

Ex: The joint p.d.f. of X and Y is

$$f_{X,Y}(x,y) = \begin{cases} 2e^{-x}e^{-2y}, & 0 < x < \infty, 0 < y < \infty. \\ 0, & \text{otherwise.} \end{cases}$$

Find (i) $P[X > 1, Y < 1]$ | (iii') $P[X < a]$
(ii) $P[X < Y]$

$$\begin{aligned}
 (1) \quad P[X > 1, Y < 1] &= P[X \in (1, \infty), Y \in (-\infty, 1)] \\
 &= P[(X, Y) \in \underbrace{(1, \infty) \times (-\infty, 1)}_B]
 \end{aligned}$$

$$= \int \int_B f_{X,Y}(x,y) dx dy$$

$$= \int_{-\infty}^1 \int_1^{\infty} f_{X,Y}(x,y) dx dy$$

$$= \int_0^1 \int_1^{\infty} f_{X,Y}(x,y) dx dy$$

$$= 2 \int_0^1 \int_1^\infty (e^{-x} e^{-2y}) dx dy$$

$$= 2 \left(\int_0^1 e^{-2y} dy \right) \left(\int_1^\infty e^{-x} dx \right)$$



$$\left[\frac{e^{-2y}}{-2} \right]_0^1$$

$$= \boxed{1 - e^{-2}}$$

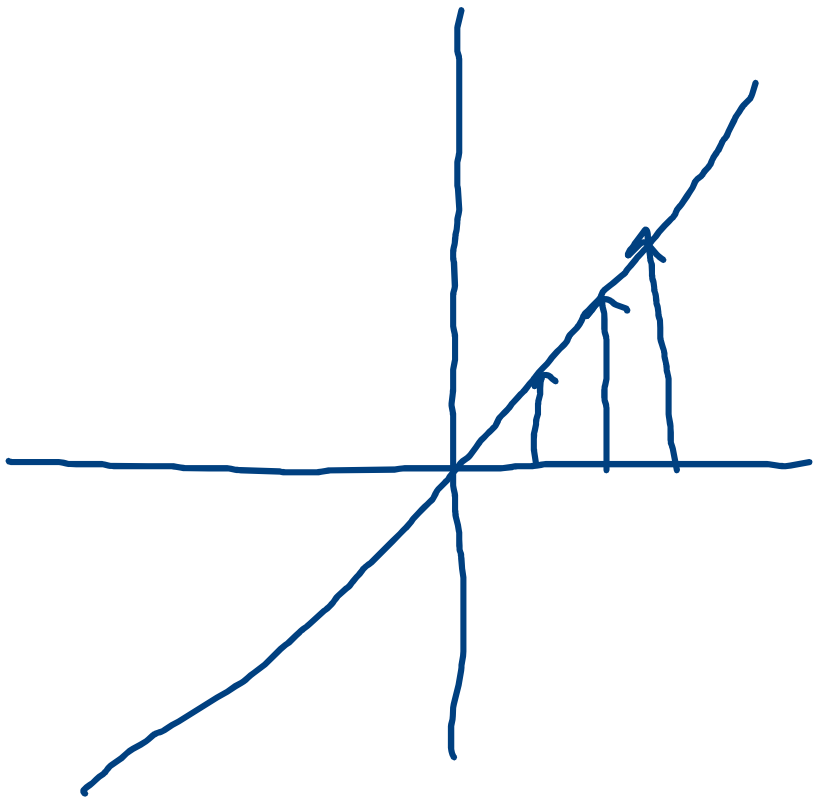


$$\lim_{a \rightarrow \infty} \int_1^a e^{-x} dx$$

$$= \lim_{a \rightarrow \infty} [e^{-x}]_1^a$$

$$= \lim_{a \rightarrow \infty} [1 - e^{-a}] = 1$$

$$(ii) P[X < Y] = \iint_{\{(x,y): x < y\}} f_{X,Y}(x,y)$$



$$= \int_0^{\infty} \int_0^y 2e^{-x} e^{-2y} dx dy$$

$$= 2 \int_0^{\infty} e^{-2y} [1 - e^{-y}] dy$$

$$= 2 \int_0^{\infty} (e^{-2y} - e^{-3y}) dy =$$

$$\int_0^{\infty} e^{-ay} dy = \lim_{t \rightarrow \infty} \int_0^t e^{-ay} dy = -\frac{1}{a} \lim_{t \rightarrow \infty} [e^{-at} - 1] = \frac{1}{a}$$

$$P[X < Y] = 2 \left[\frac{1}{2} - \frac{1}{3} \right] = 2 \cdot \frac{1}{6} = \frac{1}{3}$$

$$\begin{aligned} \text{(iii)} \quad P[X < a] &= P[X < a, Y \in (-\infty, \infty)] \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^a f_{X,Y}(x,y) dx dy \end{aligned}$$

$$F_X(a) = P[X \leq a] = P[X \leq a, Y < \infty]$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^a f_{X,Y}(x,y) dx dy$$

$$= 2 \int_0^{\infty} \int_0^a e^{-x} e^{-2y} dx dy$$

$$= 2 \left(\int_0^{\infty} e^{-2y} dy \right) \left(\int_0^a e^{-x} dx \right)$$

$$= \cancel{2} [1 - e^{-a}] \frac{1}{\cancel{2}} = 1 - e^{-a}$$

Independent r.v.

We say that r.v. X and Y are independent if

$$P[X \in A, Y \in B] = P[X \in A]P[Y \in B].$$

for any $A, B \subseteq \mathbb{R}$.

In terms of joint c.d.f.

X and Y are independent if

$$\begin{aligned} \Rightarrow F_{X,Y}(x,y) &= P[\underbrace{X \leq x}_{x \in (-\infty, x]} \underbrace{, Y \leq y}_{(-\infty, y]}] \\ &= P[X \leq x] P[Y \leq y] \\ &= F_X(x) F_Y(y) \end{aligned}$$

In terms of joint p.m.f.

X and Y are independent if

$$\begin{aligned}\rightarrow P_{X,Y}(x_i, y_j) &= P[X=x_i, Y=y_j] \\ &= P[X=x_i] P[Y=y_j] \\ &= p_X(x_i) p_Y(y_j).\end{aligned}$$

In terms of joint p.d.f.

X and Y are independent if

$$f_{X,Y}(x,y) = f_X(x) f_Y(y).$$

Ex: let X and Y be independent r.v.
that have common density funⁿ.

$$\rightarrow f_X(x) = \begin{cases} e^{-x} & \text{if } x > 0 \\ 0 & \text{otherwise.} \end{cases}$$

\rightarrow Find p.d.f. of $\boxed{X/Y}$ $\frac{X}{Y}$

$$F_{X/Y}(a) = 0 \quad \text{if} \quad a \leq 0$$

For ~~all~~ ~~a~~ $a > 0$,

$$F_{X/Y}(a) = \mathbb{P}\left[\frac{X}{Y} \leq a\right] = \mathbb{P}[X \leq aY]$$

$$= \iint_{(x,y): x \leq ay} f_{X,Y}(x,y) dx dy$$

$$= \int_0^{\infty} \int_0^{ay} [f_X(x) f_Y(y)] dx dy$$

$$F_{\cancel{X/Y}}(a) = \int_0^{\infty} \int_0^{ay} (e^{-x})(e^{-y}) dx dy$$

$$= \int_0^{\infty} e^{-y} [1 - e^{-ay}] dy$$

$$= \int_0^{\infty} [e^{-y} - e^{-(1+a)y}] dy$$

$$= 1 - \frac{1}{1+a}$$

$$\Rightarrow f_{X/Y}(a) = F'_{X/Y}(a) = \frac{1}{(1+a)^2}, \quad 0 < a < \infty$$

Expectation of Random Variable

$$\begin{aligned} & \int x f_X(x) dx \\ & \approx \sum x P[x < X < x+dx] \\ & \sum x P[x < X < x+dx] \end{aligned}$$

Let X be a discrete r.v. that assumes the values x_1, x_2, x_3, \dots

Then, the expected value of X is the weighted average of the values of X , defined by

$$E[X] = \sum_{x \in X(\Omega)} x p_X(x)$$

Ex ① $X \in \{0, 1\}$, $p_X(0) = \frac{1}{2}$, $p_X(1) = \frac{1}{2}$

$$E[X] = 0 \times \frac{1}{2} + 1 \times \frac{1}{2} = \frac{1}{2} = \frac{0+1}{2}$$

$X \in \{a, b\}$, $p_X(a) = \frac{1}{2} = p_X(b) = \frac{1}{2}$

$$E[X] = \frac{a+b}{2}$$

② $X \in \{0, 1\}$, $p_X(0) = \frac{1}{4}$, $p_X(1) = \frac{3}{4}$

$$E[X] = 0 \times \frac{1}{4} + 1 \times \frac{3}{4} = \frac{3}{4} = 0.75$$

Ex: let a r.v. X denotes the outcome when we roll a fair dice. Find $E[X]$.

$$X \in \{1, 2, 3, 4, 5, 6\}, \quad p_X(i) = \frac{1}{6}$$

$$E[X] = 1 \times \frac{1}{6} + 2 \times \frac{1}{6} + 3 \times \frac{1}{6} + \dots + 6 \times \frac{1}{6}$$

$$= \frac{1 + 2 + 3 + 4 + 5 + 6}{6} = \frac{21}{6} = \frac{7}{2}.$$

$$X^2 = i^2 \in \{1, 4, 9, 16, 25, 36\}$$

Ex: Let I is an indication ~~on~~ random variable for the event A , i.e.,

$$I = \begin{cases} 1 & \text{if } A \text{ occurs} \\ 0 & \text{if } A \text{ does not occur} \end{cases}$$

Then, $E[I] = ?$

$$= 0 \times P(A^c) + 1 \times P[A] = P[A]$$

Let X be a cont. r.v. Then, the expectation of X will be defined as

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

→ Note: For small " dx ", we have

$$f(x) dx = P[x < X < x + dx].$$

Ex: Suppose that you are expecting a message at some time past 5 P.M.

From experience you know that X , the number of hours after 5 P.M., until the message arrives, is a r.v. with p.d.f

$$f_X(x) = \begin{cases} \frac{1}{1.5} & 0 < x < 1.5 \\ 0 & \text{otherwise} \end{cases}$$

$X \sim U(a, b)$
 $f_X(x) = \begin{cases} \frac{1}{b-a} & a < x < b \\ 0 & \text{otherwise} \end{cases}$
is called
uniform random
variable.

Find the expected amount of time past 5 P.M, untill the message arrives.

$$\begin{aligned} E[X] &= \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^{1.5} x f_X(x) dx \\ &= \frac{1}{1.5} \int_0^{1.5} x \cdot 2x dx = \frac{1}{3} [2 \cdot 25 - 0] \\ &= 0.75 \approx 45 \text{ min.} \end{aligned}$$

Expectation of a function of r.v.

X is a r.v.

$$E[g(X)]$$

Ex: let X be a discrete r.v. that takes the values 0, 1, 2.

The p.m.f. of X is

$$\begin{array}{l} p_X(0) = 0.2 \\ p_X(1) = 0.5 \\ p_X(2) = 0.3 \end{array}$$

Find $E[X^2]$,
 $E[X^3]$,

Sol. $X \in \{0, 1, 2\}$, $p_X(0) = 0.2$, $p_X(1) = 0.5$, $p_X(2) = 0.3$.

$$X^2 \in \{0, 1, 4\}, \quad P[X^2=0] = P[X=0] = 0.2$$

$$P[X^2=1] = P[X=1] = 0.5$$

$$P[X^2=4] = P[X=2] = 0.3$$

$$\begin{aligned} E[X^2] &= 0 \cdot p_{X^2}(0) + 1 \times p_{X^2}(1) + 4 \times p_{X^2}(4) \\ &= 0.5 + 1.2 = 1.7 \end{aligned}$$

$$X^3 \in \{0, 1, 8\}, \quad p_{X^3}(0) = 0.2, \quad p_{X^3}(1) = 0.5, \quad p_{X^3}(8) = 0.3$$

$$E[X^3] = 0 \times 0.2 + 1 \times 0.5 + 8 \times 0.3 = 2.9.$$

$$E[X^2] = \sum_{x \in X(\Omega)} x^2 p_x(x)$$

$$E[X^3] = \sum_{x \in X(\Omega)} x^3 p_x(x)$$

Ex: The time, in hours, it takes to locate and repair an electrical breakdown in a certain factory is a r.v. with p.d.f

$$f_X(x) = \begin{cases} \frac{1}{T-0} & \text{if } 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$

If the cost involved in breakdown of duration x is x^3 , what is the expected cost of such breakdown?

$$X \sim U(0,1), \quad X^3 \sim$$

$$F_{X^3}(a) = P[X^3 \leq a] = 0 \quad \text{if } a \leq 0.$$

$$= P[X^3 \leq a] \quad \text{if } \cancel{a} a > 0.$$

$$= P[X \leq a^{1/3}]$$

$$= \int_{-\infty}^{a^{1/3}} f_X(x) dx = \int_0^{a^{1/3}} f_X(x) dx$$

$$F_{X^3}(a) = \begin{cases} 0 & \text{if } a \leq 0 \\ a^{4/3} & \text{if } 0 < a < 1 \\ 1 & \text{if } a \geq 1 \end{cases}$$

$$f_{X^3}(a) = F'_{X^3}(a) = \frac{1}{3} a^{-2/3}, \quad 0 < a < 1$$

and $f_{X^3}(a) = 0$ $a \leq 0$ and $a \geq 1$

$$\text{let } Y = X^3$$

$$E[Y] = \int_{-\infty}^{\infty} y f_Y(y) dy$$

$$= \frac{1}{3} \int_0^1 y y^{-2/3} dy$$

$$= \frac{1}{3} \int_0^1 y^{1/3} dy = \frac{1}{3} \times \frac{3}{4} \left[y^{4/3} \right]_0^1$$
$$= \frac{1}{4}$$

Verify that

$$\int_{-\infty}^{\infty} \boxed{x^3} f_x(x) dx = \frac{1}{4}.$$

for this example.

$$\checkmark \quad E[X^3] = \int_{-\infty}^{\infty} x^3 f_x(x) dx.$$

X
X²
X³
X⁴

It can be shown that

$$E[(X-\mu)^n] = \int_{-\infty}^{\infty} x^n f_X(x) dx \quad n \in \mathbb{N}$$

$$= \sum_{x \in X(\Omega)} x^n P_X(x)$$

\nwarrow
 n th moment of X .

If X is a discrete r.v. with p.m.f. $p_X(x)$, then for any real-valued function g ,

$$E[g(X)] = \sum_{x \in X(\Omega)} g(x) p_X(x).$$

$$E[\sqrt{X}] = \sum_{x \in X(\Omega)} \sqrt{x} p_X(x)$$

If X is a cont. r.v. with p.d.f. $f_X(x)$,
- then for any real-valued funⁿ g ,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x) f_X(x) dx.$$