# DS4400 Notes

#### DS4400 Notes 01/24

1. Convex functions:

A function  $f: \mathbb{R}^d \to \mathbb{R}$  is convex iff  $\forall \theta_1, \theta_2 \in \mathbb{R}^d$  and  $\forall \alpha \in [0,1]$  we have  $f(\alpha \theta_1 + (1-\alpha)\theta_2) \le \alpha f(\theta_1) + (1-\alpha)f(\theta_2)$ 

In the special case (d = 1)  $f : \mathbb{R} \to \mathbb{R}$ , f is convex iff  $\forall \theta, f''(\theta) \ge 0$ 

When the function is convex, **local min**  $\equiv$  **global min**. When the system is not convex, we might find only a **local min** but not a **global min** 

2. Dealing with non convex function:

In gradient descent:

- (a) use larger  $\rho$  in the beginning and gradually decrease  $\rho$  with interation.
- (b) Run SGD/GD with multiple random initializations  $\theta_1^{(0)}, \theta_2^{(0)}...$  and keep the best solution.
- 3.  $\operatorname{arg\,min}_{\theta} \sum_{i=1}^{N} (y_i \theta^T x_i)^2 \triangleq J(\theta)$

In linear regression,  $J(\theta)$  is convex.

- 4. Robustness of Regression to outliers:
  - (a) Run outlier detection algorithm, remove detected outliers, then run Linear Regression on remaining points.
  - (b) Robust Regression cost function.

$$\operatorname{arg\,min}_{\theta} \sum_{i=1}^{N} e_i^2$$
,  $e_i \triangleq y_i - \theta^T x_i$ 

 $e^2$  is extremly unhappy with large errors.

we might use |e| to replace the function. This might be more tolerance. Then,  $\arg\min_{\theta} \sum_{i=1}^{N} |y_i - \theta^T x_i|$ 

5. Exercise:  $D = \{(x_1, y_1 = 100)...(x_10, y_10 = 100), (x_{11}, y_{11} = 0), (x_{12}, y_{12} = 0)\}$ 

$$e^2$$
:  $10(\theta - 100)^2 + 2\theta^2 \rightarrow$   
 $\frac{\partial}{\partial \theta} = 20(\theta - 100) + 4\theta = 0 \rightarrow$   
 $\theta = 83.3$ 

$$|e| : \min_{\theta} \sum_{i=1}^{12} |\theta - y_i| = 10|\theta - 100| + 2\theta$$

$$(\theta \le 100) = \min_{\theta} 10(100 - \theta) + 2\theta$$

$$=1000-8\theta \rightarrow \theta=100$$

$$(\theta \ge 100) = \min_{\theta} 10(\theta - 100) + 2\theta$$

$$=12\theta-1000\to\theta=100$$

- 6. How to solve 11-norms cost functions?
  - (a) No closed form
  - (b) we need to be careful with gradient descent
  - (c) We need to use convex programming toolboxs (convex optimizations)
- 7. Huber loss funct

$$l_{\delta}(e) = \begin{cases} \frac{1}{2}e^{2} & |e| \leq \delta \\ \delta|e| - \frac{\delta^{2}}{2} & |e| \geq \delta \end{cases}$$

$$\frac{\partial l_{\delta}(e)}{\partial e} = \begin{cases} e & -\delta \leq ele\delta \\ \delta & e > \delta \\ -\delta & e < \delta \end{cases}$$

in huber loss function, we don't have closed form solution but we can run gredient descent now.

8. Definition: Overfitting:

Learning a system from traning data that does very well on training data itself (e.g, very low regression error on traning data), but performs poorly on test data.

9. Definition: Overfitting in Linear Regression

$$\Phi^T \Phi \theta = \Phi^T Y$$
  
$$\Rightarrow \theta^* = (\Phi^T \Phi)^{-1} \Phi^T Y$$

$$\operatorname{rank}(\Phi^T \Phi) \le \min\{rk(\Phi^T), rk(\Phi)\} = rk(\Phi) \le \min\{N, d\}$$

 $\Phi^T \Phi$  is  $d \times d$  matrix, then rank is  $\leq d$ .

Therefore, when N < d it is not invertible which means we have multiple solutions and results in overfitting.

### DS4400 Notes 01/28

1. Definition: Overfitting

Refers to situation where the learned model does well on traning data and poorly on testing data.

As d (dimension of system) increases, then training error godes down (can be exactly ZERO for sufficiently large d)

2. In Linear regression:

$$\min \sum_{i=1}^n (\theta^T \phi(x_i) - y_i)^2$$

set the derivative to 0 and we find

$$\Phi^T \Phi \theta = \Phi^T Y$$

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Then  $\theta^* = (\Phi^T \Phi)^{-1} \Phi^T Y$ 

When is it the case that  $\Phi^T\Phi$  is not invertible?

Since 
$$\Phi^T \Phi \in \mathbb{R}^{N \times d}$$

$$rk(\Phi^T\Phi) \le rk(\Phi) \le min\{N,d\}$$

 $\Phi^T \Phi \in \mathbf{R}^{d \times d}$  is invertible when  $rk(\Phi^T \Phi) = d$ . Therefore, when  $N < d, rk(\Phi^T \Phi) = N$ ,  $\Phi^T \Phi$  is not invertible. There will be infinitely many solutions for  $\theta$ .

## Generally, need sufficient # samples

## 3. Test overfitting.

If  $\Phi^T \Phi$  is not invertible,

$$\exists v \neq 0, \Phi^T \Phi v = 0$$

 $\Rightarrow \theta^* + \alpha v$  is also a solution for any  $\alpha \in R$ 

$$\Phi^T \Phi(\theta^* + \alpha v) = \Phi^T \Phi \theta^* + \Phi^T \Phi(\alpha v)$$

$$= \Phi^T \Phi \theta^* + \alpha \Phi^T \Phi v$$

$$= \Phi^T \Phi \theta^* = \Phi^T Y$$

We can find large  $\alpha$  so that  $\theta^*$  have extremly large entries.

# Generally, if the entries are very large (abs) we might have overfitting

## 4. Treat overfitting

We want to change regreession optimization to prevent  $\theta$  from very large terms.

then we change the cost function:

$$\min_{\theta} \sum_{i=1}^{N} (\theta^T \phi(x_i) - y_i)^2 + \lambda \sum_{j=1}^{d} \theta_j^2$$

 $\lambda$ : regularization parameter (> 0)

$$\sum_{j=1}^{d} \theta_{j}^{2}$$
: regularizer.

$$\lambda \rightarrow 0$$
: back to overfitting

$$\lambda \to \infty$$
:  $\theta^* = 0$ , underfitting

# (a) closed-form

$$\frac{\partial J}{\partial \theta}$$

$$= 2\Phi^{T}(\Phi\theta - Y) + \lambda \frac{\partial \sum_{j=1}^{N} \theta_{j}^{2}}{\partial \theta}$$
$$= 2\Phi^{T}(\Phi\theta - Y) + 2\lambda\theta$$

Let it be zero:

$$\Phi^T \Phi \theta + \lambda \theta = \Phi^T Y$$

$$(\Phi^T \Phi + \lambda I_d)\theta = \Phi^T Y$$

Then 
$$\theta^* = (\Phi^T \Phi + \lambda I_d)^{-1} \Phi^T Y$$

# (b) Gradient descent

Find initial 
$$\theta^{(0)}$$

$$\theta^{t} = \theta^{(t-1)} - \rho \frac{\partial J}{\partial \theta}|_{\theta^{(t-1)}}$$
  
=  $\theta^{(t-1)} - 2\Phi^{T}(\Phi\theta^{(t-1)} - Y) + 2\lambda\theta^{(t-1)}$ 

# 5. Hyperparameter Tunning

GD: set learning rate  $\rho$ 

Robust Reg: Huber loss  $\delta$ 

overfitting and regularization:  $\lambda$ 

 $\rho$ ,  $\delta$ ,  $\lambda$  = hyperparameters

How to pick hyperparameters?

#### **BAD APPROACH 1:**

(a) pick some set of possible  $\lambda_i \in \{\lambda_1, \lambda_2...\}$ Run regression with  $\lambda_i$  and find  $\theta_i^*$ Measure regression error:

$$\epsilon_{tr}(\lambda) = \sum_{i=1}^{N} ((\theta^*(\lambda))^T x_i - y_i)^2$$

To sum: just find  $\lambda$  for which  $\epsilon_{tr}(\lambda)$  is minimum

# This approach is setting $\lambda$ back to 0 Test data needed!!!

(a) We need to Train  $\lambda_i$  on **training set** to minimize the cost function

$$2\Phi^{T}(\Phi\theta - Y) + 2\lambda\theta$$

to find  $\theta_i^*$ 

(b) Measure regression error on the **hold-out set**  $D^{ho}$ 

$$\epsilon_{tr} = \sum_{x_i, y_i \in D^{ho}} (y_i - (\theta^*(\lambda))^T x_i)^2$$

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1. Hyperparameter Tunning:

$$\min_{\theta} \sum_{i=1}^{N} (\theta^T \phi(x_i) - y_i)^2 + \lambda \sum_{j=1}^{d} \theta_j^2$$

- For  $\lambda \in \{\lambda_1, \lambda_2, \dots, \lambda_p\}$ 
  - Tran using  $D^{tr}$  with  $\lambda \to \theta^*(\lambda)$
  - Measure validation error

$$\epsilon^{tr}(\lambda) = \sum_{x_i, y_i \in D^{ho}} (y_i - (\theta^*(\lambda))^T x_i)^2$$

• select  $\lambda$  which minimizes

$$\epsilon^{ho}(\lambda) \to \lambda^* = \min_{\{\lambda_1, \lambda_2, \dots, \lambda_p\}} \epsilon^{ho}(\lambda)$$

#### 2. Problems:

- Take much longer time since we are training the models multiple times
- Each training is using a subset of the data set, then each training is amplifing the problem of overfitting.

#### 3. K-fold cross validation

divide Data set to k equally large sets  $\{D_1, D_2, ..., D_k\} \in D$ 

• For 
$$\lambda \in \{\lambda_1, \lambda_2, ..., \lambda_p\}$$

- For  $i = 1, 2, ..., k$ 

\* train on  $\bigcup_{j \neq i} D^j$  and get  $\theta_i^*(\lambda)$ 

\* compute validate error on  $D^i \to \epsilon_i^{ho}(\lambda)$ 

- compute average of  $\{\epsilon_i^{ho}(\lambda)\}$ :  $\epsilon^{ho} = \frac{1}{k} \sum_{i=1}^k \epsilon^{ho}(\lambda)$ 

• select  $\lambda^* = \min_{\{\lambda_1, \lambda_2, ..., \lambda_p\}} \epsilon^{ho}(\lambda)$ 

Once we find the best  $\lambda$ , train the model on the whole set.

#### 4. PROBABILITY REVIEW

- Random Variable: a variable that takes values corresponding to outcome of a random phenomenon.
- Discrete r.v.: descrete values
- continuous r.v. continus range of values
- Condition:  $P(X|Y) = \frac{P(X \cap Y)}{P(Y)}$

$$P(X,Y) = P(X|Y)P(Y)$$

$$P(X,Y) = P(Y|X)P(X)$$

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#### Chain rule:

$$P(X_1, X_2, ..., X_n) = P(X_1)P(X_2|X_2)P(X_3|X_1, X_2)$$
  
...  $P(X_N|X_1, X_2...X_N)$ 

Marginalization

$$p(x,y)$$
 known  
 $p(x) = \sum_{y} p(x, Y = y) = \int p(x,y)dy$ 

• Bayes Rule:

$$P(X|Y) = \frac{P(Y|X)P(X)}{P(Y)} = \frac{P(X|Y)P(Y)}{P(X)}$$

• Independence:

r.v. are independent 
$$(X \perp \!\!\!\perp Y)$$
 iff  $P(X|Y) = p(X), P(Y|X) = p(Y)$  or  $P(x,y) = P(x)p(y)$ 

conditional independence example: X = height of person, Y = vocabulary, X is not independent of Y since babies may have less vocabulary and with lower heights.
 However, X = height, Y = vocab, Z = age. Then (X ⊥⊥ Y) | Z

$$P(X, Y|Z) = P(X|Z)P(Y|Z)$$
$$\Rightarrow P(X|Y,Z) = P(X|Z)$$

• Expectation:

$$E(X) = \sum xp(x) \text{ or } \int xp(x)dx$$
  
 $E(f(X)) = \sum f(x)p(x) \text{ or } \int f(x)p(x)dx$   
Given  $X \perp \perp Y$ ,  $E[XY] = E[X]E[Y]$   
hint:  $(E[XY] = E[f(x,y)]$ 

• IID r.v: independent and identically destributed  $p(X_1 = x_1, X_2 = x_2, ..., X_n = x_n) = p(X_1 = x_1)p(X_2 = x_2)...p(X_n = x_n)$  and each expriment is identical.  $P(X_1 = \theta) = P(X_2 = \theta) = \cdots = P(X_n = \theta)$ 

DS4400 Notes 02/04

#### **Maximum Likelihood Estimation**

- 1. Some distributions:
  - Gaussian Dist.

$$P(X = x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

• Laplace Dist.

$$P(X = x) = \frac{1}{2\lambda} e^{-\frac{|x-\mu|}{\lambda}}$$

• Bernoulli Dist.

$$P_{\theta}(x=1) = \theta$$
,  $P_{\theta}(x=0) = 1 - \theta$ 

2. Goal: Learn parameters of probability models. (fix the prob. model class)

In ML, we learn these parameters  $(\theta)$  using training data, D.

We want to measure  $P_{\theta}(D)$ 

MLE:  $\theta^* = \arg \max_{\theta} P_{\theta}(D)$  Under such  $\theta^*$ , the probability of observing the given dataset is maximum.

3. Exercise: Fliping a coin.

This is a Binomial Dist. (n time Bernoulli Trials)

model: 
$$p(X = x) = \theta^x (1 - \theta)^{1-x}, x = 0, 1$$

$$P_{\theta}(D) = P_{\theta}(X_1 = x_1, X_2 = x_2...)$$

Assuming that tossing coins are iids:

$$P_{\theta}(D) = P_{\theta}(x_1)P_{\theta}(x_2)\dots$$
  
=  $\theta^{\sum x_i}(1-\theta)^{\sum (1-x_i)}$ 

Then likelihood funciton:

$$L(\theta) = P_{\theta}(D)$$

Take the logrithm of both sides (simplify product to sum)

$$\theta^* = \arg\max_{\theta} logL(\theta)$$

#### **REASON:**

- 1. log is monotonically increasing
- 2. simplify the powers to scale, the product to sum.
- 3. Increase the dynamic range (working with small numbers is not accurate on computers and memory consuming)

$$\frac{\partial logL(\theta)}{\partial \theta} = \sum x_i \frac{1}{\theta} + (N - \sum x_i) \frac{-1}{1 - \theta}$$

Let the derivative equals to 0.

$$\frac{1}{\theta} \sum x_i = \frac{1}{\theta - 1} (N - \sum x_i)$$
$$\theta = \frac{\sum x_i}{N}$$

4. Exercise: People's height

Use model: normal distribution.

$$L(\theta) = P_{\sigma,\mu}(D) = \prod_{i=1}^{N} P_e(x_i)$$

$$logL(\theta) = -\frac{N}{2}log(2\pi\sigma^{2}) - \frac{\sum_{i=1}^{N}(x_{i}-\mu)^{2}}{2\sigma^{2}} = -\frac{N}{2}log(2\pi) - \frac{N}{2}log(\sigma^{2}) - \frac{\sum_{i=1}^{N}(x_{i}-\mu)^{2}}{2\sigma^{2}}$$

$$\frac{\partial logL(\theta)}{\partial \mu} = -\sum_{i=1}^{N}(\mu - x_{i})/\sigma^{2}$$

$$\Rightarrow \hat{\mu} = \frac{\sum_{i=1}^{N}x_{i}}{N}$$

$$\frac{\partial logL(\theta)}{\partial \sigma^{2}} = -\frac{N}{2}\frac{1}{\sigma^{2}} - \frac{\sum_{i=1}^{N}(x_{i}-\mu)^{2}}{2} - \frac{1}{\sigma^{4}}$$

$$= \frac{-N}{2\sigma^{2}} + \frac{\sum_{i=1}^{N}(x_{i}-\mu)^{2}}{2\sigma^{4}}$$

$$\hat{\sigma^{2}} = \frac{1}{N}\sum_{i=1}^{N}(x_{i}-\hat{\mu})^{2}$$

DS4400 Notes 02/07

Classification:

## **Binary Classification:**

input = Email → output = 'span' vs 'non-span'

## **Multiclass Classification:**

input = Image  $\rightarrow$  output = 'car', 'bike', 'stop-sign', ...

## • Classification Setup:

Given a training dataset  $D = \{(x_1, y_1) \cdots (x_n, y_n)\}$  where  $x_i \in \mathbb{R}^d$  is input feature vector and  $y_i \in \{0, 1, 2, ..., L-1\}$ , Find a mapping  $g : \mathbb{R}^d \to \{0, 1, 2, ..., L-1\}$  s.t.  $g_w(x_i) = y_i$  for many i's  $\in \{1, 2, ..., N\}$ 

## • Assumption:

Assume that there is a hyperplane  $w^T \phi(x) = 0$  that separates data into two classes. Then set:  $w^T \phi(x) > 0 \rightarrow y = 1$ , set:  $w^T \phi(x) < 0 \rightarrow y = 0$ 

## • Logistic Regression Model:

$$P_w(y=1|x) \propto e^{w^T \phi(x)/2}$$

$$P_w(y=0|x) \propto e^{-w^T \phi(x)/2}$$

Determine Z: 
$$P_w(y = 1|x) + P_w(y = 0|x) = 1$$
:  $\frac{1}{z}e^{w^T\phi(x)/2} + \frac{1}{z}e^{-w^T\phi(x)/2} = 1$ 

$$\frac{1}{7}e^{w^T\phi(x)/2} + \frac{1}{7}e^{-w^T\phi(x)/2} = 1$$

$$\Rightarrow z = e^{w^T \phi(x)/2} + e^{-w^T \phi(x)/2}$$

Model: 
$$P_w(y = 1|x) = \frac{1}{1 + e^{-w^T \phi(x)}}$$

Model: 
$$P_w(y = 1|x) = \frac{1}{1 + e^{-w^T \phi(x)}}$$
  
Model:  $P_w(y = 0|x) = 1 - \frac{1}{1 + e^{-w^T \phi(x)}}$ 

• signoid/logistic function: 
$$\sigma(z) = \frac{1}{1+e^{-z}}$$

$$\sigma(z) = \frac{1}{1 + e^{-z}}$$

# • Logistic regression:

$$P_w(y = 1|x) = \frac{1}{1 + e^{-w^T \phi(x)}} = \sigma(w^T \phi(x))$$

• Training: Learn 
$$w^*$$
 given training data D

• Testing: 
$$P_w(y^n = 1|x^n) = \frac{1}{1 + e^{-w^*T}\phi(x^n)}$$

• Assign 
$$P > 0.5 \rightarrow class1$$
,  $P \le 0.5 \rightarrow class0$ 

# • Training via MLE:

$$\max_{w} P_w(D) = \max_{w} P_w(y_1|x_1) \cdots P_w(y_N|x_N)$$

$$= \max_{w} \prod_{i=1}^{N} P_{w}(y_{i}|x_{i})$$

$$P_w(y_i|x_i) = P_w(y_i = 1|x_1)^{y_i} P_w(y_i = 0|x_1)^{1-y_i}$$

$$\max_{w} log P_{w}(D) = \max_{w} \sum_{i=1}^{N} log \left[ \left( \frac{1}{1 + e^{-w^{T} \phi(x_{i})}} \right)^{y_{i}} + \left( \frac{1}{1 + e^{w^{T} \phi(x_{i})}} \right)^{1 - y_{i}} \right]$$

$$= \max_{w} \sum_{i=1}^{N} (y_{i}) log \left( \frac{1}{1 + e^{-w^{T} \phi(x_{i})}} \right) + (1 - y_{i}) log \left( \frac{1}{1 + e^{w^{T} \phi(x_{i})}} \right)$$

$$\begin{split} &= \max_{w} \sum_{i=1}^{N} (y_{i}) [log(\frac{1}{1 + e^{-w^{T}}\phi(x_{i})}) - log(\frac{1}{1 + e^{w^{T}}\phi(x_{i})})] + log\frac{1}{1 + e^{w^{T}}\phi(x_{i})} \\ &= \max_{w} \sum_{i=1}^{N} (y_{i}) [log(\frac{1 + e^{w^{T}}\phi(x_{i})}{1 + e^{-w^{T}}\phi(x_{i})})] + log\frac{1}{1 + e^{w^{T}}\phi(x_{i})} \\ &\max_{w} \sum_{i=1}^{N} (y_{i}) [log(e^{w^{T}}\phi(x_{i}))] + log\frac{1}{1 + e^{w^{T}}\phi(x_{i})} \\ &= \max_{w} \sum_{i=1}^{N} (y_{i}w^{T}\phi(x_{i})) - log(1 + e^{w^{T}}\phi(x_{i})) \\ &\equiv \min_{w} \sum_{i=1}^{N} -y_{i}w^{T}\phi(x_{i}) + log(1 + e^{w^{T}}\phi(x_{i})) \\ &\text{Derivative:} \\ &\frac{\partial J}{\partial w} \end{split}$$

#### DS4400 Notes 02/11

1. REVIEW: logistic model:  $P_w(y=1|x) = \frac{1}{1+e^{-w^T}\phi(x)}$ 

$$\ell(x) = \log P_w(D) = \log \prod_{i=1}^N P_w(y_i|x_i)$$

$$= \sum_{i=1}^N [y_i \phi(x_i)^T w - \log(1 + e^{w^T \phi(x_i)})] \text{ maximizing } \ell(w) \equiv \min_w \sum_{i=1}^N -y_i w^T \phi(x_i) + \log(1 + e^{w^T \phi(x_i)})$$

$$= \min_w \sum_{i=1}^N -y_i w^T \phi(x_i) + \log(1 + e^{w^T \phi(x_i)})$$

$$= \min_w -y_i \phi(x_i)^T w + \log(1 + e^{w^T \phi(x_i)})$$
derivative:

$$\frac{\partial -\ell(w)}{\partial w} = \sum_{i} -y_{i}\phi(x_{i}) + \frac{1}{1+e^{w^{T}\phi(x_{i})}}(e^{w^{T}\phi(x_{i})})(\phi(x_{i}))$$

$$\frac{\partial w}{\partial x} = \sum_{i} -y_{i} \phi(x_{i}) + \frac{e^{w^{T} \phi(x_{i})} \phi(x_{i})}{1 + e^{w^{T} \phi(x_{i})}}$$

$$= \sum_{i} -y_{i} \phi(x_{i}) + \frac{e^{w^{T} \phi(x_{i})} \phi(x_{i})}{1 + e^{w^{T} \phi(x_{i})}}$$

$$= \sum_{i} -y_{i} \phi(x_{i}) + \frac{\phi(x_{i})}{1 + e^{-w^{T} \phi(x_{i})}}$$

$$= \sum_{i} (-y_{i} + \frac{1}{1 + e^{-w^{T} \phi(x_{i})}}) \phi(x_{i})$$
No closed form solution for = 0.

- 2. GD of logistic regression
  - Initialize  $w^0$

• For 
$$t = 1, 2, \dots$$
 (until converge)

$$- w^{t} = w^{t-1} - \rho \frac{\partial J}{\partial w}|_{t-1}$$

$$= w^{t-1} - \rho \sum (-y_{i} + \frac{1}{1 + e^{-w^{T} \phi(x_{i})}}) \phi(x_{i})|_{t-1}$$

3. overfitting.

overfitting: do well on training but poorly on testing. Symptom: w with large entries.

4. Regularized logistic regression:

$$\begin{aligned} \min_{w} J(w) &= -\ell(w) + \frac{\lambda}{2} ||w||_{2}^{2} \\ \text{GD: } \frac{\partial J + \frac{\lambda}{2} ||w||_{2}^{2}}{\partial w} &= \frac{\partial J}{\partial w} + \lambda w \\ &= \sum_{i} (-y_{i} + \frac{1}{1 + e^{-w^{T} \phi(x_{i})}}) \phi(x_{i}) + \lambda w^{t-1} \end{aligned}$$

5. clastering more than 2 one of the methods: Just creating n models for n type of data. each model is a i vs rest. For each model we have:

$$P(y = i|x) = \sigma(w_i^T x)$$

Then see which have the max probability.

$$y^{test} = \underset{i \in \{0,1,\dots,N\}}{\arg\max} \, \sigma(w_i^T \phi(x^{test}))$$

6. MAXIMUM a Posteriori (MAP) Estimation:

Incorporating with prior knowledge with parameters. When the data is not enough and we have some prior knowledge, we do MAP.
MAP setting:

- we start with a "prior" model on parameters of systems  $\rightarrow P_{prior}(\theta)$
- we observe a dataset  $D \to P_{\theta}(D)$
- Given D, how the prior knowledge on  $\theta$  changes  $\rightarrow P(\theta|D)$

MAP:  $\max_{\theta} P(\theta|D) \rightarrow \hat{\theta_{MAP}}$ 

DS4400 Notes 02/14

## Maximum A Posteriori (MAP) Estimation:

Incoperate prior knowledge into system(parameter) learning

1. Exercise: bernoulli expriment:  $\theta = P(x = 1) \rightarrow \hat{\theta}_{MLE}$  can be learned from training set  $X_1 = x_1 \dots$ 

Tossing a coin :  $\theta = P(X = 1) = P('H'), D = H, H \rightarrow \hat{\theta}_{MLE} = \frac{2}{2} = 1$ 

- 2. MAP setting:
  - Put a prior distribution on  $\theta$  (that encodes prior knowledge / domain expertise)
  - Observe a data set  $D = \{X_1 = x_1, ..., x_N = N\} \rightarrow P(D|\theta) = P_{\theta}(D)$
  - How much our knowledge about  $\theta$  changes after seeing data, D:  $P(\theta|D) \to \text{posterior}$  dist.

MAP:  $\max_{\theta} P(\theta|D) \equiv \frac{P(D|\theta)P(\theta)}{P(D)} = \max_{\theta} P(D|\theta)P(\theta) = \max_{\theta} L(\theta)P(\theta)$ 

3. Exercise:

$$\begin{split} X_1 &= x_1, X_2 = x_2 \dots, X_N = x_N, x_i \in \{0,1\} \\ P(x_i = 1) &= \theta, p(x_i = 0) = 1 - \theta \\ \text{Using Beta distribution: } P_{\alpha,\beta}(\theta) \propto \theta^{\alpha-1}(1-\theta)^{\beta-1} \\ \text{MAP: } \max_{\theta} P(\theta|D) &= \max P(x_1, x_2, \cdots, x_N|\theta) P_{\alpha,\beta}(\theta) \\ &= \prod_{i=1}^N P(X_i = x_i) P_{\alpha,\beta}(\theta) \\ &= \theta^{\sum x_i} (1-\theta)^{N-\sum x_i} \theta^{\alpha-1} (1-\theta)^{\beta-1} \\ &= \theta^{\sum x_i + \alpha - 1} (1-\theta)^{N-\sum x_i + \beta - 1} \\ \text{Let } \alpha' &= \sum x_i + \alpha - 1, \ \beta' &= N - \sum x_i + \beta - 1 \\ \text{Then } P(\theta|D) \propto P_{\alpha',\beta'}(\theta) \\ \max_{\theta} P(\theta|D) &= \max_{\theta} P_{\alpha',\beta'}(\theta) = \frac{\alpha'-1}{\alpha'+\beta'-2} \\ &= \frac{N}{N+\alpha+\beta-2} \frac{\sum x_1}{N} + \frac{\alpha+\beta-2}{N+\alpha+\beta-2} \frac{\alpha-1}{\alpha+\beta-2} \\ &= \eta \hat{\theta}_{MLE} + \eta \hat{\theta}_{prior\ mode} \end{split}$$

#### 4. Conclusion:

$$N \to \infty \Rightarrow \hat{\theta}_{MAP} \sim \hat{\theta}_{MLE}$$

$$N \to \hat{\theta} \Rightarrow \hat{\theta}_{MAP} \sim \hat{\theta}_{MLE}$$

$$N \to 0 \Rightarrow \hat{\theta}_{MAP} \sim \hat{\theta}_{prior\ mode}$$

having prior mode is just adding fake observations generated by prior mode. The bigger the  $1 - \eta$ , the more fake observations we add to dataset

## 5. MAP on logistic regression:

$$\rightarrow \min_{\omega} J(\omega)$$
. We might have  $\theta$  of very large terms.

Then choose 
$$p(\omega) \propto e^{-\omega^T \omega/2\sigma^2} = e^{-\|\omega\|_2^2/2\sigma^2}$$

$$\max P(\omega|D) \propto \max P(D|\omega)P(\omega)$$

$$\max_{\omega} P(\omega|D) \propto \max_{\omega} P(D|\omega)P(\omega)$$
Then  $\max_{\omega} log(P(\omega|D)) \propto \max_{\omega} logP(D|\omega) + logP(\omega)$ 

$$= \max_{\Omega} J(\omega) - \frac{1}{2\sigma^2} ||\omega||_2^2$$

$$= \max_{\omega} J(\omega) - \frac{1}{2\sigma^2} ||\omega||_2^2$$
$$\equiv \min_{\omega} -J(\omega) + \lambda ||\omega||_2^2$$

DS4400 Notes 02/18

#### Classification

## 1. Discriminative Modeling

Find a decisim boundary that seperates data into classes e.g logistic regression

Discriminative approacheds model:

$$P(y|x)$$
, y is class, x is feature vector.

e.g. 
$$P(y = 1|x) = \sigma(w^T \phi(x)) = \frac{1}{1 + e^{-w^T \phi(x)}}$$

# 2. Generative Modeling

Model distribution of data in each class as well as the distribution of classes themselves.

- $\rightarrow P(x|y)$  (Feature of class) and P(y) (class).
  - Assume we learn P(x|y), P(y) during training
  - How to classify a new test sample  $x^+$ ?

$$\Rightarrow \underset{j=0,1,...,L-1}{\operatorname{arg\,max}} P(y=j|x^t) = \underset{j=0,1,...,L-1}{\operatorname{arg\,max}} \frac{P(x^t|y=j)P(y=j)}{P(x^t)} \equiv \underset{j=0,1,...,L-1}{\operatorname{arg\,max}} P(x^t|y=j)P(y=j)$$

3. Example: email classification:  $\{(x^1, y^1), \dots, (x^N, y^N)\}$ 

probable 
$$x : \binom{"CPAS"}{"Free"} y = \{"non - spam", "spam"\}$$

Parameters to learn are

$$\theta_0^y \triangleq P(y=0) = P('non - spam')$$

$$\theta_1^y \triangleq P(y=1) = P('spam')$$

$$\theta_{\bar{x}|0}^{x|y} \triangleq P(x=\bar{x} \mid y=1) = P(x=\bar{x} \mid 'non - spam')$$

$$\theta_{\bar{x}|1}^{x|y} \triangleq P(x=\bar{x} \mid y=1) = P(x=\bar{x} \mid 'spam')$$

More generally,  $\Theta \triangleq$ 

$$\begin{cases}
\theta_j^{y} \triangleq P(y=j), \ \forall j=0,1,\dots,L-1 \\
\theta_{\bar{x}|j}^{x|y} \triangleq P(x=\bar{x} \mid y=j), \ \forall x=\bar{x}, \forall j=0,1,\dots,L-1
\end{cases}$$
(1)

4. Approach: MLE:

Approach: MLE: 
$$\text{MLE: } L(\theta) = P_{\Theta}(x1, y1, \dots, x^N, y^N) =_{iid} = \prod_j P_{\Theta}(x^i, y^i) = \prod_j P_{\Theta}(x^i|y^i) P(y^i) \\ P(y^i) = P(y^i = 0)^{1(y^i = 0)} \cdot P(y^i = 0)^{1(y^i = 0)} \cdot \dots P(y^i = 0)^{1(y^i = 0)} \\ = \theta_0^{1(y^i = 0)} \theta_0^{1(y^i = 1)} \dots \theta_0^{1(y^i = L - 1)} \\ L(\theta) = \prod_i P(x^i|y^i) \prod_i \prod_{j=0}^{L-1} P(y^i = j)^{1(j^i = j)} \\ = \prod_i P(x^i|y^i) \prod_i \prod_{j=0}^{L-1} \theta_j^{y1(j^i = j)} \\ \Rightarrow log L(\theta) = \sum_i log P(x^i|y^i) \sum_i \sum_{j=0}^{L-1} 1(j^i = j) log(\theta_j^y) \\ \text{To estimate} \\ \hat{\theta}_j^y \Rightarrow \frac{\partial Log L(\theta)}{\partial \theta_j^y} = 0 \Rightarrow \hat{\theta}_j^y = \frac{\sum_i^N 1(y^i = j)}{N} \\ \hat{\theta}_{\overline{x}|j}^{x|y} = \frac{\sum_i 1(x^i = \overline{x}, y^i = j)}{\sum_{i=1}^{N} 1(y^i = j)} \\ \text{This is called Vanilla Generative Model}.$$

$$\theta_j^y \Rightarrow L$$
 estimations  $\theta_{\bar{x}|j}^{x|y} \Rightarrow Lm^d$  estimations

(given x has d dimension and each dimension has m values) i.e.  $x = \begin{bmatrix} 0, 1, 2, ..., m-1 \\ 0, 1, 2, ..., m-1 \\ \vdots \\ 0, 1, 2, ..., m-1 \end{bmatrix}$ 

5. Problem: Document Classifications:

length of doc: |DOC|, we have possibly |DOC| features, each feature may have |DOC| of possible values. Then the estimation is  $L \cdot |DOC|^{|DOC|}$ 

6. Naive Bayes Method:

Generative model where feature are independent for a particular given class.

 $P(x = \bar{x}|y = i)$  where x has d features.

e.g. spam classifications: 
$$x = \begin{pmatrix} 'free' \\ 'caps' \\ 'call\ now' \end{pmatrix}, y = 0, 1$$

 $P(x = (1, 1, 1)^T | y = 1) = P('free' = 1 | y = 1)P('caps' = 1 | y = 1)P('call\ now' | y = 1)$ class conditional independence. O(Lmd)