Supplemenentary code to reproduce the numerical results in Di Caterina and Kosmidis (2017)

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Workspace preparation

This document provides R (R Core Team 2017) code to reproduce the results in the manuscript 'Location-adjusted Wald statistic for scalar parameters' (Di Caterina and Kosmidis 2017).

This script assumes that the current working directory has the sub-directories `code`, `results` and `lesion data` as provided in the supplementary material.

```
path <- "."
code_path <- paste(path, "code", sep = "/")
results_path <- paste(path, "results", sep = "/")
lesions_path <- paste(path, "lesion data", sep = "/")
```

The contents of the directories are as follows

```
dir(code_path)
# [1] "babies_simulation.R"   "brains_case_study.R"   "clotting_simulation.R"
# [4] "dyslexia_simulation.R" "logodds_functions.R"   "overlay2_nifti.R"
dir(results_path)
# [1] "babies_simulation.rda"   "brains_case_study.rda"
# [3] "clotting_simulation.rda" "dyslexia_simulation.rda"
dir(lesions_path)
# [1] "data_demo.dat" "images"
```

First, make sure that you have the latest version of the **waldi** R package installed.

```
waldi_version <- try(packageVersion("waldi"), silent = TRUE)
if (inherits(waldi_version, "try-error")) {
    devtools::install_github("ikosmidis/waldi")
}
```

The following code chunk loads the required packages

```
library("waldi")
library("oro.nifti")
library("boot")
library("plyr")
library("plotrix")
library("dplyr")
library("survival")
library("cond")
library("lmtest")
library("betareg")
library("enrichwith")
library("brglm2")
library("ggplot2")
```

```
library("gridExtra")
library("colorspace")
```

Pre-saved R image files

Some of the code-chunks below load objects from the pre-saved R image files in the results directory. These image files are the outputs of the script `babies_simulation.R`, `brockwell_gordon_simulation.R`, `clotting_simulation.R`, `dyslexia_simulation.R`.

Table 1

```
data("ReadingSkills", package = "betareg")
## maximum likelihood estimates and corresponding 95\% Wald confidence intervals
rs_beta_ml <- betareg(accuracy ~ dyslexia * iq | dyslexia + iq,
                      data = ReadingSkills, type = "ML", hessian = FALSE)
rs_summary_ml <- coef(summary(rs_beta_ml))
rs_ml_estimates <- do.call("rbind", lapply(rs_summary_ml,
                                           function(z) z[, c("Estimate", "Std. Error")]))
rs_ml_cis <- confint(rs_beta_ml)
## bias corrected fit and corresponding 95\% Wald confidence intervals
rs_beta_br <- update(rs_beta_ml, type = "BR")
rs_summary_br <- coef(summary(rs_beta_br))
rs_br_estimates <- do.call("rbind", lapply(rs_summary_br,
                                           function(z) z[, c("Estimate", "Std. Error")]))
rs_br_cis <- confint(rs_beta_br)
round(cbind(rs_ml_estimates, rs_br_estimates, rs_ml_cis, rs_br_cis), 3)
#             Estimate Std. Error Estimate Std. Error  2.5 % 97.5 %  2.5 % 97.5 %
# (Intercept)    1.123      0.143    1.114      0.148  0.843  1.403  0.824  1.405
# dyslexia      -0.742      0.143   -0.734      0.148 -1.021 -0.462 -1.024 -0.444
# iq             0.486      0.133    0.441      0.141  0.225  0.747  0.165  0.717
# dyslexia:iq   -0.581      0.133   -0.532      0.140 -0.841 -0.321 -0.807 -0.257
# (Intercept)    3.304      0.223    3.092      0.225  2.868  3.741  2.652  3.533
# dyslexia       1.747      0.262    1.654      0.264  1.232  2.261  1.138  2.171
# iq             1.229      0.267    1.048      0.271  0.705  1.753  0.518  1.578
```

Table 2

`dyslexia_simulation.rda` contains the outputs of `dyslexia_simulation.R` in ./code, which replicates the simulation study described in Example 1.1 of Di Caterina and Kosmidis (2017)

```
load(paste(results_path, paste0("dyslexia_simulation.rda"), sep = "/"))
rs_coverage <- results %>%
    filter(parameter %in% c("dyslexia", "iq", "dyslexia:iq", "(phi)_dyslexia", "(phi)_iq")) %>%
    filter(statistic %in% c("ml", "br", "ml_cor", "br_cor",
                            "ml_stud", "ml_cor_stud", "br_stud", "br_cor_stud",
                            "ml_cor_ses_cor", "br_cor_ses_cor")) %>%
    mutate(parameter = recode(parameter,
                              "dyslexia" = 2, "iq" = 3, "dyslexia:iq" = 4,
```

```
                              "(phi)_dyslexia" = 6, "(phi)_iq" = 7)) %>%
    mutate(level = 100 * level) %>%
    group_by(level, statistic, parameter) %>%
    summarize(coverage = round(mean(cover, na.rm = TRUE) * 100, 1)) %>%
    as.data.frame() %>%
    reshape(idvar = c("statistic", "parameter"), v.names = "coverage",
            timevar = "level",
            direction = "wide")
rs_coverage %>% filter(statistic %in% c("ml", "br")) %>%
    select(statistic, parameter, coverage.90, coverage.95, coverage.99)
#    statistic parameter coverage.90 coverage.95 coverage.99
# 1         br         2        88.1        93.4        98.2
# 2         br         3        87.2        92.9        98.0
# 3         br         4        87.3        92.9        98.0
# 4         br         6        83.8        90.2        96.7
# 5         br         7        82.7        89.2        96.1
# 6         ml         2        86.9        92.4        97.7
# 7         ml         3        84.8        91.0        97.1
# 8         ml         4        85.0        91.2        97.2
# 9         ml         6        82.4        89.1        96.1
# 10        ml         7        79.1        86.0        94.4
```

Figure 1

```
rs_cor_ml_cis <- waldi_confint(rs_beta_ml, level = 0.95, adjust = TRUE)
interpolation <- waldi_confint(rs_beta_ml, level = 0.95,
                               which = rownames(rs_cor_ml_cis),
                               adjust = TRUE,
                               return_values = TRUE,
                               length = 20)
intervals <- data.frame(low = rs_cor_ml_cis[, 1],
                        upp = rs_cor_ml_cis[, 2],
                        parameter = rownames(rs_cor_ml_cis))
interpolation <- interpolation %>%
    filter(!(parameter %in% c("(Intercept)", "(phi)_(Intercept)"))) %>%
    mutate(parameter = recode(parameter,
                              "dyslexia" = "beta[2]",
                              "iq" = "beta[3]",
                              "dyslexia:iq" = "beta[4]",
                              "(phi)_dyslexia" = "gamma[2]",
                              "(phi)_iq" = "gamma[3]"))
intervals <- intervals %>%
    filter(!(parameter %in% c("(Intercept)", "(phi)_(Intercept)"))) %>%
    mutate(parameter = recode(parameter,
                              "dyslexia" = "beta[2]",
                              "iq" = "beta[3]",
                              "dyslexia:iq" = "beta[4]",
                              "(phi)_dyslexia" = "gamma[2]",
                              "(phi)_iq" = "gamma[3]"))
ggplot(interpolation) +
    geom_point(aes(x = grid, y = value)) +
```

```
    geom_line(aes(x = grid, y = value), col = "grey") +
    geom_hline(aes(yintercept = qnorm(0.975)), col = "grey", lty = 3) +
    geom_hline(aes(yintercept = qnorm(0.025)), col = "grey", lty = 3) +
    geom_vline(data = intervals, aes(xintercept = low), col = "grey", lty = 2) +
    geom_vline(data = intervals, aes(xintercept = upp), col = "grey", lty = 2) +
    facet_grid(~ parameter, scale = "free_x", labeller = "label_parsed") +
    theme_minimal() +
    theme(axis.text.x = element_text(size = 7)) +
    labs(x = "parameter value", y = "statistic")
```

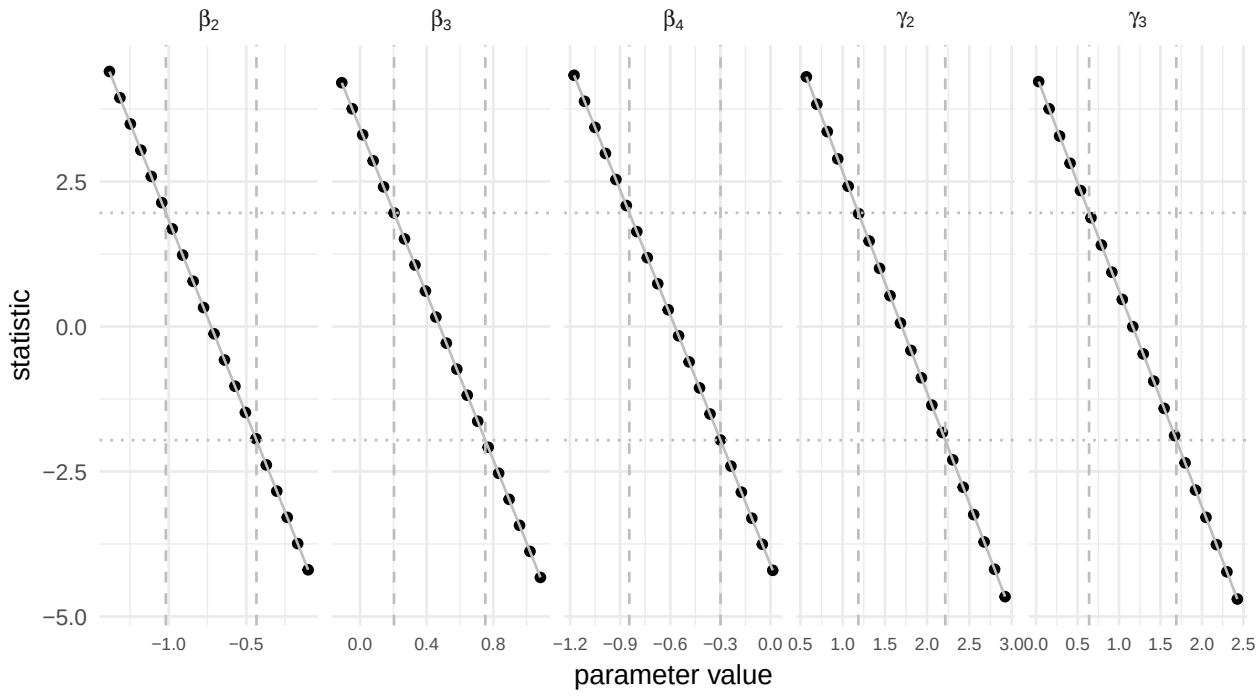


Table 3

```
## Confidence intervals based on the location-adjusted Wald statistic
rs_cor_ml_cis <- waldi_confint(rs_beta_ml, level = 0.95, adjust = TRUE, parallel = FALSE)
rs_cor_br_cis <- waldi_confint(rs_beta_br, level = 0.95, adjust = TRUE, parallel = FALSE)
## Studentized bootstrap intervals
set.seed(123)
quantiles_ml <- dyslexia_bootstrap(rs_beta_ml, R = 500, ncores = 1)$quantiles
quantiles_br <- dyslexia_bootstrap(rs_beta_br, R = 500, ncores = 1)$quantiles
rs_ml_stud_cis <- waldi_confint(rs_beta_ml, adjust = FALSE, parallel = FALSE,
                                quantiles = quantiles_ml$zstat[, c("0.025", "0.975")])
rs_br_stud_cis <- waldi_confint(rs_beta_br, adjust = FALSE, parallel = FALSE,
                                quantiles = quantiles_br$zstat[, c("0.025", "0.975")])
rs_cor_ml_stud_cis <- waldi_confint(rs_beta_ml, adjust = TRUE, parallel = FALSE,
                                    quantiles = quantiles_ml$zstat_cor[, c("0.025", "0.975")])
rs_cor_br_stud_cis <- waldi_confint(rs_beta_br, adjust = TRUE, parallel = FALSE,
                                    quantiles = quantiles_br$zstat_cor[, c("0.025", "0.975")])

round(rbind(cbind(rs_cor_ml_cis, rs_cor_br_cis),
```

```
            cbind(rs_cor_ml_stud_cis, rs_cor_br_stud_cis)), 3)
#                   2.5 % 97.5 %  2.5 % 97.5 %
# (Intercept)        0.816  1.400  0.827  1.411
# dyslexia          -1.019 -0.435 -1.031 -0.446
# iq                 0.204  0.752  0.165  0.719
# dyslexia:iq       -0.845 -0.299 -0.809 -0.257
# (phi)_(Intercept)  2.689  3.564  2.652  3.532
# (phi)_dyslexia     1.186  2.214  1.134  2.169
# (phi)_iq           0.639  1.691  0.513  1.574
# (Intercept)        0.812  1.420  0.830  1.481
# dyslexia          -1.059 -0.442 -1.091 -0.440
# iq                 0.171  0.792  0.159  0.758
# dyslexia:iq       -0.871 -0.268 -0.853 -0.264
# (phi)_(Intercept)  2.709  3.680  2.654  3.682
# (phi)_dyslexia     1.112  2.303  1.040  2.241
# (phi)_iq           0.565  1.835  0.394  1.769
rs_coverage %>% filter(statistic %in% c("ml_cor", "br_cor")) %>%
    select(statistic, parameter, coverage.90, coverage.95, coverage.99)
#    statistic parameter coverage.90 coverage.95 coverage.99
# 1     br_cor         2        88.3        93.5        98.3
# 2     br_cor         3        87.3        93.0        98.0
# 3     br_cor         4        87.5        93.0        98.0
# 4     br_cor         6        83.9        90.3        96.8
# 5     br_cor         7        82.7        89.2        96.2
# 6     ml_cor         2        88.5        93.7        98.4
# 7     ml_cor         3        87.1        92.8        98.0
# 8     ml_cor         4        87.2        92.8        98.0
# 9     ml_cor         6        83.5        90.0        96.6
# 10    ml_cor         7        81.8        88.6        95.7
rs_coverage %>% filter(statistic %in% c("ml_cor_stud", "br_cor_stud")) %>%
    select(statistic, parameter, coverage.90, coverage.95, coverage.99)
#      statistic parameter coverage.90 coverage.95 coverage.99
# 1  br_cor_stud         2        89.4        94.6        98.6
# 2  br_cor_stud         3        89.4        94.5        98.5
# 3  br_cor_stud         4        89.5        94.4        98.6
# 4  br_cor_stud         6        90.1        94.9        98.8
# 5  br_cor_stud         7        90.5        95.1        98.8
# 6  ml_cor_stud         2        89.5        94.5        98.7
# 7  ml_cor_stud         3        89.2        94.3        98.5
# 8  ml_cor_stud         4        89.3        94.3        98.5
# 9  ml_cor_stud         6        89.9        94.7        98.7
# 10 ml_cor_stud         7        90.1        94.9        98.7
```

The following chunk of code reproduces the times for the computation of the confidence intervals reports in Section~6.

```
## Intervals based on the location-adjusted Wald statistic
system.time({
    waldi_confint(rs_beta_ml, adjust = TRUE, parallel = FALSE, length = 5)
})
#    user  system elapsed
#   1.350   0.027   1.380
simu_fun <- get_simulate_function(rs_beta_ml)
generate_dyslexia <- function(data, mle) {
```

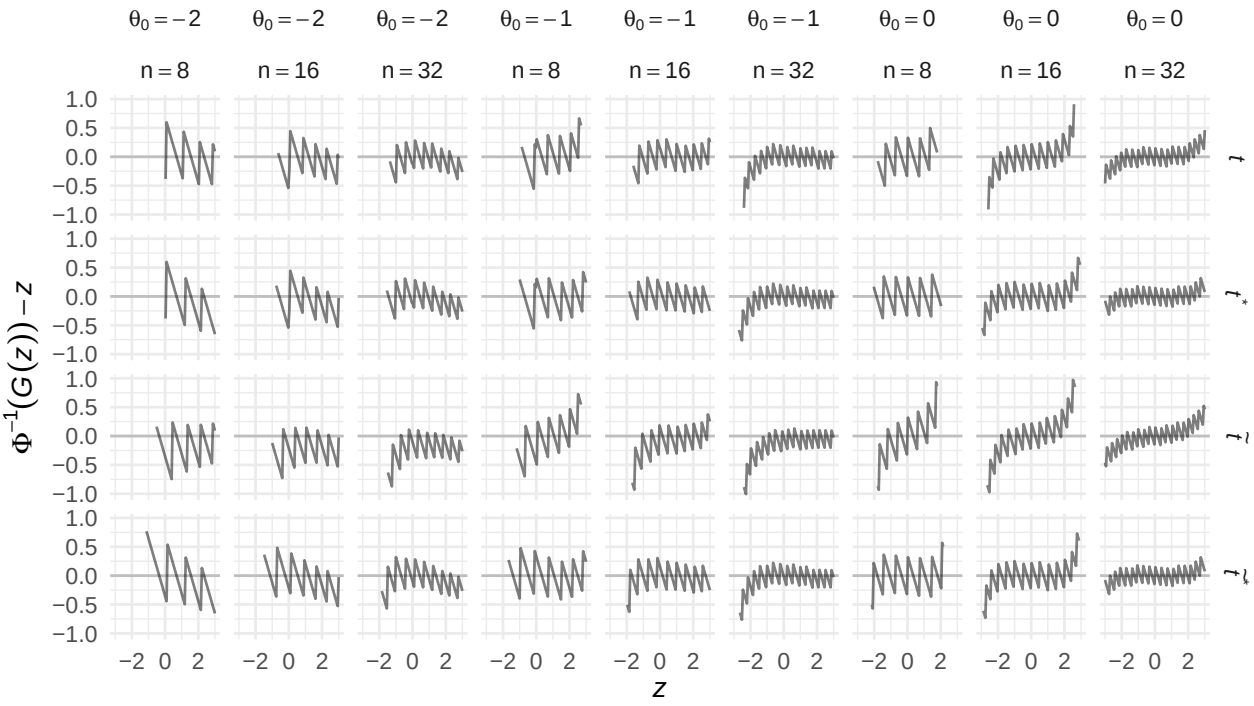
```
    simu_fun(mle)
}
stat <- function(data, psi) {
    temp <- ReadingSkills
    temp$accuracy <- data
    temp_fit <- try(update(rs_beta_ml, data = temp))
    if (inherits(temp_fit, "try-error")) {
        rep(NA, 7)
    }
    else {
        waldi(temp_fit, null = psi, adjust = TRUE)
    }
}
## Studentized bootstrap intervals
system.time({
    stats <- boot(ReadingSkills$accuracy, statistic = stat,
                  R = 500, sim = "parametric", ran.gen = generate_dyslexia,
                  mle = coef(rs_beta_ml), psi = coef(rs_beta_ml), ncpus = 1)$t
    quant <- t(apply(stats, 2, quantile, probs = c(0.025, 0.975), na.rm = TRUE))
    waldi_confint(rs_beta_br, adjust = TRUE, parallel = FALSE,
                  quantiles = quant)
})
#    user  system elapsed
# 135.690   3.006 139.245
```

Figure 2

```
source(paste0(code_path, "/", "logodds_functions.R"))
## Distribution of the statistic against normal
settings <- expand.grid(m = c(8, 16, 32), theta0 = c(-2, -1, 0))
plot_data <- NULL
for (j in seq.int(nrow(settings))) {
    setting <- settings[j, ]
    z <- seq(-3, 3, length = 100)
    dat <- t(sapply(z, dist_function, n = setting$m, theta0 = setting$theta0))
    dd <- stack(as.data.frame(dat))
    dd$z <- z
    names(dd) <- c("prob", "method", "z")
    dd$theta0 <- setting$theta0
    dd$m <- setting$m
    plot_data <- rbind(plot_data, dd)
}
plot_data$theta0 <- paste0("theta[0] == ", plot_data$theta0)
plot_data$theta0 <- factor(plot_data$theta0, levels = unique(plot_data$theta0),
                           ordered = TRUE)
plot_data$m <- paste0("n == ", plot_data$m)
plot_data$m <- factor(plot_data$m, levels = unique(plot_data$m), ordered = TRUE)
plot_data$method <- factor(plot_data$method, levels = c("ml", "a_ml", "br", "a_br"),
                           ordered = TRUE)
plot_data$method <- recode(plot_data$method,
                           "ml" = "italic(t)",
```

```
                            "a_ml" = "italic(t)^{'*'}",
                            "br" = "tilde(italic(t))",
                            "a_br" = "tilde(italic(t))^{'*'}")
ggplot(plot_data) +
    geom_abline(aes(intercept = 0, slope = 0), col = "grey") +
    geom_line(aes(z, qnorm(prob) - z), alpha = 0.5) +
    facet_grid(method ~ theta0 + m, label = label_parsed) +
    theme_minimal() +
    labs(y = expression(paste(Phi^list(-1),(italic(G)(italic(z)))-italic(z))),
         x = expression(italic(z))) +
    theme(text=element_text(size = 11))
# Warning in qnorm(prob): NaNs produced

# Warning in qnorm(prob): NaNs produced
# Warning: Removed 50 rows containing missing values (geom_path).
```

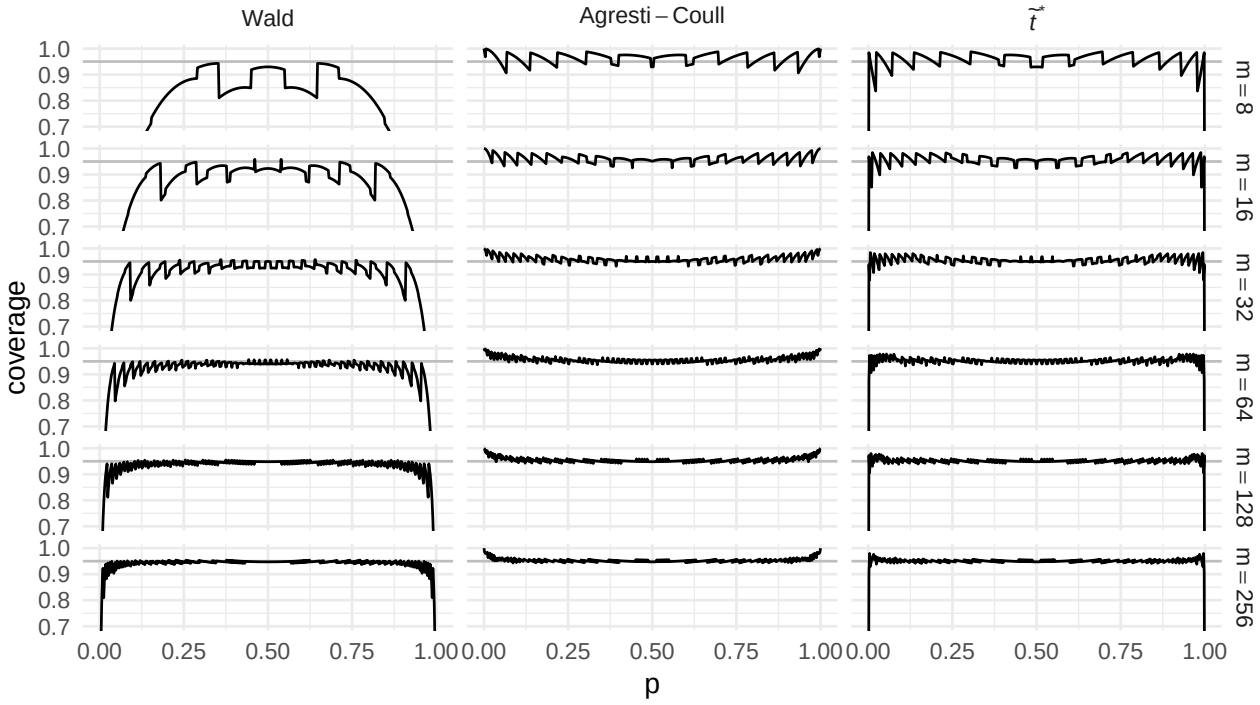


Coverage and length of confidence intervals for a binomial proportion

This section provides evidence for the stated coverage and expected length properties of confidence intervals for a binomial proportion in Section 7 of the main text. The code chunk below computes and visualised the coverage and expected length of the 95% confindence intervals $\bar{y} \pm z_{0.975}\sqrt{\bar{y}(1-\bar{y})/n}$ (Wald), $\tilde{p} \pm z_{0.975}\sqrt{\tilde{p}(1-\tilde{p})/(n+4)}$, where $\tilde{p} = (\sum y_i + 2)/(n+4)$ (Agresti-Coull; Agresti and Coull (1998) and Agresti and Caffo (2000)), and the intervals based on the transformation of the endpoints of the confidence intervals for the log-odds based on $\tilde{t}^*$.

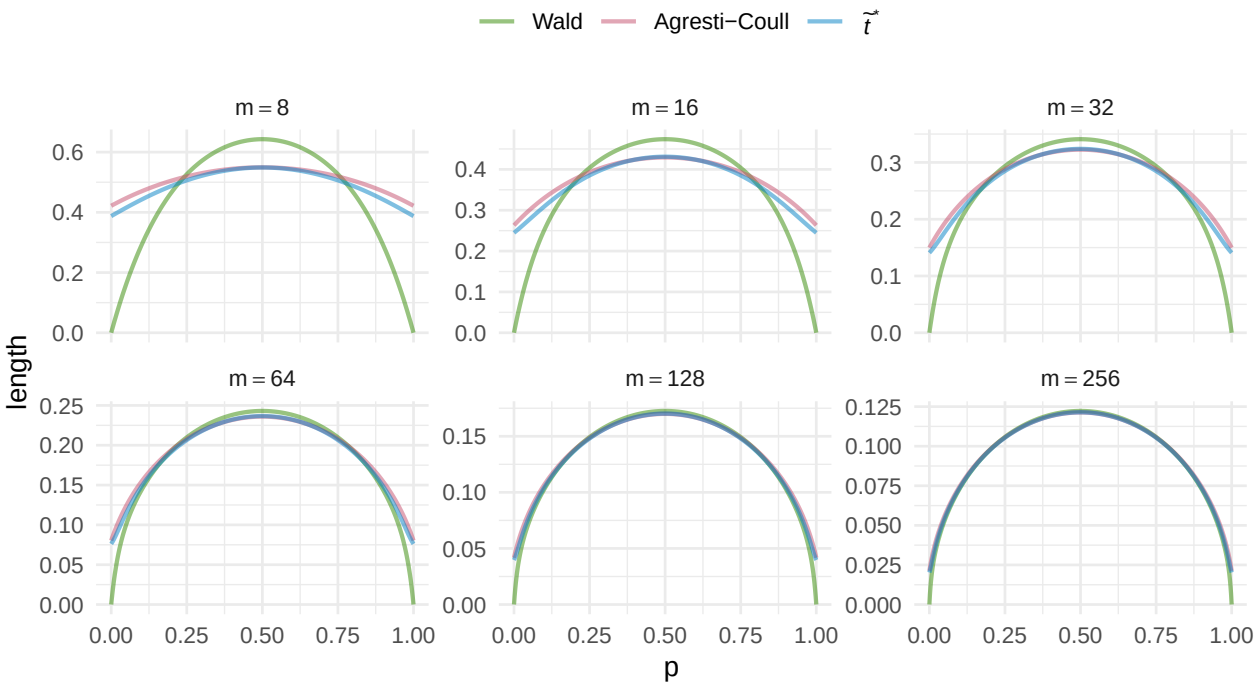
```
probs <- seq(1e-08, 1 - 1e-08, length = 500)
df <- ddply(data.frame(m = c(8, 16, 32, 64, 128, 256)), ~ m, function(x) {
    m <- x$m
```

```
    cis <- compute_cis(m, level = 0.95)
    cc <- lapply(probs, function(pp) cover_ci_prop(n = m, p = pp, level = 0.95, cis = cis))
    do.call("rbind", cc)
})
df$m <- factor(paste("m ==", df$m), levels = paste("m ==", sort(unique(df$m))),
               ordered = TRUE)
df$method <- factor(df$method, levels = c("wald", "ac", "a_br", "ml", "a_ml", "br"),
                    ordered = TRUE)
df$method <- recode(df$method,
                    "wald" = "Wald",
                    "ml" = "italic(t)[trans]",
                    "a_ml" = "italic(t)^*[trans]",
                    "br" = "tilde(italic(t))[trans]",
                    "a_br" = "tilde(italic(t))^{'*'}",
                    "ac" = "Agresti-Coull")
## coverage
ggplot(df %>% filter(method %in% c("Wald", "Agresti-Coull", "tilde(italic(t))^{'*'}"))) +
    geom_hline(aes(yintercept = 0.95), col = "grey") +
    geom_line(aes(x = p, y = coverage)) +
    facet_grid(m ~ method, label = label_parsed) +
    coord_cartesian(ylim = c(0.7, 1)) +
    theme_minimal()
```




```
## expected length
ggplot(df %>% filter(method %in% c("Wald", "Agresti-Coull", "tilde(italic(t))^{'*'}"))) +
    geom_line(aes(x = p, y = length, col = method), alpha = 0.5, size = 0.8) +
    facet_wrap(~ m, label = label_parsed, scales = "free_y", ncol = 3) +
    scale_colour_manual(values = c("#328900", "#C54E6D", "#0080C5"),
                        name = "",
                        labels = c("Wald",
                                   "Agresti-Coull",
```

```
                                bquote(tilde(italic(t))^{'*'}))) +
  theme_minimal() +
  theme(legend.position = "top")
```



Hauck and Donner effect

```
sapply(28:32, t_ml, n =  32, theta0 = 0)
# [1] 3.640465 3.740749 3.708150 3.379905 0.000000
sapply(28:32, t_adjusted_ml, n =  32, theta0 = 0)
# [1] 3.770481 3.912737 3.955360 3.816022 0.000000
sapply(28:32, t_br, n =  32, theta0 = 0)
# [1] 3.583279 3.712935 3.744298 3.587411 2.884566
sapply(28:32, t_adjusted_br, n =  32, theta0 = 0)
# [1] 3.763721 3.902155 3.935838 3.762302 2.921237
```

Table 4

```
## The clotting data set
clotting <- data.frame(
  conc = c(118,58,42,35,27,25,21,19,18,69,35,26,21,18,16,13,12,12),
  u = c(5,10,15,20,30,40,60,80,100, 5,10,15,20,30,40,60,80,100),
  lot = factor(c(rep(1, 9), rep(2, 9))))
## The maximum likelihood fit of the gamma regression model
clotting_ml <- glm(conc ~ log(u)*lot, data = clotting, family = Gamma(link = "log"))
## Maximum likelihood estimates and Wald statistics using maximum likelihood estimator
## of the dispersion parameter
dispersion_ml <- MASS::gamma.dispersion(clotting_ml)
```

```
clotting_summary_ml <- summary(clotting_ml, dispersion = dispersion_ml)
clotting_ml_estimates <- coef(clotting_summary_ml)[, c("Estimate", "z value")]
## Reduced-bias estimates and Wald statistics
clotting_summary_rb <- summary(update(clotting_ml, method = "brglmFit"))
## Maximum likelihood estimates and Wald statistics using the moment-based estimator
## of the dispersion parameter
clotting_summary_mom <-  summary(clotting_ml)
dispersion_mom <- clotting_summary_mom$dispersion
clotting_mom_estimates <- coef(clotting_summary_mom)[, c("Estimate", "t value")]
## Location-adjusted Wald statistic
clotting_waldi <- waldi(clotting_ml, null = 0, adjust = TRUE)
round(cbind(c(clotting_ml_estimates[, 1], dispersion_ml, dispersion_mom),
            c(clotting_ml_estimates[, 2], NA, NA),
            c(clotting_mom_estimates[, 2], NA, NA),
            c(clotting_waldi, NA, NA)), 3)
#                [,1]    [,2]    [,3]    [,4]
# (Intercept)  5.503  34.124  29.282  28.953
# log(u)      -0.602 -12.842 -11.020 -10.896
# lot2        -0.584  -2.563  -2.199  -2.173
# log(u):lot2  0.034   0.520   0.446   0.441
#              0.017      NA      NA      NA
#              0.024      NA      NA      NA
```

Figure 3 including rejection probabilities based on t_j^* and the Wald statistic using $\tilde{\phi}$

`clotting_simulation.rda` below is the output of `clotting_simulation.R` in ./code, which replicates the simulation study described in Section 8.3 of Di Caterina and Kosmidis (2017).

```
load(paste(results_path, "clotting_simulation.rda", sep = "/"))
## Compute type I error rates
typeI <- ddply(res, ~ statistic + parameter, function(x) {
    ## empirical <- pnorm(quantile(x$value, c(0, 1, 2.5, 5, 1)/100))
    levels <- c(0.1, 1, 2.5, 5)/100
    p_value_2sided <- 2 * pnorm(-abs(x$value))
    p_value_left <- pnorm(x$value)
    p_value_right <- 1 - pnorm(x$value)
    rate_2sided <- sapply(levels, function(alpha) mean(p_value_2sided < alpha))
    rate_left <- sapply(levels, function(alpha) mean(p_value_left < alpha))
    rate_right <- sapply(levels, function(alpha) mean(p_value_right < alpha))
    out <- data.frame(
        test = rep(c("2sided", "left", "right"), each = length(levels)),
        typeI = c(rate_2sided, rate_left, rate_right),
        level = rep(levels, times = 3))
    out
})
typeI <- typeI %>%
    filter(test != "right") %>%
    mutate(test = recode(test,
                         "2sided" = "beta[italic(j)] != beta[paste(italic(j), 0)]",
                         "left" = "beta[italic(j)] < beta[paste(italic(j), 0)]",
```

```
                          "right" = "beta[italic(j)] > beta[paste(italic(j), 0)]"),
           level_chr = paste(level*100, "~symbol('\045')"),
           upper = typeI - qnorm(1 - 0.01/2)*sqrt(typeI*(1-typeI)/nsimu),
           lower = typeI + qnorm(1 - 0.01/2)*sqrt(typeI*(1-typeI)/nsimu))
## Figure 2 in the manuscript
ggplot(typeI %>% filter(parameter != 1)) +
    geom_point(aes(parameter, typeI, pch = statistic), alpha = 0.7) +
    geom_hline(aes(yintercept = level), col = "grey", lty = 2) +
    facet_grid(test ~ level_chr, labeller = label_parsed, scales = "free") +
    scale_x_continuous(name = element_blank(),
                       breaks = c(2, 3, 4),
                       limits = c(1.8, 4.2),
                       labels = c(
                           expression(beta[2]),
                           expression(beta[3]),
                           expression(beta[4]))) +
    scale_y_continuous(name = expression(paste("Empirical rejection probability (",
                                               symbol('\045'), ")")),
                       labels = function (x) {
                           if (length(x) == 0)
                               return(character())
                           x <- round_any(x, scales:::precision(x)/100)
                           scales:::comma(x * 100)
                       }) +
    theme_bw() +
    theme(legend.position = "top",
          panel.grid.major.y = element_blank(),
          panel.grid.minor.y = element_blank(),
          panel.grid.minor.x = element_blank(),
          strip.background = element_blank())
```

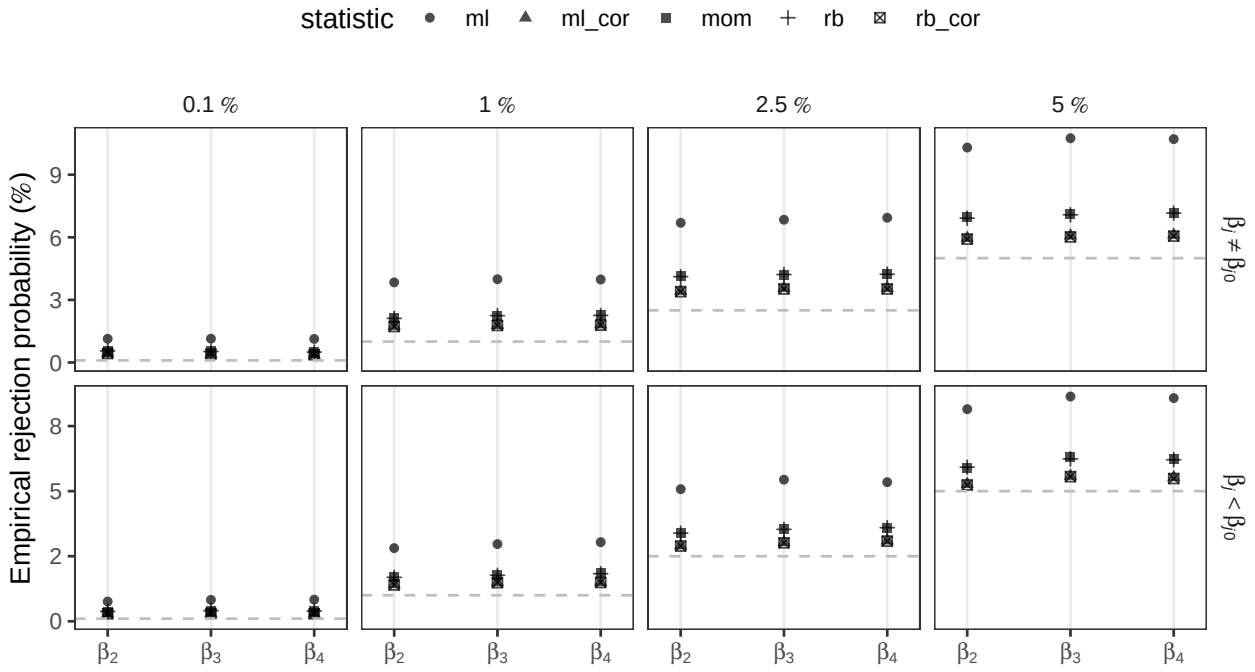


Table 5

```
data("babies", package = "cond")
## clogit understands only 0-1 so expand
babies_expand <- ddply(babies, ~ lull + day, function(z) {
    data.frame(y = rep(c(0, 1), c(z$r2, z$r1)))
})
## Maximum likelihood fit
babies_ml <- glm(formula = y ~ day + lull - 1,
                 family = binomial, data = babies_expand)
babies_rb <- update(babies_ml, method = "brglmFit")
## Maximum conditional likelihood fit
babies_cond <- clogit(y ~ strata(day) + lull, data = babies_expand)
ml <- coef(summary(babies_ml))["lullyes", ]
rb <- coef(summary(babies_rb))["lullyes", ]
mcl <- coef(summary(babies_cond))["lullyes", ]
r <- lrtest(update(babies_ml, . ~ . - lull),
            babies_ml)
rc <- summary(babies_cond)$logtest[1]
scorec <- summary(babies_cond)$sctest[1]
out1 <- c(
    ml = unname(ml["Estimate"]),
    rb = unname(rb["Estimate"]),
    mcl = unname(mcl["coef"]),
    wald_ml = unname(ml["z value"]),
    wald_mcl = unname(mcl["z"]),
    wald_rb = unname(rb["z value"]),
    r = unname(sign(ml["Estimate"]) * sqrt(r$Chisq[2])),
    rc = unname(sign(mcl["coef"]) * sqrt(rc)),
    wald_ml_adjusted = unname(waldi(babies_ml, which = 19)),
    wald_rb_adjusted = unname(waldi(babies_rb, which = 19)))
out2 <- c(
    ml_se = unname(ml["Std. Error"]),
    rb_se = unname(rb["Std. Error"]),
    mcl_se = unname(mcl["se(coef)"]),
    ml_p = ml["Pr(>|z|)"],
    mcl_p = mcl["Pr(>|z|)"],
    rb_p = rb["Pr(>|z|)"],
    r_p = 2 * pnorm(-abs(out1["r"])),
    rc_p = 2 * pnorm(-abs(out1["rc"])),
    cor_ml_p = 2 * pnorm(-abs(out1["wald_ml_adjusted"])),
    cor_rb_p = 2 * pnorm(-abs(out1["wald_rb_adjusted"])))
round(matrix(c(out1, out2), ncol = 10, byrow = TRUE,
             dimnames = list(NULL,
                             c("mle", "rb", "mcle", "wald_ml", "wald_mlc",
                               "wald_rb", "r", "rc", "wald_ml_adjusted",
                               "wald_rb_adjusted"))), 4)
#         mle     rb   mcle wald_ml wald_mlc wald_rb      r     rc
# [1,] 1.4324 1.1562 1.2561  1.9511   1.8307  1.7362 2.1596 2.0214
# [2,] 0.7341 0.6659 0.6861  0.0510   0.0671  0.0825 0.0308 0.0432
#      wald_ml_adjusted wald_rb_adjusted
# [1,]           1.9257           1.9064
# [2,]           0.0541           0.0566
```

Figure 4

`babies_simulation.rda` below is the output of `babies_simulation.R` in ./code, which replicates the simulation study described in Section 8.4 of Di Caterina and Kosmidis (2017)

```
load(paste(results_path, "babies_simulation.rda", sep = "/"))

## The bootstrap p-value for the babies data is
set.seed(123)
babies_bootstrap(babies_ml, R = 1000)$conv
# [1] 0.0230001
## Compute pvalues from the various statistics account for the existence of bootstrap
## p-values
pval <- ddply(res %>% filter(!infinite & !is.na(value) & type != "summary"),
              ~ name,
              function(data) {
    if (all(data$type == "bootstrap_statistic")) {
        data.frame(sample = pnorm(data$value),
                   test = gsub("boot_prep_|boot_conv_", "", data$name))
    }
    else {
        p2 <- 2 * pnorm(-abs(data$value))
        pl <- pnorm(data$value)
        pr <- 1 - pl
        data.frame(sample = c(p2, pl, pr),
                   test = rep(c("2sided", "left", "right"), each = length(p2))) }
})
## Get rid of left right 2sided from statistic names
pval <- pval %>% mutate(name = gsub("_left|_right|_2sided", "", name))
pval <- pval %>%
    filter(!(name %in% c("scorec", "boot_prep")) & test != "right") %>%
    mutate(test = dplyr::recode(test,
                                "2sided" = "gamma != 0",
                                "left" = "gamma < 0",
                                "right" = "gamma > 0"),
           name = factor(name,
                         levels = c("mle", "rbe", "r", "cond", "scorec", "rc",
                                    "boot_conv", "cor", "cor_rb"),
                         ordered = TRUE)) %>%
    mutate(name = factor(name,
                         levels = c("mle", "r", "boot_conv", "rbe",
                                    "cond", "scorec", "rc",
                                    "cor", "cor_rb"),
                         ordered = TRUE)) %>%
    mutate(statistic = dplyr::recode(name,
                                     "mle" = "italic(t)",
                                     "rbe" = "italic(tilde(t))",
                                     "r" = "italic(r)",
                                     "cond" = "italic(t)[c]",
                                     "scorec" = "italic(s)[c]",
                                     "rc" = "italic(r)[c]",
                                     "cor" = "italic(t)^'*'",
                                     "cor_rb" = "tilde(italic(t))^'*'",
                                     "boot_conv" = "italic(boot)"))
```

```
## Bin sample
breaks <- (0:20)/20
pval <- pval %>%
    group_by(statistic, test) %>%
    mutate(sample = cut(sample, breaks = breaks, include.lowest = TRUE)) %>%
    group_by(statistic, test, sample)
ggplot(pval) +
    geom_hline(aes(yintercept = 1)) +
    geom_bar(aes(x = sample, y = ..count../2500), fill = "darkgray", alpha = 0.5) +
    facet_grid(test ~ statistic, labeller = label_parsed) +
    theme_bw() +
    scale_x_discrete(breaks = c("[0,0.05]", "(0.25,0.3]", "(0.5,0.55]",
                                "(0.75,0.8]", "(0.95,1]"),
                     labels = c(0, 0.25, 0.5, 0.75, 1)) +
    theme(legend.position = "top",
          panel.grid.major.y = element_blank(),
          panel.grid.minor.y = element_blank(),
          panel.grid.minor.x = element_blank(),
          panel.grid.major.x = element_blank(),
          strip.background = element_blank(),
          axis.text.x = element_text(angle = 45, hjust = 1, size = 8)) +
    labs(x = expression(paste(italic(p), "-value ")), y = "Density")
```

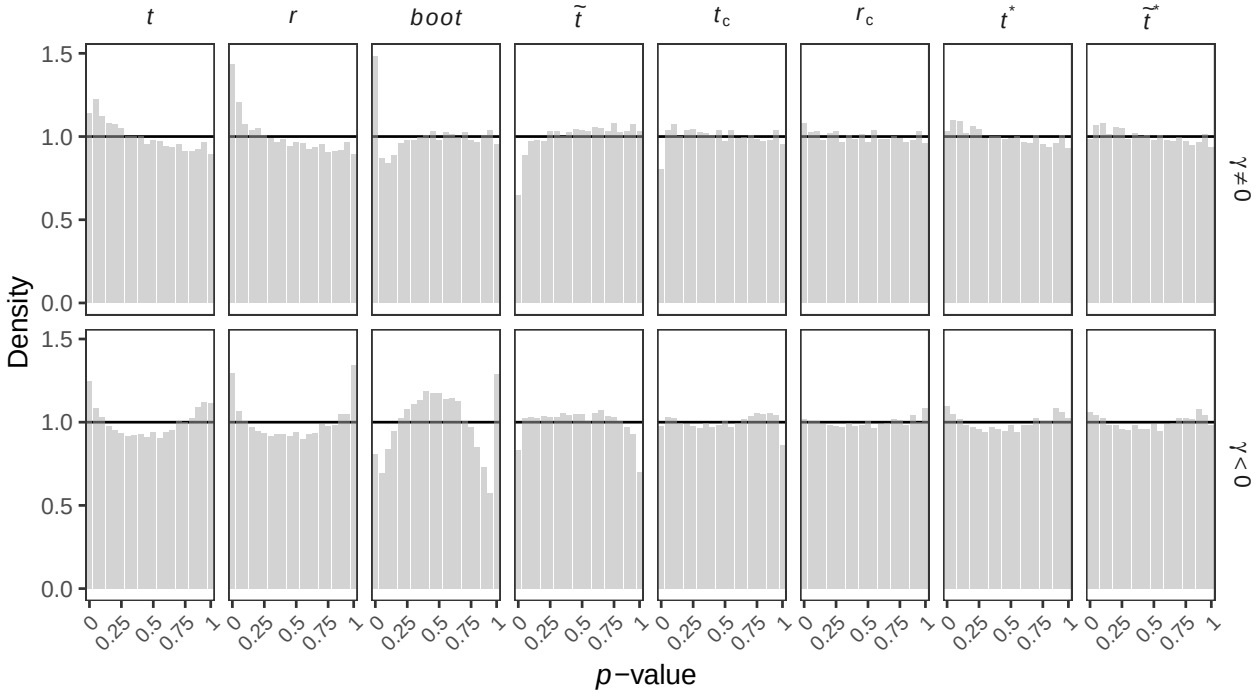
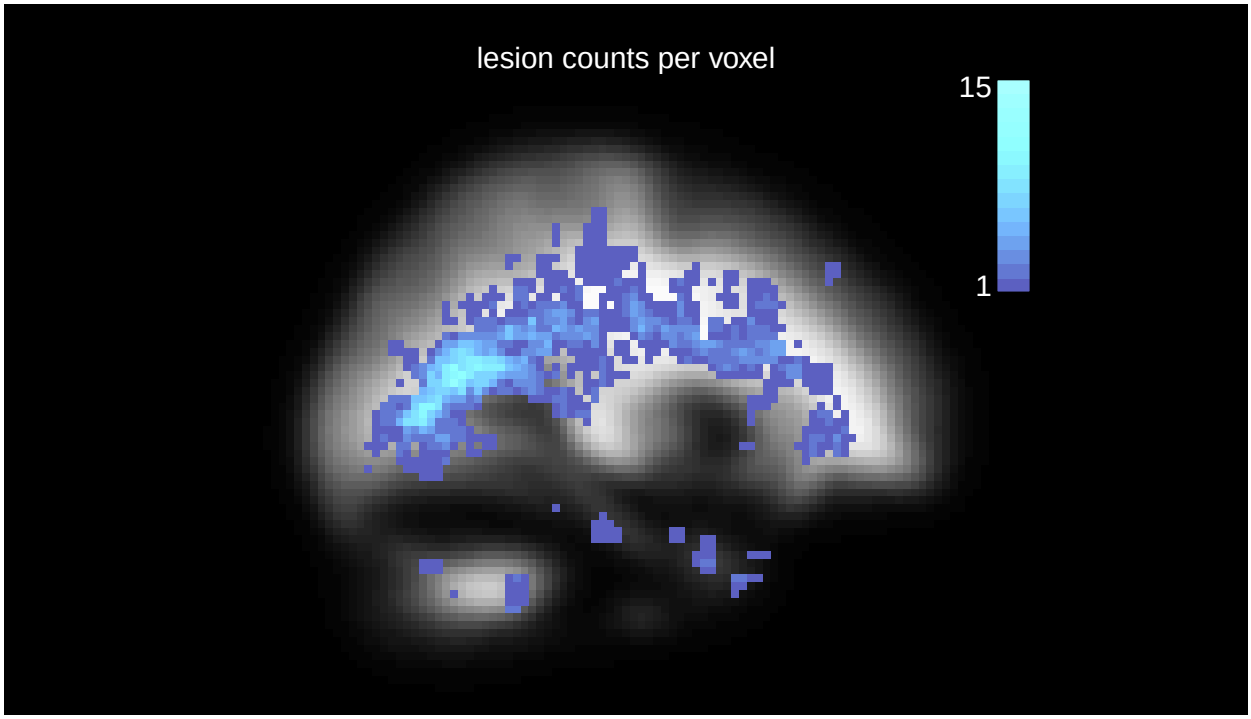


Figure 5

`brains_case_study.rda` below is the output of `brains_case_study.R` in ./code, which replicates the simulation study described in Section 9 of Di Caterina and Kosmidis (2017)

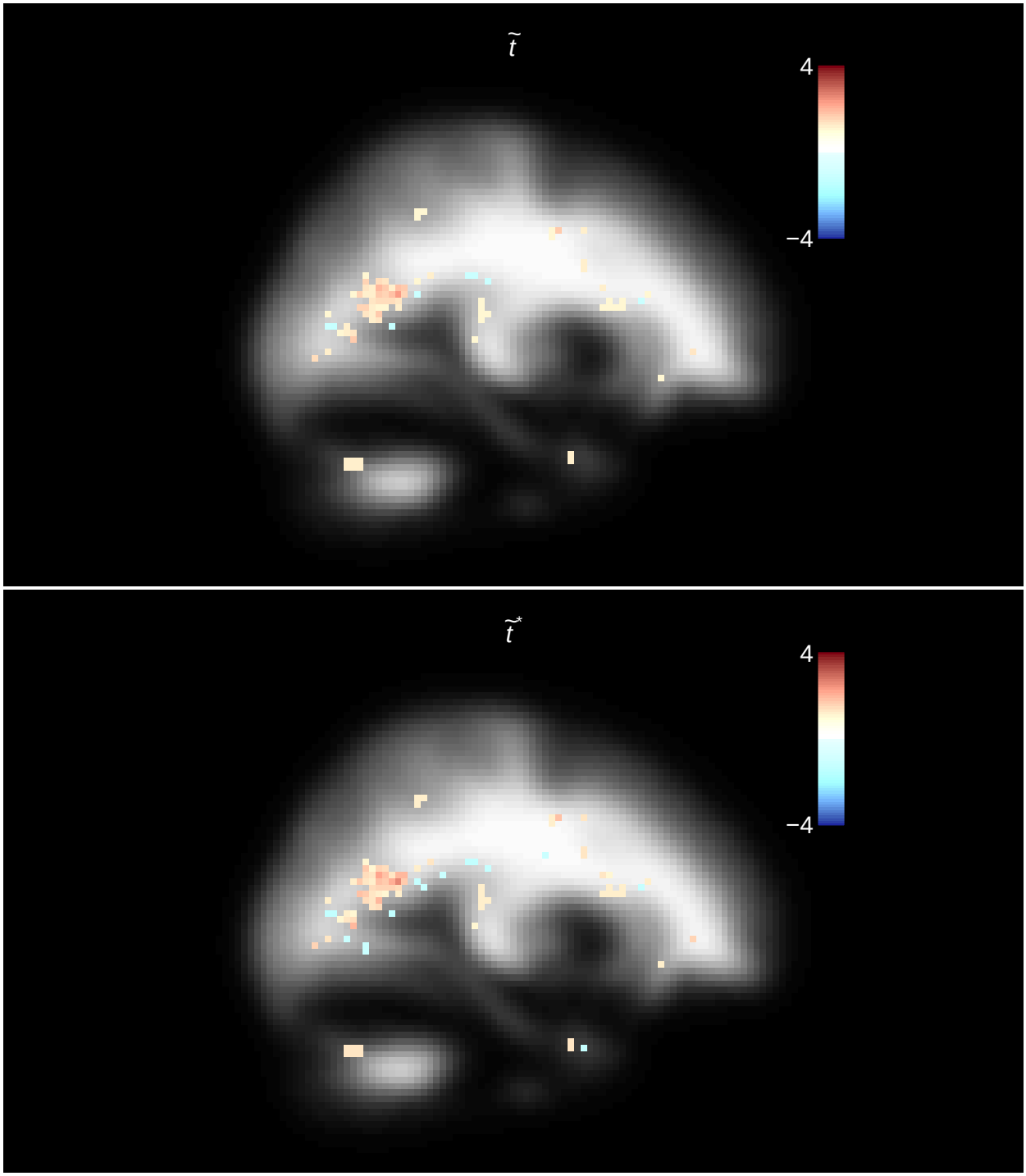
```
source(paste0(code_path, "/", "overlay2_nifti.R"))
load(paste(results_path, "brains_case_study.rda", sep = "/"))
```

```r
## Check how many times LR failed, excluding trivial voxels, and compute probability of
## infinite estimates
fits_mat %>% filter(statistic == "r" & voxel != 1) %>% group_by(parameter) %>%
    summarize(failed = 100 * sum((value == -Inf) * count) / sum(count),
              infinite = 100 * sum(infinite * count) / sum(count))
# # A tibble: 6 x 3
#   parameter failed infinite
#   <fct>      <dbl>    <dbl>
# 1 age         20.5     63.7
# 2 DD          18.1     63.7
# 3 EDSS        10.3     63.2
# 4 PASAT       16.8     63.6
# 5 sex         22.4     78.3
# 6 type2       19.2     75.5
## detections
fits_mat %>%
    group_by(parameter, statistic) %>%
    filter(statistic %in% c("z_br", "corz_br")) %>%
    summarize(detections = mean(value < -1 | value > 1) * 100)
# # A tibble: 12 x 3
# # Groups:   parameter [?]
#    parameter statistic detections
#    <fct>     <fct>          <dbl>
#  1 age       corz_br         39.2
#  2 age       z_br            33.0
#  3 DD        corz_br         24.8
#  4 DD        z_br            18.9
#  5 EDSS      corz_br         26.0
#  6 EDSS      z_br            19.8
#  7 PASAT     corz_br         37.1
#  8 PASAT     z_br            29.9
#  9 sex       corz_br         29.9
# 10 sex       z_br            22.7
# 11 type2     corz_br         22.1
# 12 type2     z_br            17.1
## Empirical lesion counts
lesion_counts <- colSums(lesions)
lesion_counts[lesion_counts == 0] <- NA
nifti_counts <- nifti(img = array(lesion_counts, dim(white_matter)))
lumin <- c(45, 100)
cols_counts <-heat_hcl(n = max(lesion_counts, na.rm = TRUE),
                       h = c(265, 200),
                       c = c(80, 50),
                       l = lumin,
                       power = c(0.7, 2))
overlay2.nifti(white_matter, y = nifti_counts, z = 32,
               plot.type = "single", plane = "sagittal",
               col.y = cols_counts, title = "lesion counts per voxel",
               col.main = "white")
```



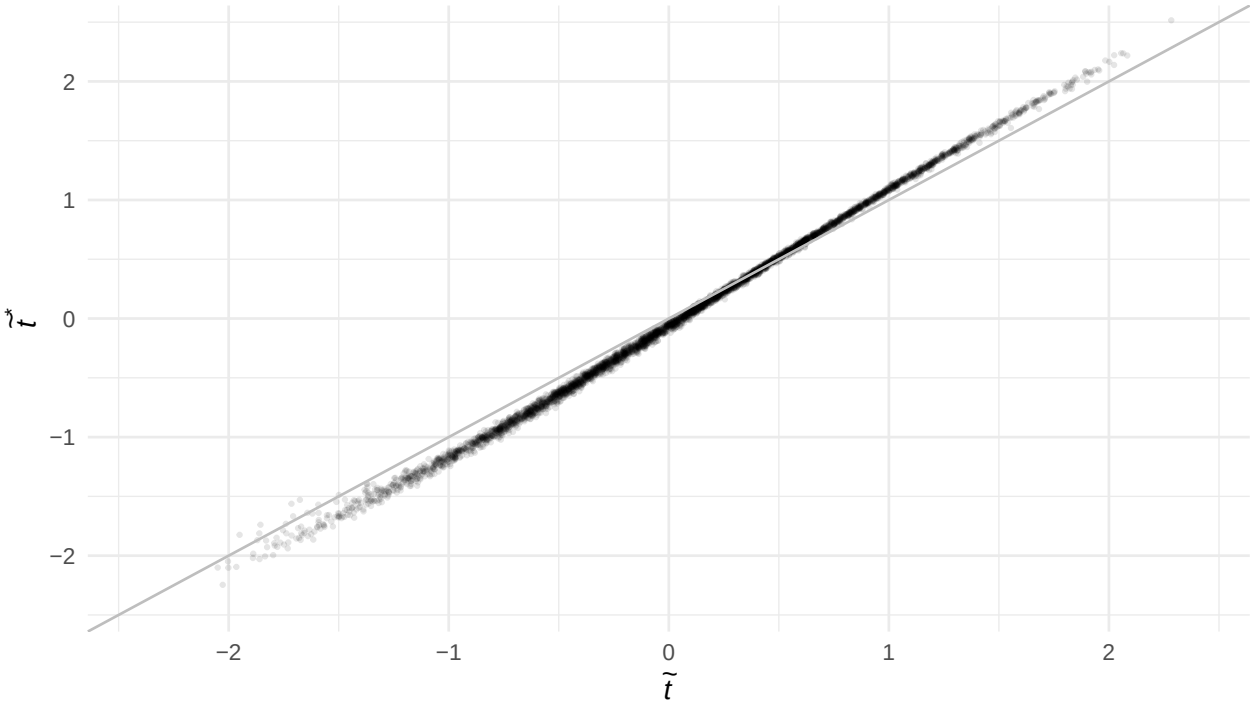

```
## Significance maps
param <- "DD"
low <- 1
upp <- 4
lumin <- c(25, 120)
cols <- c(heat_hcl(n = 32,
                   h = c(265, 200),
                   c = c(80, 50),
                   l = lumin,
                   power = c(0.7, 2)),
          rev(heat_hcl(n = 32,
                       h = c(10, 40),
                       c = c(80, 50),
                       l = lumin,
                       power = c(0.4, 1.3))))
for (stat in c("z_br", "corz_br")) {
    zz <- (fits_mat %>% filter(statistic == stat & parameter == param))
    zz <- zz$value[array_indices]
    ## Threshold as in Ge et al (2014, AOAS)
    low_ind <- abs(zz) < low
    low_ind[is.na(low_ind)] <- FALSE
    zz[low_ind] <- NA
    upp_ind <- abs(zz) >= upp
    upp_ind[is.na(upp_ind)] <- FALSE
    zz[upp_ind] <- sign(zz[upp_ind]) * upp
    nifti_z <- nifti(img = array(zz, dim(white_matter)))
    nifti_z[1,1,1] <- -upp
    nifti_z[1,1,2] <- upp
    main <- switch(stat,
                   z_br = expression(tilde(italic(t))),
                   corz_br = expression(tilde(italic(t))^'*'))
```

```
    overlay2.nifti(white_matter, y = nifti_z, z = 32, plot.type = "single",
                   plane = "sagittal", col.y = cols, title = main,
                   col.main = "white")
}
```




```
### Plot z_br vs corz_br per parameter
v1 <- fits_mat %>%
    filter(statistic == "z_br", parameter == param) %>%
```

```
    select(z_br_value = value, voxel, parameter)
v2 <- fits_mat %>%
    filter(statistic == "corz_br", parameter == param) %>%
    select(corz_br_value = value, voxel, parameter)
v <- join(v1, v2, by = c("voxel", "parameter"))
ggplot(v) +
    geom_point(aes(x = z_br_value, y = corz_br_value), alpha = 0.1, size = 0.5) +
    geom_abline(aes(intercept = 0, slope = 1), col = "grey") +
    coord_cartesian(xlim = c(-2.4, 2.4), ylim = c(-2.4, 2.4)) +
    theme_minimal() +
    labs(x = expression(tilde(italic(t))), y = expression(tilde(italic(t))^'*'))
```



References

Agresti, Alan, and Brian Caffo. 2000. "Simple and Effective Confidence Intervals for Proportions and Differences of Proportions Result from Adding Two Successes and Two Failures." *The American Statistician* 54 (4). Taylor & Francis: 280–88. https://doi.org/10.1080/00031305.2000.10474560.

Agresti, Alan, and Brent A. Coull. 1998. "Approximate Is Better Than Exact for Interval Estimation of Binomial Proportions." *The American Statistician* 52 (2). Taylor & Francis: 119–26. https://doi.org/10.1080/00031305.1998.10480550.

Di Caterina, Claudia, and Ioannis Kosmidis. 2017. "Location-Adjusted Wald Statistic for Scalar Parameters." *ArXiv E-Prints*. https://arxiv.org/abs/1710.11217.

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