Application of Information Theory, Lecture 9 Parallel Repetition of Interactive Arguments

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Part I

Interactive Proofs and Arguments

\mathcal{NP} as a Non-interactive Proofs

Definition 1 (\mathcal{NP})

 $\mathcal{L} \in \mathcal{NP}$ iff \exists and poly-time algorithm \lor such that:

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- ▶ V(x, w) = 0 for every $x \notin \mathcal{L}$ and $w \in \{0, 1\}^*$

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- Efficient verifier, efficient prover (given the witness)
- Soundness holds unconditionally

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Completeness $\forall x \in \mathcal{L}$: $Pr[(P, V)(x) = 1] \ge 2/3$.

Soundness $\forall x \notin \mathcal{L}$, and any algorithm P*: $\Pr[(P^*, V)(x) = 1] \leq 1/3$.

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- Games no-input protocols.

Section 1

Interactive Proof for Graph Non-Isomorphism

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- We will show a simple interactive proof for GNT Idea: Beer tasting...

Interactive proof for \mathcal{GNI}

Protocol 4 ((P, V)(G₀ = ([m], E₀), G₁ = ([m], E₁)))

- 1. V chooses $b \leftarrow \{0,1\}$ and $\pi \leftarrow \Pi_m$, and sends $\pi(E_b)$ to P.^a
- **2.** P send b' to V (tries to set b' = b).
- **3.** V accepts iff b' = b.
 - ${}^{a}\pi(E) = \{(\pi(u), \pi(v) : (u, v) \in E\}.$

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Claim 5

The above protocol is IP for \mathcal{GNI} , with perfect completeness and soundness error $\frac{1}{2}$.

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Hence,

$$G_0 \equiv G_1$$
: $\Pr[b' = b] \le \frac{1}{2}$.
 $G_0 \not\equiv G_1$: $\Pr[b' = b] = 1$ (i.e., *b* can be extracted from $\pi(E_i)$)



Part II

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- ▶ Public-coin interactive proof/argument in each round the verifier flips coins and sends them to the prover. To compute its output, the verifier applies some (fixed) function to the protocol's transcript.

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- ▶ In the following we focus on games (no input protocols)

Section 2

Parallel repetition of public-coin interactive argument



Theorem 6

Let $\pi = (P, V)$ be m-round, public-coin protocol with $\Pr\left[(\widetilde{P}, V) = 1\right] \le \varepsilon$ for any s-size \widetilde{P} , then $\Pr\left[(\widetilde{P^{(k)}}, V^{(k)}) = 1^k\right] \le \varepsilon^{k/4}$ for any $s \cdot \frac{\varepsilon^{k/4}}{mk^3s_V}$ -size $\widetilde{P^{(k)}}$, where s_V is V's size.

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Proof plan: Let $\widetilde{\mathsf{P}^{(k)}}$ be $s^{(k)}$ -size algorithm with $\Pr\left[(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)})=1^k\right]=\varepsilon^{(k)}$, we construct $s^{(k)}\cdot\frac{mk^3\mathsf{s}_\mathsf{V}}{\varepsilon^{(k)}}$ -size $\widetilde{\mathsf{P}}$ with $\Pr\left[(\widetilde{\mathsf{P}},\mathsf{V})=1\right]\geq(\varepsilon^{(k)})^{4/k}$.

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- ▶ Let $x_{\leq j} = x_1, ..., x_j$ (hence $R_{\leq j}$ denote the coins of first j rounds).

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- Assume for simplicity that P^(k) is deterministic
- Assume wlg. that V sends the first message in π and that in each round it sends ℓ coins.
- ▶ We view the coins of $V^{(k)}$ as a matrix $R \in \{0,1\}^{m \times (k\ell)}$, letting R_j denote the coins of the j'th round
- Let $x_{< j} = x_1, \dots, x_j$ (hence $R_{< j}$ denote the coins of first j rounds).
- ▶ Let $\mathbf{R} \sim \{0,1\}^{m \times (k\ell)}$

Algorithm $\widetilde{\mathsf{P}}$

Algorithm \widetilde{P} Let $q = k^2$.

Algorithm P

Let $q = k^2$.

Algorithm 7 (\widetilde{P})

- 1. Let $i^* \leftarrow [k]$.
- **2.** Upon getting the j'th round message r from V, do:
 - **2.1** Let $R \leftarrow \{0,1\}^{m \times (k\ell)}$, conditioned on $R_{\leq j-1} = \widetilde{R}_{\leq j-1}$ and $R_{j,j^*} = r$.
 - **2.2** If $(P^{(k)}, V^{(k)}(R)) = 1^k$:
 - **2.2.1** Set $\widetilde{R}_j = R_j$
 - **2.2.2** Send a_{j,i^*} back to V, for a_j being the j'th message $P^{(k)}$ send to $V^{(k)}$ in $(\widetilde{P^{(k)}}, V^{(k)}(R))$.

Else, GOTO Line 2.1

2.3 Abort, if overall number of sampling exceeds $\lceil 2qm/\varepsilon^{(k)} \rceil$.

Algorithm P

Let $q = k^2$.

Algorithm 7 (\widetilde{P})

- 1. Let $i^* \leftarrow [k]$.
- **2.** Upon getting the *j*'th round message *r* from V, do:
 - **2.1** Let $R \leftarrow \{0,1\}^{m \times (k\ell)}$, conditioned on $R_{\leq j-1} = \widetilde{R}_{\leq j-1}$ and $R_{j,j^*} = r$.
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- **2.3** Abort, if overall number of sampling exceeds $\lceil 2qm/\varepsilon^{(k)} \rceil$.
- ▶ Let \widetilde{P}' be the non aborting variant of \widetilde{P} , let \widetilde{R} and \widetilde{N} be the value of \widetilde{R} and # of samples done in a random execution of $(\widetilde{P}', V^{(k)})$, respectively.

Algorithm P

Let $q = k^2$.

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- 1. Let $i^* \leftarrow [k]$.
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- ▶ Let \widetilde{P}' be the non aborting variant of \widetilde{P} , let \widetilde{R} and \widetilde{N} be the value of \widetilde{R} and # of samples done in a random execution of $(\widetilde{P}', V^{(k)})$, respectively.
- $\qquad \qquad \Pr\left[(\widetilde{P},V)=1\right] \geq \Pr\left[\text{win}(\widetilde{\boldsymbol{R}},\widetilde{\boldsymbol{N}}) := (\widetilde{P^{(k)}},V^{(k)}(\widetilde{\boldsymbol{R}})) = 1^k \wedge \widetilde{\boldsymbol{N}} \leq 2qm/\varepsilon^{(k)} \right].$

Experiment 8 (P)

- 1. Let $R \leftarrow \{0,1\}^{m \times (k\ell)}$, conditioned on $R_{\leq j-1} = \widehat{R}_{\leq j-1}$.
- **2.** If $(\widetilde{P^{(k)}}, V^{(k)}(R)) = 1^k$, set $\widehat{R}_j = R_j$. Else, GOTO Line 1.

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Ideal "attacker"

Experiment 8 (\widehat{P})

For j = 1 to m:

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- Let $\hat{\mathbf{R}}$ be the value of $\hat{\mathbf{R}}$ in the end of a random execution of $\hat{\mathbf{P}}$.
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Lemma 9

$$\Pr\left[\widehat{m{N}}>qm/arepsilon^{(k)}
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- Let \hat{N} be # of samples done in \hat{P} .

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$$\Pr\left[\widehat{m{N}}>qm/arepsilon^{(k)}
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$$\implies \Pr\left[\mathsf{win}(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}) \right] = \Pr\left[(\widetilde{\mathsf{P}^{(k)}}, \mathsf{V}^{(k)}(\widehat{\mathbf{R}})) = 1^k \wedge \widehat{\mathbf{N}} \leq 2qm/\varepsilon^{(k)} \right] \geq 1 - \tfrac{1}{q}$$

Proving Lemma 9 —
$$\Pr\left[\widehat{\mathbf{N}} > qm/\varepsilon^{(k)}\right] < \frac{1}{q}$$

▶ Let $(X_1, \ldots, X_m) = \mathbf{R}$ and $(Y_1, \ldots, Y_m) = \widehat{\mathbf{R}} (\sim \mathbf{R}|_{(\widetilde{\mathbf{P}(k)} \ V^{(k)}(\mathbf{R})) = 1^k})$

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- For $\mathbf{y} \in \operatorname{Supp}(Y_{\leq j})$, let $\nu(\mathbf{y}) := \Pr\left[(\widetilde{\mathsf{P}^{(k)}}, \mathsf{V}^{(k)}(X^m) = 1^k \mid X_{\leq j} = \mathbf{y} \right]$

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Claim 10

For $j \in \{0, \dots, m-1\}$ and $\mathbf{y} \in \operatorname{Supp}(Y_{\leq j})$, it holds that $\Pr_{Y_{\leq j}}[\mathbf{y}] = \Pr_{X_{\leq j}}[\mathbf{y}] \cdot \frac{v(\mathbf{y})}{\varepsilon^{(k)}}$

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Hence,
$$E\left[\frac{1}{\nu(Y_{\leq i})}\right] = \sum_{\mathbf{y} \in \text{Supp}(Y_{\leq i})} \Pr[Y_{\leq i} = \mathbf{y}] \cdot \frac{1}{\nu(\mathbf{y})}$$

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$$= \sum_{\mathbf{y}} \mathsf{Pr}[X_{\leq j} = \mathbf{y}] \cdot \frac{v(\mathbf{y})}{\varepsilon^{(k)}} \cdot \frac{1}{v(\mathbf{y})} = \frac{1}{\varepsilon^{(k)}} \cdot \sum_{\mathbf{y} \in \mathsf{Supp}(Y_{\leq j})} \mathsf{Pr}[X_{\leq j} = \mathbf{y}] \leq \frac{1}{\varepsilon^{(k)}}. \ \Box$$

Proving Claim 10 —
$$\Pr_{Y_{\leq j}}[\mathbf{y}] = \Pr_{X^j}[\mathbf{y}] \cdot \frac{v(\mathbf{y})}{\varepsilon^{(k)}}$$

Recall $v(\mathbf{y}) := \Pr\left[(\widetilde{P^{(k)}}, V^{(k)}(X^m) = 1^k \mid X_{\leq j} = \mathbf{y}\right].$

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$$\Pr_{\mathbf{Y}_{j}|\mathbf{Y}_{\leq j-1} = \mathbf{y}_{\leq j-1}}[y_{j}] = \sum_{\ell=1}^{\infty} (1 - \nu(\mathbf{y}_{\leq j-1}))^{\ell-1} \cdot \Pr_{\mathbf{X}_{j}|\mathbf{X}_{\leq j-1} = \mathbf{y}_{\leq j-1}}[y_{j}] \cdot \nu(\mathbf{y}) \qquad (1)$$

$$= \frac{1}{\nu(\mathbf{y}_{\leq j-1})} \cdot \Pr_{\mathbf{X}_{j}|\mathbf{X}_{\leq j-1} = \mathbf{y}_{\leq j-1}}[y_{j}] \cdot \nu(\mathbf{y})$$

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$$= \Pr_{\mathbf{X} \leq j-1} [\mathbf{y}_{\leq j-1}] \cdot \frac{v(\mathbf{y}_{\leq j-1})}{\varepsilon^{(k)}} \cdot \Pr_{\mathbf{Y}_{j} \mid \mathbf{Y}_{\leq j-1} = \mathbf{y}_{\leq j-1}} [y_{j}] \tag{i.h.}$$

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$$\Pr_{Y \leq j}[\mathbf{y}] = \Pr_{X^j}[\mathbf{y}] \cdot \frac{v(\mathbf{y})}{\varepsilon^{(k)}}$$

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&\Pr_{Y \leq j}[\mathbf{y}] = \Pr_{Y \leq j-1}[\mathbf{y}_{\leq j-1}] \cdot \Pr_{Y_{j}|Y_{\leq j-1} = \mathbf{y}_{\leq j-1}}[y_{j}] \\
&= \Pr_{X \leq j-1}[\mathbf{y}_{\leq j-1}] \cdot \frac{\nu(\mathbf{y}_{\leq j-1})}{\varepsilon^{(k)}} \cdot \Pr_{Y_{j}|Y_{\leq j-1} = \mathbf{y}_{\leq j-1}}[y_{j}] \\
&= \Pr_{X \leq j-1}[\mathbf{y}_{\leq j-1}] \cdot \frac{\nu(\mathbf{y}_{\leq j-1})}{\varepsilon^{(k)}} \cdot \frac{\nu(\mathbf{y})}{\nu(\mathbf{y}_{\leq j-1})} \cdot \Pr_{X_{j}|X_{\leq j-1} = \mathbf{y}_{\leq j-1}}[y_{j}] \quad \text{(i.h.)}
\end{aligned}$$

Proving Claim 10 —
$$\Pr_{Y \leq j}[y] = \Pr_{X^j}[y] \cdot \frac{v(y)}{\varepsilon^{(k)}}$$

Recall
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$$= \frac{1}{v(\mathbf{y}_{< j-1})} \cdot \Pr_{\mathbf{X}_{j}|\mathbf{X}_{\leq j-1} = \mathbf{y}_{\leq j-1}}[y_{j}] \cdot v(\mathbf{y})$$

$$\Pr_{\mathbf{Y} \leq j} [\mathbf{y}] = \Pr_{\mathbf{Y} \leq j-1} [\mathbf{y}_{\leq j-1}] \cdot \Pr_{\mathbf{Y}_{j} \mid \mathbf{Y}_{\leq j-1} = \mathbf{y}_{\leq j-1}} [y_{j}]$$

$$= \Pr_{\mathbf{X} \leq j-1} [\mathbf{y}_{\leq j-1}] \cdot \frac{\mathbf{v}(\mathbf{y}_{\leq j-1})}{\varepsilon^{(k)}} \cdot \Pr_{\mathbf{Y}_{j} \mid \mathbf{Y}_{\leq j-1} = \mathbf{y}_{\leq j-1}} [y_{j}] \qquad (i.h.)$$

$$= \Pr_{\mathbf{X} \leq j-1} [\mathbf{y}_{\leq j-1}] \cdot \frac{\mathbf{v}(\mathbf{y}_{\leq j-1})}{\varepsilon^{(k)}} \cdot \frac{\mathbf{v}(\mathbf{y})}{\mathbf{v}(\mathbf{y}_{\leq j-1})} \cdot \Pr_{\mathbf{X}_{j} \mid \mathbf{X}_{\leq j-1} = \mathbf{y}_{\leq j-1}} [y_{j}] \qquad (Eq. (1))$$

$$= \Pr_{\mathbf{X} \leq i} [\mathbf{y}] \cdot \frac{\mathbf{v}(\mathbf{y})}{\varepsilon^{(k)}}.$$

- 1. Let $i^* \leftarrow [k]$.
- **2.** For j = 1 to m:
 - **2.1** Let $R \leftarrow \{0,1\}^{m \times (k\ell)}$, conditioned on $R_{\leq j-1} = \widehat{R}_{\leq j-1}$.
 - **2.2** If $(P^{(k)}, V^{(k)}(R)) = 1^k$, set $\hat{R}_{j,i^*} = R_{j,i^*}$. Else, GOTO Line 2.1.
 - **2.3** Let $R \leftarrow \{0,1\}^{m \times \ell}$, conditioned on $R_{\leq j-1} = \widehat{R}_{\leq j-1}$ and $R_{j,j^*} = \widehat{R}_{j,j^*}$.
 - **2.4** If $(P^{(k)}, V^{(k)}(R)) = 1^k$, set $\hat{R}_i = R_i$. Else, GOTO Line 2.3.

Experiment 11 (\hat{P})

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 - Let $\hat{\mathbf{R}}$ be the final value of $\hat{\mathbf{R}}$ in $\hat{\mathbf{P}}$.
- $\blacktriangleright |\widehat{\mathbf{R}} \sim \mathbf{R}|_{(\widetilde{\mathsf{P}^{(k)}},\mathsf{V}^{(k)}(\mathbf{R}))=1^k}$
- ▶ Let \hat{N} be the # of Step-2.3-samples done in \hat{P} .

Lemma 12 (essentially the same proof as of Lemma 9)

$$\Pr\left[\text{win}(\widehat{\boldsymbol{\textit{R}}},\widehat{\boldsymbol{\textit{N}}})\right] = \Pr\left[(\widetilde{P^{(k)}},V^{(k)}(\widehat{\boldsymbol{\textit{R}}})) = 1^k \wedge \widehat{\boldsymbol{\textit{N}}} \leq 2qm/\varepsilon^{(k)}\right] \geq 1 - \frac{1}{q}$$

Let $\widetilde{\mathbf{I}}$ be the value of i^* in $\widetilde{\mathbf{P}}$.

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$$D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}||_{\widetilde{\mathbf{I}}=i}).$$

Let $\tilde{\mathbf{I}}$ be the value of i^* in $\tilde{\mathbf{P}}$.

Claim 13

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} ||_{\widetilde{\mathbf{I}} = i}).$$

$$\sum_{i\in[k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}|_{\widetilde{\mathbf{I}}=i}) \leq D(\widehat{\mathbf{R}}||\mathbf{R}).$$

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Claim 14

$$\sum_{i\in[k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}|_{\widetilde{\mathbf{I}}=i}) \leq D(\widehat{\mathbf{R}}||\mathbf{R}).$$

1. Thm. 7 in Lecture 7 $\implies D(\widehat{\mathbf{R}}||\mathbf{R}) \leq \log \frac{1}{\Pr[(\widehat{\mathbf{P}^{(k)}}, \mathbf{V}^{(k)}(\mathbf{R})) = 1^k]} \leq \log \frac{1}{\varepsilon^{(k)}}$

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- 2. Hence, $D(\min(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})||\min(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})) \leq D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}||\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq -\frac{1}{k} \cdot \log \varepsilon^{(k)}$

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- 3. Lemma 12 $\implies \alpha := \Pr[\min(\widehat{\mathbf{R}}, \widehat{\mathbf{N}})] \ge 1 \frac{1}{q}$, and let $\beta := \Pr[\min(\widetilde{\mathbf{R}}, \widetilde{\mathbf{N}})]$.

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- **4.** By (2), $\alpha \cdot \log \frac{\alpha}{\beta} + (1 \alpha) \log \frac{1 \alpha}{1 \beta} \le -\frac{1}{k} \cdot \log \varepsilon^{(k)}$

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- **4.** By (2), $\alpha \cdot \log \frac{\alpha}{\beta} + (1 \alpha) \log \frac{1 \alpha}{1 \beta} \le -\frac{1}{k} \cdot \log \varepsilon^{(k)}$ $\implies \beta \ge 2^{\log \alpha + \frac{1 \alpha}{\alpha} \log(1 \alpha) + \frac{1}{\alpha k} \log \varepsilon^{(k)}}$

Let \tilde{I} be the value of i^* in \tilde{P} .

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$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} ||_{\widetilde{\mathbf{I}} = i}).$$

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- **4.** By (2), $\alpha \cdot \log \frac{\alpha}{\beta} + (1 \alpha) \log \frac{1 \alpha}{1 \beta} \le -\frac{1}{k} \cdot \log \varepsilon^{(k)}$ $\implies \beta > 2^{\log \alpha + \frac{1 \alpha}{\alpha} \log(1 \alpha) + \frac{1}{\alpha k} \log \varepsilon^{(k)}}$
- **5.** Since $q = k^2$: $\alpha \ge 2^{-\frac{2}{q}} \ge 2^{-\frac{1}{k}}$ and $\frac{1-\alpha}{\alpha} \log(1-\alpha) \ge -\frac{4 \log k}{k^2} \ge -\frac{1}{k}$

From ideal to real

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- 4. By (2), $\alpha \cdot \log \frac{\alpha}{\beta} + (1 \alpha) \log \frac{1 \alpha}{1 \beta} \le -\frac{1}{k} \cdot \log \varepsilon^{(k)}$ $\implies \beta \ge 2^{\log \alpha + \frac{1 \alpha}{\alpha} \log(1 \alpha) + \frac{1}{\alpha k} \log \varepsilon^{(k)}}$
- 5. Since $q=k^2$: $\alpha \geq 2^{-\frac{2}{q}} \geq 2^{-\frac{1}{k}}$ and $\frac{1-\alpha}{\alpha}\log(1-\alpha) \geq -\frac{4\log k}{k^2} \geq -\frac{1}{k}$
- **6.** We conclude that $\beta \geq 2^{\frac{4}{k}\log \varepsilon^{(k)}} = \sqrt[k]{\xi^{(k)}}$.

Proving Claim 14 — $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widetilde{\mathbf{R}}|_{\widetilde{\mathbf{I}}=i}) \leq D(\widehat{\mathbf{R}}||\mathbf{R})$

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Lemma 15

Let $Z = \{Z_{ij}\}_{(i,j) \in [k] \times [m]}$ be iids and let W be an event. For $z \in \text{Supp}(Z)$, let

$$\xi_{i}(z) := \prod_{j=1}^{m} \Pr[Z_{j,i} = z_{i,j}] \cdot \Pr[Z_{j,-i} = z_{i,j-1} | Z_{1,...,j-1} = z_{1,...,j-1} \wedge Z_{j,i} = z_{i,j} \wedge W].$$

Then $\sum_{i=1}^{k} D(Z|_{W}||\xi_{i}) \leq D(Z|_{W}||Z)$.

Proving Claim 14 — $\sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}}|_{\widetilde{\mathbf{I}} = i}) \leq D(\widehat{\mathbf{R}} || \mathbf{R})$

Lemma 15

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Then $\sum_{i=1}^{k} D(Z|_{W}||\xi_{i}) \leq D(Z|_{W}||Z)$.

Letting
$$Z = \mathbf{R}$$
 and W be the event $(\widehat{\mathsf{P}^{(k)}}, \mathsf{V}^{(k)}(\mathbf{R})) = 1^k$, Lemma 15 yields that $\sum_{i \in [k]} D(\widehat{\mathbf{R}}||\widehat{\mathbf{R}}_{\widetilde{\mathbf{l}}=i}) = \sum_{i \in [k]} D(\mathbf{R}|_W||\widehat{\mathbf{R}}_{\widetilde{\mathbf{l}}=i}) \le D(\mathbf{R}|_W||\mathbf{R}) = D(\widehat{\mathbf{R}}||\mathbf{R})$. \square

We prove for m = k = 2.

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$$Z = (X_0, X_1, Y_0, Y_1)$$
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$$\xi_{i}(x_{0}, x_{1}, y_{0}, y_{1}) := \Pr[X_{i} = x_{i}] \cdot \Pr[X_{-i} = x_{-i} \mid X_{i} = x_{i} \wedge W] \cdot \\ \Pr[Y_{i} = y_{i}] \cdot \Pr[Y_{-i} = Y_{-i} \mid Y_{i} = y_{i} \wedge (X_{0}, X_{1}) = (x_{0}, x_{1}) \wedge W].$$

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We need to prove that $D(Z|_W||Z) \ge \sum_{i=1}^2 D(Z|_W||\xi_i)$.

▶ Let $U = p_Z$ and $C = p_{Z|_W}$.

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- Let $U = p_Z$ and $C = p_{Z|_W}$. (Hence, we need to prove that $D(C||U) \ge \sum_{i=1}^2 D(Z|_W||\xi_i)$)
- ▶ Let $X = (X_0, X_1)$
- ► Let $Q(x_0, x_1, y_0, y_1) := \Pr[X_0 = x_0 | W] \cdot \Pr[X_1 = x_1 | W] \cdot \Pr[Y_0 = y_0 | W, X = (x_0, x_1)] \cdot \Pr[Y_1 = y_1 | W, X = (x_0, x_1)]$

We prove for m = k = 2.

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 - (Hence, we need to prove that $D(C||U) \ge \sum_{i=1}^{2} D(Z|_{W}||\xi_{i})$)
- ▶ Let $X = (X_0, X_1)$
- ► Let $Q(x_0, x_1, y_0, y_1) := \Pr[X_0 = x_0 | W] \cdot \Pr[X_1 = x_1 | W] \cdot \Pr[Y_0 = y_0 | W, X = (x_0, x_1)] \cdot \Pr[Y_1 = y_1 | W, X = (x_0, x_1)]$
- ► We write $\frac{C(x_0, x_1, y_0, y_1)}{U(x_0, x_1, y_0, y_1)} = \frac{\Pr[X_0 = x_0 | W] \cdot \Pr[Y_0 = y_0 | W, X = (x_0, x_1)]}{\Pr[X_0 = x_0] \cdot \Pr[Y_0 = y_0]} \cdot \frac{\Pr[X_1 = x_1 | W] \cdot \Pr[Y_1 = y_1 | W, X = (x_0, x_1)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \cdot \frac{C(x_0, x_1, y_0, y_1)}{Q(x_0, x_1, y_0, y_1)}$

$$D(C||U) = \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{\mathsf{E}} \left[\log \frac{\Pr[X_0 = x_0|W] \cdot \Pr[Y_0 = y_0|W, X = (x_0, x_1)]}{\Pr[X_0 = x_0] \cdot \Pr[Y_0 = y_0]} \right]$$

$$+ \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{\mathsf{E}} \left[\log \frac{\Pr[X_1 = x_1|W] \cdot \Pr[Y_1 = y_1|W, X = (x_0, x_1)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \right]$$

$$+ \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{\mathsf{E}} \left[\log \frac{C(x_0, x_1, y_0, y_1)}{Q(x_0, x_1, y_0, y_1)} \right].$$

$$D(C||U) = \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{E} \left[log \frac{Pr[X_0 = x_0|W] \cdot Pr[Y_0 = y_0|W, X = (x_0, x_1)]}{Pr[X_0 = x_0] \cdot Pr[Y_0 = y_0]} \right]$$

$$+ \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{E} \left[log \frac{Pr[X_1 = x_1|W] \cdot Pr[Y_1 = y_1|W, X = (x_0, x_1)]}{Pr[X_1 = x_1] \cdot Pr[Y_1 = y_1]} \right]$$

$$+ \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{E} \left[log \frac{C(x_0, x_1, y_0, y_1)}{Q(x_0, x_1, y_0, y_1)} \right].$$

It follows that

$$\begin{split} D(C||U) &= D(X_0|_W, X_1|_{W,X_0}, Y_0|_{W,X}, Y_1|_{W,X,Y_0}||X_0, X_1|_{W,X_0}, Y_0, Y_1|_{W,X,Y_0}) \\ &+ D(X_1|_W, X_0|_{W,X_1}, Y_1|_{W,X}, Y_0|_{W,X,Y_1}||X_1, X_0|_{W,X_1}, Y_1, Y_0|_{W,X,Y_1}) \\ &+ D(C||Q) \end{split}$$

$$\begin{split} D(C||U) &= \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{\mathsf{E}} \left[\log \frac{\Pr[X_0 = x_0|W] \cdot \Pr[Y_0 = y_0|W, X = (x_0, x_1)]}{\Pr[X_0 = x_0] \cdot \Pr[Y_0 = y_0]} \right] \\ &+ \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{\mathsf{E}} \left[\log \frac{\Pr[X_1 = x_1|W] \cdot \Pr[Y_1 = y_1|W, X = (x_0, x_1)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \right] \\ &+ \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{\mathsf{E}} \left[\log \frac{C(x_0, x_1, y_0, y_1)}{Q(x_0, x_1, y_0, y_1)} \right]. \end{split}$$

It follows that

$$D(C||U) = D(X_0|_W, X_1|_{W,X_0}, Y_0|_{W,X}, Y_1|_{W,X,Y_0}||X_0, X_1|_{W,X_0}, Y_0, Y_1|_{W,X,Y_0})$$

$$+ D(X_1|_W, X_0|_{W,X_1}, Y_1|_{W,X}, Y_0|_{W,X,Y_1}||X_1, X_0|_{W,X_1}, Y_1, Y_0|_{W,X,Y_1})$$

$$+ D(C||Q)$$

$$= \sum_{i=1}^{2} D(Z|_W||\xi_i) + D(C||Q)$$

$$D(C||U) = \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{\mathsf{E}} \left[\log \frac{\Pr[X_0 = x_0|W] \cdot \Pr[Y_0 = y_0|W, X = (x_0, x_1)]}{\Pr[X_0 = x_0] \cdot \Pr[Y_0 = y_0]} \right]$$

$$+ \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{\mathsf{E}} \left[\log \frac{\Pr[X_1 = x_1|W] \cdot \Pr[Y_1 = y_1|W, X = (x_0, x_1)]}{\Pr[X_1 = x_1] \cdot \Pr[Y_1 = y_1]} \right]$$

$$+ \underset{(x_0, x_1, y_0, y_1) \leftarrow C}{\mathsf{E}} \left[\log \frac{C(x_0, x_1, y_0, y_1)}{Q(x_0, x_1, y_0, y_1)} \right].$$

It follows that

$$\begin{split} D(C||U) &= D(X_0|_W, X_1|_{W,X_0}, Y_0|_{W,X}, Y_1|_{W,X,Y_0}||X_0, X_1|_{W,X_0}, Y_0, Y_1|_{W,X,Y_0}) \\ &+ D(X_1|_W, X_0|_{W,X_1}, Y_1|_{W,X}, Y_0|_{W,X,Y_1}||X_1, X_0|_{W,X_1}, Y_1, Y_0|_{W,X,Y_1}) \\ &+ D(C||Q) \\ &= \sum_{i=1}^2 D(Z|_W||\xi_i) + D(C||Q) \\ &\geq \sum_{i=1}^2 D(Z|_W||\xi_i). \Box \end{split}$$

Proving Claim 13 — $D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} ||_{\widetilde{\mathbf{I}} = i})$

Let $\widehat{\mathbf{I}}$ be the value of i^* in $\widehat{\mathbf{P}}$ (recall that $\widetilde{\mathbf{I}}$ is the value of i^* in $\widetilde{\mathbf{P}}$).

Proving Claim 13 — $D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} ||_{\widetilde{\mathbf{I}} = i})$

Let $\widehat{\mathbf{I}}$ be the value of i^* in $\widehat{\mathbf{P}}$ (recall that $\widetilde{\mathbf{I}}$ is the value of i^* in $\widetilde{\mathbf{P}}$).

$$\text{Let }(\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)})=(\widetilde{\mathbf{R}},\widetilde{\mathbf{N}})|_{\widetilde{\mathbf{I}}=i} \text{ and } (\widehat{R}_{(i)},\widehat{\mathbf{N}}_{(i)})=(\widehat{\mathbf{R}},\widehat{\mathbf{N}})|_{\widehat{\mathbf{I}}=i}. \text{ Note that } \widehat{R}_{(i)}=\widehat{\mathbf{R}}.$$

Proving Claim 13 — $D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} |_{\widetilde{\mathbf{I}} = i})$

Let \widehat{I} be the value of i^* in \widehat{P} (recall that \widetilde{I} is the value of i^* in \widetilde{P}).

Let
$$(\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)})=(\widetilde{\mathbf{R}},\widetilde{\mathbf{N}})|_{\widetilde{\mathbf{I}}=i}$$
 and $(\widehat{R}_{(i)},\widehat{\mathbf{N}}_{(i)})=(\widehat{\mathbf{R}},\widehat{\mathbf{N}})|_{\widehat{\mathbf{I}}=i}$. Note that $\widehat{R}_{(i)}=\widehat{\mathbf{R}}$.

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}, \widehat{\mathbf{I}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}, \widetilde{\mathbf{I}})$$

Proving Claim 13 —
$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} ||_{\widetilde{\mathbf{I}} = i})$$

Let \widehat{I} be the value of i^* in \widehat{P} (recall that \widetilde{I} is the value of i^* in \widetilde{P}).

$$\text{Let }(\widetilde{\textbf{R}}_{(i)},\widetilde{\textbf{N}}_{(i)})=(\widetilde{\textbf{R}},\widetilde{\textbf{N}})|_{\widetilde{\textbf{I}}=i} \text{ and } (\widehat{R}_{(i)},\widehat{\textbf{N}}_{(i)})=(\widehat{\textbf{R}},\widehat{\textbf{N}})|_{\widehat{\textbf{I}}=i}. \text{ Note that } \widehat{R}_{(i)}=\widehat{\textbf{R}}.$$

$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}}, \widehat{\mathbf{I}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}, \widetilde{\mathbf{I}})$$

(data-processing)

Proving Claim 13 —
$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} ||_{\widetilde{\mathbf{I}} = i})$$

Let \hat{I} be the value of i^* in \hat{P} (recall that \hat{I} is the value of i^* in \hat{P}).

$$\text{Let }(\widetilde{\textbf{R}}_{(i)},\widetilde{\textbf{N}}_{(i)})=(\widetilde{\textbf{R}},\widetilde{\textbf{N}})|_{\widetilde{\textbf{I}}=i} \text{ and } (\widehat{R}_{(i)},\widehat{\textbf{N}}_{(i)})=(\widehat{\textbf{R}},\widehat{\textbf{N}})|_{\widehat{\textbf{I}}=i}. \text{ Note that } \widehat{R}_{(i)}=\widehat{\textbf{R}}.$$

$$\begin{split} D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) &\leq D(\widehat{\mathbf{R}},\widehat{\mathbf{N}},\widehat{\mathbf{I}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}},\widetilde{\mathbf{I}}) \\ &= D(\widehat{\mathbf{I}}||\widetilde{\mathbf{I}}) + \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{(i)},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \end{split} \tag{data-processing}$$

Proving Claim 13 —
$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} ||_{\widetilde{\mathbf{I}} = i})$$

Let \hat{I} be the value of i^* in \hat{P} (recall that \hat{I} is the value of i^* in \hat{P}).

$$\text{Let }(\widetilde{\textbf{R}}_{(i)},\widetilde{\textbf{N}}_{(i)})=(\widetilde{\textbf{R}},\widetilde{\textbf{N}})|_{\widetilde{\textbf{I}}=i} \text{ and } (\widehat{R}_{(i)},\widehat{\textbf{N}}_{(i)})=(\widehat{\textbf{R}},\widehat{\textbf{N}})|_{\widehat{\textbf{I}}=i}. \text{ Note that } \widehat{R}_{(i)}=\widehat{\textbf{R}}.$$

$$\begin{split} D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) &\leq D(\widehat{\mathbf{R}},\widehat{\mathbf{N}},\widehat{\mathbf{I}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}},\widetilde{\mathbf{I}}) \\ &= D(\widehat{\mathbf{I}}||\widetilde{\mathbf{I}}) + \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{(i)},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \end{split} \tag{chain rule}$$

Proving Claim 13 —
$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} ||_{\widetilde{\mathbf{I}} = i})$$

Let $\widehat{\mathbf{I}}$ be the value of i^* in $\widehat{\mathbf{P}}$ (recall that $\widetilde{\mathbf{I}}$ is the value of i^* in $\widetilde{\mathbf{P}}$).

$$\text{Let }(\widetilde{\textbf{R}}_{(i)},\widetilde{\textbf{N}}_{(i)})=(\widetilde{\textbf{R}},\widetilde{\textbf{N}})|_{\widetilde{\textbf{I}}=i} \text{ and } (\widehat{R}_{(i)},\widehat{\textbf{N}}_{(i)})=(\widehat{\textbf{R}},\widehat{\textbf{N}})|_{\widehat{\textbf{I}}=i}. \text{ Note that } \widehat{R}_{(i)}=\widehat{\textbf{R}}.$$

$$\begin{split} D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) &\leq D(\widehat{\mathbf{R}},\widehat{\mathbf{N}},\widehat{\mathbf{I}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}},\widetilde{\mathbf{I}}) \\ &= D(\widehat{\mathbf{I}}||\widetilde{\mathbf{I}}) + \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{(i)},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \\ &= \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{i},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \end{split} \tag{chain rule}$$

Proving Claim 13 —
$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} |_{\widetilde{\mathbf{I}} = i})$$

Let $\widehat{\mathbf{I}}$ be the value of i^* in $\widehat{\mathbf{P}}$ (recall that $\widetilde{\mathbf{I}}$ is the value of i^* in $\widetilde{\mathbf{P}}$).

$$\text{Let }(\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)})=(\widetilde{\mathbf{R}},\widetilde{\mathbf{N}})|_{\widetilde{\mathbf{I}}=i} \text{ and } (\widehat{R}_{(i)},\widehat{\mathbf{N}}_{(i)})=(\widehat{\mathbf{R}},\widehat{\mathbf{N}})|_{\widehat{\mathbf{I}}=i}. \text{ Note that } \widehat{R}_{(i)}=\widehat{\mathbf{R}}.$$

$$\begin{split} D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) &\leq D(\widehat{\mathbf{R}},\widehat{\mathbf{N}},\widehat{\mathbf{I}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}},\widetilde{\mathbf{I}}) \\ &= D(\widehat{\mathbf{I}}||\widetilde{\mathbf{I}}) + \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{(i)},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \\ &= \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{i},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \end{split} \tag{chain rule}$$

$$D(\widehat{\boldsymbol{\mathsf{R}}}_{(i)},\widehat{\boldsymbol{\mathsf{N}}}_{(i)}||\widehat{\boldsymbol{\mathsf{R}}}_{(i)},\widetilde{\boldsymbol{\mathsf{N}}}_{(i)}) = D(\widehat{\boldsymbol{\mathsf{R}}}_{(i)}||\widehat{\boldsymbol{\mathsf{R}}}_{(i)}) + \underset{r \leftarrow \widehat{\boldsymbol{\mathsf{R}}}_{(i)}}{\mathsf{E}} \left[D(\widehat{\boldsymbol{\mathsf{N}}}_{(i)}|_{\widehat{\boldsymbol{\mathsf{R}}}_{(i)} = r}||\widehat{\boldsymbol{\mathsf{N}}}_{(i)}|_{\widehat{\boldsymbol{\mathsf{R}}}_{(i)} = r}) \right]$$

Proving Claim 13 —
$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} |_{\widetilde{\mathbf{I}} = i})$$

Let $\widehat{\mathbf{I}}$ be the value of i^* in $\widehat{\mathbf{P}}$ (recall that $\widetilde{\mathbf{I}}$ is the value of i^* in $\widetilde{\mathbf{P}}$).

$$\text{Let }(\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)})=(\widetilde{\mathbf{R}},\widetilde{\mathbf{N}})|_{\widetilde{\mathbf{I}}=i} \text{ and } (\widehat{R}_{(i)},\widehat{\mathbf{N}}_{(i)})=(\widehat{\mathbf{R}},\widehat{\mathbf{N}})|_{\widehat{\mathbf{I}}=i}. \text{ Note that } \widehat{R}_{(i)}=\widehat{\mathbf{R}}.$$

$$\begin{split} D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) &\leq D(\widehat{\mathbf{R}},\widehat{\mathbf{N}},\widehat{\mathbf{I}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}},\widetilde{\mathbf{I}}) \\ &= D(\widehat{\mathbf{I}}||\widetilde{\mathbf{I}}) + \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{(i)},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \\ &= \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{i},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \end{split} \tag{chain rule}$$

$$D(\widehat{\mathbf{R}}_{(i)}, \widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)}, \widetilde{\mathbf{N}}_{(i)}) = D(\widehat{\mathbf{R}}_{(i)}||\widetilde{\mathbf{R}}_{(i)}) + \underset{r \leftarrow \widehat{\mathbf{R}}_{(i)}}{\mathsf{E}} \left[D(\widehat{\mathbf{N}}_{(i)}|_{\widehat{\mathbf{R}}_{(i)}=r}||\widetilde{\mathbf{N}}_{(i)}|_{\widetilde{\mathbf{R}}_{(i)}=r}) \right] \quad \text{(chain rule)}$$

Proving Claim 13 —
$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} |_{\widetilde{\mathbf{I}} = i})$$

Let $\hat{\mathbf{I}}$ be the value of i^* in $\hat{\mathbf{P}}$ (recall that $\hat{\mathbf{I}}$ is the value of i^* in $\hat{\mathbf{P}}$).

$$\text{Let } (\widetilde{\textbf{R}}_{(i)},\widetilde{\textbf{N}}_{(i)}) = (\widetilde{\textbf{R}},\widetilde{\textbf{N}})|_{\widetilde{\textbf{I}}=i} \text{ and } (\widehat{R}_{(i)},\widehat{\textbf{N}}_{(i)}) = (\widehat{\textbf{R}},\widehat{\textbf{N}})|_{\widehat{\textbf{I}}=i}. \text{ Note that } \widehat{R}_{(i)} = \widehat{\textbf{R}}.$$

$$\begin{split} D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) &\leq D(\widehat{\mathbf{R}},\widehat{\mathbf{N}},\widehat{\mathbf{I}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}},\widetilde{\mathbf{I}}) \\ &= D(\widehat{\mathbf{I}}||\widetilde{\mathbf{I}}) + \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{(i)},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \\ &= \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{i},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \end{split} \tag{chain rule}$$

$$\begin{split} D(\widehat{\mathbf{R}}_{(i)},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widehat{\mathbf{N}}_{(i)}) &= D(\widehat{\mathbf{R}}_{(i)}||\widetilde{\mathbf{R}}_{(i)}) + \mathop{\mathbb{E}}_{r \leftarrow \widehat{\mathbf{R}}_{(i)}} \left[D(\widehat{\mathbf{N}}_{(i)}|_{\widehat{\mathbf{R}}_{(i)} = r}||\widetilde{\mathbf{N}}_{(i)}|_{\widetilde{\mathbf{R}}_{(i)} = r}) \right] \quad \text{(chain rule)} \\ &= D(\widehat{\mathbf{R}}_{(i)}||\widetilde{\mathbf{R}}_{(i)}) \end{split}$$

Proving Claim 13 —
$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} |_{\widetilde{\mathbf{I}} = i})$$

Let $\hat{\mathbf{I}}$ be the value of i^* in $\hat{\mathbf{P}}$ (recall that $\hat{\mathbf{I}}$ is the value of i^* in $\hat{\mathbf{P}}$).

$$\text{Let } (\widetilde{\textbf{R}}_{(i)},\widetilde{\textbf{N}}_{(i)}) = (\widetilde{\textbf{R}},\widetilde{\textbf{N}})|_{\widetilde{\textbf{I}}=i} \text{ and } (\widehat{R}_{(i)},\widehat{\textbf{N}}_{(i)}) = (\widehat{\textbf{R}},\widehat{\textbf{N}})|_{\widehat{\textbf{I}}=i}. \text{ Note that } \widehat{R}_{(i)} = \widehat{\textbf{R}}.$$

$$\begin{split} D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) &\leq D(\widehat{\mathbf{R}},\widehat{\mathbf{N}},\widehat{\mathbf{I}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}},\widetilde{\mathbf{I}}) \\ &= D(\widehat{\mathbf{I}}||\widetilde{\mathbf{I}}) + \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{(i)},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \\ &= \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{i},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \end{split} \tag{chain rule}$$

$$\begin{split} D(\widehat{\mathbf{R}}_{(i)}, \widehat{\mathbf{N}}_{(i)} || \widetilde{\mathbf{R}}_{(i)}, \widetilde{\mathbf{N}}_{(i)}) &= D(\widehat{\mathbf{R}}_{(i)} || \widetilde{\mathbf{R}}_{(i)}) + \mathop{\mathbb{E}}_{r \leftarrow \widehat{\mathbf{R}}_{(i)}} \left[D(\widehat{\mathbf{N}}_{(i)}|_{\widehat{\mathbf{R}}_{(i)} = r} || \widetilde{\mathbf{N}}_{(i)}|_{\widetilde{\mathbf{R}}_{(i)} = r}) \right] &\quad \text{(chain rule)} \\ &= D(\widehat{\mathbf{R}}_{(i)} || \widetilde{\mathbf{R}}_{(i)}) &\quad \text{(since } \widehat{\mathbf{N}}_{(i)}|_{\widehat{\mathbf{R}}_{(i)} = r} \equiv \widetilde{\mathbf{N}}_{(i)}|_{\widetilde{\mathbf{R}}_{(i)} = r}) \end{split}$$

Proving Claim 13 —
$$D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}} || \widetilde{\mathbf{R}} |_{\widetilde{\mathbf{I}} = i})$$

Let $\hat{\mathbf{I}}$ be the value of i^* in $\hat{\mathbf{P}}$ (recall that $\tilde{\mathbf{I}}$ is the value of i^* in $\tilde{\mathbf{P}}$).

Let
$$(\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)})=(\widetilde{\mathbf{R}},\widetilde{\mathbf{N}})|_{\widetilde{\mathbf{I}}=i}$$
 and $(\widehat{R}_{(i)},\widehat{\mathbf{N}}_{(i)})=(\widehat{\mathbf{R}},\widehat{\mathbf{N}})|_{\widehat{\mathbf{I}}=i}$. Note that $\widehat{R}_{(i)}=\widehat{\mathbf{R}}$.

$$\begin{split} D(\widehat{\mathbf{R}},\widehat{\mathbf{N}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}}) &\leq D(\widehat{\mathbf{R}},\widehat{\mathbf{N}},\widehat{\mathbf{I}}||\widetilde{\mathbf{R}},\widetilde{\mathbf{N}},\widetilde{\mathbf{I}}) \\ &= D(\widehat{\mathbf{I}}||\widetilde{\mathbf{I}}) + \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{(i)},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \\ &= \frac{1}{k} \sum_{i \in [k]} D(\widehat{\mathbf{R}}_{i},\widehat{\mathbf{N}}_{(i)}||\widetilde{\mathbf{R}}_{(i)},\widetilde{\mathbf{N}}_{(i)}) \end{split} \tag{chain rule}$$

For $i \in [k]$, it holds that

$$\begin{split} D(\widehat{\mathbf{R}}_{(i)}, \widehat{\mathbf{N}}_{(i)} || \widetilde{\mathbf{R}}_{(i)}, \widetilde{\mathbf{N}}_{(i)}) &= D(\widehat{\mathbf{R}}_{(i)} || \widetilde{\mathbf{R}}_{(i)}) + \mathop{\mathbb{E}}_{r \leftarrow \widehat{\mathbf{R}}_{(i)}} \left[D(\widehat{\mathbf{N}}_{(i)} ||_{\widehat{\mathbf{R}}_{(i)} = r} || \widetilde{\mathbf{N}}_{(i)} ||_{\widetilde{\mathbf{R}}_{(i)} = r}) \right] &\quad \text{(chain rule)} \\ &= D(\widehat{\mathbf{R}}_{(i)} || \widetilde{\mathbf{R}}_{(i)}) \quad \text{(since } \widehat{\mathbf{N}}_{(i)} ||_{\widehat{\mathbf{R}}_{(i)} = r} \equiv \widetilde{\mathbf{N}}_{(i)} ||_{\widetilde{\mathbf{R}}_{(i)} = r}) \end{split}$$

Hence, $D(\widehat{\mathbf{R}}, \widehat{\mathbf{N}} || \widetilde{\mathbf{R}}, \widetilde{\mathbf{N}}) \leq \frac{1}{k} \sum_{e[k]} D(\widehat{\mathbf{R}}_{(i)} || \widetilde{\mathbf{R}}_{(i)}) \square$

Similar proof to the public-coin proof we gave above.

- Similar proof to the public-coin proof we gave above.
- In each round, the attacker \widetilde{P} samples random continuations of $(\widetilde{P^{(k)}}, V^{(k)})$, till he gets an accepting execution.

- Similar proof to the public-coin proof we gave above.
- ▶ In each round, the attacker \widetilde{P} samples random continuations of $(\widetilde{P^{(k)}}, V^{(k)})$, till he gets an accepting execution.
- What fails us to extend this approach for non-public-coin interactive arguments?

Section 3

Parallel amplification for any interactive argument



Parallel amplification theorem for any protocol

Can we amplify the security of any interactive argument "in parallel"?

Parallel amplification theorem for any protocol

- Can we amplify the security of any interactive argument "in parallel"?
- Yes we can!

Relevant papers

Kai-Min Chung and Rafael Pass: Tight Parallel Repetition Theorems for Public-Coin Arguments using KL-divergence.

The proof given in class is in the spirit of this paper.