# STAT 210 Applied Statistics and Data Analysis Multiple Linear Regression 5 Multicollinearity

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# Collinearity<sup>1</sup>

 $<sup>^1</sup>$ See Chapter 15, Applied Statistics with R, David Dalpiaz, https://daviddalpiaz.github.io/appliedstats/

Two regressors  $X_1$  and  $X_2$  are **collinear** if they are linearly dependent, i.e. there exist constants  $c_1$ ,  $c_2$  and  $c_3$ , not all equal to zero, such that

$$c_1 X_1 + c_2 X_2 = c_3 \tag{1}$$

This may happen, for instance, when two variables in a regression represent the same magnitude in different scales (weight in kilograms and pounds) or when the total amount of two variables is fixed, as when two chemicals are chosen so that the sum of their weights or volumes is fixed.

In these examples, both variables have the same information, and including them in the same model causes the design matrix to be singular. Only one of them should be included in the model.

This example of exact collinearity is taken from the book by Dalpiaz.

```
collin_data = function(num_samples = 100) {
 x1 = rnorm(n = num samples, mean = 80, sd = 10)
 x2 = rnorm(n = num samples, mean = 70, sd = 5)
 x3 = 2 * x1 + 4 * x2 + 3
 v = 3 + x1 + x2 + rnorm(n = num samples,
                          mean = 0, sd = 1)
 data.frame(y, x1, x2, x3)
set.seed(123)
collin_exmpl <- collin_data()</pre>
```

```
collin.lm \leftarrow lm(y \sim x1 + x2 + x3, data = collin_exmpl)
S(collin.lm)
## Call: lm(formula = y ~ x1 + x2 + x3, data = collin_exmpl)
##
## Coefficients: (1 not defined because of singularities)
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 3.86708 1.65405 2.338 0.0214 *
## x1
         0.98668 0.01049 94.087 <2e-16 ***
## x2
         ## x3
                  NA
                            NA
                                   NA
                                           NΑ
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard deviation: 0.9513 on 97 degrees of freedom
## Multiple R-squared: 0.9912
## F-statistic: 5491 on 2 and 97 DF. p-value: < 2.2e-16
     AIC
           BIC
##
## 278.76 289.18
```

We see that the third variable has been excluded from the regression.

In this case, the design matrix is

```
X = cbind(1, as.matrix(collin_exmpl[,-1]))
```

and if we try to invert X'X

```
solve(t(X) %*% X)
```

we get a warning that says

Error in solve.default(t(X) %\*% X) : system is computationally singular: reciprocal condition number = 1.01841e-17 ...

When this happens, we have exact collinearity.

The fitted model was  $y \sim x1 + x2$  and excluded one of the variables, in this case, x3, but observe that other models would accomplish exactly the same fit

```
fit1 = lm(y \sim x1 + x2, data = collin_exmpl)
fit2 = lm(y \sim x1 + x3, data = collin_exmpl)
fit3 = lm(y \sim x2 + x3, data = collin_exmpl)
```

The fitted values for these three models are exactly the same:

all.equal(fitted(fit1), fitted(fit2))

```
## [1] TRUE
all.equal(fitted(fit2), fitted(fit3))
```

```
## [1] TRUE
```

But the estimated coefficients are not

```
coef(fit1); coef(fit2); coef(fit3)
   (Intercept)
                         x1
                                     x2.
     3.8670796
                 0.9866828
                              1.0047623
##
   (Intercept)
                         x1
                                     x3
     3.1135079
                 0.4843017
                              0.2511906
##
## (Intercept)
                         x2
                                     x3
     2.3870554 -0.9686034
##
                              0.4933414
```

However, only the first model explains the relationship between the variables.

The other models are able to predict correctly, but the coefficients are meaningless.

Approximate collinearity happens if equation (1) is approximately true

Collinearity between  $X_1$  and  $X_2$  is measured by the square of their sample correlation  $r_{12}^2$ .

Exact collinearity corresponds to  $r_{12}^2 = 1$  while non-collinearity corresponds to  $r_{12}^2 = 0$ .

If  $r_{12}^2$  is close to 1, we have approximate collinearity.



For p>2 regressors, approximate collinearity happens if there are constants  $c_0,c_1,\ldots,c_p$  not all equal to zero so that

$$c_1X_1+c_2X_2+\cdots+c_pX_p\approx c_0$$

Observe that if  $c_i \neq 0$ , then we can write  $X_i$  approximately as a linear combination of the other variables.

In this case, instead of the squared correlation, variable  $X_i$  is regressed on the X's, and the  $R^2$  for this regression is considered as the multiple correlation between  $X_i$  and the other variables and denoted by  $R_i^2$ .

If the largest  $R_i^2$  is close to 1, we have approximate collinearity.

When a set of predictors is exactly collinear, one or more predictors must be deleted to be able to estimate the coefficients for the model.

Since the information in the deleted predictor is contained in the other regressors, no information is lost in this process. However, the interpretation of the parameters may be different or more complex.

When approximate collinearity is present, the usual remedy is again to delete variables, with loss of information expected to be small.

The tricky part may be deciding which variable(s) to delete.

One important effect of a high correlation between regressors is the increased variance of the estimates.

The sampling variance of  $\hat{\beta}_j$  is

$$Var(\hat{eta}_j) = rac{1}{1 - R_j^2} rac{\sigma^2}{(n-1)S_j^2}$$

where

$$S_j^2 = \frac{1}{n-1} \sum_i (x_{ij} - \bar{x}_{\bullet j})^2$$

is the sample variance of  $X_j$ .

The term  $1/(1-R_j^2)$ , known as the variance inflation factor (VIF), indicates the effect of collinearity on the variance of  $\hat{\beta}_j$ .

This simulated example is from S. Weisberg *Applied Linear Regression*, Wiley.

We consider two models of the form

$$Y = 1 + X_1 + X_2 + 0 \cdot X_3 + 0 \cdot X_4 + \epsilon$$

where  $\epsilon \sim N(0,1)$ .

In the first model  $\mathbf{X} = (X_1, X_2, X_3, X_4)$  are independent normal random variables while in the second case the covariance matrix is

$$\begin{pmatrix}
1 & 0 & .95 & 0 \\
0 & 1 & 0 & -.95 \\
.95 & 0 & 1 & 0 \\
0 & -.95 & 0 & 1
\end{pmatrix}$$

so that  $X_1$  and  $X_3$  are highly positively correlated while  $X_2$  and  $X_3$  are highly negatively correlated.

We fit linear models in each case

```
library(mvtnorm)
sigma1 <- diag(4); sigma2 <- sigma1
sigma2[3,1] \leftarrow sigma2[1,3] \leftarrow 0.95
sigma2[4,2] \leftarrow sigma2[2,4] \leftarrow -0.95
sample1 <- rmvnorm(100, sigma = sigma1)</pre>
sample1 <- data.frame(sample1)</pre>
colnames(sample1) <- c('X1','X2','X3','X4')</pre>
sample2 <- rmvnorm(100, sigma = sigma2)</pre>
colnames(sample1) <- c('X1','X2','X3','X4')</pre>
v1 < 1 + sample1[,1] + sample1[,2] + rnorm(100)
y2 \leftarrow 1 + sample2[,1] + sample2[,2] + rnorm(100)
```

```
col1 \leftarrow lm(v1 \sim X1 + X2 + X3 + X4, data = sample1)
summary(col1)
##
## Call:
## lm(formula = v1 \sim X1 + X2 + X3 + X4, data = sample1)
##
## Residuals:
##
      Min
               10 Median
                               30
                                     Max
## -2.1377 -0.6255 -0.0358 0.4447 2.8748
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.96891 0.10085 9.607 1.14e-15 ***
## X1
              0.93469 0.10538 8.870 4.30e-14 ***
## X2
             0.72390 0.11863 6.102 2.26e-08 ***
            -0.01678 0.09660 -0.174 0.862
## X3
            -0.07702 0.09004 -0.855 0.395
## X4
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9878 on 95 degrees of freedom
## Multiple R-squared: 0.5539, Adjusted R-squared: 0.5351
## F-statistic: 29.49 on 4 and 95 DF, p-value: 6.09e-16
```

```
set.seed(7364)
sample2 <- rmvnorm(100, sigma = sigma2); sample2 <- data.frame(sample2)
colnames(sample2) <- c('X1','X2','X3','X4')
y2 <- 1 + sample2[,1] + sample2[,2] + rnorm(100)
col2a <- lm(y2 - X1 + X2 + X3 + X4, data = sample2)
summary(col2a)</pre>
```

```
##
## Call:
## lm(formula = y2 \sim X1 + X2 + X3 + X4, data = sample2)
##
## Residuals:
            1Q Median 3Q
     Min
                                   Max
## -2.4231 -0.7921 0.1726 0.8837 2.2925
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1.0315 0.1111 9.288 5.48e-15 ***
## X1
            0.9071 0.2988 3.035 0.0031 **
## X2
     0.6506 0.3526 1.846 0.0681 .
## X3 0.1342 0.3084 0.435 0.6645
## X4
         -0.1462 0.3597 -0.406 0.6854
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.083 on 95 degrees of freedom
## Multiple R-squared: 0.5786, Adjusted R-squared: 0.5609
## F-statistic: 32.62 on 4 and 95 DF, p-value: < 2.2e-16
```

```
set.seed(574597)
sample2 <- rmvnorm(100, sigma = sigma2); sample2 <- data.frame(sample2)
colnames(sample2) <- c('X1','X2','X3','X4')
y2 <- 1 + sample2[,1] + sample2[,2] + rnorm(100)
col2b <- lm(y2 - X1 + X2 + X3 + X4, data = sample2)
summary(col2b)</pre>
```

```
##
## Call:
## lm(formula = y2 \sim X1 + X2 + X3 + X4, data = sample2)
##
## Residuals:
              1Q Median 3Q
       Min
                                       Max
## -2.15898 -0.63164 0.08673 0.63820 2.40883
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1.0798 0.1007 10.728 <2e-16 ***
## X1
            0.5737 0.3641 1.576 0.118
## X2
     0.4461 0.3069 1.454 0.149
## X3 0.3938 0.3682 1.070 0.287
## X4
          -0.3111 0.3006 -1.035 0.303
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.002 on 95 degrees of freedom
## Multiple R-squared: 0.5494, Adjusted R-squared: 0.5304
## F-statistic: 28.95 on 4 and 95 DF, p-value: 9.757e-16
```

## Multiple R-squared: 0.6215, Adjusted R-squared: 0.6056
## F-statistic: 39 on 4 and 95 DF. p-value: < 2.2e-16</pre>

```
set.seed(16299125)
sample2 <- rmvnorm(100, sigma = sigma2); sample2 <- data.frame(sample2)</pre>
colnames(sample2) <- c('X1','X2','X3','X4')</pre>
y2 \leftarrow 1 + sample2[,1] + sample2[,2] + rnorm(100)
col2c <- lm(y2 ~ X1 + X2 + X3 + X4, data = sample2)
summary(col2c)
##
## Call:
## lm(formula = y2 \sim X1 + X2 + X3 + X4, data = sample2)
##
## Residuals:
               1Q Median 3Q
       Min
                                          Max
## -2.73719 -0.60276 0.03033 0.62877 2.25573
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1.0413 0.1069 9.744 5.81e-16 ***
## X1
              0.4285 0.3333 1.286 0.201684
## X2
          1.5355 0.4166 3.686 0.000379 ***
## X3
          0.5874 0.3397 1.729 0.087035 .
## X4
            0.6302 0.4048 1.557 0.122821
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.06 on 95 degrees of freedom
```

```
round(vif(col2a),3); round(vif(col2b),3); round(vif(col2c),3);
     X 1
           X2 X3 X4
##
## 7.140 9.568 7.279 9.416
##
      X 1
            X2
                   ХЗ
                          Х4
## 12.010 8.053 12.143 8.051
      X1
             X2
                   ХЗ
                          Х4
##
## 12.691 12.060 12.753 12.090
```

```
For the rat example
```

```
vif(m1); vif(m1b)

## BodyWt LiverWt Dose
## 52.101917 1.335679 51.427154

## BodyWt LiverWt Dose
## 259.449422 1.445674 253.199751

For the UN11 model
vif(lm1)
```

```
## log(ppgdp) pctUrban
## 2.272698 2.272698
```

The next example comes from https://datascienceplus.com/multicollinearity-in-r/

It is also considered in https://www.r-bloggers.com/dealing-with-the-problem-of-multicollinearity-in-r/  $\,$ 

The data file can be downloaded from StatLib at http://lib.stat.cmu.edu/datasets/CPS\_85\_Wages

```
data1 = read.table('CPS_85_Wages.txt', header = T)
head(data1, 4)
```

```
Education South Sex Experience Union Wage Age Race
##
                                        0.5.10
## 1
                                 21
                                                35
## 2
                                 42 0 4.95 57
                                        0.6.67
## 3
                                                19
                                        0 4.00
## 4
            12
                                  4
                                                22
                                                      3
##
     Occupation Sector Marr
## 1
              6
## 2
              6
## 3
              6
              6
                          0
## 4
```

'The Current Population Survey (CPS) is used to supplement census information between census years. These data consist of a random sample of 534 persons from the CPS, with information on wages and other characteristics of the workers, including sex, number of years of education, years of work experience, occupational status, region of residence and union membership. We wish to determine whether wages are related to these characteristics.'

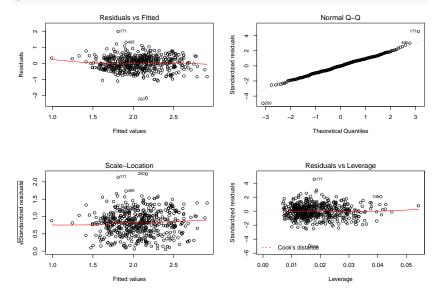
```
str(data1)
## 'data.frame': 534 obs. of 11 variables:
   $ Education : int 8 9 12 12 12 13 10 12 16 12 ...
##
##
   $ South
              : int
                    0 0 0 0 0 0 1 0 0 0 ...
##
   $ Sex
              : int 1100000000...
##
   $ Experience: int
                    21 42 1 4 17 9 27 9 11 9 ...
##
   $ Union
              : int
                    0000010000...
##
   $ Wage
              : num
                    5.1 4.95 6.67 4 7.5 ...
##
   $ Age
              : int
                    35 57 19 22 35 28 43 27 33 27 ...
##
   $ Race
              : int
                    2 3 3 3 3 3 3 3 3 ...
##
   $ Occupation: int 6 6 6 6 6 6 6 6 6 6 ...
##
   $ Sector
              : int
                     1 1 1 0 0 0 0 0 1 0 ...
##
   $ Marr
              : int 1100100010...
```

```
fit_model1 = lm(log(data1$Wage) ~., data = data1)
summary(fit_model1)
```

```
##
## Call:
## lm(formula = log(data1$Wage) ~ ., data = data1)
##
## Residuals:
                 10 Median
##
       Min
                                  3Q
                                          Max
## -2.16246 -0.29163 -0.00469 0.29981 1.98248
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1.078596 0.687514 1.569 0.117291
## Education
              0.179366 0.110756 1.619 0.105949
## South
            -0.102360 0.042823 -2.390 0.017187 *
## Sex
            -0.221997 0.039907 -5.563 4.24e-08 ***
## Experience 0.095822 0.110799 0.865 0.387531
              0.200483 0.052475 3.821 0.000149 ***
## Union
## Age
            -0.085444 0.110730 -0.772 0.440671
## Race
             0.050406 0.028531 1.767 0.077865 .
## Occupation -0.007417 0.013109 -0.566 0.571761
## Sector
              0.091458 0.038736 2.361 0.018589 *
## Marr
               0.076611 0.041931 1.827 0.068259 .
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.4398 on 523 degrees of freedom
## Multiple R-squared: 0.3185, Adjusted R-squared: 0.3054
## F-statistic: 24.44 on 10 and 523 DF, p-value: < 2.2e-16
```

Observe that four variables are not significant, Education, Experience, Age, and Occupation while two other variables are only significant at the 0.1 level, Race and Marr.

```
par(mfrow=c(2,2))
plot(fit_model1)
```



#### Variance Inflation Factors:

```
round(vif(fit_model1),3)
```

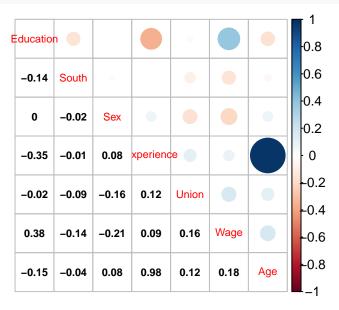
```
##
    Education
                    South
                                 Sex Experience
                                                       Union
      231.196
                    1.047
                                1.092
                                        5184.094
                                                       1.121
##
                     Race Occupation
                                          Sector
                                                        Marr
##
          Age
##
     4645,665
                    1.037
                               1.298
                                           1,199
                                                       1.096
```

Two nice plots for the correlation matrix:

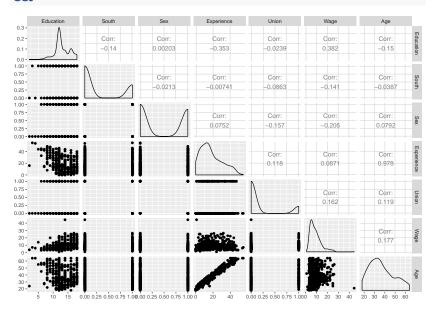
Reduce the number of variables.

```
X<-data1[,-(8:11)]
library(GGally)
library(corrplot)
cor1 = cor(X)</pre>
```

```
corrplot.mixed(cor1, lower.col = 'black', number.cex = .7, tl.cex=0.7)
```



#### ggpairs(X)



##

Variables age and experience are very highly correlated, so we only include one of them, age.

```
fit_model2 <- update(fit_model1, ~. - Experience, data = data1)
summary(fit_model2)
##
## Call:
## lm(formula = log(data1$Wage) ~ Education + South + Sex + Union +
      Age + Race + Occupation + Sector + Marr, data = data1)
##
##
## Residuals:
##
       Min
                 1Q
                     Median
                                          Max
## -2.16018 -0.29085 -0.00513 0.29985 1.97932
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 0.501358 0.164794 3.042 0.002465 **
## Education
               0.083815 0.007728 10.846 < 2e-16 ***
            -0.103186 0.042802 -2.411 0.016261 *
## South
## Sex
            -0.220100 0.039837 -5.525 5.20e-08 ***
## Union
             0.200018 0.052459 3.813 0.000154 ***
## Age
             0.010305 0.001745 5.905 6.34e-09 ***
            0.050674 0.028523 1.777 0.076210 .
## Race
## Occupation -0.006941 0.013095 -0.530 0.596309
              0.091013 0.038723 2.350 0.019125 *
## Sector
## Marr
               0.075125
                         0.041886 1.794 0.073458 .
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

## Residual standard error: 0.4397 on 524 degrees of freedom
## Multiple R-squared: 0.3175, Adjusted R-squared: 0.3058

#### round(vif(fit\_model2),3)

##	Education	South	Sex	Union	Age
##	1.126	1.046	1.088	1.121	1.154
##	Race	Occupation	Sector	Marr	
##	1.037	1.296	1.198	1.094	