# Application of Information Theory, Lecture 2 Joint & Conditional Entropy, Mutual Information

#### **Handout Mode**

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# Part I

# **Joint and Conditional Entropy**

# Joint entropy

Recall that the entropy of rv X, is defined by

$$H(X) = -\sum_{x \in \mathcal{X}} P_X(x) \log P_X(x)$$

- Shorter notation: for  $X \sim p$ , let  $H(X) = -\sum_{x} p(x) \log p(x)$  (where the summation is over the domain of X).
- ► The joint entropy of (jointly distributed) rvs X and Y with  $(X, Y) \sim p$ , is

$$H(X,Y) = -\sum_{x,y} p(x,y) \log p(x,y)$$

This is simply the entropy of the rv Z = (X, Y).

Example:

X	0	1
0	1/4	1/4
1	1 2	0

$$H(X, Y) = -\frac{1}{2} \log \frac{1}{\frac{1}{2}} - \frac{1}{4} \log \frac{1}{\frac{1}{4}} - \frac{1}{4} \log \frac{1}{\frac{1}{4}}$$
$$= \frac{1}{2} \cdot 1 + \frac{1}{2} \cdot 2 = 1\frac{1}{2}$$

# Joint entropy, cont.

▶ The joint entropy of  $(X_1, ..., X_n) \sim p$ , is

$$H(X_1,\ldots,X_n)=-\sum_{x_1,\ldots,x_n}p(x_1,\ldots,x_n)\log p(x_1,\ldots,x_n)$$

# **Conditional entropy**

- Let (X, Y) ~ p.
- ▶ For  $x \in \text{Supp}(X)$ , the random variable Y|X = x is well defined.
- ► The entropy of Y conditioned on X, is defined by

$$H(Y|X) := \mathop{\mathsf{E}}_{x \leftarrow X} H(Y|X = x) = \mathop{\mathsf{E}}_{X} H(Y|X)$$

- ▶ Measures the uncertainty in *Y* given *X*.
- Let  $p_X \& p_{Y|X}$  be the marginal & conational distributions induced by p.

$$H(Y|X) = \sum_{x \in \mathcal{X}} p_X(x) \cdot H(Y|X = x)$$

$$= -\sum_{x \in \mathcal{X}} p_X(x) \sum_{y \in \mathcal{Y}} p_{Y|X}(y|x) \log p_{Y|X}(y|x)$$

$$= -\sum_{x \in \mathcal{X}, y \in \mathcal{Y}} p(x, y) \log p_{Y|X}(y|x)$$

$$= -\sum_{x \in \mathcal{X}, y \in \mathcal{Y}} p(x, y) \log p_{Y|X}(y|x)$$

$$= -\sum_{x \in \mathcal{X}, y \in \mathcal{Y}} \log p_{Y|X}(Y|X)$$

$$= -\sum_{z \in p_{Y|X}(Y|X)} \log z$$

# Conditional entropy, cont.

#### Example

XY	0	1
0	1/4	1/4
1	1 2	0

What is H(Y|X) and H(X|Y)?

$$H(Y|X) = \mathop{\mathbb{E}}_{x \leftarrow X} H(Y|X = x)$$

$$= \frac{1}{2} H(Y|X = 0) + \frac{1}{2} H(Y|X = 1)$$

$$= \frac{1}{2} H(\frac{1}{2}, \frac{1}{2}) + \frac{1}{2} H(1, 0) = \frac{1}{2}.$$

$$H(X|Y) = \mathop{\mathsf{E}}_{y \leftarrow Y} H(X|Y = y)$$

$$= \frac{3}{4} H(X|Y = 0) + \frac{1}{4} H(X|Y = 1)$$

$$= \frac{3}{4} H(\frac{1}{3}, \frac{2}{3}) + \frac{1}{4} H(1, 0) = 0.6887 \neq H(Y|X).$$

# Conditional entropy, cont..

$$H(X|Y,Z) = \mathop{\mathsf{E}}_{(y,z)\leftarrow(Y,Z)} H(X|Y=y,Z=z)$$

$$= \mathop{\mathsf{E}}_{y\leftarrow Y} \mathop{\mathsf{E}}_{z\leftarrow Z|Y=y} H(X|Y=y,Z=z)$$

$$= \mathop{\mathsf{E}}_{y\leftarrow Y} \mathop{\mathsf{E}}_{z\leftarrow Z|Y=y} H((X|Y=y)|Z=z)$$

Let 
$$(X_y, Z_y) = (X, Z)|Y = y$$
. Then

$$H(X|Y,Z) = \mathop{\mathsf{E}}_{y \leftarrow Y} \mathop{\mathsf{E}}_{z \leftarrow Z_y} H(X_y|Z = z)$$
$$= \mathop{\mathsf{E}}_{y \leftarrow Y} \mathop{\mathsf{E}}_{z \leftarrow Z_y} H(X_y|Z_y = z)$$
$$= \mathop{\mathsf{E}}_{y \leftarrow Y} H(X_y|Z_y)$$

# Relating mutual entropy to conditional entropy

- ▶ What is the relation between H(X), H(Y), H(X, Y) and H(Y|Y)?
- Intuitively, 0 ≤ H(Y|X) ≤ H(Y)
  Non-negativity is immediate. We prove upperbound later.
- ▶ We will also see that H(Y|X) = H(Y) iff X and Y are independent.
- ► In our example,  $H(Y) = H(\frac{3}{4}, \frac{1}{4}) > \frac{1}{2} = H(Y|X)$
- ▶ Note that H(Y|X = x) might be larger than H(Y) for some  $x \in \text{Supp}(X)$ .
- ► Chain rule (proved next). H(X, Y) = H(X) + H(Y|X)
- ► Intuitively, uncertainty in (X, Y) is the uncertainty in X plus the uncertainty in Y given X.
- ► H(Y|X) = H(X, Y) H(X) is as an alternative definition for H(Y|X).

# Chain rule (for the entropy function)

#### Claim 1

For rvs X, Y, it holds that H(X, Y) = H(X) + H(Y|X).

Proof immediately follow by the grouping axiom:

X			
	$P_{1,1}$		$P_{1,n}$
		:	:
	$P_{n,1}$		$P_{n,n}$

Let 
$$q_i = \sum_{j=1}^n p_{i,j}$$
 (=  $\Pr[X = i]$   
 $H(P_{1,1}, \dots, P_{n,n})$   
 $= H(q_1, \dots, q_n) + \sum_i q_i H(\frac{P_{i,1}}{q_i}, \dots, \frac{P_{i,n}}{q_i})$   
 $= H(X) + H(Y|X).$ 

- ▶ Another proof. Let  $(X, Y) \sim p$ .
- $\implies \log p(x,y) = \log p_X(x) + \log p_{Y|X}(y|x)$
- $\implies$  E log  $p(X, Y) = E \log p_X(X) + E \log p_{Y|X}(Y|X)$
- $\implies$  H(X, Y) = H(X) + H(Y|X).

$$H(Y|X) \leq H(Y)$$

Jensen inequality: for any concave function f, values  $t_1, \ldots, t_k$  and  $\lambda_1, \ldots, \lambda_k \in [0, 1]$  with  $\sum_i \lambda_i = 1$ , it holds that  $\sum_i \lambda_i f(t_i) \leq f(\sum_i \lambda_i t_i)$ . Let  $(X, Y) \sim p$ .

$$H(Y|X) = -\sum_{x,y} p(x,y) \log p_{Y|X}(y|X)$$

$$= \sum_{x,y} p(x,y) \log \frac{p_X(x)}{p(x,y)}$$

$$= \sum_{x,y} p_Y(y) \cdot \frac{p(x,y)}{p_Y(y)} \log \frac{p_X(x)}{p(x,y)}$$

$$= \sum_{y} p_Y(y) \sum_{x} \frac{p(x,y)}{p_Y(y)} \log \frac{p_X(x)}{p(x,y)}$$

$$\leq \sum_{y} p_Y(y) \log \sum_{x} \frac{p(x,y)}{p_Y(y)} \frac{p_X(x)}{p(x,y)}$$

$$= \sum_{y} p_Y(y) \log \frac{1}{p_Y(y)} = H(Y).$$

$$H(Y|X) \leq H(Y)$$
 cont.

▶ Assume X and Y are independent (i.e.,  $p(x, y) = p_X(x) \cdot p_Y(y)$  for any x, y)

$$\implies p_{Y|X} = p_Y$$

$$\implies H(Y|X) = H(Y)$$

- ▶ Is the converse also true: H(Y|X) = H(Y) implies X and Y are independent?
  - Yes, since  $\log$  is strictly concave in the range. Equality happens iff all  $t_i$  are the same,
- which happens iff  $p(x, y) = p_X(x)p_Y(y)$  for all x, y

## Other inequalities

- ► H(X),  $H(Y) \le H(X, Y) \le H(X) + H(Y)$ . Follows from H(X, Y) = H(X) + H(Y|X).
  - Left inequality since H(Y|X) is non negative.
  - ▶ Right inequality since  $H(Y|X) \le H(Y)$ .
- ► H(X, |Z) = H(X|Z) + H(Y|X, Z) (by chain rule)
- $\vdash$   $H(X|Y,Z) \leq H(X|Y)$

Proof:

$$H(X|Y,Z) = \underset{Z,Y}{\mathsf{E}} H(X \mid Y,Z)$$

$$= \underset{YZ|Y}{\mathsf{E}} E H(X \mid Y,Z)$$

$$= \underset{YZ|Y}{\mathsf{E}} E H((X \mid Y) \mid Z)$$

$$\leq \underset{YZ|Y}{\mathsf{E}} E H(X \mid Y)$$

$$= \underset{Y}{\mathsf{E}} H(X \mid Y)$$

$$= H(X|Y).$$

# Chain rule (for the entropy function), general case

#### Claim 2

For rvs  $X_1, \ldots, X_k$ , it holds that

$$H(X_1,\ldots,X_k) = H(X_i) + H(X_2|X_1) + \ldots + H(X_k|X_1,\ldots,X_{k-1}).$$

Proof: ?

- Extremely useful property!
- Analogously to the two variables case, it also holds that:
- $H(X_i) \leq H(X_1, \dots, X_k) \leq \sum_i H(X_i)$
- $H(X_1,\ldots,X_K|Y) \leq \sum_i H(X_i|Y)$

#### **Examples**

- ▶ (from last class) Let  $X_1, ..., X_n$  be Boolean iid with  $X_i \sim (\frac{1}{3}, \frac{2}{3})$ . Compute  $H(X_1, ..., X_n)$
- ► As above, but under the condition that  $\bigoplus_i X_i = 0$ ?
  - Via chain rule?
  - Via mapping?

## **Applications**

Let  $X_1, ..., X_n$  be Boolean iids with  $X_i \sim (p, 1-p)$  and let  $X = X_1, ..., X_n$ . Let f be such that  $\Pr[f(X) = z] = \Pr[f(X) = z']$ , for every  $k \in \mathbb{N}$  and  $z, z' \in \{0, 1\}^k$ . Let K = |f(X)|. Prove that  $E K < n \cdot h(p)$ .

 $\blacktriangleright$ 

$$n \cdot h(p) = H(X_1, \dots, X_n)$$

$$\geq H(f(X), K)$$

$$= H(K) + H(f(X) \mid K)$$

$$= H(K) + E K$$

$$\geq E K$$

- Interpretation
- Upper bounds

#### Applications cont.

- How many comparisons it takes to sort n elements?
  Let A be a sorter for n elements algorithm making t comparisons.
  What can we say about t?
- Let X be a uniform random permutation of [n] and let Y₁,..., Yt be the answers A gets when sorting X.
- ► X is determined by  $Y_1, ..., Y_t$ . Namely,  $X = f(Y_1, ..., Y_t)$  for some function f.
- $H(X) = \log n!$

$$H(X) = H(f(Y_1, ..., Y_n))$$

$$\leq H(Y_1, ..., Y_n)$$

$$\leq \sum_i H(Y_i)$$

$$= t$$

 $\implies t \ge \log n! = \Theta(n \log n)$ 

# **Concavity of entropy function**

Let  $p=(p_1,\ldots,p_n)$  and  $q=(q_1,\ldots,q_n)$  be two distributions, and for  $\lambda\in[0,1]$  consider the distribution  $\tau_\lambda=\lambda p+(1-\lambda)q$ . (i.e.,  $\tau_\lambda=(\lambda p_1+(1-\lambda)q_1,\ldots,\lambda p_n+(1-\lambda)q_n)$ .

#### Claim 3

$$H(\tau_{\lambda}) \ge \lambda H(p) + (1 - \lambda)H(q)$$

#### Proof:

- Let Y over  $\{0,1\}$  be 1 wp  $\lambda$
- Let X be distributed according to p if Y = 0 and according to q otherwise.
- $H(\tau_{\lambda}) = H(X) \ge H(X \mid Y) = \lambda H(p) + (1 \lambda)H(q)$

We are now certain that we drew the graph of the (two-dimensional) entropy function right...

# Part II

# **Mutual Information**

#### **Mutual information**

► I(X; Y) — the "information" that X gives on Y

$$I(X; Y) := H(Y) - H(Y|X)$$

$$= H(Y) - (H(X, Y) - H(X))$$

$$= H(X) + H(Y) - H(X, Y)$$

$$= H(Y) - H(X|Y)$$

$$= I(Y; X).$$

- ► The mutual information that *X* gives about *Y* equals the mutual information that *Y* gives about *X*.
- $I(X;Y) \geq 0.$ 
  - When 0?

- I(X;X) = H(X)
- ▶ I(X; f(X)) = H(f(X)) (and smaller then H(X) if f is non-injective)
- ▶  $I(X; Y, Z) \ge I(X; Y), I(X; Z)$  (since  $H(X \mid Y, Z) \le H(X \mid Y), H(X \mid Z)$ )
- I(X; Y|Z) := H(Y|Z) H(Y|X,Z)  $\geq 0$
- ► I(X; Y|Z) = I(Y; X|Z) (since I(X'; Y') = I(Y'; X'))

# **Numerical example**

#### Example

X	0	1
0	1 4	$\frac{1}{4}$
1	1 2	0

$$I(X; Y) = H(X) - H(X|Y)$$

$$= 1 - \frac{3}{4} \cdot h(\frac{1}{3})$$

$$= I(Y; X)$$

$$= H(Y) - H(Y|X)$$

$$= h(\frac{1}{4}) - \frac{1}{2}h(\frac{1}{2})$$

#### Chain rule for mutual information

#### Claim 4 (Chain rule for mutual information)

For rvs  $X_1, ..., X_k, Y$ , it holds that  $I(X_1, ..., X_k; Y) = I(X; Y) + I(X_2; Y|X_1) + ... + I(X_k; Y|X_1, ..., X_{k-1})$ .

Proof: ? HW

#### **Examples**

Let  $X_1, \ldots, X_n$  be iid with  $X_i \sim (p, 1 - p)$ , under the condition that  $\bigoplus_i x_i = 0$ .

Compute  $I(X_1, ..., X_{n-1}; X_n)$ .

By chain rule

$$I(X_1,...,X_{n-1};X_n)$$
=  $H(X_1;X_n) + H(X_2;X_n|X_1) + ... + H(X_{n-1};X_n|X_1,...,X_{n-2})$   
=  $0 + 0 + ... + 1 = 1$ .

▶ Let T and F be the top and front side, respectively, of a 6-sided fair dice. Compute I(T; F).

$$I(T; F) = H(T) - H(T|F)$$
  
= log 6 - log 4  
= log 3 - 1.

# Part III

# **Data processing**

# **Data processing Inequality**

#### **Definition 5 (Markov Chain)**

Rvs  $(X, Y, Z) \sim p$  form a Markov chain, denoted  $X \to Y \to Z$ , if  $p(x, y, z) = p_X(x) \cdot p_{Y|X}(y|x) \cdot p_{Z|Y}(z|y)$ , for all x, y, z.

Example: random walk on graph.

#### Claim 6

If 
$$X \to Y \to Z$$
, then  $I(X; Y) \ge I(X; Z)$ .

- ▶ By Chain rule, I(X; Y, Z) = I(X; Z) + I(X; Y|Z) = I(X; Y) + I(X; Z|Y).
- I(X;Z|Y) = 0

  - I(X; Z|Y) = H(Z|Y) H(Z|Y, X)  $= \mathop{\mathsf{E}}_{Y} H(p_{Z|Y=y}) \mathop{\mathsf{E}}_{Y,X} H(p_{Z|Y=y,X=x})$   $= \mathop{\mathsf{E}}_{Y} H(p_{Z|Y=y}) \mathop{\mathsf{E}}_{Y} H(p_{Z|Y=y}) = 0.$
- ▶ Since  $I(X; Y|Z) \ge 0$ , we conclude  $I(X; Y) \ge I(X; Z)$ .

# Fano's Inequality

- ► How well can we guess X from Y?
- ▶ Could with no error if H(X|Y) = 0. What if H(X|Y) is small?

## Theorem 7 (Fano's inequality)

For any rvs X and Y, and any (even random) g, it holds that

$$h(P_e) + P_e \log |\mathcal{X}| \ge H(X|\hat{X}) \ge H(X|Y)$$

for 
$$\hat{X} = g(Y)$$
 and  $P_e = \Pr \left[ \hat{X} \neq X \right]$ .

- Note that  $P_e = 0$  implies that H(X|Y) = 0
- ▶ The inequality can be weekend to  $1 + P_e \log |\mathcal{X}| \ge H(X|Y)$ ,
- ▶ Alternatively, to  $P_e \ge \frac{H(X|Y)-1}{\log|\mathcal{X}|}$
- ▶ Intuition for  $\propto \frac{1}{\log |\mathcal{X}|}$
- ▶ We call  $\hat{X}$  an estimator for X (from Y).

# **Proving Fano's inequality**

Let X and Y be rvs, let  $\hat{X} = g(Y)$  and  $P_e = \Pr \left[ \hat{X} \neq X \right]$ .

$$\blacktriangleright \text{ Let } E = \left\{ \begin{array}{ll} 1, & \hat{X} \neq X \\ 0, & \hat{X} = X. \end{array} \right.$$

$$H(E, X|\hat{X}) = H(X|\hat{X}) + \underbrace{H(E|X, \hat{X})}_{=0}$$

$$= \underbrace{H(E|\hat{X})}_{\leq H(E) = h(P_e)} + \underbrace{H(X|E, \hat{X})}_{\leq P_e \log_{|\mathcal{X}|}(?)}$$

- ▶ It follows that  $h(P_e) + P_e \log |\mathcal{X}| \ge H(X|\hat{X})$
- ► Since  $X \to Y \to \hat{X}$ , it holds that  $I(X; Y) \ge I(X; \hat{X})$  $\implies H(X|\hat{X}) \ge H(X|Y)$