

# 통계학 특강 라제 1 해답

$$1. (1) L(\lambda; y) = f(y; \lambda) = \frac{\lambda^y e^{-\lambda}}{y!}$$

$$(2) \ell(\lambda; y) = \log L(\lambda; y) = y \log \lambda - \lambda - \log(y!)$$

$$(3) S(\lambda) = \frac{\partial}{\partial \lambda} \ell(\lambda; y) = \frac{y}{\lambda} - 1$$

$$(4) J(\lambda) = - \frac{\partial^2}{\partial \lambda^2} \ell(\lambda; y) = y / \lambda^2$$

$$(5) I(\lambda) = E[J(\lambda)] = \frac{1}{\lambda^2} E(y) = \frac{1}{\lambda}$$

$$(b) S(\lambda) = 0 \Rightarrow \hat{\lambda} = y$$

$$2 \quad (1) L_n(\lambda; \underline{y}) = \prod_{i=1}^n f(y_i; \lambda) = \prod_{i=1}^n \frac{\lambda^{y_i} e^{-\lambda}}{y_i!}$$

$$= \lambda^{\sum y_i} e^{-n\lambda} \prod_{i=1}^n \frac{1}{y_i!}$$

$$(2) \ell_n(\lambda; \underline{y}) = \sum y_i \log \lambda - n\lambda + \log \left[ \prod_{i=1}^n \frac{1}{y_i!} \right]$$

$$(3) S_n(\lambda) = \frac{\sum y_i}{\lambda} - n$$

$$(4) J_n(\lambda) = \frac{\sum y_i}{\lambda^2}$$

$$(5) I_n(\lambda) = \frac{n}{\lambda}$$

$$(b) \quad \hat{\lambda} = \frac{\sum y_i}{n}$$

$$3. \quad y_i \stackrel{\text{ind}}{\sim} N(\mu, \sigma^2)$$

편의를 위하여  $\sigma^2 = \tau$  로 놓은 변수벡터  $\underline{\theta} = (\mu, \tau)^T$  라고 하자

$$\begin{aligned} L_n(\underline{\theta}; \underline{y}) &= \prod_{i=1}^n (2\pi\tau)^{-1/2} \exp\left\{-\frac{(y_i - \mu)^2}{2\tau}\right\} \\ &= [2\pi\tau]^{-n/2} \exp\left\{-\frac{1}{2\tau} \sum_{i=1}^n (y_i - \mu)^2\right\} \end{aligned}$$

$$\ln L_n(\underline{\theta}; \underline{y}) = -\frac{n}{2} \log(2\pi) - \frac{n}{2} \log \tau - \frac{1}{2\tau} \sum_{i=1}^n (y_i - \mu)^2$$

MLE는 다음의 방정식을 만족하는  $\hat{\theta} = (\hat{\mu}, \hat{\tau})^T$  이다.

$$S_n(\underline{\theta}; \underline{y}) = \begin{bmatrix} \frac{\partial}{\partial \mu} \ln L_n(\underline{\theta}; \underline{y}) \\ \frac{\partial}{\partial \tau} \ln L_n(\underline{\theta}; \underline{y}) \end{bmatrix} = 0$$

$$\textcircled{1} \quad \frac{\partial}{\partial \mu} \ln L_n(\underline{\theta}; \underline{y}) = \frac{1}{\tau} \sum (y_i - \mu) = 0$$

$$\textcircled{2} \quad \frac{\partial}{\partial \tau} \ln L_n(\underline{\theta}; \underline{y}) = -\frac{n}{2\tau} + \frac{1}{2\tau^2} \sum (y_i - \mu)^2 = 0$$

따라서 위의 두 방정식을 만족하는 MLE는

$$\hat{\mu} = \frac{\sum y_i}{n} = \bar{y} \quad \hat{\tau} = \hat{\sigma}^2 = \frac{\sum (y_i - \bar{y})^2}{n}$$

$$J_n(\underline{\theta}) = - \begin{bmatrix} \frac{\partial^2}{\partial \mu^2} \ln(\underline{\theta}) & \frac{\partial^2}{\partial \mu \partial \tau} \ln(\underline{\theta}) \\ \frac{\partial^2}{\partial \tau \partial \mu} \ln(\underline{\theta}) & \frac{\partial^2}{\partial \tau^2} \ln(\underline{\theta}) \end{bmatrix} \quad \begin{matrix} \tau^{-2} \\ -2\tau \end{matrix}$$

$$= \begin{bmatrix} \frac{n}{\tau} & \frac{1}{\tau^2} \sum (y_i - \mu) \\ \frac{1}{\tau^2} \sum (y_i - \mu) & -\frac{n}{2\tau^2} + \frac{1}{\tau^3} \sum (y_i - \mu)^2 \end{bmatrix}$$

$$I_n(\underline{\theta}) = E[J_n(\underline{\theta})]$$

$$= \begin{bmatrix} \frac{n}{\tau} & 0 \\ 0 & -\frac{n}{2\tau^2} + \frac{n}{\tau^2} \end{bmatrix}$$

$$= \begin{bmatrix} \frac{n}{\tau} & 0 \\ 0 & \frac{n}{2\tau^2} \end{bmatrix} \Rightarrow$$

$$I_n^{-1}(\underline{\theta}) = \begin{bmatrix} \tau/n & 0 \\ 0 & 2\tau^2/n \end{bmatrix}$$

$\underline{\theta}_0 \stackrel{?}{=} \underline{\theta}$  의 값이 주어지면

$$\hat{\underline{\theta}} = \begin{bmatrix} \hat{\mu} \\ \hat{\tau} \end{bmatrix} \sim N(\underline{\theta}_0, I_n^{-1}(\underline{\theta}_0))$$

$$4. \quad y_i \stackrel{\text{iid}}{\sim} \exp(\lambda) \quad f(y; \lambda) = \frac{1}{\lambda} e^{-y/\lambda}, \quad \lambda > 0$$

$$L_n(\lambda) = \lambda^{-n} e^{-\sum y_i / \lambda}$$

$$\ln L_n(\lambda) = -n \log \lambda - \frac{\sum y_i}{\lambda}$$

$$S_n(\lambda) = -\frac{n}{\lambda} + \frac{\sum y_i}{\lambda^2}$$

$$S_n(\lambda) = 0 \quad \text{을 만족하는 MLE 는}$$

$$\hat{\lambda} = \frac{\sum y_i}{n}$$

$$J_n(\lambda) = -\frac{n}{\lambda^2} + \frac{2 \sum y_i}{\lambda^3}$$

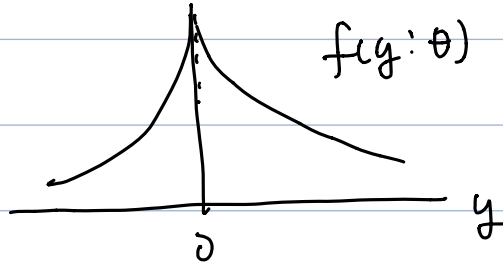
$$I_n(\lambda) = -\frac{n}{\lambda^2} + \frac{2n}{\lambda^2} = \frac{n}{\lambda^2}$$

$$\lambda_0 \text{ 는 } \lambda \text{ 의 참값이라고 하면}$$

$$\hat{\lambda} \sim N\left(\lambda_0, \frac{\lambda_0^2}{n}\right)$$

$$5 \quad y_i \stackrel{iid}{\sim} f(y|\theta) = \frac{1}{2\theta} \exp\left(-\frac{|y|}{\theta}\right)$$

$f(y|\theta)$  의 확률밀도함수는 다음과 같다.



하지만  $f(y|\theta)$  를  $\theta$  의 함수로 보면 미분가능한 함수이다.

$$L_n(\theta; \underline{y}) = (2\theta)^{-n} \exp\left(-\frac{1}{\theta} \sum_{i=1}^n |y_i|\right)$$

$$\ln L_n(\theta; \underline{y}) = -n \log 2 - n \log \theta - \frac{1}{\theta} \sum_{i=1}^n |y_i|$$

$$S_n(\theta; \underline{y}) = -\frac{n}{\theta} + \frac{1}{\theta^2} \sum_{i=1}^n |y_i|$$

$S_n(\theta; \underline{y})$  는 만족하는 MLE 는

$$\hat{\theta} = \frac{\sum |y_i|}{n}$$

$$J_n(\theta; \underline{y}) = -\frac{n}{\theta^2} + \frac{2}{\theta^3} \sum |y_i|$$

$$I_n(\theta; \underline{y}) = -\frac{n}{\theta^2} + \frac{4n}{\theta^2} = \frac{3n}{\theta^2}$$

따라서

$$E[|Y|] = \int_{-\infty}^{\infty} |y| f(y; \theta) dy$$

$$= \int_{-\infty}^{\infty} |y| \cdot \frac{1}{\theta} e^{-|y|/\theta} dy \quad \text{↗ 제곱}$$

$$= 2 \cdot \int_0^{\infty} y \frac{1}{\theta} e^{-y/\theta} dy$$

$$= 2\theta$$

↘가 평균이  
 $\theta$ 인 지수분포  
를 따르는 경우  
 $E(Y) = \theta$

따라서

$$\hat{\theta} \sim N\left(\theta_0, \frac{\theta_0^2}{3n}\right)$$