

Markov's Inequality

If X is a random variable that takes only non-negative values, then for any value $a > 0$,

$$P[X \geq a] \leq \frac{E[X]}{a}$$

Proof:

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx$$

$$= \int_0^{\infty} x f_X(x) dx$$

$$= \int_0^a x f_X(x) dx + \int_a^{\infty} x f_X(x) dx$$

$$\geq \int_a^{\infty} x f_X(x) dx \geq a \int_a^{\infty} f_X(x) dx$$

$$= a P[X \geq a]$$

$\forall a > 0,$

$$P[X \geq a] \leq \frac{E[X]}{a}.$$

Chebyshev's inequality

If X is a r.v. with mean μ_X and variance σ_X^2 , then for any $k > 0$,

$$P[\underline{|X - \mu_X| \geq k}] \leq \frac{\sigma^2}{k^2}$$

Sol: $\{ |X - \mu_X| \geq k \} \Leftrightarrow \{ (X - \mu_X)^2 \geq k^2 \}$

Now, $P[(X - \mu_x)^2 \geq k^2] \leq \frac{E[(X - \mu_x)^2]}{k^2}$

$$= \frac{\sigma_x^2}{k^2}$$

$$\Rightarrow P[|X - \mu_x| \geq k] \leq \frac{\sigma_x^2}{k^2}$$

The weak law of large numbers

Let X_1, X_2, \dots be a sequence of identically distributed independent (i.i.d.) r.v.s each having mean $\mu = E[X]$.

Then, for any $\epsilon > 0$,

$$P \left[\left| \underbrace{\frac{X_1 + X_2 + \dots + X_n}{n}}_{\text{Sample mean}} - \mu \right| > \varepsilon \right] \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

Proof:

$$\sigma_x^2 < \infty.$$

$$E \left[\frac{X_1 + \dots + X_n}{n} \right] = \frac{1}{n} [E[X_1] + \dots + E[X_n]] = \mu$$

$$\text{Var}\left(\frac{X_1 + \dots + X_n}{n}\right) = \frac{1}{n^2} \text{Var}(X_1 + \dots + X_n)$$

$$= \frac{1}{n^2} \sum_i \text{Var}(X_i) = \frac{\cancel{n} \sigma_x^2}{n^2}$$

$$= \frac{\sigma_x^2}{n}$$

Chebyshev's inequality:

$$P[|X - \mu| \geq k] \leq \frac{\sigma_X^2}{k^2}$$

Replacing

$$X = \frac{X_1 + \dots + X_n}{n}, \quad \mu = \mu, \quad \sigma_X^2 \text{ by } \frac{\sigma_X^2}{n}, \quad k = \varepsilon.$$

$$P\left[\left|\frac{X_1 + \dots + X_n}{n} - \mu\right| \geq \varepsilon\right] \leq \frac{\sigma_X^2}{n \varepsilon^2} \rightarrow 0 \text{ as } n \rightarrow \infty$$

Ex Suppose that it is known that the number of items produced in a factory during a week is a r.v. with mean 50.

(a) What can be said about the probability that this week's production will exceed 75?

$$P[X > 75] \leq \frac{E[X]}{75} = \frac{50}{75} = \frac{2}{3}.$$

(b) If the variance of a week's production is known to equal 25, then what can be said about the probability that this ~~re~~ week's production will be between 40 and 60?

$$\begin{aligned} P[40 < X < 60] &= P[|X - 50| < 10] \\ &= 1 - P[|X - 50| \geq 10] \end{aligned}$$

$$P[|X-50| \geq 10] \leq \frac{\sigma_x^2}{10^2} = \frac{25}{100} = \frac{1}{4}$$

$$\Rightarrow P[|X-50| < 10] \geq 1 - P[|X-50| \geq 10] \\ \geq 1 - \frac{1}{4} = \frac{3}{4}.$$

Ex: Suppose that a sequence of i.i.d.

[$X_1, X_2, \dots, X_n, \dots$ are independent
and share the same c.d.f.]

is performed. Let E be a fixed event
and $P(E)$ denotes the probability
that E occurs on a given trial.

Let $X_i = \begin{cases} 1 & \text{if } E \text{ occurs on a trial } i \\ 0 & \text{if } E \text{ does not occur on a trial } i \end{cases}$
 $i = 1, 2, 3, 4, \dots$

$X_1 + X_2 + \dots + X_n$: no. of times that E occurs in the 1st n trials.

$$E[X_i] = 1 P[E] + 0 P[E^c] = P[E] \quad \forall i$$

$\Rightarrow \mu = P[E].$

$$\text{Var}(X_i) = P[E] [1 - P[E]] = \sigma_x^2 \quad \forall i.$$

Now, let $\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$

$$E[\bar{X}] = \frac{1}{n} \sum_{i=1}^n E[X_i] = \mu = P[E].$$

$$\text{Var}(\bar{X}) = \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) = \frac{\sigma_x^2}{n}$$

$$P[|\bar{X} - \mu| > \varepsilon] \rightarrow 0 \text{ as } n \rightarrow \infty$$

for every $\varepsilon > 0$.

Markov's Inequality

$$P[\bar{X} > a] \leq \frac{E[X]}{a} = \frac{\mu}{a}, \quad a > 0.$$

$$a = \mu, \quad a = \sqrt{\mu};$$

Chubyshev's inequality

$$P[|\bar{X} - \mu| > k] \leq \frac{\sigma_x^2}{k^2}$$

Suppose that a trial, or an experiment, whose outcome can be classified as either a "success" or as a "failure" is performed.

Let $X = \begin{cases} 1 & \text{when the outcome is a success} \\ 0 & \text{failure} \end{cases}$

and be s.t. $P[X=0] = 1-p$, $P[X=1] = p$, $p \in [0,1]$

Def:

say X ,

A r.v. X is said to be a Bernoulli r.v. if it assumes the values 0 or 1 and its p.m.f. is given by

$$P[X=0] = 1-p$$

$$P[X=1] = p, \quad \text{for some } p \in (0,1)$$

Suppose that "n" independent Bernoulli trials are to be performed, (say $X_i, i=1, 2, \dots, n$)

Assume that each trial results in a success with probability "p" and in a failure with prob. $1-p$.

Let $X = X_1 + X_2 + \dots + X_n$
= no. of successes that occurs in the n trials.

$$X \in \{0, 1, 2, \dots, n\}$$

$X \sim B(n, p)$ = Binomial r.v. with parameters (n, p) .

no. of times
independent
trials performed.

probability of success
in each trial.

$P[X=i]$ = Probability of n outcomes
containing i -successes
and $(n-i)$ -failures.

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nC_i

$$p^i (1-p)^{n-i}$$

$i=0,1,\dots,n$

$$[p+(1-p)]^n$$

Let $X \sim B(n, p)$.. Show that
 $E[X] = np$.

& $\text{Var}(X) = np(1-p)$.

$$\left[\text{Var}(X) < E[X] \right]$$

Sol

$$E[X] = \sum_{i=0}^n i P[X=i]$$

$$= \sum_{i=1}^n i {}^n C_i p^i (1-p)^{n-i}$$

$$= \sum_{i=1}^n \frac{n! i}{(n-i)! i!} p^i (1-p)^{n-i}$$

$$= np \sum_{i=1}^n \frac{(n-1)!}{(n-i)! (i-1)!} p^{i-1} (1-p)^{n-i}$$

$$= np \sum_{i=1}^n {}^{n-1}C_{i-1} p^{i-1} (1-p)^{n-i}$$

$$= np [p + (1-p)]^{n-1}$$

$$= np$$

$$\# \quad X \sim B(n, p).$$

$$X = X_1 + X_2 + \dots + X_n$$

= Sum of n -independent Bernoulli trials with probability of success p in each trial.

$$\Rightarrow E[X] = \sum_{i=1}^n E[X_i] = np$$

$$\text{Var}(X) = \sum_{i=1}^n \text{Var}(X_i) = np(1-p).$$

Ex: It is known that disks produced by a certain company will be defective with probability 0.01 independently of each other. The company sells the disks in packages of 10 and offer a money-back guarantee that at most 1 of the 10 disks is defective. What proportion of package is returned?

Sol: Let X denotes the number of defective disks in a Package.

$$X \in \{0, 1, \dots, 10\}.$$

Since each disk will be defective with prob. 0.01 and non-defective with prob. 0.99, therefore, it is a Bernoulli trial.

Also, ~~each~~ Bernoulli trials ~~is~~ ^{are} independent of each other.

$$\Rightarrow X \sim B(10, 0.01).$$

$$P[X > 1] = 1 - P[X \leq 1]$$

$$= 1 - P[X=0] - P[X=1]$$

$$= 1 - {}^{10}C_0 (0.01)^0 (0.99)^{10} - {}^{10}C_1 (0.01)^1 (0.99)^9$$

$$\approx 0.005.$$

Since each package will, independently, have to be replaced with probability 0.005. therefore, from ~~the~~ the law of large numbers it follows that in the long ~~run~~ run 0.5% of the packages will have to be replaced.

Ques: If someone buys three packages, what is the probability that exactly one of

them will be returned?

Ans:-

Y = no. of packages that the person will have to return.

$$Y \sim B(3, 0.005)$$

$$\begin{aligned} P[Y=1] &= {}^3C_1 (0.005) (.995)^2 \\ &= 0.015. \end{aligned}$$

Ex: If X_1 and X_2 are independent binomial r.v.s such that

$$X_i \sim B(n_i, p), \quad i=1, 2.$$

then,

$$X_1 + X_2 \sim B(n_1 + n_2, p).$$