DS4400 HW3

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1. **MAP estimation.** Consider a Bernoulli random variable x with $p(x=1)=\theta$. Given a dataset $D=\{x_1,\dots,x_N\}$, assume N_1 is the number of trials where $x_i=1$, N_0 is the number of trials where $x_i=0$ and $N=N_0+N_1$ is the total number of trials. Consider the following prior, that believes the experiments is biased:

$$p(\theta)=\begin{cases} 0.2 & \text{if } \theta=0.6 \\ 0.8 & \text{if } \theta=0.8 \\ 0 & \text{otherwise} \end{cases}$$

   (a) Write down the likelihood function, i.e., $p(D|\theta)$. What is the maximum likelihood solution for θ (we already have derived this in the class)?

   Solution:

   we write $P(X=x)=\theta^x(1-\theta)^{(1-x)}$

   $P(D|\theta)=P(X_1=x_1,X_2=x_2\dots,X_N=x_N)$

   $=\prod_{x=1}^{N}\theta_i^x(1-\theta)^{(1-x_i)}$

   $=\theta^{N_1}(1-\theta)^{N_0}$

   Let $J(\theta)=logP(D|\theta)=N_1log(\theta)+N_0log(1-\theta)$

   $\frac{\partial J(\theta)}{\partial\theta}=\frac{N_1}{\theta}-\frac{N_0}{1-\theta}$. Let $\frac{\partial J(\theta)}{\partial\theta}=0$.

   Then $\hat{\theta}=\frac{N_1}{N_0+N_1}=\frac{N_1}{N}$

   Therefore, the maximum likelihood solution is $\hat{\theta}=\frac{N_1}{N}$

   (b) Consider maximizing the posterior distribution, $p(D|\theta)\times p(\theta)$, that takes advantage of the prior. What is the MAP estimation?

   Solution:

$$p(D|\theta)\times p(\theta)=\begin{cases} 0.2\cdot 0.6^{N_1}\cdot 0.4^{N_0} & \theta=0.6 \\ 0.8\cdot 0.8^{N_1}\cdot 0.2^{N_0} & \theta=0.8 \\ 0 & \text{otherwise} \end{cases}$$

   $\frac{P(D|0.6)P(0.6)}{P(D|0.8)P(0.8)}=\frac{1}{4}(\frac{3}{4})^{N_1}(2)^{N_0}=3^{N_1}4^{-N_1-1}2^{N_0}=2^{N_1log_2^3}2^{-2N_1-2}2^{N_0}=2^{N_0-(2-log_2^3)N_1-2}$

   Therefore, when $\frac{P(D|0.6)P(0.6)}{P(D|0.8)P(0.8)}\geq 1$:

   $N_0-(2-log_2^3)N_1-2\geq 0\Rightarrow N_0\geq(2-log_2^3)N_1+2$

   Therefore, $\hat{\theta}=\begin{cases} 0.6 & N_0\geq(2-log_2^3)N_1+2 \\ 0.8 & N_0<(2-log_2^3)N_1+2 \end{cases}$

2. **Naive Bayes Classifier.** Assume you have the following training set with two binary features x_1 and x_2, and a binary response/output y. Suppose you have to predict y using a naive Bayes classifier.

x_1	x_2	y
1	0	0
0	1	0
0	0	0
1	0	1
0	0	1
0	1	1
1	1	1

(a) Compute the Maximum Likelihood Estimates (MLE) for θ_j^y for $j = 0, 1$ as well as $\theta_{\bar{x}_l|y}^{x_l|y}$ for $j = 0, 1$ and for $\ell = 1, 2$.

Solution:

$$\theta_j^y = \begin{cases} \frac{3}{7} & j = 0 \\ \frac{4}{7} & j = 1 \end{cases}$$

$$\theta_{\bar{x}_\ell|j}^{x_\ell|y} = \begin{cases} \frac{1}{3} & x_1 = 1, j = 0 \\ \frac{2}{3} & x_1 = 0, j = 0 \\ \frac{1}{3} & x_2 = 1, j = 0 \\ \frac{2}{3} & x_2 = 0, j = 0 \\ \frac{1}{2} & x_1 = 1, j = 1 \\ \frac{1}{2} & x_1 = 0, j = 1 \\ \frac{1}{2} & x_2 = 1, j = 1 \\ \frac{1}{2} & x_2 = 0, j = 1 \end{cases}$$

(b) After learning via MLE is complete, what would be the estimate for $P(y = 0|x1 = 0, x2 = 1)$.

Solution:

$P(y = 0|x_1 = 0, x_2 = 1)$
$= \frac{P(x_1=0,x_2=1|y=0)P(y=0)}{P(x_1=0,x_2=1)}$
$= \frac{P(x_1=0|y=0)P(x_2=1|y=0)P(y=0)}{P(x_1=0,x_2=1)}$
$= \frac{\theta_{0|0}^{x_1|y}\theta_{1|0}^{x_2|y}\theta_0^y}{P(x_1=0,x_2=1)}$
$= \frac{2}{3} \cdot \frac{1}{3} \cdot \frac{3}{7} / \frac{2}{7}$
$= \frac{1}{3}$

(c) What would be the solution of the previous part without the naive Bayes assumption?

Solution:

Without Naive Bayes Assumption:
$P(y = 0|x_1 = 0, x_2 = 1)$
$= P(y = 0, x_1 = 0, x_2 = 1)/P(x_1 = 0, x_2 = 1)$
$= \frac{1}{2}$

3. Constrained Optimization. Consider the regression problem on a dataset $\{(x_i, y_i)\}N_i = 1$, where $x_i \in \mathbb{R}^k$ denotes the input and y_i denotes the output/response. Let $\mathbf{X} = [\mathbf{x}_1 \ \dots \ \mathbf{x}_N]^T$ and $\mathbf{y} = [\mathbf{y}_1 \ \dots \ \mathbf{y}_N]^T$. Consider the following regression optimization.

$$\min_\theta \frac{1}{2}||\mathbf{y} - \mathbf{X}\theta||_2^2$$

$$\text{s.t. } \omega^{\mathbf{T}}\theta = \mathbf{b}$$

where ω and b are given and indicate the parameters of a hyperplane on which the desired parameter vector, θ, lies.

(a) Assume that $X^T X = I^k$, where I_k denotes the identity matrix. Find the closed-form solution of the above regression problem.

Solution:

Let $L(\theta,\alpha) = \frac{1}{2}\|\mathbf{y} - \mathbf{X}\theta\|_2^2 + \alpha(\omega^T\theta - b)$

Then:

$\frac{\partial L(\theta,\alpha)}{\partial\theta} = \frac{1}{2}\frac{\partial\|\mathbf{y}-\mathbf{X}\theta\|_2^2}{\partial\theta} + \alpha\frac{\partial\omega^T\theta-b}{\partial\theta}$

$= \frac{1}{2}\frac{\partial\left\|\begin{bmatrix} y_1 - x_1\theta \\ y_2 - x_2\theta \\ \dots \\ y_N - x_N\theta \end{bmatrix}\right\|_2^2}{\partial\theta} + \alpha\omega$

$= \frac{1}{2}\frac{\partial\sum_{i=1}^N(y_i - x_i\theta)^2}{\partial\theta} + \alpha\omega$

$= \frac{1}{2}\sum_{i=1}^N 2(y_i - x_i\theta)(-x_i) + \alpha\omega$

$= \sum_{i=1}^N(-x_i y_i + (x_i)^T x_i\theta)(-x_i) + \alpha\omega$

$= -X^T y + X^T X\theta + \alpha\omega$

$= -X^T y + \theta + \alpha\omega$

$\frac{\partial L(\theta,\alpha)}{\partial\alpha} = \omega^T\theta - b.$

Let $\begin{cases} \frac{\partial L(\theta,\alpha)}{\partial\theta} = 0 \\ \frac{\partial L(\theta,\alpha)}{\partial\alpha} = 0 \end{cases}$ Then, $\begin{cases} -X^T y + \theta + \alpha\omega = 0 \\ \omega^T\theta - b = 0 \end{cases}$

Since, $\theta = X^T y - \alpha\omega$,

then $\omega^T(X^T y - \alpha\omega) = b$.

$\Rightarrow \omega^T X^T y - \omega^T\alpha\omega - b = 0.$

$\Rightarrow \omega^T X^T y - b = \omega^T\omega\alpha$

$\Rightarrow \alpha = \frac{\omega^T X^T y - b}{\omega^T\omega}.$

Therefore $\theta^* = X^T y - \frac{\omega^T X^T y - b}{\omega^T\omega}\omega$

(b) Verify if your obtained solution θ^* satisfies the constraint $\omega^T\theta^* = b$.

Solution:

$\omega^T\theta^*$

$= \omega^T(X^T y - \frac{\omega^T X^T y - b}{\omega^T\omega}\omega)$

$= \omega^T X^T y - \frac{\omega^T X^T y - b}{\omega^T\omega}\omega^T\omega$

$= \omega^T X^T y - (\omega^T X^T y - b)$

$= b$

(c) What you have been the solution of this optimization, if the constraint $\omega^T\theta = b$ was not present?

Solution:

If there is no constraint $\omega^T\theta = b$, it is just $\min_\theta \frac{1}{2}\|\mathbf{y} - \mathbf{X}\theta\|_2^2$.

Then $\frac{\partial\frac{1}{2}\|\mathbf{y}-\mathbf{X}\theta\|_2^2}{\partial\theta} = \theta - X^T y$. Let it to be 0.

Then $\theta^* = X^T y$