Foundation of Cryptography (0368-4162-01), Lecture 2 Pseudorandom Generators

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Section 1

Distributions and Statistical Distance

Distributions and Statistical Distance

Let P and Q be two distributions over a finite set \mathcal{U} . Their statistical distance (also known as, variation distance), denoted by $\mathrm{SD}(P,Q)$, is defined as

$$SD(P,Q) := \frac{1}{2} \sum_{x \in \mathcal{U}} |P(x) - Q(x)| = \max_{S \subseteq \mathcal{U}} (P(S) - Q(S))$$

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Claim 1

For any pair of (finite) distribution P and Q, it holds that such

$$SD(P,Q) = \max_{D} \left(Pr_{x \leftarrow P}[D(x) = 1] - Pr_{x \leftarrow Q}[D(x) = 1] \right),$$

where D is any algorithm.

Some useful facts

Let P, Q, R be finite distributions, then

Triangle inequality:

$$SD(P,R) \leq SD(P,Q) + SD(Q,R)$$

Repeated sampling:

$$SD((P, P), (Q, Q)) \le 2 \cdot SD(P, Q)$$

Random variables

Distribution ensembles and statistical indistinguishability

Definition 2 (distribution ensembles)

 $\mathcal{P} = \{P_n\}_{n \in \mathbb{N}}$ is a distribution ensemble, if P_n is a (finite) distribution for any $n \in \mathbb{N}$.

 \mathcal{P} is efficiently samplable (or just efficient), if $\exists \ \mathsf{PPT} \ Samp$ with $\mathsf{Sam}(\mathsf{1}^n) \equiv P_n$.

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Two distribution ensembles \mathcal{P} and \mathcal{Q} are statistically indistinguishable, if $SD(P_n, Q_n) = neg(n)$.

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Two distribution ensembles \mathcal{P} and \mathcal{Q} are statistically indistinguishable, if $SD(P_n, Q_n) = neg(n)$.

Alternatively, if $\left|\Delta_{(\mathcal{P},\mathcal{Q})}^{\mathsf{D}}(n)\right| = \mathsf{neg}(n)$, for *any* algorithm D, where

$$\Delta^{\mathsf{D}}_{(\mathcal{P},\mathcal{Q})}(n) := \mathsf{Pr}_{x \leftarrow P_n}[\mathsf{D}(1^n,x) = 1] - \mathsf{Pr}_{x \leftarrow Q_n}[\mathsf{D}(1^n,x) = 1].$$

Section 2

Computational Indistinguishability

Definition 4 (computational indistinguishability)

$$(\Delta^{\mathsf{D}}_{(\mathcal{P},\mathcal{Q})}(n) := \mathsf{Pr}_{\mathsf{x} \leftarrow P_n}[\Delta \mathsf{D}(\mathsf{1}^n,\mathsf{x}) = \mathsf{1}] - \mathsf{Pr}_{\mathsf{x} \leftarrow Q_n}[\mathsf{D}(\mathsf{1}^n,\mathsf{x}) = \mathsf{1}])$$

Definition 4 (computational indistinguishability)

Two distribution ensembles \mathcal{P} and \mathcal{Q} are *computationally indistinguishable*, if $\left|\Delta_{(\mathcal{P},\mathcal{Q})}^{\mathsf{D}}(n)\right| = \mathsf{neg}(n)$, for any PPT D.

$$(\Delta^{\mathsf{D}}_{(\mathcal{P},\mathcal{Q})}(n) := \mathsf{Pr}_{x \leftarrow P_n}[\Delta \mathsf{D}(1^n, x) = 1] - \mathsf{Pr}_{x \leftarrow Q_n}[\mathsf{D}(1^n, x) = 1])$$

• Can it be different from the statistical case?

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Question 5

Assume that \mathcal{P} and \mathcal{Q} are computationally indistinguishable, is it always true that $\mathcal{P}^2=(\mathcal{P},\mathcal{P})$ and $\mathcal{Q}^2=(\mathcal{Q},\mathcal{Q})$ are?

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Assume that $\left|\Delta^{\mathsf{D}}_{(\mathcal{P}^2,\mathcal{Q}^2)}(n)\right| = \delta(n)$ for some PPT D, we would like to prove that \exists PPT D' with $\left|\Delta^{\mathsf{D}}_{(\mathcal{P},\mathcal{Q})}(n)\right| \geq \delta(n)/2$ for every $n \in \mathbb{N}$.

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Assume that \mathcal{P} and \mathcal{Q} are computationally indistinguishable, is it always true that $\mathcal{P}^2 = (\mathcal{P}, \mathcal{P})$ and $\mathcal{Q}^2 = (\mathcal{Q}, \mathcal{Q})$ are?

Assume that $\left|\Delta^{\mathsf{D}}_{(\mathcal{P}^2,\mathcal{Q}^2)}(n)\right|=\delta(n)$ for some PPT D, we would like to prove that $\exists \ \mathsf{PPT} \ \mathsf{D}'$ with $\left|\Delta^{\mathsf{D}}_{(\mathcal{P},\mathcal{O})}(n)\right| \geq \delta(n)/2$ for every $n \in \mathbb{N}$. Indeed $\delta(n) = |\Pr_{x \leftarrow P^2}[D(x) = 1] - \Pr_{x \leftarrow Q^2}[D(x) = 1]|$ $\leq \left| \mathsf{Pr}_{x \leftarrow P_n^2}[\mathsf{D}(x) = 1] - \mathsf{Pr}_{x \leftarrow (P_n, Q_n)}[\mathsf{D}(x) = 1] \right|$ $+\left|\operatorname{Pr}_{x\leftarrow(P_n,Q_n)}[\operatorname{D}(x)=1]-\operatorname{Pr}_{x\leftarrow Q_n^2}[\operatorname{D}(x)=1]\right|$ $= \left| \Delta_{(\mathcal{P}^2,(\mathcal{P},\mathcal{O})}^{\mathsf{D}}(n) \right| + \left| \Delta_{((\mathcal{P},\mathcal{O}),\mathcal{O}^2)}^{\mathsf{D}}(n) \right|$

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Assume that \mathcal{P} and \mathcal{Q} are computationally indistinguishable, is it always true that $\mathcal{P}^2=(\mathcal{P},\mathcal{P})$ and $\mathcal{Q}^2=(\mathcal{Q},\mathcal{Q})$ are?

Assume that
$$\left|\Delta_{(\mathcal{P}^2,\mathcal{Q}^2)}^{D}(n)\right| = \delta(n)$$
 for some PPT D, we would like to prove that \exists PPT D' with $\left|\Delta_{(\mathcal{P},\mathcal{Q})}^{D}(n)\right| \geq \delta(n)/2$ for every $n \in \mathbb{N}$. Indeed
$$\delta(n) = \left|\Pr_{x \leftarrow P_n^2}[D(x) = 1] - \Pr_{x \leftarrow Q_n^2}[D(x) = 1]\right|$$

$$\leq \left|\Pr_{x \leftarrow P_n^2}[D(x) = 1] - \Pr_{x \leftarrow (P_n,Q_n)}[D(x) = 1]\right|$$

$$+ \left|\Pr_{x \leftarrow (P_n,Q_n)}[D(x) = 1] - \Pr_{x \leftarrow Q_n^2}[D(x) = 1]\right|$$

$$= \left|\Delta_{(\mathcal{P}^2,(\mathcal{P},\mathcal{Q})}^{D}(n)\right| + \left|\Delta_{((\mathcal{P},\mathcal{Q}),\mathcal{Q}^2)}^{D}(n)\right|$$
 So either $|\Delta_{(\mathcal{P}^2,\mathcal{Q})}^{D}(n)| \geq \delta(n)/2$, or $|\Delta_{((\mathcal{P},\mathcal{Q}),\mathcal{Q}^2)}^{D}(n)| \geq \delta/2$

• Assume that $\left|\Delta^{\mathbb{D}}_{(\mathcal{P}^2,\mathcal{Q}^2)}(n)\right| \geq 1/p(n)$ for some $p \in \mathsf{poly}$ and infinitely many n's, and assume wlg. that $\left|\Delta^{\mathbb{D}}_{\mathcal{P}^2,(\mathcal{P},\mathcal{Q})}(n)\right| \geq 1/2p(n)$ for infinitely many n's.

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- Can we use D to contradict the fact that P and Q are computationally close?
- ullet Assuming that ${\mathcal P}$ and ${\mathcal Q}$ are efficiently samplable
- Non-uniform settings

Repeated sampling cont.

Given
$$t = t(n) \in \mathbb{N}$$
 and a distribution ensemble $\mathcal{P} = \{P_n\}_{n \in \mathbb{N}}$, let $\mathcal{P}^t = \{P_n^{t(n)}\}_{n \in \mathbb{N}}$

Question 6

Let $t = t(n) \le \operatorname{poly}(n)$ be an eff. computable integer function. Assume that \mathcal{P} and \mathcal{Q} are eff. samplable and computationally indistinguishable, does it mean that \mathcal{P}^t and \mathcal{Q}^t are?

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Proof:

Induction?

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Proof:

- Induction?
- Hybrid

Hybrid argument

Let D be an algorithm, and for $n \in \mathbb{N}$ let

$$\delta(n) = \left| \Delta^{\mathsf{D}}_{(\mathcal{P}^{t(n)}, \mathcal{Q}^{t(n)})}(t(n)) \right|.$$

• For $i \in \{0, ..., t = t(n)\}$, let $H^i = (p_1, ..., p_i, q_{i+1}, ..., q_t)$, where the p's [resp., q's] are uniformly (and independently) chosen from P_n [resp., from Q_n].

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- Since $\delta(n) = \left| \Delta^{\mathsf{D}}_{H^n,H^0}(t) \right| = \left| \sum_{i \in [t]} \Delta^{\mathsf{D}}_{H^i,H^{i-1}}(t) \right|$, there exists $i \in [t]$ with $\left| \Delta^{\mathsf{D}}_{H^i,H^{i-1}}(t) \right| \geq \delta(n)/t(n)$

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How do we use it?

Using hybrid argument via estimation

Algorithm 7 (D')

- Find $i \in [t]$ with $\left| \Delta_{H^i, H^{i-1}}^{D}(t) \right| \geq \delta(n)/2t(n)$
- 2 Let $(p_1,\ldots,p_i,q_{i+1},\ldots,q_t) \leftarrow H^i$
- **3** Return $D(1^t, p_1, \dots, p_{i-1}, x, q_{i+1}, \dots, q_t)$,.

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- how do we find *i*?
- Easy in the non-uniform case

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- **2** Let $(p_1, ..., p_i, q_{i+1}, ..., q_t) \leftarrow H^i$
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$$\left|\Delta^{\mathsf{D}'}_{(\mathcal{P},\mathcal{Q})}(n)\right| \ = \ \left|\mathsf{Pr}_{\rho\leftarrow P_n}[\mathsf{D}'(\rho)=1] - \mathsf{Pr}_{q\leftarrow Q_n}[\mathsf{D}'(q)=1]\right|$$

Algorithm 8 (D')

- **1** Sample $i \leftarrow [t = t(n)]$
- 2 Let $(p_1, \ldots, p_i, q_{i+1}, \ldots, q_t) \leftarrow H^i$
- **3** Return $D(1^t, p_1, \dots, p_{i-1}, x, q_{i+1}, \dots, q_t)$.

$$\begin{aligned} \left| \Delta_{(\mathcal{P}, \mathcal{Q})}^{D'}(n) \right| &= \left| \mathsf{Pr}_{p \leftarrow P_n}[\mathsf{D}'(p) = 1] - \mathsf{Pr}_{q \leftarrow Q_n}[\mathsf{D}'(q) = 1] \right| \\ &= \left| \frac{1}{t} \sum_{i \in [t]} \mathsf{Pr}_{x \leftarrow H_i}[\mathsf{D}(x) = 1] - \frac{1}{t} \sum_{i \in [t]} \mathsf{Pr}_{x \leftarrow H_{i-1}}[\mathsf{D}(x) = 1] \right| \end{aligned}$$

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Using Hybrid argument via sampling

Algorithm 8 (D')

Input: 1^n and $x \in \{0, 1\}^*$

- Sample $i \leftarrow [t = t(n)]$
 - 2 Let $(p_1, \ldots, p_i, q_{i+1}, \ldots, q_t) \leftarrow H^i$
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Section 3

Pseudorandom Generators

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Definition 10 (pseudorandom generators (PRGs))

An efficiently computable function $g:\{0,1\}^n\mapsto\{0,1\}^{\ell(n)}$ is a pseudorandom generator, if

- g is length extending (i.e., $\ell(n) > n$ for any n)
- $g(U_n)$ is pseudorandom

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- Imply one-way functions (homework)

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PRGs from OWPs

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- Imply one-way functions (homework)
- Do they have any use?

Section 4

Hardcore Predicates

Building blocks in constructions of PRGS from OWF

Building blocks in constructions of PRGS from OWF

Definition 11 (hardcore predicates)

An efficiently computable function $b: \{0,1\}^n \mapsto \{0,1\}$ is a hardcore predicate of $f: \{0,1\}^n \mapsto \{0,1\}^n$, if

$$\Pr[P(f(U_n)) = b(U_n)] \le \frac{1}{2} + \operatorname{neg}(n),$$

for any PPT P.

Building blocks in constructions of PRGS from OWF

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 Does the existence of a hardcore predicate for f, implies that f is one way?

Building blocks in constructions of PRGS from OWF

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for any PPT P.

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- Fact: any PRG has HCP (homework).

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- Does the existence of a hardcore predicate for f, implies that f is one way? If f is injective?
- Fact: any PRG has HCP (homework).
- Fact: any OWF has a hardcore predicate (next class)

Section 5

PRGs from OWPs

PRGs from OWPs

OWP to PRG

Claim 12

Let $f: \{0,1\}^n \mapsto \{0,1\}^n$ be a permutation and let $b: \{0,1\}^n \mapsto \{0,1\}$ be a hardcore predicate for f, then g(x) = (f(x), b(x)) is a PRG.

Pseudorandom Generators

Claim 12

Let $f: \{0,1\}^n \mapsto \{0,1\}^n$ be a permutation and let $b: \{0,1\}^n \mapsto \{0,1\}$ be a hardcore predicate for f, then q(x) = (f(x), b(x)) is a PRG.

Proof: Assume \exists a PPT D, and infinite set $\mathcal{I} \subseteq \mathbb{N}$ and $p \in \text{poly}$ with

$$\left|\Delta_{g(U_n),U_{n+1}}^{\mathsf{D}}\right| > \varepsilon(n) = 1/p(n)$$

for any $n \in \mathcal{I}$. We use D for breaking the hardness of b.

Pseudorandom Generators

Claim 12

Let $f: \{0,1\}^n \mapsto \{0,1\}^n$ be a permutation and let $b: \{0,1\}^n \mapsto \{0,1\}$ be a hardcore predicate for f, then q(x) = (f(x), b(x)) is a PRG.

Proof: Assume \exists a PPT D, and infinite set $\mathcal{I} \subseteq \mathbb{N}$ and $p \in \text{poly}$ with

$$\left|\Delta_{g(U_n),U_{n+1}}^{\mathsf{D}}\right| > \varepsilon(n) = 1/p(n)$$

for any $n \in \mathcal{I}$. We use D for breaking the hardness of b.

 We assume wlg. that $\Pr[\mathsf{D}(g(U_n)) = 1] - \Pr[\mathsf{D}(U_{n+1}) = 1] \ge \varepsilon(n)$ for any $n \in \mathcal{I}$ (can we do it?), and fix $n \in \mathcal{I}$.

PRGs from OWPs

OWP to PRG cont.

• Let $\delta(n) = \Pr[D(U_{n+1}) = 1]$ (note that $Pr[D(g(U_n)) = 1] = \delta + \varepsilon$).

- Let $\delta(n) = \Pr[D(U_{n+1}) = 1]$ (note that $\Pr[\mathsf{D}(q(U_n))=1]=\delta+\varepsilon).$
- Compute

$$\delta = \Pr[D(f(U_n), U_1) = 1]$$

$$= \Pr[U_1 = b(U_n)] \cdot \Pr[D(f(U_n), U_1) = 1 \mid U_1 = b(U_n)]$$

$$+ \Pr[U_1 = \overline{b(U_n)}] \cdot \Pr[D(f(U_n), U_1) = 1 \mid U_1 = \overline{b(U_n)}]$$

PRGs from OWPs

Pseudorandom Generators

- Let $\delta(n) = \Pr[D(U_{n+1}) = 1]$ (note that $\Pr[\mathsf{D}(q(U_n))=1]=\delta+\varepsilon).$
 - Compute

$$\begin{array}{lll} \delta & = & \Pr[\mathsf{D}(f(U_n), U_1) = 1] \\ & = & \Pr[U_1 = b(U_n)] \cdot \Pr[\mathsf{D}(f(U_n), U_1) = 1 \mid U_1 = b(U_n)] \\ & + & \Pr[U_1 = \overline{b(U_n)}] \cdot \Pr[\mathsf{D}(f(U_n), U_1) = 1 \mid U_1 = \overline{b(U_n)}] \\ & = & \frac{1}{2}(\delta + \varepsilon) + \frac{1}{2} \cdot \Pr[\mathsf{D}(f(U_n), U_1) = 1 \mid U_1 = \overline{b(U_n)}]. \end{array}$$

- Let $\delta(n) = \Pr[D(U_{n+1}) = 1]$ (note that $\Pr[\mathsf{D}(q(U_n))=1]=\delta+\varepsilon).$
- Compute

$$\begin{array}{lll} \delta & = & \Pr[\mathsf{D}(f(U_n),U_1)=1] \\ & = & \Pr[U_1=b(U_n)] \cdot \Pr[\mathsf{D}(f(U_n),U_1)=1 \mid U_1=b(U_n)] \\ & + & \Pr[U_1=\overline{b(U_n)}] \cdot \Pr[\mathsf{D}(f(U_n),U_1)=1 \mid U_1=\overline{b(U_n)}] \\ & = & \frac{1}{2}(\delta+\varepsilon)+\frac{1}{2} \cdot \Pr[\mathsf{D}(f(U_n),U_1)=1 \mid U_1=\overline{b(U_n)}]. \end{array}$$

Hence.

$$\Pr[\mathsf{D}(f(U_n), \overline{b(U_n)}) = 1] = \delta - \varepsilon \tag{1}$$

- $Pr[D(f(U_n), b(U_n)) = 1] = \delta + \varepsilon$
- $Pr[D(f(U_n), \overline{b(U_n)}) = 1] = \delta \varepsilon$

Pseudorandom Generators

- $Pr[D(f(U_n), b(U_n)) = 1] = \delta + \varepsilon$
- $Pr[D(f(U_n), \overline{b(U_n)}) = 1] = \delta \varepsilon$
- Consider the following algorithm for predicting *b*:

Algorithm 13 (P)

Input: $y \in \{0, 1\}^n$

- Flip a random coin $c \leftarrow \{0, 1\}$.
- 2 If D(y, c) = 1 output c, otherwise, output \overline{c} .

PRGs from OWPs

OWP to PRG cont.

- $Pr[D(f(U_n), b(U_n)) = 1] = \delta + \varepsilon$
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 - It follows that

$$\Pr[P(f(U_n)) = b(U_n)] \\
= \Pr[c = b(U_n)] \cdot \Pr[D(f(U_n), c) = 1 \mid c = b(U_n)] \\
+ \Pr[c = \overline{b(U_n)}] \cdot \Pr[D(f(U_n), c) = 0 \mid c = \overline{b(U_n)}]$$

- $\Pr[\mathsf{D}(f(U_n),b(U_n))=1]=\delta+\varepsilon$
- $\Pr[D(f(U_n), \overline{b(U_n)}) = 1] = \delta \varepsilon$
- Consider the following algorithm for predicting b:

Algorithm 13 (P)

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- Flip a random coin $c \leftarrow \{0, 1\}$.
- 2 If D(y, c) = 1 output c, otherwise, output \overline{c} .
 - It follows that

$$\begin{aligned} & \Pr[\mathsf{P}(f(U_n)) = b(U_n)] \\ &= & \Pr[c = b(U_n)] \cdot \Pr[\mathsf{D}(f(U_n), c) = 1 \mid c = b(U_n)] \\ &+ \Pr[c = \overline{b(U_n)}] \cdot \Pr[\mathsf{D}(f(U_n), c) = 0 \mid c = \overline{b(U_n)}] \\ &= & \frac{1}{2} \cdot (\delta + \varepsilon) + \frac{1}{2} (1 - \delta + \varepsilon) = \frac{1}{2} + \varepsilon. \end{aligned}$$

Remark 14

Prediction to distinguishing (homework)

Remark 14

- Prediction to distinguishing (homework)
- PRG from any OWF: (1) Regular OWFs, first use pairwise hashing to convert into "almost" permutation. (2) Any OWF, harder

Section 6

PRG Length Extension

PRG Length Extension

Construction 15 (iterated function)

Given a length increasing function $g: \{0,1\}^n \mapsto \{0,1\}^\ell$ and $i \in \mathbb{N}$, define $g^i: \{0,1\}^n \mapsto \{0,1\}^{n+i(\ell-n)}$ as

$$g^{i}(x) = x_{n+1,\dots,|x^{i-1}|}^{i-1}, g(x_{1,\dots,n}^{i-1}),$$

where $x^{i-1} = g^{i-1}(x)$ and $g^0(x) = x$.

PRG Length Extension

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PRGs from OWPs

where $x^{i-1} = g^{i-1}(x)$ and $g^0(x) = x$.

Claim 16

Let $g: \{0,1\}^n \mapsto \{0,1\}^{n+1}$ be a PRG, then $a^t : \{0, 1\}^n \mapsto \{0, 1\}^{n+t(n)}$ is a PRG, for any $t \in \text{poly}$.

PRG Length Extension

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Proof: Assume \exists a PPT D, an infinite set $\mathcal{I} \subseteq \mathbb{N}$ and $p \in \text{poly}$ with

$$\left|\Delta_{g^t(U_n),U_{n+t(n)}}^{\mathsf{D}}\right| > \varepsilon(n) = 1/p(n),$$

for any $n \in \mathcal{I}$. We use D for breaking the hardness of g.

• Fix $n \in \mathbb{N}$, for $i \in \{0, \dots, t = t(n)\}$, let $H^i = U_{t-i}, g^i(U_n)$ (i.e., the distribution of H^i is $(x, g^i(x'))_{x \leftarrow \{0,1\}^{t-i}, x' \leftarrow \{0,1\}^n}$)

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- Note that $H^0 \equiv U_{n+t}$ and $H^t \equiv g^t(U_n)$.

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PRGs from OWPs

• Note that $H^0 \equiv U_{n+t}$ and $H^t \equiv g^t(U_n)$.

Algorithm 17 (D')

Input: 1^n and $y \in \{0, 1\}^{n+1}$

- \bigcirc Sample $i \leftarrow [t]$
- 2 Return D(1ⁿ, U_{t-i} , y_{n+1} , $g^{i-1}(y_{1,...,n})$).

• Fix $n \in \mathbb{N}$, for $i \in \{0, \dots, t = t(n)\}$, let $H^i = U_{t-i}, g^i(U_n)$ (i.e., the distribution of H^i is $(x, g^i(x'))_{x \leftarrow \{0,1\}^{t-i}, x' \leftarrow \{0,1\}^n}$)

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It holds that $\left|\Delta_{g(U_n),U_{n+1}}^{\mathsf{D}'}\right| > \varepsilon(n)/t(n)$

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Proof: ...