# Bayesian prediction and model selection Bayesian Computation 2 / 14

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## Objectives

In this first course, we will continue to explain how to solve the three basic questions of statistics in the Bayesian paradigm

- Prediction
- Model selection

We will also start to see that computing the posterior is tricky.

Bayesian prediction

Model selection

Conclusions

# Bayesian prediction

#### The importance of prediction

#### Two very important role for statistics:

- Predict the future.
- Reveal the unseen.

#### How? Learn the correlation between:

- Easy to access predictor variables  $X_1 \dots X_d$ .
- Hard to access variable of interest Y.

Alice wants to sell ice-cream at Ouchy during the summer. She needs to accurately forecast how much ice-cream she will sell during the day.

Assume that over the last week, she has sold the following quantities (in L):

She was never sold out (yet).

How can we try to predict the quantity of ice cream she might sell tomorrow  $Y_{n+1}$ ?

#### Two key ideas:

• The answer needs to be encoded into a probability distribution:

$$Y \sim F_Y = ??$$

 We need to derive this probability distribution through the application of Bayes' rule.

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#### Ideas?

Solution: augment the model with another variable !

#### Normal model:

Conditional model:

$$f(Y_1 \dots Y_7 | \boldsymbol{\theta})$$

• Prior model  $f(\theta)$ 

#### Augmented model:

• Add a conditional model for  $Y_{n+1}$ .

#### For example:

• Assume that all quantities are IID Gaussian:

$$Y_i \stackrel{IId}{\sim} \mathcal{N}\left(\mu, \sigma^2\right)$$

includes the observed  $Y_1 \dots Y_7$  and the unobserved  $Y_{n+1}$ .

- Two unknowns:  $\mu, \sigma^2$ .
  - ullet For technical reasons, we will work instead with the precision  $eta=\sigma^{-2}$  .
  - For technical reasons, we consider the following prior:

$$\beta \sim \Gamma (a = 1, b = 100)$$
  
 $\mu | \beta \sim \mathcal{N} (50, \beta^{-1} * 1)$ 

(A Gamma-normal hierarchical distribution)



Prior density:

$$f(\beta, \mu) \propto \beta^{1/2} \exp\left(-100 \ \beta - \frac{\beta}{2} (\mu - 50)^2\right)$$
$$\propto \beta^{s-1/2} \exp\left(-b \ \beta - \frac{\beta}{2} (\mu - 50)^2\right)$$

Likelihood:

$$f(y_1 \dots y_7 | \mu, \beta) \propto \prod_{i=1}^{7} \beta^{1/2} \exp\left(-\frac{\beta}{2} (y_i - \mu)^2\right)$$

$$\propto \beta^{7/2} \exp\left(-\frac{\beta}{2} \sum_{i=1}^{7} (y_i - \mu)^2\right)$$

$$\propto \beta^{7/2} \exp\left(-\frac{\beta}{2} \left\{7 (\bar{y} - \mu)^2 + \sum_{i=1}^{7} (y_i - \bar{y})^2\right\}\right)$$

$$\propto \beta^{7/2} \exp\left(-\frac{7 (\bar{y} - \mu)^2}{2} \beta - \frac{\beta}{2} \sum_{i=1}^{7} (y_i - \bar{y})^2\right)$$

Posterior:

$$f(\beta,\mu|d) \propto eta^{8/2} \exp\left(-\left\{100 + rac{\sum_{i=1}^{7} (y_i - \bar{y})^2}{2}
ight\} \ eta - rac{eta}{2} \left\{(\mu - 50)^2 + 7(\bar{y})^2\right\}$$

We recognize another Gamma-Normal distribution (!!?!). Defining:

$$\hat{\mu} = \frac{50 + 7\bar{y}}{1 + 7}$$

•  $\beta | d$  is marginally Gamma:

$$\beta | d \sim \Gamma \left( a = 1 + \frac{n}{2}, b = 100 + \frac{\sum_{i=1}^{7} (y_i - \hat{\mu})^2}{2} + \frac{1}{2} (\hat{\mu} - 50)^2 \right)$$

• While  $\mu | \beta, d$  is Gaussian:

$$\mu|eta,d\sim\mathcal{N}\left(\hat{\mu},(8eta)^{-1}
ight)$$

Now, we can compute the posterior of the new observation  $Y_{n+1}$ . Conditional on  $\mu, \beta, d$ , it is Gaussian:

$$Y_{n+1}|\mu,\beta,d\sim\mathcal{N}\left(\mu,\beta^{-1}\right)$$

Marginalizing out  $\mu$ :

$$Y_{n+1}|\beta, d \sim \mathcal{N}\left(\hat{\mu}, \beta^{-1} + (8\beta)^{-1}\right)$$
  
  $\sim \mathcal{N}\left(\hat{\mu}, \frac{9}{8}\beta^{-1}\right)$ 

Marginalizing out  $\beta$  is harder: it gives a student distribution:

$$Y_{n+1}|d \sim T\left(\mathsf{ddof} = n+2, \mathbb{E} = \hat{\mu}, \sigma^2 = \frac{100 + rac{\sum_{i=1}^{7} (y_i - \hat{\mu})^2}{2} + rac{1}{2} (\hat{\mu} - 50)^2}{8/9 (n+2)}\right)$$

(Don't sue me if I got it wrong)

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## Take home messages

#### Keys:

- Bayesian inference can involve quite a bit of work.
- Importantly, here, the normalization constant did not matter
- Prediction involves the addition of more variables in the model.
- We obtain a posterior over the variable to be predicted:

$$f(Y_{n+1}|d)$$

We can then construct normal Bayesian point estimates:

- mean, median, MAP
- or interval estimates:
  - Credible intervals

## Take home messages

- Magical coincidence: we recovered a posterior inside the same family as the prior.
  - This is called a conjugate family associated to a conditional model.
  - This is extremely rare.
     I chose this feature on purpose to make my life simple.
  - We'll discuss this more next week.
- Prior is partially interpretable since it plays a role comparable to the data ("Pseudo-data" interpretation):
  - For example, the prior mean ( $\mathbb{E}(\mu) = 50$ ) and the empirical mean  $\bar{y}$  play the same role in the final formula.
- This is again a property of conjugate families that we will talk about next week.

# Model selection

## Choosing the right model

In many situations, a number of qualitatively different models could explain the data. The job of the statistician then consists in determining which one is the best.

- Dependance or independance of two measured variables
- Is situation A different or identical to situation B.
- Which predictors are useful for anticipating the value of Y.

This problem of model selection is probably the hardest problem of statistics.

## Choosing the right model

#### Classical approaches:

- Neyman-Pearson:
  - Heavily biased towards scientific inference.
  - Two alternatives:  $H_0$  and  $H_1$ .
  - Asymetric: reject or conserve  $H_0$ .
- Best validation Performance:
  - For each model, find the best fit on a training data set.
  - Compute the "performance" of the best fit on a new data set.
  - Optional: correct for the number of parameters (AIC, BIC, etc)
  - Choose the model with the best validation performance.

Let's return to Alice and here ice-creams.

Assume she wants to know whether doing something different (e.g. changing the price of the ice-cream, or the recipe) modifies the amount of money she makes in a day.

First, we need to collect data. Let's assume that:

- Each day, **she chooses randomly** whether she will in condition 1 or 2.
- She has collected n observations from each case:  $X_i$  and  $Y_i$ .
- Does the intervention matter?

We want to know whether intervention matters or not:

- Once again, the answer needs to be a probability distribution.
- That is derived through applying Bayes' rule.

Ideas?

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#### Ideas?

Once again, the solution consists in augmenting the model with more variables.

A key idea of Bayesian model choice: sampling from the prior should generate realistic datasets (generative approach to priors).

- NB: sampling from the prior means:
  - ullet Choosing a random  $oldsymbol{ heta}$  from the prior.
  - Choosing a random dataset from the conditional distribution  $f(\mathcal{D} = d|\theta)$ .

Here, sampling from the prior should generate:

- Some datasets for which the intervention does nothing.
- Some datasets for which the intervention does something.

An elegant solution for this: the addition of a "Flag" variable: a discrete variable  $F \in \{0,1\}$ .

- ullet F=1 corresponds to the active model: intervention does something.
- $\bullet$  F = 0 corresponds to the inactive model: intervention does nothing.

#### Prior distribution:

- Sample  $F \sim B(p)$ .
- Conditional on F=0

$$\mu_{X} = \mu_{Y} \sim \mathcal{N}\left(300, (50)^{2}\right)$$

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ullet Conditional on F=1

$$\mu_X, \mu_Y \stackrel{\textit{IID}}{\sim} \mathcal{N}\left(300, (50)^2\right)$$

• Conditional on  $\mu_X, \mu_Y$ :

$$X_i \stackrel{IID}{\sim} \mathcal{N}(\mu_X, 1)$$
  
 $Y_i \stackrel{IID}{\sim} \mathcal{N}(\mu_Y, 1)$ 



We have defined the model. Now comes the painful part where we apply Bayes' rule.

We know how to perform the inference conditional on the value of F: it's just simple inference for a Gaussian model. We thus have:

$$f(\mu_X, \mu_Y | F = 0, d)$$
  
$$f(\mu_X, \mu_Y | F = 1, d)$$

In order to finish, we only need to characterize the marginal distribution: f(F|d).

Applying Bayes' rule to the pair F, d yields:

$$f(F|d) \propto f(F) f(d|F)$$

where f(d|F) is the distribution of d when I marginalize out  $\mu_X, \mu_Y$ .

$$f(d|F) = \int f(d \& \mu_X, \mu_Y|F) d\mu_X d\mu_Y$$

This term is precisely the normalizing constant that we obtain when we apply Bayes' rule conditional on the value of F to compute the posterior of  $\mu_X, \mu_Y$ :

$$f(\mu_X, \mu_Y|F, d) = \frac{f(\mu_X, \mu_Y|F) f(d|\mu_X, \mu_Y, F)}{f(d|F)}$$

This is why and where the normalizing constant matters in Bayesian inference: in order to perform model selection !!!

Thus, the overall logic of Bayesian model selection is the following:

• First, perform inference in each model, i.e. conditional on F=0 or F=1.

$$f(\mu_X, \mu_Y | F, d) = \dots$$

Critically, we need to evalute the normalization constant f(d|F)!!

Then, perform inference for the "Flag" variable:

$$f(F|d) \propto f(F) f(d|F)$$

Here, the normalization constant does not matter.

Thankfully, I've chosen a simple model for which the calculation of f(d|F)is simpler.

For F = 0

$$\mu = \mu_X = \mu_Y \sim \mathcal{N}\left(300, (50)^2\right)$$
 $X, Y \stackrel{\textit{IID}}{\sim} \mathcal{N}\left(\mu, 1\right)$ 

Thus, the marginal distribution of X, Y is a Gaussian with parameters:

$$\mathbb{E}(X_i) = \mathbb{E}(Y_i) = 300$$
 $Var(X_i) = Var(Y_i) = 1 + 2500$ 
 $Cov(X_i, X_j) = Cov(X_i, Y_j) = 2500$ 

Thus:

$$f(d|F=0) = \frac{(2\pi)^{2n/2}}{|\mathsf{Cov}|^{1/2}} \exp\left(-\frac{1}{2}([X,Y]-300)(\mathsf{Cov})^{-1}([X,Y]-300)\right)$$

The same logic applies for F=1 except the covariance matrix is slightly different:

$$Cov(X_i, Y_j) = 0$$

Once again:

$$f(d|F=0) = \frac{(2\pi)^{2n/2}}{|\mathsf{Cov}|^{1/2}} \exp\left(-\frac{1}{2}\left([X,Y] - 300\right)(\mathsf{Cov})^{-1}\left([X,Y] - 300\right)\right)$$

## Take home messages

#### Keys:

- Once again, notice how much work we had to do on the posterior.
- Here, the normalization constant of the intermediate variables  $\mu_X, \mu_Y$  played a key role.
- Once again, we expanded the model in order to answer the question of interest.
- Here, the normalization constant was accessible directly.
   This is very rare and occured because I chose to make my life simple.

# Conclusions

## The story so far

We now know how to answer key statistical questions of stastical inference:

- Estimation:
  - Point estimates:
    - compress the posterior into a scalar: MAP, Mean, Loss function.
  - Intervals:
    - credible intervals (loss function??).
- 2 Prediction:
  - Augment the model.
  - Obtain a posterior on desired variables.
- Model selection:
  - Augment the model.
  - Compute the normalization constants in the intermediate posterior.
  - Obtain a posterior on "Flag" variables.



## The story so far

In all examples so far, I've made my life simple: the posterior was always explicit.

This is **rare**. We'll highlight next week the necessary conditions and explain why this almost never occurs in practice.

#### Bayesian prediction

Given data  $\mathcal{D}=d$  and unobserved variable(s) of interest  $Y_{prediction}$ :

• Choose a joint model of data and variable(s) of interest:

$$f(\mathcal{D} \& Y_{prediction}|\theta)$$

- 2 Choose a prior:  $f(\theta)$ .
- **3** Compute the posterior distribution of the random variables  $\theta$ ,  $Y_{prediction}$ :

$$f(\theta, Y_{prediction}|\mathcal{D} = d)$$

• Marginalize out  $\theta$ :

$$f(Y_{prediction}|\mathcal{D}=d) = \int d\theta f(\theta, Y_{prediction}|\mathcal{D}=d)$$

 $\odot$  Return Bayesian point or interval estimates of  $Y_{prediction}$ 

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#### Bayesian model selection

Given data  $\mathcal{D} = d$  and multiple competing models:

$$f_{M_1}(\mathcal{D}|\boldsymbol{\theta}) \quad f_{M_2}(\mathcal{D}|\boldsymbol{\theta}) \quad f_{M_3}(\mathcal{D}|\boldsymbol{\theta}) \dots$$

(NB: very often, the number or interpretation of the parameters might change drastically from one model to the next):

- **1** Augment the model with a Flag variable I such that I = i means that model  $M_i$  is active.
- **2** Choose the prior:  $\mathbb{P}(I = i)$ .
- 3 For every model, compute the normalization probability of the model:

$$f_{M_i}(\mathcal{D}=d)=\int d\theta f_{M_i}(\mathcal{D}=d|\theta)$$

The posterior over I is:

$$\mathbb{P}(I=i) \propto \mathbb{P}(I=i) f_{M_i}(\mathcal{D}=d)$$

#### Methods for prior choice: pseudo-data

The following principles can guide our choice of prior:

- Vague priors:
  - One weak principle for prior choice is to use priors with very large width.
  - This encodes the common situation of having not much prior information.
- Pseudo-data interpretation:
  - Priors that are conjugate to a conditional model can be interpreted in terms of adding virtual observations. The value of these virtual observations can be deduced from the parameters of the prior.
  - e.g. Gaussian prior, Beta prior, Student prior.
- Generative priors:
  - Sampling datasets from the prior distribution  $f(\mathcal{D})$  should generate (somewhat) credible artificial datasets.
  - This principle rarely constraints the shape but can be helpful in finding the appropriate scale of the prior.