25. Future treatments as proxies for confounding Class guest: Felix Elwert University of Wisconsin, Madison

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Cornell Info 6751: Causal Inference in Observational Settings
Fall 2022

Learning goals for today

At the end of class, you will be able to:

1. Reason about when future treatments can proxy for unmeasured confounding

Note that this class is based on:

Elwert, F., & Pfeffer, F. T. (2022). The future strikes back: Using future treatments to detect and reduce hidden bias. *Sociological Methods & Research*, 51(3), 1014-1051.



Hernán & Cole 2009

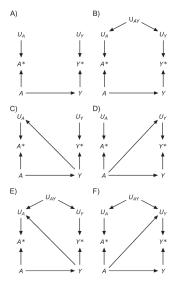


Figure 2. A structural classification of measurement error.

Hernán & Cole 2009

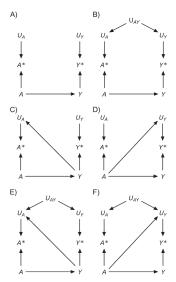
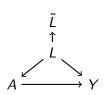


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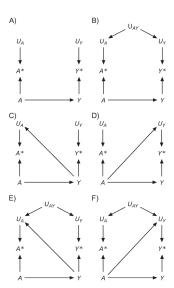
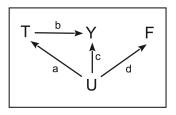


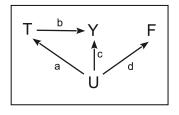
Figure 2. A structural classification of measurement error.

Elwert & Pfeffer 2022



If you measure U

$$\mathsf{E}(Y \mid A, U) = \alpha + \beta A + \gamma U$$



If you measure U

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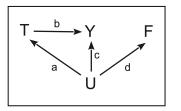
$$\beta = \frac{\text{Cov}(T, Y) - \text{Cov}(U, Y)\text{Cov}(U, A)}{1 - [\text{Cov}(U, A)]^2}$$

$$= \frac{(b + ac) - ((c + ab)a)}{1 - a^2}$$

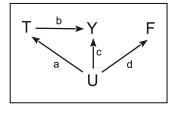
$$= \frac{b + ac - ac - a^2b}{1 - a^2}$$

$$= \frac{b(1 - a^2)}{1 - a^2}$$

$$= b$$

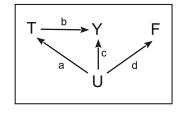


If you don't measure ${\it U}$



If you don't measure U

$$\mathsf{E}(Y\mid A,F) = \alpha + \beta A + \gamma F$$



If you don't measure U

$$\mathsf{E}(Y \mid A, F) = \alpha + \beta A + \gamma F$$

$$\beta = \frac{\operatorname{Cov}(A, Y) - \operatorname{Cov}(\tilde{L}, Y)\operatorname{Cov}(\tilde{L}, A)}{1 - \left[\operatorname{Cov}(\tilde{L}, A)\right]^2}$$

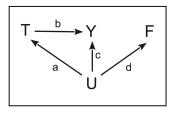
$$= \frac{\left(b + ac\right) - \left((cd + abd)ad\right)}{1 - a^2d^2}$$

$$= \frac{b + ac - acd^2 - a^2bd^2}{1 - a^2d^2}$$

$$= \frac{b(1 - a^2d^2) + ac(1 - d^2)}{1 - a^2d^2}$$

$$= b + \underbrace{ac}_{\text{Bias}}\underbrace{\frac{1 - d^2}{1 - a^2d^2}}_{\text{Blutipiler}}$$

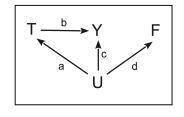
|M| < 1



Elwert & Pfeffer 2022

If you don't measure U

$$\mathsf{E}(Y\mid A,F) = \alpha + \beta A + \gamma F$$



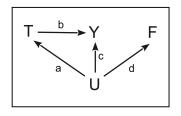
Control Estimator

$$\beta = b + \underbrace{ac}_{\substack{\text{Bias} \\ \text{without} \\ \text{control}}} \underbrace{\frac{1 - d^2}{1 - a^2 d^2}}_{\substack{\text{Bias} \\ \text{Multiplier} \\ |M| < 1}}$$

Elwert & Pfeffer 2022

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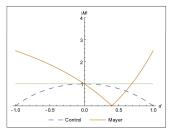
Difference (Mayer) Estimator

$$\beta - \gamma = b + \underbrace{\mathsf{ac}}_{\substack{\mathsf{Bias}\\ \mathsf{without}\\ \mathsf{control}}} \underbrace{\frac{a - d}{\underbrace{a - a^2 a}}}_{\substack{\mathsf{Bias}\\ \mathsf{Multiplier}\\ |M|}}$$

Elwert & Pfeffer 2022

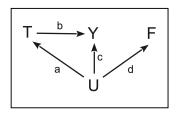
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Control Estimator

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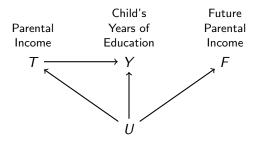
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$$\beta - \gamma = b + \underbrace{ac}_{\substack{\text{Bias} \\ \text{without} \\ \text{control}}} \underbrace{\frac{a - d}{\frac{a - a^2 d}{\frac{a - b}{\|a\|}}}}_{\substack{\text{Bias} \\ \text{Multiplier} \\ |M|}}$$

Parts of the paper we have not yet covered

Empirical example

What is the effect of log parental income on years of education?



Panel Study of Income Dynamics. Born 1956–1968. (n = 1,513)

- ► T log family income averaged at child age 13–17
- ► Y years of education by age 24
- ► F log family income at child age 25–29

Table 3. Estimating the Causal Effect of Parental Income on Children's Years of Education With and Without Future Treatments.

	(1)	(2)	(3)	(4)
Coefficients				
T: Parental income	.448 (.039)***	.319 (.049)***	.185 (.039)***	.118 (.041)**
F: Future parental income	,	`.274 [´] (.088)**	,	`.202 [´] (.076)**
X: Controls		,	Yes	`Yes
Difference in coefficients				
T-F		.045 (.126)		084 (.098)
Test of equality of coefficients of Model (1) versus (2): .0006 Model (1) versus (3): .0000 Model (3) versus (4): .0006 Model (1) versus (4): .0000	n T: þ values	` ,		,
N	1,513	1,513	1,513	1,513

Note. Standardized OLS regression coefficients (standard errors in parentheses); weighted. Significance tests for the difference between coefficients across models are using seemingly unrelated regression.

Statistical significance at $^{\dagger}p$ < .10. $^{*}p$ < .05. $^{**}p$ < .01. $^{***}p$ < .001 (two-tailed test).

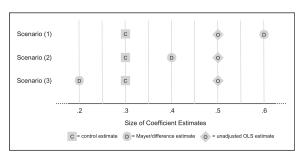
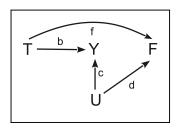
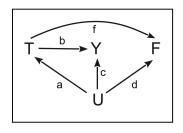


Figure 4. Illustration of the heuristic for choosing between estimates. The relative position of the control (*C*), difference (*D*), and unadjusted OLS (*O*) estimates can help the analyst decide between alternative estimates. In data generated by Figure 2, the location of the control estimate indicates the direction of unadjusted OLS bias (in this example, upward bias). In scenarios (1) and (2), the control estimate is preferred. In scenario (3), additional assumptions are needed to decide between the control and difference estimates.

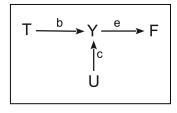
Challenge 1: True State Dependence



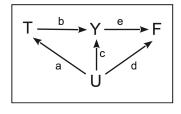
Challenge 2: Confounded True State Dependence



Challenge 3: Unconfounded Study with Selection



Challenge 4: Confounded Study with Selection



Challenge 4: Confounded Study with Selection

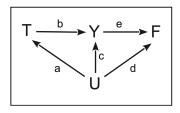


Table 2. Performance of the Control Estimator and the Mayer/Difference Estimator in the Presence of Selection and Weak to Moderate Path Parameters, $|\mathbf{p}| < .5$.

	Bias With			
Selection (e)	Control Estimator	Mayer/Difference Estimator		
Selection (e \neq 0)	Negligibly amplified or weakly reduced (see below)	Mostly amplified		
Mild selection (e ≤0.3)	Weakly reduced	Mostly amplified		

Previous results relied on a linear path model. Next results rely only on a DAG (nonparametric).

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Motivating (but incomplete) intuition:

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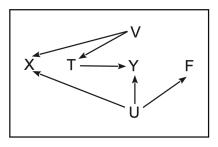
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Want to say:

If and only if Y is conditionally related to F, then $T \to Y$ is confounded. (this will require additional assumptions)

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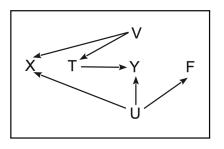
Is F related to Y? Is $T \rightarrow Y$ confounded given X?



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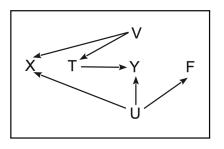
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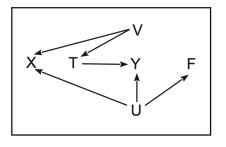
Want to say: If and only if Y is conditionally related to F, then $T \to Y$ is confounded

Is F related to Y? Yes. $Y \leftarrow U \rightarrow F$ Is $T \rightarrow Y$ confounded given X? Yes. $T \leftarrow V \rightarrow X \leftarrow U \rightarrow Y$



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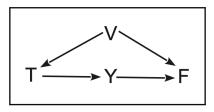
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√ Consistent with what we want to say

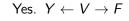
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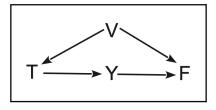
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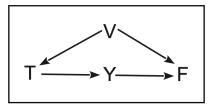
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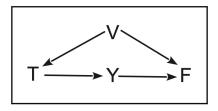
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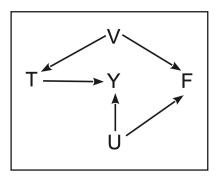
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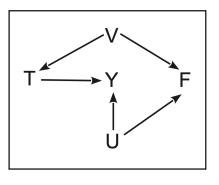
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Nonparametric results: Example 3.

Want to say: If and only if Y is conditionally related to F, then $T \to Y$ is confounded

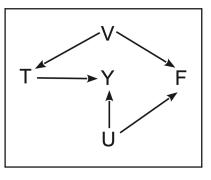
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Nonparametric results: Example 3.

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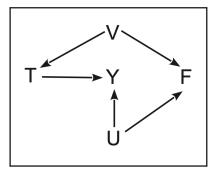
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Is F related to Y? Is $T \rightarrow Y$ confounded? Yes. $Y \leftarrow V \rightarrow F$ No.

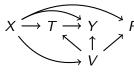


X Inconsistent with what we want to say

Two nonparametric results: A formal answer What does the relationship between F and Y tell us about confounding?

Two nonparametric results: A formal answer

What does the relationship between F and Y tell us about confounding?



Result 11. (Requires Assumption 1)

$$F \perp \!\!\! \perp Y \mid \{T,X\} \qquad \rightarrow \qquad \{Y^0,Y^1\} \perp \!\!\! \perp T \mid X$$

Result 12. (Requires Assumptions 2-3)

$$F \cancel{x} Y \mid \{T, X\} \qquad \rightarrow \qquad \{Y^0, Y^1\} \cancel{x} T \mid X$$

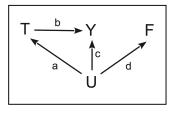
Assumption 1. There exists some unobserved V such that $V \to T$ and $V \not\!\!\!\perp\!\!\!\perp F \mid \{T,X\}$

Assumption 2. All unobserved causes of F also cause T

Assumption 3. Y does not directly or indirectly cause F

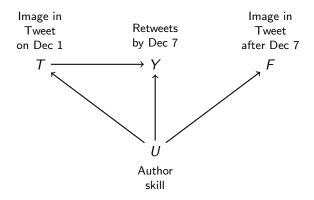
Discussion: Applied examples

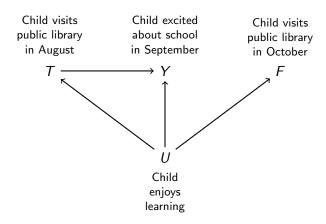
The usefulness of future treatments relies heavily on this DAG.

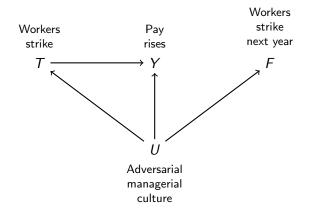


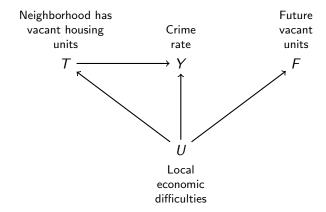
Exercise. Discuss the plausibility of this DAG in applied cases.

tinyurl.com/FutureTreatmentsExamples









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1. Reason about when future treatments can proxy for unmeasured confounding

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Feedback from class: tinyurl.com/CausalQuestions

No office hours today

Invitation: Cornell Sociology Department Colloquium

Friday, 3-4:15pm, Uris G08

Felix Elwert, University of Wisconsin, Madison
Rearranging the Desk Chairs: A Large Randomized
Field Experiment on the Effects of Close Contact on Interethnic
Relations