

Panel Data

Difference in difference
Interrupted time series
Regression discontinuity
Synthetic control

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Soc 212B
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Learning goals for today

At the end of class, you will be able to:

1. Recognize the promises and pitfalls of four methods to study the effects of treatments that turn on once
 - 1.1 Difference in difference (DID)
 - 1.2 Interrupted time series (ITS)
 - 1.3 Regression discontinuity (RD)
 - 1.4 Synthetic control (SC)

Card, D., & Krueger, A. B. (1994).

Minimum Wages and Employment: A Case Study of the Fast-Food Industry in New Jersey and Pennsylvania.

The American Economic Review, 84(4), 772-793.

Economic theory

When the minimum wage rises, how might employment change?

Economic theory

When the minimum wage rises, how might employment change?

- ▶ employees cost more

Economic theory

When the minimum wage rises, how might employment change?

- ▶ employees cost more
- ▶ employers might get by with fewer employees

The setting

The setting

- ▶ Federal minimum wage
 - ▶ \$3.80 on April 1, 1990
 - ▶ \$4.25 on April 1, 1991

The setting

- ▶ Federal minimum wage
 - ▶ \$3.80 on April 1, 1990
 - ▶ \$4.25 on April 1, 1991
- ▶ New Jersey minimum wage
 - ▶ \$5.05 on April 1, 1992

NJ introduces a high minimum wage.

How would you study the effect on employment?

Source: Wikimedia



Panel Data

Difference in Difference

Interrupted Time Series

Regression Discontinuity

Synthetic Control



Photo by James Loesch - <https://www.flickr.com/photos/jal33/49113053632/>
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New Jersey
Minimum Wage
Rose

Photo by James Loesch - <https://www.flickr.com/photos/jal33/49113053632/>
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Roy Rogers





171 stores



80 stores



99 stores



60 stores



171 stores



80 stores



99 stores



60 stores

Phone interview: Feb-Mar 1992 before minimum wage rose
 Nov-Dec 1992 after minimum wage rose



171 stores



80 stores



99 stores



60 stores

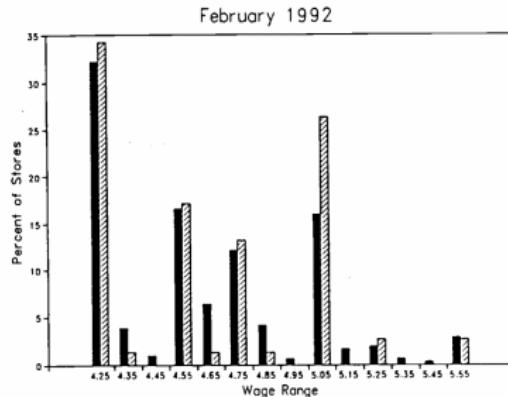
Phone interview: Feb-Mar 1992 before minimum wage rose

Nov-Dec 1992 after minimum wage rose

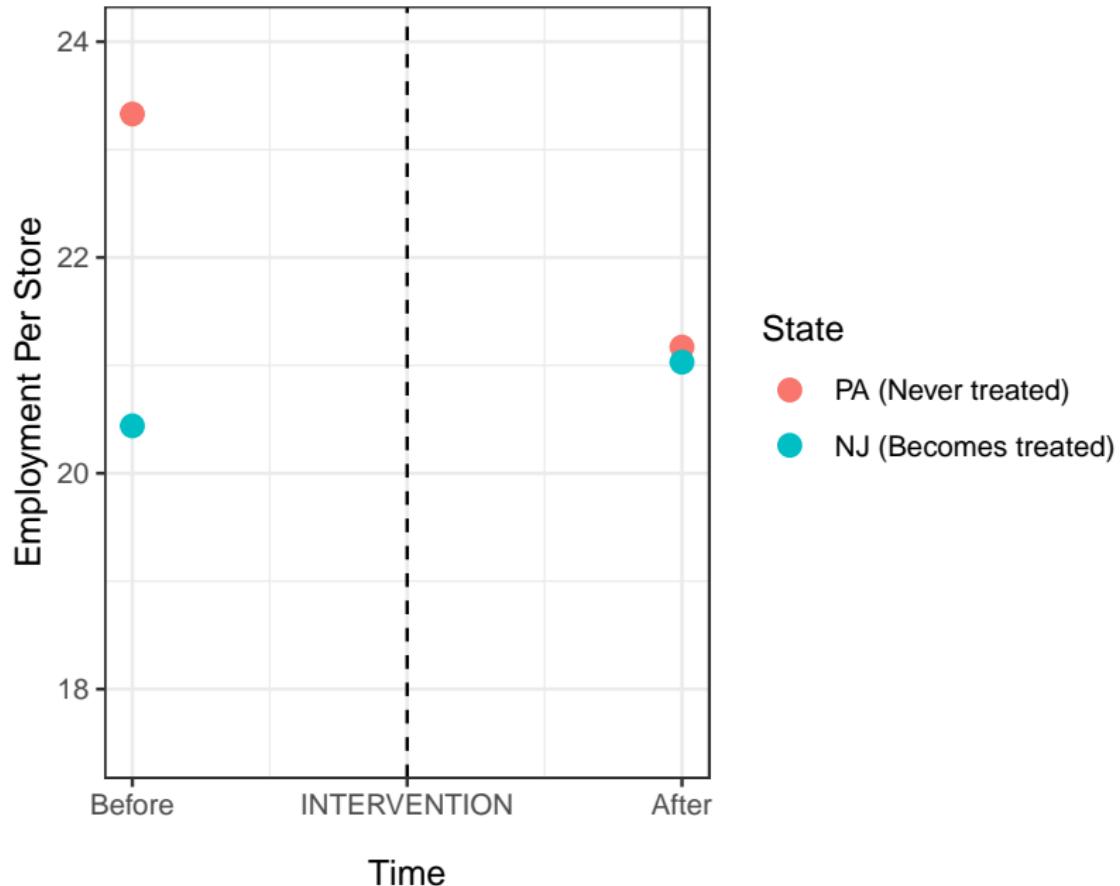
Recorded: How many full-time equivalent employees?

Did starting wages rise in NJ?

Figure 1
Distribution of Starting Wage Rates



How did employment change?



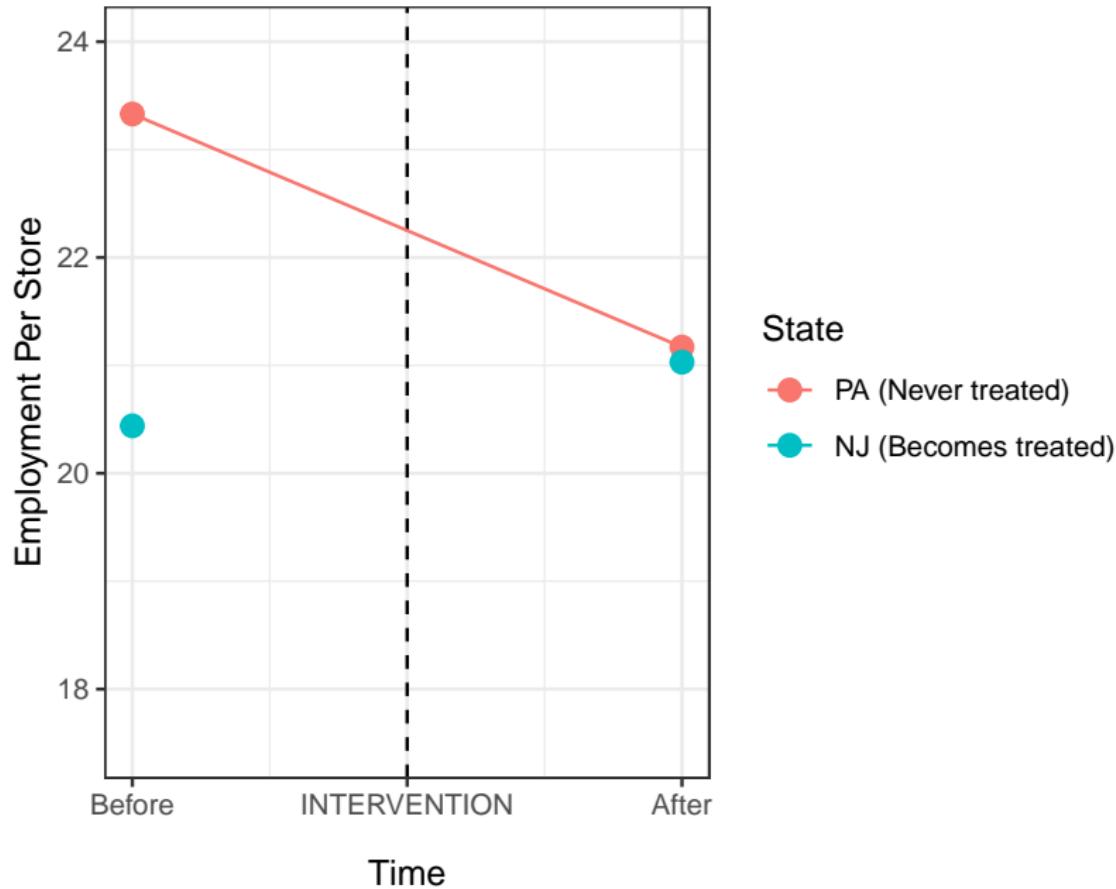
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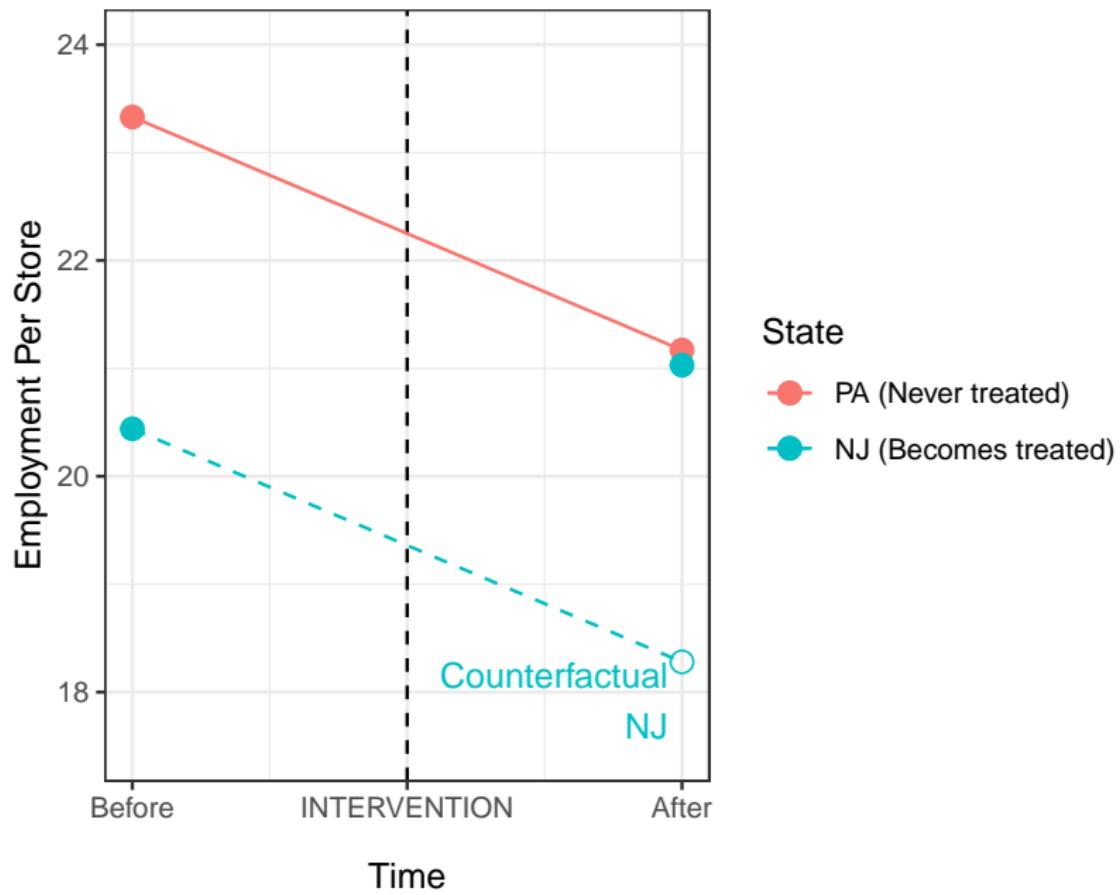
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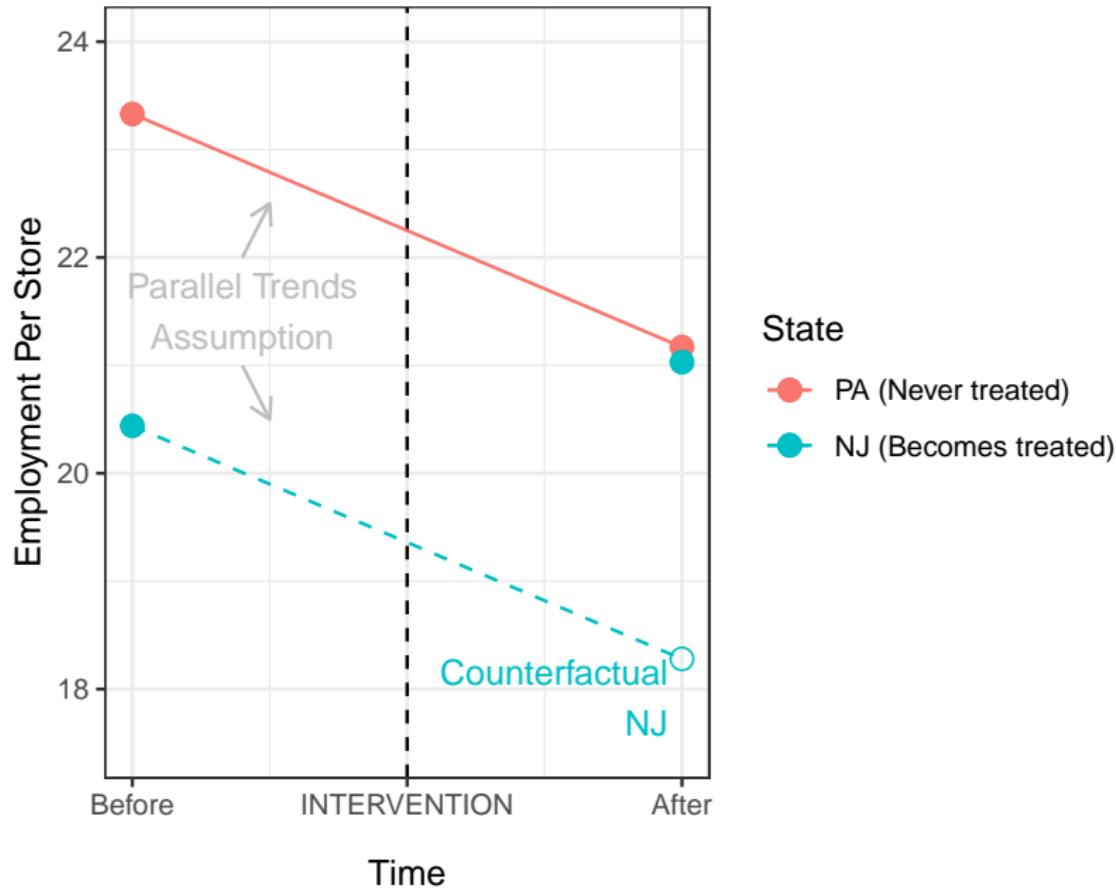
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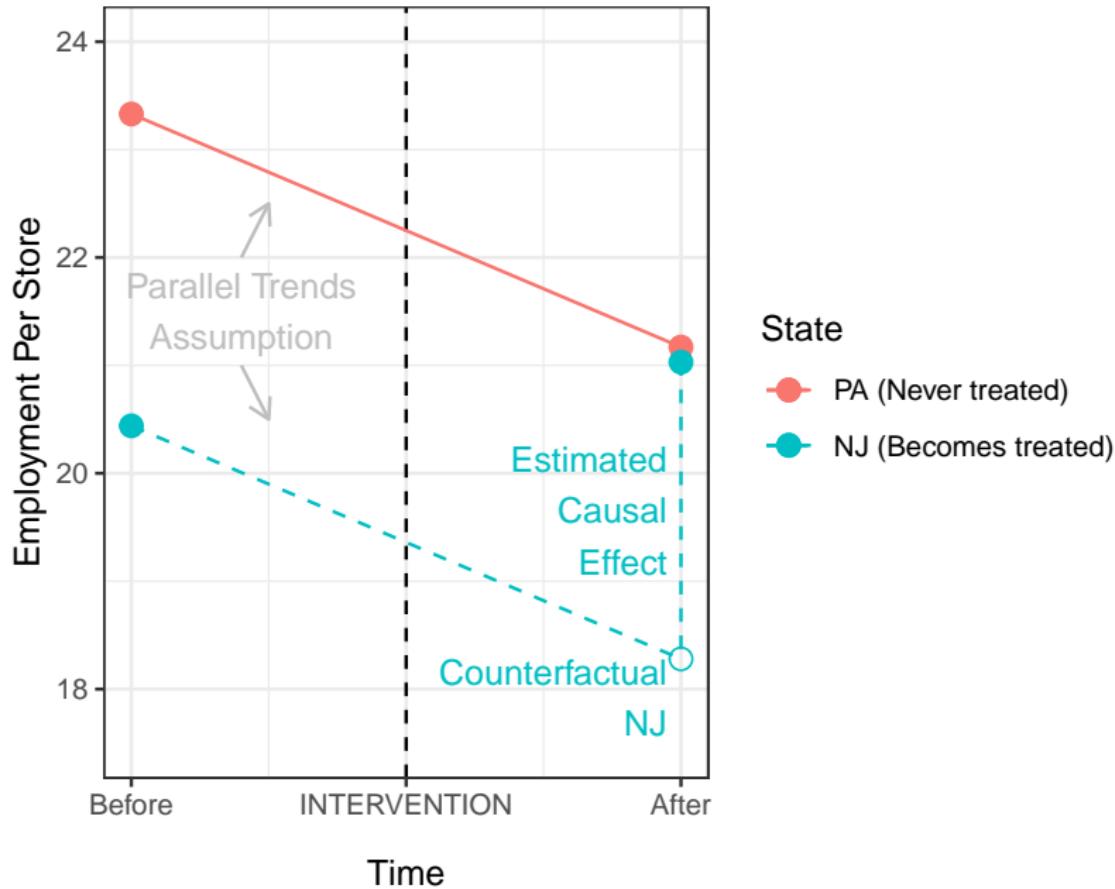
Panel Data

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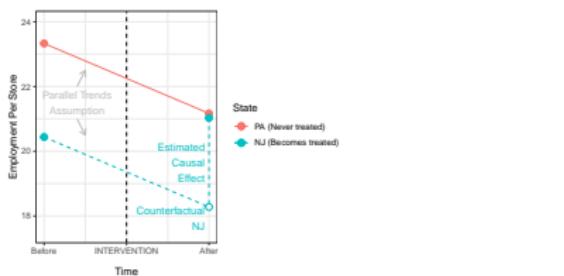
Variable	Stores by state		
	PA (i)	NJ (ii)	Difference, NJ – PA (iii)
1. FTE employment before, all available observations	23.33 (1.35)	20.44 (0.51)	-2.89 (1.44)
2. FTE employment after, all available observations	21.17 (0.94)	21.03 (0.52)	-0.14 (1.07)
3. Change in mean FTE employment	-2.16 (1.25)	0.59 (0.54)	2.76 (1.36)

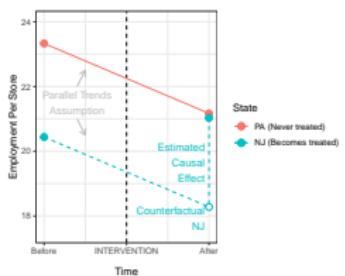
“Contrary to the central prediction of the textbook model of the minimum wage,...we find no evidence that the rise in New Jersey’s minimum wage reduced employment at fast-food restaurants in the state.”

Card & Krueger 1994, p. 792

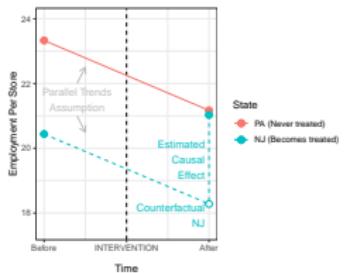
- ▶ simple study
- ▶ well-executed
- ▶ upended conventional wisdom

Key assumption: Parallel trends

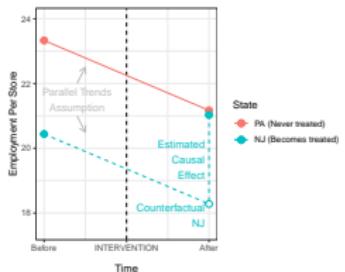




Parallel trends assumption:

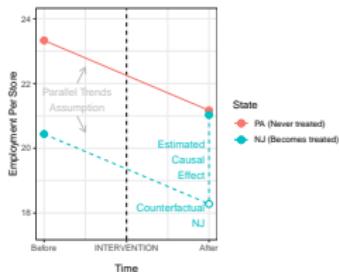


Parallel trends assumption: If no law had taken effect, then



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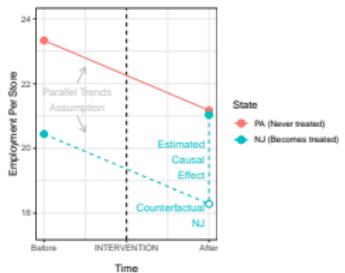
$$\underbrace{Y_{\text{NJ}, \text{After}}^0 - Y_{\text{NJ}, \text{Before}}^0}_{\text{the trend in NJ}} \underset{\equiv}{\text{would equal}} \underbrace{Y_{\text{PA}, \text{After}}^0 - Y_{\text{PA}, \text{Before}}^0}_{\text{the trend in PA}}$$



Parallel trends assumption: If no law had taken effect, then

$$\underbrace{Y_{\text{NJ}, \text{After}}^0 - Y_{\text{NJ}, \text{Before}}^0}_{\text{the trend in NJ}} \underset{\text{would equal}}{\equiv} \underbrace{Y_{\text{PA}, \text{After}}^0 - Y_{\text{PA}, \text{Before}}^0}_{\text{the trend in PA}}$$

Rearranging yields a formula for the counterfactual outcome

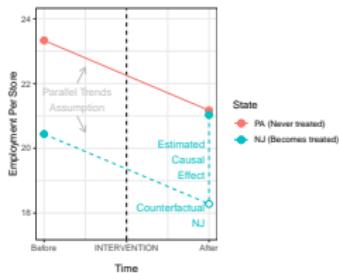


Parallel trends assumption: If no law had taken effect, then

$$\underbrace{Y_{\text{NJ}, \text{After}}^0 - Y_{\text{NJ}, \text{Before}}^0}_{\text{the trend in NJ}} \underset{\text{would equal}}{\equiv} \underbrace{Y_{\text{PA}, \text{After}}^0 - Y_{\text{PA}, \text{Before}}^0}_{\text{the trend in PA}}$$

Rearranging yields a formula for the counterfactual outcome

$$\underbrace{Y_{\text{NJ}, \text{After}}^0}_{\text{Counterfactual}} \underset{\text{By Assumption}}{\equiv} \underbrace{Y_{\text{PA}, \text{After}}^0}_{\text{Actual}}$$

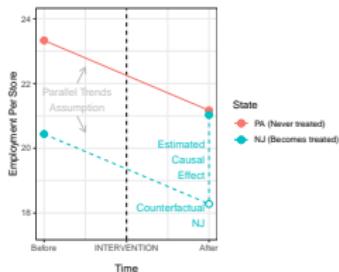


Parallel trends assumption: If no law had taken effect, then

$$\underbrace{Y_{\text{NJ}, \text{After}}^0 - Y_{\text{NJ}, \text{Before}}^0}_{\text{the trend in NJ}} \underset{\text{would equal}}{\equiv} \underbrace{Y_{\text{PA}, \text{After}}^0 - Y_{\text{PA}, \text{Before}}^0}_{\text{the trend in PA}}$$

Rearranging yields a formula for the counterfactual outcome

$$\underbrace{Y_{\text{NJ}, \text{After}}^0}_{\text{Counterfactual}} \underset{\text{By Assumption}}{\equiv} \underbrace{Y_{\text{NJ}, \text{Before}}^0 + Y_{\text{PA}, \text{After}}^0 - Y_{\text{PA}, \text{Before}}^0}_{\text{Factual}}$$



Parallel trends assumption: If no law had taken effect, then

$$\underbrace{Y_{\text{NJ}, \text{After}}^0 - Y_{\text{NJ}, \text{Before}}^0}_{\text{the trend in NJ}} \underset{\text{would equal}}{\equiv} \underbrace{Y_{\text{PA}, \text{After}}^0 - Y_{\text{PA}, \text{Before}}^0}_{\text{the trend in PA}}$$

Rearranging yields a formula for the counterfactual outcome

$$\underbrace{Y_{\text{NJ}, \text{After}}^0}_{\text{Counterfactual}} \stackrel{\text{By Assumption}}{=} \underbrace{Y_{\text{NJ}, \text{Before}}^0 + Y_{\text{PA}, \text{After}}^0 - Y_{\text{PA}, \text{Before}}^0}_{\text{Factual}}$$

$$\text{Effect in NJ} = \underbrace{Y_{\text{NJ}, \text{After}}^1}_{\text{Observed}} - \underbrace{Y_{\text{NJ}, \text{After}}^0}_{\text{Estimated by Above}}$$

Can we test the parallel trends assumption?

Assumption: The trend in NJ would equal the trend in PA

$$\overbrace{Y_{\text{NJ},\text{After}}^0 - Y_{\text{NJ},\text{Before}}^0} \underset{=} {\equiv} \overbrace{Y_{\text{PA},\text{After}}^0 - Y_{\text{PA},\text{Before}}^0}$$

Can we test the parallel trends assumption?

No.

Assumption: The trend in NJ would equal the trend in PA

$$\underbrace{Y_{\text{NJ}, \text{After}}^0 - Y_{\text{NJ}, \text{Before}}^0}_{\text{Not Observable}} \underset{=}{\wedge} \underbrace{Y_{\text{PA}, \text{After}}^0 - Y_{\text{PA}, \text{Before}}^0}$$

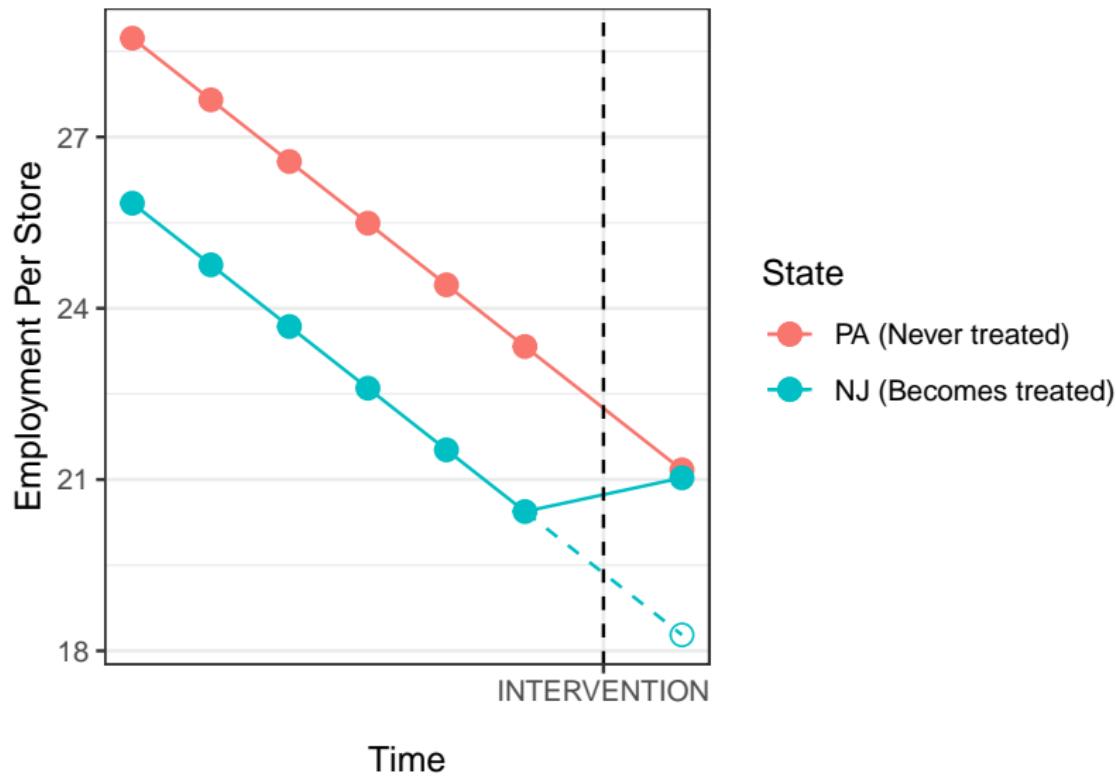
Can we test the parallel trends assumption?

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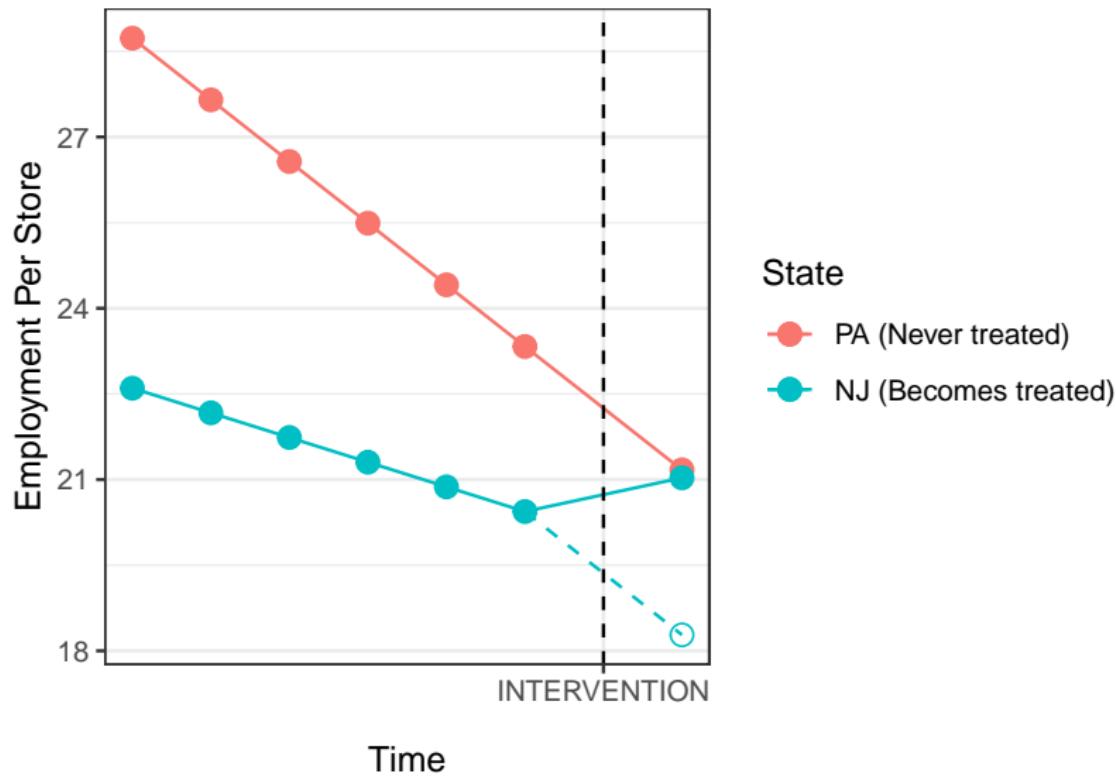
$$\underbrace{Y_{\text{NJ}, \text{After}}^0 - Y_{\text{NJ}, \text{Before}}^0}_{\text{Not Observable}} \underset{\approx}{\text{would equal}} \underbrace{Y_{\text{PA}, \text{After}}^0 - Y_{\text{PA}, \text{Before}}^0}_{\text{the trend in PA}}$$

You can make it credible by looking at many pre-treatment periods

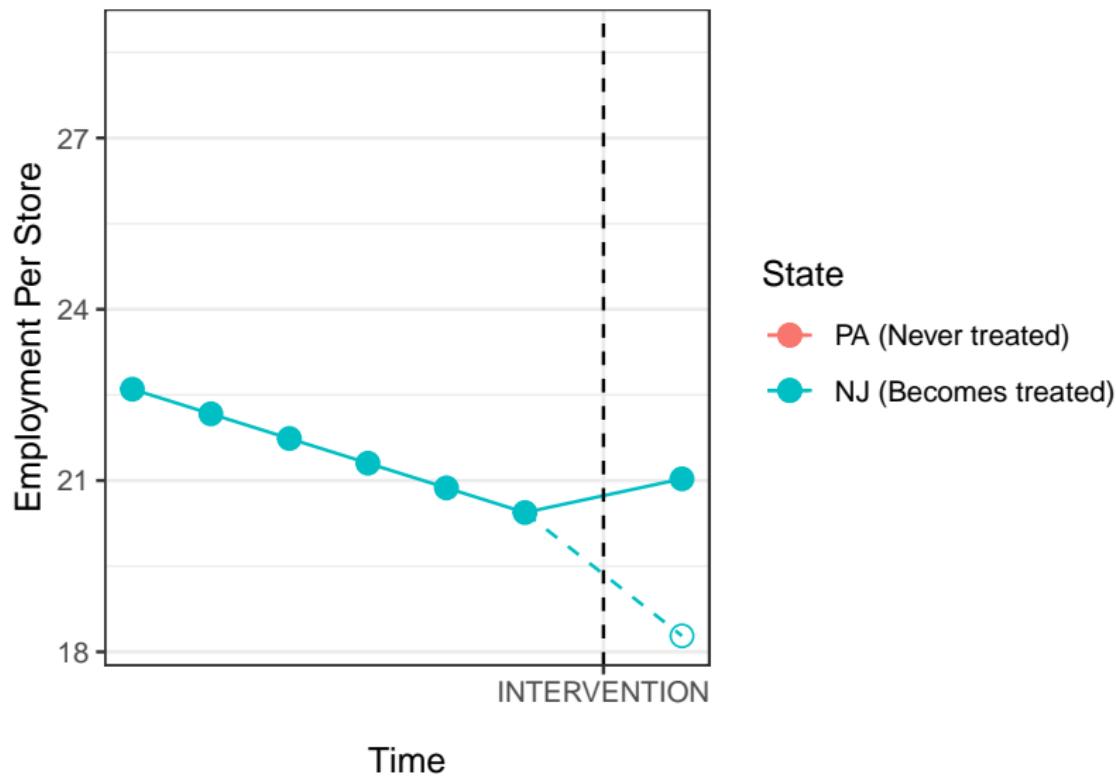
DID would be very credible



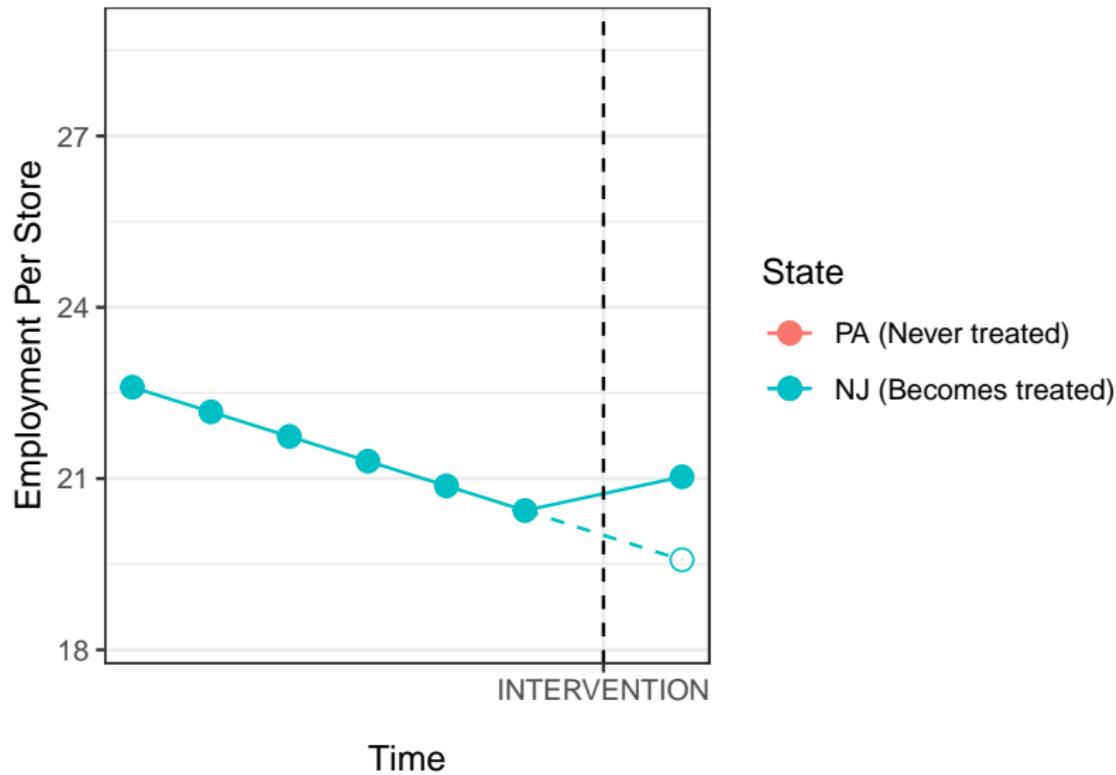
DID would be very doubtful



DID would be very doubtful



DID would be very doubtful

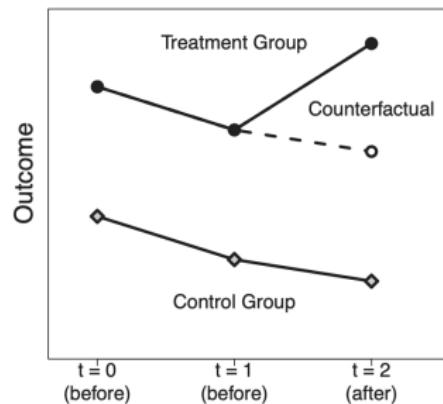


Egami, N., & Yamauchi, S. (2023).

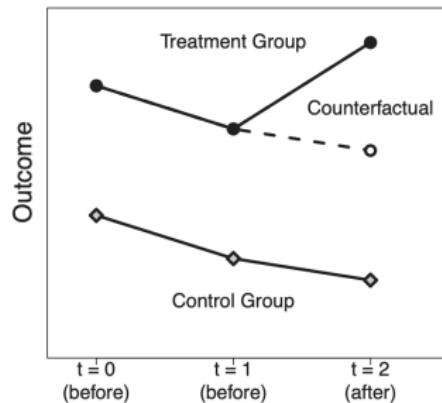
Using multiple pretreatment periods to improve
difference-in-differences and staggered adoption designs.

Political Analysis, 31(2), 195-212.

Difference in difference



Difference in difference

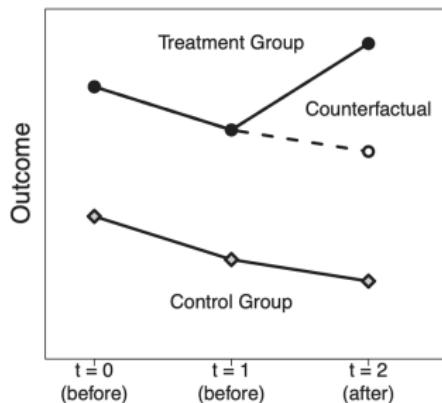


Notation

$Y_{(unit)(time)}^{\text{treatment value}}$

Example: Y_{i1}^0
is unit i at time 1
under treatment 0

Difference in difference



Parallel Trends Assumption
(untestable)

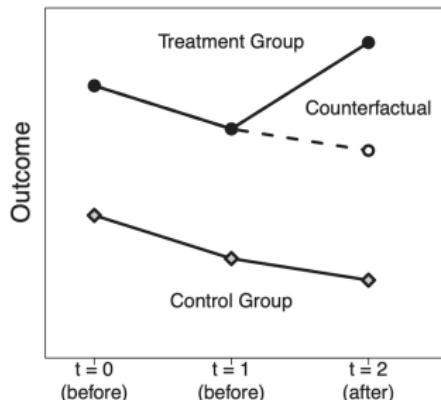
$$E(Y_{\text{Treated},2}^0 - Y_{\text{Treated},1}^0) = E(Y_{\text{Control},2}^0 - Y_{\text{Control},1}^0)$$

Notation

$Y_{(\text{unit})(\text{time})}^{\text{treatment value}}$

Example: Y_{i1}^0
is unit i at time 1
under treatment 0

Difference in difference



Notation

$Y_{(unit)(time)}$
treatment value

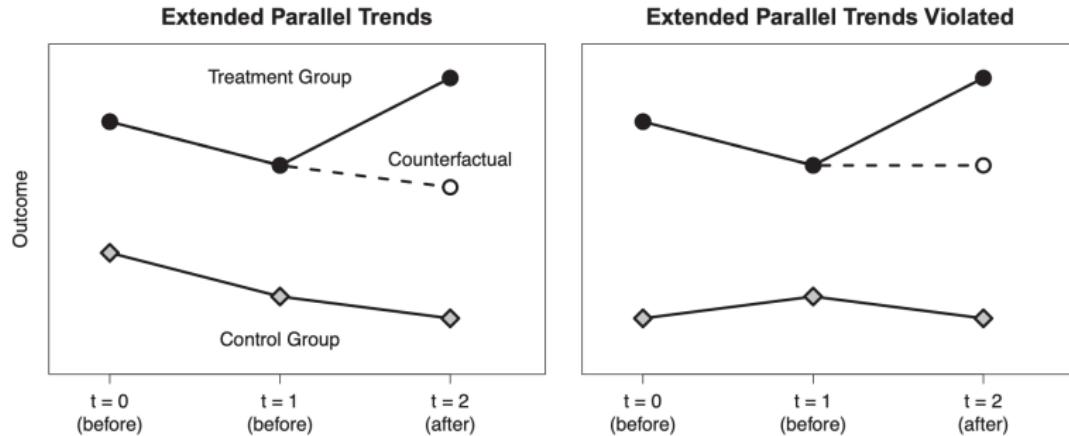
Example: Y_{i1}^0
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Parallel Trends Assumption (untestable)

$$E(Y_{\text{Treated},2}^0 - Y_{\text{Treated},1}^0) = E(Y_{\text{Control},2}^0 - Y_{\text{Control},1}^0)$$

Extended Parallel Trends (testable)

$$E(Y_{\text{Treated},1}^0 - Y_{\text{Treated},0}^0) = E(Y_{\text{Control},1}^0 - Y_{\text{Control},0}^0)$$



Malesky, E. J., Nguyen, C. V., & Tran, A. (2014).
*The impact of recentralization on public services:
A difference-in-differences analysis of the abolition of elected
councils in Vietnam.*
American Political Science Review, 108(1), 144-168.

Does government work better when it is centralized or decentralized?

Vietnam setting: A study of recentralization



Vietnam setting: A study of recentralization

Input from social scientists

Vietnam setting: A study of recentralization

Input from social scientists

1. Enough treated units to study

Vietnam setting: A study of recentralization

Input from social scientists

1. Enough treated units to study
2. Sampling stratified by region

Vietnam setting: A study of recentralization

Input from social scientists

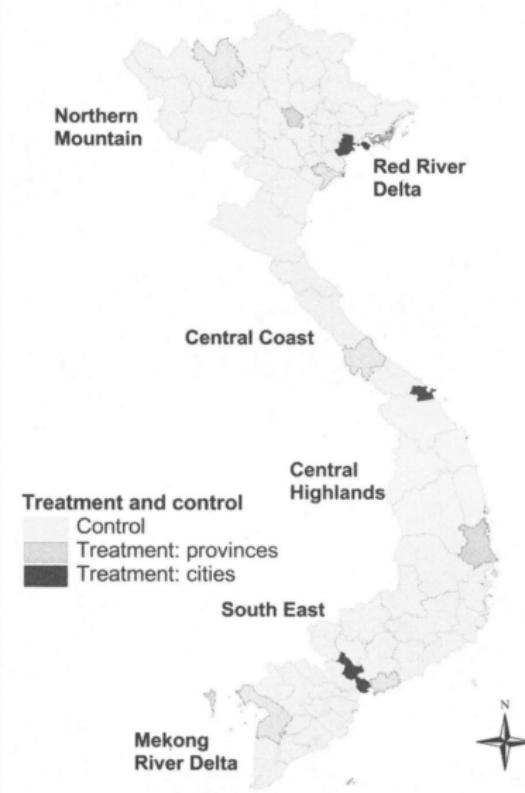
1. Enough treated units to study
2. Sampling stratified by region
3. Sampling stratified by
 - ▶ city versus rural
 - ▶ lowland versus highland
 - ▶ midland versus inter-nationally bordered land

Vietnam setting: A study of recentralization

Input from social scientists

1. Enough treated units to study
2. Sampling stratified by region
3. Sampling stratified by
 - ▶ city versus rural
 - ▶ lowland versus highland
 - ▶ midland versus inter-nationally bordered land
4. Sampling stratified by socioeconomic and public administration performance

FIGURE 2. Map of Treatment Provinces and National-Level Cities



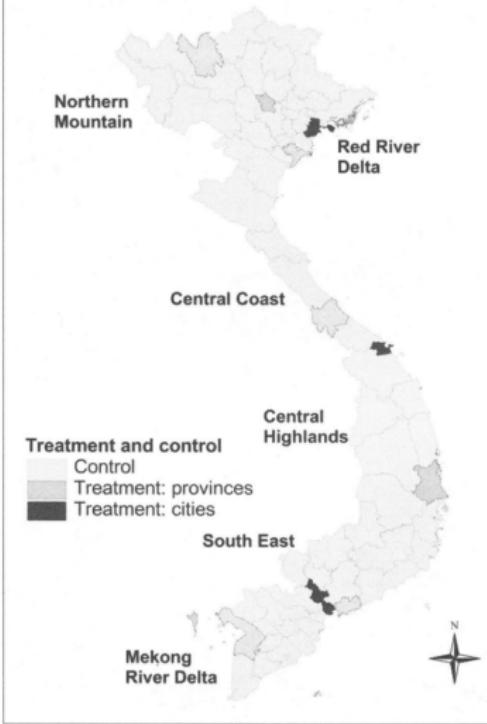
Vietnam Household Living Standards Survey
Reports by each local commune by commune leaders

- ▶ 2006 and 2008: Before DPC abolition
- ▶ 2010: After DPC abolition

One outcome we will examine:
Is there the following project in the commune?

- ▶ Investment on culture and education

FIGURE 2. Map of Treatment Provinces and National-Level Cities



Outcome 1

Education and cultural programs

Is there the following project in the commune?

Investment on culture
and education

FIGURE 2. Map of Treatment Provinces and National-Level Cities

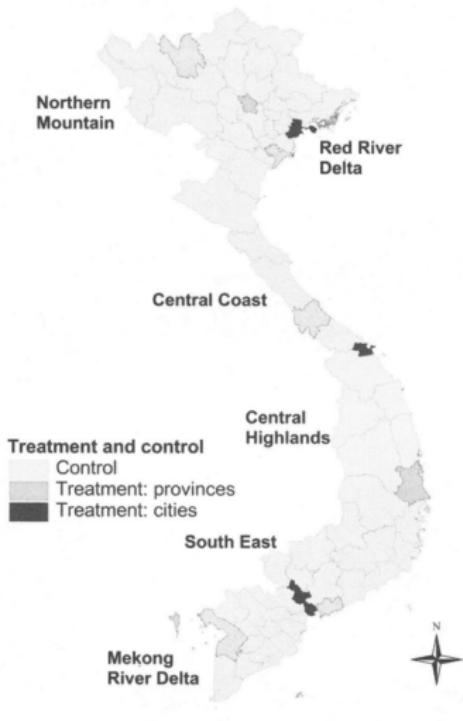
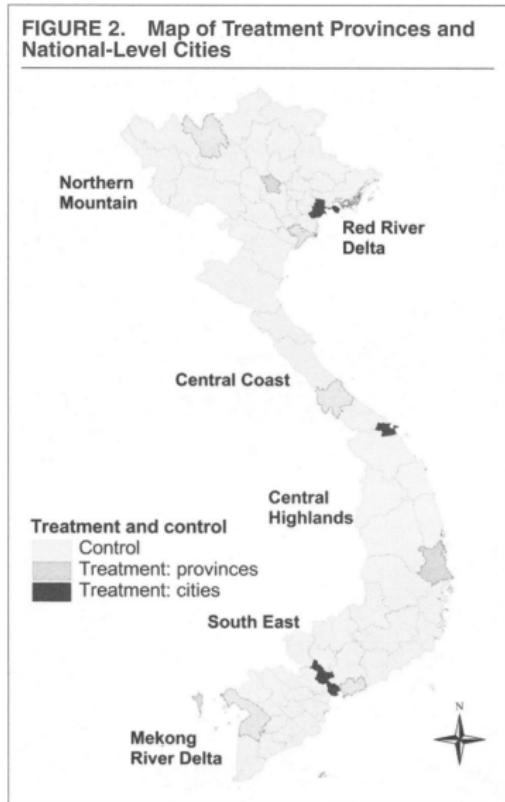


FIGURE 2. Map of Treatment Provinces and National-Level Cities



Outcome 2

Tap water

Is there the following project in the commune?

Coded 1

- Indoor private piped water
- Outdoor private piped water
- Public piped water

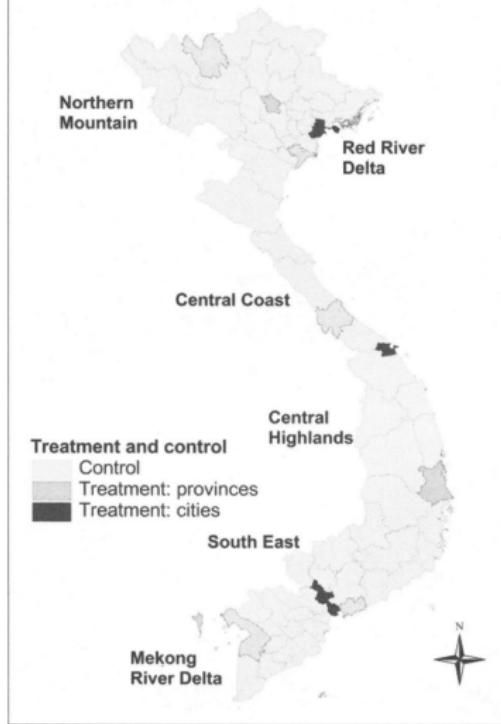
Coded 0

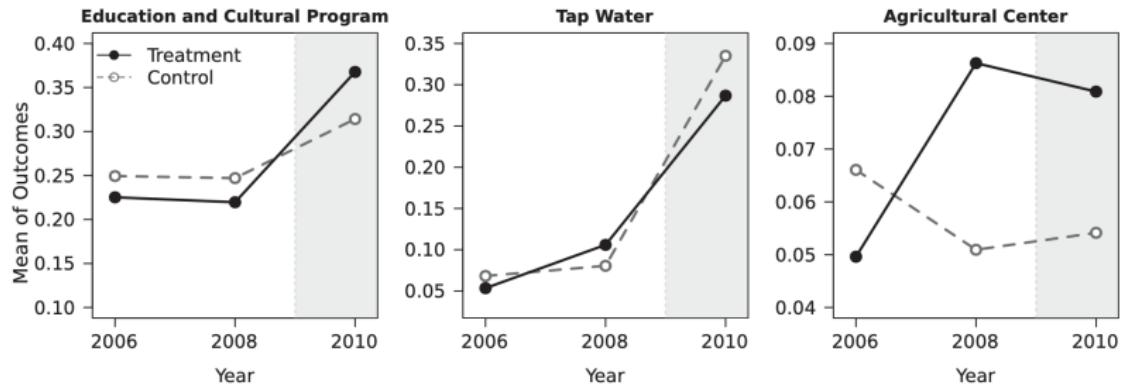
- Well water
- Well with protection walls
- Well without protection walls
- Stream water with protection
- Stream water without protection
- Rainwater
- Bottled water
- Water brought by pedicab
- Tank water
- river lake pond

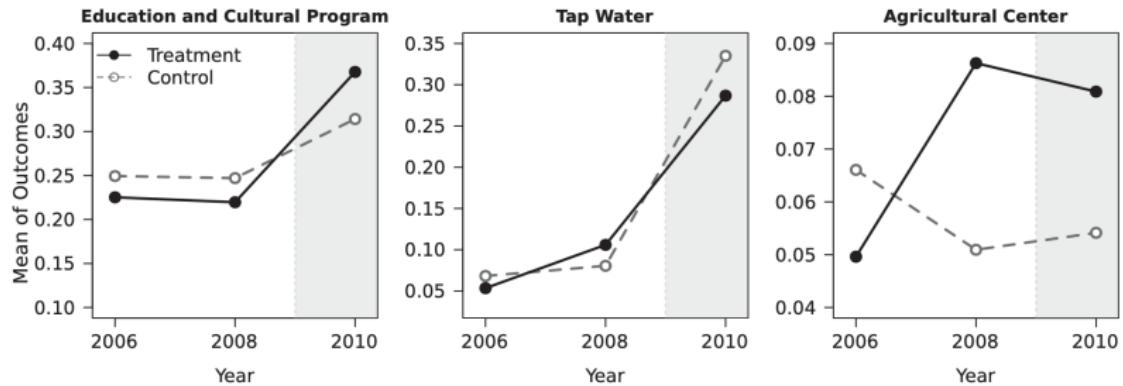
Outcome 3 Agricultural center

Is there any agriculture extension center
in this commune?

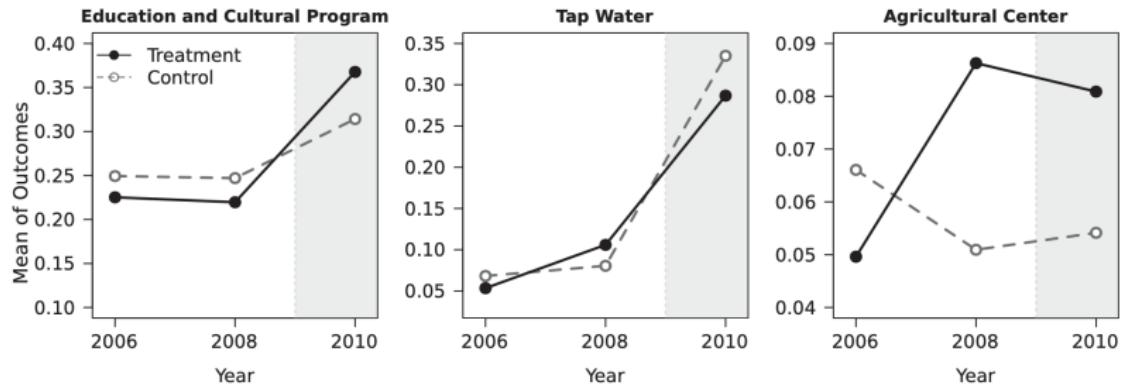
FIGURE 2. Map of Treatment Provinces and National-Level Cities







In each case, do you believe parallel trends?



In each case, do you believe parallel trends?

Table 2. Assessing underlying assumptions using the pretreatment outcomes.

	Estimate	Std. error	p-value	95% Std. equivalence CI
Education and cultural program	-0.007	0.096	0.940	[-0.166, 0.166]
Tap water	0.166	0.083	0.045	[-0.302, 0.302]
Agricultural center	0.198	0.082	0.015	[-0.332, 0.332]

Benefit 1: Assessing assumptions

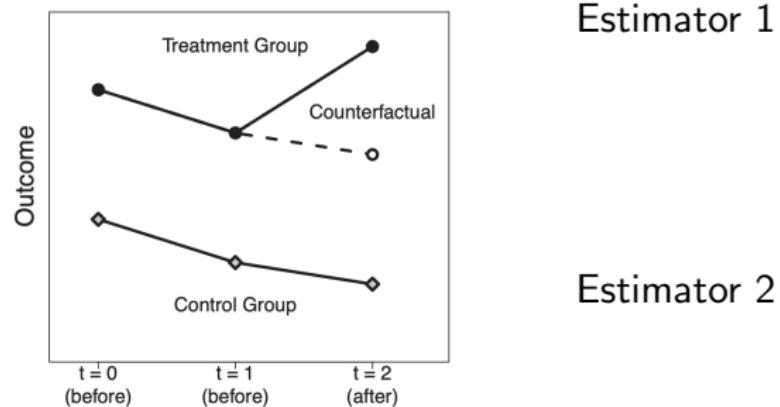
Pre-treatment periods enable us to
assess underlying assumptions

Parallel trends is untestable, but being parallel
in the pre-treatment period builds confidence

Benefit 2: Improving efficiency

Pre-treatment periods also enable us to
improve estimation accuracy
when parallel trends holds

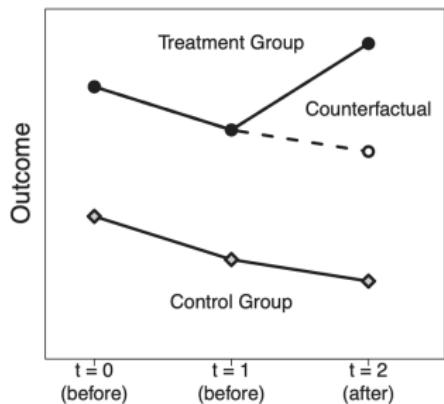
Benefit 2: Improving efficiency



Notation

$\underline{Y}^{\text{treatment value}}$
(unit)(time)

Benefit 2: Improving efficiency



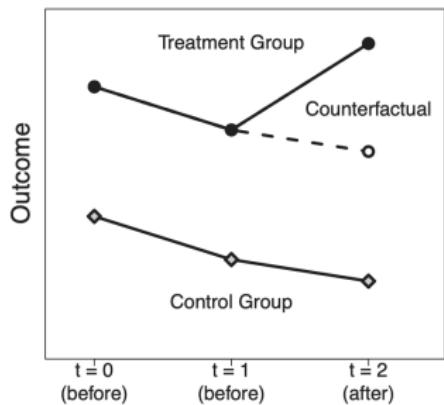
Estimator 1

$$\underbrace{(\bar{Y}_{T2}^1 - \bar{Y}_{T1}^0)}_{\text{Treatment Group Time 2 - Time 1}} - \underbrace{(\bar{Y}_{C2}^0 - \bar{Y}_{C1}^0)}_{\text{Control Group Time 2 - Time 1}}$$

Estimator 2

Notation
 $\bar{Y}_{(unit)(time)}^{\text{treatment value}}$

Benefit 2: Improving efficiency



Estimator 1

$$\underbrace{(\bar{Y}_{T2}^1 - \bar{Y}_{T1}^0)}_{\substack{\text{Treatment Group} \\ \text{Time 2 - Time 1}}} - \underbrace{(\bar{Y}_{C2}^0 - \bar{Y}_{C1}^0)}_{\substack{\text{Control Group} \\ \text{Time 2 - Time 1}}}$$

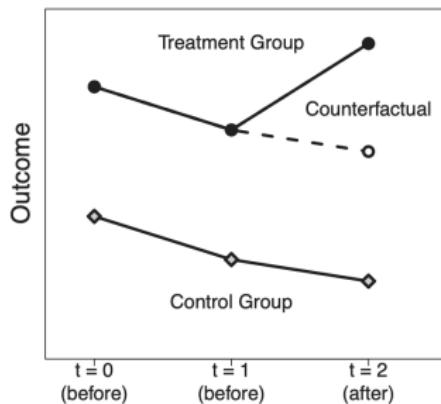
Estimator 2

$$\underbrace{(\bar{Y}_{T2}^1 - \bar{Y}_{T0}^0)}_{\substack{\text{Treatment Group} \\ \text{Time 2 - Time 0}}} - \underbrace{(\bar{Y}_{C2}^0 - \bar{Y}_{C0}^0)}_{\substack{\text{Control Group} \\ \text{Time 2 - Time 0}}}$$

Notation

$\gamma^{\text{treatment value}}$
(unit)(time)

Benefit 2: Improving efficiency



Estimator 1

$$\underbrace{(\bar{Y}_{T2}^1 - \bar{Y}_{T1}^0)}_{\text{Treatment Group Time 2 - Time 1}} - \underbrace{(\bar{Y}_{C2}^0 - \bar{Y}_{C1}^0)}_{\text{Control Group Time 2 - Time 1}}$$

Estimator 2

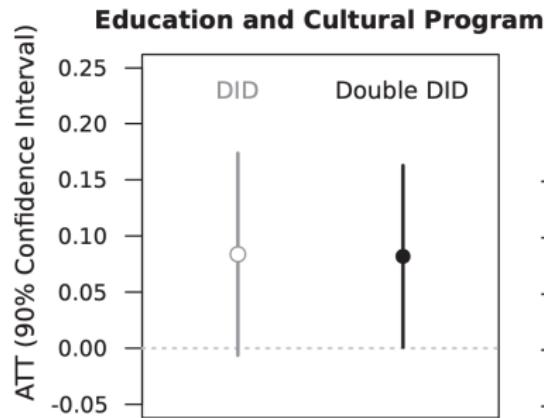
$$\underbrace{(\bar{Y}_{T2}^1 - \bar{Y}_{T0}^0)}_{\text{Treatment Group Time 2 - Time 0}} - \underbrace{(\bar{Y}_{C2}^0 - \bar{Y}_{C0}^0)}_{\text{Control Group Time 2 - Time 0}}$$

Notation

$\bar{Y}_{(unit)(time)}$
treatment value

Pooled estimator:
Average the two!

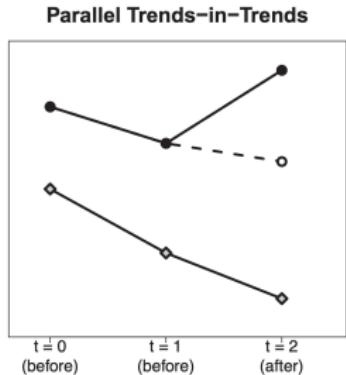
Benefit 2: Improving efficiency



Benefit 3: A more flexible assumption

Pre-treatment periods make it possible to
allow for a more flexible parallel trends assumption

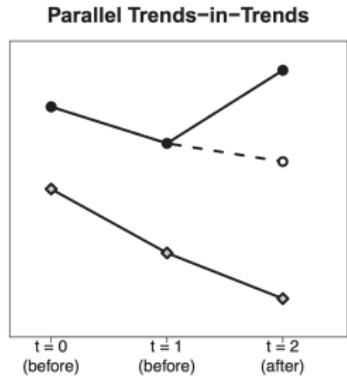
Benefit 3: A more flexible assumption



Trend of Treatment Group
 $(-2, -1)$

Trend of Control Group
 $(-3.5, -2.5)$

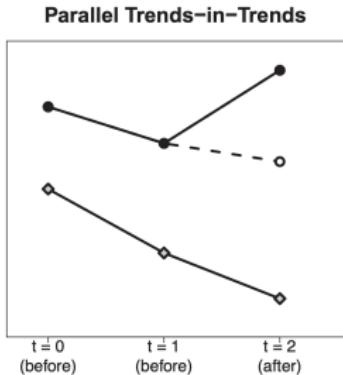
Benefit 3: A more flexible assumption



ASSUMPTION 3 (Parallel Trends-in-Trends)

$$\underbrace{\{E[Y_{i2}(0) | G_i = 1] - E[Y_{i1}(0) | G_i = 1]\}}_{\text{Trend of the treatment group from } t=1 \text{ to } t=2} - \underbrace{\{E[Y_{i1}(0) | G_i = 1] - E[Y_{i0}(0) | G_i = 1]\}}_{\text{Trend of the treatment group from } t=0 \text{ to } t=1}$$
$$= \underbrace{\{E[Y_{i2}(0) | G_i = 0] - E[Y_{i1}(0) | G_i = 0]\}}_{\text{Trend of the control group from } t=1 \text{ to } t=2} - \underbrace{\{E[Y_{i1}(0) | G_i = 0] - E[Y_{i0}(0) | G_i = 0]\}}_{\text{Trend of the control group from } t=0 \text{ to } t=1}.$$

Benefit 3: A more flexible assumption



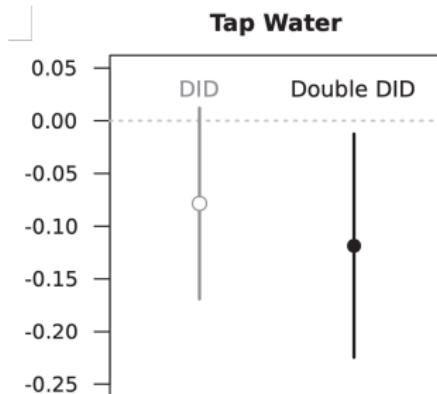
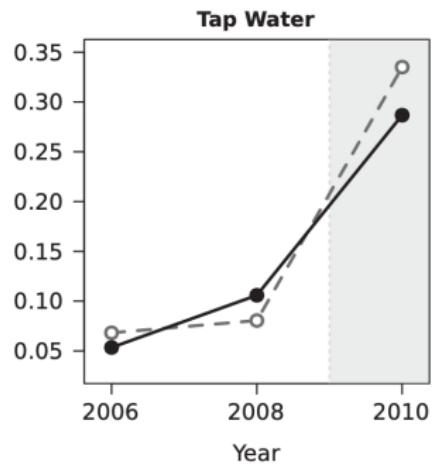
ASSUMPTION 3 (Parallel Trends-in-Trends)

$$\begin{aligned} & \underbrace{\{\mathbb{E}[Y_{i2}(0) | G_i = 1] - \mathbb{E}[Y_{i1}(0) | G_i = 1]\}}_{\text{Trend of the treatment group from } t=1 \text{ to } t=2} - \underbrace{\{\mathbb{E}[Y_{i1}(0) | G_i = 1] - \mathbb{E}[Y_{i0}(0) | G_i = 1]\}}_{\text{Trend of the treatment group from } t=0 \text{ to } t=1} \\ &= \underbrace{\{\mathbb{E}[Y_{i2}(0) | G_i = 0] - \mathbb{E}[Y_{i1}(0) | G_i = 0]\}}_{\text{Trend of the control group from } t=1 \text{ to } t=2} - \underbrace{\{\mathbb{E}[Y_{i1}(0) | G_i = 0] - \mathbb{E}[Y_{i0}(0) | G_i = 0]\}}_{\text{Trend of the control group from } t=0 \text{ to } t=1}. \end{aligned}$$

Sequential DID Estimator

$$\widehat{\tau}_{\text{s-DID}} = \left\{ \left(\frac{\sum_{i: G_i=1} Y_{i2}}{n_{12}} - \frac{\sum_{i: G_i=1} Y_{i1}}{n_{11}} \right) - \left(\frac{\sum_{i: G_i=0} Y_{i2}}{n_{02}} - \frac{\sum_{i: G_i=0} Y_{i1}}{n_{01}} \right) \right\} \\ - \left\{ \left(\frac{\sum_{i: G_i=1} Y_{i1}}{n_{11}} - \frac{\sum_{i: G_i=1} Y_{i0}}{n_{10}} \right) - \left(\frac{\sum_{i: G_i=0} Y_{i1}}{n_{01}} - \frac{\sum_{i: G_i=0} Y_{i0}}{n_{00}} \right) \right\},$$

Benefit 3: A more flexible assumption

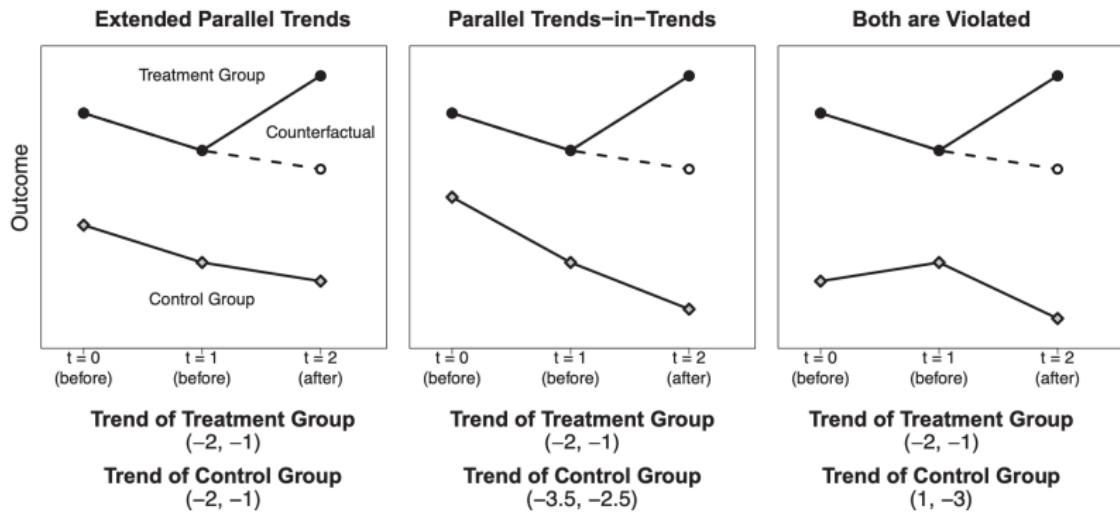


Benefits of multiple pre-treatment periods

1. assess underlying assumptions
2. improve estimation accuracy
3. allow for a more flexible parallel trends assumption

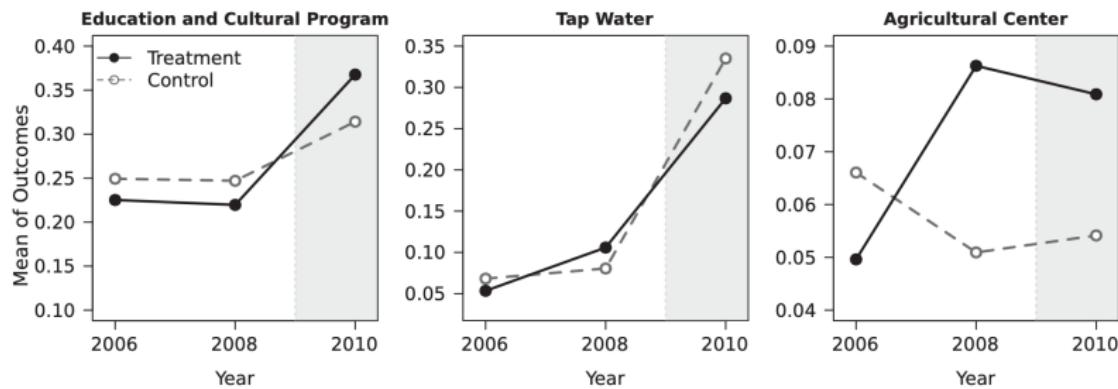
Benefits of multiple pre-treatment periods

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Benefits of multiple pre-treatment periods

1. assess underlying assumptions
2. improve estimation accuracy
3. allow for a more flexible parallel trends assumption



Interrupted time series¹

¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). International Journal of Epidemiology, 46(1), 348-355.

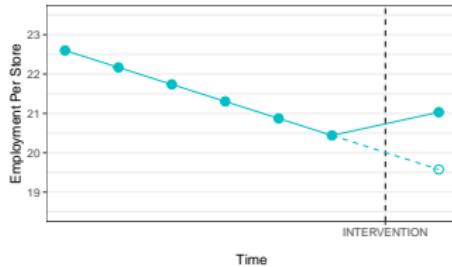
Interrupted time series¹

You study one unit. It is untreated. Then it is treated.

¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). International Journal of Epidemiology, 46(1), 348-355.

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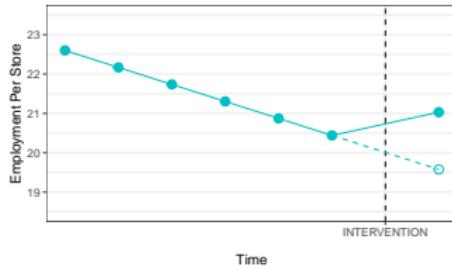


In what settings does this work well?

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Interrupted time series¹

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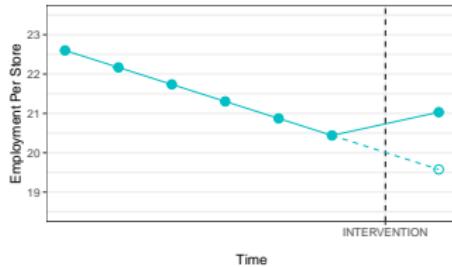
In what settings does this work well?

- When you have a strong pre-treatment trend to forecast Y_t^0

¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). *International Journal of Epidemiology*, 46(1), 348-355.

Interrupted time series¹

You study one unit. It is untreated. Then it is treated.



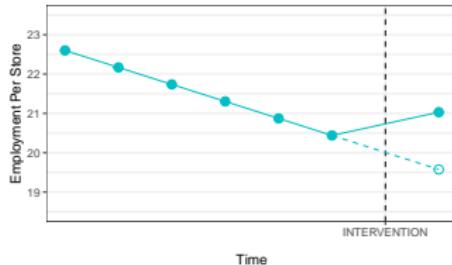
In what settings does this work well?

- When you have a strong pre-treatment trend to forecast Y_t^0
- When you don't have a comparable unit that is never treated

¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). *International Journal of Epidemiology*, 46(1), 348-355.

Interrupted time series¹

You study one unit. It is untreated. Then it is treated.



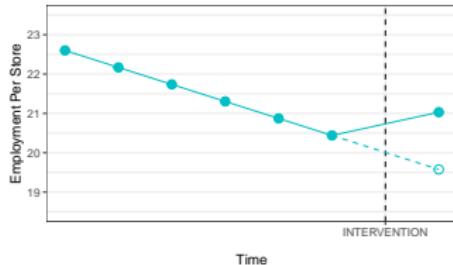
Theoretical Estimand

$$E(Y^1 - Y^0 \mid T > t_{\text{Intervention}})$$

¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). *International Journal of Forecasting*, *33*(3), 802–818. doi:10.1016/j.ijforecast.2016.09.002

Interrupted time series¹

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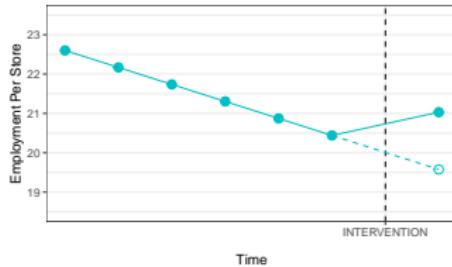


Identifying Assumption

¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). *International Journal of Epidemiology*, 46(1), 348-355.

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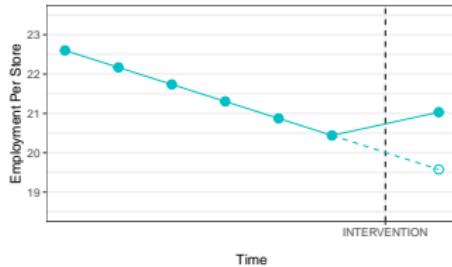
Identifying Assumption

- In the absence of the intervention,
the pre-intervention trend in Y^0 would have continued

¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). *International Journal of Epidemiology*, 46(1), 348-355.

Interrupted time series¹

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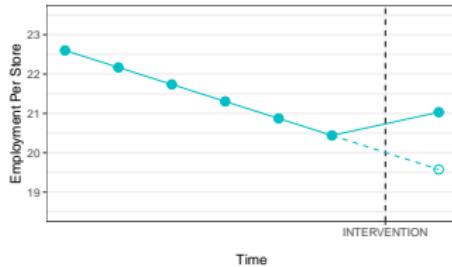


Concrete steps:

¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). International Journal of Epidemiology, 46(1), 348-355.

Interrupted time series¹

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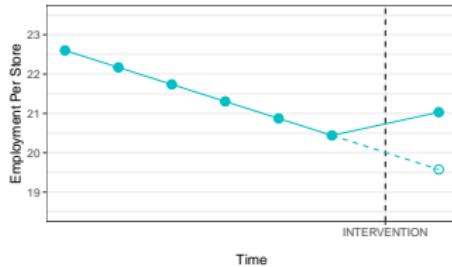
Concrete steps:

1. Learn a model on the pre-treatment period

¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). International Journal of Epidemiology, 46(1), 348-355.

Interrupted time series¹

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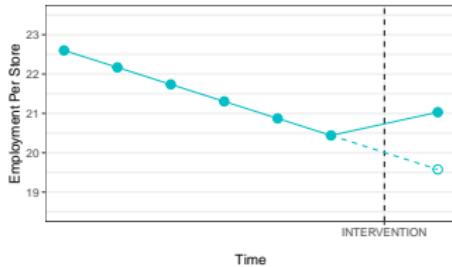
Concrete steps:

1. Learn a model on the pre-treatment period
 - Evaluation metric: Forecast within the pre-treatment period

¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). International Journal of Epidemiology, 46(1), 348-355.

Interrupted time series¹

You study one unit. It is untreated. Then it is treated.



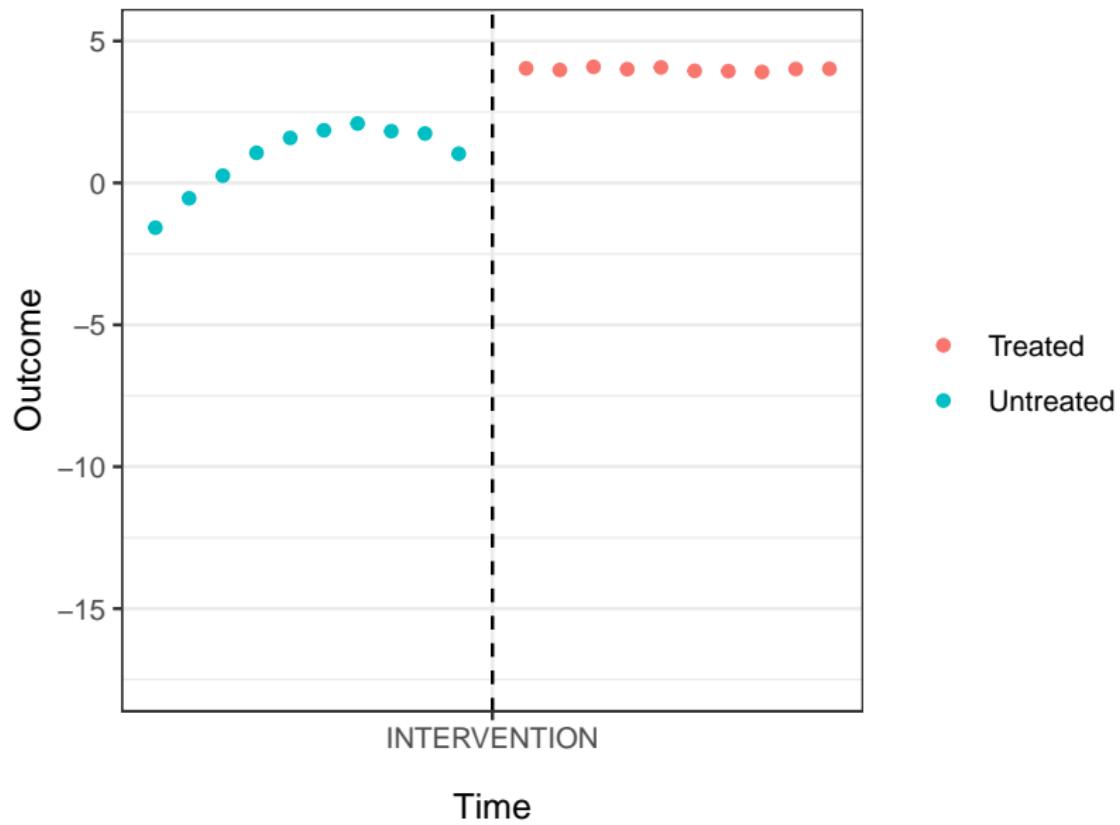
Concrete steps:

1. Learn a model on the pre-treatment period
 - Evaluation metric: Forecast within the pre-treatment period
2. Forecast Y^0 for the post-treatment period

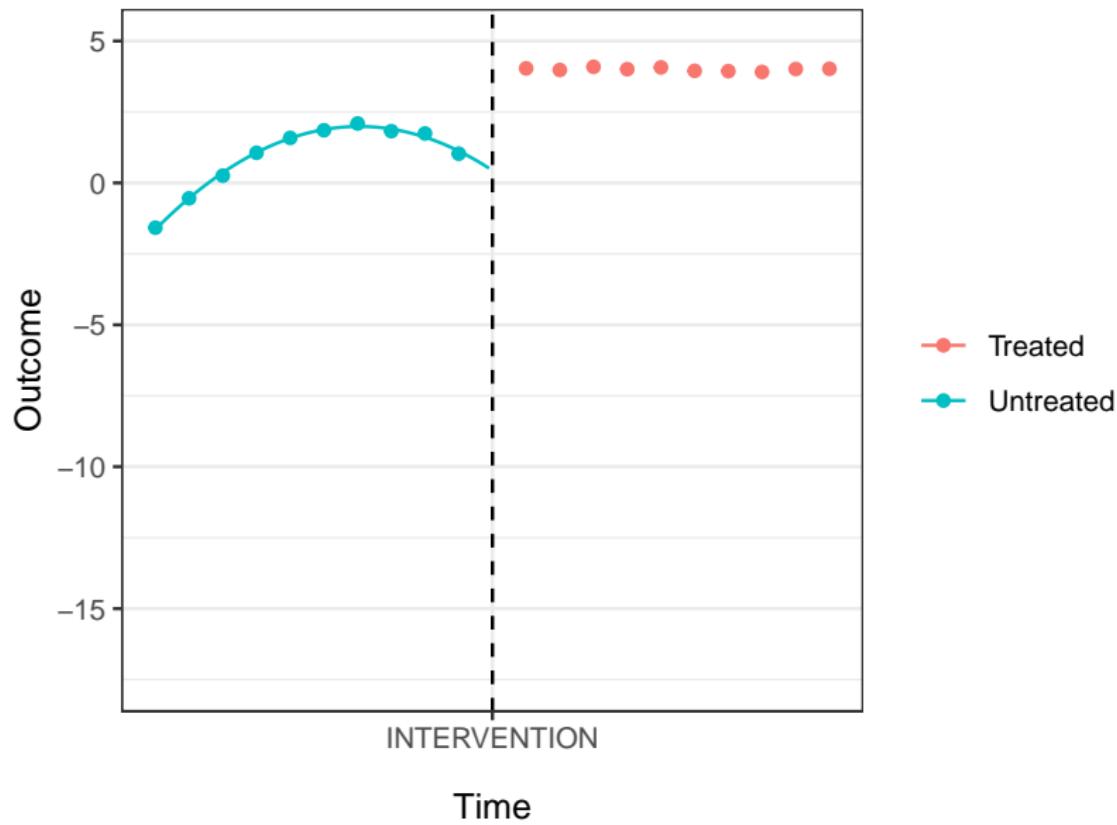
¹Bernal, J. L., Cummins, S., & Gasparrini, A. (2017). [Interrupted time series regression for the evaluation of public health interventions: A tutorial](#). International Journal of Epidemiology, 46(1), 348-355.

Interrupted time series: When it becomes doubtful

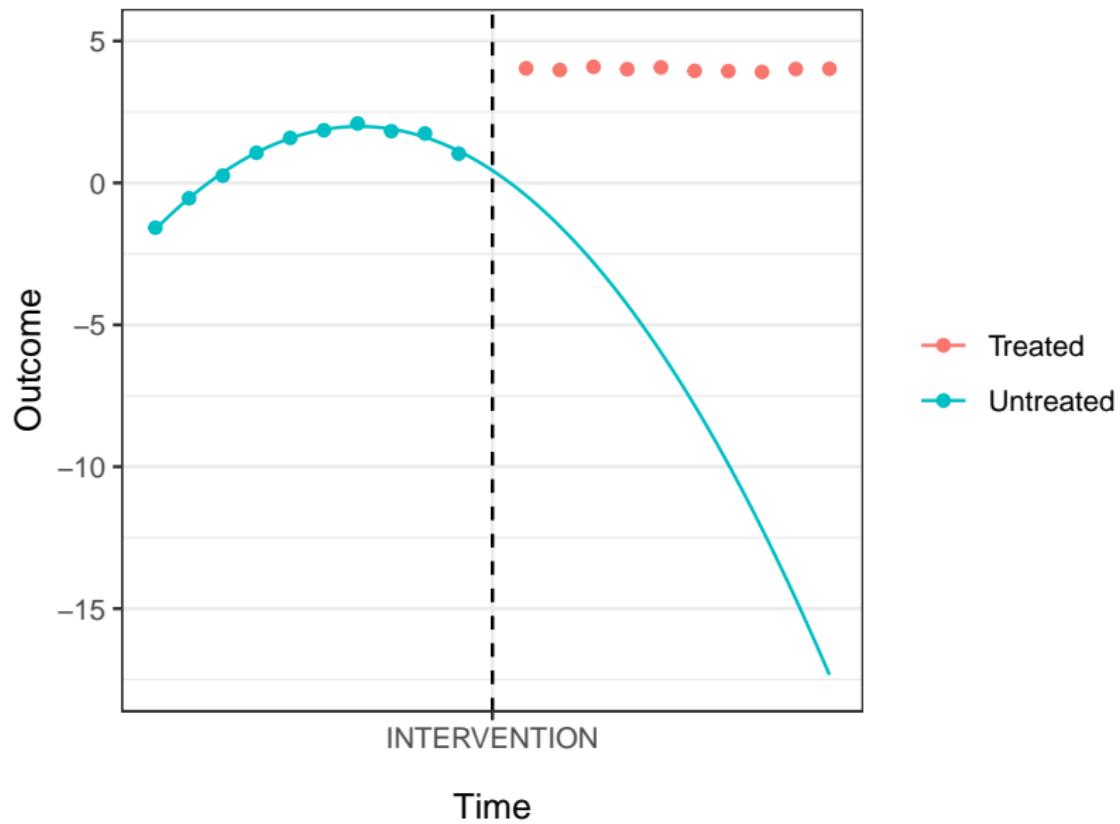
Interrupted time series: When it becomes doubtful



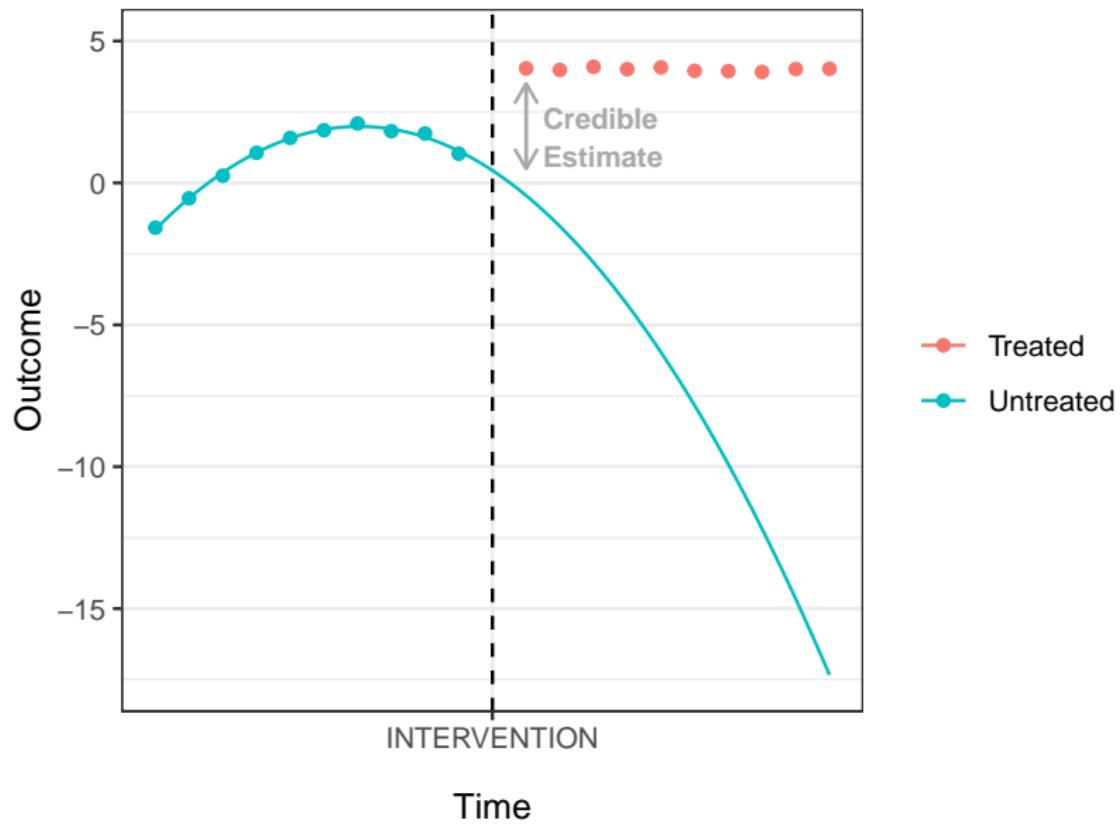
Interrupted time series: When it becomes doubtful



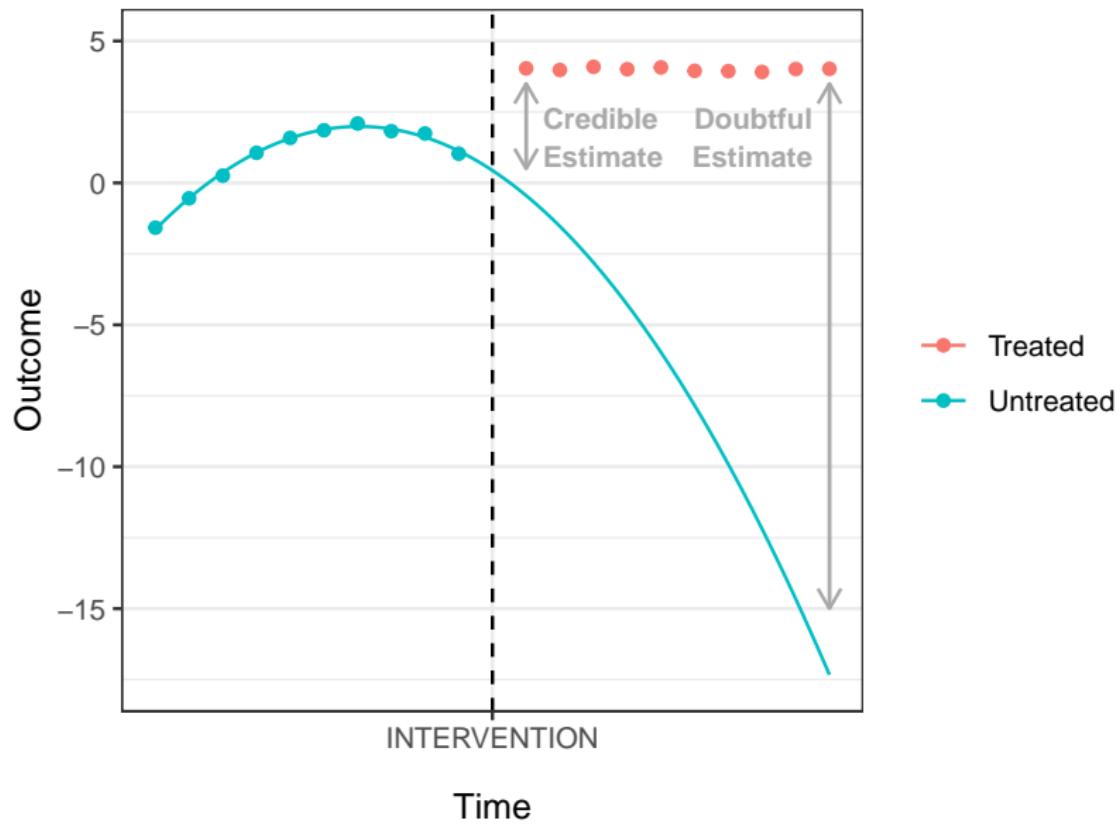
Interrupted time series: When it becomes doubtful



Interrupted time series: When it becomes doubtful



Interrupted time series: When it becomes doubtful



Interrupted time series: Recap

- ▶ ITS applies when treatment turns on at one time for all units
- ▶ ITS requires a parametric model to extrapolate
- ▶ ITS is most credible near the time when treatment turns on

When to use each method

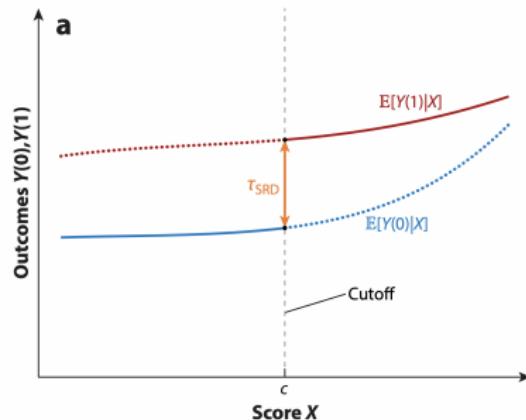
- ▶ Difference in difference
 - ▶ One unit becomes treated
 - ▶ One unit never becomes treated
 - ▶ The trends in Y^0 are parallel
 - ▶ Interrupted time series
 - ▶ Everyone becomes treated at $X = c$
 - ▶ You believe you can forecast Y^0 from $X < c$ to $X > c$
- New Jersey
Pennsylvania
- New drug
Deaths would
have been stable

Regression discontinuity²

²Cattaneo, M. D., & Titiunik, R. (2022). Regression discontinuity designs. Annual Review of Economics, 14, 821–851.

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a

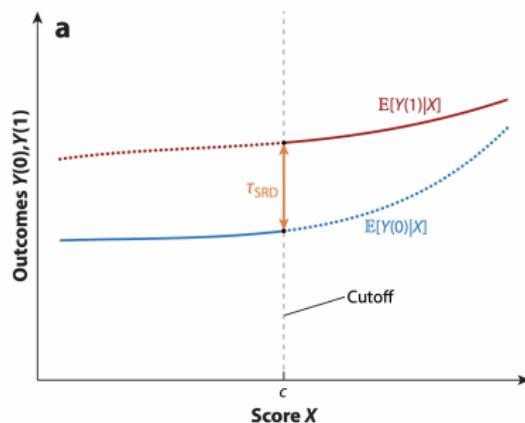


²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#).

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a

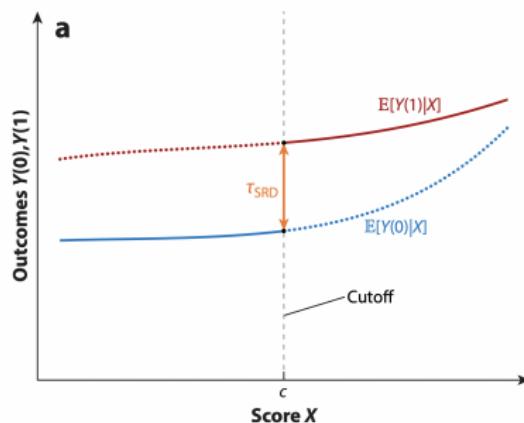
Examples



²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#).

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



Examples

X is PSAT test score

c is a score cutoff

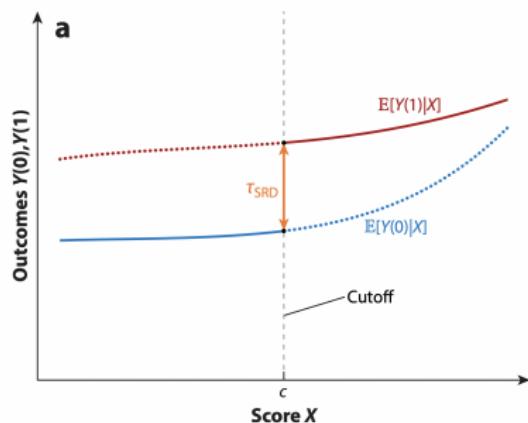
A is National Merit Scholarship

(Thistlewaite & Campbell 1960)

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Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



Examples

X is PSAT test score

c is a score cutoff

A is National Merit Scholarship

(Thistlewaite & Campbell 1960)

X is vote share

c is 50%

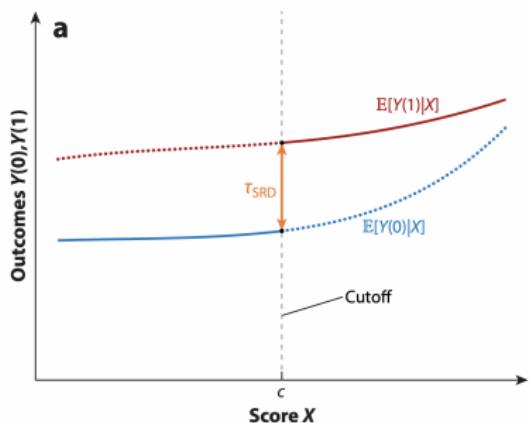
A is winning the election

(De la Cuesta & Imai 2016)

²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#).

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



Examples

X is PSAT test score

c is a score cutoff

A is National Merit Scholarship

(Thistlewaite & Campbell 1960)

X is vote share

c is 50%

A is winning the election

(De la Cuesta & Imai 2016)

X is date

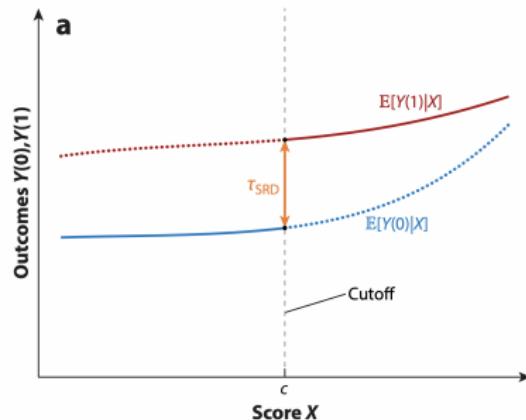
c is 2am Nov 6 2022

A is hours of PM darkness

²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#).

Regression discontinuity²

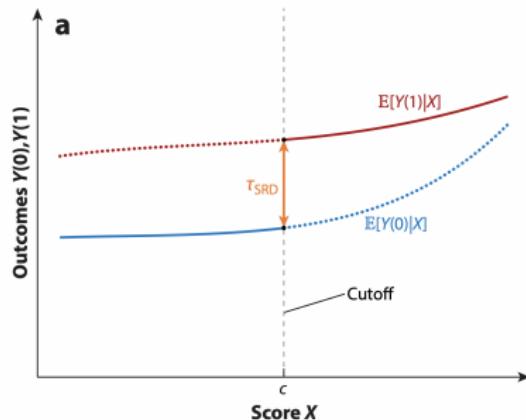
Cattaneo & Titiunik 2022 Fig 1a



²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#).

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a

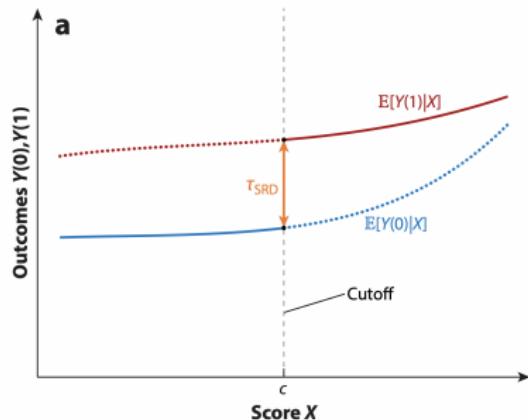


Theoretical Estimand
 $E(Y(1) - Y(0) | X = c)$

²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#). Annual Review of Economics, 14, 821–851.

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



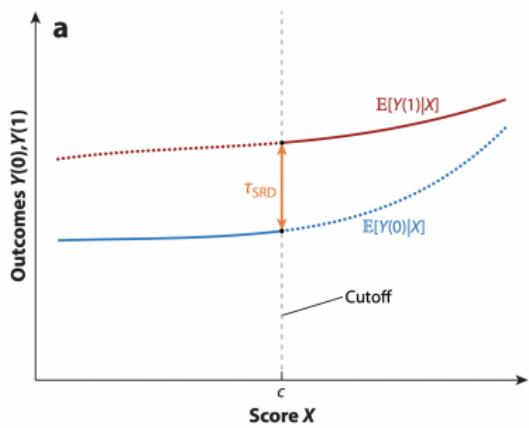
Theoretical Estimand
 $E(Y(1) - Y(0) | X = c)$

Empirical Estimand
 $\lim_{x \downarrow c} E(Y | X = x)$
—
 $\lim_{x \uparrow c} E(Y | X = x)$

²Cattaneo, M. D., & Titiunik, R. (2022). Regression discontinuity designs. Annual Review of Economics, 14, 821–851.

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



Theoretical Estimand
 $E(Y(1) - Y(0) | X = c)$

Empirical Estimand
 $\lim_{x \downarrow c} E(Y | X = x)$
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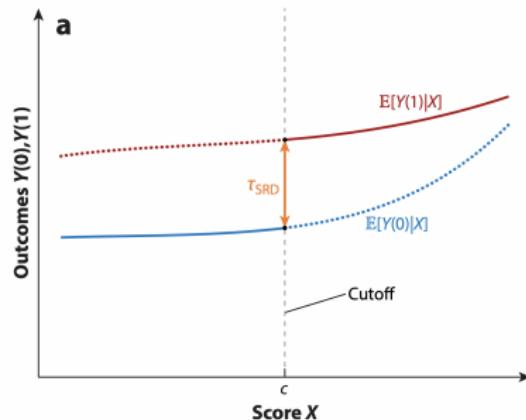
Identifying Assumptions
 $E(Y(1) | X = x)$ and
 $E(Y(0) | X = x)$ are
continuous at $x = c$

and $f_X(x) > 0$ for x near c

²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#). Annual Review of Economics, 14, 821–851.

Regression discontinuity²

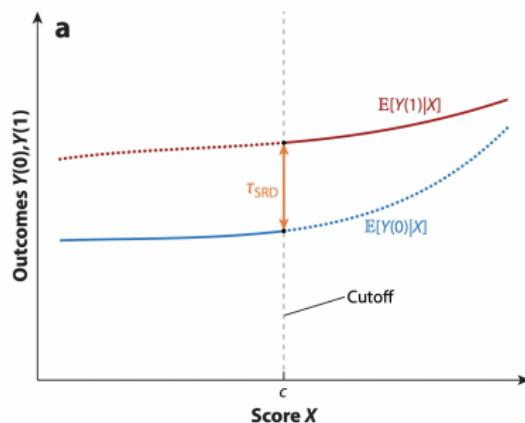
Cattaneo & Titiunik 2022 Fig 1a



²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#).

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



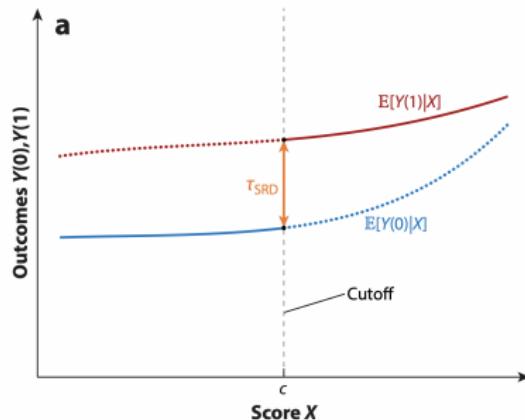
Promises of RD

Drawbacks of RD

²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#). Annual Review of Economics, 14, 821–851.

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



Promises of RD

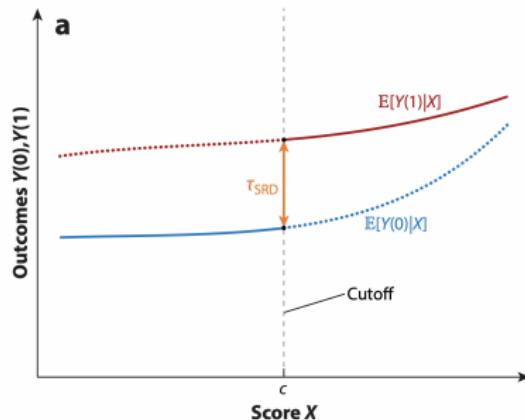
— Highly credible

Drawbacks of RD

²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#).

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



Promises of RD

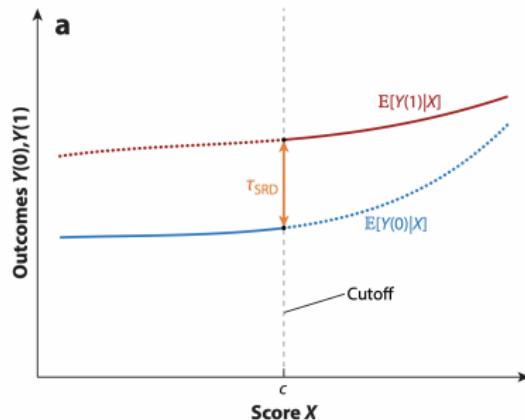
- Highly credible
- Easy to visualize

Drawbacks of RD

²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#). Annual Review of Economics, 14, 821–851.

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



Promises of RD

- Highly credible
- Easy to visualize

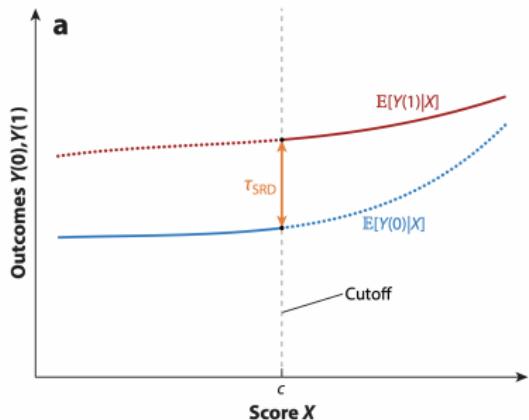
Drawbacks of RD

- Local to $X = c$

²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#).

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



Promises of RD

- Highly credible
- Easy to visualize

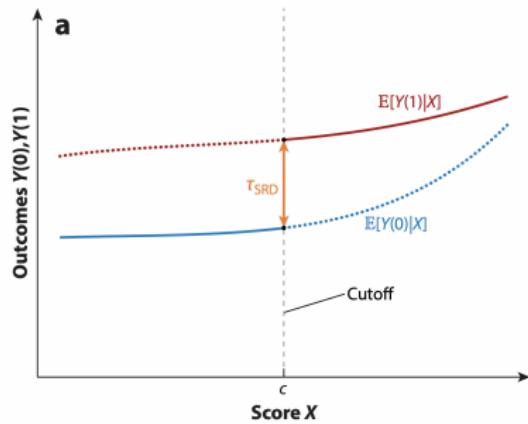
Drawbacks of RD

- Local to $X = c$
- Sensitive to sorting

²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#).

Regression discontinuity²

Cattaneo & Titiunik 2022 Fig 1a



Promises of RD

- Highly credible
- Easy to visualize

Drawbacks of RD

- Local to $X = c$
- Sensitive to sorting

(people moving
strategically over
the cutoff)

²Cattaneo, M. D., & Titiunik, R. (2022). [Regression discontinuity designs](#).

When to use each method

- ▶ Difference in difference
 - ▶ One unit becomes treated
 - ▶ One unit never becomes treated
 - ▶ The trends in Y^0 are parallel
 - ▶ Interrupted time series
 - ▶ Everyone becomes treated at $X = c$
 - ▶ You believe you can forecast Y^0 from $X < c$ to $X > c$
 - ▶ Regression discontinuity
 - ▶ Everyone becomes treated at $X = c$
 - ▶ You want a local estimate $E(Y^1 - Y^0 | X = c)$ at the cutoff
 - ▶ Y^0 and Y^1 are continuous at $X = c$
- | | |
|------------------|-------------------------------|
| New Jersey | Pennsylvania |
| New drug | Deaths would have been stable |
| Win the election | Close elections |

Synthetic control³

³Abadie, A., Diamond, A., & Hainmueller, J. (2010). [Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program](#). Journal of the American Statistical Association, 105(490), 493-505.

Synthetic control³

In 1988, California implemented a tobacco control program

³Abadie, A., Diamond, A., & Hainmueller, J. (2010). [Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program](#). Journal of the American Statistical Association, 105(490), 493-505.

Synthetic control³

- In 1988, California implemented a tobacco control program
- New tax: 25 cents per pack

³Abadie, A., Diamond, A., & Hainmueller, J. (2010). [Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program](#). Journal of the American Statistical Association, 105(490), 493-505.

Synthetic control³

In 1988, California implemented a tobacco control program

- ▶ New tax: 25 cents per pack
- ▶ Money earmarked for smoking-reduction programs

³Abadie, A., Diamond, A., & Hainmueller, J. (2010). [Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program](#). Journal of the American Statistical Association, 105(490), 493-505.

Synthetic control³

In 1988, California implemented a tobacco control program

- ▶ New tax: 25 cents per pack
- ▶ Money earmarked for smoking-reduction programs

How much did it reduce CA cigarette sales in 1990? 1995? 2000?

³Abadie, A., Diamond, A., & Hainmueller, J. (2010). [Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program](#). Journal of the American Statistical Association, 105(490), 493-505.

Synthetic control⁴

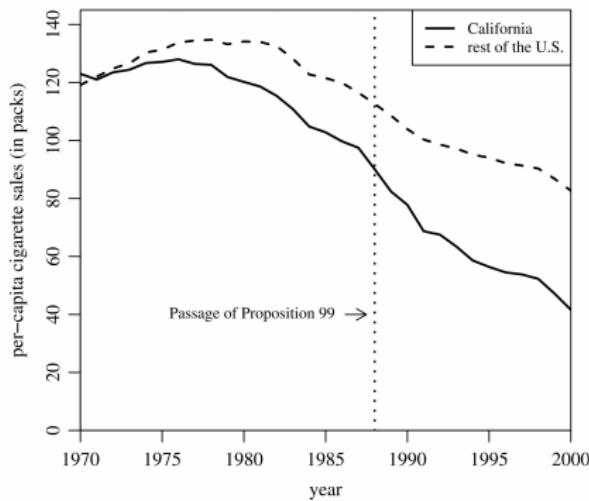


Figure 1. Trends in per-capita cigarette sales: California vs. the rest of the United States.

Can't use RD

- Effect at 1988 not of interest

Can't use ITS

- Hard to extrapolate Y^0 trend

Can't use DID

- No other state like CA

Idea: Create a
synthetic CA
to estimate
 $Y_{CA,t}^0$ for $t \geq 1988$

⁴Abadie, A., Diamond, A., & Hainmueller, J. (2010). **Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program**. Journal of the American Statistical Association, 105(490), 493-505.

Synthetic control⁵

Synthetic CA as a weighted average of other states

Table 1. Cigarette sales predictor means

Variables	California		Average of 38 control states
	Real	Synthetic	
Ln(GDP per capita)	10.08	9.86	9.86
Percent aged 15–24	17.40	17.40	17.29
Retail price	89.42	89.41	87.27
Beer consumption per capita	24.28	24.20	23.75
Cigarette sales per capita 1988	90.10	91.62	114.20
Cigarette sales per capita 1980	120.20	120.43	136.58
Cigarette sales per capita 1975	127.10	126.99	132.81

NOTE: All variables except lagged cigarette sales are averaged for the 1980–1988 period (beer consumption is averaged 1984–1988). GDP per capita is measured in 1997 dollars, retail prices are measured in cents, beer consumption is measured in gallons, and cigarette sales are measured in packs.

⁵Abadie, A., Diamond, A., & Hainmueller, J. (2010). [Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program](#). Journal of the American Statistical Association, 105(490), 493–505.

Synthetic control⁶

Synthetic CA as a weighted average of other states

Theoretical Estimand: $\tau(t) = Y_{\text{CA},t}^1 - Y_{\text{CA},t}^0$ $t \geq 1988$

⁶Abadie, A., Diamond, A., & Hainmueller, J. (2010). [Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program](#). Journal of the American Statistical Association, 105(490), 493-505.

Synthetic control⁶

Synthetic CA as a weighted average of other states

$$\text{Theoretical Estimand: } \tau(t) = Y_{\text{CA},t}^1 - Y_{\text{CA},t}^0 \quad t \geq 1988$$

$$\text{Empirical Estimand: } \theta(t) = Y_{\text{CA},t}^1 - Y_{\text{SyntheticCA},t}^0 \quad t \geq 1988$$

⁶Abadie, A., Diamond, A., & Hainmueller, J. (2010). [Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program](#). Journal of the American Statistical Association, 105(490), 493-505.

Synthetic control⁶

Synthetic CA as a weighted average of other states

Theoretical Estimand: $\tau(t) = Y_{CA,t}^1 - Y_{CA,t}^0 \quad t \geq 1988$

Empirical Estimand: $\theta(t) = Y_{CA,t}^1 - Y_{SyntheticCA,t}^0 \quad t \geq 1988$

Identifying Assumption:

$$\underbrace{Y_{CA,t}^0}_{\text{Counterfactual}} = \underbrace{Y_{SyntheticCA,t}^0}_{\text{Factual}} \quad t \geq 1988$$

⁶Abadie, A., Diamond, A., & Hainmueller, J. (2010). [Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program](#). Journal of the American Statistical Association, 105(490), 493-505.

Synthetic control⁷

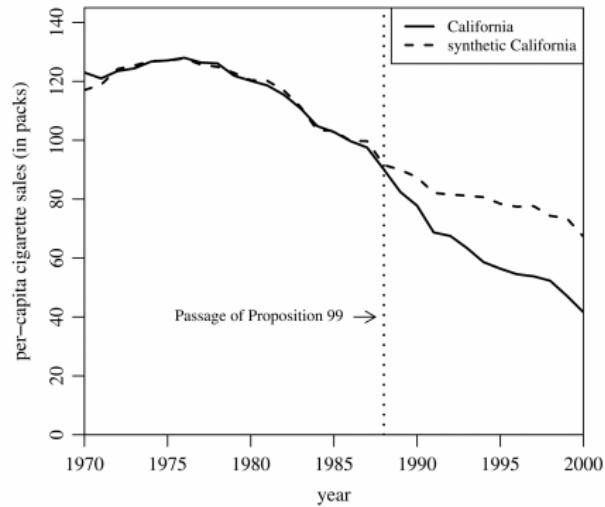


Figure 2. Trends in per-capita cigarette sales: California vs. synthetic California.

⁷Abadie, A., Diamond, A., & Hainmueller, J. (2010). [Synthetic control methods for comparative case studies: Estimating the effect of California's tobacco control program](#). Journal of the American Statistical Association, 105(490), 493-505.

When to use each method

- ▶ Difference in difference
 - ▶ One unit becomes treated New Jersey
 - ▶ One unit never becomes treated Pennsylvania
 - ▶ The trends in Y^0 are parallel
- ▶ Interrupted time series
 - ▶ Everyone becomes treated at $X = c$ New drug
 - ▶ You believe you can forecast Y^0 Deaths would from $X < c$ to $X > c$ have been stable
- ▶ Regression discontinuity
 - ▶ Everyone becomes treated at $X = c$ Win the election
 - ▶ You want a local estimate $E(Y^1 - Y^0 | X = c)$ at the cutoff Close elections
 - ▶ Y^0 and Y^1 are continuous at $X = c$
- ▶ Synthetic control
 - ▶ One unit becomes treated California
 - ▶ Many units are never treated Other states
 - ▶ You want to extrapolate far from the cutoff 1988→2000

Discussion

- ▶ What data do you need to use the method?
- ▶ What are the most likely limitations?
- ▶ How would you generalize your method to settings where many units become treated, potentially at different time points (staggered adoption)?

Learning goals for today

At the end of class, you will be able to:

1. Recognize the promises and pitfalls of four methods to study the effects of treatments that turn on once
 - 1.1 Difference in difference (DID)
 - 1.2 Interrupted time series (ITS)
 - 1.3 Regression discontinuity (RD)
 - 1.4 Synthetic control (SC)