Application of Information Theory, Lecture 2 Joint & Conditional Entropy, Mutual Information

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Nov 4, 2014

Part I

Joint and Conditional Entropy

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$$= \frac{1}{2} \cdot 1 + \frac{1}{2} \cdot 2 = 1\frac{1}{2}$$

Joint entropy, cont.

▶ The joint entropy of $(X_1, ..., X_n) \sim p$, is

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$$= -\sum_{Z = p_{Y|X}(Y|X)} \log Z$$

Conditional entropy, cont.

Example

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- ► H(Y|X) = H(X, Y) H(X) is as an alternative definition for H(Y|X).

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Claim 1

For rvs X, Y, it holds that H(X, Y) = H(X) + H(Y|X).

X			
	<i>P</i> _{1,1}		$P_{1,n}$
	::	:	:
	$P_{n,1}$		$P_{n,n}$

Let
$$q_i = \sum_{j=1}^{n} p_{i,j}$$

$$H(P_{1,1}, \dots, P_{n,n})$$

$$= H(q_1, \dots, q_n) + \sum_{i=1}^{n} q_i H(\frac{P_{i,1}}{q_i}, \dots, \frac{P_{i,n}}{q_i})$$

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$H(Y|X) \leq H(Y)$

Jensen inequality: for any concave function f, values t_1, \ldots, t_k and $\lambda_1, \ldots, \lambda_k \in [0, 1]$ with $\sum_i \lambda_i = 1$, it holds that $\sum_i \lambda_i f(t_i) \leq f(\sum_i \lambda_i t_i)$.

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$$\implies$$
 $H(Y|X) = H(Y)$

Other inequalities

► $H(X), H(Y) \le H(X, Y) \le H(X) + H(Y)$.

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$$H(X|Y,Z) = \mathop{\mathbb{E}}_{Z,Y} H(X \mid Y,Z)$$
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Proof:

$$H(X|Y,Z) = \underset{Z,Y}{\mathbb{E}} H(X \mid Y,Z)$$

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Claim 2

For rvs X_1, \ldots, X_k , it holds that

$$H(X_1,\ldots,X_k) = H(X_i) + H(X_2|X_1) + \ldots + H(X_k|X_1,\ldots,X_{k-1}).$$

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▶ (from last class) Let $X_1, ..., X_n$ be Boolean iid with $X_i \sim (\frac{1}{3}, \frac{2}{3})$. Compute $H(X_1, ..., X_n)$

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 - Via mapping?

Let X_1, \ldots, X_n be Boolean iids with $X_i \sim (p, 1-p)$ and let $X = X_1, \ldots, X_n$. Let f be such that $\Pr[f(X) = z] = \Pr[f(X) = z']$, for every $k \in \mathbb{N}$ and $z, z' \in \{0, 1\}^k$. Let K = |f(X)|. Prove that $E K \le n \cdot h(p)$.

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Interpretation

Let $X_1, ..., X_n$ be Boolean iids with $X_i \sim (p, 1-p)$ and let $X = X_1, ..., X_n$. Let f be such that $\Pr[f(X) = z] = \Pr[f(X) = z']$, for every $k \in \mathbb{N}$ and $z, z' \in \{0, 1\}^k$. Let K = |f(X)|. Prove that $E K < n \cdot h(p)$.

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$$\geq H(f(X), K)$$

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- Interpretation
- Positive results

Applications cont.

▶ How many comparisons it takes to sort *n* elements?

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Let A be a sorter for n elements algorithm making t comparisons.
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- Let X be a uniform random permutation of [n] and let Y_1, \ldots, Y_t be the answers A gets when sorting X.
- \blacktriangleright X is determined by Y_1, \ldots, Y_t .

- How many comparisons it takes to sort n elements?
 Let A be a sorter for n elements algorithm making t comparisons.
 What can we say about t?
- Let X be a uniform random permutation of [n] and let Y_1, \ldots, Y_t be the answers A gets when sorting X.
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 $\implies t > \log n! = \Theta(n \log n)$

Let $p=(p_1,\ldots,p_n)$ and $q=(q_1,\ldots,q_n)$ be two distributions, and for $\lambda\in[0,1]$ consider the distribution $\tau_\lambda=\lambda p+(1-\lambda)q$. (i.e., $\tau_\lambda=(\lambda p_1+(1-\lambda)q_1,\ldots,\lambda p_n+(1-\lambda)q_n)$.

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We are now certain that we drew the graph of the (two-dimensional) entropy function right...

Part II

Mutual Information

$$I(X; Y) := H(Y) - H(Y|X)$$

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► I(X; Y) — the "information" that X gives on Y

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- I(X; Y|Z) = I(Y; X|Z) (since I(X'; Y') = I(Y'; X'))

X	0	1
0	$\frac{1}{4}$	$\frac{1}{4}$
1	1 2	0



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0	1 4	$\frac{1}{4}$
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$$= H(Y) - H(Y|X)$$

$$= h(\frac{1}{4}) - \frac{1}{2}h(\frac{1}{2})$$

Claim 4 (Chain rule for mutual information)

For rvs $X_1, ..., X_k, Y$, it holds that $I(X_1, ..., X_k | Y) = I(X|Y) + I(X_2; Y|X_1) + ... + I(X_k; Y|X_1, ..., X_{k-1})$.

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Proof: ?

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Proof: ? HW

Let X_1, \ldots, X_n be iid with $X_i \sim (p, 1 - p)$, under the condition that $\bigoplus_i x_i = 0$. Compute $I(X_1, \ldots, X_{n-1}; X_n)$.

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Compute $I(X_1, ..., X_{n-1}; X_n)$.

By chain rule

$$I(X_1,...,X_{n-1};X_n) = H(X_1;X_n) + H(X_2;X_n|X_1) + ... + H(X_{n-1};X_n|X_1,...,X_{n-2})$$

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► Let *T* and *F* be the top and front side, respectively, of a 6-sided fair dice. Compute *I*(*T*; *F*).

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$$I(T;F) = H(T) - H(T|F)$$

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$$I(T; F) = H(T) - H(T|F)$$

= log 6 - log 4
= log 3 - 1.

Part III

Data processing

Definition 5 (Markov Chain)

Rvs $(X, Y, Z) \sim p$ form a Markov chain, denoted $X \to Y \to Z$, if $p(x, y, z) = p_X(x) \cdot p_{Y|X}(y|x) \cdot p_{Z|Y}(z|y)$, for all x, y, z.

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Example: random walk on graph.

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If $X \to Y \to Z$, then $I(X; Y) \ge I(X; Z)$.

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 - I(X; Z|Y) = H(Z|Y) H(Z|Y, X) $= \underset{Y}{\mathsf{E}} H(p_{Z|Y=y}) \underset{Y,X}{\mathsf{E}} H(p_{Z|Y=y,X=x})$

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- I(X; Z|Y) = 0

 - I(X; Z|Y) = H(Z|Y) H(Z|Y, X) $= \mathop{\mathsf{E}}_{Y} H(p_{Z|Y=y}) \mathop{\mathsf{E}}_{Y, X} H(p_{Z|Y=y, X=x})$ $= \mathop{\mathsf{E}}_{Y} H(p_{Z|Y=y}) \mathop{\mathsf{E}}_{Y} H(p_{Z|Y=y}) = 0.$

Definition 5 (Markov Chain)

Rvs $(X, Y, Z) \sim p$ form a Markov chain, denoted $X \to Y \to Z$, if $p(x, y, z) = p_X(x) \cdot p_{Y|X}(y|x) \cdot p_{Z|Y}(z|y)$, for all x, y, z.

Example: random walk on graph.

Claim 6

If
$$X \to Y \to Z$$
, then $I(X; Y) \ge I(X; Z)$.

- ▶ By Chain rule, I(X; Y, Z) = I(X; Z) + I(X; Y|Z) = I(X; Y) + I(X; Z|Y).
- I(X;Z|Y) = 0

$$I(X; Z|Y) = H(Z|Y) - H(Z|Y, X)$$

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$$= \mathop{\mathbb{E}}_{Y} H(p_{Z|Y=y}) - \mathop{\mathbb{E}}_{Y} H(p_{Z|Y=y}) = 0.$$

▶ Since $I(X; Y|Z) \ge 0$, we conclude $I(X; Y) \ge I(X; Z)$.

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Theorem 7 (Fano's inequality)

$$h(P_e) + P_e \log |\mathcal{X}| \ge H(X|\hat{X}) \ge H(X|Y)$$

for
$$\hat{X} = g(X)$$
 and $P_e = \Pr \left[\hat{X} \neq X \right]$.

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Theorem 7 (Fano's inequality)

For any rvs X and Y, and any (even random) g, it holds that

$$h(P_e) + P_e \log |\mathcal{X}| \ge H(X|\hat{X}) \ge H(X|Y)$$

for
$$\hat{X} = g(X)$$
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▶ The inequality can be weekend to $1 + P_e \log |\mathcal{X}| \ge H(X|Y)$

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- We call \hat{X} the estimator for X.

$$\blacktriangleright \text{ Let } E = \left\{ \begin{array}{ll} 1, & \hat{X} \neq X \\ 0, & \hat{X} = X. \end{array} \right.$$

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Let X and Y be rvs, let $\hat{X} = g(Y)$ and $P_e = \Pr \left[\hat{X} \neq X \right]$.

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- ► Since $X \to Y \to \hat{X}$, it holds that $I(X; Y) \ge I(X; \hat{X})$ $\implies H(X|\hat{X}) \ge H(X|Y)$