# Introduction to Bayesian inference

Bayesian Computation 1 / 14

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# Objectives

In this first course, we will:

- Explain how the course is structured
- Introduce Bayesian statistics
- Recap Bayes' formula
- Introduce the Bayesian approach to statistics
  - Bayesian point estimation.
  - Bayesian interval estimation.

- Thesaurus
- 2 Introduction
- Bayes' formula
- Bayesian statistics
- Examples
- Conclusion

# Thesaurus

### Statistics

The object of statistics: extract rational conclusions from a dataset.

Classical approach to statistics: frequentist inference.

**Key:** probability statements are made *a-priori* of the data.

Alternatively, Bayesian inference. Possible arguments:

- Conditional statements are easy to interpret.
- Honest account of bias and modeling.
- Frequentist optimality.
- Subjective point of view on probability (probability as belief).
- Probability as rationality (philosophy / neuroscience).

Worst case: it's an important part of an "homme du monde" education in statistics.

#### **Statistics**

Like every field of statistics, Bayesian inference is supported on three pillars:

- Probability theory: my analyses are mathematically sound.
   (NB: not unambiguous: requires a definition of soundness).
- Computation: my computer can perform my analysis before the heatdeath of the universe.
- Modeling: my mathematical model of my data is a close match to the truth.

This course will focus on computation, but we will still discuss Theory and Modeling issues when appropriate.

# Course objectives

The goal of the course is to give you an overview of Bayesian inference:

- What it is and it isn't.
- 2 When to use it.
- The strengths and the weaknesses.
- The computational methods which make it possible.

You will not be experts but you should become fluent.

### Structure of the course

Every week, the lecture will focus mostly on one topic / method.

The exercise session is focused on programming the methods of the lecture.

As we progress, we will shift towards spending the exercise sessions on your projects.

Going to the exercises sessions is extremely strongly recommended.

Feel free to program in any language.

# Course requirements

The requirements for the course should be low:

- Basics of probability theory and of statistics.
- Basics of analysis (no measure theory).
- Familiarity with Markov Chains (I'll review them briefly).

If there is any point which flies over your head, please let me know.

#### Exam

The course will be graded on an **oral presentation** of a project that you do during the semester.

You will implement multiple Bayesian methods on a **single dataset** and present the results.

You should work on the project during the exercise sessions and seek feedback often on your work.

### Class interaction

### The #ThreeQuestionChallenge:

- Throughout the 14 weeks, I challenge every single one of you to intervene three times.
- Interventions can be big or small.
- There are no such thing as stupid questions.

I also ask many questions during the class:

- Most of these are easy!
- Don't worry, there are no wrong answers.

Questions?

# Introduction

#### What is statistics

Statistics is the field that tries to answer the following question:

How can I draw logical / rational conclusions from random data?

It is thus a non-trivial extension of logic.

Both logic and statistics obey the same structure:

- Start from postulates / premises:
  - Logic example: definition of addition.
  - Stats example: My data is Gaussian IID
- Based on these postulates, draw a conclusion:
  - Logic: Thus, addition is commutative. 1
  - **2** Stats: Thus,  $\bar{X}$ ,  $S^2$  are great estimators of  $(\mu, \sigma^2)$ .

<sup>&</sup>lt;sup>1</sup>Cf: Poincaré: La Science et l'Hypothèse.

# The Frequentist approach

The most well-known statistical framework is the frequentist approach.

Its logic is a little tricky to follow.

- Premise: my data is distributed according to *Model*.
- Furthermore, in *Model*, the (random) output of *calculation* has good properties.
  - Eg: estimator  $\hat{m{ heta}}$  is almost equal to the truth  $m{ heta}_0$ .
- Conclusion: calculation performed on my data should also have those properties.

## Critically, the good properties are a-priori of the data.

Indeed, the data is the only random quantity. Once the data is known, we cannot make probability statements anymore!

# The Bayesian approach

The logic of Bayesian inference is quite different.

We assume that the parameters of the model are random.

We can then apply Bayes' formula:

Conditional on the data d,  $\theta$  is still random with distribution:

$$f(\theta|d) = f(\theta) \frac{f(d|\theta)}{f(d)}$$

Any statement from Bayesian inference is **conditional** / **a posteriori on** the data.

### The real reason

This difference is a very small nitpick.

In practice, both approaches lead to quite similar conclusions / estimations (and they should).

Some strengths of Bayesian inference:

- Interpretability.
- Separation of modeling and inference.
- Frequentist optimality results (Cramer-Rao bound).

We'll return to these much later.

# A little bit of history

Like the name indicates, Bayes' formula was first proposed by Thomas Bayes (1750's).

Laplace (1800's) was another major influence in the field and can be considered the true father of Bayesian statistics (and all of statistics).

Bayesian inference was originally the dominating paradigm in statistics. This changed around 1850-1900 with the development of the formal axiomatization of probability theory (Kolmogorov) and the development of statistics as an independent field (Pearson K., Fisher, Neyman, Pearson E.)

There have been historically a million attempts to assert that one statistical philosophy is superior to the other. These are extremely silly and they have exactly zero value.

Bayes' formula

# The historical problem

- We push a black ball onto a billiard table.
- We then push n white balls onto the table and record the number X of balls which stop before the black ball.

Assuming the final position of each ball is a **uniform RV**, how can we estimate the position of the black ball  $\theta$  using X?

# The historical problem

X is a binomial distribution:

$$X \sim \mathcal{B}\left(\frac{\theta}{L}, n\right)$$

A first **frequentist** answer to Bayes' problem consists in using the estimator:

$$\theta \approx \hat{\theta} = L \frac{X}{n}$$

# The historical problem

This frequentist solution ignores the additional information contained in:

"We push a black ball onto a billiard table."

Thus,  $\theta$  is itself random:

$$\theta \sim \mathcal{U}([0,L])$$

We thus have a two-step process:

- Generate  $\theta$  randomly.
- Generate X randomly (through generating the random positions of the white balls).

We want to estimate the result of the first step through observing the result of the second.

# Bayes' formula

The solution to the problem is given by Bayes' formula:

$$f(\theta|x) = f(\theta) \frac{f(x|\theta)}{f(x)}$$

Let us give the correct interpretation of the various elements of this formula:

- $\theta \to f(\theta)$  is the **prior** of the unobserved quantity  $\theta$ .
  - = what we knew before the data x was observed.
- ②  $\theta \to f(x|\theta)$  is the likelihood = a compatibility measure of  $\theta$  (variable) with the data x (fixed).
- **1** If f(x) is a normalization constant. We need to compute it:  $f(x) = \int f(\theta) f(x|\theta)$ . We will refer to:

$$\tilde{f}(\theta|x) = f(\theta) f(x|\theta)$$

as the unnormalized posterior.

# Bayes' formula

Critically, Bayes' formula says that, given our observation of x,  $\theta$  is still random and has density  $f(\theta|x)$ .

Contrast this with the usual frequentist assumption that our parameters  $\theta$  are fixed properties of the universe.

As you might have noticed, Bayes' formula isn't statistics. It's an undeniable (basic) law of probability, but it is **not an undeniable law of statistics.** 

# Bayesian statistics

### Model based statistics

What ingredients do we need to do Bayesian statistics:

- A parametric model, i.e. a conditional distribution:

$$f(\mathcal{D}=d|\boldsymbol{\theta})$$

where  $\theta$  represents parameters of the model. If  $\theta$  is finite dimensional, then the model is **parametric**.

**3** A prior distribution on the parameters  $f(\theta)$ .

Note how similar this is to other **likelihood methods** (e.g. the Maximum Likelihood Estimator): we have only added the prior distribution on top of the conditional model.

### The MLE

Let us briefly recall how the MLE approach works.

The likelihood function:

$$\theta \rightarrow f(d|\theta)$$

gives a measure of the **compatibility** of each value of  $oldsymbol{ heta}$  with the data.

ullet The MLE is the value of  $oldsymbol{ heta}$  with maximum compatibility:

$$\hat{\theta}_{ML} = \max_{oldsymbol{ heta}} \left[ f\left(d|oldsymbol{ heta}
ight) 
ight]$$

It has various great properties.

The goal of the MLE approach is to answer the most complicated question of classic statistics:

#### How do I choose my estimator?

A simple variant of the MLE consists in adding a regularization term, eg:

$$\hat{\boldsymbol{\theta}}_{ML}' = \max_{\boldsymbol{\theta}} \left[ f(\boldsymbol{d}|\boldsymbol{\theta}) - \lambda \|\boldsymbol{\theta}\|_{2}^{2} \right]_{\text{absolute of the problem}}$$

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# Bayesian statistics

With the addition of the prior  $f(\theta)$ , we have a joint distribution over  $(\theta, \mathcal{D})$ .

This joint distribution is defined as a two-step process:

- ullet First, pick the value of the parameter RV  $oldsymbol{ heta}$
- ullet Then pick the value of the data, conditional on  $oldsymbol{ heta}$
- Finally, observe one realization of the data d

We can then compute the posterior distribution of  $\theta$  given the data we got, d, using Bayes' formula:

$$f(\theta|d) = f(\theta) \frac{f(d|\theta)}{f(d)}$$

**Under our modelling assumptions**, Bayes' formula is undeniably what we should do.

Of course, one could object to our modeling.

# Bayesian statistics

Critically, under our modelling assumptions,  $\theta$  is still a random variable! It has density  $f(\theta|d)$ .

Once more, contrast this to the usual frequentist assumption that  $m{ heta}$  is fixed for all eternity.

The posterior holds all of the information we have about  $oldsymbol{ heta}$ 

The key assumption of Bayesian statistics is that

Any question that we ask should be answered on the basis of the posterior distribution of  $\theta$ .

# The three questions of statistics

Let's see how we can answer the three questions of statistics using the posterior:

- Estimation:
  - Point estimation.
  - Interval estimation.
- 4 Hypothesis testing.
- Prediction.

#### What is the difference?

- Estimation / Testing
  - Trying to understand part of the distribution of X
- Prediction:
  - Understand all / as much as we can of the distribution of X

#### Estimation

Classical estimation problem: guess the value of  $\theta$ .

In Bayesian statistics, this requires:

- Summarizing the posterior distribution  $f(\theta|\mathcal{D}=d)$ .
- Into a single scalar that represents the whole distribution.

There are many possibilities for Bayesian estimation.

#### Estimation

#### Most important ones:

- ullet Posterior mean:  $\mathbb{E}\left(oldsymbol{ heta}|\mathcal{D}=d
  ight)$
- Posterior median.
- Posterior mode (or Maximum A Posteriori value; MAP).

### A general recipe:

- **1** Choose a loss function  $L(\theta, \hat{\theta})$ .
- Minimize the expected loss under the posterior:

$$\hat{oldsymbol{ heta}}_{L} = \mathsf{argmin}_{\hat{oldsymbol{ heta}}} \left[ \mathbb{E} \left( L \left( oldsymbol{ heta}, \hat{oldsymbol{ heta}} 
ight) | \mathcal{D} = d 
ight) 
ight]$$



#### Interval estimation

Point estimates are a **very poor way** to solve the estimation question because there is no notion of precision.

Frequentist answer: Confidence Interval

Bayesian answer: Credible Interval:

$$I$$
 such that  $\mathbb{P}\left(oldsymbol{ heta} \in I | \mathcal{D} = d
ight) = 1 - lpha$ 

There are infinitely many possible choices. Usually, we either:

- Take the interval with minimal length.
- Or one centered around an interesting estimator

#### Interval estimation

For example, if the posterior is a p-dimensional Gaussian or almost Gaussian:

$$oldsymbol{ heta}|\mathcal{D}=d\sim\mathcal{N}\left(oldsymbol{\mu}_{ extit{post}}, \Sigma_{ extit{post}}
ight)$$

then, we can construct a confidence zone that is egg-shaped:

$$\left( heta - \mu_{post}
ight)^{\mathcal{T}} \left[\Sigma_{post}
ight]^{-1} \left( heta - \mu_{post}
ight) \leq q_{lpha}\left(\chi_{p}^{2}
ight)$$

# Examples

Medecine: is Alice sick?

As a simple example, consider the following problem:

- Alice is at the doctor showing possible signs of the Cold.
- The doctor does a simple blood test which comes back positive.
- Is Alice sick?

#### Bayesian answer:

- Prior probability of a patient having a cold: 0.10.
- Probability of false positive (confidence): 0.20.
- Probability of true positive (power): 1.
- Posterior probability of Alice being sick?

Medecine: is Alice sick?

#### Key lessons:

- Confidence statements do not translate to credible statements:
  - Confidence and Power of test can't be interpreted as posterior probabilities.
- Posterior probability gives us naturally a notion of certainty.
- Posterior probability can be recombined with new information.

#### Limits:

- Where did prior probability come from?
  - It seems legitimate to have some prior information.
    - In summer: no Cold epidemic.  $p \approx 0$ .
    - In winter: Cold epidemic.  $p \gg 0$ .
  - But how do we get a value?
    - Fraction of people that are sick at a given point in time?
    - Additional information: Country, Town, Demographic information?

We want to test the effect of tutoring on student grades (in [0, 100]).

We construct two groups of n students. Group 1 receives 1-on-1 tutoring while Group 2 does normal supervised exercises sessions. Is the expense worth it?

Data: 
$$X_1 \dots X_n$$
 and  $Y_1 \dots Y_n$ .

Conditional model:

$$X_i \sim \mathcal{N}\left(\mu_X, (10)^2\right)$$
  
 $Y_i \sim \mathcal{N}\left(\mu_Y, (10)^2\right)$ 

Prior model:

$$\mu \stackrel{\textit{IID}}{\sim} \mathcal{N}\left(60, (10)^2\right)$$

encoding the fact that I expect the average grade of students to not deviate too much from the typical value 60.

Posterior is independent:

$$\mu_X|x_1 \dots x_n \sim \mathcal{N}\left(\frac{n\bar{x}+60}{n+1}, \frac{(10)^2}{n+1}\right)$$

$$\mu_Y|y_1 \dots y_n \sim \mathcal{N}\left(\frac{n\bar{y}+60}{n+1}, \frac{(10)^2}{n+1}\right)$$

Posterior difference of  $\mu_X - \mu_Y$  is marginally Gaussian:

$$\mu_X - \mu_Y | d \sim \mathcal{N}\left(\frac{n}{n+1}(\bar{x} - \bar{y}), \frac{2}{n+1}(10)^2\right)$$

Estimation of the difference of means?



#### Estimation:

Posterior mean, median and mode are identical:

$$\widehat{\mu_X - \mu_Y} = \frac{n}{n+1} (\bar{x} - \bar{y})$$

• 0.95 credible region:

$$I = \frac{n}{n+1} (\bar{x} - \bar{y}) \pm 2 * 10 \sqrt{\frac{2}{n+1}}$$

#### Limits:

- Model is extremely unrealistic:
  - Variance assumed known.
  - Gaussian conditional model when grade is bounded.
  - Gaussian prior model when mean-grade is bounded too.

(it makes my life simple: explicit posterior)

- Choosing the right model?
- Justify prior information??

# Conclusion

# The story so far

Bayesian inference is very close to MLE methods. It requires one additional ingredient: the prior distribution  $f(\theta)$ .

The posterior distribution is sometimes analytically tractable, most often not.

Most of the course will focus on the computational tools which can deal with this issue.

So far, we have given **no good argument** in favor of Bayesian inference. These will come later in the course.

## Bayes formula

If  $\theta$  and  $\mathcal D$  are two random variables defined on the same probability space. Then, observing  $\mathcal D=d$  modifies the probability of  $\theta$ .

Event based formulation:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A)}{\mathbb{P}(B)}\mathbb{P}(A)$$

• If  $f(\theta)$  is a density (with respect to some base measure  $d\nu_{\theta}$ ) and, for every  $\theta$ ,  $f(\mathcal{D}=d|\theta)$  is a density (with respect to some base measure  $d\nu_{\mathcal{D}}$ ),

Then the posterior is a density (with respect to the base measure  $d\nu_{ heta}$ ):

$$f(\theta|\mathcal{D}=d) = f(\theta) \frac{f(\mathcal{D}=d|\theta)}{f(\mathcal{D}=d)}$$

Define the unnormalized posterior:

$$\tilde{f}(\theta|\mathcal{D}=d)=f(\theta)f(\mathcal{D}=d|\theta)=f(\mathcal{D}=d\ \&\ \theta)$$

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## Bayesian inference

#### Bayesian statistical inference:

- Model a dataset  $\mathcal{D} = d$  with:
  - A conditional model:  $f(\mathcal{D} = d|\theta)$
  - ullet A prior model over all parameters:  $f\left(oldsymbol{ heta}
    ight)$
- Perform inference:
  - Compute the posterior distribution:

$$f(\theta|\mathcal{D}=d) = f(\theta) \frac{f(\mathcal{D}=d|\theta)}{f(\mathcal{D}=d)}$$

Squeeze out the answer from the posterior: depends on what we want.

#### Bayesian estimation

If we want to estimate the unknown parameters heta given a posterior:

- Point estimates:
  - Posterior mode (MAP)
  - Posterior Mean
  - Posterior median
  - General: minimize loss:

$$\hat{oldsymbol{ heta}}_{L} = \operatorname{argmin}_{\hat{oldsymbol{ heta}}} \left[ \mathbb{E} \left( L \left( oldsymbol{ heta}, \hat{oldsymbol{ heta}} 
ight) | \mathcal{D} = d 
ight) 
ight]$$

- Interval estimates:
  - Credible interval:  $I_{\alpha}$  such that:

$$\mathbb{P}\left(\boldsymbol{\theta} \in I_{\alpha} | \mathcal{D} = d\right) = 1 - \alpha$$

General: minimize loss??