## STAT 210 Applied Statistics and Data Analysis Nonparametric Regression

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So far, in regression, we have assumed that the relation between the output Y and the regressors  $\mathbf{X}$  is linear  $Y = \mathbf{X}\boldsymbol{\beta}$  and the parameters have been estimated using least squares

In nonparametric regression, we do not assume a predetermined form for the regression function. The shape of the regression function is determined from the data.

We will only review one case of nonparametric regression, kernel regression, linked to the density estimation procedures we studied. Also, we only look at simple regression

The regression function we want to estimate is the expected value of Y given the value of the regressor X=x. We will consider this as a function of x

$$m(x) = E(Y|X=x)$$

To simplify the presentation we will assume that both variables are continuous and have joint density f(x, y). The marginal densities for X and Y are denoted  $f_X(x)$  and  $f_Y(y)$ .

The probability distribution of Y given that X = x is characterized by the conditional density of Y given X = x, which is defined as

$$f_{Y|X=x}(y) = \frac{f(x,y)}{f_X(x)}.$$

Using this conditional density, we can write the regression function as

$$m(x) = E(Y|X = x) = \int y \, f_{Y|X=x}(y) \, dy$$

$$= \int y \, \frac{f(x,y)}{f_X(x)} \, dy$$

$$= \frac{1}{f_X(x)} \int y \, f(x,y) \, dy$$
(1)

Therefore, given a sample  $(X_1, Y_1), (X_2, Y_2), \ldots, (X_n, Y_n)$  we may obtain a nonparametric estimate of the regression function m by estimating densities  $f_X$  and f and using them in equation (1).

For the bivariate density f(x, y) we consider the kernel density estimator based on the product of univariate kernels with bandwidths  $\mathbf{h} = (h_1, h_2)$ .

$$\hat{f}(x,y;\mathbf{h}) = \frac{1}{n} \sum_{i=1}^{n} K_{h_1}(x-x_i) K_{h_2}(y-y_i)$$
 (2)

If we use the same bandwidth for the univariate density estimation

$$\hat{f}_X(x;h_1) = \frac{1}{n} \sum_{i=1}^n K_{h_1}(x - x_j)$$
 (3)

Replacing these equation in (1) we have

$$m(x) = \frac{1}{\hat{f}_{X}(x; h_{1})} \int y \hat{f}(x, y; \mathbf{h}) dy$$

$$= \frac{\int y \frac{1}{n} \sum_{i=1}^{n} K_{h_{1}}(x - x_{i}) K_{h_{2}}(y - y_{i}) dy}{\frac{1}{n} \sum_{j=1}^{n} K_{h_{1}}(x - x_{j})}$$

$$= \frac{\frac{1}{n} \sum_{i=1}^{n} K_{h_{1}}(x - x_{i}) \int y K_{h_{2}}(y - y_{i}) dy}{\frac{1}{n} \sum_{j=1}^{n} K_{h_{1}}(x - x_{j})}$$

$$= \frac{\frac{1}{n} \sum_{i=1}^{n} K_{h_{1}}(x - x_{i}) y_{i}}{\frac{1}{n} \sum_{j=1}^{n} K_{h_{1}}(x - x_{j})}$$

$$= \sum_{i=1}^{n} \frac{K_{h_{1}}(x - x_{i})}{\sum_{i=1}^{n} K_{h_{1}}(x - x_{j})} y_{i}$$

This is known as the Nadaraya-Watson estimator for the regression function.

Therefore

$$\widehat{m}(x;0,h) = \sum_{i=1}^{n} \frac{K_{h_1}(x-x_i)}{\sum_{j=1}^{n} K_{h_1}(x-x_j)} y_i$$

$$= \sum_{j=1}^{n} W_i^0(x) y_i$$
(4)

where

$$W_i^0(x) = \frac{K_{h_1}(x - x_i)}{\sum_{j=1}^n K_{h_1}(x - x_j)}$$

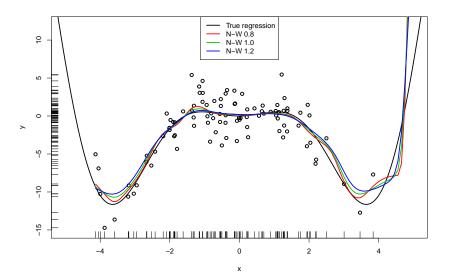
Thus, the Nadaraya - Watson estimator is a weighted average of the response values  $y_1, \ldots, y_n$  using weights  $\{W_i^0(x)\}$ , which depend upon the evaluation point.

In this sense the estimator is a local average of  $y_1, \ldots, y_n$  around X = x.

In R this nonparametric estimator of the regression function is
implemented in the function ksmooth() with the option kernel =
'normal'.

Let's see first a simulated example where the regression function is known.

```
n <- 100
m <- function(x) x^2 * cos(x)
x <- rnorm(n, sd=2); epsilon <- rnorm(n, sd=2)
y <- m(x) + epsilon
x_points <- seq(-10, 10, 1=500)</pre>
```



The Nadaraya -Watson is a particular case of a broader class of nonparametric estimators, known as the *local polynomial estimators*.

We want to find an estimator  $\widehat{m}$  of m that minimizes the residual sum of squares

$$\sum_{i=1}^{n} (y_i - \widehat{m}(x_i))^2$$
 (5)

without assuming a particular form for the true m.

In (simple) linear regression we assumed that  $m(x) = \beta_0 + \beta_1 x$ , so the model is parametrized by  $\beta = (\beta_0, \beta_1)$  and the minimization of (5) produced the least squares estimators  $\hat{\beta} = (\hat{\beta}_0, \hat{\beta}_1)$ 

When m has no available parametrization, we can introduce a local parametrization using a p-th order Taylor expansion around the observed points  $x_1, \ldots, x_n$ .

The degree p of the approximating polynomial is usually one or two.

For x close to  $x_i$ 

$$m(x) \approx m(x_i) + m'(x_i)(x - x_i) + \frac{1}{2}m''(x_i)(x - x_i)^2 + \cdots + \frac{1}{p!}m^{(p)}(x_i)(x - x_i)^p$$

The derivatives are, of course, unknown, but we can turn this into a regression problem by setting

$$\beta_j = \frac{1}{j!} m^{(j)}, j = 0, \dots, n$$

and now we want to determine the parameters  $(\beta_0, \beta_1, \dots, \beta_p)$ .

The final ingredient is to weigh each point's contribution according to the distance of  $x_i$  to x. For that, we use a kernel.

The estimated parameters are

$$\hat{\beta}_h = \arg\min \sum_{i=1}^n \left( y_i - \sum_{j=0}^p \beta_j (x - x_i)^j \right)^2 K_h(x - x_i)$$

If p = 0, this is the Nadaraya-Watson estimator.

In R, local polynomials are implemented in the function locpoly in the package KernSmooth.

The first link is to the chapter on nonparametric regression in Eduardo Garcia's book, Notes for Nonparametric Statistics.

https://bookdown.org/egarpor/NP-UC3M/kre-i-kre.html

The second link is to a Shiny app for nonparametric regression  $https://ec2-35-177-34-200.eu-west-\\ 2.compute.amazonaws.com/kreg/$ 

#### loess

Another smoother based on these ideas is loess in the base package. The smoothness in loess is determined by the proportion of points in the sample taken into consideration for the local fit around x. The default value for span is 0.75.

The contributions are then weighted using a triweight kernel instead of a normal kernel.

This is different from the polynomial regression we have reviewed but is based on the same principles.

We give an example using both functions and polynomials of degree 0 and 1.

The example comes from *Notes for Nonparametric Statistics* by Eduardo García, https://bookdown.org/egarpor/NP-UC3M/

```
set.seed(123456)
n < -100
eps \leftarrow rnorm(n, sd = 2)
m \leftarrow function(x) x^3 * sin(x)
X \leftarrow rnorm(n, sd = 1.5)
Y \leftarrow m(X) + eps
x \text{ grid} \leftarrow \text{seq}(-10, 10, 1 = 500)
# KernSmooth::locpoly fits
h < -0.25
lp0 <- KernSmooth::locpoly(x = X, y = Y, bandwidth = h, degree = 0,</pre>
                              range.x = c(-10, 10), gridsize = 500)
lp1 <- KernSmooth::locpoly(x = X, y = Y, bandwidth = h, degree = 1,</pre>
                              range.x = c(-10, 10), gridsize = 500)
# Provide the evaluation points by range.x and gridsize
# loess fits
span <- 0.25 # The default span is 0.75, which works very bad here
lo0 <- loess(Y ~ X, degree = 0, span = span)</pre>
lo1 <- loess(Y ~ X, degree = 1, span = span)</pre>
```

```
# Prediction at x = 0 and x = 2
# Prediction by locpoly
c(lp1$y[which.min(abs(lp1$x ))],lp1$y[which.min(abs(lp1$x - 2))])

## [1] 0.9694953 5.4459745
# Prediction by loess
predict(lo1, newdata = data.frame(X = c(0,2)))

## [1] 1.142635 5.379652
m(c(0,2)) # True regression

## [1] 0.000000 7.274379
```

