# Agglomeration Economies Urban Economics

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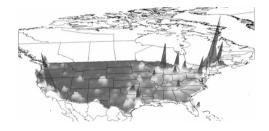
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# Agenda

- 1 Motivation
- 2 Duranton and Overmar
- 3 Aside: A Primer on Density Estimation
  - Parametric vs nonparametric methods
  - Parametric Methods
  - Non-parametric methods
    - Non-parametric methods: The Histogram
    - Non-parametric methods: the naive estimator
    - Non-parametric methods: Rosenblatt-Parzen density estimation approach
- 4 Duranton and Overman: Again

## **Agglomeration Economies**

▶ Why do we see such a remarkable clustering of human activity in a small number of urban areas?



## **Evidence of Agglomeration Economies**

- ► Three strategies to identify agglomeration economies
  - 1 Show there is too much spatial concentration to be random (Duranton and Overman, 2005)
  - 2 Compare productivity over space (Greenstone et al., 2010)
  - 3 Compare wages and rents across space (Quantitative Spatial Models, Ahlfeldt et al, 2015)

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## **Spatial Concentration**

Extremes of Localization and Dispersion



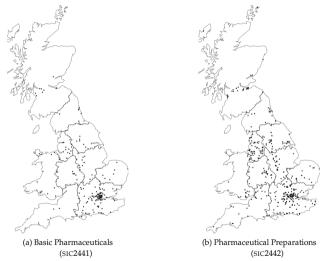
(c) Other Agricultural and Forestry Machinery (SIC2932)



(d) Machinery for Textile, Apparel and Leather Production (SIC2954)

## **Spatial Concentration**

## **Ambiguous Cases**



# Duranton & Overman Methodology

#### **1** Select Relevant Establishments:

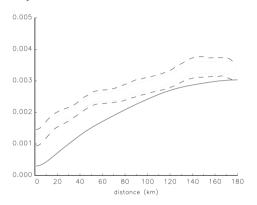
- Choose establishments based on industry and size.
- ► Consider different thresholds to assess robustness (e.g., include only those contributing to 90% of employment).

#### 2 Compute Bilateral Distances:

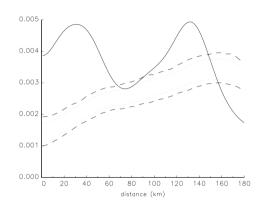
- Calculate Euclidean distances between all pairs of establishments.
- ▶ Use Kernel Density Estimation (KDE) to estimate the density of these distances.

## **Spatial Concentration**

#### **K** Density Estimates



(c) Other Agricultural and Forestry Machinery (SIC2932)

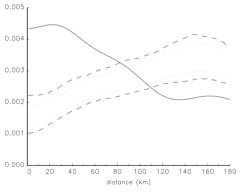


(d) Machinery for Textile, Apparel and Leather Production (SIC2954)

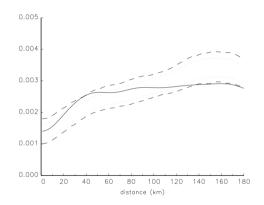
Figure 2. K-density, local confidence intervals and global confidence bands for four illustrative industries

## **Spatial Concentration**

#### **Ambiguous Cases**



(a) Basic Pharmaceuticals (SIC2441)



(b) Pharmaceutical Preparations (SIC2442)

# Duranton & Overman Methodology

#### **4** Generate Counterfactuals:

- Randomly assign establishments to locations, maintaining the number of establishments and industrial concentration.
- ► Create 1,000 simulations to construct a baseline for comparison.

### 5 Statistical Significance:

- Compare actual densities with simulated counterfactuals to determine if localization is significant.
- Use local and global confidence intervals to assess statistical significance of localization.

#### **6** Localization Metrics:

- **Define indices for localization** ( $\gamma$ ) and dispersion ( $\psi$ ) at each distance.
- ▶ Determine global localization or dispersion based on these indices over all distances.

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# Univariate density estimation: parametric vs nonparametric methods

- Let f = f(.) be the density function of the random variable X
- ▶ Let  $x_1$ ,  $x_2$ , . . .  $x_n$  be a random sample of X. Then,  $x_i \sim f$  iid
- ▶ How can we estimate the density in a particular point  $x_0$ , then  $f(x_0)$ ?

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## Parametric methods

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- Example:

$$f(x_o) = \phi(x_o) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{1}{2} \left(\frac{x_o - \mu}{\sigma}\right)^2\right]$$
 (1)

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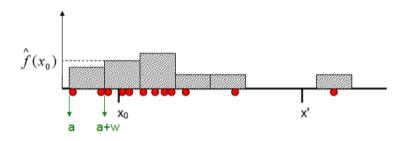
## Non-parametric methods

▶ They seek to estimate  $f(x_o)$  without assuming a particular functional form, only assuming certain regularity conditions of the density (smoothness, differentiability)



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## A rudimentary non-parametric estimator: the histogram



The height of the bars is an estimator of the density at any point in the interval

$$\hat{f}(x_o) = \frac{Nber. obs interval}{n \times w}$$
 (2)

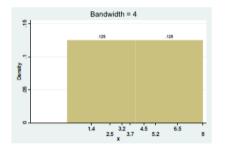
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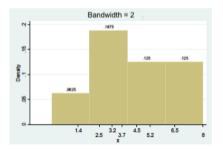
# Histogram problems: (1) depends on starting point



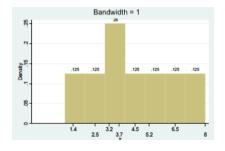


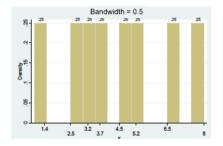
# Histogram problems: (2) depends on the bandwidth





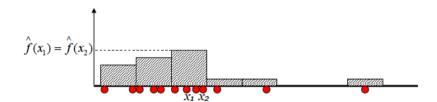
# Histogram problems: (2) depends on the bandwidth





# Histogram problems: (3) is discontinuous at the ends of the interval

- Note that  $\hat{f}(x_1) = \hat{f}(x_2)$
- ▶ But  $\hat{f}(x_2 + \epsilon) = \frac{1}{4}\hat{f}(x_2)$  for any  $\epsilon > 0$



# Another non-parametric estimator: the naive estimator

Discrete case

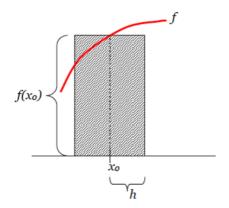
- X is a discrete RV, iid
- ► Objective: estimate  $Pr(X = x_o) = f(x_o)$
- ► The naive estimator comes from:

$$\hat{f}(x_o) = \frac{\# x_i = x_o}{n} = \frac{1}{n} \sum_{i=1}^n I(x_i = x_o)$$
(3)

## Another non-parametric estimator: the naive estimator

#### Continuous case

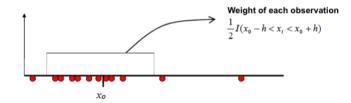
- ▶ X is a continuous RV ( $Pr(X = x_0)$ ) we evaluate the probability that X is "close" to  $x_0$ .
- ▶ We say that  $x_i$  is close to  $x_0$  if  $x_i$  belongs to the interval  $(x_0 h, x_0 + h)$



### The naive estimator

#### Notice:

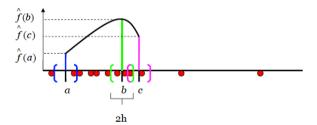
- $\triangleright$  Every possible point  $x_0$  is the center of an interval
- ▶ Observations that are within that interval (less than one h away from  $x_0$ ) are weighted by 1/2



## The naive estimator

#### Notice:

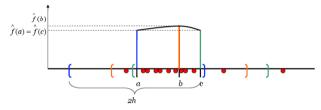
- ▶ To estimate the density at points a, b and c we construct intervals around them
- ► Then, unlike what happens in the histograms, the intervals of the naive estimator overlap (the initial point no longer matters)



#### The naive estimator

#### **Problems**

▶ It depends on h: the larger the bandwidth, the more distant observations from  $x_o$  are used to estimate  $f(x_o)$ . The higher h, the smoother the estimated density (h=smoothing parameter)



- ▶ The weight 1/2I(.) is discontinuous at the limits of each interval, generating discontinuities in the estimated density.
- ▶ The weight treats observations very close to  $x_0$  in the same way as others somewhat further away, as long as they belong to the interval of length 2h around  $x_0$

- ► The kernel estimator is a generalization of the naive estimator that overcomes some of the deficiencies of the latter.
- ▶ The weight 1/2I(.) of the naive estimator is replaced by a new weight K(.):

$$\hat{f}(x_o) = \frac{1}{n \times h} \sum_{i=1}^n K\left(\frac{x_i - x_0}{h}\right) \tag{4}$$

Rosenblatt-Parzen density estimation approach

- ▶ The weights or kernels are known functions that satisfy:
  - $K(\phi) \geq 0$
  - $\int_{-\infty}^{\infty} K(\phi) d\phi = 1$
  - $(\phi) = K(-\phi)$
  - **4**  $E(K(\phi)) = 0$

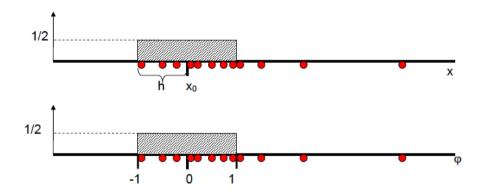
Example 1: rectangular or uniform kernel

$$K(\phi) = \begin{cases} \frac{1}{2} & |\phi| < 1\\ 0 & o.w. \end{cases}$$
 (5)

where  $\phi = \frac{x_i - x_o}{h}$  In this case, the Kernels estimator matches the naive estimator. Note that the weight is similar to a uniform density function on the interval (-1, 1)

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Example 1: rectangular or uniform kernel



#### Example 2: Gaussian kernel

► Gaussian

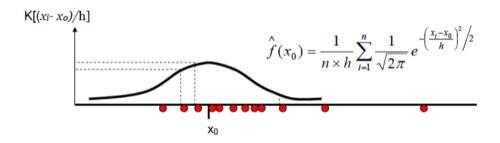
$$K(\phi) = \frac{1}{\sqrt{2}\pi} e^{-\phi^2/2}$$
 (6)

- ► The kernel function corresponds to the standard normal density function that satisfies the above assumptions.
- ► For this type of kernel, the density estimator is given by:

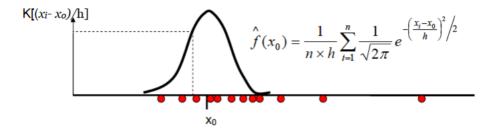
$$\hat{f}(x_o) = \frac{1}{n \times h} \sum_{i=1}^{n} \frac{1}{\sqrt{2}\pi} e^{-\left(\frac{x_i - x_o}{h}\right)^2/2}$$
 (7)

► Important: we don't assume normal distribution for any variable. The functional form is just used to weight the sample observations

Example 2: Gaussian kernel



Example 2: Gaussian kernel (smaller h)



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#### Other Kernels

Kernel type	Formula	Support
Gaussian or normal	$\frac{1}{\sqrt{2\pi}}\exp(-\varphi^2/2)$	$\varphi \in (-\infty, \infty)$
Epanechnikov	$\frac{3}{4\sqrt{5}}(1-\frac{1}{5}\varphi^2)I( \varphi <\sqrt{5})$	$\varphi \in \left(-\sqrt{5}, \sqrt{5}\right)$
Epanechnikov modificado	$\frac{3}{4}(1-\varphi^2)I( \varphi <1)$	<i>φ</i> ∈ (-1, 1)
Triangular	$(1- \varphi )I( \varphi <1)$	$arphi \in$ (-1, 1)
Uniform or rectangular	$\frac{1}{2}I( \varphi <1)$	<i>φ</i> ∈ (-1, 1)

Properties: Bias

- ► The kernel estimator is generally biased.
- ► The approximate expression for the bias is given by:

$$bias[\hat{f}(x_o)] \approx \frac{h^2}{2} f''(x_o) \int_{-\infty}^{\infty} K(\phi) \phi^2 d\phi$$
 (8)

► The approximate expression for the asymptotic variance of the kernel estimator is given by:

$$variance[\hat{f}(x_o)] \approx \frac{1}{n \times h} f(x_o) \int_{-\infty}^{\infty} K^2(\phi) d\phi$$
 (9)

Bandwidth Selection

$$MSE(\hat{f}(x)) = Var(\hat{f}(x)) + \left(bias(\hat{f}(x))\right)^{2}$$
(10)

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► We obtain a bandwidth which globally balances bias and variance by minimizing MSE with respect to h, i.e.,

$$h_{opt} = \left(\frac{\int K^2(z)dz}{(\int z^2 K(z)dz)^2 \int f'(x)^2 dx}\right)^{-1/5} n^{-1/5}$$
(11)

► There are two popular approaches to bandwidth selection,

1 Rules-of-thumb,

Bandwidth Selection

# The weighted average estimator or kernel method Rule-of-Thumb

- ▶ The rule-of-thumb for choosing the bandwidth makes assumtions about *f* and *K*
- ► For example: under a gaussian density and kernel

$$h_{opt} = 1.059\sigma n^{-1/5} \tag{12}$$

Bandwidth Selection

- ▶ There are two popular approaches to bandwidth selection,
  - 1 Rules-of-thumb,
  - 2 cross-validation methods,

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Extremes of Localization and Dispersion



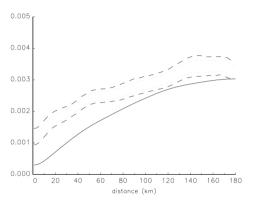
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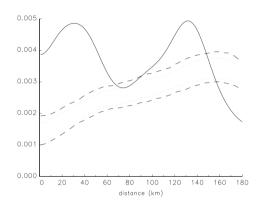
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#### K Density Estimates



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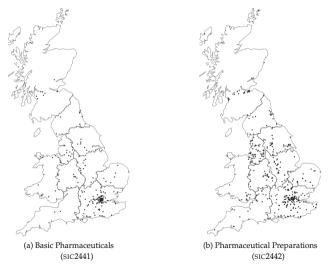


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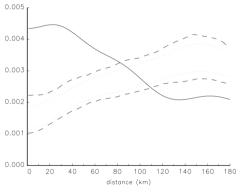
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Figure 2. K-density, local confidence intervals and global confidence bands for four illustrative industries

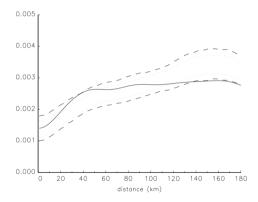
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