

<u>objedus</u>

- Maximize Between-Class Variance: Ensures that different classes are as far apart as possible.
- Minimize Within-Class Variance: Ensures that data points within the same class are as close to each other as possible.



Assumptions

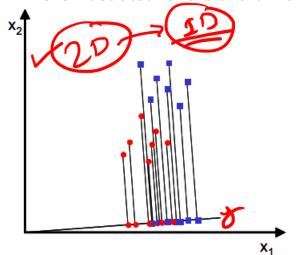
LDA operates under several key assumptions:

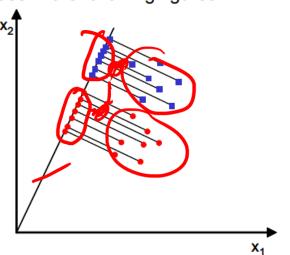
- 1. **Normal Distribution**: Features are assumed to follow a multivariate normal distribution within each class.
- 2. **Equal Covariance Matrices:** All classes share the same covariance matrix, implying that they have similar shapes in feature space.
- Linearly Separable Classes. Classes can be separated by linear boundaries.

- /-
 - The objective of LDA is to perform dimensionality reduction while preserving as much of the class discriminatory information as possible
 - Assume we have a set of D-dimensional samples {x⁽¹, x⁽², ..., x^(N)}, N₁ of which belong to class ω₁, and N₂ to class ω₂. We seek to obtain a scalar y by projecting the samples x onto a line

 $y = w^T x$

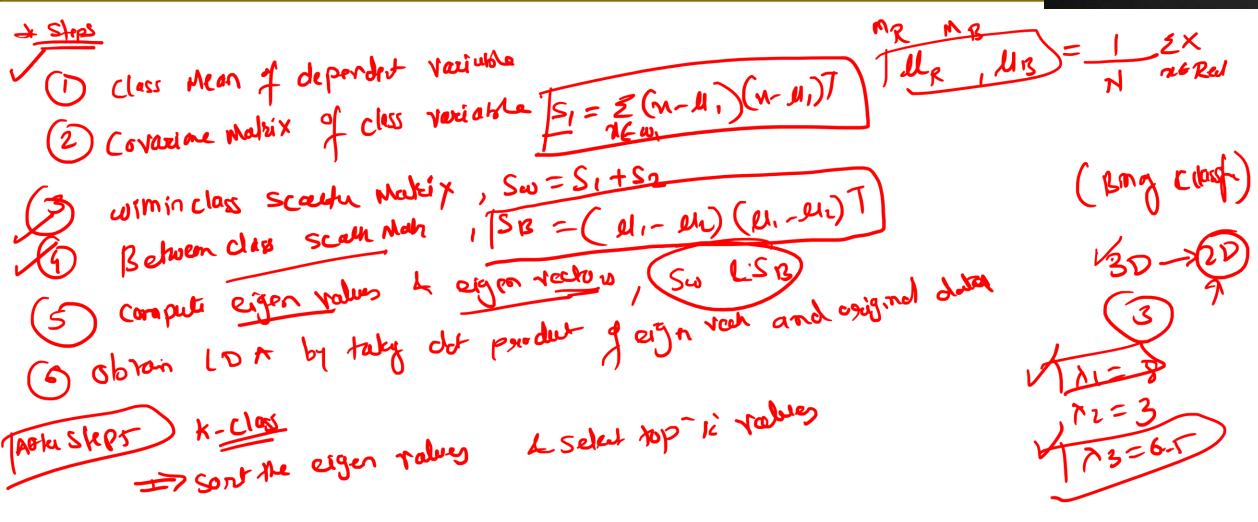
- Of all the possible lines we would like to select the one that <u>maximizes</u> the separability of the scalars
 - This is illustrated for the two-dimensional case in the following figures







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 Compute the Linear Discriminant projection for the following two-dimensional dataset

$$\begin{array}{c} \text{All} & \text{All$$

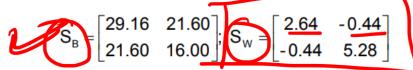


· The class statistics are:

$$\begin{bmatrix} 0.80 & -0.40 \\ -0.40 & 2.60 \end{bmatrix}; S_2 = \begin{bmatrix} 1.84 & -0.04 \\ -0.04 & 2.64 \end{bmatrix}$$

$$\begin{bmatrix} \mu_1 & [3.00 & 3.60] & \mu_2 & [8.40 & 7.60] \end{bmatrix}$$

The within- and between-class scatter are



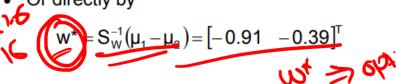
2 X_1

The LDA projection is then obtained as the solution of the generalized eigenvalue problem

$$S_{W}^{-1}S_{B}v = \lambda v \Rightarrow \begin{vmatrix} S_{W}^{-1}S_{B} - \lambda I \\ S_{W}^{-1}S_{B} - \lambda I \end{vmatrix} = 0 \Rightarrow \begin{vmatrix} 11.89 - \lambda & 8.81 \\ 5.08 & 3.76 - \lambda \end{vmatrix} = 0 \Rightarrow \lambda = 15.65$$

$$\begin{bmatrix} 11.89 & 8.81 \\ 5.08 & 3.76 \end{bmatrix} \begin{bmatrix} v_{1} \\ v_{2} \end{bmatrix} = 15.65 \begin{bmatrix} v_{1} \\ v_{2} \end{bmatrix} \Rightarrow \begin{bmatrix} v_{1} \\ v_{2} \end{bmatrix} = \begin{bmatrix} 0.91 \\ 0.39 \end{bmatrix}$$
• Or directly by

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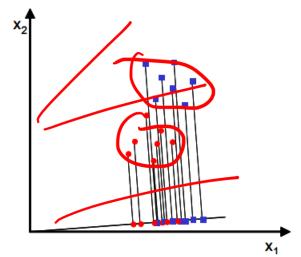


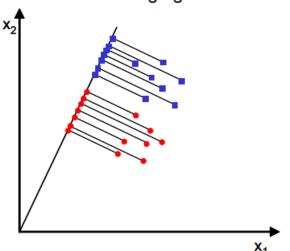
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- The objective of LDA is to perform dimensionality reduction while preserving as much of the class discriminatory information as possible
 - Assume we have a set of D-dimensional samples $\{x^{(1)}, x^{(2)}, ..., x^{(N)}\}$, N_1 of which belong to class ω_1 , and N_2 to class ω_2 . We seek to obtain a scalar \boldsymbol{y} by projecting the samples \boldsymbol{x} onto a line

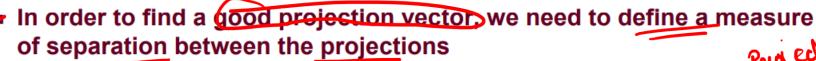
$$y = w^T x$$

- Of all the possible lines we would like to select the one that maximizes the separability of the scalars
 - This is illustrated for the two-dimensional case in the following figures









■ The mean vector of each class in x and y feature space is

w= Pagedia vector

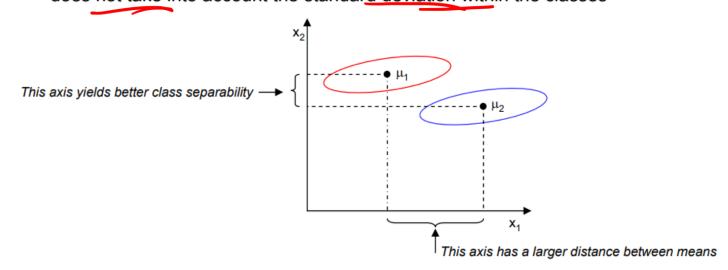
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$$\mu_i = \frac{1}{N_i} \sum_{x \in \omega_i} x \quad \text{ and } \quad \widetilde{\mu}_i = \frac{1}{N_i} \sum_{y \in \omega_i} y = \frac{1}{N_i} \sum_{x \in \omega_i} w^T x = w^T \mu_i$$

We could then choose the distance between the projected means as our objective function

$$J(N) = |\widetilde{\mu}_1 - \widetilde{\mu}_2| = |W|(\mu_1 - \mu_2)|$$

 However, the distance between the projected means is not a very good measure since it does not take into account the standard deviation within the classes





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- The solution proposed by Fisher is to maximize a function that represents the difference between the means, normalized by a measure of the within-class scatter
 - For each class we define the scatter, an equivalent of the variance, as



- where the quantity $(\tilde{S}_1^2 + \tilde{S}_2^2)$ is called the within-class scatter of the projected examples
- The Fisher linear discriminant is defined as the linear function wtx that maximizes the criterion function



In order to find the optimum projection w, we need to express J(w) as an explicit function of w



We define a measure of the scatter in multivariate feature space x, which are scatter matrices

$$S_i = \sum_{\mathbf{x} \in \mathcal{U}_i} (\mathbf{x} - \mathbf{\mu}_i) (\mathbf{x} - \mathbf{\mu}_i)^T$$

$$S_1 + S_2 = S_W$$

where S_W is called the within-class scatter matrix

The scatter of the projection y can then be expressed as a function of the scatter matrix in feature space x



$$\widetilde{\mathbf{S}}_{i}^{2} = \sum_{\mathbf{y} \in \omega_{i}} (\mathbf{y} - \widetilde{\boldsymbol{\mu}}_{i})^{2} = \sum_{\mathbf{x} \in \omega_{i}} (\mathbf{w}^{\mathsf{T}} \mathbf{x} - \mathbf{w}^{\mathsf{T}} \boldsymbol{\mu}_{i})^{2} = \sum_{\mathbf{x} \in \omega_{i}} \mathbf{w}^{\mathsf{T}} (\mathbf{x} - \boldsymbol{\mu}_{i}) (\mathbf{x} - \boldsymbol{\mu}_{i})^{\mathsf{T}} \mathbf{w} = \underline{\mathbf{w}}^{\mathsf{T}} \mathbf{S}_{i} \mathbf{w}$$

$$\widetilde{\mathbf{S}}_{1}^{2} + \widetilde{\mathbf{S}}_{2}^{2} = \mathbf{w}^{\mathsf{T}} \mathbf{S}_{w} \mathbf{w}$$

$$\widetilde{\mathbf{S}}_{1}^{2} + \widetilde{\mathbf{S}}_{2}^{2} = \mathbf{w}^{\mathsf{T}} \mathbf{S}_{w} \mathbf{w}$$

Similarly, the difference between the projected means can be expressed in terms of the means in the original feature space

$$(\widetilde{\boldsymbol{\mu}}_{\!\!1} - \widetilde{\boldsymbol{\mu}}_{\!\!2})^2 = (\boldsymbol{w}^\mathsf{T} \boldsymbol{\mu}_{\!\!1} - \boldsymbol{w}^\mathsf{T} \boldsymbol{\mu}_{\!\!2})^2 = \boldsymbol{w}^\mathsf{T} \underbrace{(\boldsymbol{\mu}_{\!\!1} - \boldsymbol{\mu}_{\!\!2}) (\boldsymbol{\mu}_{\!\!1} - \boldsymbol{\mu}_{\!\!2})^\mathsf{T}}_{\widetilde{\boldsymbol{S}}_{\!\scriptscriptstyle B}} \boldsymbol{w} = \boldsymbol{w}^\mathsf{T} \boldsymbol{S}_{\!\scriptscriptstyle B} \boldsymbol{w}$$

- The matrix S_B is called the **between-class scatter**. Note that, since S_B is the outer product of two vectors, its rank is at most one
- We can finally express the Fisher criterion in terms of S_W and S_B as

$$J(w) = \frac{w^{T}S_{B}w}{w^{T}S_{W}w}$$



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• To find the maximum of J(w) we derive and equate to zero

$$\frac{d}{dw} [J(w)] = \frac{d}{dw} \left[\frac{w^{\mathsf{T}} S_{\mathsf{B}} w}{w^{\mathsf{T}} S_{\mathsf{W}} w} \right] = 0 \implies$$

$$\Rightarrow \left[w^{\mathsf{T}} S_{\mathsf{W}} w \right] \frac{d \left[w^{\mathsf{T}} S_{\mathsf{B}} w \right]}{dw} - \left[w^{\mathsf{T}} S_{\mathsf{B}} w \right] \frac{d \left[w^{\mathsf{T}} S_{\mathsf{W}} w \right]}{dw} = 0 \implies$$

$$\Rightarrow \left[w^{\mathsf{T}} S_{\mathsf{W}} w \right] 2 S_{\mathsf{B}} w - \left[w^{\mathsf{T}} S_{\mathsf{B}} w \right] 2 S_{\mathsf{W}} w = 0$$



Dividing by w^TS_ww

$$\begin{bmatrix}
w^{\mathsf{T}} S_{\mathsf{W}} \mathbf{w} \\
w^{\mathsf{T}} S_{\mathsf{W}} \mathbf{w}
\end{bmatrix} S_{\mathsf{B}} \mathbf{w} - \begin{bmatrix}
w^{\mathsf{T}} S_{\mathsf{B}} \mathbf{w} \\
w^{\mathsf{T}} S_{\mathsf{W}} \mathbf{w}
\end{bmatrix} S_{\mathsf{W}} \mathbf{w} = 0 \implies S_{\mathsf{B}} \mathbf{w} - J S_{\mathsf{W}} \mathbf{w} = 0 \implies S_{\mathsf{$$

Solving the generalized eigenvalue problem (S_w⁻¹S_Rw=Jw) yields

 $\mathbf{w} = \underset{\mathbf{w}}{\operatorname{argmax}} \left\{ \frac{\mathbf{w}^{\mathsf{T}} \mathbf{S}_{\mathsf{B}} \mathbf{w}}{\mathbf{w}^{\mathsf{T}} \mathbf{S}_{\mathsf{W}} \mathbf{w}} \right\} = \mathbf{S}_{\mathsf{W}}^{-1} (\mu_{1} - \mu_{2}) \mathbf{v}$

 This is know as <u>Fisher's Linear Discriminant</u> (1936), although it is not a discriminant but rather a specific choice of direction for the projection of the data down to one dimension

