

Density Estimation & Nonparametric Regression

Urban Economics

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Agenda

- ① Recap: Univariate Density Estimation
- ② Kernel Density
- ③ Regressions: Non parametric estimation of conditional expectations
 - Local Constant Kernel Estimation
 - Local Linear Kernel Estimation
- ④ Further Readings

① Recap: Univariate Density Estimation

② Kernel Density

③ Regressions: Non parametric estimation of conditional expectations

- Local Constant Kernel Estimation
- Local Linear Kernel Estimation

④ Further Readings

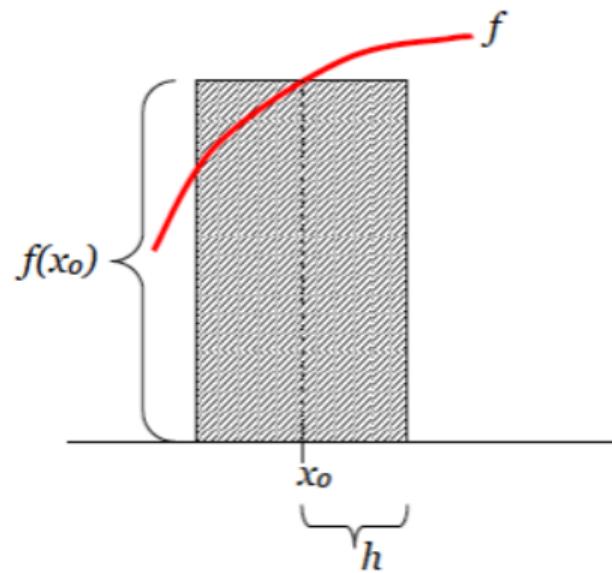
Univariate density estimation: parametric vs nonparametric methods

- ▶ Parametric Methods
- ▶ Non parametric
 - ▶ Histogram
 - ▶ Naive

Another non-parametric estimator: the naive estimator

Continuous case

- ▶ X is a continuous RV ($\Pr(X = x_0) = 0$) we evaluate the probability that X is "close" to x_0 .
- ▶ We say that x_i is close to x_0 if x_i belongs to the interval $(x_0 - h, x_0 + h)$



The naive estimator

► Then:

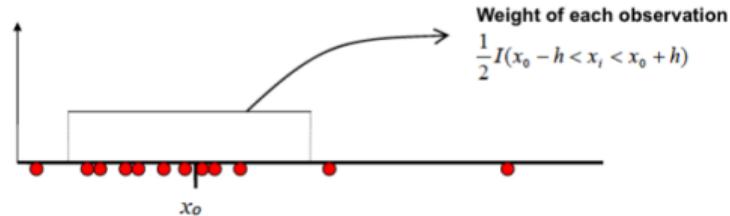
$$f(x_o) \cong \frac{1}{2h \times n} \sum_{i=1}^n I \left(-1 < \frac{x_i - x_o}{h} < 1 \right) \quad (1)$$

► $\hat{f}(x_o)$ is an estimator of the height of the rectangle in the previous graph

The naive estimator

Notice:

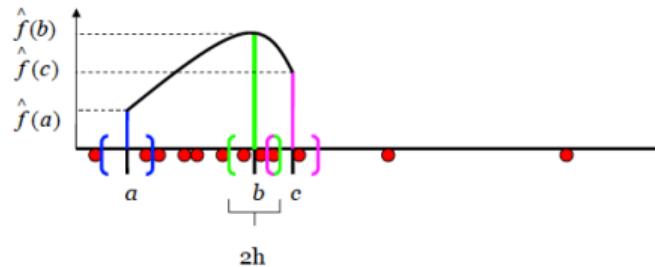
- ▶ Every possible point x_0 is the center of an interval
- ▶ Observations that are within that interval (less than one h away from x_0) are weighted by $1/2$



The naive estimator

Notice:

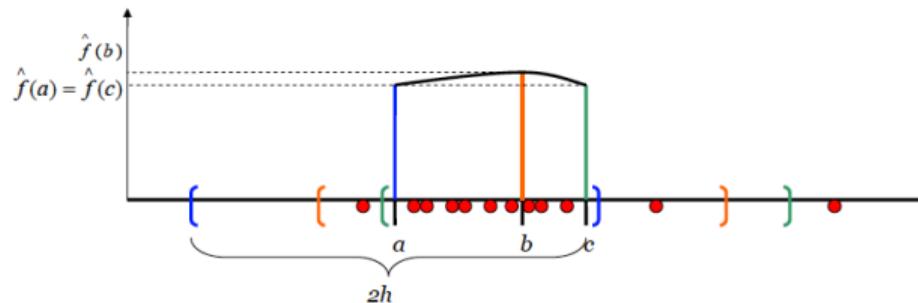
- ▶ To estimate the density at points a , b and c we construct intervals around each
- ▶ Then, unlike what happens in the histograms, the intervals of the naive estimator overlap



The naive estimator

Problems

- It depends on h : the larger the bandwidth, the more distant observations from x_0 are used to estimate $f(x_0)$. The higher h , the smoother the estimated density (h =smoothing parameter)



- The weight $1/2I(\cdot)$ is discontinuous at the limits of each interval, generating discontinuities in the estimated density.
- The weight treats observations very close to x_0 in the same way as others somewhat further away, as long as they belong to the interval of length $2h$ around x_0

① Recap: Univariate Density Estimation

② Kernel Density

③ Regressions: Non parametric estimation of conditional expectations

- Local Constant Kernel Estimation
- Local Linear Kernel Estimation

④ Further Readings

The weighted average estimator or kernel method

- ▶ The kernel estimator is a generalization of the naive estimator that overcomes some of the deficiencies of the latter.
- ▶ The weight $1/2I(\cdot)$ of the naive estimator is replaced by a new weight $K(\cdot)$:

$$\hat{f}(x_0) = \frac{1}{n \times h} \sum_{i=1}^n K\left(\frac{x_i - x_0}{h}\right) \quad (2)$$

- ▶ Rosenblatt-Parzen density estimation approach

The weighted average estimator or kernel method

- The weights or kernels are known functions that satisfy:

$$1 \quad K(\phi) \geq 0$$

$$2 \quad \int_{-\infty}^{\infty} K(\phi) d\phi = 1$$

$$3 \quad K(\phi) = K(-\phi)$$

The weighted average estimator or kernel method

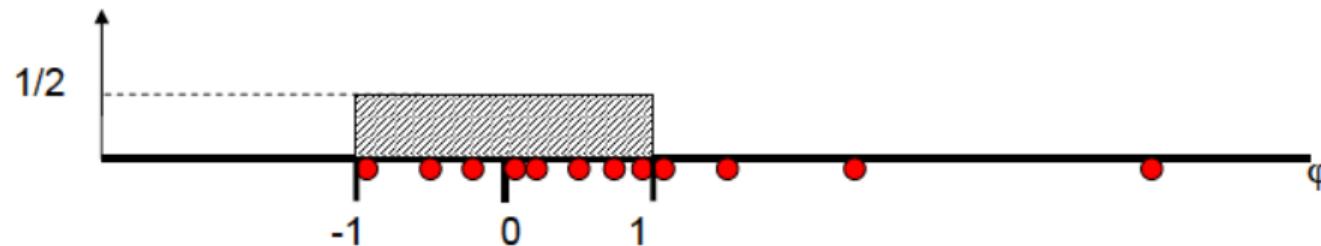
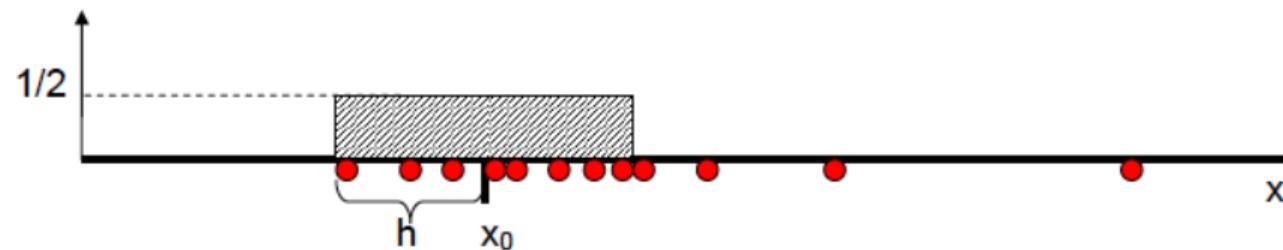
Example 1: rectangular or uniform kernel

$$K(\phi) = \begin{cases} \frac{1}{2} & |\phi| < 1 \\ 0 & o.w. \end{cases} \quad (3)$$

where $\phi = \frac{x_i - x_0}{h}$ In this case, the Kernels estimator matches the naive estimator. Note that the weight is similar to a uniform density function on the interval (-1, 1)

The weighted average estimator or kernel method

Example 1: rectangular or uniform kernel



The weighted average estimator or kernel method

Example 2: Gaussian kernel

- ▶ Gaussian

$$K(\phi) = \frac{1}{\sqrt{2\pi}} e^{-\phi^2/2} \quad (4)$$

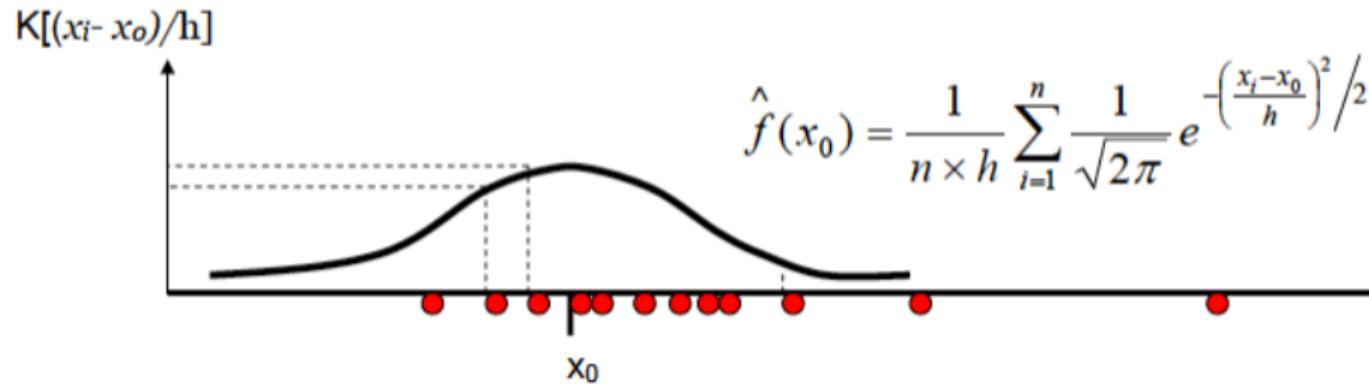
- ▶ The kernel function corresponds to the standard normal density function that satisfies the above assumptions.
- ▶ For this type of kernel, the density estimator is given by:

$$\hat{f}(x_o) = \frac{1}{n \times h} \sum_{i=1}^n \frac{1}{\sqrt{2\pi}} e^{-\left(\frac{x_i - x_o}{h}\right)^2/2} \quad (5)$$

- ▶ Important: we don't assume normal distribution for any variable. The functional form is just used to weight the sample observations

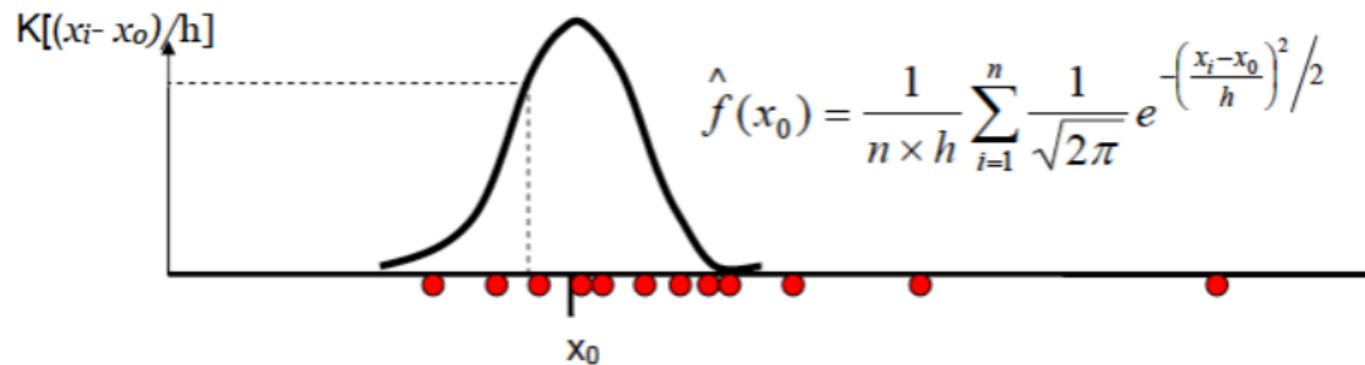
The weighted average estimator or kernel method

Example 2: Gaussian kernel



The weighted average estimator or kernel method

Example 2: Gaussian kernel (smaller h)



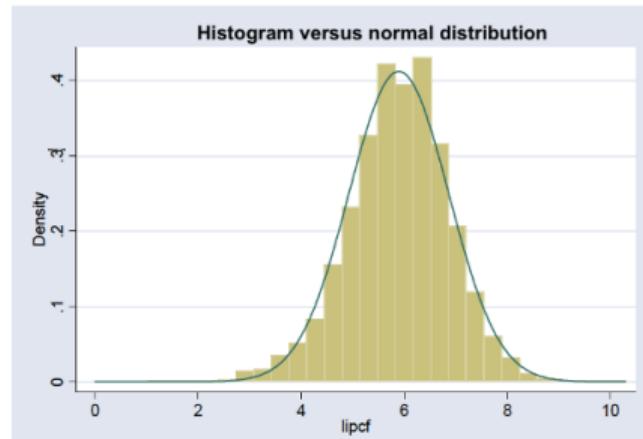
The weighted average estimator or kernel method

Other Kernels

Kernel type	Formula	Support
Gaussian or normal	$\frac{1}{\sqrt{2\pi}} \exp(-\varphi^2 / 2)$	$\varphi \in (-\infty, \infty)$
Epanechnikov	$\frac{3}{4\sqrt{5}} (1 - \frac{1}{5}\varphi^2) I(\varphi < \sqrt{5})$	$\varphi \in (-\sqrt{5}, \sqrt{5})$
Epanechnikov modificado	$\frac{3}{4} (1 - \varphi^2) I(\varphi < 1)$	$\varphi \in (-1, 1)$
Triangular	$(1 - \varphi) I(\varphi < 1)$	$\varphi \in (-1, 1)$
Uniform or rectangular	$\frac{1}{2} I(\varphi < 1)$	$\varphi \in (-1, 1)$

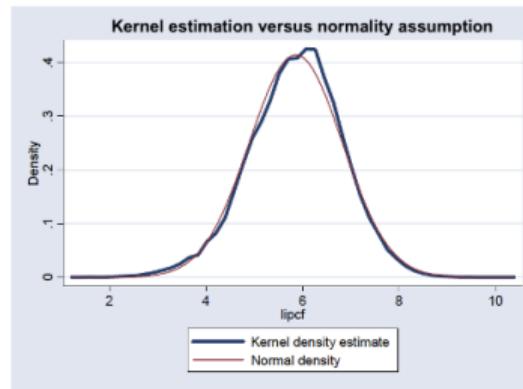
The weighted average estimator or kernel method

Example: Income Distribution



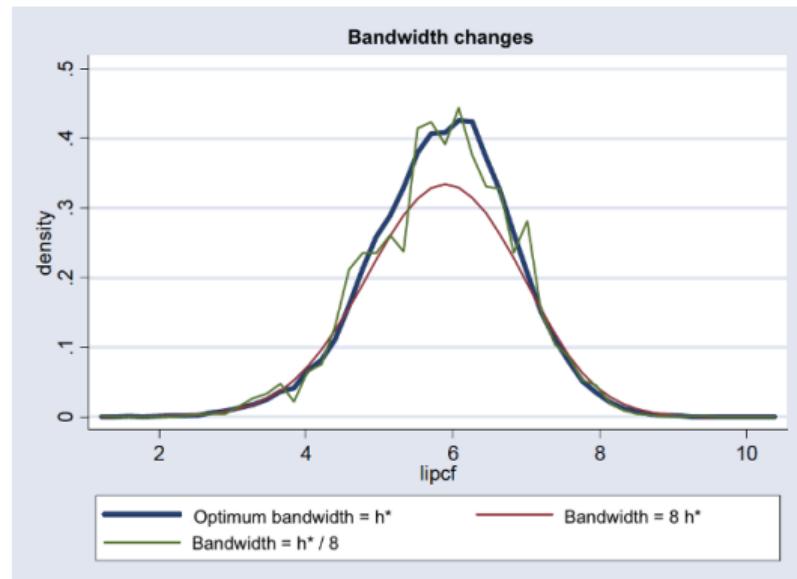
The weighted average estimator or kernel method

Example: Income Distribution



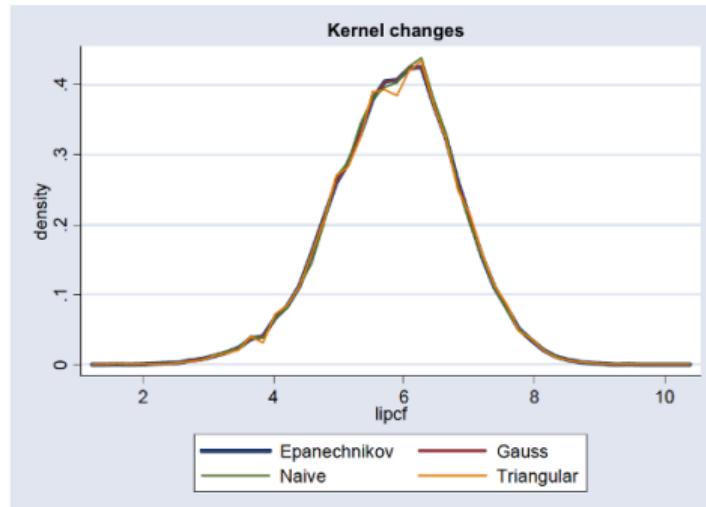
The weighted average estimator or kernel method

Example: Income Distribution



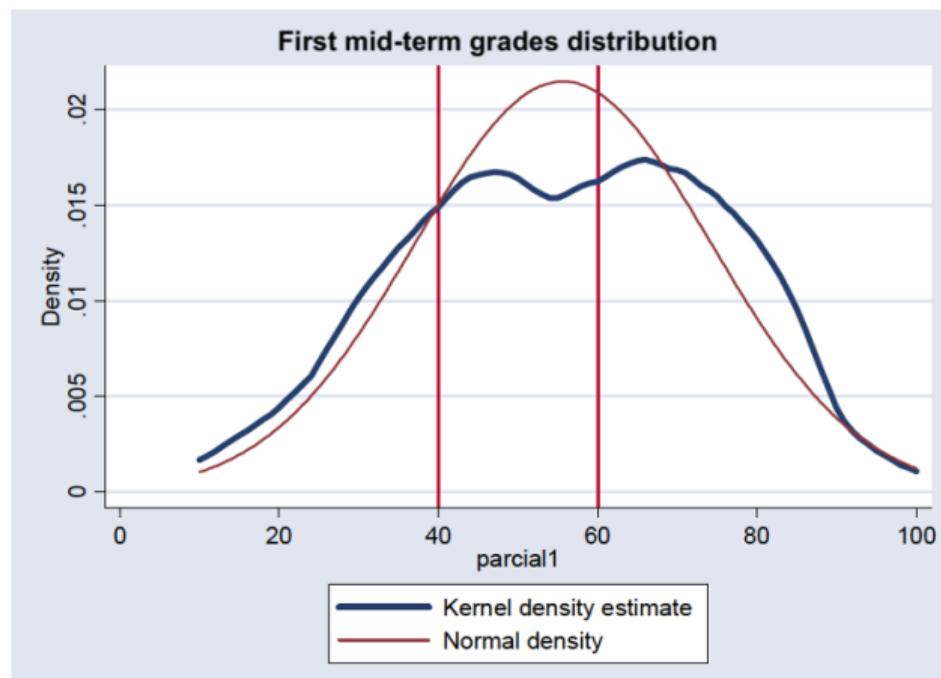
The weighted average estimator or kernel method

Example: Income Distribution



The weighted average estimator or kernel method

Example: Income Distribution



The weighted average estimator or kernel method

Properties: Bias

- ▶ The kernel estimator is generally biased.
- ▶ The approximate expression for the bias is given by:

$$bias[\hat{f}(x_o)] \approx \frac{h^2}{2} f''(x_o) \int_{-\infty}^{\infty} K(\phi) \phi^2 d\phi \quad (6)$$

The weighted average estimator or kernel method

Properties: Variance

- ▶ The approximate expression for the asymptotic variance of the kernel estimator is given by:

$$\text{variance}[\hat{f}(x_o)] \approx \frac{1}{n \times h} f(x_o) \int_{-\infty}^{\infty} K^2(\phi) d\phi \quad (7)$$

The weighted average estimator or kernel method

Properties: MSE

$$MSE(\hat{f}(x)) = Var(\hat{f}(x)) + \left(bias(\hat{f}(x)) \right)^2 \quad (8)$$

The weighted average estimator or kernel method

Properties: MSE

$$MSE(\hat{f}(x)) = Var(\hat{f}(x)) + \left(bias(\hat{f}(x)) \right)^2 \quad (8)$$

$$IMSE(\hat{f}(x)) = \int MSE(\hat{f}(x))dx \quad (9)$$

The weighted average estimator or kernel method

Bandwidth Selection

- We obtain a bandwidth which globally balances bias and variance by minimizing IMSE with respect to h , i.e.,

$$h_{opt} = \left(\frac{\int K^2(z)dz}{(\int z^2 K(z)dz)^2 \int f'(x)^2 dx} \right)^{-1/5} n^{-1/5} \quad (10)$$

The weighted average estimator or kernel method

Bandwidth Selection

- ▶ There are four general approaches to bandwidth selection,
 - 1 reference rules-of-thumb,

The weighted average estimator or kernel method

Reference Rule-of-Thumb

- ▶ The reference rule-of-thumb for choosing the bandwidth uses a standard family of distributions
- ▶ Under a gaussian density and kernel

$$h_{opt} = 1.059\sigma n^{-1/5} \quad (11)$$

The weighted average estimator or kernel method

Bandwidth Selection

- ▶ There are four general approaches to bandwidth selection,
 - 1 reference rules-of-thumb,
 - 2 plug-in methods,
 - 3 cross-validation methods,
 - 4 bootstrap methods.

Multivariate Density

- ▶ Suppose we have X and Y iid
- ▶ We want to estimate

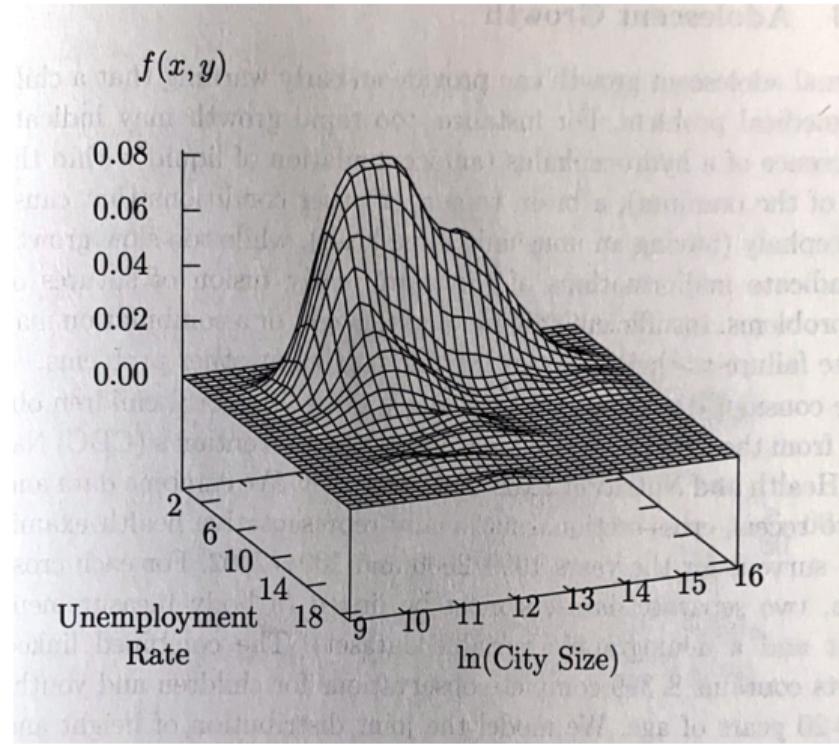
$$f(x, y) \tag{12}$$

Multivariate Density

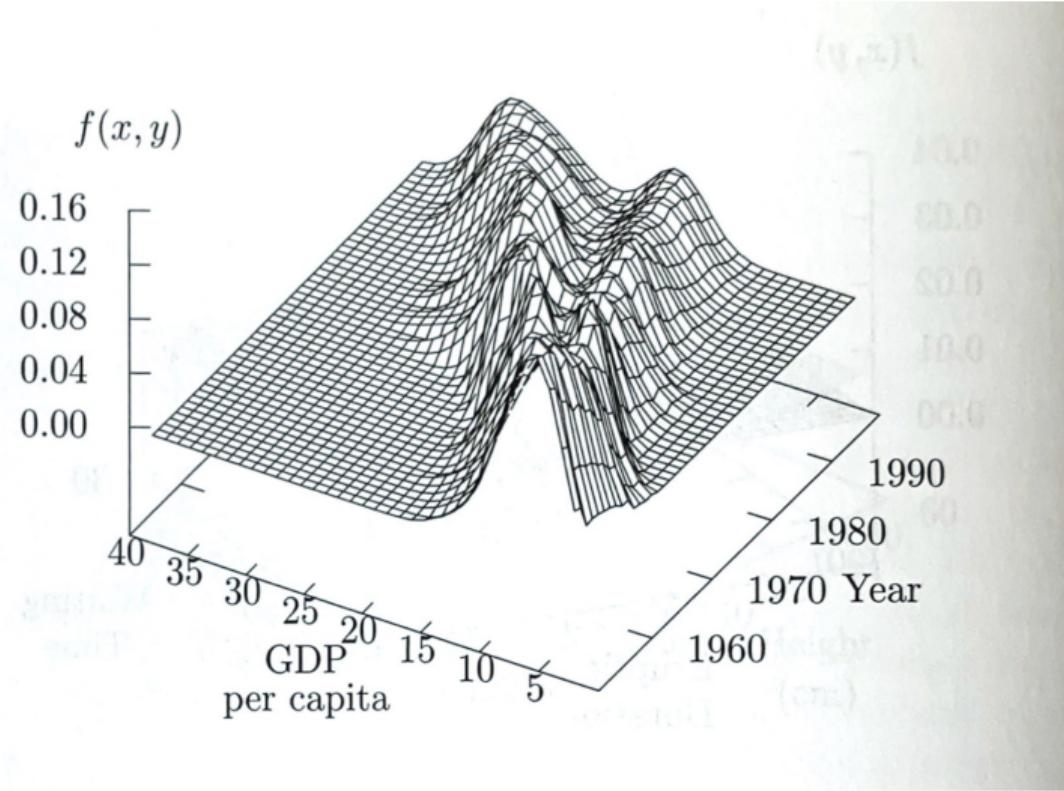
- We use a product kernel function

$$\hat{f}(x, y) = \frac{1}{nh_x h_y} \sum_{i=1}^n K\left(\frac{X_i - x}{h_x}\right) K\left(\frac{Y_i - y}{h_y}\right) \quad (13)$$

Multivariate Density



Multivariate Density



1 Recap: Univariate Density Estimation

2 Kernel Density

3 Regressions: Non parametric estimation of conditional expectations

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4 Further Readings

Non parametric estimation of conditional expectations

- ▶ In this part we are interested in the relationship between y and x where y is the dependent variable and x is the explanatory variable.

$$y = f(x) + u \tag{14}$$

- ▶ where we make no assumptions about $m(x)$
- ▶ The only assumption we add is $E(u|x) = 0$.

Non parametric estimation of conditional expectations

$$E(y|x) = m(x) \tag{15}$$

- ▶ This is true if $E(u|x) = 0$.
- ▶ The question is how we estimate this $m(x)$

Local Constant Kernel Estimation (Nadaraya–Watson)

$$E(y|x) = \int_{-\infty}^{\infty} yf(y|x)dy$$

Local Constant Kernel Estimation (Nadaraya–Watson)

$$\begin{aligned} E(y|x) &= \int_{-\infty}^{\infty} yf(y|x)dy \\ &= \int_{-\infty}^{\infty} y \frac{f(y,x)}{f(x)} dy \\ &= \frac{\int_{-\infty}^{\infty} yf(y,x)dy}{f(x)} \\ &= \frac{v(x)}{g(x)} \end{aligned}$$

Local Constant Kernel Estimation (Nadaraya–Watson)

Then the estimation we can propose is

$$\hat{E}(y|x) = \hat{m}(x) = \frac{\int_{-\infty}^{\infty} y \hat{f}_{x,y}(y, x) dy}{\hat{f}_x(x)} = \frac{\hat{v}(x)}{\hat{g}(x)}$$

then

$$\hat{f}_x(x) = \frac{1}{nh^q} \sum \underline{K}\left(\frac{x_i - x}{h}\right)$$

where $\underline{K}\left(\frac{x_i - x}{h}\right) = K\left(\frac{x_{1i} - x}{h}\right) \cdot K\left(\frac{x_{2i} - x}{h}\right) \dots \dots K\left(\frac{x_{qi} - x}{h}\right)$

Local Constant Kernel Estimation (Nadaraya–Watson)

$$\hat{f}_{x,y}(y, x) = \frac{1}{nh^{q+1}} \sum K\left(\frac{y_i - y}{h}\right) \underline{K}\left(\frac{x_i - x}{h}\right)$$

replacing

$$\hat{v}(x) = \frac{1}{nh^{q+1}} \sum \underline{K}\left(\frac{x_i - x}{h}\right) \int y K\left(\frac{y_i - y}{h}\right) dy$$

$$\hat{v}(x) = \frac{1}{nh^q} \sum \underline{K}\left(\frac{x_i - x}{h}\right) y_i$$

Local Constant Kernel Estimation (Nadaraya–Watson)

Then we have

$$\hat{m}(x) = \frac{\hat{v}(x)}{\hat{g}(x)} = \frac{\frac{1}{nh^q} \sum \underline{K}\left(\frac{x_i-x}{h}\right) y_i}{\frac{1}{nh^q} \sum \underline{K}\left(\frac{x_i-x}{h}\right)} = \frac{\sum y_i \underline{K}\left(\frac{x_i-x}{h}\right)}{\sum \underline{K}\left(\frac{x_i-x}{h}\right)}$$

Note that we can write $w_i = \frac{\underline{K}\left(\frac{x_i-x}{h}\right)}{\sum \underline{K}\left(\frac{x_i-x}{h}\right)}$

$$\hat{m}(x) = \sum y_i w_i$$

Local Constant Kernel Estimation (Nadaraya–Watson)

We can show that this is the same to solving the followign problem

$$\underset{a}{\text{Min}} \sum_{i=1}^n w_i (y_i - a)^2$$

where $\hat{a}(x) = \hat{m}(x)$.

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Local Linear Kernel Estimation

We can think to fit a line instead of a constant

$$\underset{\{a,b\}}{\text{Min}} \sum_{i=1}^n w_i \{y_i - [a + b(x_i - x)]\}^2$$

Intuition

- ▶ The question is how we can write $m(x_i)$ in terms of x .

Intuition

- ▶ The question is how we can write $m(x_i)$ in terms of x .
- ▶ With a linear approximation

$$m(x_i) \approx \underbrace{m(x)}_a + \underbrace{m'(x)}_b (x_i - x)$$

- ▶ We get a marginal effect

Local Linear Kernel Estimation

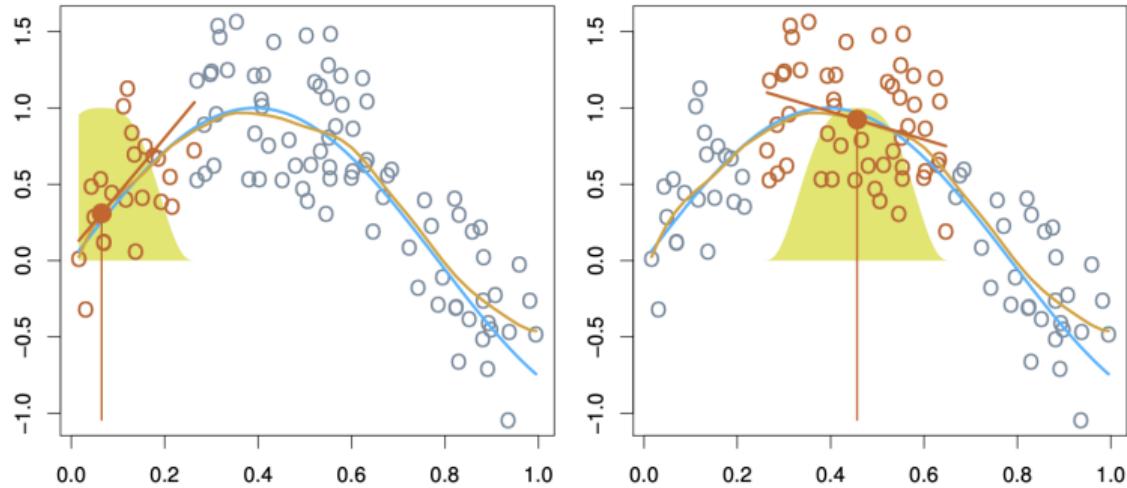
We can show that

$$\hat{a}(x) = y_w - \hat{b} \underbrace{(\bar{x}_w - x)}_{\text{weighted local mean}}$$

$$\hat{b}(x) = \frac{\sum (y_i - \bar{y}_w) (x_i - \bar{x}_w) w_i}{\sum (x_i - \bar{x}_w) w_i}$$

with $\bar{x}_w = \frac{\sum x_i w_i}{\sum w_i}$ y $\bar{y}_w = \frac{\sum y_i w_i}{\sum w_i}$ are the weighted means

Intuition



Example: Geographical Weighted Regressions

The determinant of educational achievement in Georgia (Fotheringham et al., 2002)

$$Bachelor\ or\ higher\ degrees_{county} = f(Rural, Eld, FB, Pov, Black) + u \quad (16)$$

Example: Geographical Weighted Regressions

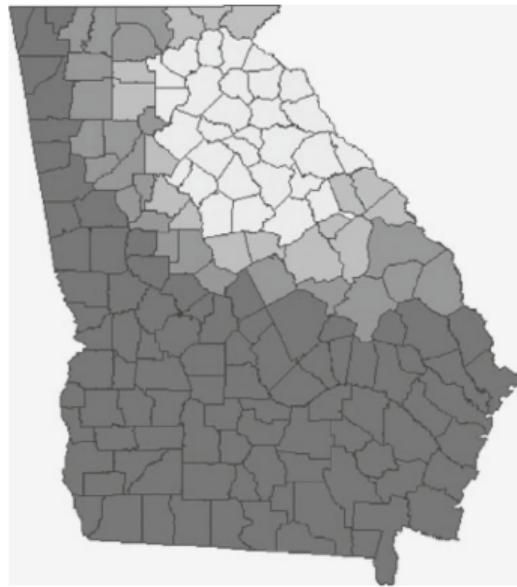
The determinant of educational achievement in Georgia (Fotheringham et al., 2002)

$$Bachelor\ or\ higher\ degrees_{county} = f(Rural, Eld, FB, Pov, Black) + u \quad (16)$$

Parameter	Minimum	First Quartile	Median	Third Quartile	Maximum	Range	Global (OLS)
Intercept	14.170000	15.350000	17.050000	18.200000	18.860000	4.69	17.2437
Rural	-0.081350	-0.073480	-0.064850	-0.055110	-0.051080	0.03027	-0.0703
Eld	-0.191200	-0.094630	-0.065330	-0.032360	0.012500	0.2037	0.0114
FB	0.854300	1.282000	2.031000	2.796000	3.138000	2.2837	1.8525
Pov	-0.304800	-0.258100	-0.196100	-0.115100	-0.034210	0.27059	-0.2552
Black	-0.016900	0.006347	0.031610	0.060620	0.087210	0.07031	0.0491

Example: Geographical Weighted Regressions

The determinant of educational achievement in Georgia (Fotheringham et al., 2002)



Further Readings

- ▶ Li, Q., & Racine, J. S. (2007). Nonparametric econometrics: theory and practice. Princeton University Press.
- ▶ Pagan, A. & Ullah (1999). Nonparametric Econometrics. Cambridge University Press.
- ▶ Silverman (1998). Density Estimation for Statistics and Data Analysis. Chapman & Hall/CRC.