)}(\boldsymbol{\mathsf{R}}))=1^k}$
- ► In particular,  $\Pr\left[(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\widehat{\mathbf{R}})=1^k\right]=1$
- Let  $\hat{N}$  be # of samples done in  $\hat{R}$ .

# Experiment 8 (P)

For j = 1 to m:

- 1. Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned that  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
- **2.** If  $(\widetilde{P^{(k)}}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 1.
  - Let  $\hat{\mathbf{R}}$  be the value of  $\hat{\mathbf{R}}$  in the end of a random execution of  $\hat{\mathbf{P}}$ .
- $\blacktriangleright \ \widehat{\boldsymbol{\mathsf{R}}} \sim \boldsymbol{\mathsf{R}}|_{(\widetilde{\boldsymbol{\mathsf{P}}^{(k)}},\boldsymbol{\mathsf{V}}^{(k)}(\boldsymbol{\mathsf{R}}))=1^k}$
- ► In particular,  $\Pr\left[(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\widehat{\mathsf{R}})=1^k\right]=1$
- Let  $\hat{N}$  be # of samples done in  $\hat{R}$ .

#### Lemma 9

$$\Pr\left[\widehat{\mathbf{N}} \leq qm/arepsilon^{(k)}
ight] \geq 1 - rac{1}{q}$$

▶ Let  $(X_1, ..., X_m) = \mathbf{R}$  and  $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$ 

- ► Let  $(X_1, ..., X_m) = \mathbf{R}$  and  $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$
- ▶  $v(\mathbf{y} = (y_1, ..., y_j)) := \Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(X^m)) = 1^k \mid X^j = \mathbf{y}\right]$ (letting  $X^j = (X_1, ..., X_j)$ )

- ▶ Let  $(X_1, ..., X_m) = \mathbf{R}$  and  $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$
- ►  $v(\mathbf{y} = (y_1, ..., y_j)) := \Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(X^m)) = 1^k \mid X^j = \mathbf{y}\right]$ (letting  $X^j = (X_1, ..., X_j)$ )
- Conditioned on  $Y^j = \mathbf{y} = (y_1, \dots, y_j)$ , the expected # of samples done in (j+1)'th round of  $\widehat{P}$  is  $\frac{1}{V(\mathbf{y})}$ .

- ► Let  $(X_1, ..., X_m) = \mathbf{R}$  and  $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$
- ▶  $v(\mathbf{y} = (y_1, ..., y_j)) := \Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(X^m)) = 1^k \mid X^j = \mathbf{y}\right]$ (letting  $X^j = (X_1, ..., X_j)$ )
- Conditioned on  $Y^j = \mathbf{y} = (y_1, \dots, y_j)$ , the expected # of samples done in (j+1)'th round of  $\widehat{P}$  is  $\frac{1}{\nu(\mathbf{y})}$ .
- ▶ We prove Lemma 9 showing that  $E\left[\frac{1}{\nu(Y^j)}\right] = \frac{1}{\varepsilon^{(k)}}$  for every  $j \in \{0, \dots, m-1\}$

- ► Let  $(X_1, ..., X_m) = \mathbf{R}$  and  $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$
- ▶  $v(\mathbf{y} = (y_1, ..., y_j)) := \Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(X^m)) = 1^k \mid X^j = \mathbf{y}\right]$ (letting  $X^j = (X_1, ..., X_j)$ )
- Conditioned on  $Y^j = \mathbf{y} = (y_1, \dots, y_j)$ , the expected # of samples done in (j+1)'th round of  $\widehat{P}$  is  $\frac{1}{\nu(\mathbf{y})}$ .
- ▶ We prove Lemma 9 showing that  $E\left[\frac{1}{\nu(Y^j)}\right] = \frac{1}{\varepsilon^{(k)}}$  for every  $j \in \{0, \dots, m-1\}$

- ► Let  $(X_1, ..., X_m) = \mathbf{R}$  and  $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$
- ▶  $v(\mathbf{y} = (y_1, ..., y_j)) := \Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(X^m)) = 1^k \mid X^j = \mathbf{y}\right]$ (letting  $X^j = (X_1, ..., X_j)$ )
- ► Conditioned on  $Y^j = \mathbf{y} = (y_1, \dots, y_j)$ , the expected # of samples done in (j+1)'th round of  $\widehat{P}$  is  $\frac{1}{\nu(\mathbf{y})}$ .
- ▶ We prove Lemma 9 showing that  $\mathsf{E}\left[\frac{1}{\nu(Y^j)}\right] = \frac{1}{\varepsilon^{(k)}}$  for every  $j \in \{0, \dots, m-1\}$

#### Claim 10

For  $j \in \{0, \dots, m-1\}$  and  $\mathbf{y} \in \operatorname{Supp}(Y^j)$  it holds that  $\Pr_{Y^j}[\mathbf{y}] = \Pr_{X^j}[\mathbf{y}] \cdot \frac{\nu(\mathbf{y})}{\varepsilon^{(k)}}$ 

- ► Let  $(X_1, ..., X_m) = \mathbf{R}$  and  $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$
- ►  $v(\mathbf{y} = (y_1, ..., y_j)) := \Pr\left[(\widetilde{\mathbf{P}^{(k)}}, V^{(k)}(X^m)) = 1^k \mid X^j = \mathbf{y}\right]$ (letting  $X^j = (X_1, ..., X_j)$ )
- Conditioned on  $Y^j = \mathbf{y} = (y_1, \dots, y_j)$ , the expected # of samples done in (j+1)'th round of  $\widehat{P}$  is  $\frac{1}{\nu(\mathbf{y})}$ .
- ▶ We prove Lemma 9 showing that  $\mathsf{E}\left[\frac{1}{\nu(\mathsf{Y}^j)}\right] = \frac{1}{\varepsilon^{(k)}}$  for every  $j \in \{0, \dots, m-1\}$

#### Claim 10

For  $j \in \{0, \dots, m-1\}$  and  $\mathbf{y} \in \operatorname{Supp}(Y^j)$  it holds that  $\Pr_{Y^j}[\mathbf{y}] = \Pr_{X^j}[\mathbf{y}] \cdot \frac{\nu(\mathbf{y})}{\varepsilon^{(k)}}$ 

Hence, 
$$\mathsf{E}_{Y^j}\left[\frac{1}{v(Y^j)}\right] = \sum_{\mathbf{y} \in \mathsf{Supp}(Y^j)} \mathsf{Pr}[Y^j = \mathbf{y}] \cdot \frac{1}{v(\mathbf{y})}$$

- ► Let  $(X_1, ..., X_m) = \mathbf{R}$  and  $(Y_1, ..., Y_m) = \widehat{\mathbf{R}}$
- ►  $v(\mathbf{y} = (y_1, ..., y_j)) := \Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(X^m)) = 1^k \mid X^j = \mathbf{y}\right]$ (letting  $X^j = (X_1, ..., X_j)$ )
- Conditioned on  $Y^j = \mathbf{y} = (y_1, \dots, y_j)$ , the expected # of samples done in (j+1)'th round of  $\widehat{P}$  is  $\frac{1}{\nu(\mathbf{y})}$ .
- ▶ We prove Lemma 9 showing that  $\mathsf{E}\left[\frac{1}{\nu(\mathsf{Y}^j)}\right] = \frac{1}{\varepsilon^{(k)}}$  for every  $j \in \{0, \dots, m-1\}$

#### Claim 10

For  $j \in \{0, \dots, m-1\}$  and  $\mathbf{y} \in \operatorname{Supp}(Y^j)$  it holds that  $\Pr_{Y^j}[\mathbf{y}] = \Pr_{X^j}[\mathbf{y}] \cdot \frac{\nu(\mathbf{y})}{\varepsilon^{(k)}}$ 

Hence, 
$$\mathsf{E}_{\mathsf{Y}^j}\left[\frac{1}{\nu(\mathsf{Y}^j)}\right] = \sum_{\mathbf{y} \in \mathsf{Supp}(\mathsf{Y}^j)} \mathsf{Pr}[\mathsf{Y}^j = \mathbf{y}] \cdot \frac{1}{\nu(\mathbf{y})}$$

$$= \sum_{\mathbf{y}} \mathsf{Pr}[\mathsf{X}^j = \mathbf{y}] \cdot \frac{\nu(\mathbf{y})}{\varepsilon^{(k)}} \cdot \frac{1}{\nu(\mathbf{y})} = \frac{1}{\varepsilon^{(k)}} \cdot \sum_{\mathbf{y} \in \mathsf{Supp}(\mathsf{Y}^j)} \mathsf{Pr}[\mathsf{X}^j = \mathbf{y}] = \frac{1}{\varepsilon^{(k)}}. \ \Box$$

#### Note that

$$\Pr_{Y_{j}|Y^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] = \sum_{\ell=1}^{\infty} (1 - v(\mathbf{y}_{1,...,j-1}))^{\ell-1} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] \cdot v(\mathbf{y}) \qquad (1)$$

$$= \frac{1}{v(\mathbf{y}_{1,...,j-1})} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] \cdot v(\mathbf{y})$$

Note that

$$\Pr_{Y_{j}|Y^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] = \sum_{\ell=1}^{\infty} (1 - \nu(\mathbf{y}_{1,...,j-1}))^{\ell-1} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] \cdot \nu(\mathbf{y}) \qquad (1)$$

$$= \frac{1}{\nu(\mathbf{y}_{1,...,j-1})} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] \cdot \nu(\mathbf{y})$$

Note that

$$\Pr_{Y_{j}|Y^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] = \sum_{\ell=1}^{\infty} (1 - v(\mathbf{y}_{1,...,j-1}))^{\ell-1} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] \cdot v(\mathbf{y}) \qquad (1)$$

$$= \frac{1}{v(\mathbf{y}_{1,...,j-1})} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] \cdot v(\mathbf{y})$$

$$\Pr_{\mathbf{y}^{j}}[\mathbf{y}] = \Pr_{\mathbf{y}^{j-1}}[\mathbf{y}_{1\dots,j-1}] \cdot \Pr_{\mathbf{y}_{j}|\mathbf{y}^{j-1}=\mathbf{y}_{1\dots,j-1}}[\mathbf{y}_{j}]$$

Note that

$$\Pr_{Y_{j}|Y^{j-1}=\mathbf{y}_{1...,j-1}}[y_{j}] = \sum_{\ell=1}^{\infty} (1 - \nu(\mathbf{y}_{1...,j-1}))^{\ell-1} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1...,j-1}}[y_{j}] \cdot \nu(\mathbf{y}) \qquad (1)$$

$$= \frac{1}{\nu(\mathbf{y}_{1...,j-1})} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1...,j-1}}[y_{j}] \cdot \nu(\mathbf{y})$$

$$\Pr_{Y_{j}}[\mathbf{y}] = \Pr_{Y_{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \Pr_{Y_{j}|Y^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] 
= \Pr_{X_{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \frac{v(\mathbf{y}_{1...,j-1})}{\varepsilon^{(k)}} \cdot \Pr_{Y_{j}|Y^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}]$$
(i.h.)

Note that

$$\Pr_{Y_{j}|Y^{j-1}=\mathbf{y}_{1...,j-1}}[y_{j}] = \sum_{\ell=1}^{\infty} (1 - \nu(\mathbf{y}_{1...,j-1}))^{\ell-1} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1...,j-1}}[y_{j}] \cdot \nu(\mathbf{y}) \qquad (1)$$

$$= \frac{1}{\nu(\mathbf{y}_{1...,j-1})} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1...,j-1}}[y_{j}] \cdot \nu(\mathbf{y})$$

$$\begin{aligned}
&\Pr_{Y_{j}}[\mathbf{y}] = \Pr_{Y_{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \Pr_{Y_{j}|Y^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \\
&= \Pr_{X^{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \frac{v(\mathbf{y}_{1...,j-1})}{\varepsilon^{(k)}} \cdot \Pr_{Y_{j}|Y^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \\
&= \Pr_{X^{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \frac{v(\mathbf{y}_{1...,j-1})}{\varepsilon^{(k)}} \cdot \frac{v(\mathbf{y})}{v(\mathbf{y}_{1...,j-1})} \cdot \Pr_{X_{j}|X^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \quad \text{(Eq. (1))}
\end{aligned}$$

Note that

$$\Pr_{Y_{j}|Y^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] = \sum_{\ell=1}^{\infty} (1 - v(\mathbf{y}_{1,...,j-1}))^{\ell-1} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] \cdot v(\mathbf{y}) \qquad (1)$$

$$= \frac{1}{v(\mathbf{y}_{1,...,j-1})} \cdot \Pr_{X_{j}|X^{j-1}=\mathbf{y}_{1,...,j-1}}[y_{j}] \cdot v(\mathbf{y})$$

$$\begin{aligned}
&\Pr_{Y_{j}}[\mathbf{y}] = \Pr_{Y_{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \Pr_{Y_{j}|Y^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \\
&= \Pr_{X^{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \frac{v(\mathbf{y}_{1...,j-1})}{\varepsilon^{(k)}} \cdot \Pr_{Y_{j}|Y^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \\
&= \Pr_{X^{j-1}}[\mathbf{y}_{1...,j-1}] \cdot \frac{v(\mathbf{y}_{1...,j-1})}{\varepsilon^{(k)}} \cdot \frac{v(\mathbf{y})}{v(\mathbf{y}_{1...,j-1})} \cdot \Pr_{X_{j}|X^{j-1} = \mathbf{y}_{1...,j-1}}[y_{j}] \quad \text{(Eq. (1))} \\
&= \Pr_{Y_{j}}[\mathbf{y}] \cdot \frac{v(\mathbf{y})}{\varepsilon^{(k)}}.
\end{aligned}$$

- 1. Let  $i^* \leftarrow [k]$ .
- **2.** For for j = 1 to m:
  - **2.1** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
  - **2.2** If  $(P^{(k)}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_{j,j^*} = R_{j,j^*}$ . Else, GOTO Line 2.1.
  - **2.3** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,...,j-1} = \widehat{R}_{1,...,j-1}$  and  $R_{j,i^*} = \widehat{R}_{j,i^*}$ .
  - **2.4** If  $(P^{(k)}, V^{(k)}(R)) = 1^k$ , set  $\hat{R}_i = R_i$ . Else, GOTO Line 2.3.

- **1.** Let  $i^* \leftarrow [k]$ .
- **2.** For for j = 1 to m:
  - **2.1** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
  - **2.2** If  $(P^{(k)}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_{j,j^*} = R_{j,j^*}$ . Else, GOTO Line 2.1.
  - **2.3** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,...,j-1} = \widehat{R}_{1,...,j-1}$  and  $R_{j,i^*} = \widehat{R}_{j,i^*}$ .
  - **2.4** If  $(P^{(k)}, V^{(k)}(R)) = 1^k$ , set  $\hat{R}_i = R_i$ . Else, GOTO Line 2.3.
- Let  $\widehat{\mathbf{R}}$  be the final value of  $\widehat{\mathbf{R}}$  in  $\widehat{\mathbf{P}}$ .

- **1.** Let  $i^* \leftarrow [k]$ .
- **2.** For for j = 1 to m:
  - **2.1** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
  - **2.2** If  $(\widehat{P^{(k)}}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_{j,j^*} = R_{j,j^*}$ . Else, GOTO Line 2.1.
  - **2.3** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,...,j-1} = \widehat{R}_{1,...,j-1}$  and  $R_{j,j^*} = \widehat{R}_{j,j^*}$ .
  - **2.4** If  $(\widehat{P^{(k)}}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 2.3.
- Let  $\widehat{\mathbf{R}}$  be the final value of  $\widehat{\mathbf{R}}$  in  $\widehat{\mathbf{P}}$ .
- $\blacktriangleright \ \widehat{\mathbf{R}} \sim (\mathbf{R}|(\widehat{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\mathbf{R})) = 1^k)$

- **1.** Let  $i^* \leftarrow [k]$ .
- **2.** For for j = 1 to m:
  - **2.1** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
  - **2.2** If  $(P^{(k)}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_{j,j^*} = R_{j,j^*}$ . Else, GOTO Line 2.1.
  - **2.3** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,...,j-1} = \widehat{R}_{1,...,j-1}$  and  $R_{j,j,*} = \widehat{R}_{j,j,*}$ .
  - **2.4** If  $(P^{(\overline{k})}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 2.3.
- Let  $\hat{\mathbf{R}}$  be the final value of  $\hat{\mathbf{R}}$  in  $\hat{\mathbf{P}}$ .
- $ightharpoonup \widehat{\mathbf{R}} \sim (\mathbf{R}|(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\mathbf{R})) = 1^k)$
- Let  $\hat{N}$  be the # of Step-2.3-samples done in  $\hat{P}$ .

- **1.** Let  $i^* \leftarrow [k]$ .
- **2.** For for j = 1 to m:
  - **2.1** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
  - **2.2** If  $(P^{(k)}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_{j,j^*} = R_{j,j^*}$ . Else, GOTO Line 2.1.
  - **2.3** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,...,j-1} = \widehat{R}_{1,...,j-1}$  and  $R_{j,j,*} = \widehat{R}_{j,j,*}$ .
  - **2.4** If  $(P^{(\overline{k})}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 2.3.
- Let  $\hat{\mathbf{R}}$  be the final value of  $\hat{\mathbf{R}}$  in  $\hat{\mathbf{P}}$ .
- $ightharpoonup \widehat{\mathbf{R}} \sim (\mathbf{R}|(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\mathbf{R})) = 1^k)$
- Let  $\hat{N}$  be the # of Step-2.3-samples done in  $\hat{P}$ .

# Experiment 11 (P)

- 1. Let  $i^* \leftarrow [k]$ .
- **2.** For for j = 1 to m:
  - **2.1** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$ .
  - **2.2** If  $(\widehat{P^{(k)}}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_{j,j^*} = R_{j,j^*}$ . Else, GOTO Line 2.1.
  - **2.3** Let  $R \leftarrow \{0,1\}^{m \times \ell}$ , conditioned on  $R_{1,\dots,j-1} = \widehat{R}_{1,\dots,j-1}$  and  $R_{j,j^*} = \widehat{R}_{j,j^*}$ .
  - **2.4** If  $(P^{(k)}, V^{(k)}(R)) = 1^k$ , set  $\widehat{R}_j = R_j$ . Else, GOTO Line 2.3.
  - Let  $\hat{\mathbf{R}}$  be the final value of  $\hat{\mathbf{R}}$  in  $\hat{\mathbf{P}}$ .
  - $\blacktriangleright \ \widehat{\mathsf{R}} \sim (\mathsf{R}|(\mathsf{P}^{(k)},\mathsf{V}^{(k)}(\mathsf{R})) = \mathsf{1}^k)$
- Let  $\hat{N}$  be the # of Step-2.3-samples done in  $\hat{P}$ .

#### Lemma 12

$$\Pr\left[\operatorname{win}(\widehat{\boldsymbol{R}},\widehat{\boldsymbol{N}})\right] \geq 1 - \frac{1}{q}$$

Let 
$$\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$$
 (=  $\widehat{\mathbf{R}}$ ) and  $\widetilde{\mathbf{R}}_i|_{i^*=i}$ .

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### **Proposition 13**

$$D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i||\widetilde{\mathbf{R}}_i).$$

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$$D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i||\widetilde{\mathbf{R}}_i).$$

Proof:

Let 
$$\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$$
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## **Proposition 13**

$$D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i||\widetilde{\mathbf{R}}_i).$$

Proof: HW.

Let 
$$\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$$
 (=  $\widehat{\mathbf{R}}$ ) and  $\widetilde{\mathbf{R}}_i|_{i^*=i}$ .

## **Proposition 13**

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

Proof: HW.

$$\sum_{i\in[k]}D(\widehat{\boldsymbol{R}}||\widetilde{\boldsymbol{R}}_i)\leq D(\widehat{\boldsymbol{R}}||\boldsymbol{R})$$

Let 
$$\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$$
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## **Proposition 13**

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

Proof: HW.

$$\sum_{i \in [k]} D(\widehat{\textbf{\textit{R}}}||\widetilde{\textbf{\textit{R}}}_i) \leq D(\widehat{\textbf{\textit{R}}}||\textbf{\textit{R}})$$

$$\blacktriangleright \ \, \text{By Lecture 7 (Thm. 7), } \, D(\widehat{\mathbf{R}}||\mathbf{R}) \leq \log \frac{1}{\Pr\left[(\widehat{P^{(k)}}, V^{(k)}(\mathbf{R})) = 1^k\right]} = \log \frac{1}{\varepsilon^{(k)}}$$

Let 
$$\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$$
 (=  $\widehat{\mathbf{R}}$ ) and  $\widetilde{\mathbf{R}}_i|_{i^*=i}$ .

## **Proposition 13**

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

Proof: HW.

$$\sum_{i \in [k]} D(\widehat{\textbf{\textit{R}}}||\widetilde{\textbf{\textit{R}}}_i) \leq D(\widehat{\textbf{\textit{R}}}||\textbf{\textit{R}})$$

- ▶ By Lecture 7 (Thm. 7),  $D(\widehat{\mathbf{R}}||\mathbf{R}) \le \log \frac{1}{\Pr[\widehat{P^{(k)}}, V^{(k)}(\mathbf{R})) = 1^k]} = \log \frac{1}{\varepsilon^{(k)}}$
- ▶ Let  $\alpha := \Pr[\min(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})]$  and  $\beta := \Pr[\min(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})]$ .

Let 
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## **Proposition 13**

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

Proof: HW.

$$\sum_{i \in [k]} D(\widehat{\textbf{\textit{R}}}||\widetilde{\textbf{\textit{R}}}_i) \leq D(\widehat{\textbf{\textit{R}}}||\textbf{\textit{R}})$$

- ▶ By Lecture 7 (Thm. 7),  $D(\widehat{\mathbf{R}}||\mathbf{R}) \leq \log \frac{1}{\Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(\mathbf{R})) = 1^k\right]} = \log \frac{1}{\varepsilon^{(k)}}$
- ▶ Let  $\alpha := \Pr[\min(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})]$  and  $\beta := \Pr[\min(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})]$ .
- ▶ Lemma 14  $\implies \alpha \cdot \log \frac{\alpha}{\beta} \le -\frac{1}{k} \cdot \log \varepsilon^{(k)} \implies \beta \ge 2^{\log \alpha + \frac{1}{\alpha k} \log \varepsilon^{(k)}}$

Let 
$$\widehat{\mathbf{R}}_i := \widehat{\mathbf{R}}|_{i^*=i}$$
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## **Proposition 13**

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

Proof: HW.

$$\sum_{i \in [k]} D(\widehat{\boldsymbol{R}}||\widetilde{\boldsymbol{R}}_i) \leq D(\widehat{\boldsymbol{R}}||\boldsymbol{R})$$

- ▶ By Lecture 7 (Thm. 7),  $D(\widehat{\mathbf{R}}||\mathbf{R}) \leq \log \frac{1}{\Pr[\widehat{P^{(k)}}, V^{(k)}(\mathbf{R})) = 1^k]} = \log \frac{1}{\varepsilon^{(k)}}$
- ▶ Let  $\alpha := \Pr[\min(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})]$  and  $\beta := \Pr[\min(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})]$ .
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- ▶ Lemma 12  $\implies \alpha \ge 1 \frac{1}{q} \ge 2^{-\frac{2}{q}} \ge 2^{\frac{\log \varepsilon^{(k)}}{k}}$

Let 
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## **Proposition 13**

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_i || \widetilde{\mathbf{R}}_i).$$

Proof: HW.

$$\sum_{i\in[k]}D(\widehat{\boldsymbol{R}}||\widetilde{\boldsymbol{R}}_i)\leq D(\widehat{\boldsymbol{R}}||\boldsymbol{R})$$

- ▶ By Lecture 7 (Thm. 7),  $D(\widehat{\mathbf{R}}||\mathbf{R}) \le \log \frac{1}{\Pr[\widehat{P^{(k)}}, V^{(k)}(\mathbf{R})) = 1^k]} = \log \frac{1}{\varepsilon^{(k)}}$
- ▶ Let  $\alpha := \Pr[\min(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})]$  and  $\beta := \Pr[\min(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})]$ .
- ▶ Lemma 14  $\implies \alpha \cdot \log \frac{\alpha}{\beta} \le -\frac{1}{k} \cdot \log \varepsilon^{(k)} \implies \beta \ge 2^{\log \alpha + \frac{1}{\alpha k} \log \varepsilon^{(k)}}$
- ▶ Lemma 12  $\implies \alpha \ge 1 \frac{1}{q} \ge 2^{-\frac{2}{q}} \ge 2^{\frac{\log e^{(k)}}{k}}$
- ▶ We conclude that  $\beta \ge 2^{\frac{4}{k}\log \varepsilon^{(k)}} = \sqrt[k]{\varepsilon^{(k)}}$ .

#### Lemma 15

Let 
$$Z = \{Z_{ij}\}_{(i,j) \in [k] \times [m]}$$
 be iids, let  $W$  be an event, and let  $D_i(z) = \prod_{j=1}^m \Pr\left[Z_{j,i} = z_{i,j}|Z_{1,...,j-1} = z_{1,...,j-1}\right] \cdot \Pr\left[Z_{j,-i} = z_{i,j-1}|Z_{1,...,j-1} = z_{1,...,j-1} \wedge Z_{j,i} = z_{i,j} \wedge W\right]$   
Then  $\sum_{i=1}^k D(Z_W||D_i) \leq D(Z_W||Z)$ .

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Then  $\sum_{i=1}^k D(Z_W || D_i) \leq D(Z_W || Z)$ .

Letting 
$$Z = \mathbf{R}$$
 and  $W$  be the event  $(\widehat{P^{(k)}}, V^{(k)}(\mathbf{R})) = 1^k$ , Lemma 15 yields that  $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widetilde{\mathbf{R}}_i) \le D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$ .  $\square$ 

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Then  $\sum_{i=1}^k D(Z_W || D_i) \leq D(Z_W || Z)$ .

Letting 
$$Z = \mathbf{R}$$
 and  $W$  be the event  $(\mathsf{P}^{(k)},\mathsf{V}^{(k)}(\mathbf{R})) = 1^k$ , Lemma 15 yields that  $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widetilde{\mathbf{R}}_i) \leq D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$ .  $\square$ 

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Letting 
$$Z = \mathbf{R}$$
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Proof: (of Lemma 15) We prove for m = k = 2.

▶ Let  $X = Z_1$  and  $Y = Z_2$ 

#### Lemma 15

Let 
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 be iids, let  $W$  be an event, and let  $D_i(z) = \prod_{j=1}^m \Pr\left[Z_{j,i} = z_{i,j}|Z_{1,...,j-1} = z_{1,...,j-1}\right] \cdot \Pr\left[Z_{j,-i} = z_{i,j-1}|Z_{1,...,j-1} = z_{1,...,j-1} \wedge Z_{j,i} = z_{i,j} \wedge W\right]$   
Then  $\sum_{i=1}^k D(Z_W||D_i) \leq D(Z_W||Z)$ .

Letting 
$$Z = \mathbf{R}$$
 and  $W$  be the event  $(\mathsf{P}^{(k)},\mathsf{V}^{(k)}(\mathbf{R})) = 1^k$ , Lemma 15 yields that  $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widetilde{\mathbf{R}}_i) \leq D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$ .  $\square$ 

- ▶ Let  $X = Z_1$  and  $Y = Z_2$
- $U(x_1, x_2, y_1, y_2) := \Pr_{(X,Y)} [(x_1, x_2, y_1, y_2)]$

#### Lemma 15

Let 
$$Z = \{Z_{ij}\}_{(i,j) \in [k] \times [m]}$$
 be iids, let  $W$  be an event, and let  $D_i(z) = \prod_{j=1}^m \Pr\left[Z_{j,i} = z_{i,j}|Z_{1,...,j-1} = z_{1,...,j-1}\right] \cdot \Pr\left[Z_{j,-i} = z_{i,j-1}|Z_{1,...,j-1} = z_{1,...,j-1} \wedge Z_{j,i} = z_{i,j} \wedge W\right]$   
Then  $\sum_{i=1}^k D(Z_W||D_i) \leq D(Z_W||Z)$ .

Letting 
$$Z = \mathbf{R}$$
 and  $W$  be the event  $(\mathsf{P}^{(k)},\mathsf{V}^{(k)}(\mathbf{R})) = 1^k$ , Lemma 15 yields that  $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widetilde{\mathbf{R}}_i) \le D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$ .  $\square$ 

- ▶ Let  $X = Z_1$  and  $Y = Z_2$
- $U(x_1, x_2, y_1, y_2) := \Pr_{(X,Y)} [(x_1, x_2, y_1, y_2)]$
- $ightharpoonup C(x_1, x_2, y_1, y_1) := (X|_W)(x_1, x_2, y_1, y_1)$

#### Lemma 15

Let 
$$Z = \{Z_{ij}\}_{(i,j) \in [k] \times [m]}$$
 be iids, let  $W$  be an event, and let  $D_i(z) = \prod_{j=1}^m \Pr\left[Z_{j,i} = z_{i,j}|Z_{1,...,j-1} = z_{1,...,j-1}\right] \cdot \Pr\left[Z_{j,-i} = z_{i,j-1}|Z_{1,...,j-1} = z_{1,...,j-1} \wedge Z_{j,i} = z_{i,j} \wedge W\right]$   
Then  $\sum_{i=1}^k D(Z_W||D_i) \leq D(Z_W||Z)$ .

Letting 
$$Z = \mathbf{R}$$
 and  $W$  be the event  $(\mathsf{P}^{(k)},\mathsf{V}^{(k)}(\mathbf{R})) = 1^k$ , Lemma 15 yields that  $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widetilde{\mathbf{R}}_i) \leq D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$ .  $\square$ 

- ▶ Let  $X = Z_1$  and  $Y = Z_2$
- $U(x_1, x_2, y_1, y_2) := \Pr_{(X,Y)} [(x_1, x_2, y_1, y_2)]$
- $C(x_1, x_2, y_1, y_1) := (X|_W)(x_1, x_2, y_1, y_1)$
- $Pr[X_1, x_2, y_1, y_1) := Pr[X_1 = x_1 | W] \cdot Pr[X_2 = x_2 | W] \cdot Pr[Y_1 = y_1 | W, X = (x_1, x_2)] \cdot Pr[Y_2 = y_2 | W, X = (x_1, x_2)]$

#### Lemma 15

Let 
$$Z = \{Z_{ij}\}_{(i,j) \in [k] \times [m]}$$
 be iids, let  $W$  be an event, and let  $D_i(z) = \prod_{j=1}^m \Pr\left[Z_{j,i} = z_{i,j}|Z_{1,...,j-1} = z_{1,...,j-1}\right] \cdot \Pr\left[Z_{j,-i} = z_{i,j-1}|Z_{1,...,j-1} = z_{1,...,j-1} \wedge Z_{j,i} = z_{i,j} \wedge W\right]$   
Then  $\sum_{i=1}^k D(Z_W||D_i) \leq D(Z_W||Z)$ .

Letting 
$$Z = \mathbf{R}$$
 and  $W$  be the event  $(\mathbf{P}^{(\overline{k})}, \mathbf{V}^{(k)}(\mathbf{R})) = 1^k$ , Lemma 15 yields that  $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}_i) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widetilde{\mathbf{R}}_i) \le D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$ .  $\square$ 

- ▶ Let  $X = Z_1$  and  $Y = Z_2$
- $ightharpoonup C(x_1, x_2, y_1, y_1) := (X|_W)(x_1, x_2, y_1, y_1)$
- $Pr[X_1, x_2, y_1, y_1) := Pr[X_1 = x_1 | W] \cdot Pr[X_2 = x_2 | W] \cdot Pr[Y_1 = y_1 | W, X = (x_1, x_2)] \cdot Pr[Y_2 = y_2 | W, X = (x_1, x_2)]$
- ► We write  $\frac{C(x_1, x_2, y_1, y_1)}{U(x_1, x_2, y_1, y_1)} = \frac{\Pr[X_1 = x_1 | W] \cdot \Pr[Y_1 = y_1 | W, X = (x_1, x_2)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \cdot \frac{\Pr[X_2 = x_2 | W] \cdot \Pr[Y_2 = y_2 | W, X = (x_1, x_2)]}{\Pr[X_2 = x_2] \cdot \Pr[Y_2 = y_2]} \cdot \frac{C(x_1, x_2, y_1, y_1)}{Q(x_1, x_2, y_1, y_1)}$

$$\begin{split} D(C||U) &= \underset{(x_1, x_2, y_1, y_2) \leftarrow C}{\mathsf{E}} \left[ \log \frac{\mathsf{Pr}\left[X_1 = x_1 | W\right] \cdot \mathsf{Pr}\left[Y_1 = y_1 | W, X = (x_1, x_2)\right]}{\mathsf{Pr}\left[X_1 = x_1\right] \cdot \mathsf{Pr}\left[Y_1 = y_1\right]} \right] \\ &+ \underset{(x_1, x_2, y_1, y_2) \leftarrow C}{\mathsf{E}} \left[ \log \frac{\mathsf{Pr}\left[X_2 = x_2 | W\right] \cdot \mathsf{Pr}\left[Y_2 = y_2 | W, X = (x_1, x_2)\right]}{\mathsf{Pr}\left[X_2 = x_2\right] \cdot \mathsf{Pr}\left[Y_2 = y_2\right]} \right] \\ &+ \underset{(x_1, x_2, y_1, y_2) \leftarrow C}{\mathsf{E}} \left[ \log \frac{C(x_1, x_2, y_1, y_2)}{Q(x_1, x_2, y_1, y_2)} \right]. \end{split}$$

$$\begin{split} D(C||U) &= \mathop{\mathbb{E}}_{(x_1,x_2,y_1,y_2) \leftarrow C} \left[ \log \frac{\Pr[X_1 = x_1|W] \cdot \Pr[Y_1 = y_1|W,X = (x_1,x_2)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \right] \\ &+ \mathop{\mathbb{E}}_{(x_1,x_2,y_1,y_2) \leftarrow C} \left[ \log \frac{\Pr[X_2 = x_2|W] \cdot \Pr[Y_2 = y_2|W,X = (x_1,x_2)]}{\Pr[X_2 = x_2] \cdot \Pr[Y_2 = y_2]} \right] \\ &+ \mathop{\mathbb{E}}_{(x_1,x_2,y_1,y_2) \leftarrow C} \left[ \log \frac{C(x_1,x_2,y_1,y_2)}{Q(x_1,x_2,y_1,y_2)} \right]. \end{split}$$

#### It follows that

$$\begin{split} D(C||U) &= D(X_1|_W, X_2|_{W,X_1}, Y_1|_{W,X}, Y_2|_{W,X,Y_1}||X_1, X_2|_{W,X_1}, Y_1, Y_2|_{W,X,Y_1}) \\ &+ D(X_2|_W, X_1|_{W,X_2}, Y_2|_{W,X}, Y_1|_{W,X,Y_2}||X_2, X_1|_{W,X_2}, Y_2, Y_1|_{W,X,Y_2}) \\ &+ D(C||Q), \end{split}$$

$$\begin{split} D(C||U) &= \mathop{\mathsf{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[ \log \frac{\Pr[X_1 = x_1 | W] \cdot \Pr[Y_1 = y_1 | W, X = (x_1, x_2)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \right] \\ &+ \mathop{\mathsf{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[ \log \frac{\Pr[X_2 = x_2 | W] \cdot \Pr[Y_2 = y_2 | W, X = (x_1, x_2)]}{\Pr[X_2 = x_2] \cdot \Pr[Y_2 = y_2]} \right] \\ &+ \mathop{\mathsf{E}}_{(x_1, x_2, y_1, y_2) \leftarrow C} \left[ \log \frac{C(x_1, x_2, y_1, y_2)}{Q(x_1, x_2, y_1, y_2)} \right]. \end{split}$$

It follows that

$$D(C||U) = D(X_1|_W, X_2|_{W,X_1}, Y_1|_{W,X}, Y_2|_{W,X,Y_1}||X_1, X_2|_{W,X_1}, Y_1, Y_2|_{W,X,Y_1})$$

$$+ D(X_2|_W, X_1|_{W,X_2}, Y_2|_{W,X}, Y_1|_{W,X,Y_2}||X_2, X_1|_{W,X_2}, Y_2, Y_1|_{W,X,Y_2})$$

$$+ D(C||Q),$$

and the proof follows since  $D(\cdot||\cdot) \geq 0$ .  $\square$ 

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- Why fails us to extend this approach for non-public-coin interactive arguments?

# Section 3

# Parallel amplification for any interactive argument



# Parallel amplification theorem for any protocol

Can we amplify the security of any interactive argument "in parallel"?

# Parallel amplification theorem for any protocol

- Can we amplify the security of any interactive argument "in parallel"?
- Yes we can!