

# Bayesian Statistics : Assignment 2

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November 14, 2023

This assignment is part of a project to validate the course. Failure in returning one or more assignments implies failure of the class.

Instructions :

- 1). Students have to work in **groups of 3 students** (same groups as for the Assignment 1). Only one document per group has to be returned.
- 2). The due date is **December 5, 2023** (assignment returned after this day will not be considered and the whole project will not be validated).
- 3). Each group has to upload the document containing the solution to the assignment on <https://app.compilatio.net/v5/document-submission/J23-E55-778> in the pdf format (no other format will be considered).
- 4). The pdf has to be named in the following way : `FAMILYNAME1_FAMILYNAME2_FAMILYNAME3_Ass_2`.
- 5). The solution has to be typed (not handwritten).
- 6). For each question, please provide all the computational details.
- 7). Question 2 has to be developed with the statistical software that you prefer among Matlab, R or Python.

We consider the same model as in the Assignment 1 : for  $t = 1, \dots, T$ ,

$$y_t = u_t' \phi + x_t' \beta + \varepsilon_t, \quad \varepsilon_t \sim^{i.i.d.} \mathcal{N}(0, \sigma^2), \quad (0.1)$$

where  $x_t$  is a vector of regressors of dimension  $k$  and  $u_t$  is a vector of regressors of dimension  $\ell$ . For simplicity, the variance of each regressor is normalized to 1.

The prior for  $\theta := (\sigma^2, \phi', \beta') \in \mathbb{R}^{1+\ell+k}$  is specified as in the Assignment 1 :  $\pi(\theta) = \pi(\sigma^2) \times \pi(\phi) \times \pi(\beta|\sigma^2, \gamma^2, q)$  with

$$\begin{aligned} \pi(\sigma^2) &= \frac{1}{\sigma^2} \\ \pi(\phi) &= \text{flat} \\ \pi(\beta_i|\sigma^2, \gamma^2, q) &\sim^{i.i.d.} \begin{cases} \mathcal{N}(0, \sigma^2 \gamma^2) & \text{with probability } q \\ 0 & \text{with probability } 1 - q \end{cases} \quad \text{for } i = 1, \dots, k \end{aligned}$$

and  $\gamma^2, q$  are hyperparameters taking respectively nonnegative values and values in  $[0, 1]$ .

In addition, we specify the prior for the hyperparameters  $q$  and  $\gamma^2$  in the following way. First, we specify the mapping  $(\gamma^2, q) \mapsto R^2(\gamma^2, q) := \frac{qk\gamma^2\bar{v}_x}{qk\gamma^2\bar{v}_x+1}$ , where  $\bar{v}_x := \frac{1}{k} \sum_{j=1}^k \widehat{Var}(x_{t,j})$  is the average sample variance of the predictors. Then, the prior for  $q$  and  $R^2$  is

$$\begin{aligned} \pi(q) &\sim \text{Beta}(a, b) \\ \pi(R^2) &\sim \text{Beta}(A, B). \end{aligned}$$

Introduce a latent variable  $z_i$  that takes the value 1 if the corresponding covariate  $x_i$  is included in the model (that is, is active) and its coefficient  $\beta_i$  is non-zero, and it takes the value 0 otherwise. The random variable  $z_i$  is *Bernoulli*( $q$ ), where  $q$  denotes the probability of success. Let  $\tilde{\beta}$  be the vector of the non-zero coefficients (i.e. those associated with  $z_i = 1$ ) and  $\tilde{X}$  be the corresponding regressors. We can recover  $\beta$  from  $z := (z_1, \dots, z_k)'$  and  $\tilde{\beta}$  in the following way :  $\beta_i = z_i \tilde{\beta}_{\bar{z}_i}$ , where  $\bar{z}_i = \sum_{\ell \leq i} z_\ell$ .

**Question 1.** Show that  $\pi(\beta|\sigma^2, R^2, q)$  (which is  $\prod_{i=1}^k [q\mathcal{N}(0, \sigma^2 \gamma^2) + (1 - q)\delta_0(\beta_i)]$ , where  $\delta_0(\cdot)$  denotes a Dirac in 0) is equal to

$$\prod_{i=1}^k q^{z_i} (1 - q)^{1-z_i} \prod_{i=1}^k [\mathcal{N}(0, \sigma^2 \gamma^2)]^{z_i} [\delta_0(\beta_i)]^{1-z_i}.$$

**Question 2.** Here you are asked to simulate from the joint posterior distribution. To do so, you have to implement the following Gibbs sampling with blocks (where we denote  $z = (z_1, \dots, z_k)'$ ,  $Y = (y_1, \dots, y_T)'$ ,  $U = [u_1, \dots, u_T]'$  ( $U$  is  $T \times \ell$  matrix) and  $X = [x_1, \dots, x_T]'$  ( $X$  is  $T \times k$  matrix)). The blocks are  $(R^2, q)$ ,  $\phi$  and  $(\beta, \sigma^2)$  :

(I). Draw from the conditional posterior of  $(R^2, q)$  given by :

$$\pi(R^2, q|Y, U, X, \theta, z) \propto \left[ e^{-\frac{1}{2\sigma^2} \frac{k\bar{v}_x q(1-R^2)}{R^2} \beta' \text{diag}(z)\beta} \right] \\ \times q^{s(z) + \frac{s(z)}{2} + a - 1} (1 - q)^{k - s(z) + b - 1} (R^2)^{A - 1 - s(z)/2} (1 - R^2)^{s(z)/2 + B - 1},$$

where  $s(z) := \sum_{i=1}^k z_i$ . To draw from this distributions you have to sample from a discrete approximation of this distribution. More specifically, let us discretize the support of  $R^2$  and  $q$  by interlacing two grids defined over the unit interval as follows : the grid for  $q$  is  $[0 : .001 : .1 \ .11 : .01 : .9 \ .901 : .001 : 1]$  and for  $R^2$  is  $[0 : .001 : .1 \ .11 : .01 : .9 \ .901 : .001 : 1]$  (each grid has increments of 0.01, and finer increments of 0.001 near the boundaries).

(II). Draw from the conditional posterior of  $\phi$  given by :

$$\pi(\phi|Y, U, X, z, \beta, R^2, q, \sigma^2) = \pi(\phi|Y, U, X, \beta, \sigma^2) \\ = \mathcal{N}((U'U)^{-1}U'(Y - X\beta), \sigma^2(U'U)^{-1}).$$

(III). To sample from the conditional posterior of  $z$  observe that

$$\pi(z|Y, U, X, \phi, R^2, q) \propto q^{s(z)}(1-q)^{k-s(z)} \left( \frac{1}{\gamma^2} \right)^{\frac{s(z)}{2}} |\widetilde{W}|^{-1/2} \left[ \frac{\widetilde{Y}'\widetilde{Y} - \widetilde{\beta}'\widetilde{W}\widetilde{\beta}}{2} \right]^{-T/2} \Gamma\left(\frac{T}{2}\right),$$

where  $\widetilde{\beta} := \widetilde{W}^{-1}\widetilde{X}'\widetilde{Y}$ ,  $\widetilde{W} := (\widetilde{X}'\widetilde{X} + I_{s(z)}/\gamma^2)$ , and  $\widetilde{Y} := Y - U\phi$ . Therefore, to sample from the posterior of  $z|Y, U, X, \phi, R^2, q$  we can draw iteratively from the distribution of  $z_i|Y, U, X, \phi, R^2, q, z_{-i}$ ,  $i = 1, \dots, k$ , using a Gibbs sampling.

(IV). Draw from the conditional posterior of  $\sigma^2$  given by :

$$\pi(\sigma^2|Y, U, X, \phi, R^2, q, z) = I\Gamma\left(\frac{T}{2}, \frac{\tilde{Y}'\tilde{Y} - \tilde{\beta}'(\tilde{X}'\tilde{X} + I_{s(z)}/\gamma^2)\tilde{\beta}}{2}\right).$$

(V.) Let  $\tilde{\beta}$  be the subvector of  $\beta$  containing only the non-zero coefficients (notice that we know the identity of the non-zero coefficients thanks to the vector  $z$ ). Draw from the conditional posterior of  $\tilde{\beta}$  given by :

$$\begin{aligned}\pi(\tilde{\beta}|Y, U, X, \phi, R^2, q, \sigma^2, z) \\ = \mathcal{N}\left(\left(I_{s(z)}/\gamma^2 + \tilde{X}'\tilde{X}\right)^{-1} \tilde{X}'(Y - U\phi), \sigma^2 \left(I_{s(z)}/\gamma^2 + \tilde{X}'\tilde{X}\right)^{-1}\right).\end{aligned}$$

where  $\tilde{X}$  are the regressors corresponding to  $\tilde{\beta}$  and the other  $\beta_i$ 's are equal to 0.

You will repeat these steps for 110 000 iterations and you will eliminate the first 10000 iterations (burn-in period).

The data are generated according to model (0.1) with  $\ell = 0$ ,  $k = 100$ , and  $T = 200$ . The predictors  $x_t$  are drawn from a zero-mean Normal distribution with a Toeplitz correlation matrix with  $\text{corr}(x_{t,i}, x_{t,j}) = \rho^{|i-j|}$ , where we set  $\rho = 0.75$ .

We set  $k - s$  elements of  $\beta$  (chosen at random) at zero and draw the remaining  $s$  from independent  $\mathcal{N}(0, 1)$  distributions. The parameter  $s$  is set equal to 5, 10 and 100.

The error term  $\varepsilon_t$  is *i.i.d.*  $\mathcal{N}(0, \sigma^2)$  with  $\sigma^2 = (1/R_y - 1) \sum_{t=1}^T (\beta' x_t)^2 / T$  and where  $R_y = 2\%, 25\%, 50\%$  denotes the ratio between explained and total variance.

The values for the hyperparameters of the priors are :  $a = b = A = B = 1$ .

You have to :

1. Simulate 100 data sets, standardize the data.
2. For each value of  $s \in \{5, 10, 100\}$  and of  $R_y \in \{2\%, 25\%, 50\%\}$  :
  - (a) For each data set, compute the posterior median of  $q$  and plot an histogram (or kernel) approximation of the distribution of the posterior median of  $q$  across the 100 data sets.
  - (b) For one data set, plot the marginal posterior distribution of  $q$  (by using the histogram).

- (c) Comment and compare the results for the different values of  $s$  and  $R_y$  by recalling that  $q$  is the posterior of the probability of inclusion.