# Section 8 Convex Optimisation 2

# Lagrangian

#### Consider a general optimization problem

$$\min_{\boldsymbol{x}} f(\boldsymbol{x})$$
subject to  $h_i(\boldsymbol{x}) \leq 0, \ i = 1, \dots, m,$ 

$$\ell_j(\boldsymbol{x}) = 0, \ j = 1, \dots, r.$$

The objective function f needs not to be convex. Of course we pay special attention to the convex case.

# Definition 8.1 (Lagrangian)

$$L\left(\boldsymbol{x},\boldsymbol{u},\boldsymbol{v}\right) = f\left(\boldsymbol{x}\right) + \sum_{i=1}^{m} u_{i}h_{i}\left(\boldsymbol{x}\right) + \sum_{j=1}^{r} v_{j}\ell_{j}\left(\boldsymbol{x}\right).$$
 Here  $\boldsymbol{u} \in \mathbb{R}^{m} \ \boldsymbol{v} \in \mathbb{R}^{r}$  and  $\boldsymbol{u} > \boldsymbol{0}$ 

Here  $\boldsymbol{u} \in \mathbb{R}^m$ ,  $\boldsymbol{v} \in \mathbb{R}^r$ , and  $\boldsymbol{u} > \boldsymbol{0}$ .

# Lagrange Dual Function

Uhanstrained opt. Problem
Definition 8.2 (Lagrange Dual Function) respect to  $\chi$ .

$$g\left(\boldsymbol{u},\boldsymbol{v}\right):=\min_{\boldsymbol{x}\in\mathbb{R}^{n}}L\left(\boldsymbol{x},\boldsymbol{u},\boldsymbol{v}\right)=\min_{\boldsymbol{x}\in\mathbb{R}^{n}}\left[f\left(\boldsymbol{x}\right)+\sum_{i=1}^{m}u_{i}h_{i}\left(\boldsymbol{x}\right)+\sum_{j=1}^{r}v_{j}\ell_{j}\left(\boldsymbol{x}\right)\right].$$

▶ For every feasible x ( $x \in \mathcal{X}$ ),  $L(x, u, v) \leq f(x)$ 

$$L\left(\boldsymbol{x},\boldsymbol{u},\boldsymbol{v}\right) = f\left(\boldsymbol{x}\right) + \underbrace{\sum_{i=1}^{m} u_{i} h_{i}\left(\boldsymbol{x}\right)}_{\leq 0} + \underbrace{\sum_{j=1}^{r} v_{j} \ell_{j}\left(\boldsymbol{x}\right)}_{=0}.$$

Let  $\mathcal{X}$  denote the primal feasible set.

$$g\left(\boldsymbol{u},\boldsymbol{v}\right) = \min_{\boldsymbol{x} \in \mathbb{R}^n} L\left(\boldsymbol{x},\boldsymbol{u},\boldsymbol{v}\right) \leq \min_{\boldsymbol{x} \in \mathcal{X}} L\left(\boldsymbol{x},\boldsymbol{u},\boldsymbol{v}\right) \leq f\left(\boldsymbol{x}\right).$$
 (13)

# Concavity of Lagrange Dual Function

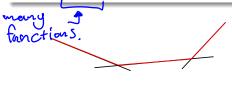
#### Lemma 8.3

The Lagrange dual function

$$g\left(\boldsymbol{u},\boldsymbol{v}\right) = \min_{\boldsymbol{x} \in \mathbb{R}^n} L\left(\boldsymbol{x},\boldsymbol{u},\boldsymbol{v}\right) = \min_{\boldsymbol{x} \in \mathbb{R}^n} f\left(\boldsymbol{x}\right) + \sum_{i=1}^m u_i h_i\left(\boldsymbol{x}\right) + \sum_{j=1}^r v_j \ell_j\left(\boldsymbol{x}\right)$$
 is concave in  $(\boldsymbol{u},\boldsymbol{v})$ .

#### Lemma 8.4

- ▶ Let  $f_{\alpha}(x)$  be concave functions. Then  $g(x) = \inf_{\alpha} f_{\alpha}(x)$  is concave.
- ▶ Let  $f_{\alpha}\left(x\right)$  be convex functions. Then  $g\left(x\right)=\sup_{\alpha}\ f_{\alpha}\left(x\right)$  is convex.





#### **Proofs**

Proof of Lemma 8.4: For any 
$$\lambda \in [0,1]$$
, 
$$g\left(\lambda x + (1-\lambda)\,y\right) = \inf_{\alpha} \, f_{\alpha}\left(\lambda x + (1-\lambda)\,y\right) \quad \text{for is} \\ \geq \inf_{\alpha} \, \lambda f_{\alpha}\left(x\right) + (1-\lambda)\,f_{\alpha}\left(y\right) \\ \geq \lambda \inf_{\alpha} \, f_{\alpha}\left(x\right) + (1-\lambda)\inf_{\alpha} \, f_{\alpha}\left(y\right).$$

Proof of Lemma 8.3: For any given x, L(x, u, v) is linear in (u, v), and hence concave in (u, v). The minimum of concave functions is concave based on Lemma 8.4.

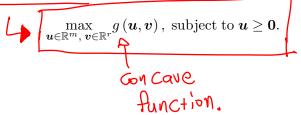
# Lagrange Dual Problem

#### Given the primal problem

$$\min_{\boldsymbol{x}} f(\boldsymbol{x})$$
subject to  $h_i(\boldsymbol{x}) \leq 0, \ i = 1, \dots, m,$ 

$$\ell_j(\boldsymbol{x}) = 0, \ j = 1, \dots, r.$$

#### Its Lagrange dual problem is



# Weak and Strong Duality

Weak duality: the dual optimal value  $q^*$  satisfies

$$f^{\star} \geq g^{\star}$$
.

This is a direct consequence of (13).

Strong duality is referred to as the case that  $f^* = g^*$ . If the potential way  $f^* = g^*$ . We equal

Slater's condition: if the primal is a convex problem (i.e., f and  $g_i$ 's are convex and  $\ell_i$ 's are affine), and there exists at least one strictly feasible  $oldsymbol{x} \in \mathbb{R}^n$  satisfying

$$h_i(\boldsymbol{x}) < 0, \ \forall i \in [m], \ \text{and} \ \ell_j(\boldsymbol{x}) = 0, \ \forall j \in [r],$$

then strong duality holds. (Proof is omitted.)

#### Karush-Kuhn-Tucker conditions

Given the optimization problem

Fucker conditions ation problem 
$$cowex$$
 functions  $f(x)$  minimize  $f(x)$   $i=1,\dots,m$ 

subject to 
$$h_i(\mathbf{x}) \leq 0, i = 1, \dots, m,$$

The Karush-Kuhn-Tucker (KKT) conditions are:

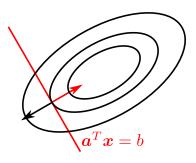
▶ 
$$\mathbf{0} \in \partial f(\mathbf{x}) + \sum_{i=1}^{m} u_i \partial h_i(\mathbf{x}) + \sum_{j=1}^{r} v_j \partial \ell_j(\mathbf{x}).$$
 (stationarity)

$$ullet u_i h_i\left(oldsymbol{x}
ight) = 0, \ orall i.$$
 Length useful, (complementary slackness)

$$lacksymbol{h}_{i}\left(oldsymbol{x}
ight)\leq0,\;\ell_{j}\left(oldsymbol{x}
ight)=0,\;orall i,\;orall j.\;$$
  $lacksymbol{\psi}_{i}$  (primal feasibility)

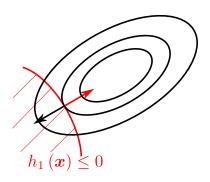
Necessary under strong duality.

## Geometric Intuition: Equality Constraints



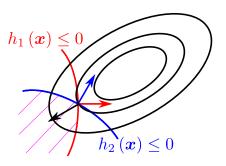
 $\partial f(x)$  is a linear combination of  $\partial \ell_i(x)$ 's.

# Geometric Intuition: One Inequality Constraint



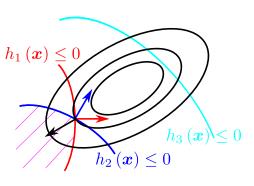
$$\partial f(\mathbf{x}) + u_1 \partial h_1(\mathbf{x}) = \mathbf{0}.$$
  
 $h_1(\mathbf{x}) = 0.$ 

# Geometric Intuition: Inequality Constraints



$$\partial f(\mathbf{x}) + \sum_{i=1}^{2} u_i \partial h_i(\mathbf{x}) = \mathbf{0}.$$
  
$$h_1(\mathbf{x}) = 0, h_2(\mathbf{x}) = 0.$$

# Geometric Intuition: Inequality Constraints



$$\begin{aligned} \partial f\left(\boldsymbol{x}\right) + \sum_{i=1}^{3} u_{i} \partial h_{i}\left(\boldsymbol{x}\right) &= \boldsymbol{0}. \\ h_{1}\left(\boldsymbol{x}\right) &= 0, \ h_{2}\left(\boldsymbol{x}\right) &= 0, \\ h_{3}\left(\boldsymbol{x}\right) &< 0 \text{ but } u_{3} &= 0 \text{ so that } u_{3}h_{3}\left(\boldsymbol{x}\right) &= 0 \end{aligned}$$

# Sufficiency

If  $x^{\star}, u^{\star}, v^{\star}$  satisfy the KKT conditions, then  $x^{\star}$  and  $u^{\star}, v^{\star}$  are primal and dual solutions.

If  $x^\star, u^\star, v^\star$  satisfy the KKT conditions, then

$$g(\boldsymbol{u}^{\star}, \boldsymbol{v}^{\star}) = f(\boldsymbol{x}^{\star}) + \sum_{i=1}^{m} u_{i}^{\star} h_{i}(\boldsymbol{x}^{\star}) + \sum_{j=1}^{r} v_{i}^{\star} \ell_{i}(\boldsymbol{x}^{\star})$$
$$= f(\boldsymbol{x}^{\star}),$$

where the first equality follows from stationarity, and the second follows from complementary slackness. This equality suggests the duality gap is zero. Hence,  $\boldsymbol{x}^{\star}$ ,  $\boldsymbol{u}^{\star}$  and  $\boldsymbol{v}^{\star}$  are primal and dual optimal.

## Necessity

Suppose that the strong duality holds and that  $x^*$  and  $u^*$ ,  $v^*$  are primal and dual solutions. Then  $x^*$ ,  $u^*$ ,  $v^*$  satisfy the KKT conditions.

Due to the strong duality, one has

$$f(\boldsymbol{x}^{\star}) = g(\boldsymbol{u}^{\star}, \boldsymbol{v}^{\star})$$

$$= \min_{\boldsymbol{x} \in \mathbb{R}^{n}} f(\boldsymbol{x}) + \sum_{i=1}^{m} u_{i}^{\star} h_{i}(\boldsymbol{x}) + \sum_{j=1}^{r} v_{j}^{\star} \ell_{j}(\boldsymbol{x})$$

$$\leq f(\boldsymbol{x}^{\star}) + \sum_{i=1}^{m} u_{i}^{\star} h_{i}(\boldsymbol{x}^{\star}) + \sum_{j=1}^{r} v_{j}^{\star} \ell_{j}(\boldsymbol{x}^{\star})$$

$$\leq f(\boldsymbol{x}^{\star}).$$

In other words, all the inequalities are actually equalities.

# Quadratic Programming with Equality Constraints

Let  $\mathbf{Q} \succeq 0$ .

$$\min_{\boldsymbol{x}} \frac{1}{2} \boldsymbol{x}^T \boldsymbol{Q} \boldsymbol{x} + \boldsymbol{c}^T \boldsymbol{x} \text{ subject to } \boldsymbol{A} \boldsymbol{x} = \boldsymbol{0}.$$

By KKT conditions, x is the minimizer if and only if

$$\left[egin{array}{cc} m{Q} & m{A}^T \ m{A} & m{0} \end{array}
ight] \left[egin{array}{cc} m{x} \ m{u} \end{array}
ight] = \left[egin{array}{cc} -m{c} \ m{0} \end{array}
ight],$$

where the first set of linear equations come from the stationarity and the second set follows from the primal feasibility.

The optimal  $x^*$  can be obtained by solving the linear inverse problem.

Water Filling

r Filling channel capacity of the 
$$\min_{\boldsymbol{x}} - \sum_{i=1}^{n} \log (\alpha_i + x_i) \text{ subject to } \boldsymbol{x} \geq \boldsymbol{0}, \ \boldsymbol{1}^T \boldsymbol{x} = 1.$$

$$L = -\sum_{i=1}^{n} \log (\alpha_i + x_i) + \sum_{i=1}^{n} (\alpha_i +$$

By KKT conditions

$$-1/(\alpha_i + x_i) - u_i + v = 0, \forall i$$

$$u_i x_i = 0, \forall i$$

$$x > 0$$
,  $\mathbf{1}^T x = 1$ ,  $u > 0$ .

Eliminate u. The first two conditions become

$$1/(\alpha_i + x_i) \le v$$
, and  $x_i(v - 1/(\alpha_i + x_i)) = 0$ ,  $\forall i$ .

Therefore, the solution:

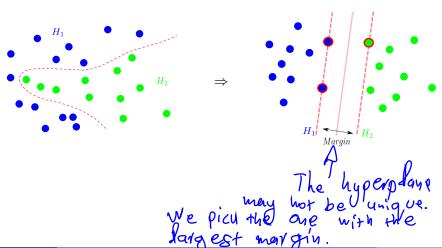
$$x_i = \max(0, 1/v - \alpha_i)$$

where v is chosen such that

$$\sum_{i=1}^{n} \max(0, 1/v - \alpha_i) = 1.$$

# Section 9 Support Vector Machine

#### Idea of SVM

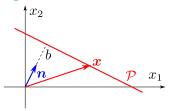


## A Hyperplane

A hyperplane in  $\mathbb{R}^n$  can be defined using its normal vector  $n \in \mathbb{R}^n$ :

$$\mathcal{P} = \left\{ oldsymbol{x} : \quad oldsymbol{n}^T oldsymbol{x} = b 
ight\}.$$

▶ Usually we assume  $\|n\|_2 = 1$ .



The projection  $\|\operatorname{Proj}(\boldsymbol{x}, \operatorname{span}(\boldsymbol{n}))\|_2 = b$ .

▶ If  $||n||_2 \neq 1$ , then

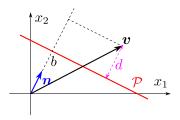
$$\mathcal{P} = \left\{ oldsymbol{x} : \quad oldsymbol{n}^T oldsymbol{x} = b 
ight\} = \left\{ oldsymbol{x} : \quad oldsymbol{n}^T oldsymbol{x} / \left\| oldsymbol{n} 
ight\|_2 = b / \left\| oldsymbol{n} 
ight\|_2 
ight\}.$$

## Distance to a Hyperplane

Define a hyperplane  $\mathcal{P} = \{x : n^T x = b\}$  where  $||n||_2 = 1$ . Let v be an arbitrary point.

The distance between v and  $\mathcal{P}$  is given by

$$d = d(\mathbf{v}, \mathcal{P}) = |\mathbf{n}^T \mathbf{v} - b|.$$
(14)



When  $\|\boldsymbol{n}\|_2 \neq 1$ ,

$$d = \left| \frac{\boldsymbol{n}^T}{\|\boldsymbol{n}\|_2} \boldsymbol{v} - \frac{b}{\|\boldsymbol{n}\|_2} \right| = \frac{\left| \boldsymbol{n}^T \boldsymbol{v} - b \right|}{\|\boldsymbol{n}\|_2} \quad \text{withing the } (15)$$

# SVM: Separate Points from Two Different Classes

Given training dataset  $\{x_i, y_i\}$  where the labels  $y_i \in \{-1, 1\}$ , want to find  $\boldsymbol{\beta}$  and b s.t.

$$\boldsymbol{\beta}^T \boldsymbol{x}_i + b \ge +1$$
 for  $y_i = +1$ ,  
 $\boldsymbol{\beta}^T \boldsymbol{x}_i + b \le -1$  for  $y_i = -1$ .

or equivalently

$$y_i \left( \boldsymbol{\beta}^T \boldsymbol{x}_i + b \right) - 1 \ge 0, \quad \forall i.$$

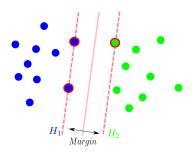
In other words, find a hyperplane  $\{x: \beta^T x - b\}$  s.t.



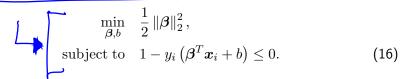


Distance from one class to the hyperplane is  $1/\|\boldsymbol{\beta}\|_2$ .
Distance between the two classes (along direction  $\boldsymbol{\beta}$ ) is  $2/\|\boldsymbol{\beta}\|_2$ .

# SVM: Best Separation



### SVM: a convex optimization problem:



W. Dai (IC)

## Lagrange Dual Problem of SVM

Lagrangian of the SVM primal optimization problem:

$$L = \frac{1}{2} \|\boldsymbol{\beta}\|^2 + \sum_{i} \lambda_i \left(1 - y_i \left(\boldsymbol{\beta}^T \boldsymbol{x}_i + b\right)\right), \tag{17}$$

where  $\lambda_i \geq 0$ .

#### Lagrange Dual Problem



#### The Dual Function

To solve  $\min_{\beta,b} L$ , set  $\partial L/\partial \beta$  and  $\partial L/\partial b$  to zero:

$$\frac{\partial L}{\partial \boldsymbol{\beta}} = \boldsymbol{\beta} - \sum_{i} \lambda_{i} y_{i} \boldsymbol{x}_{i} = 0 \implies \boldsymbol{\beta} = \sum_{i} \lambda_{i} y_{i} \boldsymbol{x}_{i}.$$
 (18)

$$\frac{\partial L}{\partial b} = \sum_{i} \lambda_i y_i = 0. \tag{19}$$

Substitute (18) and (19) into the Lagrangian (17). It holds that 
$$L_D = \sum \lambda_i - \frac{1}{2} \|\boldsymbol{\beta}\|_2^2 = \sum_i \lambda_i - \frac{1}{2} \sum_{i,j} \lambda_i \lambda_j y_i y_j \boldsymbol{x}_i^T \boldsymbol{x}_j$$

$$= -\frac{1}{2} \boldsymbol{\lambda}^T \boldsymbol{K} \boldsymbol{\lambda} + \boldsymbol{1}^T \boldsymbol{\lambda}, \tag{20}$$

where  $K_{i,j} = y_i \boldsymbol{x}_i^T \boldsymbol{x}_j y_j$ .

#### The Dual Problem

The dual problem becomes:

$$\max_{\lambda} -\frac{1}{2} \lambda^{T} K \lambda + \mathbf{1}^{T} \lambda,$$
subject to  $\lambda_{i} \geq 0, \quad \forall i,$ 

$$\sum_{i} \lambda_{i} y_{i} = 0.$$
(21)

#### The KKT Condition

$$\frac{\partial L}{\partial \boldsymbol{\beta}} = \boldsymbol{\beta} - \sum_{i} \lambda_{i} y_{i} \boldsymbol{x}_{i} = 0, \tag{22}$$

$$\frac{\partial L}{\partial b} = \sum_{i} \lambda_{i} y_{i} = 0, \tag{23}$$

$$1 - y_i \left( \boldsymbol{\beta}^T \boldsymbol{x}_i + b \right) \leq 0, \tag{24}$$

$$\lambda_i \geq 0, \tag{25}$$

$$\frac{\lambda_{i} \left(1 - y_{i} \left(\boldsymbol{\beta}^{T} \boldsymbol{x}_{i} + b\right)\right) = 0.}{\boldsymbol{\beta} \cdot \boldsymbol{\beta} \cdot \boldsymbol{\beta} \cdot \boldsymbol{\beta} \cdot \boldsymbol{\beta} \cdot \boldsymbol{\beta} \cdot \boldsymbol{\beta}}$$
(26)

of VIUT conditions,

SVM Classifier: Support Vectors

26) implies 
$$\begin{cases} \text{this slide I will } \\ \text{have understood the} \\ \text{if } \lambda_i \neq 0 \\ \text{if } 1 \neq y_i \left( \boldsymbol{\beta}^T \boldsymbol{x}_i + b \right) \end{cases} \text{ then } 1 = y_i \left( \boldsymbol{\beta}^T \boldsymbol{x}_i + b \right), \text{ concept.}$$

Hence from (22),

Condition (26) implies

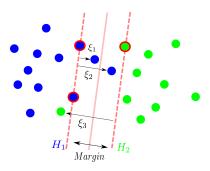
$$\boldsymbol{\beta} = \sum_{i \in \mathcal{I}} \lambda_i y_i \boldsymbol{x}_i, \quad \boldsymbol{\mathcal{I}} = \left\{i: \ y_i \left(\boldsymbol{\beta}^T \boldsymbol{x}_i + b\right) = 1 \quad (\text{or } \lambda_i \neq 0)\right\}.$$
 new test data  $\boldsymbol{x}^{\text{new}}$ , boundary points.

For a new test data  $x^{\text{new}}$ ,

$$y^{\text{new}} = \text{sign}\left(\sum_{i \in \mathcal{I}} \lambda_i y_i \boldsymbol{x}_i^T \boldsymbol{x}^{\text{new}} + b\right).$$

The classifier only uses the boundary points (sparsity!).

# SVM for Overlapping Classes



# Primal Problem for Overlapping Classes

The constraints:

$$\begin{split} \boldsymbol{\beta}^T \boldsymbol{x}_i + b &\geq +1 - \xi_i \quad \text{for } y_i = +1, \\ \boldsymbol{\beta}^T \boldsymbol{x}_i + b &\leq -1 + \xi_i \quad \text{for } y_i = -1, \end{split}$$
 where  $\xi_i \geq 0$ ,  $\forall i$ . We have data points  $\mathcal{T}_i = \mathcal{O}_i$ 

#### SVM Primal Problem:

$$\min_{\boldsymbol{\beta},b,\boldsymbol{\xi}} \quad \frac{1}{2} \|\boldsymbol{\beta}\|_{2}^{2} + C \left(\sum_{i} \xi_{i}\right)^{k}$$
subject to 
$$1 - \xi_{i} - y_{i} \left(\boldsymbol{\beta}^{T} \boldsymbol{x}_{i} + b\right) \leq 0,$$
$$-\xi_{i} \leq 0, \quad \forall i,$$

where C>0 is a constant and k is a positive integer. Usually k=1.

#### **Dual Function**

#### The Lagrangian

$$L = \frac{1}{2} \|\beta\|_{2}^{2} + C \sum \xi_{i} + \sum \lambda_{i} (1 - \xi_{i} - y_{i} (\beta^{T} x_{i} - b)) - \sum u_{i} \xi_{i},$$

where  $\lambda_i > 0$ ,  $u_i > 0$  are Lagrange multipliers.

The dual function

$$L_D = \min_{\boldsymbol{\beta}, b, \boldsymbol{\xi}} L.$$

To find the dual function

$$\begin{split} \frac{dL}{d\boldsymbol{\beta}} &= 0 \quad \Rightarrow \quad \boldsymbol{\beta} = \sum \lambda_i y_i \boldsymbol{x}_i. \\ \frac{dL}{db} &= 0 \quad \Rightarrow \quad \sum \lambda_i y_i = 0. \\ \frac{dL}{d\boldsymbol{\xi}} &= 0 \quad \Rightarrow \quad C - \lambda_i - u_i = 0 \quad \Rightarrow \quad \lambda_i = C - u_i \leq C. \end{split}$$

#### The Dual Problem

The dual problem:

$$\max_{\lambda} \quad \sum \lambda_i - \frac{1}{2} \|\boldsymbol{\beta}\|_2^2 = -\frac{1}{2} \boldsymbol{\lambda}^T \boldsymbol{K} \boldsymbol{\lambda} + \mathbf{1}^T \boldsymbol{\lambda}$$
subject to  $0 \le \lambda_i \le C$ , 
$$\sum \lambda_i y_i = 0$$
,

where  $K_{i,j} = y_i \boldsymbol{x}_i^T \boldsymbol{x}_i y_i$ .

The only difference is that now  $\lambda_i$ 's are upper bounded by C.

Again, only boundary points are involved.

$$\boldsymbol{\beta} = \sum_{i \in \mathcal{I}} \lambda_i y_i \boldsymbol{x}_i, \quad \mathcal{I} = \{i : \lambda_i \neq 0\},$$

which comes from the KKT condition  $\lambda_i \left(1 - \xi_i - y_i \left(\boldsymbol{\beta}^T \boldsymbol{x}_i + b\right)\right) = 0.$ 

#### The General Case

- ► Two classes ⇒ multiple classes
  - Regression
- ▶ Data space ⇒ feature space Define a kernel function  $\varphi: \mathbb{R}^n \to \mathcal{H}$  and work on the space of  $\varphi(x_i)$ .

In SVM, what really matters is  $x_i^T x_i$ . In the general case (kernel method), what matters is

$$\kappa\left(\boldsymbol{x}_{i}, \boldsymbol{x}_{j}\right) = \varphi^{T}\left(\boldsymbol{x}_{i}\right) \varphi\left(\boldsymbol{x}_{j}\right).$$

Example of nonlinear features:

$$\kappa (\mathbf{x}, \mathbf{y}) = (\mathbf{x}^T \mathbf{y})^2 = (x_1 y_1 + x_2 y_2)^2 = x_1^2 y_2^2 + 2x_1 y_1 x_2 y_2 + x_2^2 y_2^2.$$

$$\qquad \qquad \kappa\left(\boldsymbol{x},\boldsymbol{y}\right) = \varphi^{T}\left(\boldsymbol{x}\right)\varphi\left(\boldsymbol{y}\right) \text{ with } \varphi\left(\boldsymbol{x}\right) = \left[x_{1}^{2},\sqrt{2}x_{1}x_{2},x_{2}^{2}\right]^{T}.$$

$$\kappa\left(\boldsymbol{x},\boldsymbol{y}\right) = \exp\left(-\left\|\boldsymbol{x}-\boldsymbol{y}\right\|_{2}^{2}/2\sigma^{2}\right). \quad \text{or small = b watch the heavest neighbours.}$$

## SVM for Regression

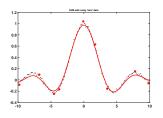
Regression problem: find  $\beta$  and b s.t.

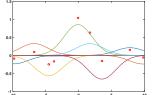
$$y_{i} = f(\mathbf{x}_{i}) = \boldsymbol{\beta}^{T} \varphi(\mathbf{x}_{i}) + b$$

$$= \sum_{j} \lambda'_{j} \varphi^{T}(\mathbf{x}_{j}) \varphi(\mathbf{x}_{i}) + b$$

$$= \sum_{j} \lambda'_{j} \kappa(\mathbf{x}_{i}, \mathbf{x}_{j}) + b.$$

We do not care about 4, but Obout w, the inner product.





# The Primal Optimization Problem

Let  $\epsilon > 0$  be the error tolerance. Then one has

$$\min \quad \frac{1}{2} \|\boldsymbol{\beta}\|_{2}^{2}$$
subject to 
$$\left| y_{i} - \boldsymbol{\beta}^{T} \varphi \left( \boldsymbol{x}_{i} \right) - b \right| \leq \epsilon.$$

The constraints are equivalent to

$$y_i - \boldsymbol{\beta}^T \varphi(\boldsymbol{x}_i) - b \le \epsilon,$$
  
 $\boldsymbol{\beta}^T \varphi(\boldsymbol{x}_i) + b - y_i \le \epsilon.$ 

Now if we allow additional noise, represented by  $\xi_i \geq 0$  and  $\xi_i^{\star} \geq 0$ . Then

$$\min \frac{1}{2} \|\boldsymbol{\beta}\|_{2}^{2} + C \sum_{i} (\xi_{i} + \xi_{i}^{\star})$$
subject to 
$$y_{i} - \boldsymbol{\beta}^{T} \varphi(\boldsymbol{x}_{i}) - b \leq \epsilon + \xi_{i},$$

$$\boldsymbol{\beta}^{T} \varphi(\boldsymbol{x}_{i}) + b - y_{i} \leq \epsilon + \xi_{i}^{\star},$$

$$-\xi_{i} \leq 0, \quad -\xi_{i}^{\star} \leq 0.$$

## Lagrangian

$$L = \frac{1}{2} \|\boldsymbol{\beta}\|_{2}^{2} + C \sum_{i} (\xi_{i} + \xi_{i}^{*}) - \sum_{i} (u_{i}\xi_{i} + \sum_{i} u_{i}^{*}\xi_{i}^{*})$$
$$+ \lambda_{i} (y_{i} - \boldsymbol{\beta}^{T} \varphi (\boldsymbol{x}_{i}) - b - \epsilon - \xi_{i})$$
$$+ \lambda_{i}^{*} (\boldsymbol{\beta}^{T} \varphi (\boldsymbol{x}_{i}) + b - y_{i} - \epsilon - \xi_{i}^{*}),$$

where  $u_i, u_i^{\star}, \xi_i, \xi_i^{\star} \geq 0$  are Lagrange multiplier. To minimize L,

$$dL/d\boldsymbol{\beta} = \mathbf{0} \quad \Rightarrow \quad \boldsymbol{\beta} = \sum_{i} (\lambda_{i} - \lambda_{i}^{\star}) \, \varphi \left(\boldsymbol{\xi}_{i}\right),$$
 
$$dL/db = \mathbf{0} \quad \Rightarrow \quad \sum_{i} \lambda_{i} = \sum_{i} \lambda_{i}^{\star},$$
 
$$dL/d\xi_{i} = 0, \ dL/d\xi_{i}^{\star} = 0 \quad \Rightarrow \quad \lambda_{i} \leq C, \ \lambda_{i}^{\star} \leq C.$$

#### The Dual Problem

The objective function of the dual problem

$$L_D = -\epsilon \sum_{i,j} (\lambda_i + \lambda_i^*) + y_i \sum_{i,j} (\lambda_i - \lambda_i^*) - \frac{1}{2} \sum_{i,j} (\lambda_i - \lambda_i^*) (\lambda_j - \lambda_j^*) \kappa(\mathbf{x}_i, \mathbf{x}_j),$$

$$\|\boldsymbol{\beta}\|_2^2$$

where 
$$\kappa\left(\boldsymbol{x}_{i}, \boldsymbol{x}_{j}\right) = \varphi^{T}\left(\boldsymbol{x}_{i}\right) \varphi\left(\boldsymbol{x}_{j}\right)$$
 .

The optimizatoin constraints are

$$\sum_{i} (\lambda_i - \lambda_i^*) = 0,$$
  
$$0 \le \lambda_i, \lambda_i^* \le C.$$

## KKT Condition and Support Vectors

Part of the KKT condition is that  $\forall i$ ,

$$\begin{cases} \lambda_{i} \left( y_{i} - \boldsymbol{\beta}^{T} \boldsymbol{\varphi} \left( \boldsymbol{x}_{i} \right) - b - \epsilon - \xi_{i} \right) = 0, \\ \lambda_{i}^{\star} \left( \boldsymbol{\beta}^{T} \boldsymbol{\varphi} \left( \boldsymbol{x}_{i} \right) + b - y_{i} - \epsilon - \xi_{i}^{\star} \right) = 0. \end{cases}$$

- ▶ Interior points:  $\left|y_i \boldsymbol{\beta}^T \varphi\left(\boldsymbol{x}_i\right) b\right| < \epsilon + \xi_i$ .
  - ▶ Both  $\lambda_i$  and  $\lambda_i^*$  are zero.
- ▶ Boundary points:  $|y_i \boldsymbol{\beta}^T \varphi(\boldsymbol{x}_i) b| = \epsilon + \xi_i$ .
  - One of  $\lambda_i$  and  $\lambda_i^{\star}$  is zero.
  - $\lambda_i \neq \lambda_i^{\star}.$

#### The Standard Form

Let  $\gamma_i = \lambda_i$  and  $\gamma_{i+n} = \lambda_i^{\star}$  (Merge  $\lambda$  and  $\lambda^{\star}$  into a single vector). The dual problem becomes

$$\min_{\boldsymbol{\gamma}} \quad \frac{1}{2} \boldsymbol{\gamma}^T \boldsymbol{Q} \boldsymbol{\gamma} + \boldsymbol{v}^T \boldsymbol{\gamma},$$
subject to  $0 \le \gamma_i \le C, \quad \sum_{i=1}^n \gamma_i - \sum_{i=n+1}^{2n} \gamma_i = 0.$ 

The boundary points are given by  $\mathcal{I} = \{i : \gamma_i - \gamma_{i+n} \neq 0\}.$ 

For a new data point  $oldsymbol{x}^{\mathrm{new}}$ , the regression is

$$f(\boldsymbol{x}^{\mathrm{new}}) = \sum_{i} (\gamma_i - \gamma_{i+n}) \kappa(\boldsymbol{x}_i, \boldsymbol{x}^{\mathrm{new}}) + b.$$

## Section 10 Gaussian Distribution

Statistical learning.

#### Gaussian Random Vectors

A random vector  $m{X} \in \mathbb{R}^n$  is Gaussian distributed  $m{X} \sim \mathcal{N}\left(m{\mu}, m{\Sigma}\right)$  if its pdf is given by

$$p(\mathbf{x}) = |2\pi \mathbf{\Sigma}|^{-\frac{1}{2}} \exp\left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right),$$

where  $\mu \in \mathbb{R}^n$  and  $\Sigma \in \mathcal{S}^n_+$  (the set of  $n \times n$  symmetric positive semidefinite matrices).

Here, we have assumed that  $\Sigma$  is invertible (of full rank).

#### Gaussian Random Vectors: Characteristic Function

 $\mathsf{PDF} \begin{tabular}{l}{\mathbf{Fourier}}{\mathbf{Transform}}\\ \rightleftarrows\\ \mathbf{Inverse}\ \mathbf{Fourier}\ \mathbf{Transform}\\ \end{bmatrix} \mathsf{Characteristic}\ \mathsf{function}\ \mathbf{E}\left[e^{i\langle \pmb{\lambda}, \pmb{X}\rangle}\right].$ 

$$oldsymbol{X} \sim \mathcal{N}\left(oldsymbol{\mu}, oldsymbol{\Sigma}
ight)$$
 if

$$\mathrm{E}\left[e^{i\langle oldsymbol{\lambda}, oldsymbol{X}
angle}
ight] = \exp\left(i\,\langle oldsymbol{\lambda}, oldsymbol{\mu}
angle - rac{1}{2}oldsymbol{\lambda}^Toldsymbol{\Sigma}oldsymbol{\lambda}
ight).$$

It is well defined even when  $\Sigma$  is not invertible.

Quadratic form in both the pot and characteristic function.

#### Affine Transformation

#### Lemma 10.1

Let  $X \sim \mathcal{N}\left(oldsymbol{\mu}, oldsymbol{\Sigma}
ight)$ . Then for any  $oldsymbol{A} \in \mathbb{R}^{m imes n}$  and  $oldsymbol{b} \in \mathbb{R}^m$ ,

$$AX + b \sim \mathcal{N}\left(A\mu + b, A\Sigma A^T\right).$$

#### Proof:

$$\begin{split} \mathbf{E}\left[e^{i\langle\boldsymbol{\lambda},\boldsymbol{A}\boldsymbol{X}+\boldsymbol{b}\rangle}\right] &= \mathbf{E}\left[e^{i\langle\boldsymbol{A}^T\boldsymbol{\lambda},\boldsymbol{X}\rangle+i\langle\boldsymbol{\lambda},\boldsymbol{b}\rangle}\right] \\ &= \exp\left(i\langle\boldsymbol{A}^T\boldsymbol{\lambda},\boldsymbol{\mu}\rangle - \frac{1}{2}\left(\boldsymbol{A}^T\boldsymbol{\lambda}\right)^T\boldsymbol{\Sigma}\left(\boldsymbol{A}^T\boldsymbol{\lambda}\right)\right)e^{i\langle\boldsymbol{\lambda},\boldsymbol{b}\rangle} \\ &= \exp\left(i\langle\boldsymbol{\lambda},\boldsymbol{A}^T\boldsymbol{\mu}+\boldsymbol{b}\rangle - \frac{1}{2}\boldsymbol{\lambda}^T\left(\boldsymbol{A}\boldsymbol{\Sigma}\boldsymbol{A}^T\right)\boldsymbol{\lambda}\right). \end{split}$$

## Gaussian Conditioning Lemma

#### Lemma 10.2

Let  $X \sim \mathcal{N}\left(\mathbf{0}, \mathbf{\Sigma}\right)$ .

Let  $m{X}_{\mathcal{A}}$  and  $m{X}_{\mathcal{B}}$  be two subvectors of  $m{X}$ , i.e.,  $m{X} = m{\left[ m{X}_{\mathcal{A}}^T, m{X}_{\mathcal{B}}^T 
ight]}^T$ .

Let 
$$K:=\mathbf{\Sigma}^{-1}=\left[egin{array}{cc} K_{\mathcal{A}\mathcal{A}} & K_{\mathcal{A}\mathcal{B}} \ K_{\mathcal{B}\mathcal{A}} & K_{\mathcal{B}\mathcal{B}} \end{array}
ight]$$
 be the precision matrix.

Then  $X_{\mathcal{A}}|X_{\mathcal{B}} \sim P_{X_{\mathcal{A}}|X_{\mathcal{B}}} = \mathcal{N}\left(-K_{\mathcal{A}\mathcal{A}}^{-1}K_{\mathcal{A}\mathcal{B}}X_{\mathcal{B}}, K_{\mathcal{A}\mathcal{A}}^{-1}\right)$ . In other words,

$$X_{\mathcal{A}} = -K_{\mathcal{A}\mathcal{A}}^{-1}K_{\mathcal{A}\mathcal{B}}X_{\mathcal{B}} + \epsilon,$$

where  $\epsilon \sim \mathcal{N}\left(0, \boldsymbol{K}_{\mathcal{A}\mathcal{A}}^{-1}\right)$  is independent of  $\boldsymbol{X}_{\mathcal{B}}$ .

Remark:  $K_{\mathcal{A}\mathcal{A}}^{-1} \neq \Sigma_{\mathcal{A}\mathcal{A}}$ .

#### Matrix Identities

Block matrix inverse (BMI)

$$\begin{bmatrix} A & B \\ C & D \end{bmatrix}^{-1} = \begin{bmatrix} (A - BD^{-1}C)^{-1} & -A^{-1}B(D - CA^{-1}B)^{-1} \\ -D^{-1}C(A - BD^{-1}C)^{-1} & (D - CA^{-1}B)^{-1} \end{bmatrix}.$$
(27)

Woodbury matrix identity (WMI)

$$(A + UCV)^{-1} = A^{-1} - A^{-1}U(C^{-1} + VA^{-1}U)^{-1}VA^{-1}.$$
 (28)

## Proof of Gaussian Conditioning Lemma

By Bayes rule,  $p\left(\boldsymbol{x}_{\mathcal{A}}|\boldsymbol{x}_{\mathcal{B}}\right) = p\left(\boldsymbol{x}_{\mathcal{A}},\boldsymbol{x}_{\mathcal{B}}\right)/p\left(\boldsymbol{x}_{\mathcal{B}}\right)$ . Then

$$\ln p\left(\boldsymbol{x}_{\mathcal{A}}|\boldsymbol{x}_{\mathcal{B}}\right) = \ln p\left(\boldsymbol{x}_{\mathcal{A}},\boldsymbol{x}_{\mathcal{B}}\right) - \ln p\left(\boldsymbol{x}_{\mathcal{B}}\right)$$
$$= c - \frac{1}{2}\boldsymbol{x}_{\mathcal{A}}^{T}\boldsymbol{K}_{\mathcal{A}\mathcal{A}}\boldsymbol{x}_{\mathcal{A}} - \boldsymbol{x}_{\mathcal{A}}^{T}\boldsymbol{K}_{\mathcal{A}\mathcal{B}}\boldsymbol{x}_{\mathcal{B}} - \frac{1}{2}\boldsymbol{x}_{\mathcal{B}}^{T}\left(\boldsymbol{K}_{\mathcal{B}\mathcal{B}} - \boldsymbol{\Sigma}_{\mathcal{B}\mathcal{B}}^{-1}\right)\boldsymbol{x}_{\mathcal{B}},$$

where c is a constant. By (27),

$$\Sigma_{\mathcal{B}\mathcal{B}}^{-1} = K_{\mathcal{B}B} - K_{\mathcal{B}A}K_{\mathcal{A}A}^{-1}K_{\mathcal{A}B}.$$

One has

$$\ln p\left(\boldsymbol{x}_{\mathcal{A}}|\boldsymbol{x}_{\mathcal{B}}\right) = c - \frac{1}{2}\left(\boldsymbol{x}_{\mathcal{A}} + \boldsymbol{K}_{\mathcal{A}\mathcal{A}}^{-1}\boldsymbol{K}_{\mathcal{A}\mathcal{B}}\boldsymbol{x}_{\mathcal{B}}\right)^{T}\boldsymbol{K}_{\mathcal{A}\mathcal{A}}\left(\boldsymbol{x}_{\mathcal{A}} + \boldsymbol{K}_{\mathcal{A}\mathcal{A}}^{-1}\boldsymbol{K}_{\mathcal{A}\mathcal{B}}\boldsymbol{x}_{\mathcal{B}}\right).$$

That is,  $X_{\mathcal{A}}|X_{\mathcal{B}} \sim \mathcal{N}\left(-K_{\mathcal{A}\mathcal{A}}^{-1}K_{\mathcal{A}\mathcal{B}}X_{\mathcal{B}},K_{\mathcal{A}\mathcal{A}}^{-1}\right)$ .

## A Signal Processing Application

#### The problem:

Given

$$Y = AX + W$$

where  $X \sim \mathcal{N}\left(\mathbf{0}, \mathbf{\Sigma}_{x}\right)$  and  $W \sim \mathcal{N}\left(\mathbf{0}, \mathbf{\Sigma}_{w}\right)$ .

Given observation  $\boldsymbol{y}$ , want to find  $\hat{\boldsymbol{x}} = f\left(\boldsymbol{y}\right)$  s.t. the mean squared error  $\mathrm{E}\left[\|\boldsymbol{x} - \hat{\boldsymbol{x}}\|_2^2\right]$  is minimized (MMSE solution).

Fact: The general MMSE solution is given by

$$\hat{\boldsymbol{x}} = \mathrm{E}\left[\boldsymbol{X}|\boldsymbol{Y} = \boldsymbol{y}\right] = \int \boldsymbol{x} \cdot p_{\boldsymbol{X}|\boldsymbol{Y}}\left(\boldsymbol{x}|\boldsymbol{y}\right) d\boldsymbol{x}.$$

Hence for Gaussian random variables, Gaussian conditioning lemma can be used.

## Finding the MMSE Solution

1. Y = AX + W is Gaussian distributed  $\mathcal{N}\left(\mathbf{0}, A\Sigma_x A^T + \Sigma_w\right)$ .

2.

$$\left[egin{array}{c} oldsymbol{X} \ oldsymbol{Y} \end{array}
ight] \sim \mathcal{N} \left(oldsymbol{0}, \left[egin{array}{ccc} oldsymbol{\Sigma}_x & oldsymbol{\Sigma}_x oldsymbol{A}^T \ oldsymbol{A} oldsymbol{\Sigma}_x & oldsymbol{A} oldsymbol{\Sigma}_x oldsymbol{A}^T + oldsymbol{\Sigma}_w \end{array}
ight]
ight).$$

3. Find the precision matrix from  $\Sigma$ :

$$\boldsymbol{K} = \left[ \begin{array}{cc} \boldsymbol{\Sigma}_x^{-1} + \boldsymbol{A}^T \boldsymbol{\Sigma}_w^{-1} \boldsymbol{A} & -\boldsymbol{A}^T \boldsymbol{\Sigma}_w^{-1} \\ -\boldsymbol{\Sigma}_w^{-T} \boldsymbol{A} & \text{sth} \end{array} \right]$$

4.  $X|Y \sim \mathcal{N}\left(-K_{\mathcal{A}\mathcal{A}}^{-1}K_{\mathcal{A}\mathcal{B}}Y,K_{\mathcal{A}\mathcal{A}}^{-1}\right)$  by Gaussian Conditioning Lemma.

We use the conditional mean as the estimate  $\hat{x}$ :

$$\hat{\boldsymbol{x}} = \left(\boldsymbol{\Sigma}_x^{-1} + \boldsymbol{A}^T \boldsymbol{\Sigma}_w^{-1} \boldsymbol{A}\right)^{-1} \boldsymbol{A}^T \boldsymbol{\Sigma}_w^{-1} \boldsymbol{y}. \tag{29}$$

$$\Sigma_{X|Y} = \boldsymbol{K}_{AA}^{-1} = \left(\Sigma_x^{-1} + \boldsymbol{A}^T \Sigma_w^{-1} \boldsymbol{A}\right)^{-1}.$$
 (30)

#### Calculation of The K Matrix

$$K_{\mathcal{A}\mathcal{A}} \stackrel{\mathrm{BMI}(27)}{=} \left( \mathbf{\Sigma}_{x} - \mathbf{\Sigma}_{x} \mathbf{A}^{T} \left( \mathbf{A} \mathbf{\Sigma}_{x} \mathbf{A}^{T} + \mathbf{\Sigma}_{w} \right)^{-1} \mathbf{A} \mathbf{\Sigma}_{x} \right)^{-1}$$

$$\stackrel{\mathrm{WMI}(28)}{=} \left( \left( \mathbf{\Sigma}_{x}^{-1} + \mathbf{A}^{T} \mathbf{\Sigma}_{w}^{-1} \mathbf{A} \right)^{-1} \right)^{-1}$$

$$= \mathbf{\Sigma}_{x}^{-1} + \mathbf{A}^{T} \mathbf{\Sigma}_{w}^{-1} \mathbf{A}.$$

$$egin{aligned} oldsymbol{K}_{\mathcal{A}\mathcal{B}} &\overset{ ext{BMI}(27)}{=} -oldsymbol{\Sigma}_x^{-1} \left(oldsymbol{\Sigma}_x oldsymbol{A}^T
ight) \left(oldsymbol{A}oldsymbol{\Sigma}_x oldsymbol{A}^T + oldsymbol{\Sigma}_w - oldsymbol{A}oldsymbol{\Sigma}_x oldsymbol{\Sigma}_x^{-1} oldsymbol{\Sigma}_x oldsymbol{A}^T
ight)^{-1} \ &= -oldsymbol{A}^T oldsymbol{\Sigma}_w^{-1}. \end{aligned}$$

Hence 
$$\mathbf{\Sigma}_{X|Y} = \left(\mathbf{\Sigma}_x^{-1} + \mathbf{A}^T\mathbf{\Sigma}_w^{-1}\mathbf{A}\right)^{-1}$$
 and  $\mathbf{L} = \mathbf{\Sigma}_{X|Y}\mathbf{A}^T\mathbf{\Sigma}_w^{-1}$ .

# Section 11 Sparse Bayesian Learning

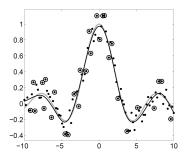
Didn't do it.

## Sparse Bayesian Learning (SBL) via Relevance Vector Machine (RVM)

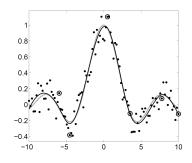
M. E. Tipping, "Sparse Bayesian learning and the relevance vector machine," J. Mach. Learn. Res., JMLR.org, 2001, 1, 211-244.

#### Examples from Tipping's paper:

To approximate the sinc function



Support vector approx.



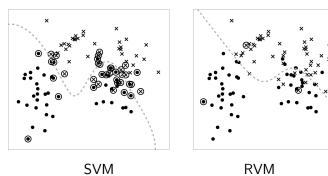
Relevance vector approx.

## Sparse Bayesian Learning (SBL) via Relevance Vector Machine (RVM)

M. E. Tipping, "Sparse Bayesian learning and the relevance vector machine," J. Mach. Learn. Res., JMLR.org, 2001, 1, 211-244.

#### Examples from Tipping's paper:

Classify two classes



#### RVM: A Hierarchical Gaussian Model

Target: Solve the sparse linear inverse problem:

$$y = \Phi x + w$$
.

Signal model: signal components are i.i.d. Gaussian

$$p(X_i|\alpha_i) = \mathcal{N}(X_i; 0, \alpha_i^{-1}).$$

- $ightharpoonup lpha_i > 0$ : precision  $(\alpha_i^{-1} = \sigma_i^2)$ , for mathematical convenience)
- $ightharpoonup lpha_i = \infty$ :  $X_i \sim \mathcal{N}\left(0,0\right)$  and  $X_i = 0$  with prob. one.
- $\bullet$   $\alpha_i \in \mathbb{R}^+$ :  $X_i \sim \mathcal{N}\left(0, \alpha_i^{-1}\right)$  and  $X_i \neq 0$  with prob. one.

### Signal Estimation with Known lpha

If we know lpha, then the posterior of x can be derived by Bayes' rule

$$p\left(\boldsymbol{x}|\boldsymbol{y}\right) = \frac{p\left(\boldsymbol{x}\right)p\left(\boldsymbol{y}|\boldsymbol{x}\right)}{p\left(\boldsymbol{y}\right)} = \mathcal{N}\left(\boldsymbol{\mu}, \boldsymbol{\Sigma}\right),$$

where

$$\boldsymbol{\mu} = \sigma_w^{-2} \boldsymbol{\Sigma} \boldsymbol{\Phi}^T \boldsymbol{y}, \ \boldsymbol{\Sigma} = \left( \boldsymbol{A} + \sigma_w^{-2} \boldsymbol{\Phi}^T \boldsymbol{\Phi} \right)^{-1}, \ \boldsymbol{A} = \operatorname{diag} \left( \boldsymbol{\alpha} \right).$$

See (29) and (30) for details.

Set  $\hat{x} = \mu$ . This is the MMSE solution as well.

## Signal Recovery with Unknown lpha

#### The key is to estimate $\alpha$ :

- $lacksquare oldsymbol{Y} \sim \mathcal{N}\left(oldsymbol{0}, \sigma_w^2 oldsymbol{I} + oldsymbol{\Phi} oldsymbol{A}^{-1} oldsymbol{\Phi}^T
  ight)$
- lacktriangle Linear combination of Gaussian is Gaussian (y=Ax+w).
- Maximum likelihood (ML) estimate:

$$\begin{split} \mathcal{L}\left(\boldsymbol{\alpha}\right) &= \log p\left(\boldsymbol{y}|\boldsymbol{\alpha},\sigma_w^2\right) = -\frac{1}{2}\left[N\log 2\pi + \log |\boldsymbol{C}| + \boldsymbol{y}^T\boldsymbol{C}^{-1}\boldsymbol{y}\right], \\ \text{where } \boldsymbol{C} &= \sigma_w^2\boldsymbol{I} + \boldsymbol{\Phi}\boldsymbol{A}^{-1}\boldsymbol{\Phi}^T. \end{split}$$

► Marginal likelihood maximization: For a given *i*, we solve

$$\max_{\alpha_i} \mathcal{L}\left(\boldsymbol{\alpha}\right)$$

by fixing all  $\alpha_j$ 's,  $j \neq i$ .

## Marginal Likelihood Maximization (1)

$$\begin{split} \boldsymbol{C} &= \sigma^2 \boldsymbol{I} + \boldsymbol{\Phi} \boldsymbol{A}^{-1} \boldsymbol{\Phi} \\ &= \sigma^2 \boldsymbol{I} + \sum_{j \neq i} \alpha_j^{-1} \boldsymbol{\phi}_j \boldsymbol{\phi}_j^T + \alpha_i^{-1} \boldsymbol{\phi}_i \boldsymbol{\phi}_i^T \\ &= \boldsymbol{C}_{-i} + \alpha_i^{-1} \boldsymbol{\phi}_i \boldsymbol{\phi}_i^T. \end{split}$$

One has

$$|C| = |C_{-i}| \cdot \left| I + \alpha_i^{-1} C_{-i}^{-1/2} \phi_i \phi_i^T C_{-i}^{-1/2} \right|$$
  
=  $|C_{-i}| \cdot \left| 1 + \alpha_i^{-1} \phi_i^T C_{-i}^{-1} \phi_i \right|.$ 

The last line comes from  $\left| I + MM^T \right| = \left| I + M^TM \right|$ , which can be easily verified by the SVD of M.

$$C^{-1} \stackrel{\text{WMI(28)}}{=} C_{-i}^{-1} - \frac{C_{-i}^{-1}\phi_i\phi_i^T C_{-i}^{-1}}{\alpha_i + \phi_i^T C_{-i}^{-1}\phi_i}.$$

## Marginal Likelihood Maximization (2)

$$\mathcal{L}(\boldsymbol{\alpha}) = -\frac{1}{2} \left[ N \log 2\pi + \log |\boldsymbol{C}| + \boldsymbol{y}^T \boldsymbol{C}^{-1} \boldsymbol{y} \right],$$

$$= c - \frac{1}{2} \left[ \log |\boldsymbol{C}_{-i}| + \boldsymbol{y}^T \boldsymbol{C}_{-i}^{-1} \boldsymbol{y} \right]$$

$$- \log \alpha_i + \log \left( \alpha_i + \boldsymbol{\phi}_i^T \boldsymbol{C}_{-i}^{-1} \boldsymbol{\phi}_i \right) - \frac{\left( \boldsymbol{\phi}_i^T \boldsymbol{C}_{-i}^{-1} \boldsymbol{y} \right)^2}{\alpha_i + \boldsymbol{\phi}_i^T \boldsymbol{C}_{-i}^{-1} \boldsymbol{\phi}_i} \right]$$

$$= \mathcal{L}(\boldsymbol{\alpha}_{-i}) + \frac{1}{2} \left[ \log \alpha_i - \log \left( \alpha_i + s_i \right) - \frac{q_i^2}{\alpha_i + s_i} \right]$$

$$= \mathcal{L}(\boldsymbol{\alpha}_{-i}) + \ell(\alpha_i),$$

where for simplification of expressions, we have defined

$$s_i := \boldsymbol{\phi}_i^T \boldsymbol{C}_{-i}^{-1} \boldsymbol{\phi}_i, \text{ and } q_i := \boldsymbol{\phi}_i^T \boldsymbol{C}_{-i}^{-1} \boldsymbol{y}.$$

Set the derivative of  $\ell(\alpha_i)$  to zero. One obtains the closed form solution for the optimal  $\alpha_i$ :

$$\alpha_i = \begin{cases} \frac{s_i^2}{q_i^2 - s_i} & \text{if } q_i^2 > s_i, \\ \infty & \text{if } q_i^2 \le s_i, \end{cases}$$

## Algorithms for RVM

#### A sequential algorithm: [Tipping & Faul, 2003]

In each iteration:

- ▶ Scan  $i \in [N]$  and find the i that  $\mathcal{L}(\alpha_i^*) \mathcal{L}(\alpha_i)$  is maximized.
- $\blacktriangleright \mathsf{Set} \ \alpha_i = \alpha_i^*.$
- ▶ Update  $\mathcal{L}$  and other parameters (preparation for the next iteration).

## Connections to Greedy Algorithms

- Similar to greedy algorithms
  - The sequential algorithm is very similar to OMP.
  - ► "Subspace pursuit" type of RVM: [Karseras & D., 2013]
  - Performance guarantees: sufficient conditions based on mutual coherence or RIP [Karseras & D., 2013]
- Different from greedy algorithms
  - Noise variance is considered.
  - ▶ Statistical information: estimation quality information

#### A Demo

#### Track global Ozone density

- Spatial sparsity:
  - Each frame is sparse under DCT transform
- ► Temporal sparsity: correlations across days
  - Support innovation.
  - Nonzero component magnitude innovation.

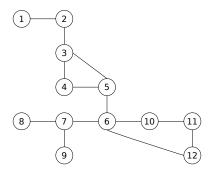
Let  $x^t$  be the DCT coefficient vector at the t-th day:

$$egin{aligned} oldsymbol{x}^t &= oldsymbol{x}^{t-1} + oldsymbol{u}^t, \ oldsymbol{y}^t &= oldsymbol{A} oldsymbol{D} oldsymbol{x}^t + oldsymbol{w}^t, \end{aligned}$$

where  $u^t \sim \prod \mathcal{N}\left(0, \alpha_i^{-1}\right)$ , A is the sampling matrix, and  $w^t$  is the white Gaussian noise.

## Section 12 Gaussian Graphic Model

## Motivation: Gaussian Graphic Model



Encoding the conditional dependencies between n random variables  $X_1, \dots, X_n$  by a graph.

## Correlation and Conditional Independence

Sneeze — Catch Cold — Weather Change

Observation: "Weather Change" and "Sneeze" are correlated.

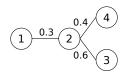
- "Weather Change" and "Catch Cold" are highly correlated.
- "Catch Cold" and "Sneeze" are highly correlated.

However, given the status of "Catch Cold", "Weather Change" and "Sneeze" are independent.

- ► Given that "Catch Cold" is false, "Sneeze" is likely to be false, independent of whether "Weather Change" is true or not.
- ► Given that "Catch Cold" is true, "Sneeze" is likely to be true, independent of whether "Weather Change" is true or not.

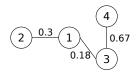
#### Other Examples

Suppose that  $\rho\left(X_1,X_2\right)=0.3$ ,  $\rho\left(X_1,X_3\right)=0.18$ , and  $\rho\left(X_1,X_4\right)=0.12$ . Suppose that on one day,  $X_2\uparrow 0.2$ ,  $X_3\downarrow 0.1$ , and  $X_4\downarrow 0.5$ . Find the expected change of  $X_1$ .



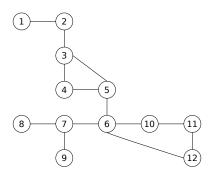
$$E[\Delta X_1] = 0.2 \times 0.3 = 0.06.$$

$$\mathrm{E}\left[\Delta X_1\right] = 0.2 \times 0.3 - 0.1 \times 0.18 - 0.5 \times 0.12 \\ = -0.018.$$



$$\mathrm{E}\left[\Delta X_1\right] = 0.2 \times 0.3 - 0.1 \times 0.18$$
  
= 0.042.

## Nondirected Graphical Model



The distribution of the Gaussian random vector  $\boldsymbol{X} = [X_1, \cdots, X_n]^T$  is a graphic model according to the graph g if

for all  $a: X_a \perp \{X_b: b \notin \operatorname{ne}(a), b \neq a\}$  given  $\{X_c: c \in \operatorname{ne}(a)\}$ .

Or, given  $X_c$ 's,  $c \in ne(a)$ ,  $X_a$  and  $X_b$ 's are independent for all b not in the neighborhood.

## Consequence of Gaussian Conditioning

Recall the Gaussian conditioning lemma (Lemma 10.2). Let  $\boldsymbol{K}$  be the precision matrix of  $\boldsymbol{X}$ .

#### Corollary 12.1

For any  $a \in [n]$ ,

$$X_a = -\sum_{b: b \neq a} \frac{K_{ab}}{K_{aa}} X_b + \epsilon_a,$$

where  $\epsilon_a \sim \mathcal{N}\left(0, K_{aa}^{-1}\right)$  is independent of  $\{X_b: b \neq a\}$ .

Proof: Apply Lemma 10.2 with  $A=\{a\}$  and  $B=[n]\setminus\{a\}=A^c$ .

Remark: Find the neighboring points.

#### Conditional Correlation

#### Corollary 12.2

$$\operatorname{cor}(X_a, X_b | \mathbf{X}_{\mathcal{C}}) = -\frac{K_{ab}}{\sqrt{K_{aa}K_{bb}}}.$$

Proof: From Gaussian Conditioning (Lemma 10.2), it holds that

$$\operatorname{cov}\left(\boldsymbol{X}_{\{a,b\}}|\boldsymbol{X}_{\mathcal{C}}\right) = \begin{bmatrix} K_{aa} & K_{ab} \\ K_{ba} & K_{bb} \end{bmatrix}^{-1} = \frac{1}{K_{aa}K_{bb} - K_{ab}^{2}} \begin{bmatrix} K_{bb} & -K_{ba} \\ -K_{ab} & K_{aa} \end{bmatrix}.$$

Plug this formula into the definition of conditional correlation. Corollary 12.2 is proved.

Remark: Find the correlation between neighboring points.

#### Estimate the Precision Matrix

From the definition  $K=\Sigma^{-1}$ , the computation seems straightforward. However, the commonly used fact

$$\frac{1}{m}\sum (X - \bar{X})(X - \bar{X})^T \to \Sigma$$
 (31)

is based on the assumption that n is fixed and  $m \to \infty$ .

In reality, we may not have sufficient data m. Hence (31) may not be applicable.

Assumption: K is sparse.

## Estimation via Regression (1)

Define the matrix  $\Theta$  by  $\theta_{ab} = -K_{ab}/K_{bb}$  for  $b \neq a$  and  $\theta_{aa} = 0$ . Then Corollary 12.1 implies

$$E[X_a|X_b: b \neq a] = \sum_b \theta_{ba} X_b.$$

Hence we need to find  $\theta_{ba}$ 's  $(b \neq a)$  to minimize

$$\mathrm{E}\left[\left(X_a - \sum_b \theta_{ba} X_b\right)^2\right].$$

Or in matrix format

$$\hat{\boldsymbol{\Theta}} = \arg \min_{\boldsymbol{\Theta} \in \boldsymbol{\Theta}} \mathbf{E} \left[ \left\| \boldsymbol{X} - \boldsymbol{\Theta}^T \boldsymbol{X} \right\|_2^2 \right],$$

where  $\Theta = \{ \Theta : \operatorname{diag}(\Theta) = \mathbf{0} \}.$ 

## Estimation via Regression (2)

The objective function can be rewritten as

$$E\left[\left\|\boldsymbol{X} - \boldsymbol{\Theta}^{T} \boldsymbol{X}\right\|_{2}^{2}\right] \approx \frac{1}{m} \sum \left(\boldsymbol{x} - \boldsymbol{\Theta}^{T} \boldsymbol{x}\right)^{T} \left(\boldsymbol{x} - \boldsymbol{\Theta}^{T} \boldsymbol{x}\right)$$

$$= \frac{1}{m} \left\| \begin{bmatrix} \boldsymbol{x}_{(1)}^{T} \\ \vdots \\ \boldsymbol{x}_{(m)}^{T} \end{bmatrix} - \begin{bmatrix} \boldsymbol{x}_{(1)}^{T} \\ \vdots \\ \boldsymbol{x}_{(m)}^{T} \end{bmatrix} \boldsymbol{\Theta} \right\|_{F}^{2}$$

$$= \frac{1}{m} \left\|\boldsymbol{X} - \boldsymbol{X} \boldsymbol{\Theta}\right\|_{F}^{2}.$$

Note that the X on this slide is the data matrix and the X on previous slides are random vectors.

## Estimation via Regression (3)

The overall optimization problem:

$$\min_{\boldsymbol{\Theta} \in \boldsymbol{\Theta}} \quad \frac{1}{m} \left\| \boldsymbol{X} - \boldsymbol{X} \boldsymbol{\Theta} \right\|_F^2 + \lambda \sum_{a \neq b} \left| \theta_{ab} \right|,$$

Or

$$\min_{\boldsymbol{\Theta} \in \boldsymbol{\Theta}} \quad \frac{1}{m} \left\| \boldsymbol{X} - \boldsymbol{X} \boldsymbol{\Theta} \right\|_F^2 + \lambda \sum_{a < b} \sqrt{\theta_{ab}^2 + \theta_{ba}^2}.$$