Changes in extremes

Detection and consequences

Ilaria Prosdocimi

13 July 2020

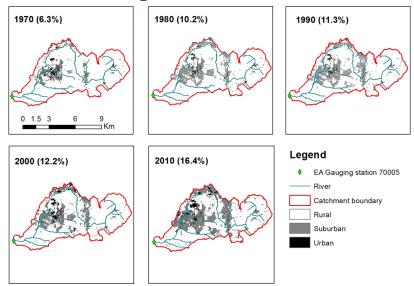
Change (?)

Increasing interest in assessing changes in extremes related to natural hazards.

Many studies investigate changes in extreme rainfall and extreme flows.

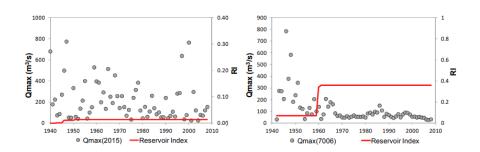
Changes in magnitude/frequencies: infrastructures are designed to with stand extreme events of some magnitude. Problematic if these become more (or less!) frequent.

What causes change



from Prosdocimi et al. (2015), WRR, doi:10.1002/2015WR017065

What causes change

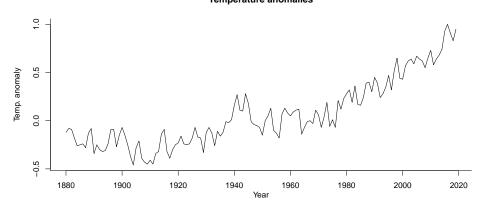


from Lopez Frances (2013), HESS, doi:10.5194/hess-17-3189-2013

What causes change

Implicit assumption:





NOAA National Centers for Environmental information, Climate at a Glance: Global Time Series, published June 2020, retrieved on July 5, 2020 from https://www.ncdc.noaa.gov/cag/

Why study change?

- Understand if process of interest (river flow, rainfall, etc) is evolving in time
- Understand how process of interest is affected by external drivers
- Assess risk connected to a certain hazard and its evolution
- If this is changing, how to account for this

Detection,

Why study change?

- Understand if process of interest (river flow, rainfall, etc) is evolving in time
- Understand how process of interest is affected by external drivers
- Assess risk connected to a certain hazard and its evolution
- If this is changing, how to account for this

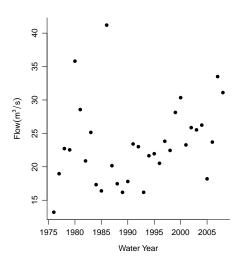
Detection, attribution

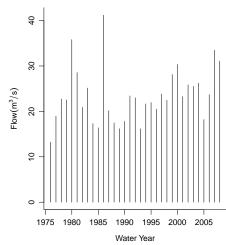
Why study change?

- Understand if process of interest (river flow, rainfall, etc) is evolving in time
- Understand how process of interest is affected by external drivers
- Assess risk connected to a certain hazard and its evolution
- If this is changing, how to account for this

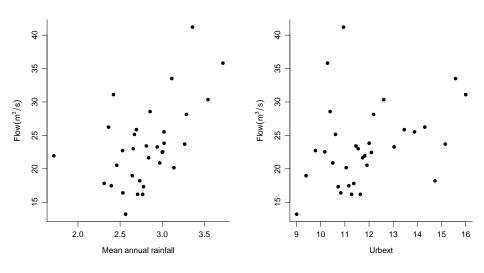
Detection, attribution and management.

The Lostock at Littlewood Bridge





The Lostock at Littlewood Bridge



Statistical tools

We assume that $\mathbf{y} = (y_1, \dots, y_n)$ is a random sample from a population.

We are interested in discovering some property of the population.

Statistical tools

We assume that $\mathbf{y} = (y_1, \dots, y_n)$ is a random sample from a population.

We are interested in discovering some property of the population.

Inference framework:

- Parametric: assume that y_i is a realisation of some distribution described by parameters θ ($f(y_i; \theta)$)
- Non-parametric: no assumption on the distribution of f(y) is made (well, less assumptions. . .)

Parameteric framework

Advantage of parametric framework:

- Describe the whole distribution (including, for example, quantiles)
- A very general framework
- Easy to extend to very complex models (but estimation can be complicated)

Parameteric framework

Advantage of parametric framework:

- Describe the whole distribution (including, for example, quantiles)
- A very general framework
- Easy to extend to very complex models (but estimation can be complicated)

The parametric framework:

- ullet Assume that each member of the sample y_i comes from some distribution Y_i
- ullet Often assumed: (Y_1,\ldots,Y_n) are independent and identically distributed (iid)
- ullet Assume that Y_i follows a known distribution parametrised by $oldsymbol{ heta}$
- (for example $Y_i \sim N(\mu, \sigma)$, with $\theta = (\mu, \sigma)$)
- ullet Find estimates $\hat{oldsymbol{ heta}}$ based on the sample

Estimation methods

- Method of moments
- Maximum likelihood
- Bayesian approaches

Estimation methods

- Method of moments
- Maximum likelihood
- Bayesian approaches

Choice of framework and estimation method should depend on:

- Actual data properties
- Main inferential question (and importance of uncertainty assessment)
- Computational hurdle
- Model complexity
- Presence of prior information (which can be formalised)

Maximum likelhood estimation

The likelihood function is defined as

$$L(\boldsymbol{\theta}; \boldsymbol{y}) = \prod_{i=1}^{n} f(y_i, \boldsymbol{\theta}),$$

but calculations typically employ the log-likelihood

$$I(\boldsymbol{\theta}; \mathbf{y}) = \sum_{i=1}^{n} \log f(y_i, \boldsymbol{\theta}).$$

 $\hat{\boldsymbol{\theta}}_{\mathit{ML}}$ is the value that maximises $I(\boldsymbol{\theta}; \boldsymbol{y}).$

Maximum likelhood estimation

The likelihood function is defined as

$$L(\boldsymbol{\theta}; \boldsymbol{y}) = \prod_{i=1}^{n} f(y_i, \boldsymbol{\theta}),$$

but calculations typically employ the log-likelihood

$$l(\boldsymbol{\theta}; \mathbf{y}) = \sum_{i=1}^{n} \log f(y_i, \boldsymbol{\theta}).$$

 $\hat{\boldsymbol{\theta}}_{ML}$ is the value that maximises $I(\boldsymbol{\theta}; \boldsymbol{y})$.

Asymptotically $(n \to \infty)$ we have that $\hat{\theta}_{ML} \sim N(\theta, I_E(\theta)^{-1})$ where $I_E(\theta)$ is the expected information matrix, with elements

$$e_{i,j}(\theta) = E\left[-\frac{d^2I(\theta)}{d\theta_id\theta_j}\right]$$

Typically $I_E(\theta)$ is unknown: use the observed information matrix evaluated at $\hat{\theta}$.

Parametric models for change

- Assume Y_i comes from a distribution $f(\theta_i, y_i)$
- Assume $\theta_i = g(\mathbf{x}_i)$
- So $Y_i = (Y|X = x_i)$ with $f(g(\mathbf{x}_i), y_i)$

Example. Linear regression (with two explanatory variables):

- $Y_i \sim N(\mu_i, \sigma)$; $\theta_i = (\mu_i, \sigma)$
- $\mu_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i}$ linear relationship
- \bullet σ is constant
- As a consequence: $E[Y_i] = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i}$, $V[Y_i] = \sigma^2$

Parametric models for change

Linear regression likelihood:

$$I(\boldsymbol{\theta}; \boldsymbol{y}) = \sum_{i=1}^{n} \log f(y_i, \boldsymbol{\theta}) \propto -n \log(\sigma) - \frac{(y - \beta_0 - \beta_1 x_{1i} - \beta_2 x_{2i})^2}{2\sigma^2}$$

ML estimates can be derived analytically: $(\hat{\beta}_0, \hat{\beta}_1, \hat{\beta}_2, \hat{\sigma})$.

And we have, for example, $\hat{\beta}_i \sim N(\beta_i, \hat{\sigma}_{\beta_i})$.

From this once can construct confidence intervals for β_i or perform a test such as:

$$H_0: \beta_0 \geq \tilde{\beta}$$
 VS $H_0: \beta_0 < \tilde{\beta}$

By default $\tilde{\beta}=0$, but one can test for any value $\tilde{\beta}$ and any type of statistical test (equality, larger than or equal, smaller than or equal).

Notice that if x_j is a factor one can account for step changes (change points).

Parametric models of change in extremes

Describing extremes is a different task than describing the typical behaviour.

 (y_1,\ldots,y_n) is a sample of extremes: what is a reasonable assumption for Y?

Extreme Value Theory gives theoretical derivation, but practice is often different.

Regardless of the choice of $f(y, \theta)$ - parametric models of change for extremes can be easily constructed assuming $Y_i = (Y|X = x_i)$ and $\theta_i = g(\mathbf{x}_i)$.

Parametric models of change in extremes

Describing extremes is a different task than describing the typical behaviour.

 (y_1, \ldots, y_n) is a sample of extremes: what is a reasonable assumption for Y?

Extreme Value Theory gives theoretical derivation, but practice is often different.

Regardless of the choice of $f(y, \theta)$ - parametric models of change for extremes can be easily constructed assuming $Y_i = (Y|X = x_i)$ and $\theta_i = g(\mathbf{x}_i)$.

What is an extreme?

- Largest event over a certain amount of time (eg water year, season)
- Events larger than a certain high threshold (independent events?)

Traditional (asymptotic) results based on extremes of stationary series:

- Block maxima: $Y \sim GEV(\mu, \sigma, \xi)$
- ullet Threshold exceedance magnitude: $Y \sim \textit{GP}(\sigma, \xi)$
- Threshold exceedance frequency: $N \sim Pois(\lambda)$

¹Smith, Statist. Sci., doi:10.1214/ss/1177012400

Traditional (asymptotic) results based on extremes of stationary series:

- Block maxima: $Y \sim GEV(\mu, \sigma, \xi)$
- Threshold exceedance magnitude: $Y \sim GP(\sigma, \xi)$
- Threshold exceedance frequency: $N \sim Pois(\lambda)$

Exceedances frequency and magnitude traditionally modelled as separate processes.

Frequency and magnitude of threshold exceedances can be modelled in a unique framework using a Point Process representation of extremes ¹.

¹Smith, Statist. Sci., doi:10.1214/ss/1177012400

Traditional (asymptotic) results based on extremes of stationary series:

- Block maxima: $Y \sim GEV(\mu, \sigma, \xi)$
- Threshold exceedance magnitude: $Y \sim GP(\sigma, \xi)$
- Threshold exceedance frequency: $N \sim Pois(\lambda)$

Exceedances frequency and magnitude traditionally modelled as separate processes.

Frequency and magnitude of threshold exceedances can be modelled in a unique framework using a Point Process representation of extremes ¹. This representation is under-utilised in hydrology.

¹Smith, Statist. Sci., doi:10.1214/ss/1177012400

Traditional (asymptotic) results based on extremes of stationary series:

- Block maxima: $Y \sim GEV(\mu, \sigma, \xi)$
- Threshold exceedance magnitude: $Y \sim \textit{GP}(\sigma, \xi)$
- Threshold exceedance frequency: $N \sim Pois(\lambda)$

Exceedances frequency and magnitude traditionally modelled as separate processes.

Frequency and magnitude of threshold exceedances can be modelled in a unique framework using a Point Process representation of extremes ¹. This representation is under-utilised in hydrology.

Using exceedances typically results in larger samples (so less variability in estimates).

¹Smith, Statist. Sci., doi:10.1214/ss/1177012400

Traditional (asymptotic) results based on extremes of stationary series:

- Block maxima: $Y \sim GEV(\mu, \sigma, \xi)$
- Threshold exceedance magnitude: $Y \sim GP(\sigma, \xi)$
- Threshold exceedance frequency: $N \sim Pois(\lambda)$

Exceedances frequency and magnitude traditionally modelled as separate processes.

Frequency and magnitude of threshold exceedances can be modelled in a unique framework using a Point Process representation of extremes ¹. This representation is under-utilised in hydrology.

Using exceedances typically results in larger samples (so less variability in estimates).

In practice other distributions are often assumed for Flow maxima

¹Smith, Statist. Sci., doi:10.1214/ss/1177012400

Generalised Extreme Value distribution

The GEV CDF:
$$F(y, \theta) = \exp \left\{ -\left(1 + \xi \frac{y - \mu}{\sigma}\right)^{-1/\xi} \right\}$$

$$\boldsymbol{\theta} = (\mu, \sigma, \xi)$$
:

- $\mu \in \mathbb{R}$: location parameter
- $\sigma > 0$; scale parameter
- $\xi \in \mathbb{R}$: shape parameter.

 $Y \sim GEV(\mu, \sigma, \xi)$ is defined on $y: 1 + \xi(y - \mu)/\sigma > 0$, this means:

- $y \in [\mu \sigma/\xi, \infty)$, if $\xi > 0$ (Frechet)
- $y \in (-\infty, \mu \sigma/\xi]$, if $\xi < 0$ (Weibull)
- $y \in (-\infty, \infty)$, if $\xi = 0$ (Gumbel)

BUT! In engineering/hydrology $Y \sim GEV(\xi, \alpha, \kappa)$ and $\kappa = -\xi$. Software can use different parametrisation.

Generalised Extreme Value distribution

The GEV CDF:
$$F(y, \theta) = \exp \left\{ -\left(1 + \xi \frac{y - \mu}{\sigma}\right)^{-1/\xi} \right\}$$

$$\boldsymbol{\theta} = (\mu, \sigma, \xi)$$
:

- $\mu \in \mathbb{R}$: location parameter
- $\sigma > 0$; scale parameter
- $\xi \in \mathbb{R}$: shape parameter.

 $Y \sim GEV(\mu, \sigma, \xi)$ is defined on $y: 1 + \xi(y - \mu)/\sigma > 0$, this means:

- $y \in [\mu \sigma/\xi, \infty)$, if $\xi > 0$ (Frechet)
- $y \in (-\infty, \mu \sigma/\xi]$, if $\xi < 0$ (Weibull)
- $y \in (-\infty, \infty)$, if $\xi = 0$ (Gumbel)

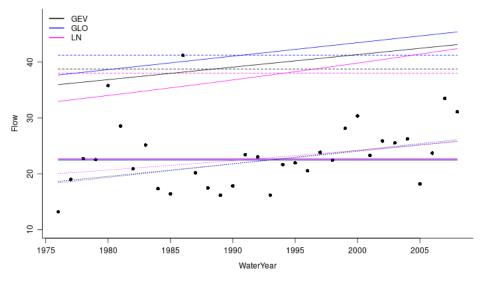
BUT! In engineering/hydrology $Y \sim GEV(\xi, \alpha, \kappa)$ and $\kappa = -\xi$. Software can use different parametrisation.

Quantile function (for $\xi \neq 0$):

$$q(y,\theta) = \mu + \frac{\sigma}{\xi} \left[(-\log(1-p))^{-\xi} - 1 \right]$$

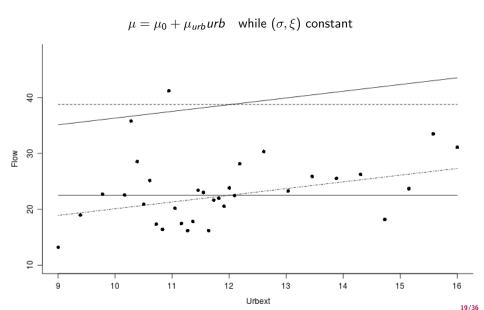
Changes in annual maxima - choice of distribution

The Lostock at Littlewood Bridge: median and effective 50-yrs event.



Changes in annual maxima

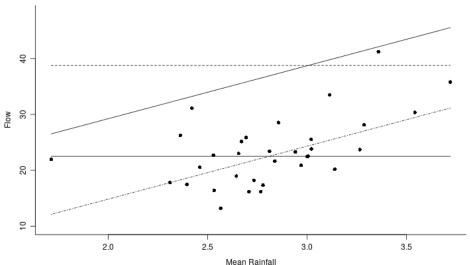
Time is not a cause for change, but land cover changes impact peak flow.



Changes in annual maxima

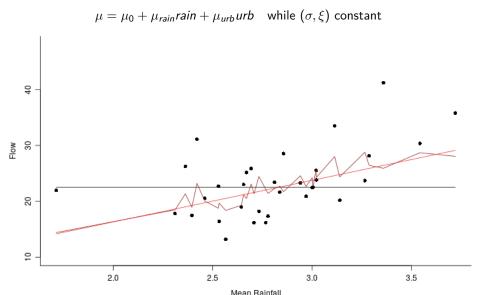
Time is not a cause for change, but soil wetness impact peak flow.

$$\mu = \mu_0 + \mu_{\it rain} rain$$
 while (σ, ξ) constant



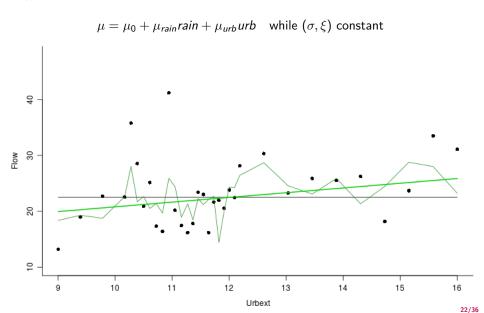
Changes in amax - effect of rain given Urbext

Separate effect of rain and urbanisation:



Changes in amax - effect of Urbext given rain

Separate effect of rain and urbanisation:



Changes in annual maxima - estimated parameters

Rain as covariate (log-lik: -98.39)

	μ_0	$\mu_{\it urb}$	$\mu_{\it rain}$	σ	ξ
MLE	-5.604	-	9.479	4.042	0.003
se	8.158	-	2.830	0.622	0.168

Urbext as covariate (log-lik: -100.0004)

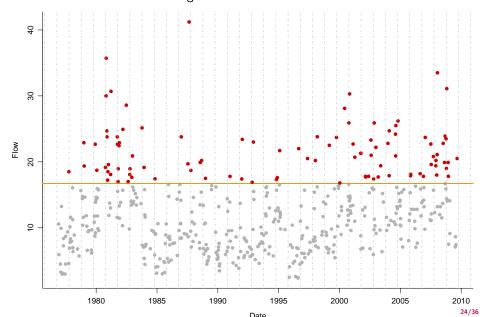
	μ_0	$\mu_{\it urb}$	$\mu_{\it rain}$	σ	ξ
MLE	6.53	1.20	-	4.17	0.04
se	4.79	0.40	-	0.60	0.13

Rain and urbext as covariate (log-lik: -96.47)

	μ_{0}	$\mu_{\it urb}$	$\mu_{\it rain}$	σ	ξ
MLE	-9.767	0.845	7.449	3.862	-0.016
se	7.344	0.422	2.668	0.580	0.153

Changes is peaks over threshold

Extract observations above a high threshold



Generalised Pareto Distribution

Y is taken to be the observations above a high threshold u (Y = (X|X > u)).

GP is the limiting distribution for the magnitude of exceedances.

$$F(y, u, \boldsymbol{\theta}) = 1 - \left(1 + \xi \frac{y - u}{\tilde{\sigma}}\right)^{-1/\xi}$$

u is a constant, $\theta = (\sigma, \xi)$:

- $\sigma > 0$; scale parameter
- $\xi \in \mathbb{R}$: shape parameter.

The domain changes depending on the sign of ξ :

- $y \in [u, \infty)$, if $\xi \ge 0$
- $y \in (-\infty, u \sigma/\xi]$, if $\xi < 0$

Quantile function:
$$q(p,u, heta)=u+rac{\sigma}{\xi}(p^{-\xi}-1)$$

Generalised Pareto Distribution

Y is taken to be the observations above a high threshold u (Y = (X|X > u)).

GP is the limiting distribution for the magnitude of exceedances.

$$F(y, u, \boldsymbol{\theta}) = 1 - \left(1 + \xi \frac{y - u}{\tilde{\sigma}}\right)^{-1/\xi}$$

u is a constant, $\theta = (\sigma, \xi)$:

- $\sigma > 0$; scale parameter
- $\xi \in \mathbb{R}$: shape parameter.

The domain changes depending on the sign of ξ :

- $y \in [u, \infty)$, if $\xi \ge 0$
- $y \in (-\infty, u \sigma/\xi]$, if $\xi < 0$

Quantile function:
$$q(p,u,oldsymbol{ heta})=u+rac{\sigma}{\xi}(p^{-\xi}-1)$$

Modelling change: $\sigma_0 + \sigma_1 x$

Generalised Pareto Distribution

Y is taken to be the observations above a high threshold u (Y = (X|X > u)).

GP is the limiting distribution for the magnitude of exceedances.

$$F(y, u, \boldsymbol{\theta}) = 1 - \left(1 + \xi \frac{y - u}{\tilde{\sigma}}\right)^{-1/\xi}$$

u is a constant, $\theta = (\sigma, \xi)$:

- $\sigma > 0$; scale parameter
- $\xi \in \mathbb{R}$: shape parameter.

The domain changes depending on the sign of ξ :

- $y \in [u, \infty)$, if $\xi \ge 0$
- $y \in (-\infty, u \sigma/\xi]$, if $\xi < 0$

Quantile function:
$$q(p,u,oldsymbol{ heta})=u+rac{\sigma}{arepsilon}(p^{-\xi}-1)$$

Modelling change: $\sigma_0 + \sigma_1 x$

Frequency of exceedances: typically modelled as a Poisson (indep. of magnitude).

Point Process representation of extremes

 $N = \{\text{no. Exceedance in a Year}\}. N \sim Pois(\lambda)$

P(no. Exceedance in a Year) is linked to magnitudes.

Express this using GEV-parameters:

$$\log \lambda = -\frac{1}{\xi} \log \left[1 + \xi \frac{u - \mu}{\sigma} \right]$$

Point Process representation of extremes

 $N = \{\text{no. Exceedance in a Year}\}. N \sim Pois(\lambda)$

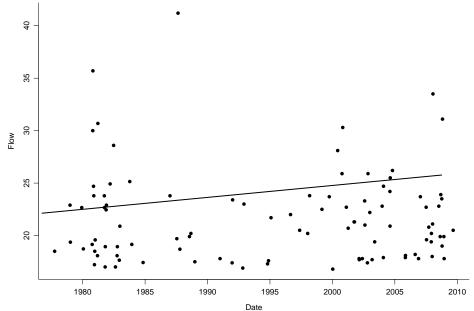
P(no. Exceedance in a Year) is linked to magnitudes.

Express this using GEV-parameters:

$$\log \lambda = -\frac{1}{\xi} \log \left[1 + \xi \frac{u - \mu}{\sigma} \right]$$

Express changes in magnitude and frequency in the same model Same meaning as GEV models of change

Changes in Peaks - Point Process



Changes in extremes - comparing the models

Rain and urbext as covariate - GEV:

	μ_0	$\mu_{\it urb}$	$\mu_{\it rain}$	σ	ξ
MLE	-9.767	0.845	7.449	3.862	-0.016
se	7.344	0.422	2.668	0.580	0.153

Rain and urbext as covariate - PP:

	μ_{0}	$\mu_{\it urb}$	$\mu_{\it rain}$	σ	ξ
MLE	-12.139	0.930	8.007	4.622	-0.184
se	6.757	0.320	1.723	0.368	0.064

Changes in extremes - comparing the models

Rain and urbext as covariate - GEV:

	μ_0	$\mu_{\it urb}$	$\mu_{\it rain}$	σ	ξ
MLE	-9.767	0.845	7.449	3.862	-0.016
se	7.344	0.422	2.668	0.580	0.153

Rain and urbext as covariate - PP:

	μ_0	$\mu_{\it urb}$	$\mu_{\it rain}$	σ	ξ
MLE	-12.139	0.930	8.007	4.622	-0.184
se	6.757	0.320	1.723	0.368	0.064

Larger sample size leads to more precise estimation (statistically)

Tail estimate is quite different

Changes in extremes

Parametric approaches: easy to include predictors and test for significance

Changes in extremes

Parametric approaches: easy to include predictors and test for significance This might be a bug and not a feature

Changes in extremes

Parametric approaches: easy to include predictors and test for significance

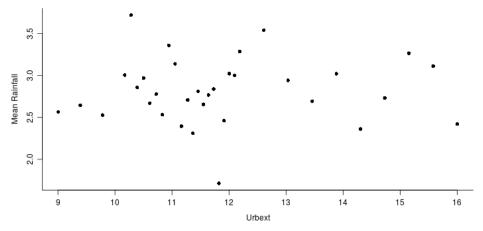
This might be a bug and not a feature

The assumption is that $Y_i = (Y|X = x_i)$ follows $f(y; \theta)$ - goodness of fit should be carried out on **residuals**

Statistical EVT and practice are not aligned

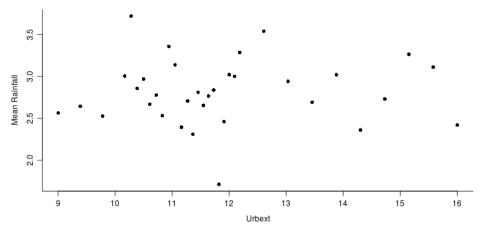
Changes in extremes - attribution

Kendall's $\hat{\tau}(\text{Urbext, Rain}) = 0.068$.



Changes in extremes - attribution

Kendall's $\hat{\tau}(Urbext, Rain) = 0.068$.



Reality is complex: linear models are a (over-simplified!) representation.

Detection

Methods sometimes chosen because of data availability

Statistical models rely on assumption of iid random observations

Short records: hard to identify complex evolutions

Short records: hard to observe a good range of the explantory variable

When detecting "change": what are we detecting?²

²Merz et al, HESS, doi:10.5194/hess-16-1379-2012

Attribution

Golden standard of causality is randomised trials: what about observational studies?

Climate sciences reproduce the treatment/placebo framework with numerical experiments (how good for extremes?).

Some numerical experiments done in hydrology - but systems are complex.

Causality: a cascade of impacts (with feedback³)

 $^{^3}$ Zhang et al, Nature, doi:10.1038/s41586-018-0676-z

Changes in annual maxima - uncertainity

Structures are designed for the "T-Year" event: estimated as the 1-1/T quantile. If the distribution is changing so is the quantile.

Changes in annual maxima - uncertainity

Structures are designed for the "T-Year" event: estimated as the 1-1/T quantile. If the distribution is changing so is the quantile.

	Q100	95% lb	95% ub	width
no-change	30.514	41.837	53.159	11.322
Rain = max(Rain)	33.676	48.403	63.130	14.727

Adding parameters adds variation to the estimates - is it worth it?

Changes in annual maxima - uncertainity

Structures are designed for the "T-Year" event: estimated as the 1-1/T quantile. If the distribution is changing so is the quantile.

	Q100	95% lb	95% ub	width
no-change	30.514	41.837	53.159	11.322
Rain = max(Rain)	33.676	48.403	63.130	14.727

Adding parameters adds variation to the estimates - is it worth it? Bias-variance trade-off and parsimonious models.

Changes in extremes - consequences

How to quantify risk under change?⁴

Choice of distribution has an impact on estimates of rare events

Today I used "effective design events": $q(p; \hat{\theta})$. So at $X = x^*$: $q(p; \hat{\theta}(x^*))$.

Choice of distribution/model has an impact on estimates of rare events.

Choice of model has an impact of description of change. Compare effective return levels for x^* and x_0 :

$$q(p; \hat{\theta}(x^*)) - q(p; \hat{\theta}(x_0)) = \mu_1(x^* - x_0)$$

⁴Volpi, Wires Water, doi:10.1002/wat2.1340

(Statistical) reccomended reading

Coles, S (2001), An introduction to statistical modeling of extreme values, Springer

Katz, R.W., Parlange, M.B. and Naveau, P., 2002. Statistics of extremes in hydrology. Advances in water resources, 25(8-12), pp.1287-1304.

Katz, Richard (2013) Statistical Methods for Nonstationary Extremes, Chapter 2 in A. AghaKouchak et al. (eds.), Extremes in a Changing Climate, Water Science and Technology Library 65, DOI 10.1007/978-94-007-4479-0 2,

Doing science the right way

Reproducibility crisis in several fields - open science movement as a results.

Replicability (i.e. being able to re-run the analysis) should be a given.

Start any project in a replicable way: literate programming and programmatic interfaces In R (and Python) this is increasingly feasible

Slides code at github.com/ilapros