

# Robust Convex Clustering Analysis

## Supplemental Materials

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### APPENDIX

Here we show the proof of Theorem 1.

Assume a data matrix  $X \in \mathbb{R}^{d \times n}$  with  $n$  samples and  $d$  features satisfy:

$$X = P^* + Q^* + \delta, \quad (12)$$

where  $P^* \in \mathbb{R}^{d \times n}$  and  $Q^* \in \mathbb{R}^{d \times n}$  are the underlying true decomposition of  $X$ .  $P^*$  is the component containing cluster information of  $X$ .  $Q^*$  is the component containing outliers.  $\delta \in \mathbb{R}^{d \times n}$  is the stochastic noise matrix. For the  $i^{th}$  row of  $\delta$ ,  $\delta_{i,j} \sim \mathcal{N}(0, \sigma^2)$ , where  $i \in \mathbb{N}_d$  and  $j \in \mathbb{N}_n$ , i.e., for each feature there exists a normal distributed noise for all data points.

The optimization problem in Eq. (2) is:

$$(\hat{P}, \hat{Q}) = \underset{P, Q}{\operatorname{argmin}} \frac{1}{2} \|X - (P + Q)\|_F^2 + \alpha \|P\|_{FU} + \beta \|Q\|_{2,1}, \quad (13)$$

where  $\hat{P} \in \mathbb{R}^{d \times n}$  and  $\hat{Q} \in \mathbb{R}^{d \times n}$  are the optimal solution pair obtained by solving Eq. (13) and  $\|P\|_{FU}$  is defined as:

$$\|P\|_{FU} = \sum_{i < j} w_{i,j} \|P_i - P_j\|_2.$$

We first present a theorem for the optimal solution pair which are important for our following theoretical analysis.

**Theorem 2.** Consider the optimization problem in Eq. (2) for  $n, d > 1$ . Take the regularization parameters  $\alpha$  and  $\beta$  as:

$$\alpha > \lambda, \beta > \frac{\lambda}{\sqrt{n}}, \lambda = \sigma \sqrt{n(d+t)}, \quad (14)$$

where  $t$  is a positive scalar and  $\sigma$  is the standard deviation for each row of  $\delta$ . Then with a probability of at least  $1 - \exp(-\frac{1}{2}(t - d \log(1 + \frac{t}{d})))$ , for a global minimizer  $\hat{P}, \hat{Q}$  in Eq. (13) we have:

$$\begin{aligned} \frac{1}{2} \|(P^* + Q^*) - (\hat{P} + \hat{Q})\|_F^2 &\leq \alpha(gn + 1) \|(\hat{P} - P^*)^T\|_{2,1} \\ &\quad + 2\beta \|\hat{Q} - Q^*\|_{2,1}, \end{aligned} \quad (15)$$

where  $P^*$  and  $Q^*$  are the ground truth that generates the data matrix  $X$  and  $g$  is the maximum value of weight  $w_{i,j}$ .

*Proof.* Since  $\hat{P}$  and  $\hat{Q}$  are global minimizers, from Eq. (13) we have:

$$\begin{aligned} \frac{1}{2} \|X - (\hat{P} + \hat{Q})\|_F^2 + \alpha \|\hat{P}\|_{FU} + \beta \|\hat{Q}\|_{2,1} \\ \leq \frac{1}{2} \|X - (P^* + Q^*)\|_F^2 + \alpha \|P^*\|_{FU} + \beta \|Q^*\|_{2,1}, \end{aligned}$$

By substituting the assumption we made about  $X$  in Eq. (12) we obtain:

$$\begin{aligned} \frac{1}{2} \|(P^* + Q^*) + \delta - (\hat{P} + \hat{Q})\|_F^2 + \alpha \|\hat{P}\|_{FU} + \beta \|\hat{Q}\|_{2,1} \\ \leq \frac{1}{2} \|\delta\|_F^2 + \alpha \|P^*\|_{FU} + \beta \|Q^*\|_{2,1}. \end{aligned} \quad (16)$$

Next we expand the first term in left hand side as:

$$\begin{aligned} \frac{1}{2} \|(P^* + Q^*) + \delta - (\hat{P} + \hat{Q})\|_F^2 \\ = \frac{1}{2} \|(P^* + Q^*) - (\hat{P} + \hat{Q})\|_F^2 + \frac{1}{2} \|\delta\|_F^2 \\ + \sum_{i,j} ((P^* + Q^*) - (\hat{P} + \hat{Q})) \delta_{i,j}. \end{aligned} \quad (17)$$

By substituting Eq. (17) into Eq. (16) and rearranging all terms we obtain:

$$\begin{aligned} \frac{1}{2} \|(P^* + Q^*) - (\hat{P} + \hat{Q})\|_F^2 \\ \leq \alpha \|P^*\|_{FU} + \beta \|Q^*\|_{2,1} - \alpha \|\hat{P}\|_{FU} - \beta \|\hat{Q}\|_{2,1} \\ + \sum_{i,j} (\hat{P}_{i,j} - P_{ij}^*) \delta_{i,j} + \sum_{i,j} (\hat{Q}_{i,j} - Q_{ij}^*) \delta_{i,j} \\ \leq \alpha \|P^*\|_{FU} + \beta \|Q^*\|_{2,1} - \alpha \|\hat{P}\|_{FU} - \beta \|\hat{Q}\|_{2,1} \\ + \sum_j \|P_j^* - \hat{P}_j\|_2 \|\delta_j\|_2 + \sum_i \|Q_i^* - \hat{Q}_i\|_2 \|\delta_i\|_2, \end{aligned} \quad (18)$$

where the last line we used Cauchy-Schwartz inequality.  $P_j^*, \hat{P}_j, \delta_j$  are the  $j^{th}$  column of  $P^*, \hat{P}$  and  $\delta$  respectively and  $Q_i^*, \hat{Q}_i, \delta_i$  are the  $i^{th}$  row of  $Q^*, \hat{Q}$  and  $\delta$  respectively.

Next we compute the upper bounds for  $\|\delta_i\|_2$  where  $i$  is the row index with the following Lemma about chi-squared random variable.

**Lemma 1.** Let  $\chi^2(d)$  be a chi-squared random variable with  $d$  degrees of freedom. Then the following holds [8]:

$$Pr(\chi^2(d) \geq d + t) \leq \exp(-\frac{1}{2}(t - d \log(1 + \frac{t}{d}))), t > 0. \quad (19)$$

Hence with the probability of at least  $1 - \exp(-\frac{1}{2}(t - d \log(1 + \frac{t}{d})))$ , we have  $\|\delta_i\|_2 \leq \sigma \sqrt{(d+t)} = \lambda/(\sqrt{n})$ .

Similarly, we can compute the upper bound for  $\|\delta_j\|_2$  where  $j$  is the column index:

$$\|\delta_j\|_2 \leq \sqrt{\|\delta\|_F^2} \leq \sigma \sqrt{n(d+t)} = \lambda. \quad (20)$$

Substitute the upper bound of  $\|\delta_i\|_2$  and  $\|\delta_j\|_2$  into Eq. (18) we obtain:

$$\begin{aligned} \frac{1}{2} \|(P^* + Q^*) - (\hat{P} + \hat{Q})\|_F^2 \\ \leq \alpha \|P^*\|_{FU} + \beta \|Q^*\|_{2,1} - \alpha \|\hat{P}\|_{FU} - \beta \|\hat{Q}\|_{2,1} \\ + \lambda \sum_j \|P_j^* - \hat{P}_j\|_2 + \frac{\lambda}{\sqrt{n}} \sum_i \|Q_i^* - \hat{Q}_i\|_2. \end{aligned}$$

Now take the regularization parameter as Eq. (14), we obtain:

$$\begin{aligned} \frac{1}{2} \|(P^* + Q^*) - (\hat{P} + \hat{Q})\|_F^2 \\ \leq \alpha \|P^*\|_{FU} + \beta \|Q^*\|_{2,1} - \alpha \|\hat{P}\|_{FU} - \beta \|\hat{Q}\|_{2,1} \\ + \alpha \|(\hat{P} - P^*)^T\|_{2,1} + \beta \|\hat{Q} - Q^*\|_{2,1}. \end{aligned} \quad (21)$$

We divide the right hand side of Eq. (21) into two parts, the terms that only contain  $P$  and the terms that only contain

$Q$ , and bound both term accordingly. First we bound the part only contain  $P$ :

$$\begin{aligned}
& \alpha \|P^*\|_{FU} - \alpha \|\hat{P}\|_{FU} + \alpha \|(P^* - \hat{P})^T\|_{2,1} \\
&= \alpha \sum_{i < j} w_{i,j} (\|P_i^* - P_j^*\|_2 - \|\hat{P}_i - \hat{P}_j\|_2) + \alpha \|(P^* - \hat{P})^T\|_{2,1} \\
&\leq \alpha \sum_{i < j} w_{i,j} \|P_i^* - P_j^* - \hat{P}_i + \hat{P}_j\|_2 + \alpha \|(P^* - \hat{P})^T\|_{2,1} \\
&\leq \alpha \sum_{i < j} w_{i,j} (\|P_i^* - \hat{P}_i\|_2 + \|\hat{P}_j - P_j^*\|_2) + \alpha \|(P^* - \hat{P})^T\|_{2,1} \\
&\leq \alpha \sum_{i < j} g (\|P_i^* - \hat{P}_i\|_2 + \|\hat{P}_j - P_j^*\|_2) + \alpha \|(P^* - \hat{P})^T\|_{2,1} \\
&\leq \alpha g n \|(P^* - \hat{P})^T\|_{2,1} + \alpha \|(P^* - \hat{P})^T\|_{2,1} \\
&\leq \alpha (g n + 1) \|(P^* - \hat{P})^T\|_{2,1}, \tag{22}
\end{aligned}$$

where indices  $i$  and  $j$  are both used to denote the column of  $P$ . Next the following terms that only contain  $Q$  can be bounded:

$$\beta \|Q^*\|_{2,1} - \beta \|\hat{Q}\|_{2,1} + \beta \|\hat{Q} - Q^*\|_{2,1} \leq 2\beta \|\hat{Q} - Q^*\|_{2,1}. \tag{23}$$

Combine Eq. (22), Eq. (23) and Eq. (21), we obtain:

$$\begin{aligned}
\frac{1}{2} \|(P^* + Q^*) - (\hat{P} + \hat{Q})\|_F^2 &\leq \alpha (g n + 1) \|(P^* - \hat{P})^T\|_{2,1} \\
&\quad + 2\beta \|\hat{Q} - Q^*\|_{2,1}. \tag{24}
\end{aligned}$$

This completes the proof of Theorem 2.  $\square$

Before we move on, we introduce two notations:  $\mathcal{J}(\cdot)$  and  $\mathcal{J}_\perp(\cdot)$  denote the index sets for nonzero rows and zero rows and the definitions are:

$$\mathcal{J}(Q) = \{i | Q_i \neq 0\}, \mathcal{J}_\perp(Q) = \{i | Q_i = 0\}. \tag{25}$$

Next we make the following assumptions about the  $X$ ,  $P^*$  and  $Q^*$  and give a theoretical bound between  $P^* + Q^*$  and  $\hat{P} + \hat{Q}$  based on these assumptions:

**Assumption 1.** For a matrix pair  $\Gamma_P \in \mathbb{R}^{d \times n}$  and  $\Gamma_Q \in \mathbb{R}^{d \times n}$ , let  $r$  and  $c$  ( $1 \leq r \leq n$ ,  $1 \leq c \leq d$ ) be the upper bounds of  $|\mathcal{J}((P^*)^T)|$  and  $|\mathcal{J}(Q^*)|$ , respectively. Let  $\gamma_1$  and  $\gamma_2$  be positive scalars. We assume there exists positive scalars  $\kappa_1(r)$  and  $\kappa_2(c)$  such that:

$$\kappa_1(r) = \min_{\Gamma_P, \Gamma_Q \in R(r,c)} \frac{\|\Gamma_P + \Gamma_Q\|_F}{\|(\Gamma_P^T)^{\mathcal{J}(P^T)}\|_{2,1}} > 0, \tag{26}$$

$$\kappa_2(c) = \min_{\Gamma_P, \Gamma_Q \in R(r,c)} \frac{\|\Gamma_P + \Gamma_Q\|_F}{\|(\Gamma_Q)^{\mathcal{J}(Q)}\|_{2,1}} > 0, \tag{27}$$

where  $R(r, c)$  is defined as:

$$\begin{aligned}
R(r, c) = \{ & \Gamma_P \in \mathbb{R}^{d \times n}, \Gamma_Q \in \mathbb{R}^{d \times n} | \Gamma_P \neq 0, \Gamma_Q \neq 0, | \\
& |\mathcal{J}(P^T)| \leq r, |\mathcal{J}(Q)| \leq c, \\
& \|(\Gamma_P^T)^{\mathcal{J}_\perp(P^T)}\|_{2,1} \leq \gamma_1 \|(\Gamma_P^T)^{\mathcal{J}(P^T)}\|_{2,1}, \\
& \|(\Gamma_Q)^{\mathcal{J}_\perp(Q)}\|_{2,1} \leq \gamma_2 \|(\Gamma_Q)^{\mathcal{J}(Q)}\|_{2,1} \},
\end{aligned}$$

where  $\mathcal{J}(\cdot)$  and  $\mathcal{J}_\perp(\cdot)$  are defined in Eq. (25) and  $|\mathcal{J}|$  denotes the number of elements in the set  $\mathcal{J}$ .

**Lemma 2.** Under Assumption 1 the following results hold with probability of at least  $1 - \exp(-\frac{1}{2}(t - d \log(1 + \frac{t}{d})))$  ( $t > 0$ ):

$$\|(P^* - \hat{P})^T\|_{2,1} \leq \frac{(\gamma_1 + 1)\sqrt{r}}{\kappa_1(r)} \left( \frac{2\alpha\sqrt{r}}{\kappa_1(r)} + \frac{2\beta\sqrt{c}}{\kappa_2(c)} \right), \tag{28}$$

$$\|Q^* - \hat{Q}\|_{2,1} \leq \frac{(\gamma_2 + 1)\sqrt{c}}{\kappa_2(c)} \left( \frac{2\alpha\sqrt{r}}{\kappa_1(r)} + \frac{2\beta\sqrt{c}}{\kappa_2(c)} \right). \tag{29}$$

The proof of a similar lemma can be found in [8].

Theorem 1 can be obtained by substituting Eq. (28) and Eq. (29) into Eq. (15).