Prediction and Linear Regression

Big Data y Machine Learning para Economía Aplicada

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Agenda

- Best Predictor
 - Statistical Properties
 - Numerical Properties
- 2 Calculating the OLS coefficients
 - Traditional Computation
 - Gradient Descent
- 3 Review

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Best Predictor

▶ In the population, the best predictor of Y given W

$$g(W) = E[Y|W] \tag{1}$$

Best Predictor

► The CEF solves the best prediciton problem

$$min_{m(W)}E[(Y - m(W))^2] = \int (Y - m(W))^2 Pr(dW, dY)$$
 (2)

conditioning on W we have that

$$E_W E_{Y|X}[(Y - m(W))^2|W]$$
(3)

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The Best Linear Prediction Problem in Finite Samples

$$(Y_i, X_i)_{i=1}^n = ((Y_1, X_1), \dots, (Y_n, X_n))$$
 (4)

 \blacktriangleright We assume that this sample is a random sample from the distribution of (Y, X)

The Best Linear Prediction Problem in Finite Samples

ightharpoonup Replace *E* with \mathbb{E}

$$\sum_{j=1}^{k} \hat{\beta}_j X_j = \hat{\beta}' X \tag{5}$$

 $\hat{\beta}$ is any solution to the Best Linear Prediction Problem in the Sample, also known as Ordinary Least Squares (OLS)

Statistical Properties

Under certain assumptions HW Review the Assumption from Econometrics

- ► Small Sample (Gauss-Markov Theorem)
 - ▶ Unbiased: $E(\hat{\beta}) = \beta$
 - ▶ Minimum Variance: $Var(\tilde{\beta}) Var(\hat{\beta})$ is positive semidefinite matrix Proof: HW. Remember: a matrix $M_{p \times p}$ is positive semi-definite iff $c'Mc \ge 0 \ \forall c \in \mathbb{R}^p$
- ► Large Sample
 - ► Consistency: $\hat{\beta} \rightarrow_p \beta$
 - ► Asymptotically Normal: $\sqrt{N}(\hat{\beta} \beta) \sim_a N(0, S)$

Gauss Markov Theorem

- ► Gauss Markov Theorem that says OLS is BLUE is perhaps one of the most famous results in statistics.
 - \triangleright $E(\hat{\beta}) = \beta$
 - $ightharpoonup Var(\hat{\beta}) = \sigma^2(X'X)^{-1}$
- ightharpoonup and implies that \hat{y} is an unbiased predictor and minimum variance, from the class of unbiased linear predictors (BLUP) H.W. proof

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- ightharpoonup and implies that \hat{y} is an unbiased predictor and minimum variance, from the class of unbiased linear predictors (BLUP) H.W. proof
- ▶ However, it is essential to note the limitations of the theorem.
 - Correctly specified with exogenous Xs,
 - ► The term error is homoscedastic
 - ► No serial correlation.
 - ▶ Nothing about the OLS estimator being the more efficient than any other estimator one can imagine.



Statistical Properties

- ightharpoonup The fundamental statistical issue is that we are trying to estimate k parameters
- ► We need many observations per parameter
- ightharpoonup n/p should be large, or, equivalently that p/n should be small

Analysis of Variance

- ► Involves the decomposition of the variation of *Y* into explained and unexplained parts.
- ► Explained variation is a measure of the predictive performance of a model.
- ► Can be conducted both in the population and in the sample.

Analysis of Variance

$$Y = \beta' X + \epsilon$$

$$E[\epsilon \mid X] = 0$$
,

Idea: decompose the variation in *Y* into the sum of explained variation and residual variation.

$$E[Y^2] = E[(\beta'X)^2] + E[\epsilon^2]$$

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Numerical Properties

- ▶ Numerical properties have nothing to do with how the data was generated
- ► These properties hold for every data set, just because of the way that $\hat{\beta}$ was calculated
- ▶ Davidson & MacKinnon, Greene y Ruud have nice geometric interpretations

Projection

OLS Residuals:

$$\hat{\epsilon} = y - X(X'X)^{-1}X'y$$

$$\hat{\epsilon} = y - \hat{y}$$
$$= y - X\hat{\beta}$$

$$\hat{eta}$$

(6)

(8)

(9)

replacing $\hat{\beta}$

$$\hat{\epsilon}$$
 =

$$= (I - X(X'X)^{-1}X')y$$

Define two matrices

- ▶ Projection matrix $P_X = X(X'X)^{-1}X'$
- Annihilator (residual maker) matrix $M_X = (I P_X)$

Projection

- ► Both are symmetric
- ▶ Both are idempotent (A'A) = A
- $ightharpoonup P_X X = X \Rightarrow \text{projection matrix}$
- ► $M_X X = 0 \Rightarrow$ annihilator matrix

Frisch-Waugh-Lovell (FWL) Theorem

- ▶ Lineal Model: $Y = X\beta + u$
- ► Split it: $Y = X_1\beta_1 + X_2\beta_2 + u$

Theorem

11 The OLS estimates of β_2 from these equations

$$y = X_1 \beta_1 + X_2 \beta_2 + u \tag{10}$$

$$M_{X_1}y = M_{X_1}X_2\beta_2 + residuals \tag{11}$$

are numerically identical

2 the OLS residuals from these regressions are also numerically identical



Applications

- ▶ Why FWL is useful in the context of big volume of data?
- ► An computationally inexpensive way of
 - Removing nuisance parameters
 - ► E.g. the case of multiple fixed effects. The traditional way is either apply the within transformation with respect to the FE with more categories then add one dummy for each category for all the subsequent FE
 - Not feasible in certain instances.
 - ightharpoonup Computing certain diagnostic statistics: Leverage, R^2 , LOOCV.
 - Helps with online updating

Applications: Fixed Effects

► For example: Carneiro, Guimarães, & Portugal (2012) AEJ: Macroeconomics

$$ln w_{ijft} = x_{it}\beta + \lambda_i + \theta_j + \gamma_f + u_{ijft}$$
(12)

$$Y = X\beta + D_1\lambda + D_2\theta + D_3\gamma + u \tag{13}$$

- ▶ Data set 31.6 million observations, with 6.4 million individuals (i), 624 thousand firms (f), and 115 thousand occupations (j), 11 years (t).
- ▶ Storing the required indicator matrices would require 23.4 terabytes of memory
- From their paper
 "In our application, we first make use of the Frisch-Waugh-Lovell theorem to remove the influence of the
 three high- dimensional fixed effects from each individual variable, and, in a second step, implement the
 final regression using the transformed variables. With a correction to the degrees of freedom, this approach
 yields the exact least squares solution for the coefficients and standard errors"

Applications: Outliers and High Leverage Data

► App

$$\hat{\beta} = (X'X)^{-1}X'y \tag{14}$$

Applications: Outliers and High Leverage Data

Consider a dummy variable e_j which is an n-vector with element j equal to 1 and the rest is 0. Include it as a regressor

$$y = X\beta + \alpha e_j + u \tag{15}$$

using FWL we can do

$$M_{e_j} y = M_{e_j} X \beta + r \tag{16}$$

- \triangleright β and *residuals* from both regressions are identical
- Same estimates as those that would be obtained if we deleted observation j from the sample. We are going to denote this as $\beta^{(j)}$

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Linear Regression

▶ Using matrix algebra, the loss function:

$$\tilde{\epsilon}'\tilde{\epsilon} = (y - X\tilde{\beta})'(y - X\tilde{\beta}) \tag{17}$$

- ► $SSR(\tilde{\beta})$ is the aggregation of squared errors if we choose $\tilde{\beta}$ as an estimator.
- ► The **least squares estimator** $\hat{\beta}$ will be

$$\hat{\beta} = \underset{\tilde{\beta}}{\operatorname{argmin}} SSR(\tilde{\beta}) \tag{18}$$

QR decomposition

- \blacktriangleright To avoid inverting X'X we can use matrix decomposition: QR decomposition
- ► Most software use it

Theorem If $A \in \mathbb{R}^{n \times k}$ then there exists an orthogonal $Q \in \mathbb{R}^{n \times k}$ and an upper triangular $R \in \mathbb{R}^{k \times k}$ so that A = OR

- Orthogonal Matrices:

 - ▶ Def: Q'Q = QQ' = I and $Q' = Q^{-1}$ ▶ Prop: product of orthogonal is orthogonal, e.g A'A = I and B'B = I then (AB)'(AB) = B'(A'A)B = B'B = I
- ▶ (Thin OR) If $A \in \mathbb{R}^{n \times k}$ has full column rank then $A = O_1 R_1$ the OR factorization is unique, where $O_1 \in \mathbb{R}^{n \times k}$ and R is upper triangular with positive diagonal entries

QR decomposition

$$X = \begin{bmatrix} 1 & 2 \\ 1 & 2 \\ 1 & 3 \end{bmatrix} y = \begin{pmatrix} 1 \\ 4 \\ 2 \end{pmatrix} \tag{19}$$

1. QR factorization X=QR

$$Q = \begin{bmatrix} -0.57 & -0.41 \\ -0.57 & -0.41 \\ -0.57 & 0.82 \end{bmatrix} R = \begin{bmatrix} -1.73 & -4.04 \\ 0 & 0.81 \end{bmatrix}$$
 (20)

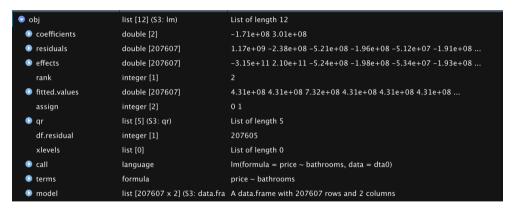
- 2. Calculate Q'y = [-4.04, -0.41]'
- 3. Solve

$$\begin{bmatrix} -1.73 & -4.04 \\ 0 & 0.81 \end{bmatrix} \begin{bmatrix} \beta_1 \\ \beta_2 \end{bmatrix} = \begin{bmatrix} -4.04 \\ -0.41 \end{bmatrix}$$
 (21)

Solution is (3.5, -0.5)

QR decomposition

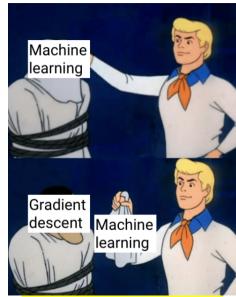
This is actually what R does under the hood



Note that R's 1m also returns many objects that have the same size as X and y

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Prediction and Linear Regression

- ► Gradient Descent is a very generic optimization algorithm capable of finding optimal solutions to a wide range of problems.
- ► The general idea of Gradient Descent is to tweak parameters iteratively in order to minimize a loss function.

$$\min_{f} E[L(y_i, f(X_i))] \tag{22}$$

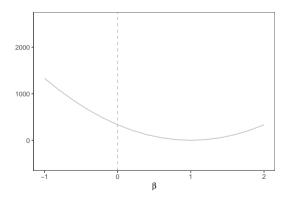
Linear regression

▶ The problem boils down to estimating the coefficients of vector β which minimize an objective function:

$$\arg\min_{\beta} \sum_{i=1}^{n} \frac{1}{n} \left(y_i - \beta_0 + \sum_{k=1}^{K} X_k \beta_k \right)^2$$
 (23)

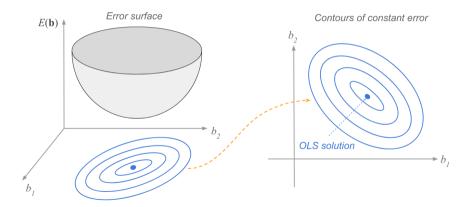
Linear regression

► Intuition: Loss Function 1 dimension



Linear regression

► Intuition: Loss Function 2 dimensions



▶ In a more general context, when at a point $\beta \in \mathbb{R}^k$, at any step j, the gradient descent algorithm tries to move in a direction $\delta\beta$ such that:

$$L(\beta^{(j)} + \delta\beta) < L(\beta^{(j)}) \tag{24}$$

► The choice of $\delta \beta$ is made such that $\delta \beta = -\epsilon \nabla_{\beta} L(\beta^{(j)})$:

$$\beta^{(t+1)} = \beta^{(j)} - \epsilon \nabla_{\beta} L(\beta^{(j)}) \tag{25}$$

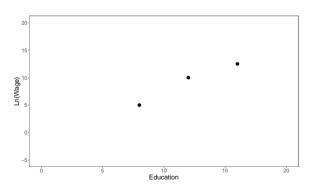
▶ In other words, you need to calculate how much the cost function will change if you change β just a little bit.

Gradient Descent

Algorithm

- 1 Randomly pick starting values for the parameters
- 2 Compute the gradient of the objective function at the current value of the parameters using all the observations from the training sample
- 3 Update the parameters
- 4 Repeat from step 2 until a fixed number of iteration or until convergence.

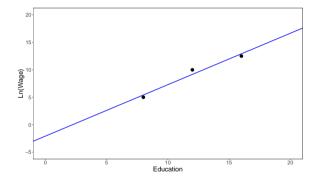
log(wage)	Education (years)		
5	8		
10	12		
12.5	16		



Education (years)		
8		
12		
16		

$$\hat{\beta} = (X'X)^{-1}X'y$$

$$y = -2.0833 + 0.9375 \times Educ$$



$$R(\alpha, \beta) = \frac{1}{n} \sum_{i=1}^{n} (y_i - (\alpha + \beta x_i))^2$$

The Gradient

$$\nabla R(\alpha, \beta) = \begin{pmatrix} \frac{\partial R}{\partial \alpha} \\ \frac{\partial R}{\partial \beta} \end{pmatrix} = \begin{pmatrix} -\frac{2}{n} \sum_{i=1}^{n} (y_i - \alpha - \beta x_i) \\ -\frac{2}{n} \sum_{i=1}^{n} x_i (y_i - \alpha - \beta x_i) \end{pmatrix}$$

Updating

$$\alpha^{(j+1)} = \alpha^{(j)} - \epsilon \frac{\partial R}{\partial \alpha}$$

$$\beta^{(j+1)} = \beta^{(j)} - \epsilon \frac{\partial R}{\partial \beta}$$

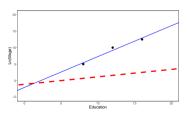


First Iteration

log(wage)	Education (years)
5	8
10	12
12.5	16

Start with an initial guess: $\alpha=-1; \beta=2$, and a learning rate ($\epsilon=0.005$). Then we have

$$\alpha^1 = -1.1384$$
 $\beta^1 = 0.2266$

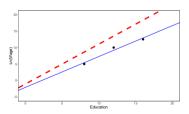


Second Iteration

log(wage)	Education (years)
5	8
10	12
12.5	16

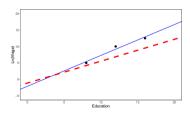
Start with an initial guess: $\alpha=-1; \beta=2$, and a learning rate ($\epsilon=0.005$). Then we have

$$\alpha^2 = -1.0624$$
$$\beta^2 = 1.212689$$



Third Iteration

$$\alpha^3 = -1.0624$$
$$\beta^3 = 1.212689$$

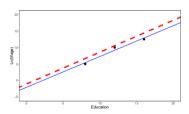


Fourth Iteration

log(wage)	Education (years)
5	8
10	12
12.5	16

$$\alpha^4 = -1.082738$$

$$\beta^4 = 0.9693922$$



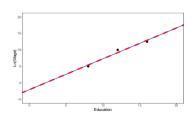
7211 Iteration

log(wage)	Education (years)
5	8
10	12
12.5	16

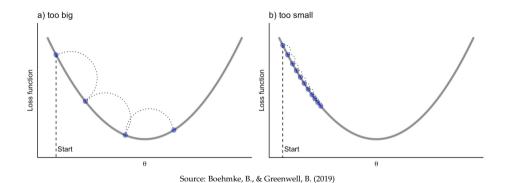
$$\alpha^{7211} = -2.076246$$

 $\beta^{7211} = 0.9369499$

$$y^{ols} = -2.0833 + 0.9375 \times Educ$$



The learning rate



- We can choose ϵ in several different ways:
 - \blacktriangleright Set ϵ to a small constant.
 - Use varying learning rates.

Table 1: Comparison of algorithms for Linear Regression

Algorithm	Large k	Large n	Hyperparams
QR	Fast	Slow	0
SVD	Fast	Slow	0
Batch GD	Slow	Fast	2

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Review

- ► These two Weeks: The predictive paradigm and linear regression
 - ightharpoonup BLP: E(y|X)
 - ► Inner workings of linear regression
- ► Next Module: Uncertainty