

1 GZ model

1.1 Mathematical model

The latent state transitions are

$$\begin{aligned} x_{a,t,i} &= \phi_a x_{a,t-1,i} + \mathbf{z}'_{a,t,i} \boldsymbol{\beta}_{\mathbf{Z}_a} + \varepsilon_{x_a}, \varepsilon_{x_a} \sim \mathcal{N}(0, \sigma_{x_a}^2), \\ x_{b,t,i} &= \phi_b x_{b,t-1,i} + \mathbf{z}'_{b,t,i} \boldsymbol{\beta}_{\mathbf{Z}_b} + \varepsilon_{x_b}, \varepsilon_{x_b} \sim \mathcal{N}(0, \sigma_{x_b}^2), \\ x_{p,t,i} &= \phi_p x_{p,t-1,i} + \mathbf{z}'_{p,t,i} \boldsymbol{\beta}_{\mathbf{Z}_p} + \varepsilon_{x_p}, \varepsilon_{x_p} \sim \mathcal{N}(0, \sigma_{x_p}^2), \\ x_{q,t,i} &= \phi_q x_{q,t-1,i} + \mathbf{z}'_{q,t,i} \boldsymbol{\beta}_{\mathbf{Z}_q} + \varepsilon_{x_q}, \varepsilon_{x_q} \sim \mathcal{N}(0, \sigma_{x_q}^2). \end{aligned}$$

The parameters of the state processes can be grouped as

$$\begin{aligned} \boldsymbol{\theta}_a &= (\phi_a, \boldsymbol{\beta}_{\mathbf{Z}_a}, \sigma_{x_a}^2), \\ \boldsymbol{\theta}_b &= (\phi_b, \boldsymbol{\beta}_{\mathbf{Z}_b}, \sigma_{x_b}^2), \\ \boldsymbol{\theta}_p &= (\phi_p, \boldsymbol{\beta}_{\mathbf{Z}_p}, \sigma_{x_p}^2), \\ \boldsymbol{\theta}_q &= (\phi_q, \boldsymbol{\beta}_{\mathbf{Z}_q}, \sigma_{x_q}^2). \end{aligned}$$

The measurement equation takes the general, highly nonlinear form, of

$$\mathbf{y}_{t,i} = g(\exp(x_{a,t,i}), \exp(x_{b,t,i}), \exp(x_{p,t,i}), \exp(x_{q,t,i})).$$

This implies $\forall t = 1, \dots, T$ and $\forall i = 1, \dots, N$

$$\begin{aligned} p_{\boldsymbol{\theta}_a}(x_{a,t,i} | x_{a,t-1,i}, \mathbf{z}'_{a,t,i}) &= \mathcal{N}(x_{a,t,i} | \phi_a x_{a,t-1,i} + \mathbf{z}'_{a,t,i} \boldsymbol{\beta}_{\mathbf{Z}_a}, \sigma_{x_a}^2), \\ p_{\boldsymbol{\theta}_b}(x_{b,t,i} | x_{b,t-1,i}, \mathbf{z}'_{b,t,i}) &= \mathcal{N}(x_{b,t,i} | \phi_b x_{b,t-1,i} + \mathbf{z}'_{b,t,i} \boldsymbol{\beta}_{\mathbf{Z}_b}, \sigma_{x_b}^2), \\ p_{\boldsymbol{\theta}_p}(x_{p,t,i} | x_{p,t-1,i}, \mathbf{z}'_{p,t,i}) &= \mathcal{N}(x_{p,t,i} | \phi_p x_{p,t-1,i} + \mathbf{z}'_{p,t,i} \boldsymbol{\beta}_{\mathbf{Z}_p}, \sigma_{x_p}^2), \\ p_{\boldsymbol{\theta}_q}(x_{q,t,i} | x_{q,t-1,i}, \mathbf{z}'_{q,t,i}) &= \mathcal{N}(x_{q,t,i} | \phi_q x_{q,t-1,i} + \mathbf{z}'_{q,t,i} \boldsymbol{\beta}_{\mathbf{Z}_q}, \sigma_{x_q}^2), \\ p(\mathbf{y}_{t,i} | x_{a,t,i}, x_{b,t,i}, x_{p,t,i}, x_{q,t,i}) &= \mathcal{MNL}(\mathbf{y}_{t,i} | \pi_{it}^{(1)}, \dots, \pi_{it}^{(M_{it})}), \end{aligned}$$

where $\pi_{it}^{(k)} = (F_{\text{GB2}}(c_{it}^{(k)}; x_{a,t,i}, x_{b,t,i}, x_{p,t,i}, x_{q,t,i}) - F_{\text{GB2}}(c_{it}^{(k-1)}; x_{a,t,i}, x_{b,t,i}, x_{p,t,i}, x_{q,t,i}))$.

The income distribution function is a four-parameter GB2:

$$\begin{aligned} F(c_{it}^{(k)}; x_{a,t,i}, x_{b,t,i}, x_{p,t,i}, x_{q,t,i}) &= B(d_{t,i}^{(k)}; x_{p,t,i}, x_{q,t,i}) = \frac{\int_0^{d_{t,i}^{(k)}} t^{\exp(x_{p,t,i})-1} (1-t)^{\exp(x_{q,t,i})-1} dt}{B(\exp(x_{p,t,i}), \exp(x_{q,t,i}))}, \\ d_{t,i}^{(k)} &= \frac{(c_{it}^{(k)} / \exp(x_{b,t,i}))^{\exp(x_{a,t,i})}}{1 + (c_{it}^{(k)} / \exp(x_{b,t,i}))^{\exp(x_{a,t,i})}}. \end{aligned}$$

Finally, we need prior assumptions on $(\boldsymbol{\theta}_a, \boldsymbol{\theta}_b, \boldsymbol{\theta}_p, \boldsymbol{\theta}_q)$:

$$\begin{aligned}
\sigma_{x_a}^2 &\sim \mathcal{IG}(a_{x_a}, b_{x_a}) , a_{x_a} = b_{x_a} = 0.001 , \\
\sigma_{x_b}^2 &\sim \mathcal{IG}(a_{x_b}, b_{x_b}) , a_{x_b} = b_{x_b} = 0.001 , \\
\sigma_{x_p}^2 &\sim \mathcal{IG}(a_{x_p}, b_{x_p}) , a_{x_p} = b_{x_p} = 0.001 , \\
\sigma_{x_q}^2 &\sim \mathcal{IG}(a_{x_q}, b_{x_q}) , a_{x_q} = b_{x_q} = 0.001 , \\
(\phi_a, \boldsymbol{\beta}_{\mathbf{Z}_a}) &\sim \mathcal{N}(0, \mathbf{I}_a) , \\
(\phi_b, \boldsymbol{\beta}_{\mathbf{Z}_b}) &\sim \mathcal{N}(0, \mathbf{I}_b) , \\
(\phi_p, \boldsymbol{\beta}_{\mathbf{Z}_p}) &\sim \mathcal{N}(0, \mathbf{I}_p) , \\
(\phi_q, \boldsymbol{\beta}_{\mathbf{Z}_q}) &\sim \mathcal{N}(0, \mathbf{I}_q) ,
\end{aligned}$$

where $\mathbf{I}_a, \mathbf{I}_b, \mathbf{I}_p, \mathbf{I}_q$ are identity matrices of appropriate dimension. Alternatively, instead of \mathbf{I}_a , one can model the variances of the parameters explicitly i.e. using $\sigma_{\boldsymbol{\beta}_{\mathbf{Z}_a}, \phi_a}^2 \times \mathbf{I}$ with an additional (hierarchical) $\mathcal{IG}(,)$ -prior on $\sigma_{\boldsymbol{\beta}_{\mathbf{Z}_a}, \phi_a}^2$ (and similarly for $\mathbf{I}_b, \mathbf{I}_p, \mathbf{I}_q$).

Let $x_{2:T}$ be a generic state proces for some fixed $i = 1, \dots, N$ i.e. either $\mathbf{x}_{a,2:T}, \mathbf{x}_{b,2:T}, \mathbf{x}_{p,2:T}$ or $\mathbf{x}_{q,2:T}$. Note that the model can be written in matrix form as:

$$\begin{aligned}
x_{2:T} &= x_{1:T-1}\phi + \mathbf{z}_{2:T}\boldsymbol{\beta}_{\mathbf{Z}} + \boldsymbol{\epsilon}_{x,2:T} , \\
x_{2:T} &= \mathbf{Z}_{2:T} \times (\phi, \boldsymbol{\beta}'_{\mathbf{Z}})' + \boldsymbol{\epsilon}_{x,2:T} ,
\end{aligned}$$

where $\mathbf{Z}_{2:T}$ is a matrix containing as first column $x_{1:T-1}$ and the remaining K regressors in $\mathbf{z}_{2:T}$. This makes it easier to calculate the Gibbs block for ϕ in the next sections.

1.2 Gibbs-Part: univariate (for one $i = 1, \dots, N$)

We derive $p(\boldsymbol{\theta}|x_{0:T}, \mathbf{y}_{1:T})$ for a particular $x_t \in \{x_{a,t,i}, x_{b,t,i}, x_{p,t,i}, x_{q,t,i}\}$ i.e. fixing the cross sectional unit i and picking one of the four GB2 parameters. The full probabilistic model with $\boldsymbol{\theta} = (\sigma_X^2, \phi, \boldsymbol{\beta}_Z)$ can then be factorized according to

$$\begin{aligned} p(\boldsymbol{\theta}, x_{0:T}, \mathbf{y}_{1:T}) &= p(\mathbf{y}_{1:T}|\boldsymbol{\theta}, x_{0:T}) p(x_{0:T}, \boldsymbol{\theta}) = \prod_{t=1}^T p(\mathbf{y}_t|x_t, \boldsymbol{\theta}) \prod_{t=1}^T p(x_t|x_{t-1}, \boldsymbol{\theta}) p(x_0|\boldsymbol{\theta}) p(\boldsymbol{\theta}) , \\ &= \prod_{t=1}^T (\mathbf{y}_t|\boldsymbol{\theta}, x_t) \\ &\times \frac{1}{(2\pi\sigma_X^2)^{T/2}} \prod_{t=1}^T \exp\left(-\frac{(x_t - \phi x_{t-1} - \mathbf{z}_t' \boldsymbol{\beta}_Z)^2}{2\sigma_X^2}\right) \\ &\times p(\boldsymbol{\theta}) . \end{aligned}$$

Then, the conditional parameter distributions are conjugate and given as

$$\begin{aligned} p(\sigma_X^2|x_{0:T}, \mathbf{y}_{1:T}) &= \frac{1}{(2\pi\sigma_X^2)^{T/2}} \prod_{t=1}^T \exp\left(-\frac{(x_t - \phi x_{t-1} - \mathbf{z}_t' \boldsymbol{\beta}_Z)^2}{2\sigma_X^2}\right) \\ &\times \frac{b_X^{a_X}}{\Gamma(a_X)} (\sigma_X^2)^{-a_X-1} \exp\left(-\frac{b_X}{\sigma_X^2}\right) \\ &\propto (\sigma_X^2)^{-(a_X+T/2)-1} \times \exp\left(-\frac{1}{\sigma_X^2} \left(b_X + \frac{\sum_{t=1}^T (x_t - \phi x_{t-1} - \mathbf{z}_t' \boldsymbol{\beta}_Z)^2}{2}\right)\right) . \end{aligned}$$

With e.g. $a_X = b_X = 0.001$, we have

$$\sigma_X^2 \sim \mathcal{IG}(a_X^*, b_X^*) , \quad a_X^* = a_X + T/2 , \quad b_X^* = b_X + \frac{\sum_{t=1}^T (x_t - \phi x_{t-1} - \mathbf{z}_t' \boldsymbol{\beta}_Z)^2}{2} .$$

For $\boldsymbol{\beta}_Z^* = (\phi, \boldsymbol{\beta}_Z')'$ with a normal prior $\boldsymbol{\beta}_Z^* \sim \mathcal{N}_{K+1}(\underline{\boldsymbol{\beta}}_Z^*, \underline{\boldsymbol{\Omega}}_Z)$ and the previous $x_{2:T} = \mathbf{Z}_{2:T} \boldsymbol{\beta}_Z^* + \boldsymbol{\varepsilon}_{x,2:T}$, we have

$$\begin{aligned} p(\boldsymbol{\beta}_Z^*|x_{2:T}, \mathbf{Z}_{2:T}, \sigma_X^2) &\propto \exp\left\{-\frac{1}{2} (x_{2:T} - \mathbf{Z}_{2:T} \boldsymbol{\beta}_Z^*)' \boldsymbol{\Omega}_{\boldsymbol{\varepsilon}_{x,2:T}}^{-1} (x_{2:T} - \mathbf{Z}_{2:T} \boldsymbol{\beta}_Z^*)\right\} \\ &\times \exp\left\{-\frac{1}{2} (\boldsymbol{\beta}_Z - \boldsymbol{\beta}_Z^*)' \underline{\boldsymbol{\Omega}}_Z^{-1} (\boldsymbol{\beta}_Z - \boldsymbol{\beta}_Z^*)\right\} \end{aligned}$$

Because we have

$$\begin{aligned}
(x_{2:T} - \mathbf{Z}_{2:T}\beta_{\mathbf{Z}}^*)' \boldsymbol{\Omega}_{\varepsilon_{x,2:T}}^{-1} (x_{2:T} - \mathbf{Z}_{2:T}\beta_{\mathbf{Z}}^*) &= \beta_{\mathbf{Z}}^{*'} \mathbf{Z}_{2:T}' \boldsymbol{\Omega}_{\varepsilon_{x,2:T}}^{-1} \mathbf{Z}_{2:T} \beta_{\mathbf{Z}}^* - 2\beta_{\mathbf{Z}}^{*'} \mathbf{Z}_{2:T}' \boldsymbol{\Omega}_{\varepsilon_{x,2:T}}^{-1} x_{2:T} \\
&\quad + x_{2:T}' \boldsymbol{\Omega}_{\varepsilon_{x,2:T}}^{-1} x_{2:T} \\
(\beta_{\mathbf{Z}}^* - \underline{\beta}_{\mathbf{Z}}^*)' \underline{\boldsymbol{\Omega}}_{\mathbf{Z}}^{-1} (\beta_{\mathbf{Z}}^* - \underline{\beta}_{\mathbf{Z}}^*) &= \beta_{\mathbf{Z}}^{*'} \underline{\boldsymbol{\Omega}}_{\mathbf{Z}}^{-1} \beta_{\mathbf{Z}}^* - 2\beta_{\mathbf{Z}}^{*'} \underline{\boldsymbol{\Omega}}_{\mathbf{Z}}^{-1} \underline{\beta}_{\mathbf{Z}}^* \\
&\quad + \underline{\beta}_{\mathbf{Z}}^{*'} * \underline{\boldsymbol{\Omega}}_{\mathbf{Z}}^{-1} \underline{\beta}_{\mathbf{Z}}^*
\end{aligned}$$

we obtain

$$\begin{aligned}
p(\beta_{\mathbf{Z}}^* | x_{2:T}, \mathbf{Z}_{2:T}, \sigma_X^2) &= \mathcal{N}_{K+1}(\overline{\beta}_{\mathbf{Z}}, \overline{\boldsymbol{\Omega}}_{\mathbf{Z}}) \\
\overline{\boldsymbol{\Omega}}_{\mathbf{Z}} &= [\mathbf{Z}_{2:T}' \boldsymbol{\Omega}_{\varepsilon_{x,2:T}}^{-1} \mathbf{Z}_{2:T} + \underline{\boldsymbol{\Omega}}_{\mathbf{Z}}^{-1}]^{-1} \\
\overline{\beta}_{\mathbf{Z}} &= \overline{\boldsymbol{\Omega}}_{\mathbf{Z}} \times [\mathbf{Z}_{2:T}' \boldsymbol{\Omega}_{\varepsilon_{x,2:T}}^{-1} x_{2:T} + \underline{\boldsymbol{\Omega}}_{\mathbf{Z}}^{-1} \underline{\beta}_{\mathbf{Z}}^*]
\end{aligned}$$

In Detail:

1.3 Gibbs-Part: multivariate (full cross section $\forall i = 1, \dots, N$)

We now consider vector valued processes stacked along the cross sectional dimension as e.g. $x_{t,1:N}, \mathbf{y}_{t,1:N}$. All the corresponding state transition and measurement equations factorize along the time dimension and given as

$$\begin{aligned} p_{\theta_a}(x_{a,t,1:N} | x_{a,t-1,1:N}, \mathbf{z}_{a,t,1:N}) &= \mathcal{N}_{1:N}(x_{a,t,1:N} | \phi_a x_{a,t-1,1:N} + \mathbf{z}_{a,t,1:N} \beta_{\mathbf{z}_a}, \sigma_{x_a}^2 \mathbf{I}_N) , \\ p_{\theta_b}(x_{b,t,1:N} | x_{b,t-1,1:N}, \mathbf{z}_{b,t,1:N}) &= \mathcal{N}_{1:N}(x_{b,t,1:N} | \phi_b x_{b,t-1,1:N} + \mathbf{z}_{b,t,1:N} \beta_{\mathbf{z}_b}, \sigma_{x_b}^2 \mathbf{I}_N) , \\ p_{\theta_p}(x_{p,t,1:N} | x_{p,t-1,1:N}, \mathbf{z}_{p,t,1:N}) &= \mathcal{N}_{1:N}(x_{p,t,1:N} | \phi_p x_{p,t-1,1:N} + \mathbf{z}_{p,t,1:N} \beta_{\mathbf{z}_p}, \sigma_{x_p}^2 \mathbf{I}_N) , \\ p_{\theta_q}(x_{q,t,1:N} | x_{q,t-1,1:N}, \mathbf{z}_{q,t,1:N}) &= \mathcal{N}_{1:N}(x_{q,t,1:N} | \phi_q x_{q,t-1,1:N} + \mathbf{z}_{q,t,1:N} \beta_{\mathbf{z}_q}, \sigma_{x_q}^2 \mathbf{I}_N) , \\ p(\mathbf{y}_{t,i} | x_{a,t,i}, x_{b,t,i}, x_{p,t,i}, x_{q,t,i}) &= \mathcal{MNL}(\mathbf{y}_{t,i} | \pi_{it}^{(1)}, \dots, \pi_{it}^{(M_{it})}) , \end{aligned}$$

with $\pi_{it}^{(k)} = (F_{\text{GB2}}(c_{it}^{(k)}; x_{a,t,i}, x_{b,t,i}, x_{p,t,i}, x_{q,t,i}) - F_{\text{GB2}}(c_{it}^{(k-1)}; x_{a,t,i}, x_{b,t,i}, x_{p,t,i}, x_{q,t,i}))$ and income distribution function as a four-parameter GB2

$$\begin{aligned} F(c_{it}^{(k)}; x_{a,t,i}, x_{b,t,i}, x_{p,t,i}, x_{q,t,i}) &= B(d_{t,i}^{(k)}; x_{p,t,i}, x_{q,t,i}) = \frac{\int_0^{d_{t,i}^{(k)}} t^{\exp(x_{p,t,i})-1} (1-t)^{\exp(x_{q,t,i})-1} dt}{B(\exp(x_{p,t,i}), \exp(x_{q,t,i}))} , \\ d_{t,i}^{(k)} &= \frac{(c_{it}^{(k)} / \exp(x_{b,t,i}))^{\exp(x_{a,t,i})}}{1 + (c_{it}^{(k)} / \exp(x_{b,t,i}))^{\exp(x_{a,t,i})}} . \end{aligned}$$

However, as all state transitions share the same structure, we derive them for a particular $x_{t,1:N} \in \{x_{a,t,1:N}, x_{b,t,1:N}, x_{p,t,1:N}, x_{q,t,1:N}\}$. Now, to obtain $p(\boldsymbol{\theta} | x_{0:T,1:N}, \mathbf{y}_{0:T,1:N})$, consider the full probabilistic model with $\boldsymbol{\theta} = (\sigma_X^2, \phi, \beta_{\mathbf{z}})$ as

$$\begin{aligned} p(\boldsymbol{\theta}, x_{0:T,1:N}, \mathbf{y}_{0:T,1:N}) &= p(\mathbf{y}_{0:T,1:N} | \boldsymbol{\theta}, x_{0:T,1:N}) p(x_{0:T,1:N} | \boldsymbol{\theta}) \\ &= \prod_{t=1}^T p(\mathbf{y}_{t,1:N} | x_{t,1:N}, \boldsymbol{\theta}) \prod_{t=1}^T p(x_{t,1:N} | x_{t-1,1:N}, \boldsymbol{\theta}) p(x_{0,1:N} | \boldsymbol{\theta}) p(\boldsymbol{\theta}) \\ &= \prod_{t=1}^T p(\mathbf{y}_{t,1:N} | x_{t,1:N}) \times \prod_{t=1}^T \frac{1}{(2\pi)^{N/2} (\det(\sigma_X^2 \mathbf{I}_N))^{1/2}} \\ &\quad \times \prod_{t=1}^T \exp\left(-\frac{1}{2\sigma_X^2} (x_{t,1:N} - \phi x_{t-1,1:N} - \mathbf{z}_{t,1:N} \beta_{\mathbf{z}})' (x_{t,1:N} - \phi x_{t-1,1:N} - \mathbf{z}_{t,1:N} \beta_{\mathbf{z}})\right) \\ &\quad \times p(\boldsymbol{\theta}) \\ &= \prod_{t=1}^T p(\mathbf{y}_{t,1:N} | x_{t,1:N}) \times (2\pi\sigma_X^2)^{-NT/2} \\ &\quad \times \exp\left(-\frac{1}{2\sigma_X^2} \sum_{t=1}^T (x_{t,1:N} - \boldsymbol{\mu}_{x,t})' (x_{t,1:N} - \boldsymbol{\mu}_{x,t})\right) \\ &\quad \times p(\boldsymbol{\theta}) , \quad \boldsymbol{\mu}_{x,t} = \phi x_{t-1,1:N} + \mathbf{z}_{t,1:N} \beta_{\mathbf{z}} . \end{aligned}$$

Then, the conditional parameter distributions are conjugate and given as

$$\begin{aligned}
p\left(\sigma_X^2|x_{0:T,1:N}\right) &= \left(2\pi\sigma_X^2\right)^{-NT/2} \times \exp\left(-\frac{1}{2\sigma_X^2} \sum_{t=1}^T \left(x_{t,1:N} - \boldsymbol{\mu}_{x,t}\right)' \left(x_{t,1:N} - \boldsymbol{\mu}_{x,t}\right)\right) \\
&\quad \times \frac{\underline{b}_X^{\underline{a}_X}}{\Gamma(\underline{a}_X)} \left(\sigma_X^2\right)^{-\underline{a}_X-1} \exp\left(-\frac{\underline{b}_X}{\sigma_X^2}\right) \\
&\propto \left(\sigma_X^2\right)^{-(\underline{a}_X+NT/2)-1} \times \exp\left(-\frac{1}{\sigma_X^2} \left(\underline{b}_X + \frac{\sum_{t=1}^T \left(x_{t,1:N} - \boldsymbol{\mu}_{x,t}\right)' \left(x_{t,1:N} - \boldsymbol{\mu}_{x,t}\right)}{2}\right)\right).
\end{aligned}$$

With e.g. $\underline{a}_X = \underline{b}_X = 0.001$, we have

$$\sigma_X^2|x_{0:T,1:N} \sim \mathcal{IG}(\overline{a}_X, \overline{b}_X^*) \text{ , } \overline{a}_X = \underline{a}_X + NT/2 \text{ , } \overline{b}_X = \underline{b}_X + \frac{\sum_{t=1}^T \left(x_{t,1:N} - \boldsymbol{\mu}_{x,t}\right)' \left(x_{t,1:N} - \boldsymbol{\mu}_{x,t}\right)}{2} .$$

For $\boldsymbol{\beta}_Z^* = (\phi, \boldsymbol{\beta}'_Z)'$ with a normal prior $\boldsymbol{\beta}_Z^* \sim \mathcal{N}_{K+1}(\underline{\boldsymbol{\beta}}_Z^*, \underline{\boldsymbol{\Omega}}_Z)$ and $x_{t,1:N} = \phi x_{t-1,1:N} + \mathbf{z}_{t,1:N} \boldsymbol{\beta}_Z + \boldsymbol{\varepsilon}_{x_{t,1:N}} = \mathbf{Z}_{t,1:N} \boldsymbol{\beta}_Z^* + \boldsymbol{\varepsilon}_{x_{t,1:N}}$, we have

$$\begin{aligned}
p\left(\boldsymbol{\beta}_Z^*|x_{0:T,1:N}, \mathbf{Z}_{0:T,1:N}, \sigma_X^2\right) &\propto \exp\left\{-\frac{1}{2} \sum_{t=1}^T \left(x_{t,1:N} - \mathbf{Z}_{t,1:N} \boldsymbol{\beta}_Z^*\right)' \boldsymbol{\Omega}_{\boldsymbol{\varepsilon}_{x_{t,1:N}}}^{-1} \left(x_{t,1:N} - \mathbf{Z}_{t,1:N} \boldsymbol{\beta}_Z^*\right)\right\} \\
&\quad \times \exp\left\{-\frac{1}{2} \left(\boldsymbol{\beta}_Z^* - \underline{\boldsymbol{\beta}}_Z^*\right)' \underline{\boldsymbol{\Omega}}_Z^{-1} \left(\boldsymbol{\beta}_Z^* - \underline{\boldsymbol{\beta}}_Z^*\right)\right\} ,
\end{aligned}$$

which can as a whole expression be written as

$$\exp\left\{-\frac{1}{2} \left[\sum_{t=1}^T \left(x_{t,1:N} - \mathbf{Z}_{t,1:N} \boldsymbol{\beta}_Z^*\right)' \boldsymbol{\Omega}_{\boldsymbol{\varepsilon}_{x_{t,1:N}}}^{-1} \left(x_{t,1:N} - \mathbf{Z}_{t,1:N} \boldsymbol{\beta}_Z^*\right) + \left(\boldsymbol{\beta}_Z^* - \underline{\boldsymbol{\beta}}_Z^*\right)' \underline{\boldsymbol{\Omega}}_Z^{-1} \left(\boldsymbol{\beta}_Z^* - \underline{\boldsymbol{\beta}}_Z^*\right)\right]\right\}$$

Because we have for every $t = 1, \dots, T$

$$\begin{aligned}
\left(x_{t,1:N} - \mathbf{Z}_{t,1:N} \boldsymbol{\beta}_Z^*\right)' \boldsymbol{\Omega}_{\boldsymbol{\varepsilon}_{x_{t,1:N}}}^{-1} \left(x_{t,1:N} - \mathbf{Z}_{t,1:N} \boldsymbol{\beta}_Z^*\right) &= \boldsymbol{\beta}_Z^{*'} \mathbf{Z}_{t,1:N}' \boldsymbol{\Omega}_{\boldsymbol{\varepsilon}_{x_{t,1:N}}}^{-1} \mathbf{Z}_{t,1:N} \boldsymbol{\beta}_Z^* - 2 \boldsymbol{\beta}_Z^{*'} \mathbf{Z}_{t,1:N}' \boldsymbol{\Omega}_{\boldsymbol{\varepsilon}_{x_{t,1:N}}}^{-1} x_{t,1:N} \\
&\quad + x_{t,1:N}' \boldsymbol{\Omega}_{\boldsymbol{\varepsilon}_{x_{t,1:N}}}^{-1} x_{t,1:N}
\end{aligned}$$

and prior $(\beta_{\mathbf{Z}}^* - \underline{\beta}_{\mathbf{Z}}^*)' \underline{\Omega}_{\mathbf{Z}}^{-1} (\beta_{\mathbf{Z}}^* - \underline{\beta}_{\mathbf{Z}}^*) = \beta_{\mathbf{Z}}^{*'} \underline{\Omega}_{\mathbf{Z}}^{-1} \beta_{\mathbf{Z}}^* - 2\beta_{\mathbf{Z}}^{*'} \underline{\Omega}_{\mathbf{Z}}^{-1} \underline{\beta}_{\mathbf{Z}}^* + \underline{\beta}_{\mathbf{Z}}^{*'} \underline{\Omega}_{\mathbf{Z}}^{-1} \underline{\beta}_{\mathbf{Z}}^*$
we obtain

$$\begin{aligned}
p(\beta_{\mathbf{Z}}^* | x_{0:T,1:N}, \mathbf{Z}_{0:T,1:N}, \sigma_X^2) &= \mathcal{N}_{K+1}(\overline{\beta}_{\mathbf{Z}}, \overline{\Omega}_{\mathbf{Z}}) \\
\overline{\Omega}_{\mathbf{Z}} &= \left[\sum_{t=1}^T \mathbf{Z}_{t,1:N}' \Omega_{\varepsilon_{x_{t,1:N}}}^{-1} \mathbf{Z}_{t,1:N} + \underline{\Omega}_{\mathbf{Z}}^{-1} \right]^{-1} \\
\overline{\beta}_{\mathbf{Z}} &= \overline{\Omega}_{\mathbf{Z}} \times \left[\sum_{t=1}^T \mathbf{Z}_{t,1:N}' \Omega_{\varepsilon_{x_{t,1:N}}}^{-1} x_{t,1:N} + \underline{\Omega}_{\mathbf{Z}}^{-1} \beta_{\mathbf{Z}} \right]
\end{aligned}$$

In Detail:

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ALGORITHM: *Conditional BPF*

START **I. Initiliaz** ($t = 0$) :

For $i = 1, \dots, N$:

1. Sample $x_0^i \sim p(x_0)$
3. Set $w_0^i = \frac{1}{N}$
2. Set $x_0^N = x_0^{\mathcal{R}}$ (conditioning)

II. For $t = 1$ **to** T :

For $i = 1, \dots, N$:

1. Draw $a_t^i \sim \mathcal{C}(\{w_t^i\}_{j=1}^N)$
2. Sample $x_t^i \sim q(x_t|x_{t-1}^{a_t^i}, \boldsymbol{\theta}) \underbrace{=} p(x_t|x_{t-1}^{a_t^i}, \boldsymbol{\theta})$
BPF

For $i = N$:

3. Set $x_t^N = x_t^{\mathcal{R}}$ (conditioning)
4. Sample $a_t^N \in \{1, \dots, N\}$ with probability

$$\mathbb{P}(a_t^N = i \propto w_{t-1}^i p(x_t^{\mathcal{R}}|x_{t-1}^i, \boldsymbol{\theta})) \quad , \quad (\text{AS-step})$$

For $i = 1, \dots, N$:

5. Set $\tilde{w}_t^i = p(y_t|x_t^i)$
6. Normalize weights $w_t^i = \frac{\tilde{w}_t^i}{\sum_{t=1}^T \tilde{w}_t^i}$

III. For $t=T$:

1. Draw $b \sim \mathcal{C}(\{w_T^i\}_{i=1}^N)$ and compute $x_{0:T}^b$

END Output $x_{0:T}^{\mathcal{R}} = x_{0:T}^b$

ALGORITHM: *PGAS with conditional BPF*

START **Initiliaz** ($m=1$):

1. Set $x_{0:T}[1]$ and $\boldsymbol{\theta}[1]$ arbitrarily

For $m = 1, \dots, N$:

2. Draw $\boldsymbol{\theta}[m] \sim p(\boldsymbol{\theta}|x_{0:T}[m-1], x_{0:T}^{\mathcal{R}})$
3. Draw $x_{0:T}[m] \sim \kappa_{N, \boldsymbol{\theta}[m]}(x_{0:T}[m-1], x_{0:T}^{\mathcal{R}})$

END Output $\boldsymbol{\theta}[1:m]$ and $x_{0:T}[1:m]$