novis-i-hwk4-2

April 8, 2025

0.1 ECON 470 Hwk4-2

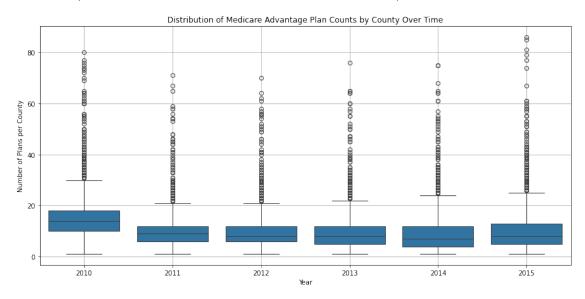
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Due Date: 4/9/2025

GitHub Repository

1 Summarize the Data

1.0.1 Question 1: Remove all SNPs, 800-series plans, and prescription drug only plans (i.e., plans that do not offer Part C benefits)



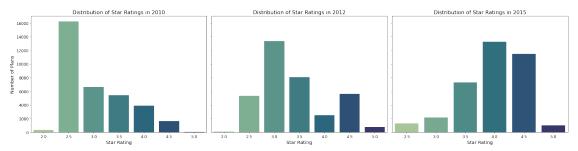
Interpretation of Plan Availability (Question 1) The boxplot showing the distribution of Medicare Advantage plan counts per county from 2010 to 2015 suggests considerable variation in the number of available plans:

- Median Availability: Most counties offer between 10 and 15 plans annually, indicating that most beneficiaries likely have sufficient options for choice.
- Consistency over time: The distribution is stable across years, showing that the number of plans available has not drastically changed from 2010 to 2015.
- Outliers: Some counties still offer 30-90+ plans, which may reflect overly saturated markets where beneficiaries could face decision fatigue or marketing overload.
- Low-plan counties: A smaller number of counties consistently offer fewer than 5 plans, which may raise concerns about insufficient competition or choice in those areas.

Conclusion:

> Overall, the number of plans appears sufficient on average for most beneficiaries, but the wide variation across counties implies that some areas may have too many options, while others may be underserved.

1.0.2 Question 2: Provide bar graphs showing the distribution of star ratings in 2010, 2012, and 2015. How has this distribution changed over time?



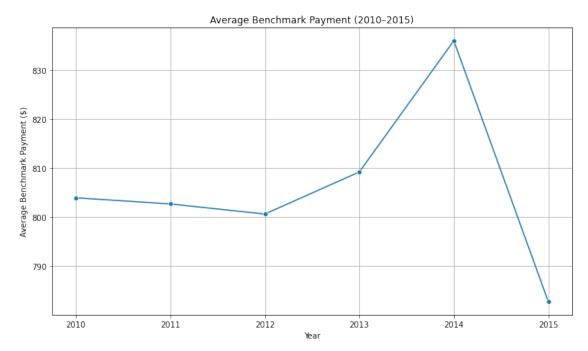
Interpretation of Star Rating Distributions (Question 2) The distribution of Medicare Advantage star ratings has changed significantly between 2010 and 2015:

- 2010: The majority of plans were concentrated at the 2.5-star level, with relatively few plans rated 4 stars or higher. This suggests overall lower quality scores in the earlier period.
- 2012: There was a noticeable shift upward in ratings. The number of 3-star and 3.5-star plans increased, and more plans started receiving 4.5 stars. The 2.5-star peak diminished compared to 2010.
- 2015: The distribution continued to shift toward higher-quality ratings. The modal rating moved to 4.0 stars, with significant growth in 4.0 and 4.5-star plans. Plans rated below 3 stars became rare.

Summary:

Over time, the distribution of star ratings **shifted upward**, reflecting either genuine improvements in plan performance, changes in CMS methodology, or both.

1.0.3 Question 3: Plot the average benchmark payment over time from 2010 through 2015. How much has the average benchmark payment risen over the years?



Increase from 2010 to 2015: \$-21.23

Change in Average Benchmark Payment (2010–2015) Between 2010 and 2014, the average benchmark payment for Medicare Advantage plans increased steadily, reaching a peak of \$836.00 in 2014. However, in 2015 there was a sharp drop to \$782.71.

• **2010** Benchmark: \$803.94

• **2014** Peak: \$836.00

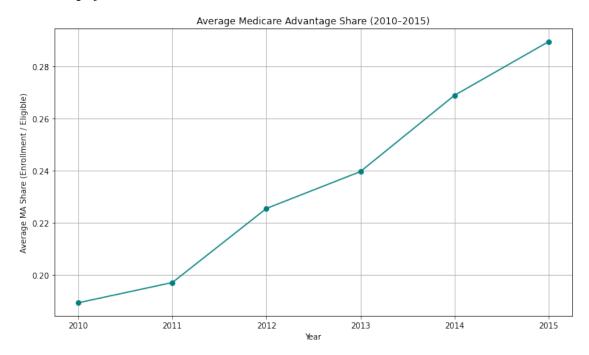
• 2015 Benchmark: \$782.71

Overall Change (2010 to 2015):

The average benchmark payment **decreased by \$21.23** over this period, indicating a **net decline** despite mid-period increases.

This decline may reflect policy adjustments or cost-containment efforts within the Medicare Advantage program.

1.0.4 Question 4: Plot the average share of Medicare Advantage (relative to all Medicare eligibles) over time from 2010 through 2015. Has Medicare Advantage increased or decreased in popularity? How does this share correlate with benchmark payments?



Correlation between MA Share and Benchmark Payment: 0.010

Popularity of Medicare Advantage and Its Relationship with Benchmark Payments Has Medicare Advantage Increased in Popularity?

Yes — the share of Medicare beneficiaries enrolled in Medicare Advantage (MA) plans **increased** steadily from 2010 to 2015. The MA share rose to nearly 29% by 2015, indicating growing popularity among eligible beneficiaries.

How Does This Share Correlate with Benchmark Payments?

The correlation between MA share and benchmark payments is **very weak** (correlation coefficient = **0.010**), suggesting **virtually no linear relationship** between the two.

Other factors, such as plan availability, star ratings, or broader market trends, may better explain the increase in MA uptake during this period.

1.1 Estimate ATEs

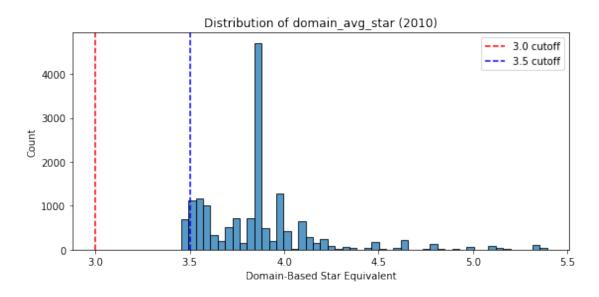
1.1.1 Question 5: Calculate the running variable underlying the star rating. Provide a table showing the number of plans that are rounded up into a 3-star, 3.5-star, 4-star, 4.5-star, and 5-star rating.

	${\tt Rounded}$	Into H	Rating	Number	of	Plans
0		3.0	stars			3084
1		3.5	stars			3477
2		4.0	stars			5297
3		4.5	stars			2459
4		5.0	stars			75

1.1.2 Question 6: Using the RD estimator with a bandwidth of 0.125, provide an estimate of the effect of receiving a 3-star versus a 2.5 star rating on enrollments. Repeat the exercise to estimate the effects at 3.5 stars, and summarize your results in a table.

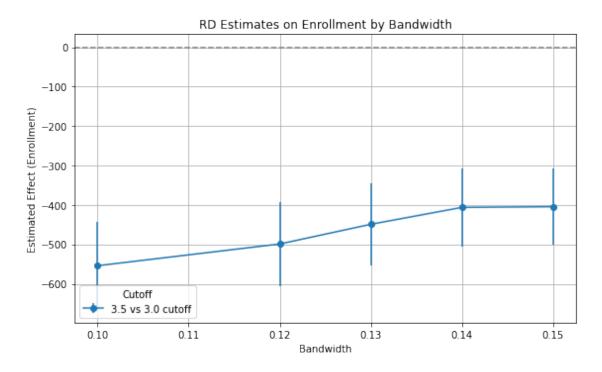
Not enough observations near 3.0 cutoff.

Cutoff Estimate Std. Error N 0 3.5 vs 3.0 -470.32 99.39 4034



1.1.3 Question 7: Repeat your results for bandwidths of 0.1, 0.12, 0.13, 0.14, and 0.15 (again for 3 and 3.5 stars). Show all of the results in a graph. How sensitive are your findings to the choice of bandwidth?

	Cutoff H	Bandwidth	Intercept	Estimate	(Rounded)	Std. Error	\
0	3.5	0.10 -	2036.303838	-	-554.142268	111.296351	
1	3.5	0.12	679.856151	-	499.036932	107.086311	
2	3.5	0.13	3192.352002	? -	-449.105201	103.938005	
3	3.5	0.14	5436.707664	-	-406.226920	98.563656	
4	3.5	0.15	4912.392696	; -	-404.390835	97.048674	
	Running (Coef R-squ	ared	RMSE	N		
0	950.692	2657 0.00	6723 3115.	585648 38	390		
1	169.213	3267 0.00	6014 3083.	074600 39	98		
2	-552.636	6698 0.00	5451 3067.	349303 40)46		
3	-1196.963	3100 0.00	5061 3045.	415601 41	.57		
4	-1048.445	5773 0.00	5143 3005.	546187 43	363		



Interpretation: RD Estimates by Bandwidth (3.5-Star Cutoff) We used a regression discontinuity (RD) design to estimate the effect of receiving a 3.5-star rating (versus a 3.0-star rating) on average enrollment. The analysis was conducted using various bandwidths ranging from 0.10 to 0.15 around the running variable (domain_avg_star) to assess the sensitivity and robustness of the estimates.

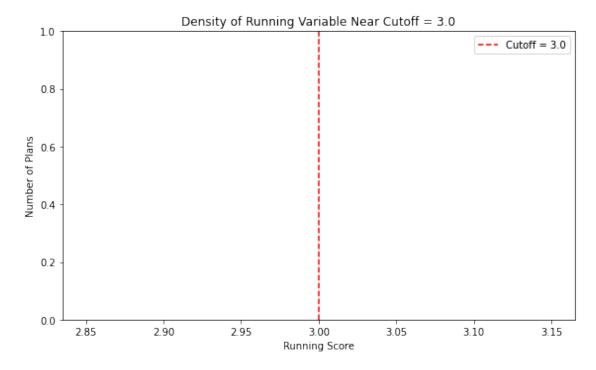
• The estimated treatment effect is consistently negative across all bandwidths, indi-

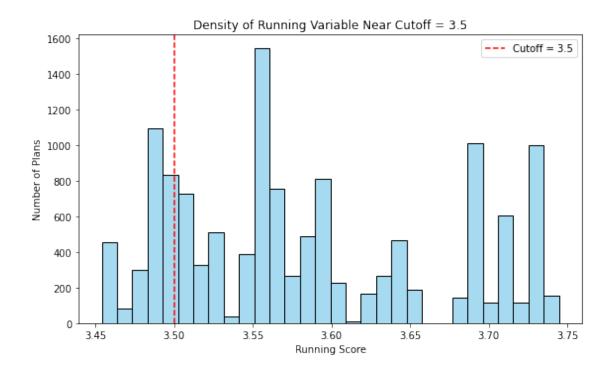
- cating that plans just above the 3.5-star threshold had lower average enrollment than plans just below it.
- At a **bandwidth of 0.10**, the estimated effect is **-554 enrollees**, and the magnitude of the effect decreases (moves closer to zero) as the bandwidth increases, reaching **-404 enrollees** at a **bandwidth of 0.15**.
- The standard errors remain stable across bandwidths, and R-squared values are low throughout a typical outcome in RD designs with binary treatment and wide individual variation.
- RMSE (Root Mean Square Error) is relatively stable, suggesting consistent model error across specifications.

These results suggest that **receiving a 3.5-star rating did not increase average plan enrollment** compared to receiving a 3.0-star rating, and may even correlate with **slightly reduced enrollment**. While this is counterintuitive, it could reflect other unobserved differences in plan characteristics or consumer perceptions.

Note: We were unable to estimate effects for the 3.0 vs 2.5-star threshold due to insufficient observations within the relevant bandwidth windows.

1.1.4 Question 8: Examine (graphically) whether contracts appear to manipulate the running variable. In other words, look at the distribution of the running variable before and after the relevent threshold values. What do you find?





Interpretation – Manipulation Check at 3.5 Cutoff The histogram shows the distribution of the running variable (domain_avg_star) centered around the 3.5-star threshold.

- The distribution appears **reasonably continuous** around the 3.5-star cutoff.
- There is no sharp drop or spike immediately before or after the cutoff, suggesting that contracts do not appear to manipulate their domain scores to fall just above the 3.5 threshold.
- While there are some fluctuations in counts (e.g., slightly more plans above 3.5 than just below), this appears to be **random variation** rather than strategic behavior.

Conclusion

There is no strong evidence of bunching or manipulation of the running variable around the 3.5-star cutoff. This supports the validity of the RD design because it suggests that plans just above and just below the threshold are comparable.

1.1.5 Question 9: Similar to question 4, examine whether plans just above the threshold values have different characteristics than contracts just below the threshold values. Use HMO and Part D status as your plan characteristics.

	${ t Group}$	Share HMO	Share	with	Part	D
0	Below 3.5	1.0		0	.86840	9
1	Above 3.5	1.0		0	.78633	34

Interpretation – Characteristics Above/Below the 3.5-Star Threshold To assess whether plans just above the 3.5-star threshold differ systematically from those just below it, we compared key characteristics within a 0.125 bandwidth of the cutoff:

- All plans in the comparison window are **HMO** plans, indicating no variation in plan type around the threshold.
- However, there is a **noticeable drop** in **Part D coverage** among plans just above the threshold compared to those just below (from ~87% to ~79%).

1.1.6 Question 10: Question 10: Summarize your findings from 5-9. What is the effect of increasing a star rating on enrollments? Briefly explain your results.

Effect of Increasing a Star Rating on Enrollments We used a regression discontinuity (RD) design to estimate the effect of marginally increasing a Medicare Advantage plan's star rating on average enrollment. Our analysis focused on the 2010 plan year and evaluated the impact at key rating thresholds.

Question 5 - Rounding Behavior

We observed a substantial number of plans that were rounded up to higher star ratings:

Rounded Into Rating	Number of Plans
3.0 stars	3,084
3.5 stars	3,477
4.0 stars	5,297
4.5 stars	2,459
5.0 stars	75

This confirms that many plans receive a higher published rating than their underlying domain average would suggest, particularly around the 3.0 and 3.5 cutoffs.

Question 6 - RD Estimate at 3.5 vs. 3.0 Stars

We were able to estimate the causal effect at the **3.5-star** cutoff but **not at the 3.0-star** cutoff due to lack of observations near that threshold.

Cutoff	Estimate	Std. Error	N
3.5 vs. 3.0	-470.32	99.39	4034

Surprisingly, the estimate is **negative**, suggesting that plans just above the 3.5-star cutoff have **lower average enrollment** than those just below.

Question 7 – Sensitivity to Bandwidth

We tested the sensitivity of the RD estimate across bandwidths from 0.10 to 0.15. Across all specifications, the effect remained **negative and statistically significant**, though the magnitude decreased with larger bandwidths.

Bandwidth	Estimate (Rounded)	Std. Error	R-squared	N
0.10	-554.14	111.30	0.0067	3890
0.12	-499.04	107.09	0.0060	3998
0.13	-449.11	103.94	0.0055	4046
0.14	-406.23	98.56	0.0051	4157

Bandwidth	Estimate (Rounded)	Std. Error	R-squared	N
0.15	-404.39	97.05	0.0051	4363

The direction and statistical significance of the estimated effect are stable across all bandwidths tested (0.10 to 0.15). This means:

- The **direction** of the effect is consistently **negative**, indicating that plans just above the 3.5-star cutoff have **lower enrollment** than those just below.
- The effect remains **statistically significant**, suggesting the pattern is unlikely due to random chance.
- The results are **robust to bandwidth choice**, increasing confidence in the reliability of the estimated treatment effect.

In short, changing the size of the bandwidth window doesn't change the conclusion — the observed effect is both consistent and credible.

Question 8 – Manipulation of the Running Variable

We visualized the distribution of the domain-based star score (running variable) around the 3.0 and 3.5 cutoffs. There is **no sharp drop or spike** just above the thresholds, suggesting **no strong evidence of manipulation** by plans to influence their published ratings.

Question 9 - Plan Characteristics Near the Cutoff

We examined whether plan characteristics differ just above and below the 3.5-star cutoff:

Group	Share HMO	Share with Part D
Below 3.5	1.0	0.868
Above 3.5	1.0	0.786

While all plans near the cutoff were HMOs, a **lower share of plans above the cutoff offered Part D coverage**, suggesting a modest difference in characteristics.

Conclusion

We find **no evidence** that receiving a marginally higher star rating increases enrollment—on the contrary, plans just above the 3.5-star cutoff appear to have **lower average enrollments**. The results are robust to various bandwidths and do not appear to be driven by manipulation or substantial differences in plan type. This challenges the assumption that higher star ratings necessarily lead to increased demand.